

Integrated Science Assessment for Sulfur Oxides – Health Criteria



Integrated Science Assessment for Sulfur Oxides—Health Criteria

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ACRONYMS AND ABBREVIATIONS

Acronym/ Abbreviation	Meaning	Acronym/ Abbreviation	Meaning
(NH ₄) ₂ SO ₄	ammonium sulfate	ATD	atmospheric transport and dispersion
³⁵ SO ₂	radiolabeled SO ₂	ATS	American Thoracic Society
α	alpha, exposure factor	avg	average
a.m.	ante meridiem (before noon)	AZ	Arizona
AA	adenine-adenine genotype	β	beta
AB	Alberta	B[a]P	benzo[a]pyrene
ACS	American Cancer Society	BAL	bronchoalveolar lavage
ADMS	Advanced Dispersion Modeling System	BALF	bronchoalveolar lavage fluid
AER	air exchange rate; Atmospheric and Environmental Research	bax	B-cell lymphoma 2-like protein 4
AERMOD	American Meteorological Society/U.S. EPA Regulatory Model	BC	black carbon
AG	adenine-guanine genotype	Bcl-2	B-cell lymphoma 2
AHR	airway hyperresponsiveness	BK	Bangkok
AIRES	Aerosol Research Inhalation Epidemiology Study	BMI	body mass index
AIRS	Aerometric Information Retrieval System; Atmospheric Infrared Sounder	BP	blood pressure
AL	Alabama	BS	black smoke
ALRI	acute lower respiratory infection	BTEX	benzene, toluene, ethylbenzene, xylene
APEX	Air Pollution Exposure model	C	degrees Celsius; the product of microenvironmental concentration; carbon
APHEA	Air Pollution and Health: A European Approach study	C1	sulfur dioxide + nitrogen dioxide
APIMS	atmospheric pressure ionization mass spectrometry	C2	sulfur dioxide + PM ₁₀
AQCD	air quality criteria document	C3	sulfur dioxide + ozone
AQMEII	Air Quality Model Evaluation International Initiative	CA	California
AQS	air quality system	Ca	central site ambient SO ₂ concentration
ARDS	Acute Respiratory Distress Syndrome	CAA	Clean Air Act
ARIES	Aerosol Research Inhalation Epidemiology Study	CALPUFF	California Puff Model
ARP	Acid Rain Program	CAMP	Childhood Asthma Management Program
AT	Atascadero	CAPES	China Air Pollution and Health Effects Study
		CASAC	Clean Air Scientific Advisory Committee

Acronym/ Abbreviation	Meaning	Acronym/ Abbreviation	Meaning
CASTNet	Clean Air Status and Trends Network	D.C.	District of Columbia
CBSA	core-based statistical area	DEcCBP	diesel exhaust particle extract-coated carbon black particles
CDC	Centers for Disease Control and Prevention	DEN	diethylnitrosamine
CFR	Code of Federal Regulations	DEP	diesel exhaust particles
cGMP	cyclic guanosine monophosphate	df	degrees of freedom
CH ₃ SH	methyl mercaptan	DMDS	dimethyl disulfide
CHAD	Consolidated Human Activity Database	DMS	dimethyl sulfide
CHD	coronary heart disease	DNA	deoxyribonucleic acid
CHF	congestive heart failure	DOAS	differential optical absorption spectroscopy
CHIMERE	regional chemistry transport model	DPB	diastolic blood pressure
CI(s)	confidence interval(s)	DVT	deep vein thrombosis
CIMS	chemical ionization mass spectroscopy	e.g.	exempli gratia (for example)
cIMT	carotid intima-media thickness	Ea	exposure to SO ₂ of ambient origin
Cl	chlorine radical	EBC	exhaled breath condensate
CMAQ	Community Multiscale Air Quality	EC	elemental carbon
CO	carbon monoxide; Colorado	ECA	Emissions Control Areas
CO ₂	carbon dioxide	ECG	electrocardiographic
COH	coefficient of haze	ED	emergency department
Conc	concentration	EGF	epidermal growth factor
Cong.	congress	EGFR	epidermal growth factor receptor
COPD	chronic obstructive pulmonary disease	EGU	electric power generating unit
COX-2	cyclooxygenase-2	EIB	exercise-induced bronchospasm
C-R	concentration-response (relationship)	EKG	electrocardiogram
CRDS	cavity ring-down spectroscopy	ELF	epithelial lining fluid
CRP	c-reactive protein	EMSA	electrophoretic mobility shift assay
CS ₂	carbon disulfide	Ena	exposure to SO ₂ of nonambient origin
CSAPR	Cross-State Air Pollution Rule	eNO	exhaled nitric oxide
CT	Connecticut	EP	entire pregnancy
CTM	chemical transport models	EPA	U.S. Environmental Protection Agency
CVD	cardiovascular disease	ET	total exposure over a time period of interest
D.C. Cir	District of Columbia Circuit	ET	extrathoracic
DBP	diastolic blood pressure	EWPM	emission-weighted proximity model

Acronym/ Abbreviation	Meaning	Acronym/ Abbreviation	Meaning
F	female	HERO	Health and Environmental Research Online
FB	fractional bias	HF	high frequency component of HRV
FEF _{25-75%}	forced expiratory flow at 25–75% of exhaled volume	HI	Hawaii
FEF _{50%}	forced expiratory flow at 50% of forced vital capacity	HK	Hong Kong
FEF _{75%}	forced expiratory flow at 75% of forced vital capacity	HO ₂	hydroperoxyl radical
FEF _{max}	maximum forced expiratory flow	HPDM	Hybrid Plume Dispersion Model
FEM	federal equivalent method	HR	hazard ratio(s); heart rate
FEV	forced expiratory volume	HRV	heart rate variability
FEV ₁	forced expiratory volume in 1 second	HS	hemorrhagic stroke
FOXp3	forkhead box P3	HSC	Harvard Six Cities
FPD	flame photometric detection	HSO ₃	bisulfate radical
FR	Federal Register	HSO ₃ ⁻	bisulfite
FRC	functional residual capacity	HSC	Harvard Six Cities
FRM	federal reference method	Hz	hertz
FVC	forced vital capacity	i.e.	id est (that is)
g	gram	i.p.	intraperitoneal
GA	Georgia	ICAM-1	intercellular adhesion molecule 1
GALA II	Genes-Environments and Admixture in Latino Americans	ICC	intraclass correlation coefficient
GG	guanine-guanine genotype	ICD	International Classification of Diseases; implantable cardioverter defibrillators
GIS	geographic information system	ICS	inhaled corticosteroid
GP	general practice	IDW	inverse distance weighting
GPS	global positioning system	IFN-γ	interferon gamma
GSTM1	glutathione S-transferase Mu 1	IgE	immunoglobulin E
GSTP1	glutathione S-transferase Pi 1	IgG	immunoglobulin G
h	hour(s)	IHD	ischemic heart disease
H ⁺	hydrogen ion	IKKβ	inhibitor of nuclear factor kappa-B kinase subunit beta
H ₂ O	water	IL	Illinois
H ₂ O ₂	hydrogen peroxide	IL-4	interleukin-4
H ₂ S	hydrogen sulfide	IL-5	interleukin-5
H ₂ SO ₃	sulfurous acid	IL-6	interleukin-6
H ₂ SO ₄	sulfuric acid	IN	Indiana
HC	hydrocarbon	IQR	interquartile range

Acronym/ Abbreviation	Meaning	Acronym/ Abbreviation	Meaning
IRP	Integrated Review Plan	MCh	methacholine
ISA	Integrated Science Assessment	MD	Maryland
ISAAC	International Study of Asthma and Allergies in Children	MDL	method detection limit
IUGR	intrauterine growth restriction	ME	Maine
I κ B α	nuclear factor of kappa light polypeptide gene enhancer in B-cells inhibitor, alpha	MG	geometric mean
j	microenvironment	mg	milligram
JE	joint model estimate	MI	myocardial infarction (“heart attack”); Michigan
k	reaction rate; decay constant derived from empirical data; rate of SO ₂ loss in the microenvironment	min	minimum; minute
kg	kilogram(s)	MINAP	Myocardial Ischaemia National Audit Project
km	kilometer(s)	MISA	Meta-analysis of the Italian studies on short-term effects of air pollution
L	liter(s)	mL	milliliter(s)
LBW	low birth weight	mm	millimeters
LDL	lower detection limit	MMEF	maximum midexpiratory flow
LF	low-frequency component of HRV	MMFR	maximal midexpiratory flow rate
LIF	laser induced fluorescence	mmHg	millimeters of mercury
LOESS	locally weighted scatterplot smoothing	mmol	millimole
Lp-PLA ₂	lipoprotein-associated phospholipase A ₂	MN	micronuclei formation, Minnesota
LRS	lower respiratory symptoms	MNPCE	polychromatophilic erythroblasts of the bone marrow
LUR	land use regression	mo	month(s)
LX	lung adenoma-susceptible mouse strain	MO	Missouri
m	meter	MOA	mode(s) of action
M	male	MODIS	Moderate Resolution Imaging Spectroradiometer
M1	Month 1	mRNA	messenger ribonucleic acid
M2	Month 2	ms	millisecond
M3	Month 3	MUC5AC	mucin 5AC glycoprotein
MA	Massachusetts	n	sample size; total number of microenvironments that the individual has encountered
MACC	Modeling Atmospheric Composition and Climate	N	population number
max	maximum	N ₂	molecular nitrogen
MAX-DOAS	multiaxis differential optical absorption spectroscopy	NAAQS	National Ambient Air Quality Standards
		NaCl	sodium chloride

Acronym/ Abbreviation	Meaning	Acronym/ Abbreviation	Meaning
NALF	nasal lavage fluid	OVA	ovalbumin
NASA	National Aeronautics and Space Administration	p	probability
NBP	NO _x Budget Program	P	Pearson correlation
NCore	National Core network	p.m.	post meridiem (after noon)
NEI	National Emissions Inventory	P53	tumor protein 53
NFκB	nuclear factor kappa-light-chain-enhancer of activated B cells	PA	Pennsylvania
NH	New Hampshire	PAH(s)	polycyclic aromatic hydrocarbon(s)
NH ₃	ammonia	PAPA	Public Health and Air Pollution in Asia
NH ₄ ⁺	ammonium ion	Pb	lead
NHAPS	National Human Activity Pattern Survey	PC	provocative concentration
NHLBI	National Heart, Lung, and Blood Institute	PC(SO ₂)	provocative concentration of SO ₂
NJ	New Jersey	PE	pulmonary embolism
nm	nanometer	PEF	peak expiratory flow
NMMAAPS	The National Morbidity Mortality Air Pollution Study	PEFR	peak expiratory flow rate
NO	nitric oxide	Penh	enhanced pause
NO ₂	nitrogen dioxide	PM	particulate matter
NO ₃	nitrate radical	PM ₁₀	In general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; a measurement of thoracic particles (i.e., that subset of inhalable particles thought small enough to penetrate beyond the larynx into the thoracic region of the respiratory tract). In regulatory terms, particles with an upper 50% cutpoint of 10 ± 0.5 μm aerodynamic diameter (the 50% cutpoint diameter is the diameter at which the sampler collects 50% of the particles and rejects 50% of the particles) and a penetration curve as measured by a reference method based on Appendix J of 40 CFR Part 50 and designated in accordance with 40 CFR Part 53 or by an equivalent method designated in accordance with 40 CFR Part 53.
non-HS	non-hemorrhagic stroke		
NO _x	the sum of NO and NO ₂		
NR	not reported		
NY	New York		
O ₃	ozone		
obs	observations		
OC	organic carbon		
OCD	Off-shore and Coastal Dispersion model		
OCS	carbonyl sulfide		
OH	hydroxide; Ohio		
OHCA	out-of-hospital cardiac arrests		
OMI	Ozone Monitoring Instrument		
ON	Ontario		
OR	odds ratio(s)		

Acronym/ Abbreviation	Meaning	Acronym/ Abbreviation	Meaning
PM _{10-2.5}	In general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm and greater than a nominal 2.5 µm; a measurement of thoracic coarse particulate matter or the coarse fraction of PM ₁₀ . In regulatory terms, particles with an upper 50% cutpoint of 10 µm aerodynamic diameter and a lower 50% cutpoint of 2.5 µm aerodynamic diameter (the 50% cutpoint diameter is the diameter at which the sampler collects 50% of the particles and rejects 50% of the particles) as measured by a reference method based on Appendix O of 40 CFR Part 50 and designated in accordance with 40 CFR Part 53 or by an equivalent method designated in accordance with 40 CFR Part 53.	Q5	5th quintile
		QT interval	time between start of Q wave and end of T wave in ECG
		r	correlation coefficient
		R ²	square of the correlation coefficient
		Raw	airway resistance
		REA	Risk and Exposure Assessment
		redox	reduction-oxidation
		RH	relative humidity
		RHC	robust highest concentration
		RIOPA	Relationship Among Indoor, Outdoor, and Personal Air
		RMSE	root mean squared error
PM _{2.5}	In general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; a measurement of fine particles. In regulatory terms, particles with an upper 50% cutpoint of 2.5 µm aerodynamic diameter (the 50% cutpoint diameter is the diameter at which the sampler collects 50% of the particles and rejects 50% of the particles) and a penetration curve as measured by a reference method based on Appendix L of 40 CFR Part 50 and designated in accordance with 40 CFR Part 53, by an equivalent method designated in accordance with 40 CFR Part 53, or by an approved regional method designated in accordance with Appendix C of 40 CFR Part 58.	rMSSD	root-mean-square of successive differences
		RR	risk ratio(s), relative risk
		RSP	respirable suspended particles
		RSV	respiratory syncytial virus
		RT	total respiratory resistance
		S	sulfur
		S. Rep	Senate Report
		S ₂ O	disulfur monoxide
		SBP	systolic blood pressure
		sCI	stabilized Criegee intermediate
		SD	standard deviation
		SDCCE	simulated downwind coal combustion emissions
PMR	peak-to-mean ratio	SDNN	standard deviation of all normal-to-normal intervals
PNC	particle number concentration	SEARCH	Southeast Aerosol Research Characterization
ppb	parts per billion	sec	second(s)
ppm	parts per million	SES	socioeconomic status
PWEI	Population Weighted Emissions Index	Sess.	session
Q1	1st quartile or quintile	SGA	small for gestational age
Q2	2nd quartile or quintile	sGAW	specific airway conductance
Q3	3rd quartile or quintile	SH	Shanghai
Q4	4th quartile or quintile		

Acronym/ Abbreviation	Meaning	Acronym/ Abbreviation	Meaning
SHEDS	Stochastic Human Exposure and Dose Simulation	TSP	total suspended solids
SHEEP	Stockholm Heart Epidemiology Programme	TX	Texas
SLAMS	state and local air monitoring stations	U.K.	United Kingdom
SO	sulfur monoxide	U.S.	United States of America
SO ₂	sulfur dioxide	U.S.C.	U.S. Code
SO ₃ ²⁻	sulfite	μ	mu; micro
SO ₃	sulfur trioxide	UFP	ultrafine particulate matter
SO ₄	sulfur tetroxide	μg/m ³	micrograms per cubic meter
SO ₄ ²⁻	sulfate	URS	upper respiratory symptoms
SOM	self-organizing map	UT	Utah
SO _x	sulfur oxides	UV	ultraviolet
SPE	single-pollutant model estimate	UVF	ultraviolet fluorescence
SPM	source proximity model; suspended particulate matter	VE	minute volume
sRaw	specific airway resistance	V _{max}	maximal flow of expired vital capacity
ST segment	segment of the electrocardiograph between the end of the S wave and beginning of the T wave	V _{max25}	maximal expiratory flow rate at 25%
STN	Speciation Trends Network	V _{max50}	maximal expiratory flow rate at 50%
SYP	synaptophysin	V _{max75}	maximal expiratory flow rate at 75%
t	fraction of time spent in a microenvironment across an individual's microenvironmental exposures, time	VOC	volatile organic compound
T1	first trimester	vs	versus
T2	second trimester	VSGA	very small for gestational age
T3	third trimester	VTE	venous thromboembolism
TB	tracheobronchial	WBC	white blood cell
TBARS	thiobarbituric acid reactive substances (species)	WH	Wuhan
TC	total hydrocarbo	WHO	World Health Organization
Th1	T-helper 1	wk	week
Th2	T-helper 2	WHI	Women's Health Initiative
TIA	transient ischemic attack	WI	Wisconsin
TNF-α	tumor necrosis factor alpha	WRF	Weather Research and Forecasting
tpy	tons per year	yr	year(s)
		μg	microgram

PREFACE

Legislative Requirements for the Review of the National Ambient Air Quality Standards

Two sections of the Clean Air Act (CAA) govern the establishment, review, and revision of the National Ambient Air Quality Standards (NAAQS). Section 108 [42 U.S. Code (U.S.C.) 7408] directs the Administrator to identify and list certain air pollutants and then to issue air quality criteria for those pollutants. The Administrator is to list those air pollutants that in his or her “judgment, cause or contribute to air pollution which may reasonably be anticipated to endanger public health or welfare,” “the presence of which in the ambient air results from numerous or diverse mobile or stationary sources,” and “for which ... [the Administrator] plans to issue air quality criteria ...” [42 U.S.C. 7408(a)(1); [CAA, 1990a](#)]. Air quality criteria are intended to “accurately reflect the latest scientific knowledge useful in indicating the kind and extent of all identifiable effects on public health or welfare, which may be expected from the presence of [a] pollutant in the ambient air ...” [42 U.S.C. 7408(a)(2)]. Section 109 [42 U.S.C. 7409; [CAA, 1990b](#)] directs the Administrator to propose and promulgate “primary” and “secondary” NAAQS for pollutants for which air quality criteria are issued.

Section 109(b)(1) defines a primary standard as one “the attainment and maintenance of which in the judgment of the Administrator, based on such criteria and allowing an adequate margin of safety, are requisite to protect the public health.”¹ A secondary standard, as defined in Section 109(b)(2), must “specify a level of air quality the attainment and maintenance of which, in the judgment of the Administrator, based on such criteria, is requisite to protect the public welfare from any known or anticipated adverse effects associated with the presence of [the] air pollutant in the ambient air.”²

The requirement that primary standards provide an adequate margin of safety was intended to address uncertainties associated with inconclusive scientific and technical information available at the time of standard setting. It was also intended to provide a

¹ The legislative history of Section 109 indicates that a primary standard is to be set at “... the maximum permissible ambient air level ... which will protect the health of any [sensitive] group of the population,” and that for this purpose “reference should be made to a representative sample of persons comprising the sensitive group rather than to a single person in such a group” S. Rep. No. 91:1196, 91st Cong., 2d Sess. 10 (1970).

² Section 302(h) of the Act [42 U.S.C. 7602(h)] provides that all language referring to effects on welfare includes, but is not limited to, “effects on soils, water, crops, vegetation, man-made materials, animals, wildlife, weather, visibility and climate, damage to and deterioration of property, and hazards to transportation, as well as effects on economic values and on personal comfort and well-being ...” ([CAA, 2005](#)).

reasonable degree of protection against hazards that research has not yet identified.¹ Both kinds of uncertainty are components of the risk associated with pollution at levels below those at which human health effects can be said to occur with reasonable scientific certainty. Thus, in selecting primary standards that provide an adequate margin of safety, the Administrator is seeking not only to prevent pollution levels that have been demonstrated to be harmful but also to prevent lower pollutant levels that may pose an unacceptable risk of harm, even if the risk is not precisely identified as to nature or degree. The CAA does not require the Administrator to establish a primary NAAQS at a zero-risk level or at background concentration levels, but rather at a level that reduces risk sufficiently so as to protect public health with an adequate margin of safety.² In so doing, protection is provided for both the population as a whole and those groups and lifestages potentially at increased risk for health effects from exposure to the air pollutant for which each NAAQS is set.

In addressing the requirement for an adequate margin of safety, the U.S. Environmental Protection Agency (U.S. EPA) considers such factors as the nature and severity of the health effects involved, the size of the sensitive group(s), and the kind and degree of the uncertainties. The selection of any particular approach to providing an adequate margin of safety is a policy choice left specifically to the Administrator's judgment.³

In setting standards that are "requisite" to protect public health and welfare as provided in Section 109(b), the U.S. EPA's task is to establish standards that are neither more nor less stringent than necessary for these purposes. In so doing, the U.S. EPA may not consider the costs of implementing the standards.⁴ Likewise, "[a]ttainability and technological feasibility are not relevant considerations in the promulgation of national ambient air quality standards."⁵

Section 109(d)(1) requires that "not later than December 31, 1980, and at 5-year intervals thereafter, the Administrator shall complete a thorough review of the criteria published under Section 108 and the national ambient air quality standards ... and shall make such revisions in such criteria and standards and promulgate such new standards as may be

¹ See *Lead Industries Association v. EPA*, 647 F.2d 1130, 1154 [District of Columbia Circuit (D.C. Cir.) 1980]; *American Petroleum Institute v. Costle*, 665 F.2d 1176, 1186 (D.C. Cir. 1981); *American Farm Bureau Federation v. EPA*, 559 F. 3d 512, 533 (D.C. Cir. 2009); *Association of Battery Recyclers v. EPA*, 604 F. 3d 613, 617–18 (D.C. Cir. 2010).

² See *Lead Industries v. EPA*, 647 F.2d at 1156 n.51; *Mississippi v. EPA*, 744 F. 3d 1334, 1339, 1351, 1353 (D.C. Cir. 2013).

³ See *Lead Industries Association v. EPA*, 647 F.2d at 1161–62; *Mississippi v. EPA*, 744 F. 3d at 1353.

⁴ See generally, *Whitman v. American Trucking Associations*, 531 U.S. 457, 465–472, 475–476 (2001).

⁵ See *American Petroleum Institute v. Costle*, 665 F. 2d at 1185.

appropriate” Consistent with this provision, this final ISA contains the air quality criteria addressing the human health effects of SO_x for the current review and reflects the U.S. EPA’s periodic review of those criteria. Section 109(d)(2) requires that an independent scientific review committee “shall complete a review of the criteria ... and the national primary and secondary ambient air quality standards ... and shall recommend to the Administrator any new ... standards and revisions of existing criteria and standards as may be appropriate” Since the early 1980s, this independent review function has been performed by the Clean Air Scientific Advisory Committee (CASAC).¹

Overview and History of the Reviews of the Primary National Ambient Air Quality Standard for Sulfur Oxides

NAAQS are defined by four basic elements: indicator, averaging time, level, and form. The indicator defines the pollutant to be measured in the ambient air for the purpose of determining compliance with the standard. The averaging time defines the time period over which air quality measurements are to be obtained and averaged or cumulated, considering evidence of effects associated with various time periods of exposure. The level of a standard defines the air quality concentration used (i.e., an ambient air concentration of the indicator pollutant) in determining whether the standard is achieved. The form of the standard defines the air quality statistic, the value of which is compared to the level of the standard in determining whether an area attains the standard. For example, the form of the current primary 1-h sulfur oxides standard is the 3-yr avg of the 99th percentile of the annual distribution of 1-h daily maximum sulfur dioxide (SO₂) concentrations. The Administrator considers these four elements collectively in evaluating the protection to public health provided by the primary NAAQS.

The U.S. EPA considers the term sulfur oxides to refer to multiple gaseous oxidized sulfur species such as sulfur dioxide (SO₂) and sulfur trioxide (SO₃). In setting the current standard in 2010, SO₂ was chosen as the indicator for sulfur oxides because as in previous reviews, it was recognized as the most abundant sulfur oxide species in the atmosphere, and there is a large body of health effects evidence associated with SO₂. The atmospheric chemistry, exposure, and health effects associated with sulfur compounds present in particulate matter (PM) were most recently considered in the U.S. EPA’s 2012 review of the NAAQS for PM. The ecological effects of sulfur oxides are being considered in a separate Integrated Science Assessment (ISA) for Oxides of Nitrogen, Oxides of Sulfur, and Particulate Matter—Ecological Criteria ([U.S. EPA, 2017a](#)), while

¹ Lists of CASAC members and of members of the CASAC Augmented for Sulfur Oxides Panel are available at: <http://yosemite.epa.gov/sab/sabproduct.nsf/WebCASAC/CommitteesandMembership?OpenDocument>.

the visibility, climate, and materials damage-related welfare effects of particulate sulfur compounds are being evaluated in the ISA for particulate matter, as described in the Integrated Review Plan for the National Ambient Air Quality Standards for Particulate Matter ([U.S. EPA, 2016c](#)).

The U.S. EPA issued the air quality criteria for sulfur oxides in 1969 [34 Federal Register (FR) 1988; ([HEW, 1969](#))]. Based on these criteria, the U.S. EPA promulgated NAAQS for sulfur oxides in 1971, establishing the indicator as SO₂ [36 FR 8186; ([U.S. EPA, 1971](#))]. The 1971 primary standards were set at 365 µg/m³ [equal to 0.14 parts per million (ppm)] averaged over a 24-h period, not to be exceeded more than once per year, and at 80 µg/m³ (equal to 0.03 ppm) annual arithmetic mean. Since then, the Agency has completed multiple reviews of the air quality criteria and standards, as summarized in [Table I](#).

Table I History of the primary National Ambient Air Quality Standards for sulfur oxides since 1971.

Final Rule/Decisions	Indicator	Averaging Time	Level	Form
1971 36 FR 8186 Apr 30, 1971	SO ₂	24 h 1 yr	140 ppb ^a 30 ppb ^a	One allowable exceedance Annual arithmetic average
1996 61 FR 25566 May 22, 1996	Both the 24-h and annual average standards retained without revision.			
2010 75 FR 35520 June 22, 2010	SO ₂	1 h	75 ppb	3-yr average of the 99th percentile of the annual distribution of daily maximum 1-h concentrations
24-h and annual SO ₂ standards revoked.				

FR = Federal Register; SO₂ = sulfur dioxide.

^aThe level of the 24-h SO₂ standard was 365 µg/m³ or 0.14 parts per million (ppm) [equivalent to 140 parts per billion (ppb)]. The level of the annual SO₂ standard was 80 µg/m³ or 0.03 ppm (30 ppb) (36 FR 8186). The levels are presented in ppb for ease of comparison with the 1-h standard issued in 2010 (75 FR 35520).

In 1982, the U.S. EPA published the Air Quality Criteria for Particulate Matter and Sulfur Oxides ([U.S. EPA, 1982a](#)) along with an addendum of newly published controlled human exposure studies, which updated the scientific criteria upon which the initial standards were based ([U.S. EPA, 1982b](#)). In 1986, a second addendum was published presenting newly available evidence from epidemiologic and controlled human exposure

studies ([U.S. EPA, 1986](#)). In 1988, the U.S. EPA published a proposed decision not to revise the existing standards (53 FR 14926). However, the U.S. EPA specifically requested public comment on the alternative of revising the current standards and adding a new 1-h primary standard of 0.4 ppm SO₂ to protect against short-term peak exposures.

As a result of public comments on the 1988 proposal and other post-proposal developments, the U.S. EPA published a second proposal on November 15, 1994 (59 FR 58958). The 1994 re-proposal was based in part on a supplement to the second addendum of the criteria document, which evaluated new findings on the respiratory effects of short-term SO₂ exposures in individuals with asthma ([U.S. EPA, 1994](#)). As in the 1988 proposal, the U.S. EPA proposed to retain the existing 24-h and annual standards. The U.S. EPA also solicited comment on three regulatory alternatives to further reduce the health risk posed by exposure to high 5-minute peaks of SO₂ if additional protection were judged to be necessary. The three alternatives were: (1) revising the existing primary NAAQS for sulfur oxides by adding a new 5-minute standard of 0.60 ppm SO₂; (2) establishing a new regulatory program under Section 303 of the Act to supplement protection provided by the existing NAAQS, with a trigger level of 0.60 ppm SO₂ with one expected exceedance; and (3) augmenting implementation of existing standards by focusing on those sources or source types likely to produce high 5-minute concentrations of SO₂.

Based upon an exposure analysis conducted by the U.S. EPA, the Administrator concluded that short-term (e.g., 5-minute) exposure of individuals with asthma to SO₂ at levels that can reliably elicit adverse health effects was likely to be a rare event when viewed in the context of the entire population of individuals with asthma. Thus, the Administrator judged that high 5-minute SO₂ concentrations did not pose a broad public health problem when viewed from a national perspective, and a 5-minute standard was not promulgated. In addition, no other regulatory alternative was finalized, and the 24-h and annual average primary SO_x standards were retained in 1996 (61 FR 25566).

The American Lung Association and the Environmental Defense Fund challenged the U.S. EPA's decision not to establish a 5-minute standard. On January 30, 1998, the Court of Appeals for the District of Columbia ("D.C. Circuit") found that the U.S. EPA had failed to adequately explain its determination that no revision to the SO₂ NAAQS was appropriate and remanded the decision back to the U.S. EPA for further explanation.¹ Specifically, the court found that the U.S. EPA had not provided adequate rationale to support the Agency judgment that exposures to high 5-minute concentrations of SO₂ do not pose a public health problem from a national perspective even though these peaks will likely cause adverse health impacts in a subset of individuals with asthma. Following

¹ See *American Lung Ass'n v. EPA*, 134 F. 3d 388 (D.C. Cir. 1998).

the remand, the U.S. EPA requested that states voluntarily submit 5-minute SO₂ monitoring data to be used to conduct air quality analyses in order to gain a better understanding of the magnitude and frequency of high, 5-minute peak SO₂ concentrations. The data submitted by states and the analyses based on this data helped inform the last review of the SO_x NAAQS, which ultimately addressed the issues raised in the 1998 remand.

The last review of the health-related air quality criteria for sulfur oxides and the primary SO_x standard was initiated in May 2006 (71 FR 28023).^{1,2} The Agency's plans for conducting the review were presented in the Integrated Review Plan (IRP) for the Primary National Ambient Air Quality Standards for Sulfur Oxides ([U.S. EPA, 2007a](#)), which included consideration of comments received during a CASAC consultation as well as public comment on a draft IRP. The science assessment for the review was described in the 2008 ISA for Sulfur Oxides—Health Criteria ([U.S. EPA, 2008d](#)), multiple drafts of which received review by CASAC and the public. The U.S. EPA also conducted quantitative human risk and exposure assessments after having consulted with CASAC and receiving public comment on a draft analysis plan ([U.S. EPA, 2007b](#)). These technical analyses were presented in the Risk and Exposure Assessment (REA) to Support the Review of the SO₂ Primary National Ambient Air Quality Standards ([U.S. EPA, 2009c](#)), multiple drafts of which were reviewed by CASAC and the public.

On June 22, 2010, the U.S. EPA revised the primary SO₂ NAAQS to provide requisite protection of public health with an adequate margin of safety (75 FR 35520). Specifically, after concluding that the then-existing 24-h and annual standards were inadequate to protect public health with an adequate margin of safety, the U.S. EPA established a new 1-h SO₂ standard at a level of 75 parts per billion (ppb), based on the 3-yr avg of the annual 99th percentile of 1-h daily maximum concentrations. This standard was promulgated to provide protection against SO₂-related health effects associated with short-term exposures ranging from 5 min to 24 h. More specifically, the U.S. EPA concluded that a 1-h SO₂ standard at 75 ppb would substantially limit exposures associated with the adverse respiratory effects (e.g., decrements in lung function and/or respiratory symptoms) reported in exercising asthmatics following 5–10 min exposures in controlled human exposure studies, as well as the more serious health associations (e.g., respiratory-related emergency department visits and hospitalizations) reported in epidemiologic studies that mostly used daily metrics

¹ Documents related to reviews completed in 2010 and 1996 are available at: <https://www.epa.gov/naaqs/sulfur-dioxide-so2-primary-air-quality-standards>.

² The U.S. EPA conducted a separate review of the secondary SO₂ NAAQS jointly with a review of the secondary NAAQS for oxides of nitrogen. The Agency retained those secondary standards, without revision, to address the direct effects on vegetation of exposure to oxides of nitrogen and sulfur (77 FR 20218).

(1-h daily max and 24-h avg). In the last review, the U.S. EPA also revoked the then-existing 24-h and annual primary standards based largely on the recognition that the new 1-h standard at 75 ppb would generally maintain 24-h and annual SO₂ concentrations well below the NAAQS, as well as the lack of evidence indicating the need for such longer-term standards (75 FR 35549-50). The decision to set a 1-h standard at 75 ppb—in part to substantially limit exposure to 5-min concentrations of SO₂ resulting in adverse respiratory effects in exercising asthmatic individuals—also addressed the issues raised in the remand by the D.C. Circuit in 1998.

As mentioned above, the U.S. EPA's last review placed considerable weight on substantially limiting health effects associated with high 5-min SO₂ concentrations. Thus, as part of the final rulemaking, the U.S. EPA for the first time required that state and local agencies operating continuous SO₂ analyzers report either the highest 5-min concentration for each hour of the day, or all twelve 5-min concentrations for each hour of the day. The rationale for this requirement was that such data were recognized as critical in the 2010 review and additional monitoring data were anticipated to be valuable for informing future health studies and NAAQS reviews (75 FR 35522).

After publication of the final rule, a number of industry groups and states filed petitions for review arguing that the U.S. EPA failed to follow notice-and-comment rulemaking procedures, and that the decision to establish the 1-h SO₂ NAAQS at 75 ppb was arbitrary and capricious because it was lower than statutorily authorized. The D.C. Circuit rejected these challenges, thereby upholding the standard in its entirety.¹

¹ See *National Environmental Development Association's Clean Air Project v. EPA*, 686 F. 3d 803 (D.C. Cir. 2012), cert. denied *Asarco LLC v. EPA*, 133 S. Ct. 983 (Jan. 22, 2013).

EXECUTIVE SUMMARY

Purpose and Scope of the Integrated Science Assessment

This Integrated Science Assessment (ISA) is a comprehensive evaluation and synthesis of policy-relevant science aimed at characterizing (1) exposures to sulfur oxides (SO_x) in ambient air, for which sulfur dioxide (SO₂) is currently the primary atmospheric indicator, and (2) the health effects associated with these exposures.¹ Thus, this ISA serves as the scientific foundation for the review of the primary (health-based) National Ambient Air Quality Standard (NAAQS) for SO_x. The indicator² for the current standard is SO₂ because at the time the standard was set it was identified as the most prevalent species of SO_x in the atmosphere and the one for which there is a large body of scientific evidence on health effects. The health effects of sulfate and other particulate sulfur compounds are considered as part of the review of the NAAQS for particulate matter [e.g., in the 2009 Integrated Science Assessment for Particulate Matter ([U.S. EPA, 2009a](#))].³ The ecological effects of sulfur oxides are being considered in a separate ISA for Oxides of Nitrogen, Oxides of Sulfur, and Particulate Matter—Ecological Criteria ([U.S. EPA, 2017a](#)), while the visibility, climate, and materials damage-related welfare effects of particulate sulfur compounds are being evaluated in the ISA for particulate matter ([U.S. EPA, 2016c](#)).

In 2010, the U.S. Environmental Protection Agency (U.S. EPA) established a new 1-hour SO₂ primary standard of 75 ppb as a 3-year avg of the 99th percentile of each year's 1-hour daily max concentrations (75 FR 35520).⁴ The 1-hour standard was established to protect against an array of respiratory effects associated with short-term exposures in at-risk populations, such as people with asthma. This standard was based on direct evidence of SO₂-related effects in controlled human exposure studies of exercising individuals with asthma, as well as epidemiologic evidence of associations between SO₂ concentrations in ambient air and respiratory-related emergency department visits and hospitalizations. The U.S. EPA also revoked the existing 24-hour and annual primary SO₂ standards of 140 and 30 ppb, respectively, based largely on the recognition that the new 1-hour standard would generally maintain 24-hour and annual SO₂ concentrations

¹ The general process for developing an ISA, including the framework for evaluating weight of evidence and drawing scientific conclusions and causal judgments, is described in a companion document, *Preamble to the Integrated Science Assessments* ([U.S. EPA, 2015b](#)), <https://www.epa.gov/isa>.

² The four components to a NAAQS are: (1) indicator (e.g., SO₂), (2) level (e.g., 75 ppb), (3) averaging time (e.g., 1 hour), and (4) form (e.g., 3 year avg of the 99th percentile of each year's daily 1-hour max concentrations).

³ In this ISA, the blue electronic links can be used to navigate to cited chapters, sections, tables, figures, and studies.

⁴ The legislative requirements and history of the SO₂ NAAQS are described in detail in the [Preface](#) to this ISA.

well below the then-existing standards and on the lack of evidence indicating the need for such longer-term standards (75 FR 35549-50). The U.S. EPA also began requiring that state and local agencies operating continuous SO₂ analyzers to report either the highest 5-minute avg SO₂ concentrations for each hour of the day or all twelve 5-minute avg SO₂ concentrations for each hour of the day.

This ISA updates the 2008 ISA for Sulfur Oxides [([U.S. EPA, 2008d](#)), hereafter referred to as the 2008 SO_x ISA] with studies and reports published from January 2008 through August 2016. The U.S. EPA conducted in-depth searches to identify peer-reviewed literature on relevant topics such as health effects, atmospheric chemistry, ambient air concentrations, and exposure. Information was also solicited from subject-matter experts and the public during a kick-off workshop held at the U.S. EPA in June 2013. Additionally, drafts of this ISA were reviewed by the CASAC at public meetings held in January 2016 and March 2017. Members of the public also had an opportunity to comment on drafts of the ISA. To fully describe the state of available science, The U.S. EPA also included in this ISA the most relevant studies from previous assessments.

As in the 2008 SO_x ISA, this ISA determines the causal nature of relationships with health effects only for SO₂ ([Chapter 5](#)). It does not make causality determinations for health effects of other SO_x species because SO₂ is the most abundant SO_x species in the atmosphere ([Chapter 2](#)), transformation products of SO_x such as sulfate are considered in the ISA for Particulate Matter ([U.S. EPA, 2009a](#)), and the health literature is focused on SO₂. Key to interpreting the health effects evidence is understanding the sources, chemistry, and distribution of SO₂ in the ambient air ([Chapter 2](#)) that influence exposure ([Chapter 3](#)), the uptake of inhaled SO₂ in the respiratory tract, and what biological mechanisms may subsequently be affected ([Chapter 4](#)). Further, the ISA aims to characterize the independent effect of SO₂ on health ([Chapter 5](#)). The ISA also informs policy-relevant issues ([Chapter 1](#) and [Chapter 6](#)), such as (1) exposure durations and patterns associated with health effects; (2) concentration-response relationship(s), including evidence of potential thresholds for effects; and (3) populations or lifestyles at increased risk for health effects related to SO₂ exposure ([Section 1.7.4](#) and [Chapter 6](#)).

Sources and Human Exposure to Sulfur Dioxide

As explained above, this ISA characterizes health effects related to ambient air SO₂ exposure. This requires understanding the factors that affect both the exposure to ambient SO₂ and the uncertainty in estimating exposure. These factors include spatial variability in SO₂ concentrations, exposure to copollutants, and uncharacterized time-activity patterns.

Emissions of SO₂ have decreased by approximately 79% from 1990 to 2014 subsequent to several federal air quality regulatory programs. Coal-fired electricity generation units are the dominant sources, emitting 3.2 million tons of SO₂ in 2014, nearly 5 times more than the next largest source (coal-fired boilers for industrial fuel combustion; [Section 2.2](#)). Preliminary data suggest power plant emissions have continued to decline through 2015–2016. In addition to emission rate, important factors that affect SO₂ concentrations at downwind locations include source characteristics (e.g., height of emissions, temperature, emission rate), local meteorology (e.g., wind, atmospheric stability, humidity, and cloud/fog cover), and chemistry in the plume ([Section 2.3](#)).

The national avg daily 1-hour max SO₂ concentration reported during 2013–2015 was 5.3 ppb with a 99th percentile concentration of 64 ppb ([Section 2.5](#)). However, 1-hour daily max SO₂ concentrations were 75 ppb or higher during this 3-year period at some monitoring sites located near point sources, such as power plants or metals processing facilities, or natural sources, such as volcanoes (which can produce hourly concentrations in excess of 2,000 ppb). The national 99th percentile 5-minute hourly max concentrations during 2013–2015 was 24 ppb, suggesting concentrations above 100 ppb are rare, although monitoring sites near large sources had high concentrations (100 ppb or more), including four monitoring sites near smelters in Gila County, AZ with 99th percentile 5-minute hourly max concentrations ranging from 116 to 252 ppb. Hourly 5-minute max concentrations tracked closely with their corresponding 1-h avg concentrations, with 75% of sites having a correlation above 0.9, indicating that fluctuations in 5-minute hourly max concentrations are well represented by changes in 1-h avg concentrations. The ratio of 5-minute hourly max concentrations to their corresponding 1-h avg concentrations was generally in the range of 1–3, although higher ratios were also observed during some hours. Background SO₂ concentrations due to natural sources and man-made sources located outside the U.S. are very low across most of the country (annual avg less than 0.03 ppb) except in areas influenced by local cross-border sources and areas affected by volcanoes, such as Hawaii and parts of the West Coast.

Air quality models are used to estimate SO₂ concentrations over various averaging times in locations without ambient SO₂ monitors ([Section 2.6](#)). As part of the implementation program for the 2010 primary NAAQS for SO_x, air quality modeling may be used to characterize air quality for determining compliance with the standard where existing SO₂ monitors may not capture peak 1-hour concentrations (75 FR 35520). The widely used dispersion model American Meteorological Society/U.S. EPA Regulatory Model (AERMOD) is based on Gaussian dispersion models with enhancements to improve modeling of SO₂ plumes. Model evaluations conducted over averaging times from 1 hour to 1 year indicate that AERMOD is relatively unbiased in estimating upper-percentile 1-hour concentration values. Lagrangian puff dispersion models, such as CALPUFF

(California Puff Model), have been developed as an alternative to Gaussian dispersion models. Uncertainties in model predictions are influenced by uncertainties in model inputs, particularly emissions data and meteorological conditions.

Correlations between ambient air concentrations of SO₂ and other criteria pollutants are generally low (<0.4), although they vary across location, study, and SO₂ averaging time and are greater than 0.7 at some monitoring sites ([Section 3.4.3](#)). Median correlations of 1-hour daily max and 24-h avg SO₂ concentrations with particulate matter, nitrogen dioxide (NO₂), and carbon monoxide (CO) during 2013–2015 ranged from 0.2–0.4, while for ozone (O₃) the median daily copollutant correlation with SO₂ was less than 0.1 ([Figure 3-5](#)).

Estimating exposure concentrations of ambient SO₂ for use in epidemiologic studies can be done in multiple ways. Air quality monitoring data from a limited number of fixed-site monitors, which are assumed to represent population exposure, are frequently used, but these monitors may not capture the spatial variation in SO₂ concentrations across an urban area, which can be relatively high in areas affected by large point sources. Modeling approaches combining air quality data with geographic information or time-activity patterns, or both, can provide estimates of local ambient concentration or exposure concentration, although more complex approaches need more detailed inputs and have the potential for uncertainty related to missing sources, overly smooth concentration gradients, and other factors.

“Exposure error,” which refers to the bias and uncertainty associated with using exposure metrics to represent the actual exposure of an individual or population, can contribute to error in health effect estimates in epidemiologic studies ([Section 3.4.4](#)). Several exposure-related factors (including uncharacterized time-activity patterns, spatial and temporal variability of SO₂ concentrations, and distance of individuals and populations from air quality monitors used in the statistical analyses) contribute to error in estimating exposure to SO₂. Variation in activity patterns across individuals and over time results in corresponding variations in exposure concentration. Uncharacterized spatial variability in SO₂ concentrations can contribute to exposure error that tends to add uncertainty and reduce the magnitude of effect estimates in daily time-series epidemiologic studies. For long-term (e.g., annual) studies, the effect estimate may be increased or reduced by using fixed-site monitoring data, depending on the relative locations of sources, monitors, and exposed people. The exposure error associated with using fixed-site monitors is generally expected to widen confidence intervals so that the nominal coverage is below 95% for exposure effect estimates.

Dosimetry and Mode of Action of Inhaled Sulfur Dioxide

Understanding the absorption and fate of SO₂ in the body (dosimetry) and the biological pathways that potentially underlie health effects (mode of action) is crucial to characterizing the biological plausibility of SO₂ exposure as the cause of observed health effects.

Inhaled SO₂ is readily absorbed in the nasal passages of resting humans and laboratory animals ([Section 4.2](#)). As physical activity increases, there is an increase in breathing rate and a shift to breathing through the mouth, resulting in greater SO₂ penetration into the lower airways. Relative to healthy adults, children, and individuals with asthma or allergic rhinitis have an increased amount of oral breathing, and thus, may be expected to have greater SO₂ penetration into the lungs.

The distribution and clearance of inhaled SO₂ from the respiratory tract involves several chemical transformations, particularly the formation of sulfite and S-sulfonates. Sulfite is metabolized into sulfate, which is rapidly excreted through the urine, while S-sulfonates are cleared more slowly from the circulation over a period of days. Although SO₂-derived products have been found in the blood and urine within minutes of an inhalation exposure, a substantial portion of these products appear to be retained within the upper airways, particularly during nasal breathing, with only slow absorption into the blood.

Although inhaled SO₂ produces sulfite that may be distributed through the circulation, overall sulfite levels are heavily influenced by production within the body (endogenous production) and by eating food with sulfur-containing amino acids or sulfite itself ([Section 4.2.6](#)). For both adults and children, metabolism of sulfur-containing amino acids produces much more sulfite than is ingested as food additives. Sulfite produced endogenously generates levels two or more orders of magnitude higher than inhalation-derived sulfite levels for both children and adults, even for full-day exposures to 75 ppb SO₂ (i.e., the level of the 1-hour NAAQS). Sulfite ingestion from food additives varies widely, but is generally expected to exceed sulfite intake from inhalation in both adults and children, even for full-day exposures to 75 ppb SO₂. However, an important distinction is that inhalation-derived SO₂ products can accumulate in the respiratory tract, whereas sulfite from ingestion or endogenous production does not.

SO₂ inhalation produces bronchoconstriction in both healthy adults and those with asthma ([Section 4.3](#)), but the underlying processes are somewhat different. The response to SO₂ in healthy adults occurs primarily from activation of sensory nerves in the respiratory tract resulting in neural reflex responses through the vagus nerve and occurs at higher concentrations than the response in people with asthma. In adults with asthma, the response is only partly due to this neural reflex response, with inflammatory

mediators also being involved. Inhalation of SO₂ increases allergic inflammation in adults with asthma and in animals with allergic airways disease, which shares many features with asthma. Furthermore, SO₂ inhalation increases allergic sensitization in animals not already allergic, and once allergic, these animals respond to an allergen challenge with greater allergic inflammation and airway obstruction (likely due to bronchoconstriction) compared to animals who were not exposed to SO₂. These findings suggest that allergic inflammation and increased airway responsiveness due to short-term SO₂ exposure (minutes up to 1 month) may be linked to asthma exacerbation seen in epidemiologic studies.

For long-term SO₂ exposure (more than 1 month to years), animal studies provide additional evidence of airway inflammation, airway remodeling, airway hyperresponsiveness (AHR), and allergic sensitization. In animals that are not allergic, SO₂ inhalation leads to airway inflammation and allergic sensitization. In animals with allergic airway disease, SO₂ exposure increases airway responsiveness and airway remodeling. Thus, inhalation of SO₂ may lead to the development and worsening of allergic airway disease. The development of AHR may link long-term exposure to SO₂ to the epidemiologic outcome of physician-diagnosed asthma (new onset asthma).

While there is some evidence for extrapulmonary effects of inhaled SO₂, the potential modes of action underlying these responses are uncertain. Controlled human exposure studies provide evidence suggesting activation of sensory nerves in the respiratory tract resulting in a neural reflex response by SO₂ exposure, which could lead to changes in heart rate or heart rate variability. Additionally, the transport of sulfite into the circulation could result in redox stress, but this is likely to only occur at elevated or prolonged exposures due to the body's efficient metabolism of sulfite to sulfate.

Health Effects of Sulfur Dioxide Exposure

This ISA integrates information on SO₂ exposure and health effects from controlled human exposure, epidemiologic, and toxicological studies to form conclusions about the causal nature of relationships between SO₂ exposure and health effects. For most health effect categories, with the exception of reproductive and developmental effects, effects are evaluated separately for short-term exposures and long-term exposures. Health effects are considered in relation to the full range of SO₂ concentrations relevant to ambient conditions. Based on upper-percentile ambient air concentrations ([Section 2.5](#)) and the emphasis in the ISA on ambient-relevant exposures within one to two orders of magnitude of current conditions [Preamble to the ISAs ([U.S. EPA, 2015b](#)), Section 5c],

SO₂ concentrations up to 2,000 ppb¹ are defined to be ambient-relevant. A consistent and transparent framework [Preamble to the ISAs ([U.S. EPA, 2015b](#)), Table II] is applied to classify the health effects evidence according to a five-level hierarchy:

1. Causal relationship
2. Likely to be a causal relationship
3. Suggestive of, but not sufficient to infer, a causal relationship
4. Inadequate to infer a causal relationship
5. Not likely to be a causal relationship

The causal determinations presented in [Table ES-1](#) are informed by recent findings and whether these recent findings, integrated with information from the 2008 SO_x ISA, support a change in causal conclusions. Important considerations include: (1) determining whether laboratory studies of humans and animals demonstrate an independent health effect of SO₂ exposure and what the potential underlying biological mechanisms are; (2) determining whether there is consistency in epidemiologic evidence across various methods used to estimate SO₂ exposure; (3) examining epidemiologic studies of the potential influence of factors that could bias associations observed with SO₂ exposure; (4) determining the coherence of findings integrated across controlled human exposure, epidemiologic, and toxicological studies; and (5) making judgments regarding error and uncertainty in the collective body of available studies.

¹ The 2,000-ppb upper limit applies mostly to animal toxicological studies and also a few controlled human exposure studies. Experimental studies examining SO₂ exposures greater than 2,000 ppb were included if they provided information on the uptake of SO₂ in the respiratory tract or on potential biological mechanisms.

Table ES-1 Causal determinations for relationships between sulfur dioxide exposure and health effects from the 2008 and 2017 Integrated Science Assessment for Sulfur Oxides.

Health Effect Category ^a and Exposure Duration ^b	Causal Determination	
	2008 SO _x ISA ^c	2017 SO _x ISA
Respiratory effects—short-term exposure Section 5.2.1, Table 5-21	Causal relationship	Causal relationship
Respiratory effects—long-term exposure Section 5.2.2, Table 5-24	Inadequate to infer a causal relationship	Suggestive of, but not sufficient to infer a causal relationship
Cardiovascular effects—short-term exposure Section 5.3.1, Table 5-34	Inadequate to infer a causal relationship	Inadequate to infer a causal relationship
Cardiovascular effects—long-term exposure Section 5.3.2, Table 5-35	Not included	Inadequate to infer a causal relationship
Reproductive and developmental effects ^d Section 5.4, Table 5-38	Inadequate to infer a causal relationship	Inadequate to infer a causal relationship
Total mortality—short-term exposure Section 5.5.1, Table 5-41	Suggestive of, but not sufficient to infer, a causal relationship	Suggestive of, but not sufficient to infer a causal relationship
Total mortality—long-term exposure Section 5.5.2, Table 5-43	Inadequate to infer a causal relationship	Inadequate to infer a causal relationship
Cancer—long-term exposure Section 5.6, Table 5-44	Inadequate to infer a causal relationship	Inadequate to infer a causal relationship

ISA = Integrated Science Assessment; SO_x = sulfur oxides.

^aAn array of outcomes is evaluated as part of a broad health effect category: physiological measures (e.g., airway responsiveness), clinical outcomes (e.g., hospital admissions), and cause-specific mortality. Total mortality includes all nonaccidental causes of mortality and is informed by findings for the spectrum of morbidity effects (e.g., respiratory, cardiovascular) that can lead to mortality. The sections and tables referenced include a detailed discussion of the evidence that supports the causal determinations and the SO₂ concentrations with which health effects have been associated.

^bShort-term exposure refers to time periods of minutes up to 1 mo, while long-term exposures are more than 1 mo to yr.

^cPrevious causal determinations taken from the 2008 SO_x ISA ([U.S. EPA, 2008d](#)).

^dReproductive and developmental effects studies consider a wide range of exposure durations.

Sulfur Dioxide Exposure and Respiratory Effects

As in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), the current ISA concludes that there is a causal relationship between short-term SO₂ exposure and respiratory effects, particularly in individuals with asthma ([Section 5.2.1](#)). This determination is based on consistent, coherent, and biologically plausible evidence for asthma exacerbation due to SO₂ exposure. The clearest evidence for this conclusion comes from controlled human exposure studies available at the time of the 2008 SO_x ISA showing lung function decrements and respiratory symptoms in individuals with asthma exposed to SO₂ for 5–10 minutes at elevated breathing rates. The effects observed in these studies are consistent with the processes leading to asthma exacerbation described in the mode of action section ([Section 4.3](#)). Epidemiologic evidence, including recent studies not available at the time of the 2008 SO_x ISA, also supports a causal relationship, primarily due to studies reporting positive associations for asthma hospital admissions and emergency department visits with short-term SO₂ exposures, specifically for children. This is coherent with studies showing that children have increased airway responsiveness to a trigger and have a greater fraction of oral breathing relative to adults, suggesting they will have a greater response to SO₂ exposure than adults. Hospital admissions and emergency department visit studies that examined potential copollutant confounding reported associations were generally unchanged in copollutant models. Additional support comes from studies reporting positive associations between short-term SO₂ exposures and respiratory symptoms in children with asthma, although the evidence from respiratory symptoms studies in adults with asthma is less consistent. Finally, epidemiologic studies that report consistent positive associations between short-term SO₂ concentrations and respiratory mortality indicate a potential continuum of effects.

For long-term SO₂ exposure and respiratory effects the evidence is suggestive of, but not sufficient to infer, a causal relationship ([Section 5.2.2](#)). This conclusion is based on coherence among findings of a limited number of new epidemiologic studies showing associations between long-term SO₂ exposure and increases in asthma incidence among children and results of animal toxicological studies that provide a pathophysiologic basis for the development of asthma. However, uncertainty remains regarding the influence of other pollutants or mixtures of pollutants on the observed associations with SO₂ because these new epidemiologic studies have not examined the potential for copollutant confounding. Some epidemiologic evidence regarding respiratory symptoms and/or respiratory allergies among children also provides limited support for a possible relationship between long-term SO₂ exposure and the development of asthma. This represents a change in the causal determination made in the 2008 SO_x ISA from inadequate to suggestive, based on a limited body of new evidence.

Sulfur Dioxide Exposure and Other Health Effects

There is more uncertainty regarding relationships between SO₂ exposure and health effects outside of the respiratory system. SO₂ itself is unlikely to enter the bloodstream; however, its reaction products, such as sulfite, may do so. The amount of circulating sulfite due to inhalation of SO₂ at concentrations relevant to ambient air is far less than the contribution from metabolism of sulfur-containing amino acids.

For short-term SO₂ exposure and total mortality, the current ISA reaches the same conclusion as the 2008 SO_x ISA ([U.S. EPA, 2008d](#)): that the evidence is suggestive of, but not sufficient to infer, a causal relationship ([Section 5.5.1](#)). This conclusion is based on previous and recent multicity epidemiologic studies providing consistent evidence of positive associations. While recent multicity studies have analyzed some key uncertainties and data gaps identified in the 2008 SO_x ISA, questions remain regarding the potential for SO₂ to have an independent effect on mortality, considering issues such as the limited number of studies that examined copollutant confounding, evidence for a decrease in the size of SO₂-mortality associations in copollutant models with NO₂ and PM₁₀, and the lack of a potential biological mechanism for mortality following short-term exposures to SO₂.

For the remaining health effect categories (short-term and long-term SO₂ exposure and cardiovascular effects, long-term exposure and total mortality, reproductive and developmental effects, and long-term exposure and cancer), the evidence is inadequate to infer a causal relationship, mainly due to inconsistent evidence across specific outcomes and uncertainties regarding exposure measurement error, copollutant confounding, and potential modes of action. These conclusions are consistent with those made in the 2008 SO_x ISA, as illustrated in [Table ES-1](#).

Policy-Relevant Considerations for Health Effects Associated with Sulfur Dioxide Exposure

This section describes issues relevant for considering the potential importance of impacts of ambient SO₂ exposure on public health, including exposure durations observed to cause health effects, the shape of the concentration-response relationship, regional differences, and at-risk populations and lifestyles.

Evidence from controlled human exposure studies of respiratory effects after exposures of 5–10 minutes indicates a rapid onset of SO₂-related effects ([Section 5.2.1](#)).

Epidemiologic studies of asthma hospital admissions and emergency department visits using daily exposure metrics (24-h avg and 1-h daily max) show positive associations that are generally unchanged in copollutant models, although these associations could be

due to very short duration exposures (5–10 minutes) experienced during the day. The rapid onset of effects is also coherent with the limited number of epidemiologic studies that examined lag structures and reported associations within the first few days of exposure. These aspects of the evidence with regard to exposure duration were cited in establishing the 1-hour averaging time for the current primary NAAQS for SO_x.

Substantial inter-individual variability was observed in controlled human exposure studies of SO₂ and respiratory effects, but there was a clear increase in the magnitude of respiratory effects with increasing exposure concentrations between 200 and 1,000 ppb during 5–10 minute SO₂ exposures ([Section 5.2.1.2](#)). Both the number of affected individuals with asthma and the severity of the response increased as SO₂ concentrations increased. Epidemiologic studies evaluating the shape of the ambient air concentration-response function have found no evidence for a population-level threshold or nonlinearity, although the evidence is limited.

SO₂ concentrations in ambient air are highly spatially heterogeneous, with SO₂ concentrations at some monitors possibly not highly correlated with the community average concentration ([Section 3.4.2.2](#)). The predominance of point sources results in an uneven distribution of SO₂ concentrations across an urban area. This spatial and temporal variability in SO₂ concentrations can contribute to exposure error in epidemiologic studies, whether the studies rely on fixed-site monitor data or concentration modeling for exposure assessment.

Consistent with the findings of the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), this ISA concludes there is adequate evidence that people with asthma, particularly children, are at increased risk for SO₂-related health effects compared with those without asthma ([Chapter 6](#)). This conclusion is based on the evidence for short-term SO₂ exposure and respiratory effects (specifically lung function decrements), for which a causal relationship has been determined. The ISA concludes there is suggestive evidence that children are at increased risk for SO₂-related health effects, based on their increased ventilation rates relative to body mass and increased oral breathing, together with some epidemiologic evidence of increased associations between SO₂ and respiratory effects relative to adults, even though recent epidemiologic evidence is less consistent. There is also evidence from epidemiologic studies of respiratory hospitalizations, particularly among adults older than 75 years, suggestive of increased risk of SO₂-related health effects for older adults relative to other lifestages.

CHAPTER 1 INTEGRATIVE SYNTHESIS OF THE ISA

1.1 Purpose and Overview of the Integrated Science Assessment

The Integrated Science Assessment (ISA) is a comprehensive evaluation and synthesis of the policy-relevant science “useful in indicating the kind and extent of all identifiable effects on public health or welfare which may be expected from the presence of [a] pollutant in the ambient air,” as described in Section 108 of the Clean Air Act ([CAA, 1990a](#)).¹ This ISA communicates critical science judgments of the health-related air quality criteria for the broad category of sulfur oxides (SO_x). As such, this ISA serves as the scientific foundation for the review of the current primary (health-based) National Ambient Air Quality Standard (NAAQS) for SO_x. Consistent with Section 109(d)(1) of the Clean Air Act, this final ISA contains the air quality criteria addressing the human health effects of SO_x for the current review and reflects the EPA’s periodic review of those criteria. SO_x include several related gaseous compounds such as sulfur dioxide (SO₂) and sulfur trioxide (SO₃) ([Section 2.3](#)). SO₂ was chosen as the indicator² for the current NAAQS because as in previous reviews, it was identified as the most abundant sulfur oxide species in the atmosphere ([U.S. EPA, 1996b](#); [HEW, 1969](#)),³ and the one for which there is a large body of evidence on health effects following exposure to SO₂ (75 FR 35536). In addition, the 2010 Final Rule concluded that “measures leading to reductions in population exposures to SO₂ can generally be expected to lead to reductions in population exposures to SO_x.” (75 FR 35536). Health effects of particulate sulfur-containing species (e.g., sulfate) are being considered in the current review of the NAAQS for particulate matter (PM) and were previously evaluated in the 2009 ISA for PM ([U.S. EPA, 2009a](#)). The ecological effects of sulfur oxides are being considered in a separate ISA for Oxides of Nitrogen, Oxides of Sulfur, and Particulate Matter-Ecological Criteria ([U.S. EPA, 2017a](#)), while the visibility, climate, and materials-related welfare effects of particulate sulfur compounds are evaluated in the ISA for particulate matter ([U.S. EPA, 2016c](#)).

¹ The general process for developing an ISA, including the framework for evaluating weight of evidence and drawing scientific conclusions and causal judgments, is described in a companion document, *Preamble to the Integrated Science Assessments* ([U.S. EPA, 2015b](#)).

² The four components to a NAAQS are (1) indicator (e.g., SO₂); (2) level (e.g., 75 ppb); (3) averaging time (e.g., 1 h), and (4) form (e.g., 3-year avg of the 99th percentile of each year’s 1-h daily max concentrations).

³ In this ISA, the blue electronic links can be used to navigate to cited chapters, sections, tables, figures, and studies.

This ISA evaluates relevant scientific literature published since the 2008 ISA for Sulfur Oxides [([U.S. EPA, 2008d](#)), or 2008 SO_x ISA], integrating key information and judgments contained in the 2008 SO_x ISA and the 1982 *Air Quality Criteria Document (AQCD) for Particulate Matter and Sulfur Oxides* ([U.S. EPA, 1982a](#)) and its Addenda ([U.S. EPA, 1994, 1986, 1982b](#)). Thus, this ISA updates the state of the science that was available for the 2008 SO_x ISA, which informed decisions on the primary SO₂ NAAQS in the review completed in 2010. In 2010, the U.S. Environmental Protection Agency (U.S. EPA) established a new 1-hour standard of 75 parts per billion (ppb) SO₂ as a 3-year avg of the 99th percentile of each year's 1-hour daily max concentrations.¹ The 1-hour standard was established to protect against an array of respiratory effects associated with short-term exposures in potential at-risk populations such as people with asthma. This standard was based on direct evidence of SO₂-related effects in controlled human exposure studies of exercising individuals with asthma, as well as epidemiologic evidence of associations between SO₂ concentrations in ambient air and respiratory-related emergency department (ED) visits and hospitalizations. The U.S. EPA also revoked the existing 24-hour and annual primary SO₂ standards of 140 and 30 ppb, respectively. This decision was largely based on the recognition that the new 1-hour standard at 75 ppb would generally maintain 24-hour and annual SO₂ concentrations well below the then-existing standards and on the lack of evidence indicating the need for such longer-term standards (75 FR 35549-50).

This new review of the primary NAAQS for SO_x is guided by several policy-relevant questions that are identified in The *Integrated Review Plan for the Primary National Ambient Air Quality Standard for Sulfur Dioxide* ([U.S. EPA, 2014b](#)). To address these questions and update the scientific judgments in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), this ISA aims to:

- Characterize the evidence for health effects associated with short-term (minutes up to 1 month) and long-term (more than 1 month to years) exposure to SO_x by integrating findings across scientific disciplines and across related health outcomes (e.g., respiratory effects) and by considering important uncertainties identified in interpreting the scientific evidence, including the role of SO₂ within the broader mixture of pollutants in the ambient air.
- Inform policy-relevant issues related to quantifying health risks, such as exposure concentrations, durations, and patterns associated with health effects; concentration-response (C-R) relationships and existence of thresholds below which effects do not occur; and populations and lifestages potentially with increased risk of health effects related to exposure to SO_x.

Sulfur dioxide is the most abundant species of SO_x in the atmosphere, with other SO_x species being much less prevalent ([Section 2.1](#)). Nearly all studies on the health effects of

¹ The legislative requirements and history of the SO₂ NAAQS are described in detail in the [Preface](#) to this ISA.

SO_x focus on SO₂. In evaluating the health evidence, this ISA considers possible influences of other atmospheric pollutants, including interactions of SO₂ with co-occurring pollutants, such as PM (including particulate sulfur compounds), nitrogen oxides (NO_x), carbon monoxide (CO), and ozone (O₃).

In addressing policy-relevant questions, this ISA aims to characterize the independent health effects of SO₂. As described in this ISA, recent evidence continues to support a causal relationship between short-term SO₂ exposure and respiratory effects based on the consistency of findings; coherence among evidence from controlled human exposure, epidemiologic, and toxicological studies; and biological plausibility for effects specifically related to asthma exacerbation. The information summarized in this ISA will serve as the scientific foundation for the review of the current primary 1-hour SO₂ NAAQS.

1.2 Process for Developing The Integrated Science Assessment

The U.S. EPA uses a structured and transparent process to evaluate scientific information and determine the causal nature of relationships between air pollution exposures and health effects [details provided in the Preamble to the *Integrated Science Assessments* ([U.S. EPA, 2015b](#))]. The ISA development process describes approaches for literature searches, criteria for selecting and evaluating relevant studies, and a framework for evaluating the weight of evidence and forming causal determinations. As part of this process, the ISA is reviewed by the Clean Air Scientific Advisory Committee (CASAC), which is a formal independent panel of scientific experts, and by the public. As this ISA informs the review of the primary SO₂ NAAQS, it integrates and synthesizes information characterizing exposure to SO₂ and potential relationships with health effects. Relevant studies include those examining atmospheric chemistry, spatial and temporal trends, and exposure assessment, as well as U.S. EPA analyses of air quality and emissions data. Relevant health research includes epidemiologic, controlled human exposure, and toxicological studies on health effects, as well as studies on dosimetry and modes of action.

The U.S. EPA initiated the current review of the primary NAAQS for SO_x in August 2013 with a call for information from the public ([U.S. EPA, 2013c](#)). Thereafter, the U.S. EPA routinely conducted literature searches to identify relevant peer-reviewed studies published since the previous ISA (i.e., from January 2008 through August 2016). Multiple search methods were used [Preamble to the ISAs ([U.S. EPA, 2015b](#)), Section 2], including searches in the PubMed and Web of Science databases. Subject-area experts and the public were also able to recommend studies and reports during a science/policy

issue “kick-off” workshop held by the U.S. EPA in June 2013. The U.S. EPA identified additional studies considered to be definitive on particular topics from previous assessments to include in this ISA. Studies that did not address a topic described in the preceding paragraph based on title were excluded. Studies that were judged to be potentially relevant based on review of the abstract or full text and “considered” for inclusion in the ISA are documented and can be found at the Health and Environmental Research Online (HERO) website. The HERO project page for this ISA (<https://hero.epa.gov/hero/sulfur-oxides>) contains the references that are cited in the ISA, the references that were considered for inclusion but not cited, and electronic links to bibliographic information and abstracts.

Categories of health effects were considered for evaluation in this ISA if they were examined in previous U.S. EPA assessments for SO_x or in multiple recent studies. For other categories of health effects, literature searches were conducted to determine the extent of available health evidence. These searches identified a few recently published epidemiologic studies on outcomes such as migraine/headache, depression, suicide, eye irritation/conjunctivitis, rheumatic disease, and gastrointestinal disorders [Supplemental Table 5S-1 ([U.S. EPA, 2016f](#))]. Literature searches have also identified a few recently published toxicological studies on hematological effects, mRNA and protein expression in the brain, sensory symptoms, and effects in other organs (e.g., liver, spleen) [Supplemental Table 5S-2 ([U.S. EPA, 2015c](#))]. These health effects are not evaluated in the current ISA because relationship is lacking between the toxicological and epidemiological health effects examined in these studies and because the results have a large potential for publication bias (i.e., a greater likelihood of publication for studies showing effects compared with those showing no effect).

The Preamble to the ISAs ([U.S. EPA, 2015b](#)) describes the general framework for evaluating scientific information, including criteria for assessing study quality and developing scientific conclusions. Aspects specific to evaluating studies of SO_x are described in the Annex to Chapter 5, Table A-1. For epidemiologic studies, emphasis is placed on studies that (1) characterize quantitative relationships between SO₂ and health effects, (2) examine exposure metrics that well represent the variability in concentrations in the study area, (3) consider the potential influence of other air pollutants and factors correlated with SO₂, (4) examine potential at-risk populations and lifestages, or (5) combine information across multiple cities. With respect to the evaluation of controlled human exposure and animal toxicological studies, emphasis is placed on studies that examine effects relevant to humans and SO₂ concentrations relevant to ambient exposures (i.e., exposures to SO₂ in ambient air). Based on peak ambient air concentrations ([Section 2.5](#)) and the ISA’s emphasis on ambient-relevant exposures within one to two orders of magnitude of current ambient concentrations, SO₂

concentrations of 2,000 ppb¹ or less are defined to be ambient-relevant. Experimental studies with higher exposure concentrations were included if they contributed to an understanding of dosimetry or potential modes of action. For the evaluation of human exposure to ambient SO₂, emphasis is placed on studies that examine the quality of data sources used to assess exposures, such as fixed-site monitors, personal exposure monitors, and dispersion models. The ISA also emphasizes studies that examine factors that influence exposure, such as time-activity patterns and building ventilation characteristics.

The ISA draws conclusions about relationships between SO₂ exposure and health effects by integrating information across scientific disciplines and related health outcomes and synthesizing evidence from previous and recent studies. Determinations are made about causation, not just association, and are based on judgments of aspects such as the consistency, coherence, and biological plausibility of observed effects (i.e., evidence for effects on key events in the mode of action) as well as related uncertainties. The ISA uses a formal causal framework [Table II of the Preamble to the ISAs ([U.S. EPA, 2015b](#))] to classify the weight of evidence according to the five-level hierarchy summarized below.

- **Causal relationship:** the consistency and coherence of evidence integrated across scientific disciplines and related health outcomes are sufficient to rule out chance, confounding, and other biases with reasonable confidence.
- **Likely to be a causal relationship:** there are studies in which results are not explained by chance, confounding, or other biases, but uncertainties remain in the evidence overall. For example, the influence of other pollutants is difficult to address, or evidence across scientific disciplines may be limited or inconsistent.
- **Suggestive of, but not sufficient to infer, a causal relationship:** evidence is generally supportive but not entirely consistent or is limited overall. Chance, confounding, and other biases cannot be ruled out.
- **Inadequate to infer a causal relationship:** there is insufficient quantity, quality, consistency, or statistical power of results from studies.
- **Not likely to be a causal relationship:** several adequate studies, examining the full range of anticipated human exposure concentrations and potential at-risk populations and lifestages, consistently show no effect.

¹ The 2,000-ppb upper limit applies largely to animal toxicological studies but also a few controlled human exposure studies.

1.3 Organization of the Integrated Science Assessment

This ISA comprises the [Preface](#) (legislative requirements of the NAAQS and history of the primary NAAQS for SO_x), [Executive Summary](#), and six chapters. This chapter ([Chapter 1](#)) synthesizes the scientific evidence that best informs policy-relevant questions that frame this review of the primary NAAQS for SO_x. [Chapter 2](#) characterizes the sources, atmospheric processes involving SO_x, and trends in ambient air concentrations. [Chapter 3](#) describes methods to estimate human exposure to SO_x and the impact of error in estimating exposure on relationships with health effects. [Chapter 4](#) describes the dosimetry and modes of action for SO₂. [Chapter 5](#) evaluates and integrates epidemiologic, controlled human exposure, and toxicological evidence for health effects related to short-term and long-term exposure to SO_x. [Chapter 6](#) evaluates information on potential at-risk populations and lifestyles. In addition, the Preamble to the ISAs ([U.S. EPA, 2015b](#)) describes the general process for developing an ISA.

The purpose of this chapter is not to summarize each of the chapters but to synthesize the key findings for each topic that helped characterize SO₂ exposure and relationships with health effects. This chapter also integrates information across the ISA to inform policy-relevant issues such as SO₂ exposure metrics associated with health effects, concentration-response relationships, and the public health impact of SO₂-related health effects ([Section 1.7](#)). A key consideration in the health effects assessment is the extent to which evidence indicates that SO₂ exposure independently causes health effects. To that end, this chapter draws upon information about the sources, distribution, and exposure to ambient SO₂ and identifies pollutants and other factors related to the distribution of or exposure to SO₂ in the ambient air that can potentially influence epidemiologic associations observed between health effects and SO₂ exposure ([Section 1.4](#)). The chapter also summarizes information on the dosimetry and mode of action of inhaled SO₂ that can provide biological plausibility for observed health effects ([Section 1.5](#)). The discussions of the health effects evidence and causal determinations ([Section 1.6](#)) describe the extent to which epidemiologic studies accounted for factors that may influence epidemiologic study results and the extent to which findings from controlled human exposure and animal toxicological studies support independent relationships between SO₂ exposure and health effects.

1.4 From Emissions Sources to Exposure to Sulfur Dioxide

Characterizing human exposure is key to understanding the relationships between ambient SO₂ exposure and health effects. The sources of SO_x and the transformations that occur in ambient air influence the spatial and temporal pattern of SO₂ concentrations.

These patterns have implications for variation in exposure in the population, the adequacy of methods used to estimate exposure, and in turn, the strength of inferences that can be drawn about health effects related to SO₂ exposure.

1.4.1 Emission Sources and Distribution of Ambient Concentrations

Emissions of SO₂ have declined by approximately 79% for all sources from 1990 to 2014 subsequent to several U.S. air quality regulatory programs. Coal-fired electricity generation units (EGUs) remain the dominant sources by nearly fivefold above the next highest source (industrial fuel combustion), emitting 3.2 million tons of SO₂ annually, according to the 2014 National Emissions Inventory (NEI; [Section 2.2](#)). Preliminary estimates through 2016 suggest further declines in emissions, particularly for EGUs.

In addition to source characteristics, such as emission rate, stack height, and plume temperature, the two important variables that determine the concentration of SO₂ downwind of the source are the photochemical and other removal processes (e.g., formation of particle-phase reduced sulfur compounds) occurring in the emissions plume and the local meteorology, including wind, atmospheric stability, humidity, and cloud/fog cover ([Section 2.3](#)). The primary gas-phase photochemical SO₂ oxidation mechanism requires the hydroxyl radical (OH). Another oxidation mechanism involves a Criegee intermediate biradical that participates in converting SO₂ to SO₃, which rapidly reacts with water vapor to form sulfuric acid (H₂SO₄). The Criegee-based SO₂ oxidation mechanism may amplify the rate of SO₂ removal and formation of organosulfur compounds in areas with high concentrations of Criegee precursors (i.e., low-molecular-weight organic gases, such as biogenic compounds and unsaturated hydrocarbons present downwind of industrial sites and refineries). Aqueous-phase oxidation of SO₂ is also an important removal mechanism. Clouds and fog can reduce local SO₂ concentrations by converting it to H₂SO₄ in the droplet phase.

Changes were undertaken to the existing U.S. EPA monitoring network as a result of decisions in the 2010 NAAQS review ([Section 2.4](#)). First, the automated pulsed ultraviolet fluorescence (UVF) method, which is most commonly used by state and local monitoring agencies for NAAQS compliance, was designated as a federal reference method (FRM). Second, new SO₂ monitoring guidelines require state and local agencies operating continuous SO₂ analyzers to report either the highest 5-minute concentration for each hour of the day or all twelve 5-minute concentrations for each hour of the day. Analysis of environmental concentrations of SO₂ data reported in [Section 2.5](#) reflect the monitoring network changes, particularly the analysis of the recent 5-minute data.

On a nationwide basis, the average 1-h daily max SO₂ reported during 2013–2015 was 5.3 ppb with a 99th percentile concentration of 64 ppb ([Section 2.5](#)). However, 99th percentile 1-h daily max SO₂ concentrations were greater than 75 ppb at some monitoring sites located near large anthropogenic sources (e.g., power plants or smelters). Volcanoes, a large natural source of SO₂, can produce nearby hourly concentrations over 2,000 ppb. Nationally, the 99th percentile 5-minute hourly max concentration for 2013–2015 was 24.0 ppb, suggesting that concentrations of 100 ppb or more are relatively rare nationwide, although monitoring sites near large sources had 99th percentile 5-min hourly max concentrations above 100 ppb. For example, the four monitors in Gila County, AZ located near smelters had 99th percentile 5-min hourly max concentrations ranging from 116 to 252 ppb. Correlations between hourly 5-minute max SO₂ concentrations and their corresponding 1-h avg concentrations were high, with approximately 75% of sites having correlations greater than 0.9. Peak-to-mean ratios (PMRs) between the two metrics were generally less than 3, although higher PMRs were observed during some hours ([Section 2.5.4](#)). Background concentrations of SO₂ from natural sources and sources outside the U.S. are very low across most of the country (annual average less than 0.03 ppb), accounting for less than 1% of ambient air concentrations except in areas influenced by local cross-border sources and areas where volcanic emissions are important, such as Hawaii and parts of the West Coast ([Section 2.5.5](#)).

SO₂ concentrations are highly variable across urban spatial scales, exhibiting moderate to poor correlations between SO₂ measured at different monitoring sites across a metropolitan area. This high degree of urban spatial variability may not be fully captured by fixed-site monitors used in epidemiologic studies, and thus, has implications for the interpretation of human exposure and health effects data ([Sections 2.5.2.2](#) and [3.4.4](#)).

Air quality models, especially dispersion models, can be used to estimate SO₂ concentrations over various averaging times in locations where monitoring is not practical or sufficient ([Section 2.6](#)). Because existing monitors may not be sited in locations to capture peak 1-hour concentrations, the implementation program for the 2010 primary NAAQS for SO_x allows for air quality modeling to be used to characterize air quality for informing designation decisions (75 FR 35520). In addition, modeling is critical for assessing the impact of future sources or proposed modifications where monitoring cannot inform and for designing and implementing mitigation techniques. Dispersion models have also been used to estimate SO₂ exposure concentrations in epidemiologic studies, particularly in long-term studies ([Section 3.3.2.4](#), [Chapter 5](#)). The widely used dispersion model, the American Meteorological Society/U.S. EPA Regulatory Model (AERMOD), is based on Gaussian dispersion models but includes advancements such as the ability to incorporate boundary layer scaling formulations,

surface and elevated emission points, interactions of plumes with buildings and terrain, and source geometry. Several evaluations of the performance of AERMOD against field study data over averaging times from 1 hour to 1 year found the model was relatively unbiased in estimating upper percentage 1-hour concentration values. Lagrangian puff dispersion models, such as CALPUFF, have been developed as an alternative to Gaussian dispersion models. CALPUFF models SO₂ as a tracer and then uses a Lagrangian step algorithm to model non-steady-state dynamics, using time-varying winds specified by meteorological models. CALPUFF simulations were found to improve in accuracy with increasing integration times. In some cases, CALPUFF predictions are closer to measured SO₂ concentrations compared with AERMOD, but other evaluations have found larger bias with CALPUFF than AERMOD. Uncertainties in model predictions are influenced by uncertainties in model input data, particularly emissions and meteorological conditions (e.g., wind).

1.4.2 Assessment of Human Exposure

Multiple techniques can be used to assign exposure for epidemiologic studies. Evaluation of data from fixed-site or personal SO₂ monitors is commonly used to derive an estimate of exposure. Various modeling approaches may also be used ([Section 3.3](#)). Each has strengths and limitations, as summarized in [Table 3-1](#). Fixed-site monitors may be intended to represent population exposure, although some monitors are located near sources to capture high concentrations locally and are not typically used as the primary data source in urban-scale epidemiologic studies. Fixed-site monitors may provide a continuous record of SO₂ concentrations over many years, but due to limited spatial coverage, they may not fully capture the relatively high spatial variability in SO₂ concentration across an urban area. Personal SO₂ monitors can capture the study participants' activity-related exposure across different microenvironments, but low ambient SO₂ concentrations often result in a substantial fraction of the samples being below the limit of detection for averaging times of 24 hours or less. The time and expense involved to deploy personal monitors make them more suitable for panel epidemiologic studies than for large-scale time-series or cohort studies. Models can be used to estimate exposure for individuals and large populations when personal exposure measurements are unavailable. In general, more complex approaches provide more detailed exposure estimates but may require additional input data, assumptions, and/or computational resources. Depending on the model type, there is the potential for bias and reduced precision due to model misspecification, missing sources, smoothing of concentration gradients, and complex topography. Evaluation of model results helps demonstrate the suitability of that approach for particular applications.

New studies of the relationship between indoor and outdoor SO₂ concentrations have focused on publicly owned buildings rather than residences ([Section 3.4.1.2](#)). The results of these studies are consistent with results of previous studies showing that indoor:outdoor ratios and indoor-outdoor regression slopes cover an extremely wide range, from near zero to near one. Differences in results among studies are due to building characteristics (e.g., forced ventilation, building age, and building type), personal activities such as opening windows and doors, and SO₂ measurement limitations. Due to indoor deposition and a relative lack of indoor sources of SO₂, indoor concentrations are often much lower than outdoor SO₂ concentrations. These low indoor concentrations also contribute to low personal exposure concentrations due to time spent indoors. When reported, correlations between indoor and outdoor concentrations were relatively high (>0.75), suggesting that variations in outdoor concentration drive indoor concentrations, particularly considering the lack of indoor SO₂ sources. These high correlations were observed across seasons and geographic locations. The bulk of the evidence for personal-ambient SO₂ relationships was available at the time of the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) and shows a wide range of correlations between ambient concentration and personal exposure, in part due to a large fraction of samples below the method detection limit (MDL) in several studies ([Section 3.4.1.3](#)). When nearly all of the personal samples are below the MDL, no correlation can be observed. However, when the bulk of the personal samples are above the MDL, personal exposure is moderately correlated ($r = 0.66$) with ambient concentration.

“Exposure error” refers to the bias and uncertainty associated with using concentration metrics to represent the actual exposure of an individual or population [([Lipfert and Wyzga, 1996](#)) [Section 3.2](#)]. Exposure error has two components: (1) exposure measurement error derived from uncertainty in the metric being used to represent exposure and (2) use of a surrogate target parameter of interest in the epidemiologic study in lieu of the true exposure, which may be unobservable ([Section 3.2.1](#)). Factors that could contribute to error in estimating exposure to ambient SO₂ include time-location-activity patterns, spatial and temporal variability in SO₂ concentrations, and proximity of populations to monitoring sites and sources ([Section 3.4.2](#)). Activity patterns vary both among and within individuals, resulting in corresponding variations in exposure across a population and over time. Variation in SO₂ concentrations among different microenvironments means that the amount of time spent in each location, as well as exertion level, will influence an individual’s exposure to ambient SO₂. Time spent in different locations has also been found to vary by age, with younger and older age groups spending a greater percentage of time outdoors than adults of typical working age (18–64 years). These variations in activity pattern contribute to differences in exposure and, if uncharacterized, introduce error into population-averaged exposure estimates.

Uncharacterized spatial and temporal variability in SO₂ concentrations can contribute to exposure error in epidemiologic studies. SO₂ has low to moderate spatial correlations among ambient air monitoring sites across urban geographic scales; thus, using fixed-site monitor data for epidemiologic exposure assessment introduces exposure error into the resulting health effect estimate. Spatial variability in the magnitude of concentrations may affect cross-sectional and large-scale cohort studies by assigning exposures from one or a small number of sites that do not capture all of the spatial variability within a city. Some models are designed to address this situation by improving the characterization of spatial and temporal variability. This issue may be less important for time-series studies, which rely on day-to-day temporal variability in concentrations to evaluate health effects.

Proximity of populations to ambient air monitoring sites may influence how well human exposure is represented by measurements at the monitors, although factors other than distance play an important role as well. Many ambient SO₂ monitoring sites are located near dense population centers, but other near-source sites may not be near population centers. Use of monitoring sites in epidemiologic studies introduces exposure error into health effect estimates. The literature has shown that exposure error and related bias in the health effect estimate is reduced by using averaging schemes in lieu of a single fixed-site monitor ([Section 3.4.2.2](#)).

Exposure to copollutants, such as other criteria pollutants, may result in confounding of health effect estimates. For SO₂, daily concentrations generally exhibit low correlations (median <0.4) with other daily NAAQS pollutant concentrations at collocated monitors ([Figure 3-5](#), [Section 3.4.3](#)). However, a wide range of copollutant correlations has been observed across different monitoring sites, from moderately negative to moderately positive. Data showing high daily SO₂ correlations with nitrogen dioxide (NO₂) and CO may have been collected before the rule to reduce sulfur content in diesel fuel (66 FR 5002) took effect in 2006 and 2007. The minority of sites with stronger correlations may introduce a greater degree of confounding into epidemiologic results. A similar impact is expected for epidemiologic studies of long-term SO₂ exposure, which also report a wide range of copollutant correlations.

Exposure error can influence epidemiologic study results by biasing effect estimates either toward or away from the null and widening confidence intervals beyond the nominal coverage that would be produced if the true exposure had been used ([Section 3.4.4](#)). The exposure error varies according to the study design, especially regarding the study's spatial and temporal aspects. For example, in time-series and panel studies, low personal-ambient correlations tend to bias the effect estimate toward the null, while spatial variation in personal-ambient correlations across an urban area contributes to widening of the confidence interval around the effect estimate so that the nominal

coverage of the confidence intervals is below 95% for exposure effect estimates. For long-term studies, bias of the health effect estimate may occur in either direction depending on whether the monitor is over- or underestimating exposure for the population of interest. In all study types, use of fixed-site monitors is expected to decrease precision of the health effect estimate because spatial variation in personal-ambient correlations across an urban area contributes to widening of the confidence interval around the effect estimate so that the nominal coverage is below 95% for exposure effect estimates.

Choice of exposure estimation method also influences the impact of exposure error on epidemiologic study results. Fixed-site monitors offer a convenient source of time-series data, but fixed-site measurements do not account for the effects of spatial variation in SO₂ concentration, differences between indoor and outdoor exposure to ambient SO₂, and varying activity patterns on personal exposure to SO₂. Personal exposure measurements, such as those made in panel epidemiologic studies, provide accurate and specific exposure estimates, but sample size is often small, and only a limited set of health outcomes can be studied. Modeled concentrations or exposures offer alternatives to measurements, with the advantage of estimating exposures over a wide range of scales, populations, and scenarios, particularly for locations lacking monitoring data. However, depending on the model type, there is the potential for bias and reduced precision due to model misspecification, missing sources, smoothing of concentration gradients, and complex topography. Model estimates for unmonitored locations or time periods are most informative when the model output has been compared to an independent set of measured concentrations or exposures. The various sources of exposure error and their potential impact are considered in evaluating the epidemiologic study results in [Chapter 5](#) of this ISA.

1.5 Dosimetry and Mode of Action of Sulfur Dioxide

Dosimetry of inhaled SO₂, including the processes of absorption, distribution, metabolism, and elimination, as well as information on the mode of action of inhaled SO₂, covering the processes by which inhaled SO₂ initiates a cascade of molecular and cellular responses and subsequent organ-level responses, is covered in [Chapter 4](#). Together, these sections provide the foundation for understanding how exposure to inhaled SO₂ can lead to health effects. This understanding enables our characterization of the biological plausibility of SO₂ exposure as the cause of health effects that may be observed in epidemiologic studies.

1.5.1 Dosimetry of Inhaled Sulfur Dioxide

Dosimetry of inhaled SO₂ refers to the measurement or estimation of the amount of SO₂ and its reaction products reaching and/or persisting at specific sites within the respiratory tract and systemically after exposure. Factors affecting the transport and fate of SO₂ in the respiratory tract include respiratory tract morphology, respiratory functional parameters, and the physicochemical properties of SO₂ and epithelial lining fluid (ELF). Health effects may be due to inhaled SO₂ or its chemical reaction products, including sulfite and S-sulfonates. Few studies have investigated SO₂ dosimetry since the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), with most studies conducted prior to the 1982 AQCD ([U.S. EPA, 1982a](#)) and the 1986 Second Addendum ([U.S. EPA, 1986](#)).

Because SO₂ is highly soluble in water, it is readily absorbed in the nasal passages of both humans and laboratory animals under resting conditions ([Section 4.2.2](#)). During nasal breathing, the majority of available data suggests 95% or greater SO₂ absorption occurs in the nasal passages, even under ventilation levels comparable to that during exercise. With increasing physical activity, there is an increase in ventilatory rate and a shift from nasal to oronasal breathing, resulting in greater SO₂ penetration into the lower respiratory tract. Even at rest, differences have been observed by age, sex, disease status, and body mass index in the fraction of oral versus nasal breathing ([Section 4.1.2](#)). Children inhale a larger fraction of air through their mouth than adults, and males tend to inhale a larger fraction of air through their mouth than females (across all ages). Individuals with allergies or upper respiratory infections experience increased nasal resistance, and thus, increased fraction of oral breathing. Obesity, especially in boys, may also contribute to increased nasal resistance and an increased oral fraction of breathing relative to normal weight children. Due to their increased amount of oral breathing, these individuals may be expected to have greater SO₂ penetration into the lower respiratory tract than healthy, normal weight adults. Children may also be expected to have a greater intake dose of SO₂ per body mass than adults.

Following absorption in the respiratory tract, SO₂ rapidly forms a mixture of bisulfite and sulfite, with the latter predominating. As much as 15–18% of the absorbed SO₂ may be desorbed and exhaled following cessation of exposure. Although some SO₂ products rapidly move from the respiratory tract into the blood and are distributed about the body, experiments using radiolabeled ³⁵S indicate that the majority of sulfur in SO₂-derived products in the body at any given time following exposure is found in the respiratory tract and may be detected there for up to a week following inhalation ([Section 4.2.3](#)). The distribution and clearance of inhaled SO₂ from the respiratory tract may involve several intermediate chemical reactions and transformations, particularly the formation of sulfite and S-sulfonates. Sulfite is metabolized into sulfate, primarily in the liver, which has

higher sulfite oxidase levels than the lung or other body tissues ([Section 4.2.4](#)). Sulfite oxidase activity is highly variable between species, with liver sulfite oxidase activity in rats being 10–20 times greater than in humans. Urinary excretion of sulfate is rapid and proportional to the concentration of SO₂ products in the blood ([Section 4.2.5](#)). S-sulfonates are cleared more slowly from the circulation with a clearance half-time of days.

Sulfite levels in the body are predominately influenced by endogenous production and ingestion of sulfite in food ([Section 4.2.6](#)). The primary endogenous contribution of sulfite is from the catabolism of sulfur-containing amino acids (namely, cysteine and methionine). Endogenous sulfite from ingested sulfur-containing amino acids far exceeds exogenous sulfite from ingestion of food additives [by 140 and 180 times in adult (19–50 years) females and males, respectively, and by 500 times or more in young children (1–3 years)]. Endogenous sulfite production is two or more orders of magnitude higher than inhalation-derived sulfite levels for both children and adults, even for full day exposures to 75 ppb SO₂ (the level of the 1-hour NAAQS). Ingestion rates of sulfite added to foods vary widely; however, in general, sulfite ingestion is expected to exceed sulfite intake from inhalation in adults and children even for full day exposures to 75 ppb SO₂. However, sulfite and sulfate from ingestion or endogenous production do not accumulate primarily in respiratory tract tissues, as is the case for inhalation-derived SO₂ products.

1.5.2 Mode of Action of Inhaled Sulfur Dioxide

Mode of action refers to a sequence of key events, endpoints, and outcomes that result in a given toxic effect. The mode of action discussion in [Section 4.3](#) of this ISA updates the basic concepts derived from the SO₂ literature presented in the 1982 AQCD ([U.S. EPA, 1982a](#)) and the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) and introduces the recent relevant literature. The main effects of SO₂ inhalation are seen at the sites of absorption (i.e., the respiratory tract) and include (1) activation of sensory nerves in the respiratory tract resulting in a neural reflex response, (2) injury to airway mucosa, and (3) increased airway responsiveness and allergic inflammation. Effects outside the respiratory tract may occur at very high concentrations of inhaled SO₂.

Reactive products formed as a result of SO₂ inhalation are responsible for a variety of downstream key events, which may include activation of sensory nerves in the respiratory tract, release of inflammatory mediators, and modulation of allergic inflammation or sensitization. These key events may collectively lead to several endpoints, including bronchoconstriction and increased airway responsiveness. A

characteristic feature of individuals with asthma is an increased propensity of their airways to narrow in response to bronchoconstrictive stimuli relative to nonasthmatic individuals without asthma. Thus, bronchoconstriction is characteristic of an asthma attack. However, individuals without asthma may also experience bronchoconstriction in response to SO₂ inhalation; generally this occurs at higher concentrations (>1,000 ppb) than in an individual with asthma. Additionally, SO₂ exposure may increase airway responsiveness to subsequent exposures of other stimuli such as allergens or methacholine. These pathways may play a role in the epidemiologic study outcome of asthma exacerbation.

The strongest evidence for the mode of action for respiratory effects following short-term exposure comes from controlled human exposure studies. SO₂ exposure resulted in increased airway resistance due to bronchoconstriction in adults, both with and without asthma. In adults without asthma, this response occurred primarily as a result of activation of sensory nerves in the respiratory tract resulting in neural reflex responses ([Section 4.3.1](#)). This is mediated by cholinergic parasympathetic pathways involving the vagus nerve. However, in adults with asthma, evidence indicates that the response is only partially due to vagal pathways and that inflammatory mediators such as histamine and leukotrienes also play an important role. Studies in experimental animals also demonstrate that SO₂ exposure activates reflexes that are mediated by cholinergic parasympathetic pathways involving the vagus nerve. However, noncholinergic mechanisms (i.e., neurogenic inflammation) may also be involved.

Evidence demonstrates that SO₂ exposure modulates allergic inflammatory responses ([Section 4.3.2](#)). Enhancement of allergic inflammation (i.e., leukotriene-mediated increases in numbers of sputum eosinophils) has been observed in adults with asthma who were exposed for 10 minutes to 750 ppb SO₂. In an animal model of allergic airway disease, repeated exposure to 2,000 ppb SO₂ led to an enhanced inflammatory response, including allergic inflammation. In naive animals, repeated exposure to SO₂ (as low as 100 ppb) over several days promoted allergic sensitization, inflammation, and AHR when animals were subsequently sensitized and challenged with an allergen. Thus, allergic inflammation and increased airway responsiveness may also link short-term SO₂ exposure to asthma exacerbation.

Evidence for the mode of action for respiratory effects due to long-term SO₂ exposure comes from studies in both naive and allergic experimental animals, which demonstrate allergic sensitization, allergic inflammation, AHR, and morphologic changes suggestive of airway remodeling following exposure to SO₂ (i.e., 2,000 ppb) over several weeks ([Section 4.3.3](#)). These changes, however, are mild compared to histopathological changes, such as mucous cell metaplasia and intramural fibrosis, which are generally

observed following chronic exposure of naive animals to extremely high SO₂ concentrations (10,000 ppb and higher). However, in allergic animals, exposure to SO₂ over several weeks leads to morphologic responses indicative of airway remodeling and to AHR ([Section 4.3.6](#)). Thus, repeated exposure to SO₂ may lead to the development of allergic airway disease, which shares many features with asthma, and to the worsening of the allergic airway disease. The development of AHR may link long-term exposure to SO₂ to the epidemiologic outcome of new onset asthma.

Although some evidence suggests that SO₂ inhalation results in extrapulmonary effects, there is uncertainty regarding the mode of action underlying these responses ([Section 4.3.4](#)). Evidence from controlled human exposure studies points to SO₂ exposure-induced activation/sensitization of neural reflexes, possibly leading to altered heart rate (HR) or heart rate variability (HRV). Evidence also points to transport of sulfite into the circulation. Sulfite is highly reactive and may be responsible for redox stress (possibly through autoxidation or peroxidase-mediated reactions to produce free radicals) in the circulation and extrapulmonary tissues. However, this stress is likely to occur only at very high SO₂ concentrations or during prolonged exposures because circulating sulfite is efficiently metabolized to sulfate in a reaction catalyzed by hepatic sulfite oxidase.

1.6 Health Effects of Sulfur Dioxide

This ISA evaluates relationships between an array of health effects and short-term and long-term exposures to SO₂ as examined in epidemiologic, controlled human exposure, and animal toxicological studies. Short-term exposures are defined as those with durations of minutes up to 1 month, with most studies examining effects related to exposures in the range of 1 hour to 1 week. Long-term exposures are defined as those with durations of more than 1 month to years. Drawing from the health effects evidence described in detail in [Chapter 5](#), information on dosimetry and modes of action presented in [Chapter 4](#), as well as issues regarding exposure assessment and potential confounding described in [Chapter 3](#), the subsequent sections and [Table 1-1](#) present the key evidence that informed the causal determinations for relationships between SO₂ exposure and health effects.

1.6.1 Respiratory Effects

1.6.1.1 Respiratory Effects Associated with Short-Term Exposure to Sulfur Dioxide

Strong scientific evidence indicates that there is a causal relationship between short-term SO₂ exposure and respiratory morbidity, particularly in individuals with asthma, which is consistent with the conclusions of the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). This determination is based on the consistency of findings within disciplines, coherence among evidence from controlled human exposure, epidemiologic, and toxicological studies, and biological plausibility for effects specifically related to asthma exacerbation ([Table 5-21](#)).

This conclusion is primarily based on controlled human exposure studies included in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) that showed lung function decrements and respiratory symptoms in individuals with asthma exposed to SO₂ for 5–10 minutes under increased ventilation conditions. No new controlled human exposure studies have since been conducted to evaluate the effect of SO₂ on respiratory morbidity among individuals with asthma. The existing studies consistently demonstrate that individuals with asthma experience a moderate or greater decrement in lung function, defined as a $\geq 100\%$ increase in specific airway resistance (sRaw) or $\geq 15\%$ decrease in forced expiratory volume in 1 sec (FEV₁), frequently accompanied by respiratory symptoms, following peak exposures of 5–10 minutes with elevated ventilation rates at concentrations of 400–600 ppb ([Section 5.2.1.2](#)). A fraction of individuals with asthma (~5–30%) was observed in these studies to have moderate decrements in lung function (i.e., $\geq 100\%$ increase in sRaw or $\geq 15\%$ decrease in FEV₁) at lower SO₂ concentrations (200–300 ppb; [Table 5-2](#)). Lung function decrements at these lower concentrations are less likely to be accompanied by respiratory symptoms. Some studies have evaluated the influence of asthma severity on response to SO₂, but the most severe asthmatics have not been tested, so their response is unknown. Adults with moderate to severe asthma demonstrated larger absolute changes in lung function during exercise in response to SO₂ than adults with mild asthma, although this difference was attributed to a larger response to the exercise component of the protocol rather than to SO₂ itself. While adults with moderate to severe asthma may have similar responses to SO₂ as those with mild asthma, they have less reserve capacity to deal with an insult compared with individuals with mild asthma; therefore, the impact of SO₂-induced decrements in lung function is likely greater in individuals with severe asthma than those with mild asthma. Although there are no laboratory studies of children younger than 12 exposed to SO₂, a number of studies have evaluated airway responsiveness of children and adults to a bronchoconstrictive stimulus.

These studies indicate that school-aged children, particularly boys, are expected to have greater responses (i.e., greater lung function decrements) following exposure to SO₂ than adolescents and adults.

These findings are consistent with the current understanding of dosimetry and modes of action ([Section 1.5](#)). Due to their increased fraction of oral breathing, individuals with asthma may be expected to have greater SO₂ penetration into the lower respiratory tract than healthy adults. Reactive products formed as a result of SO₂ inhalation, particularly sulfites and S-sulfonates, are responsible for a variety of downstream key events, which may include activation of sensory nerves in the respiratory tract resulting in a neural reflex response, release of inflammatory mediators, and modulation of allergic inflammation. These key events may lead to several endpoints including bronchoconstriction and increased airway responsiveness, resulting in the outcome of asthma exacerbation.

Epidemiologic evidence also provides support for a causal relationship, including additional studies that add to the evidence provided by the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). Studies of asthma hospital admissions and ED visits report positive associations with short-term SO₂ exposures, particularly for children (i.e., <18 years of age), with evidence from a limited number of studies that examine potential copollutant confounding indicating that associations remain positive, though are in some instances attenuated in magnitude, in copollutant models involving PM and other criteria pollutants ([Section 5.2.1.2](#)). There is also some supporting evidence for positive associations between short-term SO₂ exposures and respiratory symptoms among children with asthma ([Section 5.2.1.2](#)). Epidemiologic evidence of associations between short-term SO₂ exposures and lung function or respiratory symptoms among adults with asthma is less consistent ([Section 5.2.1.2](#)). Epidemiologic studies of cause-specific mortality that report consistent positive associations between short-term SO₂ exposures and respiratory mortality provide support for a potential continuum of effects ([Section 5.2.1.8](#)).

There is some support for other SO₂-related respiratory effects including exacerbation of chronic obstructive pulmonary disease (COPD) in individuals with COPD and other respiratory effects including respiratory infection, aggregated respiratory conditions, and respiratory mortality in the general population ([Sections 5.2.1.5](#), [5.2.1.6](#), [5.2.1.7](#), and [5.2.1.8](#)). The limited and inconsistent evidence for these nonasthma-related respiratory effects does not contribute substantially to the causal determination.

1.6.1.2 Respiratory Effects Associated with Long-Term Exposure to Sulfur Dioxide

Overall, the evidence is suggestive of, but not sufficient to infer, a causal relationship between long-term SO₂ exposure and respiratory effects, mainly the development of asthma in children ([Section 5.2.2](#)). This represents a change from the conclusion in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) that the evidence was “inadequate to infer a causal association.” There is a limited number of recent longitudinal epidemiologic studies that evaluate associations between asthma incidence among children and long-term SO₂ exposures, with the overall body of evidence lacking consistency. The evidence from longitudinal studies showing increases in asthma incidence is coherent with findings from animal toxicological studies that provide a pathophysiologic basis for the development of asthma. In naive newborn animals, repeated SO₂ exposure over several weeks resulted in immune responses and airway inflammation, key steps in allergic sensitization. In allergic newborn animals, studies with several days or several weeks of repeated SO₂ exposure found enhanced airway inflammation and some evidence of airway remodeling and AHR. The combined epidemiologic and animal toxicological evidence provides support for an independent effect of long-term exposure to SO₂ on the development of asthma in children, but key uncertainties remain, including exposure measurement error and the potential for copollutant confounding. Some evidence of a link between long-term exposure to SO₂ and respiratory symptoms or respiratory allergies among children further supports a possible relationship between long-term SO₂ exposure and the development of asthma. Details of the causal determination are provided in [Table 5-24](#).

1.6.2 Health Effects beyond the Respiratory System

1.6.2.1 Cardiovascular Effects Associated with Short-Term Exposure to Sulfur Dioxide

Overall, the available evidence is inadequate to infer a causal relationship between short-term exposure to SO₂ and cardiovascular health effects ([Table 5-31](#), [Section 5.3.1](#)). This conclusion is consistent with that of the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), which concluded “the evidence as a whole is inadequate to infer a causal relationship.” Although multiple epidemiologic studies report positive associations between short-term exposure to SO₂ and a variety of cardiovascular outcomes, the results are inconsistent across the specific cardiovascular outcomes, and the associations are generally attenuated after copollutant adjustment. There is some experimental evidence in

humans and animals for SO₂-induced effects on the autonomic nervous system and inflammation and other effects in tissues distal to the absorption site. However, the limited and inconsistent evidence from the available experimental studies does not demonstrate potentially biologically plausible mechanisms for, and is not coherent with, cardiovascular effects such as triggering a myocardial infarction. Evidence for other cardiovascular and related metabolic effects is inconclusive.

1.6.2.2 Cardiovascular Effects Associated with Long-Term Exposure to Sulfur Dioxide

Overall, the evidence is inadequate to infer a causal relationship between long-term exposure to SO₂ and cardiovascular health effects ([Table 5-35](#), [Section 5.3.2](#)). The relationship between long-term SO₂ exposure and cardiovascular outcomes was not evaluated in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). Despite a number of epidemiologic studies that report positive associations of long-term exposure to SO₂ concentrations with cardiovascular disease and stroke, the evidence for any one endpoint is limited and inconsistent. Exposure measurement error and the potential for copollutant confounding are uncertainties in the interpretation of the evidence. Additionally, the experimental evidence is insufficient to provide coherence or biological plausibility for an independent effect of long-term exposure to SO₂ on cardiovascular health.

1.6.2.3 Reproductive and Developmental Effects

Overall the evidence is inadequate to infer a causal relationship between exposure to SO₂ and reproductive and developmental outcomes ([Table 5-38](#), [Section 5.4](#)), consistent with the conclusion reached in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)).

There are several recent well-designed, well-conducted studies that indicate an association between SO₂ and reproductive and developmental health outcomes, including fetal growth metrics, preterm birth, birth weight, and fetal and infant mortality. However, a number of uncertainties are associated with the observed relationship between exposure to SO₂ and birth outcomes, such as timing of exposure windows, exposure error, and spatial and temporal heterogeneity. Few studies have examined other health outcomes, such as fertility, effects on pregnancy (e.g., pre-eclampsia, gestational diabetes), and developmental effects, and there is little coherence or consistency among epidemiologic and toxicological studies for these outcomes. There is limited toxicological evidence at relevant dose ranges of SO₂, making it difficult to evaluate the potential modes of action for reproductive and developmental effects of SO₂. Studies published since the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) have not substantially reduced any of the uncertainties identified

in the previous ISA, including exposure measurement error and the potential for copollutant confounding; therefore, the evidence is inadequate to infer a causal relationship between exposure to SO₂ and reproductive and developmental outcomes.

1.6.2.4 Total Mortality Associated with Short-Term Exposure to Sulfur Dioxide

Multicity studies evaluated since the completion of the 2008 ISA for Sulfur Oxides continue to provide consistent evidence of positive associations between short-term SO₂ exposures and total mortality ([Section 5.5.1](#)). Although the body of evidence is larger than at the time of the last review, key uncertainties and data gaps still remain, which contribute to the conclusion that the evidence for short-term SO₂ exposures and total mortality is suggestive of, but not sufficient to infer, a causal relationship ([Table 5-41](#)). This conclusion is consistent with that reached in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Overall, recent multicity studies have further informed key uncertainties and data gaps in the SO₂-mortality relationship identified in the 2008 SO_x ISA including confounding, modification of the SO₂-mortality relationship, potential seasonal differences in SO₂-mortality associations, and the shape of the SO₂-mortality C-R relationship. However, questions remain regarding whether SO₂ has an independent effect on mortality, and these lingering questions can be attributed to the limited number of studies that examined potential copollutant confounding, the relative lack of copollutant analyses with PM_{2.5}, and the evidence indicating attenuation of SO₂-mortality associations in copollutant models with NO₂ and PM₁₀. Additionally, a biological mechanism has not been characterized to date that could lead to mortality as a result of short-term SO₂ exposures.

1.6.2.5 Total Mortality Associated with Long-Term Exposure to Sulfur Dioxide

The overall evidence is inadequate to infer a causal relationship between long-term exposure to SO₂ and total mortality among adults ([Table 5-43](#), [Section 5.5.2](#)), consistent with the conclusion reached in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). Recent evidence is generally consistent with the evidence included in the 2008 SO_x ISA, although some recent cohort epidemiologic studies provide evidence for improved consistency in the association between long-term exposure to SO₂ and both respiratory and total mortality. However, none of these recent studies help to resolve the uncertainties identified in the 2008 SO_x ISA related to exposure measurement error, copollutant confounding, or the geographic scale of the analysis.

1.6.2.6 Cancer

The overall evidence for long-term SO₂ exposure and cancer is inadequate to infer a causal relationship ([Table 5-44](#), [Section 5.6](#)), the same conclusion reached in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). Recent studies include evidence on lung cancer as well as other cancer types. Although some studies of SO₂ concentrations and lung cancer mortality have reported null results, other studies that included various cofounders and copollutants reported positive associations. Positive associations were also observed in a study of SO₂ concentrations and bladder cancer mortality but not in ecological studies of bladder cancer incidence. Limited supportive evidence for mode of action is available from genotoxicity and mutagenicity studies, but animal toxicological studies provide no coherence with epidemiologic findings.

Table 1-1 Key evidence contributing to causal determinations for sulfur dioxide exposure and health effects evaluated in the current Integrated Science Assessment for Sulfur Oxides.

Health Effect Category, ^a Exposure Duration ^b , and Causal Determination	SO ₂ Concentrations Associated with Effects
<p>Respiratory Effects and Short-Term Exposure (Section 5.2.1): <u>Causal relationship</u> <i>No change in causal determination from the 2008 SO_x ISA (U.S. EPA, 2008d); new evidence is consistent with previous determination.</i></p>	
<p>Key evidence (Table 5-21)</p>	<p>Strongest evidence is for effects on asthma exacerbation. There is consistent evidence from multiple high-quality controlled human exposure studies ruling out chance, confounding, and other biases. These studies show decreased lung function and increased respiratory symptoms following peak exposures of 5–10 min in exercising individuals with asthma. Additional consistent evidence from multiple high-quality epidemiologic studies at relevant SO₂ concentrations shows an increase in asthma hospital admissions and ED visits in single- and multicity studies and in studies examining individuals of all ages, including children and older adults. These associations are generally unchanged in copollutant models involving PM and other criteria pollutants. Additionally, there is some supporting epidemiologic evidence of associations with respiratory symptoms among children with asthma. Evidence is available for activation of sensory nerves in the respiratory tract resulting in a neural reflex and/or inflammation leading to bronchoconstriction and allergic inflammation leading to increased airway responsiveness. Enhanced allergic sensitization, allergic inflammation, and airway responsiveness were observed in guinea pigs exposed to SO₂ repeatedly over several days and subsequently sensitized and challenged with an allergen. This evidence represents key events or endpoints in the proposed mode of action linking short-term SO₂ exposure and asthma exacerbation.</p>
<p>Overall study means: <i>Controlled human exposure studies of decreased lung function:</i> 200–600 ppb, with a subset analysis of responders showing statistically significant responses at 300 ppb <i>Controlled human exposure studies of increased respiratory symptoms:</i> 400–1,000 ppb <i>Epidemiologic studies:</i> 1-h max: 9.6–11 ppb 24-h avg: 1.0–37 ppb <i>Animal studies:</i> 100 ppb</p>	
<p>Respiratory Effects and Long-Term Exposure (Section 5.2.2): <u>Suggestive of, but not sufficient to infer, a causal relationship</u> <i>Change in causal determination from the 2008 SO_x ISA (U.S. EPA, 2008d) (inadequate to infer a causal relationship) due to new, but limited, evidence, although uncertainty remains regarding the potential for copollutant confounding.</i></p>	
<p>Key evidence^c (Table 5-24)</p>	<p>Evidence from epidemiologic studies is generally supportive but not entirely consistent for increases in asthma incidence and prevalence related to SO₂ exposure. Uncertainty remains regarding potential copollutant confounding, so chance, confounding, and other biases cannot be ruled out. The limited animal toxicological evidence provides biological plausibility and coherence across lines of evidence. There is some evidence for a mode of action involving inflammation and allergic sensitization.</p>
<p>Overall epidemiologic study means: 2–4 ppb Animal toxicological studies: 2,000 ppb</p>	

Table 1-1 (Continued): Key evidence contributing to causal determinations for sulfur dioxide exposure and health effects evaluated in the current Integrated Science Assessment for Sulfur Oxides.

Health Effect Category, ^a Exposure Duration ^b , and Causal Determination		SO ₂ Concentrations Associated with Effects
Cardiovascular Effects and Short-Term Exposure (Section 5.3.1) <i>Inadequate to infer a causal relationship</i> <i>No change in causal determination from the 2008 SO_x ISA (U.S. EPA, 2008d); new evidence is consistent with previous determination.</i>		
Key evidence ^c (Table 5-31)	There is some evidence of increased hospital admissions and ED visits among adults for IHD, MI, and all CVD; coherence with ST-segment depression in adults with pre-existing coronary heart disease; and increased risk of cardiovascular mortality. However, results are inconsistent across outcomes, and the associations are generally attenuated after copollutant adjustment. Recent studies have not reduced uncertainties identified in the previous ISA, including exposure measurement error and copollutant confounding. There is insufficient evidence from epidemiologic panel studies and experimental studies for clinical cardiovascular effects and to identify key events in a mode of action linking short-term SO ₂ exposure and cardiovascular effects.	Overall epidemiologic study 24-h avg means: 1.2–30 ppb
Cardiovascular Effects and Long-Term Exposure (Section 5.3.2) <i>Inadequate to infer a causal relationship</i> <i>Not included in the 2008 SO_x ISA (U.S. EPA, 2008d).</i>		
Key evidence ^c (Table 5-35)	Results of epidemiologic studies of long-term SO ₂ concentrations and MI, CVD, and stroke events are limited and inconsistent. There is limited coherence with evidence for cardiovascular mortality and weak evidence to identify key events in a mode of action linking long-term SO ₂ exposure and cardiovascular effects. Recent studies have not reduced uncertainties identified in the previous ISA, including exposure measurement error and copollutant confounding.	Overall epidemiologic study means: 1.3–1.7 ppb
Reproductive and Developmental Effects and Exposure (Section 5.4) <i>Inadequate to infer a causal relationship</i> <i>No change in causal determination from the 2008 SO_x ISA (U.S. EPA, 2008d); new evidence is consistent with previous determination.</i>		
Key evidence ^c (Table 5-38)	Consistent positive associations are observed with near-birth exposures to SO ₂ and preterm birth. Although limited evidence is available, positive associations are also reported for fetal growth metrics, birth weight, and infant and fetal mortality. There is insufficient evidence from epidemiologic studies to support an association of SO ₂ exposure with detrimental effects on fertility or pregnancy. Thus, the available studies are of insufficient consistency across outcomes. Recent studies have not reduced uncertainties identified in the previous ISA, including exposure measurement error and copollutant confounding. Limited evidence is available for an understanding of key reproductive and developmental events in mode of action.	Overall epidemiologic study means: 1.9–13 ppb

Table 1-1 (Continued): Key evidence contributing to causal determinations for sulfur dioxide exposure and health effects evaluated in the current Integrated Science Assessment for Sulfur Oxides.

Health Effect Category, ^a Exposure Duration ^b , and Causal Determination		SO ₂ Concentrations Associated with Effects
Total Mortality and Short-Term Exposure (Section 5.5.1) <u>Suggestive of, but not sufficient to infer, a causal relationship</u> <i>No change in causal determination from the 2008 SO_x ISA (U.S. EPA, 2008d); new evidence is consistent with previous determination.</i>		
Key evidence ^c (Table 5-41)	There is consistent epidemiologic evidence from multiple high-quality studies at relevant SO ₂ concentrations demonstrating increases in mortality in multicity studies conducted in the U.S., Canada, Europe, and Asia. There is limited coherence and biological plausibility with cardiovascular and respiratory morbidity evidence and uncertainty regarding a biological mechanism that would explain the continuum of effects leading to SO ₂ -related mortality; thus, chance, confounding, and other biases cannot be ruled out.	Overall epidemiologic study 24-h avg means: <i>U.S., Canada, South America, Europe:</i> 0.4–28 ppb <i>Asia:</i> 0.7–>200 ppb
Total Mortality and Long-Term Exposure (Section 5.5.2) <u>Inadequate to infer a causal relationship</u> <i>No change in causal determination from the 2008 SO_x ISA (U.S. EPA, 2008d); new evidence is consistent with previous determination.</i>		
Key evidence ^c (Table 5-43)	Some epidemiologic studies report positive associations, but results are not entirely consistent, with some studies reporting null associations. Additionally, there is no evidence for associations between SO ₂ exposure and long-term respiratory or cardiovascular health effects to support an association with mortality from these causes. Recent studies have not reduced uncertainties identified in the previous ISA, including exposure measurement error and copollutant confounding.	Overall epidemiologic study means: 1.6–24 ppb

Table 1-1 (Continued): Key evidence contributing to causal determinations for sulfur dioxide exposure and health effects evaluated in the current Integrated Science Assessment for Sulfur Oxides.

Health Effect Category, ^a Exposure Duration ^b , and Causal Determination	SO ₂ Concentrations Associated with Effects
<p>Cancer and Long-Term Exposure (Section 5.6) <u>Inadequate to infer a causal relationship</u> <i>No change in causal determination from the 2008 SO_x ISA (U.S. EPA, 2008d); new evidence is consistent with previous determination.</i></p>	
<p>Key evidence^c (Table 5-44)</p>	<p>Among a small body of evidence, some epidemiologic studies report associations in lung cancer and bladder cancer mortality. There is also some evidence identifying mutagenesis and genotoxicity as key events in a proposed mode of action linking long-term SO₂ exposure and cancer; however, toxicological studies provide limited coherence with epidemiologic studies. Recent studies have not reduced uncertainties identified in the previous ISA, including exposure measurement error and copollutant confounding.</p> <p>Overall epidemiologic study means: 1.5–28 ppb. Toxicological studies: 5,000, 10,700, 21,400, 32,100 ppb</p>

CVD = cardiovascular disease; ED = emergency department; IHD = ischemic heart disease; ISA = Integrated Science Assessment; MI = myocardial infarction; PM = particulate matter; SO₂ = sulfur dioxide; SO_x = sulfur oxides.

^aA large spectrum of outcomes is evaluated as part of a broad health effect category including physiological measures (e.g., airway responsiveness, lung function), clinical outcomes (e.g., respiratory symptoms, hospital admissions), and cause-specific mortality. Total mortality includes all nonaccidental causes of mortality and is informed by the nature of the evidence for the spectrum of morbidity effects (e.g., respiratory, cardiovascular) that can lead to mortality. The sections and tables referenced include a detailed discussion of the available evidence that informed the causal determinations.

^bShort-term exposure refers to time periods of minutes up to 1 mo, while long-term exposures are more than one mo to years.

^cUncertainties remain for many of the studies included as key evidence. Uncertainty remains in some epidemiologic studies. Exposure assessments in epidemiologic studies using fixed-site monitors may not fully capture spatial variability of SO₂. Spatial and temporal heterogeneity may introduce exposure error in long-term effects. For studies of reproductive and developmental outcomes, associations with exposure to SO₂ at particular windows during pregnancy are inconsistent between studies. Additionally, although SO₂ is generally poorly to moderately correlated with other National Ambient Air Quality Standards pollutants at collocated monitors, copollutant confounding by these and other pollutants cannot be ruled out.

1.7 Policy-Relevant Considerations

As described in the Preamble to the ISAs ([U.S. EPA, 2015b](#)) and [Section 1.1](#), this ISA informs policy-relevant issues that are aimed at characterizing relationships between SO₂ exposure and health effects and the impact of these relationships on public health. To that end, this section integrates information from the ISA to describe SO₂ exposure durations and patterns related to health effects, the shape of the concentration-response relationship, regional heterogeneity in relationships, the adverse nature of health effects, and at-risk populations and lifestages. In addressing these policy-relevant issues, this section focuses on respiratory effects associated with short-term exposures, for which the evidence indicates there is a causal relationship.

1.7.1 Durations and Lag Structure of Sulfur Dioxide Exposure Associated with Health Effects

Effects have been observed in controlled human exposure studies after SO₂ exposures during exercise as brief as 5–10 minutes with limited evidence for increased airway responsiveness to subsequent allergen challenge for at least 48 hours following SO₂ exposure in combination with a copollutant (i.e., NO₂). Consistent associations between SO₂ concentrations and asthma hospital admissions and ED visits that are generally unchanged in copollutant models have been demonstrated in epidemiologic studies using daily exposure metrics (24-h avg and 1-h daily max), although the observed effects could be related to very short duration (5–10 minutes) peak exposures experienced during the day.

Regarding the lag in effects, the findings from controlled human exposure studies provide evidence of a rapid onset of effects. The limited number of epidemiologic studies that examined lag structures reported associations within the first few days of exposure.

1.7.2 Concentration-Response Relationships and Thresholds

Characterizing the shape of concentration-response relationships for health effects associated with SO₂ helps quantify the public health impact of SO₂ exposure. A key issue is often whether the relationship is linear across the full range of policy-relevant concentrations or whether there are deviations from linearity, and if so, at what concentrations they occur. Another important issue is whether there is evidence of a

potential threshold, indicating exposures below which adverse health outcomes are not observed. The lack of a discernable threshold in the evidence for the health effects of interest (i.e., respiratory effects associated with short-term exposure) precludes the identification of an exposure level below which there is no risk of effects.

Both controlled human exposure and epidemiologic studies provide some information with respect to the concentration-response relationship between SO₂ exposures and respiratory effects. Results from controlled human exposure studies indicate wide interindividual variability in response to SO₂ exposures, with peak (5–10 minutes) exposures at concentrations as low as 200–300 ppb eliciting lung function decrements in some individuals with asthma. A clear increase in the magnitude of lung function decrements was observed with increasing exposure concentrations between 200 and 1,000 ppb during 5–10 minute SO₂ exposures.

Controlled human exposure studies provide information on the direct relationship between exposure and response over short-duration exposures (i.e., 5–10 minutes) and changes in response with different exposure concentrations. Epidemiologic studies evaluate whether the risk of respiratory effects changes at different ambient concentrations, and these studies are limited to consideration of longer exposure durations (i.e., 1-hr daily max and 24-h avg). The few epidemiologic studies that focus on the SO₂-respiratory effects concentration-response relationship examine pediatric asthma ED visits. The limited epidemiologic evidence to date does not provide evidence for a deviation from linearity or a discernable population-level threshold in the range of ambient concentrations typically observed. However, epidemiologic studies have not been conducted that provide a thorough empirical evaluation of alternatives to linearity. The interpretation of epidemiologic study results is further complicated by potential measurement error due to spatial and temporal variability in SO₂ concentrations.

1.7.3 Regional Heterogeneity in Effect Estimates

The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) discussed spatial variability in SO₂ concentrations and its impact on effect estimates from epidemiologic studies. Correlations between monitors ranged from very low to very high, suggesting that SO₂ concentrations at some monitoring sites may not be highly correlated with the community average concentration. Of particular concern for SO₂ is that most emissions are from point sources, resulting in an uneven distribution of SO₂ concentrations across an urban area. Factors contributing to differences among monitoring sites include source characteristics (e.g., stack height), proximity to sources, terrain features, and uncertainty regarding the measurement of low SO₂ concentrations.

Inability to fully characterize spatial and temporal variability in SO₂ concentrations can contribute to exposure error in epidemiologic studies, whether such studies rely on fixed-site monitor data or concentration modeling for exposure assessment. Studies using 24-h avg concentrations may not capture short-term peak exposures known to produce health effects in controlled human exposure studies. SO₂ has low to moderate spatial correlations between ambient air monitoring sites across urban geographic scales; thus, using fixed-site monitor data for epidemiologic exposure assessment introduces exposure error into the resulting effect estimate. Spatial variability in the magnitude of concentrations may affect cross-sectional and large-scale cohort studies by undermining the assumption that intraurban concentration and exposure differences are less important than interurban differences. This issue may be less important for time-series studies, which rely on day-to-day temporal variability in concentrations to evaluate health effects. Low correlations between monitors contribute to exposure error in time-series studies, including bias toward the null and wider confidence intervals.

1.7.4 Public Health Significance

The public health significance of air pollution-related health effects is informed by the adverse nature of the health effects that are observed, the size of the population exposed to the air pollutant or affected by the health outcome, and the presence of populations or lifestages with higher exposure or increased risk of air pollution-related health effects.

1.7.4.1 Characterizing Adversity of Health Effects

Both the World Health Organization (WHO) and the American Thoracic Society (ATS) have provided guidance in describing what health effects may be considered adverse. WHO defines health as “the state of complete physical, mental, and social well-being and not merely the absence of disease or infirmity” ([WHO, 1948](#)). By this definition, changes in health outcomes that are not severe enough to result in a diagnosis of a clinical effect or condition can be considered adverse if they affect the well-being of an individual. ATS also has considered a wide range of health outcomes in defining adverse effects. Distinguishing between individual and population risk, ATS described its view that small air pollution-related changes in an outcome observed in individuals might be considered adverse on a population level. This is because a shift in the distribution of population responses resulting from higher air pollution exposure might increase the proportion of the population with clinically important effects or at increased risk of a clinically important effect that could be caused by another risk factor ([ATS, 2000b](#)). Increases in ambient SO₂ concentrations are associated with a broad spectrum of health effects related

to asthma, including those characterized as adverse by ATS such as ED visits and hospital admissions.

1.7.4.2 At-Risk Populations and Lifestages for Health Effects Related to Sulfur Dioxide Exposure

The primary NAAQS are intended to protect public health with an adequate margin of safety. In so doing, protection is provided for both the population as a whole and those groups potentially at increased risk for health effects from exposure to the air pollutant for which each NAAQS is set ([Preface](#) to this ISA). Hence, the public health significance of health effects related to SO₂ exposure also is informed by whether specific lifestages or groups in the population are identified as being at increased risk of SO₂-related health effects.

At-risk populations or lifestages can be characterized by specific biological, sociodemographic, or behavioral factors, among others. Since the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), the U.S. EPA has used a framework for drawing conclusions about the role of such factors in modifying risk of health effects of air pollution exposure [Table III of the Preamble to the ISAs ([U.S. EPA, 2015b](#))]. Similar to the causal framework, conclusions about at-risk populations are based on judgments of the consistency and coherence of evidence within and across disciplines ([Chapter 6](#)). Briefly, the evaluation is based on studies that compared exposure or health effect relationships among groups that differ according to a particular factor (e.g., people with and without asthma) and studies conducted in a population or animal model with a particular factor or pathophysiological condition. Where available, information on exposure, dosimetry, and modes of action is evaluated to assess coherence with health effect evidence and inform how a particular factor may contribute to SO₂-related risk of health effects (e.g., by increasing exposure, increasing biological effect for a given dose).

There is adequate evidence that people with asthma are at increased risk for SO₂-related health effects ([Section 6.3.1](#)), which is consistent with the findings of the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). The conclusions are based on findings for short-term SO₂ exposure and respiratory effects (specifically lung function decrements), for which a causal relationship has been determined ([Section 5.2.1.9](#)). There are a limited number of epidemiologic studies evaluating SO₂-related respiratory effects that include stratification by asthma status, but there is evidence for respiratory-related hospital admissions and ED visits ([Section 5.2.1.2](#)). Further support for increased risk in individuals with asthma is provided by biological plausibility drawn from modes of action. Children with asthma may be particularly at increased risk relative to adults with asthma due to their increased responsiveness to methacholine, increased ventilation rates relative to body mass,

increased time spent outdoors, and increased proportion of oral breathing, particularly among boys. Among children in the U.S., asthma is the leading chronic illness (9.5% prevalence) and largest reason for missed school days.

There is also evidence suggestive of increased risk for children and older adults relative to other lifestages ([Section 6.5.1](#)). Although the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) discussed several studies indicating stronger associations between SO₂ and respiratory outcomes for these lifestages, the recent evidence is not entirely consistent with previous studies. For children, studies comparing SO₂-associated respiratory outcomes reported mixed results. For adults, recent evidence generally found similar associations for SO₂-related respiratory outcomes or mortality across age groups, although those over 75 years of age were more consistently at increased risk. In addition, there was insufficient toxicological evidence regarding the effect of lifestage on respiratory responses to SO₂ to support observations made across epidemiologic studies that evaluated lifestage.

1.7.4.3 Summary of Public Health Significance of Health Effects Related to Sulfur Dioxide Exposure

Several aspects of the current evidence are important for considering the public health significance of SO₂-related health effects. One aspect is adversity of the health effects, which may include health effects that are clearly adverse such as ED visits and hospital admissions for asthma and asthma exacerbation. The magnitude of the affected population is also important. As noted above, in the case of SO₂-related health effects, the potentially affected population is large, given the number of people with asthma in the U.S. The roles of co-occurring risk factors or combined higher SO₂ exposure and health risk in influencing the risk of SO₂-related health effects is not well understood. The large proportions of children and older adults in the U.S. population and the high prevalence of asthma in children may translate into a large number of people affected by SO₂, and thus, magnify the public health impact of ambient SO₂ exposure.

1.8 Summary and Health Effects Conclusions

This ISA is a comprehensive evaluation and synthesis of the policy-relevant science regarding the potential health effects of sulfur oxides in the ambient air, focusing on SO₂. The ISA development process involves review of the scientific literature, selecting and evaluating relevant studies, and evaluating the weight of evidence to reach causal determinations regarding the likelihood of independent health effects of SO₂. Information

is included in the ISA on sources of SO₂, atmospheric chemistry of SO₂ and other sulfur-containing compounds, ambient air concentrations of SO₂ nationwide and in urban areas, and modeling approaches for estimating SO₂ concentrations. Approaches for characterizing exposure to ambient SO₂, including monitoring and modeling, together with factors affecting ambient exposure, are described in terms of their potential impact on epidemiologic study results. Dosimetry of SO₂ and potential modes of action are discussed to provide context in considering the potential health effects of SO₂, including respiratory effects, cardiovascular effects, reproductive and developmental effects, cancer, and mortality.

Consistent with the findings of the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), recent studies support the conclusion that there is a causal relationship between short-term SO₂ exposure and respiratory effects. This causal determination is based on consistency of findings within disciplines, coherence among multiple lines of evidence, and biological plausibility indicating that there is a causal relationship between short-term SO₂ exposure and respiratory effects in individuals with asthma. The primary evidence for this conclusion comes from controlled human exposure studies that showed lung function decrements and respiratory symptoms in individuals with asthma exposed to SO₂ for 5–10 minutes under increased ventilation conditions. Supporting evidence was provided by epidemiologic studies that reported positive associations between short-term SO₂ exposures and asthma hospital admissions and ED visits that were generally unchanged in copollutant models involving PM and other criteria pollutants.

For both long-term exposure and respiratory effects, as well as short-term exposure and total mortality, the evidence is suggestive of, but not sufficient to infer, a causal relationship. In both cases, there is some evidence of an association between SO₂ exposure and health outcomes, but the evidence is inconsistent and uncertainties remain, including exposure error and copollutant confounding. The evidence was considered to be inadequate to infer a causal relationship for other health effects, including cardiovascular morbidity (short- and long-term exposure), reproductive and developmental effects, total mortality (long-term exposure), and cancer. For these outcome categories, the evidence generally was not consistent across specific outcomes, showed a potential for copollutant confounding, and was lacking in biological plausibility.

In considering the effects of SO₂ on various populations and lifestages, there is adequate evidence that people with asthma are at increased risk for SO₂-related health effects, as well as suggestive evidence for increased risk among children and older adults. The large proportions of children and older adults in the U.S. population and the high prevalence of

asthma in children may translate into a large number of people affected by SO₂, and thus, magnify the public health impact of ambient SO₂ exposure.

CHAPTER 2 **ATMOSPHERIC CHEMISTRY AND AMBIENT CONCENTRATIONS OF SULFUR DIOXIDE AND OTHER SULFUR OXIDES**

2.1 **Introduction**

Sulfur oxides, in the context of the National Ambient Air Quality Standards (NAAQS), are a group of closely related sulfur-containing gas-phase compounds [e.g., sulfur dioxide (SO₂), sulfur monoxide (SO), disulfur monoxide (S₂O), and sulfur trioxide (SO₃)]. Sulfur oxides also appear in the particle phase, as components of particulate matter (PM), and particle phase sulfur compounds are discussed separately as part of the Integrated Science Assessment for PM ([U.S. EPA, 2016c](#)).

The NAAQS for SO_x are currently set using SO₂ as the indicator species. Of the sulfur oxides, SO₂ is the most abundant in the atmosphere, the most important in atmospheric chemistry, and the one most clearly linked to human health effects ([U.S. EPA, 2008d](#), [1996b](#); [HEW, 1969](#)). Therefore, the emphasis in this chapter is on SO₂. Note that the mechanism of particle-phase SO₄²⁻ formation is briefly described in [Section 2.3](#) [for more detail, see [Seinfeld and Pandis \(2006\)](#), [Finlayson-Pitts and Pitts \(2000\)](#), and other atmospheric chemistry texts]. The health effects of sulfate and other sulfur compounds in the particle phase are discussed in the ISA for Particulate Matter ([U.S. EPA, 2009a](#)).

Sulfur dioxide is both a primary gas-phase pollutant (when formed during fuel combustion) and a secondary pollutant [the product of atmospheric gas- or aqueous-phase oxidation of reduced sulfur compounds (sulfides)]. Fossil fuel combustion is the main anthropogenic source of primary SO₂, while volcanoes and landscape fires (wildfires as well as prescribed burns) are the main natural sources of primary SO₂. Industrial chemical and pulp and paper production, smelter and steel mill operations, natural biological activity (plants, fungi, and prokaryotes), and volcanoes are among many sources of reduced sulfur compounds that ultimately lead, through various oxidation reactions in the atmosphere, to the formation of secondary SO₂.

This chapter describes the main sources of sulfur dioxide and other gas-phase sulfur oxides found in the atmosphere, as well as reduced sulfur gases that serve as precursors for SO₂ ([Section 2.1](#)), source emissions ([Section 2.2](#)), atmospheric chemistry and fate ([Section 2.3](#)), measurement methods ([Section 2.4](#)), environmental concentrations ([Section 2.5](#)), and atmospheric modeling of sulfur dioxide concentrations ([Section 2.6](#)). This material is provided as a prologue for detailed discussions on exposure and health

effects evidence in the subsequent chapters, and as a source of information to help interpret that evidence in the context of relevant ambient concentrations.

2.2 Anthropogenic and Natural Sources of Sulfur Dioxide

This section briefly describes the main U.S. anthropogenic and natural sources of SO₂ emissions. Emissions estimates for natural and anthropogenic sulfide emissions for the U.S. alone are not available in the literature. Therefore, a brief discussion of the sulfur cycle and estimates of the contribution of sulfides at the global scale, all of which can be found in the literature, are provided. [Section 2.2.1](#) describes the main categories of anthropogenic SO₂ emissions, while [Section 2.2.2](#) presents the geographic distribution of SO₂ sources across the U.S. The declining trend in anthropogenic SO₂ emissions is discussed in [Section 2.2.3](#). Natural sources of SO₂ are discussed in [Section 2.2.4](#). Indirect production of SO₂ through oxidation of reduced sulfur compounds emitted from geologic and biological sources is discussed in [Section 2.2.5](#).

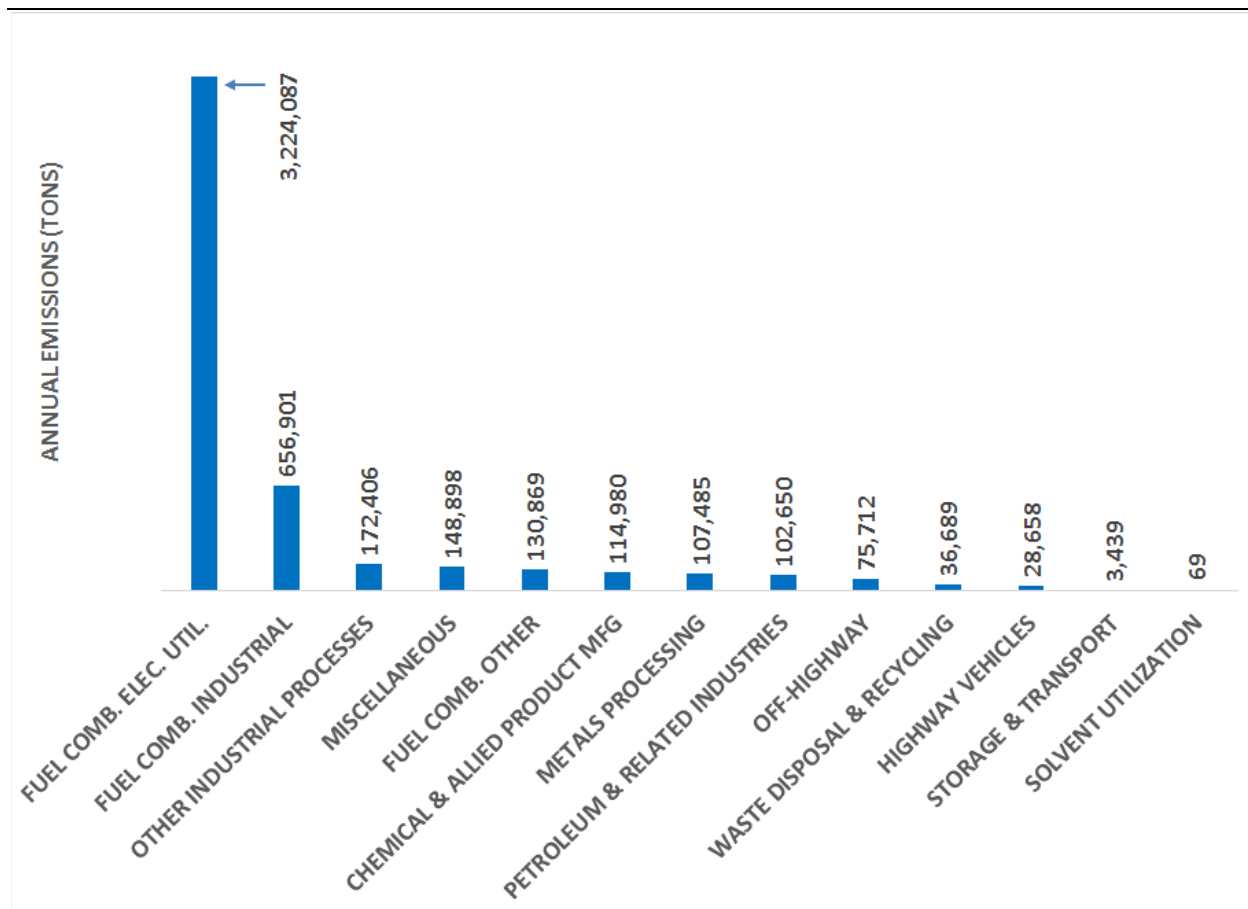
Sulfur is present to some degree in all fossil fuels, especially coal, and occurs as reduced organosulfur compounds. Coal also contains sulfur in mineral form (pyrite or other metallo-sulfur minerals) and in elemental form ([Calkins, 1994](#)). Of the most common types of coal (anthracite, bituminous, subbituminous, and lignite), sulfur content varies between 0.4 and 4% by mass. Sulfur in fossil fuels is almost entirely converted to SO₂ (or SO₃) during combustion, making accurate estimates of SO₂ combustion emissions possible based on fuel composition and combustion rates.

The mass of sulfur released into the environment by anthropogenic sources is comparable to natural sources ([Brimblecombe, 2003](#)). In addition to volcanic and other geologic SO₂ emissions, naturally occurring SO₂ is derived from the oxidation of sulfides emitted by low flux “area” sources, such as the oceans and moist soils. Anthropogenic emissions of sulfur are primarily in the form of SO₂, emerging from point sources in quantities that substantially affect local and regional air quality.

2.2.1 U.S. Anthropogenic Sources

The largest SO₂-emitting sector within the U.S. is electricity generation based on coal combustion (3,224,087 tons), although emissions from this sector have been declining in recent years due to fuel substitution and emissions controls. The mass of emissions produced by the fuel combustion in electrical utilities sector [i.e., coal-fired electric generating units (EGUs)] exceeds those produced by the next largest sector [the fuel

combustion—industrial sector (i.e., coal-fired boilers)] by nearly a factor of 5, and EGUs emit approximately twice as much SO₂ as all other sources combined. [Figure 2-1](#) provides a sector comparison of annual emissions [in tons] found in the U.S. EPA 2014 National Emissions Inventory (NEI) ([U.S. EPA, 2016b](#)).



COMB = combustion; ELEC = electric; MFG = manufacturing; UTIL = utilities.

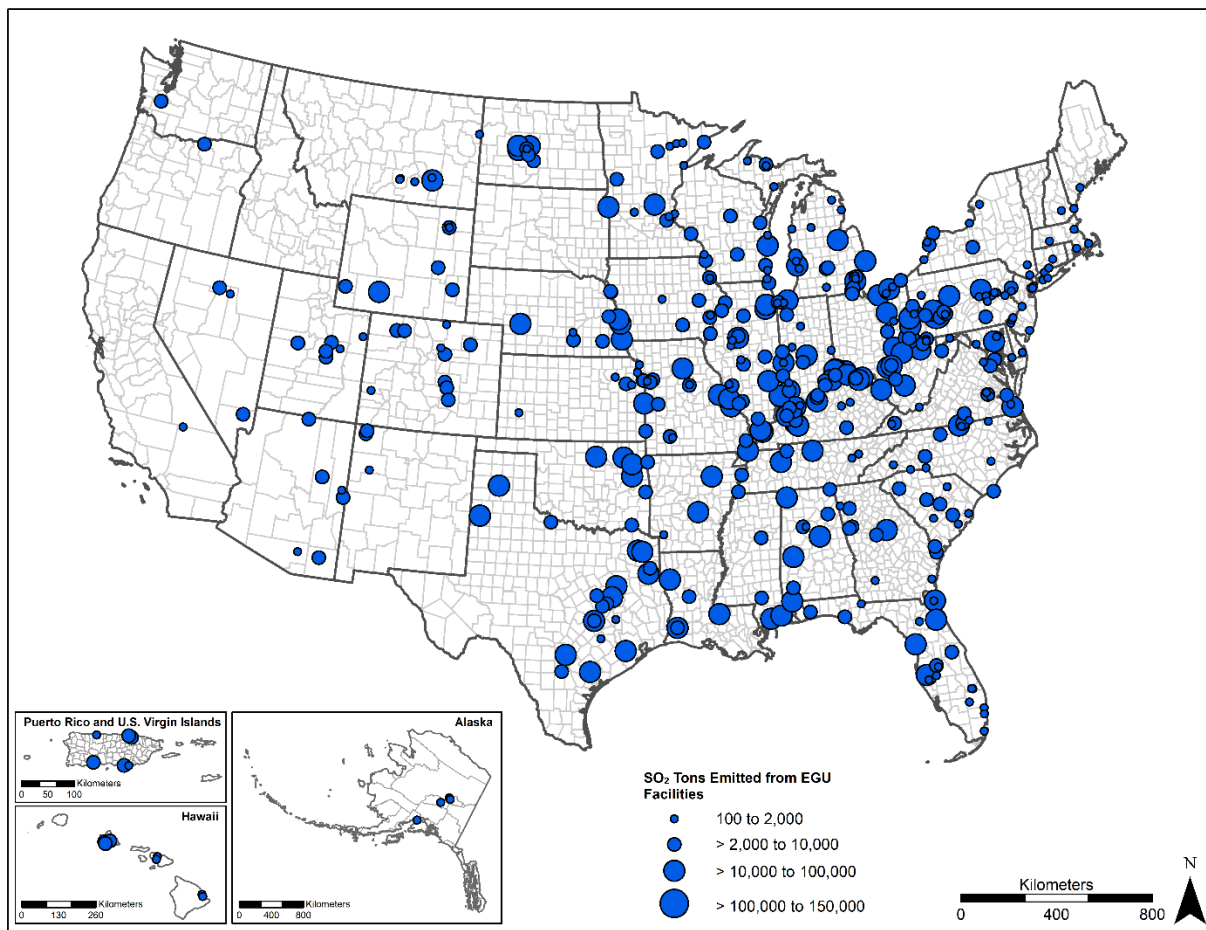
Note: “Fuel Comb. —Other” includes commercial, institutional, and residential sources. Metals Processing” includes copper smelting (22,792 tpy SO₂) and combined iron and steel mill (28,247 tpy SO₂) facilities.

Source: <https://www.epa.gov/air-emissions-inventories/air-pollutant-emissions-trends-data> (U.S. EPA, 2016b).

Figure 2-1 Sulfur dioxide emissions by sector in tons, 2014.

Because EGUs comprise the largest NEI source category, the spatial distribution of SO₂-emitting EGUs is presented here ([U.S. EPA, 2016b](#)). Most EGU sources are located in the eastern half of the continental U.S., as indicated in [Figure 2-2](#). There is a particularly high concentration of EGUs in the Ohio River valley, upper Midwest, and

along the Atlantic coast. Many of the monitoring sites with elevated SO₂ concentrations are located in these same areas ([Section 2.5.2](#)).



Note: EGU = electric power generating unit; SO₂ = sulfur dioxide.
Source: <https://www.epa.gov/air-emissions-inventories>; (U.S. EPA, 2016b).

Figure 2-2 Distribution of electric power generating units (438), scaled according to annual sulfur dioxide emissions across the U.S. based on the 2014 National Emissions Inventory.

Industrial fuel combustion is the second largest source nationwide, emitting 656,901 tpy, followed by other fuel combustion (172,406 tpy). Miscellaneous (148,898 tpy) is the fourth-largest source and includes SO₂ emissions by fire used in landscape management and agriculture as well as wildfires (U.S. EPA, 2016b). Wildfires, as a natural source of SO₂ emissions, are discussed in [Section 2.2.4.3](#). The metals processing sector includes

copper smelting (22,792 tpy SO₂) and combined iron and steel mill (28,247 tpy SO₂) facilities. Highway vehicles emit <1% of the combined emissions shown in [Figure 2-1](#). The commercial marine sector falls within the off-highway category (75,712 tpy) ([U.S. EPA, 2016b](#)). Using data from 2002, [Wang et al. \(2007\)](#) modeled SO₂ emissions from commercial marine activity by combining historical shipping data and marine traffic predictions based on port sizes and probable routes. Within a 200-nautical-mile boundary around the marine, lake, and river international borders of the U.S., the study authors estimated that 38% of emissions related to commercial marine shipping occurred along the East Coast of the U.S. Twenty percent of emissions were estimated for the West Coast and 26% for the Gulf Coast. Smaller quantities were estimated elsewhere (10% for Alaska, 3% for Hawaii, and 2% for the Great Lakes). Interior waterway activity was not included in the [Wang et al. \(2007\)](#) study. In 2010, the International Maritime Organization introduced Emissions Control Areas (ECA) around U.S., Canadian, and French waters under the International Convention for the Prevention of Pollution from Ships ([Office of Transportation and Air Quality, 2010](#)). The ECA is a 200-nautical-mile buffer around the maritime borders, within which fuels cannot contain more than 1,000 ppm sulfur as of 2015. These reductions are expected to be accomplished by having maritime vessels switch fuel sources when crossing the 200-nautical-mile buffer to approach their port. The U.S. EPA's [Office of Transportation and Air Quality \(2010\)](#) estimated that this reduction in the amount of sulfur in marine fuels used within the 200 nautical mile buffer would result in an 85% reduction in SO₂ emissions from the commercial marine sector. [Monitors located at the Port of Los Angeles and the Port of Long Beach reflect these reductions, with latest reports from these two ports showing SO₂ concentrations well below the NAAQS ([Leidos Inc, 2016](#))].

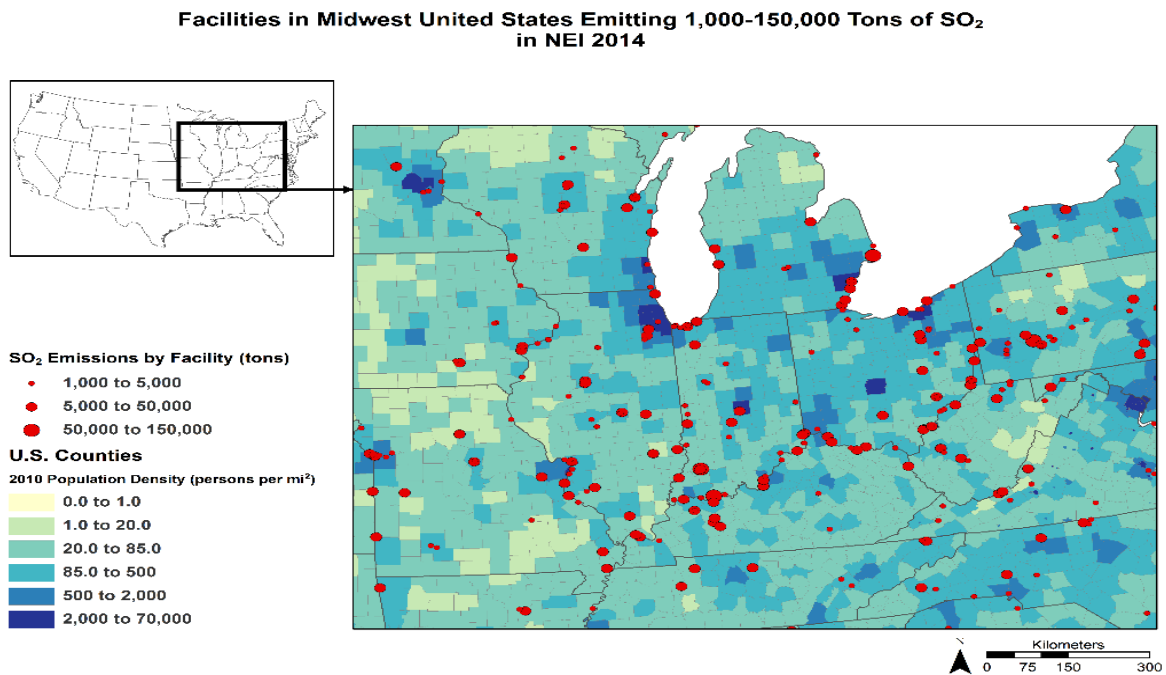
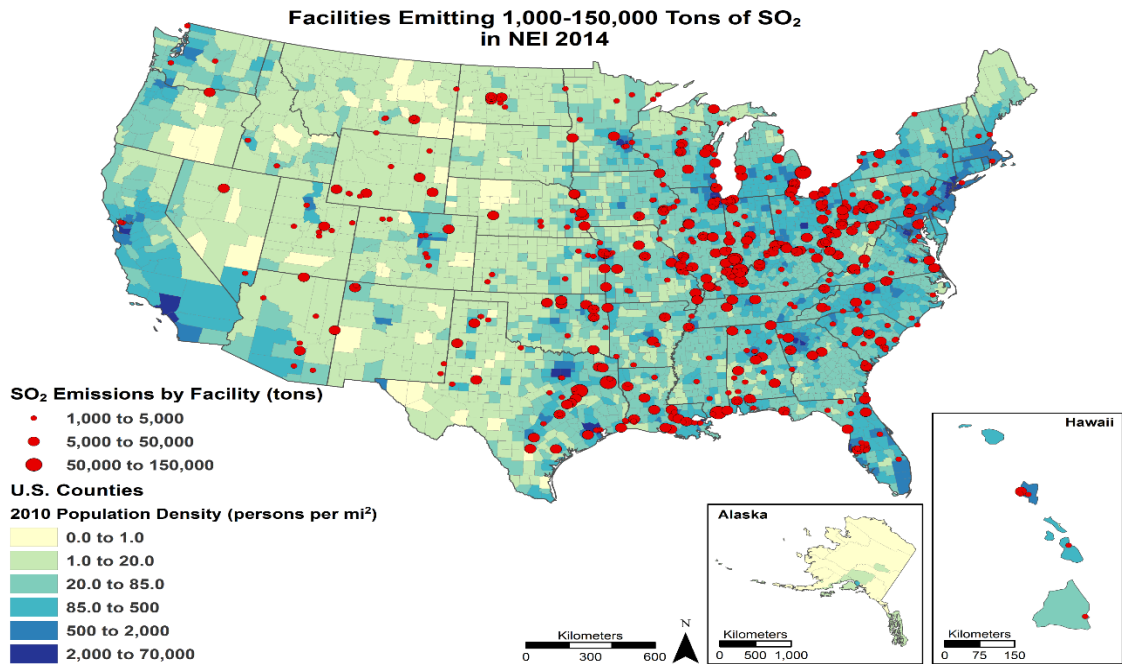
National SO₂ emissions sector summaries cannot offer insight concerning the local influence of individual SO₂-emitting facilities. Although they may be fewer in number than fossil fuel-fired EGUs, other types of large emissions facilities that may substantially impact local air quality include copper smelters, kraft pulp mills, Portland Cement plants, iron and steel mill plants, sulfuric acid plants, petroleum refineries, and chemical processing plants. For example, the metals processing sector represents less than 2.3% of total emissions from the 2014 NEI ([U.S. EPA, 2016b](#)), but monitoring sites that have recorded some of the highest 1-hour daily max SO₂ concentrations in the U.S. are located near copper smelters in Arizona ([Sections 2.5.2 and 2.5.4; Figure 2-11](#)).

2.2.2 National Geographic Distribution of Large Sources

Some industrial facilities are complex sources, with SO₂ emissions related to multiple processes. [Figure 2-3](#) shows the geographic distribution of continental U.S. facilities emitting more than 1,000 tpy SO₂, with an enlargement of the midwestern states including the Ohio River Valley, where a large number of these SO₂-emitting sources are located.

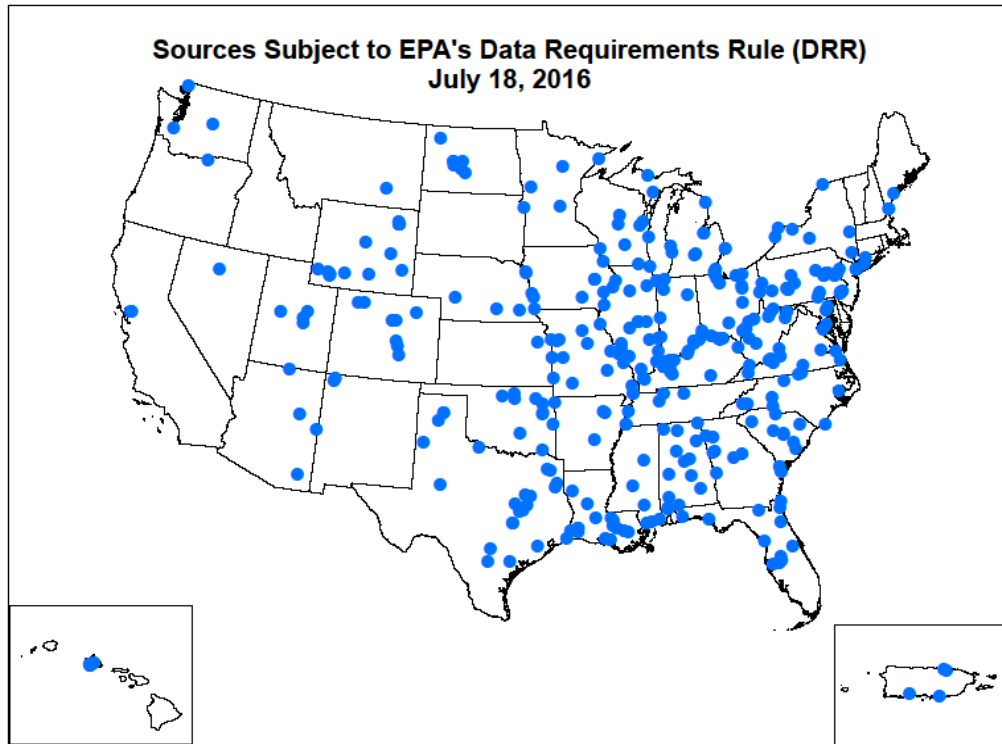
U.S. EPA Sulfur Dioxide Data Requirements Rule

Another information resource regarding air quality near large sources of SO₂ is the data produced by air quality monitors required by the SO₂ Data Requirements Rule (40 CFR 51.1202–51.1203; 80 FR50152, August 21, 2015), which was enacted in support of the SO₂ NAAQS. This regulation requires that, at a minimum, air agencies must characterize air quality around sources that emit 2,000 tons per year or more of SO₂. An air agency may avoid the requirement for air quality characterization near a source by adopting enforceable emission limits that ensure that the source will not emit more than 2,000 tpy. This final rule gives air agencies the flexibility to characterize air quality using either modeling of actual source emissions or using appropriately sited ambient air quality monitors. Under this data requirements rule, in 2016, air agencies submitted to their relevant U.S. EPA Regional Administrator a final list identifying the sources in the state around which SO₂ air quality is to be characterized. The list included sources with emissions above 2,000 tpy SO₂. The U.S. EPA Regional Offices or air agencies included additional sources on this list that they deemed necessary. The national list contained 377 sources (<https://www.epa.gov/so2-pollution/so2-data-requirements-rule-source-list>). [Figure 2-4](#) shows the locations of those sources.



Source: <https://www.epa.gov/air-emissions-inventories>; U.S. EPA (2016b).

Figure 2-3 Geographic distribution of (top) continental U.S. facilities (526) emitting more than 1,000 tpy SO₂, with (bottom) an enlargement of the midwestern states, including the Ohio River Valley, where a large number of these sources are concentrated.



DRR = Data Requirements Rule; U.S. EPA = U.S. Environmental Protection Agency.
Source: U.S. EPA Office of Air Quality Planning and Standards.

Figure 2-4 Sulfur dioxide sources (377) identified by state/local air agencies under the U.S. Environmental Protection Agency's Data Requirements Rule, as of July 18, 2016.

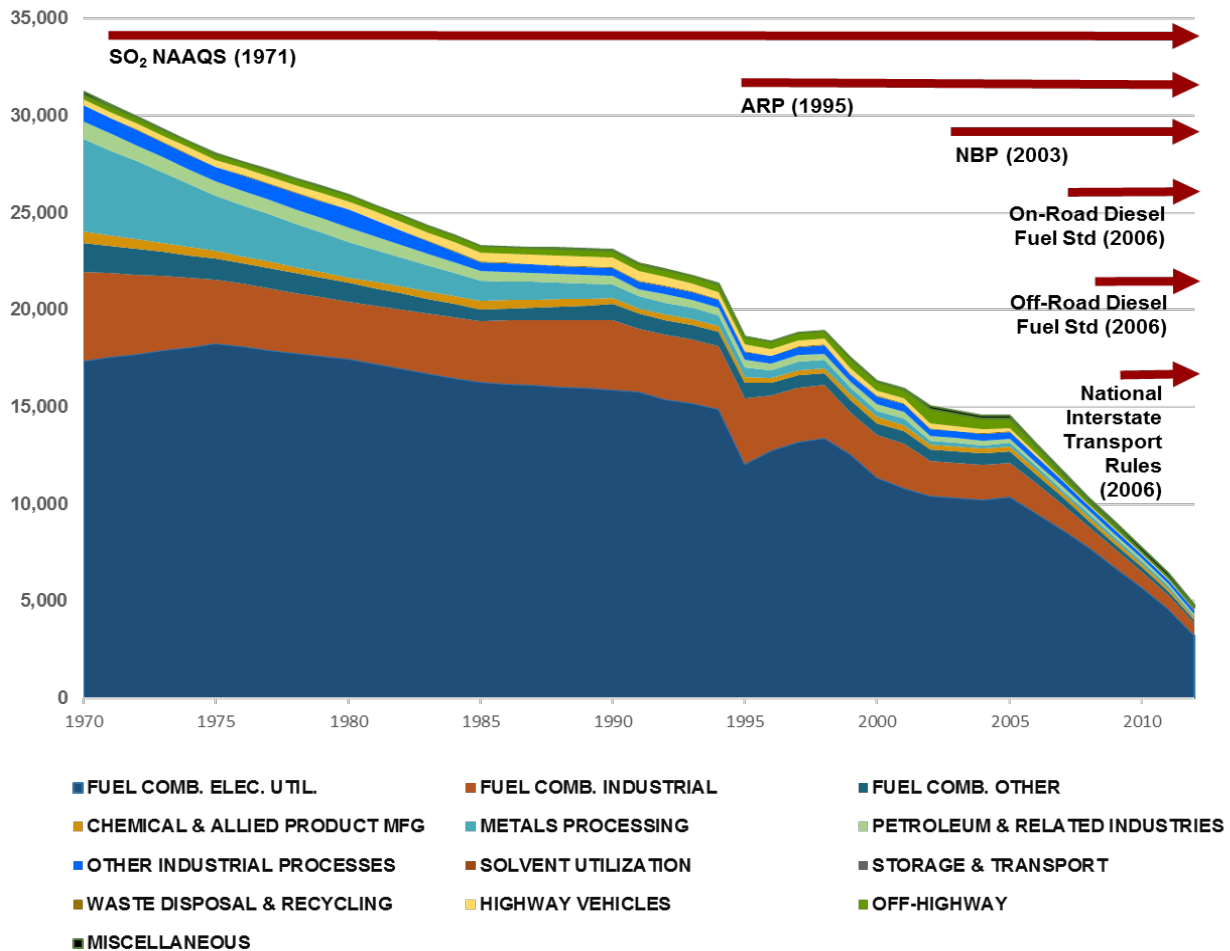
2.2.3 U.S. Anthropogenic Emission Trends

Anthropogenic emissions of SO₂ in the U.S. have shown dramatic declines since the 1970s, and emissions reductions have accelerated since the 1990 amendments to the Clean Air Act were enacted (USC Title 42 Chapter 85). [Table 2-1](#) gives the annual SO₂ emissions, percentage of the U.S. SO₂ total emissions, and change in emissions rate from 2004 to 2014. [Figure 2-5](#) illustrates the emissions trends by sector from 1970 to 2014 in relation to the timeline over which the NAAQS for SO_x and the Clean Air Act control programs [Acid Rain Program (ARP), NO_x Budget Program (NBP), on- and off-road diesel emissions standards, and Cross-State Air Pollution Rule (CSAPR) and other national interstate transport rules] have been implemented ([U.S. EPA, 2016b](#)). Exceptions to the steep decline in SO₂ emissions in the listed sectors are the increases in emissions from the waste disposal and recycling sectors, commercial storage and transport sector, and from miscellaneous sources (e.g., landscape fires). Waste disposal and recycling contributes only 0.8% of total 2014 SO₂ emissions. Landscape fires are a larger contributor to the NEI (3%) and are discussed further in [Section 2.2.4.3](#).

[Hand et al. \(2012\)](#) studied reductions in EGU-related annual SO₂ emissions during the 2001–2010 period. They found that emissions decreased throughout the U.S. by 6.2% per year, with the largest reductions in the western U.S. at 20.1% per year. The smallest reduction (1.3% per year) occurred in the Great Plains states.

Table 2-1 Summary of current U.S. Environmental Protection Agency sulfur dioxide trends data by emissions sector. Values shown in bold indicate increased emissions.

Source Type	Tons SO ₂ Emitted, 2014	Percentage of Total*	Percent Change, 2004-2014
Fuel combustion, electric utilities	3,224,087	67%	-69%
Fuel combustion, industrial	656,901	14%	-63%
Other industrial processes	172,406	3.6%	-51%
Miscellaneous	148,898	3.1%	2.6%
Fuel comb. other	130,869	2.7%	-77%
Chemical & allied product mfg	114,980	2.4%	-55%
Metals processing	107,485	2.2%	-43%
Petroleum & related industries	102,650	2.1%	-53%
Off-highway	75,712	1.6%	-87%
Waste disposal & recycling	36,689	0.8%	32%
Highway vehicles	28,658	0.6%	-86%
Storage & transport	3,439	0.1%	2.6%



ARP = Acid Rain Program; COMB = combustion; ELEC = electric; MFG = manufacturing; NAAQS = National Ambient Air Quality Standards; NBP = NO_x Budget Program; SO₂ = sulfur dioxide; STD = Standards; UTIL = utilities.

Source: <https://www.epa.gov/air-emissions-inventories/air-pollutant-emissions-trends-data>; U.S. EPA (2016b).

Figure 2-5 National sulfur dioxide emissions trends by sector (10³ tpy), 1970–2014.

2.2.4 Natural Sources

This section provides an overview of the major natural sources of SO₂ and reduced sulfur compounds that are oxidized in the atmosphere to form SO₂. [Section 2.2.4.1](#) briefly

describes the elements of the global sulfur cycle. [Section 2.2.4.2](#) briefly discusses volcanic sources of SO₂ within the U.S. [Section 2.2.4.3](#) discusses SO₂ emissions by U.S. wildfires. The section concludes with a brief summary of both anthropogenic and natural emissions of reduced sulfur gases that can serve as precursors to SO₂.

2.2.4.1 The Global Sulfur Cycle

The total budget for sulfur, in all its forms, at Earth's surface is on the order of 1.1×10^{16} tons S ([Schlesinger, 1997](#)). The sulfur cycle comprises the many chemical and biological processes that continuously interconvert the element among its four main oxidation states (-2, 0, +4, +6). The reduced form of sulfur is present in the environment as hydrogen sulfide, hydrogen disulfide, and a number of organic compounds. Oxidized sulfur is present primarily as SO₂ and sulfate (SO₄²⁻).

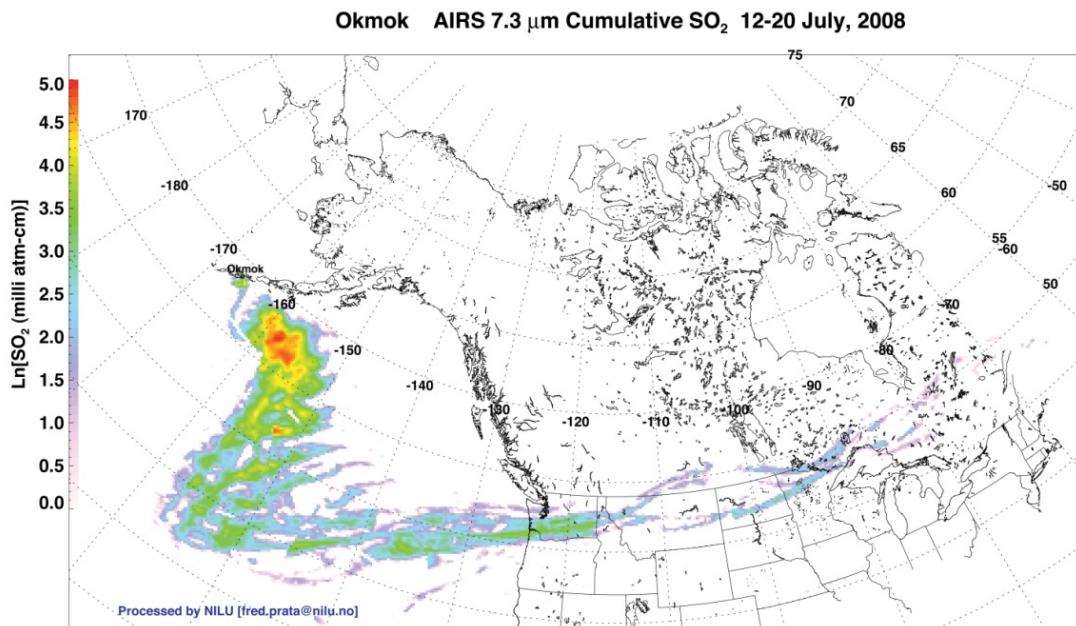
Volcanoes and wildfires are nonbiological natural sources that directly emit SO₂ to the atmosphere. Biological natural sources, together with volcanoes, emit reduced sulfur compounds that subsequently oxidize in the atmosphere to form SO₂. Under anaerobic conditions, various species of plants, fungi, and prokaryotes convert oxidized sulfur into its reduced forms ([Madigan et al., 2006](#)). Photosynthetic green and purple bacteria and some chemolithotrophs oxidize sulfides to form elemental sulfur. Some microorganisms oxidize elemental sulfur to form SO₄²⁻ and SO₂; others reduce elemental sulfur to sulfides (*dissimilative sulfur reduction*), while others are capable of reducing SO₄²⁻ all the way down to sulfide (*dissimilative SO₄²⁻ reduction*).

2.2.4.2 Volcanoes as a Natural Source of Sulfur Dioxide

Geologic activity, including fumaroles, geysers, and metamorphic degassing, emits a number of gases, including SO₂, carbon dioxide (CO₂), hydrogen sulfide (H₂S), hydrochloric acid, and chlorine ([Simpson et al., 1999](#)). Eruptive and noneruptive volcanoes are the most important sources of geologic SO₂ emissions. Noneruptive, but geothermally active, volcanoes outgas at relatively constant rates and appear to be more significant sources of SO₂ than burst emissions that occur during eruptions. The emissions from volcanic eruptions are sporadic, and therefore, vary from year to year ([Simpson et al., 1999](#)).

The western U.S. border of the continental United States (CONUS) is near and in some cases over the boundary between the North American, the Pacific, and the Juan de Fuca tectonic plates. The region is subject to ongoing volcanic activity. In Alaska, the Aleutian

volcanic arc, part of the state of Alaska, comprises 75 volcanic centers. Volcanoes in this chain have erupted once or twice per year on average over the past 100 years with impacts on local communities (Power, 2013). Figure 2-6 shows an image derived from data collected by the Atmospheric Infrared Sounder (AIRS) instrument aboard NASA's Aqua satellite during the July 12–20, 2008 eruption of the Okmok Volcano in Alaska's Aleutian Islands. The image shows sulfur dioxide at altitudes around 16 km (10 miles) released by the volcano over that time span, with red indicating the highest concentrations, and pale pink indicating the lowest (Prata et al., 2010). Sulfur dioxide has infrared absorption features at 4 and 7.3 μm , which allowed Prata et al. (2010) to calculate the total mass of SO_2 emitted during the eruption as $319,670 \pm 11,023$ tons.



AIRS = Atmospheric Infrared Sounder; SO_2 = sulfur dioxide.

Source: Image courtesy of Fred Prata of the Norwegian Institute for Air Research (NILU); [NASA \(2008a\)](#).

Figure 2-6 Sulfur dioxide released during the July 12–20, 2008 eruption of the Okmok Volcano in Alaska's Aleutian Islands (image derived from data collected by the Atmospheric Infrared Sounder instrument aboard the National Aeronautics and Space Administration Aqua satellite).

The line of volcanoes long the western side of North American extends from the Aleutian Islands in Alaska south and east through the states of Washington, Oregon, California, Arizona, and New Mexico, with outlying geologically active sites in Idaho (Craters of the

Moon) and Wyoming (Yellowstone). [Figure 2-7](#) shows the geographic location and activity potential for these sites within the continental U.S.

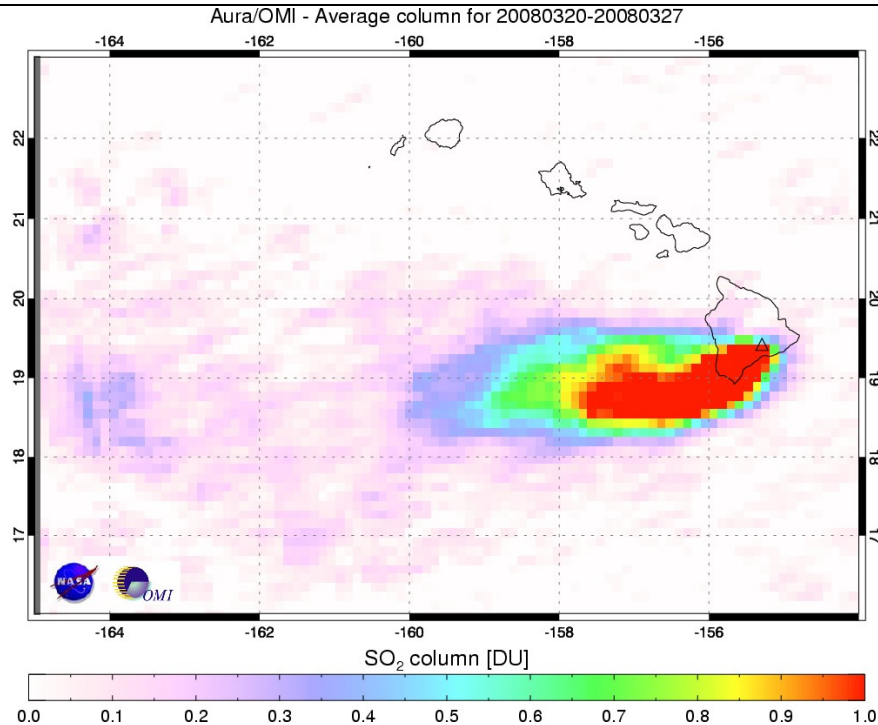


Source: [USGS \(1999\)](#). Map courtesy of Lyn Topinka (1999, USGS/CVO), Modified from Steve Brantley (USGS 1994), Volcanoes of the United States, USGS General Interest Publication.

Figure 2-7 Geographic location of volcanoes and other potentially active volcanic areas within the continental U.S.

The state of Hawaii, located over a “hot spot” in the north-central portion of the Pacific tectonic plate, is a series of volcanic islands with one of the world’s most active volcanoes, Kīlauea, located on the Big Island of Hawaii. Kīlauea typically emits SO₂ at a steady rate. In mid-March of 2008, the volcano exhibited a small, explosive eruption, followed by a two- to fourfold increase in SO₂ emissions. The Ozone Monitoring Instrument (OMI) aboard the NASA Aura satellite detected this increase in SO₂ emissions. [Figure 2-8](#) shows the average concentration of SO₂ in the evolving plume for

the March 20–27, 2008 period. Persistent easterly trade winds moved the plume westward, away from populated areas.



DU = Dobson units; OMI = Ozone Monitoring Instrument; SO₂ = sulfur dioxide.

Note: A DU is approximately equivalent to a total column concentration of 1 ppbv of SO₂. Horizontal axis is longitude with respect to Greenwich, U.K. Vertical axis is latitude with respect to the equator.

Source: [NASA \(2008b\)](#).

Figure 2-8 National Aeronautics and Space Administration/Ozone Monitoring Instrument image of the Kilauea sulfur dioxide plume during its March 20–27, 2008 eruption.

In another study using SO₂ column densities derived from the Global Ozone Monitoring Experiment-2 satellite measurements for the period 2007–2012, [Beirle et al. \(2013\)](#) determined Kilauea’s monthly mean SO₂ emission rate and effective SO₂ lifetime. For the March through November, 2008 period, the authors reported Kilauea’s SO₂ emission rate as 8,818–20,943 tons/day and the effective SO₂ lifetime as 1–2 days. Several studies have estimated the global SO₂ emissions of sulfur by volcanoes to be in the range of 7.7×10^6 – 2.0×10^7 tpy ([Chin et al., 2000](#); [Feichter et al., 1996](#); [Pham et al., 1996](#); [Langner and Rodhe, 1991](#)).

2.2.4.3 Wildfires as a Natural Source of Sulfur Dioxide

Sulfur is a component of amino acids in vegetation and is released during combustion, mainly in the form of SO₂. [Wiedinmyer et al. \(2006\)](#) determined fire location and timing, fuel loadings, and emission factors by using the literature and satellite data from various sources, including the Moderate Resolution Imaging Spectroradiometer (MODIS) Thermal Anomalies Product, the Global Land Cover Characteristics 2000 data set, and the MODIS Vegetation Continuous Fields Product. The study authors estimated SO₂ emissions from fires for the U.S. as 176,370 tons in the year 2004. Canadian fires emitted 121,254 tons, and Mexican fires emitted 55,116 tons of SO₂ for the same period. However, wildfire emissions do vary from year to year. Emissions estimates for SO₂ derived from global modeling studies of wildfire range between 5.1×10^6 and 6.3×10^6 tpy SO₂ ([Chin et al., 2000](#); [Feichter et al., 1996](#); [Pham et al., 1996](#); [Langner and Rodhe, 1991](#)). For comparison, the 2014 NEI also includes an estimate for U.S. agricultural and prescribed burning emissions at 75,643 tpy, which is comparable to the estimated SO₂ emissions from wildfires at 71,448 tpy ([U.S. EPA, 2016b](#)).

Projected increases in wildfire frequency and intensity under warming climate conditions imply increasing wildfire-related SO₂ emissions. However, these estimates are highly uncertain due to the lack of data on the sensitivity of emissions composition with respect to the effects of climate change on landscape species composition and burning conditions.

2.2.5 Reduced Sulfur Compounds as Indirect Sources of Sulfur Dioxide

Sulfides, including H₂S, carbonyl sulfide (OCS), carbon disulfide (CS₂), methylmercaptan (CH₃SH), dimethyl sulfide (DMS), and dimethyl disulfide (DMDS), are emitted from energy production, industrial activities, agriculture, and various ecosystems, especially coastal wetland systems, inland soils, and oceans. In addition to SO₂, volcanoes release sulfides, specifically H₂S, OCS, and CS₂. As described in [Section 2.3](#), all of these gases, except OCS, have short atmospheric lifetimes, given their high rates of reaction with hydroxyl and nitrate (NO₃) radicals, with SO₂ as a reaction product. [Table 2-2](#) provides a list of the natural and anthropogenic sources of the five main organosulfides. Dimethyl sulfide is particularly important for the large role it plays as a source of atmospheric sulfur.

Table 2-2 Largest global sulfide emissions sources, ranked according to total sulfur emissions (10³ Tpy).

Anthropogenic Sources	OCS	CS₂	CH₃SH	DMS	DMDS
Pulp and paper industry	107	86.5	1,852	1,612	301
Manure	NR	NR	364	728	728
Rayon/cellulosics manufacture	NR	435	56.7	39.2	NR
Pigment industry	81.6	226	NR	NR	NR
Biomass burning	47.4	1.98	NR	6.61	131
Oxidation	179	NR	NR	NR	NR
Wastewater	0.024	1.14	56.2	3.75	23.6
Biofuel combustion	80.6	3.20	NR	NR	NR
Coal combustion	66.1	0.364	NR	NR	NR
Paddy fields	0.419	29.7	0.838	27.6	0.628
Aluminum industry	33.1	4.41	NR	NR	NR
Food processing	0.694	NR	NR	4.38	31.9
Shipping	33.1	1.65	NR	NR	NR
Tire wear	12.9	17.1	NR	NR	NR
Coke production	9.93	15.4	NR	NR	N
Gas industry	0.772	NR	5.29	0.926	0.110
Vehicles	6.61	0.331	NR	NR	NR
Landfill and waste	0.087	0.209	0.375	0.287	0.009
Tire combustion	0.071	NR	0.047	NR	NR
Brickmaking	NR	0.033	NR	NR	NR
Total anthropogenic sulfur emissions, by compound	659	823	2,336	2,422	1,216
Natural Sources	OCS	CS₂	CH₃SH	DMS	DMDS
Saline and ocean water	978	268	5,223	31,071	235
Vegetation and soils	NR	77	1,913	3,825	957
Oxidation	174	NR	NR	NR	NR
Volcanoes	12	19	NR	NR	NR
Total natural sulfur emissions, by compound	1,164	364	7,135	34,896	1,190

CH₃SH = methylmercaptan; CS₂ = carbon disulfide; DMDS = dimethyl disulfide; DMS = dimethylsulfide; NR = not reported; OCS = carbonyl sulfide.

Adapted from [\(Lee and Brimblecombe, 2016\)](#).

Dimethyl sulfide (DMS) is the most abundant reduced sulfur gas. It has appreciable anthropogenic sources (pulp and paper production, agricultural operations), but these sources emit much less than the quantity emitted by natural biological activity. Natural emissions of dimethyl sulfide are due to the breakdown of dimethyl sulfoniopropionate, a metabolite of the amino acid methionine, which is produced by marine organisms, particularly in areas of oceanic upwelling and coastal zones, and by anaerobic bacteria in marshes and estuaries. The oxidation of dimethyl sulfide contributes to low-level background SO₂ concentrations in coastal environments. [Lee and Brimblecombe \(2016\)](#) provide a literature-derived global estimate of DMS emissions from seawater and marshland of 3.1×10^7 tpy S. Earlier estimates for seawater DMS emissions range widely from 6.1×10^6 to 2.4×10^7 tpy ([Liu et al., 2005](#); [Chin et al., 2000](#); [Feichter et al., 1996](#); [Pham et al., 1996](#); [Langner and Rodhe, 1991](#)). A warming climate may have a complex feedback effect on DMS emissions, influencing both ocean surface temperatures and currents controlling nutrient dispersion and affecting the population and location of DMS producing phytoplankton ([Kloster et al., 2007](#)).

2.3 Atmospheric Chemistry and Fate of Sulfur Dioxide and Other Sulfur Oxides

The important gas-phase sulfur oxides in the troposphere are SO₂ and H₂SO₄ ([U.S. EPA, 2008d](#)). SO₃ is known to be present in the emissions of coal-fired power plants, factories, and refineries, but it reacts with water vapor in the stacks or immediately after release into the atmosphere within seconds to form H₂SO₄. This short atmospheric residence time makes SO₃ difficult to detect in the ambient atmosphere. Gas-phase H₂SO₄, the product of both SO₂ and SO₃ oxidation, quickly condenses onto existing atmospheric particles or participates in new particle formation ([Finlayson-Pitts and Pitts, 2000](#)). Of these species, only SO₂ is present at concentrations in the gas phase that are relevant for chemistry in the atmospheric boundary layer and troposphere, and for human exposures. Other sulfur oxides, including both S(IV) and S(VI) compounds, appear in the atmosphere due to direct emissions and as the products of the oxidation of more reduced forms of sulfur. Gas-phase precursors to SO₂ include the sulfides ([Section 2.2.5](#)) and partially oxidized sulfur-containing organic compounds. Early research on industrial air pollution reported observations of partially oxidized sulfur compounds in industrial emissions plumes and the ambient atmosphere as both gases and particulate matter. These compounds included bis-hydroxymethyl sulfone ([Eatough and Hansen, 1984](#)), dimethyl- and monomethyl-sulfate ([Eatough et al., 1986](#); [Eatough et al., 1981](#)), sulfonic acids ([Panter and Penzhorn, 1980](#)), and particle-bound iron sulfite complexes ([Eatough et al., 1978](#)). No other more recent studies are available that quantify these intermediate sulfur oxides.

However, energy generation, industrial coal combustion, and refinement emissions control technologies reduce the contributions of anthropogenic sources of these compounds. Some more recent detailed studies of the atmospheric chemistry of these compounds also indicate that these compounds may have short residence times, further reducing their importance for human exposure.

The following subsections provide an overview of the primary atmospheric chemistry and removal processes for SO₂ that are relevant to atmospheric concentrations at urban scales. [Section 2.3.1](#) describes the photochemical reactions that remove SO₂ from the atmosphere by converting it into compounds that condense into the particle or cloud water phase. [Section 2.3.2](#) describes the aqueous-phase oxidation of SO₂, the major oxidation mechanism in the atmosphere, as well as dry and wet deposition of SO₂. [Section 2.3.3](#) summarizes the available research on the atmospheric chemistry of sulfur oxides other than SO₂ and H₂SO₄.

2.3.1 Photochemical Removal of Atmospheric Sulfur Dioxide

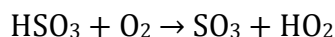
The global atmospheric lifetime (τ) of SO₂ with respect to reactions with the OH radical in the troposphere is 7.2 days. The rate constant for the reaction between SO₂ and NO₃ radical is too small to be important in lowering SO₂ concentrations at urban or regional scales. The same is true for the reaction between SO₂ and the hydroperoxyl (HO₂) radical ([Sander et al., 2011](#)).

In the stepwise oxidation of SO₂ by OH, SO₂ is oxidized to form SO₃, taking the sulfur atom from the S(IV) to S(VI) oxidation state, producing the bisulfite radical (HSO₃):



Equation 2-1

where M is an unreactive gas molecule that absorbs excess destabilizing energy from the SO₂-OH transition state. This reaction is followed by



Equation 2-2

An alternative route involves a stabilized Criegee intermediate (sCI):



Equation 2-3

A Criegee intermediate is a carbonyl oxide zwitterion, i.e. a molecule with separate positive and negative charge centers, derived from the oxidation of an alkene gas molecule by ozone. The unspecified “products” of this reaction are other organic radicals that result from the degradation of the Criegee intermediate ([Berndt et al., 2012](#); [Mauldin et al., 2012](#); [Welz et al., 2012](#)). Rate coefficients for the reaction of sCIs with SO₂ have been reported as 4×10^{-15} cm³/sec ([Johnson et al., 2001](#)), approximately 3.5×10^{-11} cm³/sec ([Liu et al., 2014b](#)), and 3.9×10^{-11} cm³/sec ([Welz et al., 2012](#)). Recent studies report rate coefficients greater than 3×10^{-11} cm³/sec ([Friedman et al., 2016](#); [Lee, 2015](#); [Berndt et al., 2012](#)). These reaction rate coefficients far exceed those of the reactions between these intermediates and H₂O. However, hydrolysis of SO₂ could be limited if sCIs that are potential SO₂ oxidants are hydrolyzed via competing reactions ([Kim et al., 2015](#)). The efficiency of Criegee radical hydrolysis is sensitive to the molecular structure of the alkene. Bimolecular hydrolysis rate constants vary by a factor of 1,000 between syn- and anti-substituted low molecular weight alkenes ([Lin and Takahashi, 2016](#)).

Criegee radicals are produced during both night and day. The relative importance of the OH and sCI pathways depends in large measure on the local concentration of alkenes, such as the low-molecular-weight ones emitted by motor vehicles and industrial processes, as well terpenoids emitted by trees. The importance of this mechanism as a sink for SO₂ is supported by observations that areas adjacent to SO₂ sources, with high biogenic or industrial volatile organic compound (VOC) concentrations, have elevated organic PM concentrations ([Friedman et al., 2016](#)). However, limited information on the identity and concentrations of alkenes at urban scales prevents estimates of the impact of this reaction pathway on urban SO₂ concentrations.

The SO₃ that is generated by either oxidation mechanism (i.e., reaction with OH or via the Criegee reaction) is highly reactive. Water vapor is sufficiently abundant in the troposphere to ensure that SO₃ is quickly converted to gas-phase sulfuric acid, as shown in the equation below ([Loerting and Liedl, 2000](#)).



Equation 2-4

Because H₂SO₄ is extremely water soluble, gaseous H₂SO₄ will rapidly dissolve into the aqueous phase of aerosol particles and cloud droplets. In a study of SO₂ plume transport in and out of foggy conditions, [Eatough et al. \(1984\)](#) observed that roughly 30% of the SO₂ converts to H₂SO₄ particulate each hour when inside a fog bank and roughly 3.1% per hour outside a fog bank. [Khoder \(2002\)](#) observed that conversion from SO₂ to H₂SO₄ increases with increasing relative humidity and O₃ levels, based on a sampling campaign in an urban area of Egypt. Pearson correlation of SO₂-to-H₂SO₄ conversion ratio with

relative humidity was 0.81 in the winter and 0.89 in the summer. [Hung and Hoffmann \(2015\)](#) recently conducted spray chamber experiments of SO₂ to H₂SO₄ conversion. They observed that SO₂ deposited to the surfaces of water microdroplets and then underwent rapid oxidation, first to HSO₃⁻ and bisulfate (HSO₄⁻), and then to SO₄²⁻. Acidic conditions promoted more rapid oxidation of SO₂.

2.3.2 Heterogeneous Oxidation of Sulfur Dioxide

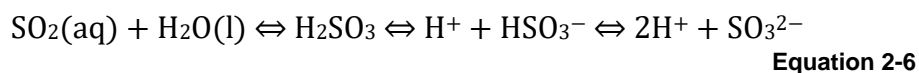
Major sulfur-containing species in clouds include the HSO₃⁻ and SO₃²⁻ (sulfite) ions that form when SO₂ dissolves in cloud droplets and subsequently undergoes acid dissociation. Both species exist in the S(IV) oxidation state and readily oxidize in the presence of aqueous-phase oxidizing agents to form the S(VI) anions, HSO₄⁻ and SO₄²⁻. The major species capable of oxidizing S(IV) to S(VI) in cloud water are O₃, peroxides [either hydrogen peroxide (H₂O₂) or organic peroxides], OH radicals, and transition metal ions such as Fe(II) and Cu(II) that catalyze the oxidation of S(IV) to S(VI) by O₂.

The basic mechanism of the aqueous-phase oxidation of SO₂ can be found in numerous texts on atmospheric chemistry [e.g., ([Seinfeld and Pandis, 2006](#); [Jacobson, 2002](#); [Finlayson-Pitts and Pitts, 2000](#); [Jacob, 1999](#))]. Similar initial steps occur in the fluids lining the airways ([Section 4.2.1](#)). The steps involved are summarized below ([Jacobson, 2002](#)).

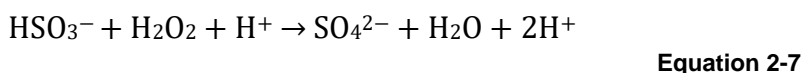
Dissolution of SO₂ occurs first,



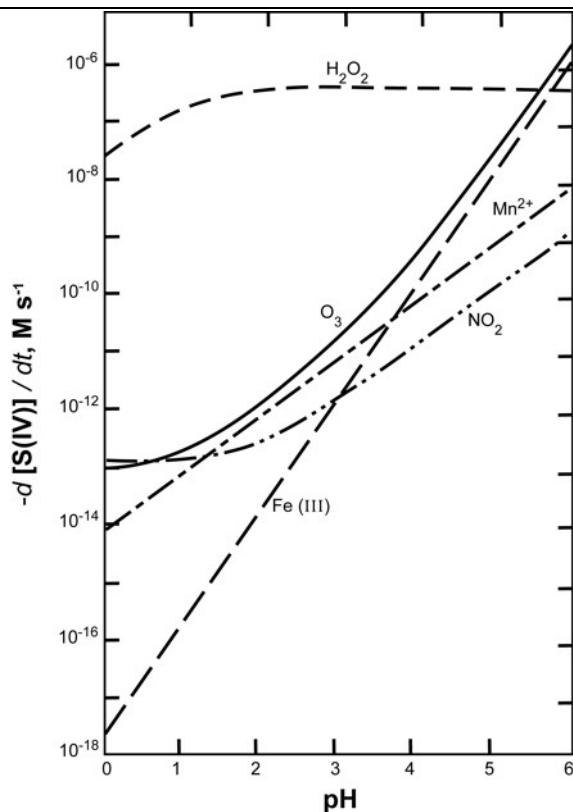
followed by the formation and dissociation of sulfurous acid (H₂SO₃).



In the pH range commonly found in rainwater (2 to 6), H₂O₂ will oxidize HSO₃⁻ to SO₄²⁻.



The rates of aqueous-phase oxidation of S(IV) to S(VI) as a function of pH are shown in [Figure 2-9](#). For pH values up to about 5.3, H₂O₂ is the predominant oxidant; above pH 5.3, O₃, followed by Fe(III), becomes predominant.



Fe = iron; H₂O₂ = hydrogen peroxide; Mn²⁺ = manganese ion; NO₂ = nitrogen dioxide; O₃ = ozone; S = sulfur.

Note: The rate of conversion of aqueous (droplet)-phase S(IV) to S(VI) is shown as a function of pH. Conditions assumed are: [SO₂(g)] = 5 ppb; [NO₂(g)] = 1 ppb; [H₂O₂(g)] = 1 ppb; [O₃(g)] = 50 ppb; [Fe(III)(aq)] = 0.3 μM; [Mn(II)(aq)] = 0.3 μM.

Source: [Seinfeld and Pandis \(2006\)](#). Reprinted with permission of Wiley.

Figure 2-9 The effect of pH on the rates of aqueous-phase sulfur (IV) oxidation by various oxidants.

Ambient ammonia (NH₃) vapor readily dissolves in acidic cloud droplets to form ammonium (NH₄⁺). Because NH₄⁺ is very effective in scavenging H⁺, which shifts the SO₂ oxidation equilibrium, amplifying the rate of oxidation of S(IV) to S(VI) and the rate of dissolution of SO₂ in particles and cloud droplets. Therefore, in environments where NH₃ is abundant, SO₂ is subject to fast removal by cloud and fog droplets and ultimately forms ammonium sulfate [(NH₄)₂SO₄].

Higher pH levels are expected to be found mainly in marine aerosols. In marine aerosols, the chlorine radical-catalyzed oxidation of S(IV) may be more important ([Hoppel and Caffrey, 2005](#); [Zhang and Millero, 1991](#)).

In the same way that it is removed from the gas phase by dissolving into cloud droplets, SO₂ can be removed by dry deposition onto wet surfaces ([Shadwick and Sickles, 2004](#);

[Clarke et al., 1997](#)). For example, in the eastern U.S., more than 85% of sulfur (as SO₂) is removed by dry deposition ([Sickles and Shadwick, 2007](#)). However, aqueous SO₄²⁻ may be removed through occult deposition of large fog or cloud droplets ([Lillis et al., 1999](#); [Pandis and Seinfeld, 1989](#); [Dollard et al., 1983](#)). Scavenging by rain (wet deposition) serves as another removal route. Modeling studies have shown that slightly more than half of SO₂ is lost by gas- and aqueous-phase oxidation, with the remainder of SO₂ loss accounted for by wet and dry deposition ([Long et al., 2013](#); [Liu et al., 2012a](#)).

Sulfur dioxide is known to adhere to and then react on dust particles. Very recent investigations have shown that, for some mineral compositions, SO₂ uptake on dust particles is sensitive to relative humidity, the mineral composition of the particle, and the availability of H₂O₂, the relevant oxidant ([Huang et al., 2015b](#)). Once SO₂ is oxidized to H₂SO₄ on the particle surface, glyoxal, one of the most prevalent organic compounds in the atmosphere, will adhere to the surface and react to form oligomers and organosulfate compounds. This process is enhanced under high humidity conditions ([Shen et al., 2016](#)).

2.3.3 Secondary Gas-phase and Particle-phase Sulfur Oxides

Little information is available in the peer-reviewed literature concerning the atmospheric chemistry of intermediate gas- and particle-phase sulfur oxidation products. One study evaluated the atmospheric residence time of dimethyl sulfate against oxidation by the gas-phase oxidants (O₃, OH, and Cl), along with NH₃ and H₂O ([Japar et al., 1990](#)). Under typical atmospheric conditions, dimethyl sulfate removal will occur in less than 2 days by reaction with water vapor (other removal rates ranged from 23 days to 33 years). Most of these species can be expected to partition into the aqueous particle phase due to their high polarity ([Barnes et al., 2006](#)). Rapid oxidation to H₂SO₄ would be expected to occur in the aqueous phase ([Japar et al., 1990](#)).

Given the technological improvements in burning coal and in controlling emissions in recent decades, one may infer that these species are unlikely to exist in concentrations significant for human exposure. However, the species reported by [Eatough et al. \(1986\)](#) have since been identified as intermediates in the oxidation of dimethyl sulfide ([Barnes et al., 2006](#)), a ubiquitous naturally emitted sulfide species associated with coastal waters and wetlands, vegetation, and soils ([Table 2-2](#)).

Particle-phase inorganic and organic sulfur compounds have been identified in early studies ([Eatough and Hansen, 1983, 1980](#); [Lee et al., 1980](#); [Eatough et al., 1979](#); [Eatough et al., 1978](#); [Smith et al., 1976](#)). These studies identified inorganic SO₃²⁻ complexed with Fe(III) in the particles emitted by a smelter near Salt Lake City, UT. In a detailed spectroscopic study of the transient complexes that form between SO₂, a source of S(IV)

in particles, and Fe(III) in the aqueous phase, [Kraft and van Eldik \(1989a\)](#) and [Kraft and van Eldik \(1989b\)](#) reported that the oxidation of S(IV) by Fe(III) to form SO_4^{2-} occurs on the order of seconds to minutes and is further accelerated by low pH. Sulfuric acid is well known to absorb water at even low ambient relative humidity ([Seinfeld and Pandis, 2006](#)). The highly acidic aqueous conditions that arise once smelter plume particles equilibrate with the ambient atmosphere ensure that S(IV)-Fe(III) complexes have a small probability of persisting and becoming a matter of concern for human exposure. Substantial effort in recent years has been applied to understanding the mechanism for the formation of organic sulfur compounds, and the results of this effort are described in [Section 2.3.1](#).

2.4 Measurement Methods

This section discusses the Federal Reference Method (FRM) and Federal Equivalent Method (FEM) used for NAAQS compliance as well as the state, local, and tribal monitoring networks across the U.S. used for NAAQS compliance monitoring. Detailed information about monitoring methods, including accuracy, precision, limits of detection, and other operational parameters was published in the 1982 Air Quality Criteria for Particulate Matter and Sulfur Oxides Volume II ([U.S. EPA, 1982a](#)) and then updated in Appendix B.6 of the 2008 ISA for Sulfur Oxides—Health Criteria ([U.S. EPA, 2008d](#)). The List of Designated Reference and Equivalent Methods ([U.S. EPA, 2016e](#)) includes all monitors approved as FRMs or FEMs and provides monitor specifications. A brief summary of that information, along with a discussion of more recent studies evaluating FRMs and FEMs for monitoring SO_2 concentration ([Section 2.4.1](#)) or alternative SO_2 monitoring methods ([Section 2.4.2](#)), is provided. [Section 2.4.3](#) describes the sampling network.

2.4.1 Federal Reference and Equivalent Methods

Currently, there are two FRMs for the measurement of SO_2 —the manual pararosaniline wet-chemistry method and the automated pulsed ultraviolet fluorescence (UVF) method. The manual method was approved as an FRM in the 1970s and was quickly supplanted by the flame photometric detection (FPD) method, an FEM, because the manual method was too complex and had a slow response even in automated form. The UVF method was designated as an FEM in the late 1970s and ultimately supplanted the FPD method. The UVF method became the dominant method in routine monitoring networks because it is inherently linear and relatively safe, whereas the FPD method requires highly flammable hydrogen gas. The UVF method has been the most commonly used method by state and

local monitoring agencies since the 1980s. It was promoted to FRM status during the promulgation of the new 1-hour SO₂ primary NAAQS established in 2010 (75 FR 35520) in light of its reliability and well-documented performance. The UVF method supports the need for a continuous monitoring method, providing highly time-resolved data, such as 5-minute data as well as routine 1-hour SO₂ measurements. 1-hour SO₂ measurements. The existing pararosaniline manual method was retained as an FRM, and although cumbersome, the method can provide hourly measurements to support the 1-hour NAAQS.

In the UVF method, SO₂ molecules absorb ultraviolet (UV) light at one wavelength, elevating the molecule to a higher energy electronic state. Once electronically excited, the molecule loses a portion of its energy by colliding with another gas molecule. The molecule relaxes back to its electronic ground state by emitting a photon of light at a longer wavelength (i.e. lower energy) than the light used to excite the molecule. The intensity of the emitted light is, therefore, proportional to the number of SO₂ molecules in the sample gas. In commercial analyzers, light from a high-intensity UV lamp passes through an optical bandwidth filter that allows only photons with wavelengths around the SO₂ absorption peak (near 214 nm) to enter the optical chamber. The light passing through the bandwidth filter is collimated using a UV lens and passes through the optical chamber to a reference photomultiplier detector. The detector is set perpendicular to the illumination path to maximize the collection of SO₂ fluorescence. An optical bandwidth filter designed to block higher energy frequencies is added to further protect the detector from stray UV lamp light. Quartz lenses are positioned between the filter and the detector to focus the SO₂ fluorescence photons onto the active area of the detector, optimizing the fluorescence signal.

Studies have compared UVF to sampled SO₂ from impregnated filters for quality assurance. Comparison of 24-hour avg concentration measurements obtained with the UVF method and with impregnated filters showed annual average differences within ± 0.07 ppb, based on data obtained between 1993 and 2001 from four Finnish cities ([Leppänen et al., 2005](#)). [Ferek et al. \(1997\)](#) evaluated the Teco model UVF (developed at the University of Washington) against carbonate-impregnated filters for measurement of SO₂ concentration in laboratory studies. The Teco UVF measured SO₂ concentrations down to 16 ppt and on average, produced a positive difference of 7% compared with the filter. The Teco UVF analyzed data at a frequency as high as 1 Hz, but noise was curtailed by averaging up to 10 minutes. The [Ferek et al. \(1997\)](#) study highlighted the Teco UVF but also included other SO₂ measurement techniques in the SO₂ monitor comparison, including gas spectrometry/mass spectrometry, high performance liquid chromatography, and a mist chamber, which produced a maximum of 30% differences for filter-measured SO₂ concentrations of 3–4 ppb averaged over 90 minutes.

2.4.1.1 Minimum Performance Specifications

During the 2010 SO₂ NAAQS review, minimum performance specifications [contained in 40 Code of Federal Regulations (CFR) Part 53] were updated and became more stringent for any new FRM and FEM automated method. The new specifications are provided in [Table 2-3](#). The previous specifications were based on the older, manual, wet-chemistry FRM and were updated to reflect current technology and improved performance in SO₂ instrumentation. The lower detection limit (LDL) for a routine, automated SO₂ analyzer is required to be 2 ppb. As part of the National Core (NCore) monitoring network, new trace-level SO₂ instruments have been developed and added to State and Local Air Monitoring Sites (SLAMS). These new trace-level (i.e., low LDL) instruments have LDLs of 0.2 ppb or lower. Note that FRMs and FEMs may have more stringent performance characteristics than the minimum performance specifications presented in [Table 2-3](#).

Depending on design, instrument settings, and operator preferences, SO₂ instruments can have different time constants or averaging times [i.e., the combined rise and fall time of the signal detection system (75 FR 35520)] ([Ferek et al., 1997](#)). Instruments operated with longer time constants will respond more slowly to short-lived, high concentration spikes in SO₂, reporting a broader peak with a lower maximum concentration compared with instruments operated with shorter time constants. The delayed decline to baseline SO₂ concentrations can influence concentration measurements at adjacent 5-minute intervals. Short concentration spikes may also not be fully detected or may be time-shifted due to the division of each hour into twelve discrete 5-minute intervals.

Table 2-3 Minimum performance specifications for sulfur dioxide established in 40 Code of Federal Regulations Part 53, Subpart B.

Performance Parameter	Specification
Range ¹	0–0.5 ppm (500 ppb)
Noise	0.001 ppm (1 ppb)
Lower detection limit (<i>two times the noise</i>)	0.002 ppm (2 ppb)
Interference equivalent <ul style="list-style-type: none"> • Each interferent • Total, all interferents 	±0.005 ppm (5 ppb) —
Zero drift (12 and 24 h)	±0.004 ppm (4 ppb)
Span drift (24 h) <ul style="list-style-type: none"> • 20% of upper range limit • 80% of upper range limit 	— ±3.0%
Lag time	2 min
Rise time	2 min
Fall time	2 min
Precision <ul style="list-style-type: none"> • 20% of upper range limit • 80% of upper range limit 	2.0% 2.0%

¹The CFR also provides for testing at lower ranges with special performance limit requirements.

2.4.1.2 Positive and Negative Interferences

The UVF method has a number of positive and negative interferences. The most frequent source of positive interference is other gases that fluoresce at the same wavelength as SO₂. The most common gases include volatile organic compounds (e.g., xylenes, benzene, toluene) and polycyclic aromatic hydrocarbons (PAHs; e.g., naphthalene). To reduce this source of positive interference, high-sensitivity SO₂ analyzers are equipped with scrubbers or “kickers” to remove these compounds from the air stream prior to entering the optical chamber. [Luke \(1997\)](#) evaluated a modified pulsed fluorescence SO₂ detector and found positive interference from nitric oxide (NO), CS₂, and several highly fluorescent aromatic hydrocarbons such as benzene, toluene, *o*-xylene, *m*-xylene, *p*-xylene, *m*-ethyltoluene, ethylbenzene, and 1,2,4-trimethylbenzene. The positive

artifacts could be virtually eliminated by using a hydrocarbon “kicker” membrane. At a flow rate of 300 standard cm³/minute and a pressure drop of 645 torr across the membrane, the interference from ppm levels of many aromatic hydrocarbons can be eliminated.

Another source of positive interference is NO, which fluoresces in a region close to that of SO₂. However, in high-sensitivity SO₂ analyzers, the bandpass filter in front of the detector is specifically designed to prevent detection of NO fluorescence at the detector. Care must be exercised when using multicomponent calibration gases containing both NO and SO₂, so that the NO rejection ratio of the SO₂ analyzer is sufficient to prevent NO interference.

The most common source of positive bias in high-sensitivity SO₂ analyzers is stray light in the optical chamber. Because SO₂ can be excited by a broad range of UV wavelengths, any stray light entering the optical chamber with an appropriate wavelength can excite SO₂ in the air stream and increase the fluorescence signal. Additionally, stray light entering the optical chamber with a similar wavelength of SO₂ fluorescence may impinge on the detector and increase the fluorescence signal. Stray light is also minimized with changes in instrument design such as use of light filters, dark surfaces, and opaque tubing.

H₂O is a common source of negative interference in high-sensitivity SO₂ monitors. When excited SO₂ molecules collide with water vapor as well as other common molecules in air (e.g., nitrogen and oxygen), nonradiative deactivation (quenching) can occur. During collisional quenching, the excited SO₂ molecule transfers energy, kinetically allowing the SO₂ molecule to return to a lower energy state without emitting a photon. Collisional quenching decreases the SO₂ fluorescence and results in an underestimation of SO₂ concentration in the air sample. Of particular concern is the variable water vapor content of air. [Luke \(1997\)](#) reported that the response of the detector could be reduced by an amount of approximately 7 to 15% at water vapor mixing ratios of 1 to 1.5 mole percent [relative humidity (RH) = 35 to 50% at 20 to 25°C and 1 atmosphere for a modified pulsed fluorescence detector (Thermo Environmental Instruments, Model 43s)]. Systems may minimize this interference by maintaining instrument and sampling lines at markedly higher operational temperatures than the expected dew point and to within a few degrees of the controlled optical bench temperature. Condensation of water vapor in sampling lines must be avoided, as water on the inlet surfaces can absorb SO₂ from the sample air. Condensation is normally prevented by heating sampling lines to a temperature above the expected dew point and to within a few degrees of the controlled optical bench temperature. Some monitors are equipped with a dryer system to remove moisture from the sample gas before it reaches the particulate filter.

2.4.2 Alternative Sulfur Dioxide Measurements

A number of optical methods for measuring SO₂ are available. They include laser induced fluorescence (LIF), cavity ring-down spectroscopy (CRDS), differential optical absorption spectrometry (DOAS), and UV light absorption. There are also methods based on mass spectroscopy or mass spectrometry [e.g., chemical ionization mass spectroscopy (CIMS) and atmospheric pressure ionization mass spectrometry (APIMS)]. These methods are often too expensive and complex for routine monitoring applications and are more suitable for source monitoring and research. However, approaches to reduce interferences and increase SO₂ selectivity could be extended to FRM and FEM instrumentation. The LIF, CRDS, and DOAS methods will be discussed below as they have the potential to provide trace-level SO₂ measurements or have shown good agreement with UVF instrumentation.

LIF is a technique that can provide high sensitivity for ambient SO₂ measurements and reduces interferences with species that fluoresce at the same wavelength as SO₂. Both tunable and nontunable laser sources have been evaluated. [Matsumi et al. \(2005\)](#) evaluated a LIF method using a tunable laser at an SO₂ absorption peak at 220.6 nm and trough at 220.2 nm. The difference between the signals at the two wavelengths is used to estimate the SO₂ concentration. This technique has a sensitivity of 5 ppt in 60 sec. [Simeonsson et al. \(2012\)](#) evaluated a direct LIF technique using a nontunable laser source at an absorption wavelength of 223 nm, which coincides with the SO₂ absorption peak. This technique has a high sensitivity with LDL of 0.5 ppb. Both the tunable and nontunable instruments have low LDL (≤ 0.5 ppb); therefore, they can provide trace-level SO₂ measurements.

CRDS is an optical absorption method based on measurement of the rate of light absorption through a sample. CRDS has successfully been used to measure ambient NO₂ and NO with high sensitivity. [Medina et al. \(2011\)](#) compared a CRDS-tunable laser method to the routinely used pulsed UVF method for measuring SO₂. At an absorption wavelength of 308 nm, the CRDS had an LDL of 3.5 ppb, which was higher than those for routine and trace-level UVF SO₂ monitors (e.g., Thermo Scientific 43i and Thermo Scientific 43i-TLE). However, the response time was faster compared to the UVF methods (a few seconds vs. 80 seconds). To reduce interferences, a ferrous sulfate scrubber was used to remove NO₂ and O₃, and a denuder was used to zero SO₂ levels. Improvements could be made to increase the sensitivity to about 1 ppb by changing the placement of the mirrors to optimize laser light reaching the cavity or using a better detection system. Additionally, improving the mirror reflectivity could improve the sensitivity to about 0.1 ppb, similar to the detection levels of trace-level SO₂ monitors.

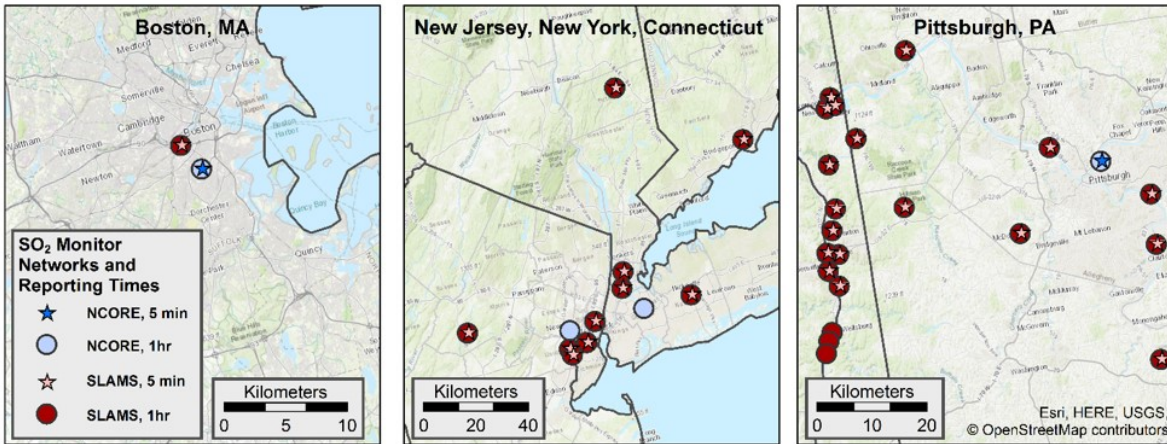
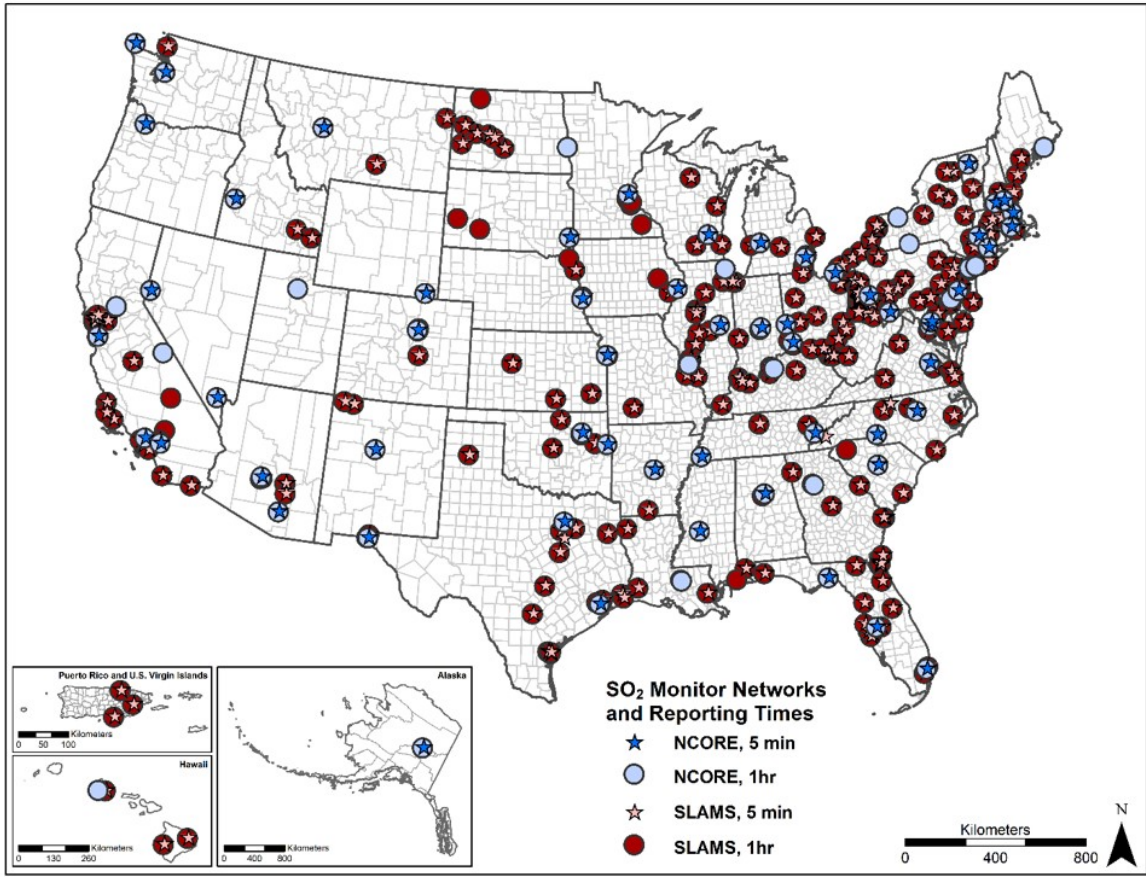
DOAS is an optical remote sensing method based on the absorption of light in the UV-visible wavelength region to measure atmospheric pollutants. [Kim and Kim \(2001\)](#) compared SO₂ concentrations measured using a DOAS system with daily mean SO₂ concentrations measured by an in situ monitor in Seoul, Korea during a 13-month period. In this study, the DOAS typically reported SO₂ concentrations around 10–40% above the in situ technique, but SO₂ concentrations measured by the DOAS were sometimes 100–200% below those measured with the in situ monitor. Across all measurements, the daily mean SO₂ concentration was 36% higher from the DOAS compared with the in situ monitor. Discrepancies between the two methods were attributed to ability to respond to meteorological factors. The DOAS was reported to have an LDL of 0.07 ppb, compared with 1 ppb reported for the in situ method. A newer technique called multiaxis differential optical absorption spectroscopy (MAX-DOAS) has been developed that offers increased sensitivity in measuring SO₂ ([Honninger et al., 2004](#)). MAX-DOAS is based on the measurement of scattered sunlight at multiple viewing directions and can obtain both surface concentrations and vertical column density of SO₂. [Wang et al. \(2014b\)](#) compared MAX-DOAS SO₂ column measurements in the 305 to 317.5 nm absorption wavelength to surface SO₂ measurements from a modified UVF SO₂ monitor (Thermo Environmental Instruments Model 43C) and found good agreement ($r = 0.81$, slope = 0.90).

Remote sensing by satellites (e.g., OMI, infrared atmospheric sounding interferometer, etc.) is an emerging technique for measuring SO₂, as well as other pollutants. This technique can be used for a variety of applications, including air quality management (e.g., augmenting ground-based monitors, assessing emissions inventories), studying pollutant transport, assessing emissions reductions, and evaluating air quality models. Remote sensing methods employ a retrieval system using a combination of solar backscatter or thermal infrared emission spectra and mathematical algorithms to estimate pollutant concentrations. Remote sensing from space is particularly challenging for SO₂ measurements for two reasons: (1) air scattering causes SO₂ to have a low optical thickness (three orders of magnitude lower than O₃), so that only large SO₂ sources can be observed ([Bogumil et al., 2003](#)) and (2) emissions reductions programs have led to lower SO₂ emissions from stationary sources, making it more difficult to see anthropogenic SO₂ emissions ([Streets et al., 2014](#)). The majority of remote sensing studies have focused on large natural sources (e.g., volcanoes), large anthropogenic sources (e.g., coal-burning power plants, smelters), fuel extraction from oil sands, and newly constructed coal-burning facilities with high, uncontrolled SO₂ emissions ([Boynard et al., 2014](#); [McCormick et al., 2014](#); [Streets et al., 2014](#); [Clarisse et al., 2012](#); [McLinden et al., 2012](#); [Fioletov et al., 2011](#); [Nowlan et al., 2011](#); [Bobrowski et al., 2010](#); [Li et al., 2010](#); [Khokhar et al., 2008](#); [Carn et al., 2007](#)).

2.4.3 Ambient Sampling Network Design

Compliance with NAAQS is primarily carried out through the SLAMS network, although modeling may also be used to characterize air quality for implementation purposes (75 FR 35520). There were 427 SLAMS sites that reported 1-hour SO₂ concentrations in 2015 to the Air Quality System (AQS), U.S. EPA's repository for detailed air pollution data that is subject to quality control and assurance procedures. In addition to their use in compliance evaluations, some of these sites function as central monitoring sites for use in epidemiologic studies. The SLAMS network also reports either the maximum 5-minute concentration in the hour (one of twelve 5-minute periods within an hour) or all twelve 5-minute average SO₂ concentrations within the hour. Siting requirements for monitors in the SLAMS network can be found in 40 CFR Part 58, Appendix E.

The SLAMS network includes the NCore monitoring network, which began January 1, 2011 and consisted of 74 mostly urban sites in 2015. NCore is a multipollutant measurement network and includes SO₂ measurements as well as measurements for other gaseous pollutants (O₃, CO, NO_x, oxides of nitrogen), PM_{2.5}, PM_{10-2.5}, and meteorology. NCore is focused on characterizing trends in pollutants, understanding pollutant transport in urban and rural areas, and evaluating data with respect to the NAAQS. [Figure 2-10](#) shows the locations of these monitoring networks across the U.S. The Clean Air Status and Trends Network (CASTNet) also measures ambient SO₂. However, these data are not used for NAAQS compliance purposes and are obtained predominantly in National Parks or other ecologically sensitive sites. Because CASTNet monitors are not deployed in populated areas, they are not useful in evaluating the health effects of SO₂. This network provides weekly averages of total sulfur (dry SO₂, dry SO₄²⁻, and wet SO₄²⁻) in about 90 sites located in or near rural locations to assess long-term trends in acidic deposition associated with emissions reduction programs. CASTNet data are presented in the Integrated Science Assessment for Oxides of Nitrogen and Sulfur—Ecological Criteria ([U.S. EPA, 2008b](#)).



NCORE = National Core; SLAMS = State and Local Air Monitoring Sites; SO₂ = sulfur dioxide.

Figure 2-10 Routinely operating sulfur dioxide monitoring networks: National Core and State and Local Air Monitoring Sites, reporting 1 hour and 5 minute sulfur dioxide concentration data. Maps for the Boston, MA, New York City, and Pittsburgh, PA metropolitan areas are provided as examples of the variation in monitor placement in the US Northeast.

The minimum monitoring requirements for the SLAMS network are outlined in 40 CFR Part 58, Appendix D. Sulfur dioxide monitors at SLAMS sites represent four main spatial scales: (1) microscale—areas in close proximity, up to 100 m from a SO₂ point or area source, (2) middle scale—areas up to several city blocks, with linear dimensions of about 100 to 500 m, (3) neighborhood scale—areas with linear dimensions of 0.5 to 4 km, and (4) urban scale—urban areas with linear dimensions of 4 to 50 km. Maximum hourly SO₂ concentrations are established based upon measurements taken at the 3 smaller scales [i.e. micro-, middle- and neighborhood] to account for near-source and neighborhood-scale concentrations. Urban-scale sites are sometimes used as central monitoring sites to characterize population exposures and trends, such as in epidemiologic studies ([Section 3.2.1](#)). Urban-scale sites can also be used to determine background concentrations in areas where monitors are located upwind of a local source. There are also a number of regional-scale monitoring sites, representing length scales of tens to hundreds of kilometers, typically in rural areas of uniform geography without large SO₂ sources. These sites can be used to determine the amount of regional pollution transport and to support secondary NAAQS.

Stationary sources are the primary emission sources of SO₂. Prior to the enactment of the revised SO₂ primary NAAQS in 2010, U.S. EPA evaluated about 488 SO₂ monitoring sites in operation during 2008 and found that the network was insufficient to support the proposed NAAQS ([U.S. EPA, 2009e](#)). To partially address this deficiency, U.S. EPA promulgated minimum monitoring requirements based on a near-source monitoring approach. The Population Weighted Emissions Index (PWEI), which is based on population and emissions inventory data at the core-based statistical area (CBSA) level, was introduced to assign the appropriate number of monitoring sites in a given CBSA (75 FR 35520). The PWEI accounts for SO₂ exposure by requiring monitor placement in urban areas where population and emissions may lead to higher potential for population exposure to maximum hourly SO₂ concentrations. The PWEI value is calculated by multiplying the population of each CBSA by the total amount of SO₂ emissions (in tons per year) in a given CBSA, using the most recent census data (or estimates) and combining the most recent county-level emissions data (from the National Emissions Inventory) for each county in each CBSA, respectively. This value is then divided by 1 million, resulting in a PWEI value with units of million person-tons per year. A minimum of three SO₂ monitoring sites is required for any CBSA with a PWEI value greater than or equal to 1,000,000. For any CBSA with a PWEI value greater than or equal to 100,000 but less than 1,000,000, a minimum of two SO₂ monitoring sites is required. Lastly, a minimum of one SO₂ monitoring site is required for any CBSA with a PWEI value greater than or equal to 5,000 but less than 100,000. The monitors sited within a CBSA based on the PWEI criterion should also be, at minimum, one of the

following monitoring site types: population exposure, highest concentration, source impacted, general background, or regional transport.

2.5 Environmental Concentrations

This section provides an overview of SO₂ ambient and background concentrations. Analyses are focused on characterizing recent SO₂ concentration data from the U.S. rather than the influence of atmospheric stability and meteorological conditions on concentration distributions. Information on previous SO₂ concentrations can be found in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) and earlier documents. SO₂ data discussed in this section were obtained from the AQS. [Section 2.5.1](#) introduces different SO₂ metrics used for NAAQS compliance and epidemiologic applications. Ambient concentrations of SO₂ are then discussed on various spatial and temporal scales. Spatial variability is discussed in [Section 2.5.2](#), which is divided into two sections discussing large-scale variability (i.e., nationwide) and small-scale variability (i.e., urban areas). Temporal variability is then discussed in [Section 2.5.3](#), extending from multiyear trends to subhourly variations. The relationships between 5-minute hourly max and 1-hour concentrations are described in [Section 2.5.4](#). Background SO₂ concentrations from natural sources are subsequently discussed in [Section 2.5.5](#).

2.5.1 Sulfur Dioxide Metrics and Averaging Time

Different metrics are used to represent ambient SO₂ concentrations for epidemiologic analysis and NAAQS compliance. As discussed in [Section 2.5.4](#), hourly and 5-minute concentration data are routinely reported to U.S. EPA's AQS data repository by state, local, and tribal agencies. Metrics can be derived from these hourly and 5-minute data to represent concentration and exposure levels on different time scales. [Table 2-4](#) provides information on how different SO₂ metrics are derived. Daily metrics include the 24-hour avg SO₂ concentration and the 1-hour daily max SO₂ concentration. Hourly metrics include the 5-minute hourly max concentration reported during a given hour and the 1-hour avg concentration. Metrics derived using maximum concentration statistics (i.e., 1-hour daily max or 5-minute hourly max) provide insight about peak ambient concentrations occurring over a given hour or day.

The following sections include national and urban statistics on daily and hourly metrics. When interpreting the statistics, it is important to consider the aggregation time when comparing the magnitude and range of ambient concentrations related to different metrics.

Table 2-4 Summary of sulfur dioxide metrics and averaging times.

Metric	Aggregation Time	Averaging Time Description
24-h avg	Daily	Daily mean of 1-h avg SO ₂ concentrations
1-h daily max	Daily	Maximum 1-h SO ₂ concentration reported during the day
1-h avg	Hourly	Hourly mean SO ₂ concentrations reported during the day
5-min hourly max	Hourly	Maximum 5-min SO ₂ concentration reported during 1-h

Avg = average; max = maximum; SO₂ = sulfur dioxide.

AQS SO₂ data used to compute national statistics meet the data quality and completeness criteria listed in [Table 2-5](#). Three additional criteria were applied for the 5-minute data to reduce the influence of outliers: (1) the 5-minute data had to correspond to an hourly data concentration, (2) the mean of the 5-minute data could be no more than 120% of the hourly mean, and (3) the 5-minute hourly max concentration had to fall within 1 to 12 times the 1-hour avg concentration. The AQS, by convention, accepts values that fall within the range defined by the positive and negative of the absolute value of the instrument's lower detection limit. This analysis included those values. Based on these criteria, statistics were computed for data from a total of 380 sites across the U.S. for 5-minute hourly max SO₂ concentrations and for data from a total of 438 sites for the 1-hour daily max, 24-hour avg, and 1-hour avg SO₂ metrics. Thirteen percent of sites did not have 5-minute data for comparison with 1-hour data.

2.5.2 Spatial Variability

This section provides a brief overview of national- and urban-scale SO₂ spatial variability and discusses how variations in ambient SO₂ concentrations influence human exposure in different geographical regions.

Table 2-5 Summary of sulfur dioxide data sets originating from the Air Quality System database.

AQS SO ₂ data used to compute national statistics (to meet the data quality and completeness criteria)	
Years	2013–2015
Months	January–December
Completeness criteria	75% of 5-min periods in an hour (where 5-min data are available)
	75% of h in day
	75% of days in calendar quarter
	3 of 4 quarters of the yr
Number of monitoring sites meeting completeness criteria	380 sites reporting 5-min data (2013–2015) 438 sites reporting 1-h data (2013–2015)

AQS = Air Quality System; SO₂ = sulfur dioxide.

2.5.2.1 Nationwide Spatial Variability

In the previous ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), 24-hour avg, 1-hour daily max, 1-hour avg, and 5-minute hourly max SO₂ concentrations measured at AQS monitoring sites during 2003–2005 were reported. Nationwide statistics of 5-minute hourly max SO₂ data were limited in the previous assessment by a scarcity of monitoring sites reporting such data. From 2003–2005 nationwide, central statistics (mean and median) of 1-hour daily max and 24-hour avg SO₂ concentrations were generally low (less than 15 ppb), while concentrations in the upper range of the distribution (e.g., 99th percentile) were substantially higher (23–116 ppb), particularly for 1-hour daily max concentrations (99th percentile: 116 ppb). In addition, 1-hour avg SO₂ concentrations exhibited low mean concentrations (4 ppb), with 99th percentile concentrations near 34 ppb. Relatively high concentrations were typically observed at sites near stationary anthropogenic sources (e.g., EGUs).

SO₂ summary data provide a snapshot of recent concentrations and, compared with those presented in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), allow for identification of possible trends. As shown by [Table 2-6](#), nationwide concentrations for 2013–2015 were slightly lower than concentrations reported in the 2008 SO_x ISA. For all 24-hour avg, 1-hour daily max, 1-hour avg, and 5-minute hourly max data pooled nationwide, mean statistics were below 6 ppb, median statistics (50th percentile) were 2 ppb or below, and SO₂

concentrations in the upper range of the distribution (99th percentile) covered a wide range of concentrations but were never greater than the primary NAAQS level of 75 ppb. The 99th percentile 5-minute hourly max concentration was 23.8 ppb, suggesting that the occurrence of very high 5-minute peak values is rare on the national scale. Across all metrics, large differences were observed between mean and 99th percentile concentrations, particularly for the SO₂ 1-hour daily max and 5-minute hourly max data. Such large differences between mean and 99th percentile concentrations are consistent with the highly variable nature of SO₂, which is characterized by periodic peak concentrations superimposed on a relatively low background concentration. Higher concentrations in the 1-hour daily max distribution compared with the 5-minute hourly max distribution were likely attributable to the omission of high 5-minute concentrations from the 58 monitoring sites without 5-minute data.

The absolute highest 1-hour daily max SO₂ concentration in 2013–2015 was 2,071 ppb. Ninety-ninth percentile 1-hour daily max concentrations over 200 ppb were reported at this site and other sites near active volcanoes in Hawaii ([Table 2-6](#)), which are discussed further in [Section 2.5.5](#). Other reports of 99th percentile, 1-hour daily max concentrations greater than 200 ppb occurred at three monitoring sites near a copper smelter in Gila County, AZ, as mentioned in [Section 2.2.2](#). In addition, sites where the 99th percentile 1-hour daily max concentration was greater than 75 ppb were located in North Dakota, Illinois, Iowa, Wisconsin, Arizona, Missouri, Indiana, Tennessee, Ohio, Kentucky, Louisiana, and Pennsylvania, often near coal-fired EGUs, although the site in North Dakota is likely influenced by a nearby shale gas processing facility. As shown in the nationwide map in [Figure 2-11](#), the majority of monitoring sites across the U.S. report 99th percentile, 1-hour daily max concentrations below the primary NAAQS level of 75 ppb. The 99th percentile of 24-hour avg concentrations, which are often used as exposure metrics in epidemiologic studies, followed a similar pattern, with most elevated values located in the industrial Midwest ([Figure 2-12](#)).

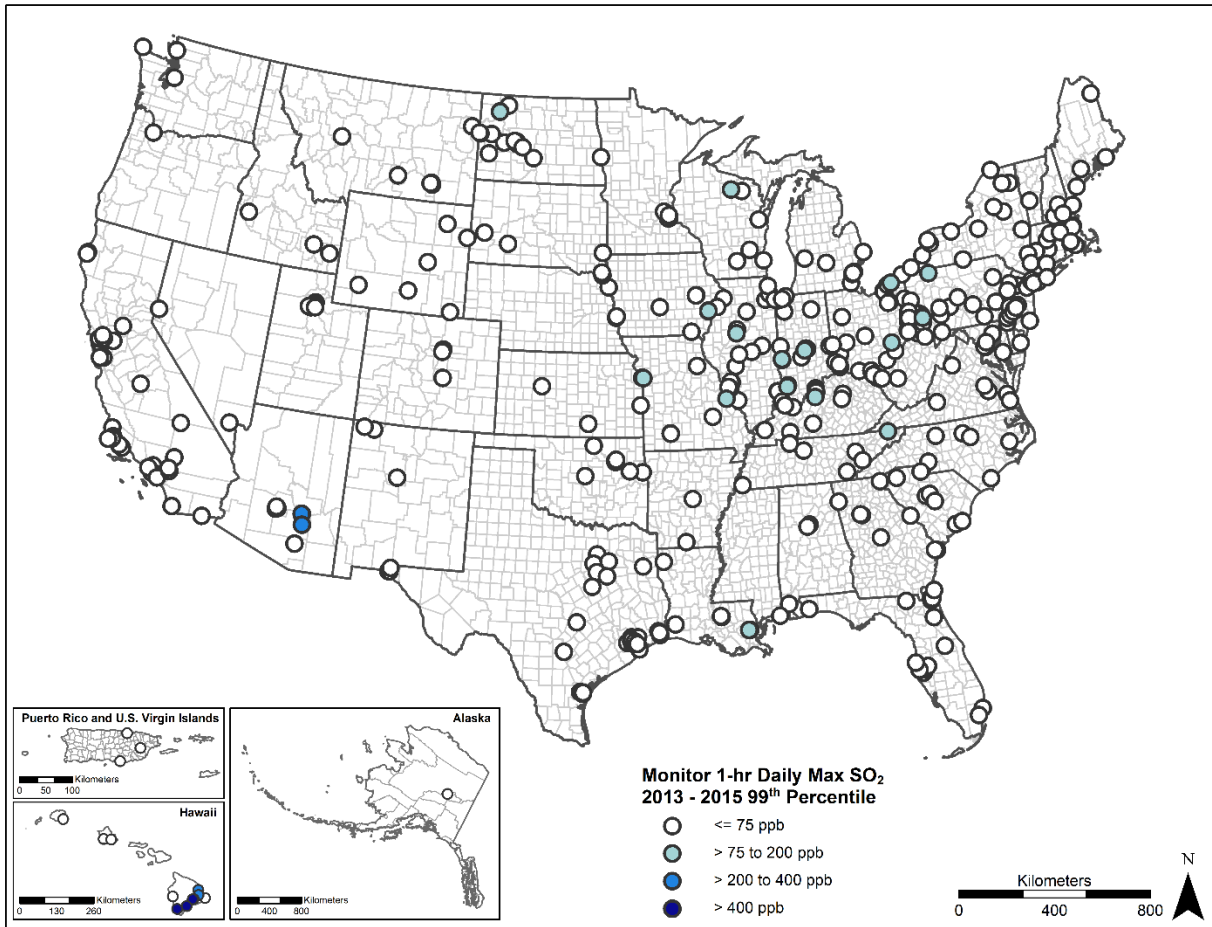
Table 2-6 National statistics of sulfur dioxide concentrations (parts per billion) from Air Quality System monitoring sites, 2013–2015.

Year	N of Obs	Mean	5%	10%	25%	50%	75%	90%	95%	98%	99%	Max	AQS Max ID ^a
5-min hourly max													
2013	3,105,078	2.2	0.0	0.0	0.1	1.0	2.0	4.0	7.0	15.0	25.3	1,441.4	160050004
2014	3,047,302	2.2	0.0	0.0	0.2	1.0	2.0	4.0	7.0	14.7	25.0	4,208.0	160050004
2015	2,997,344	1.8	0.0	0.0	0.1	0.8	1.5	3.0	5.2	11.5	20.0	1,678.0	160050004
2013–2015	9,149,724	2.1	0.0	0.0	0.1	0.9	2.0	4.0	6.5	13.8	23.8	4,208.0	160050004
1-h avg													
2013	3,105,078	1.6	0.0	0.0	0.0	0.8	1.7	3.1	5.0	9.0	15.4	2,071.0	150010007
2014	3,047,302	1.6	0.0	0.0	0.0	0.7	1.5	3.0	5.0	9.3	15.7	1,830.0	150010007
2015	2,997,344	1.3	0.0	0.0	0.0	0.6	1.1	2.5	4.0	8.0	13.0	1,779.0	150010007
2013–2015	9,149,724	1.5	0.0	0.0	0.0	0.7	1.4	3.0	4.9	9.0	15.0	2,071.0	150010007
1-h daily max													
2013	134,705	5.6	0.0	0.0	0.9	2.0	4.4	10.3	18.9	37.0	62.2	2,071.0	150010007
2014	132,228	5.7	0.0	0.0	0.8	2.0	4.4	11.0	19.7	40.7	68.0	1,830.0	150010007
2015	129,789	4.7	0.0	0.0	0.6	1.4	3.3	8.1	15.7	34.2	60.0	1,779.0	150010007
2013–2015	396,722	5.3	0.0	0.0	0.7	1.8	4.0	10.0	18.0	37.4	63.5	2,071.0	150010007
2-h avg													
2013	134,705	1.6	0.0	0.0	0.2	0.9	1.8	3.5	5.1	8.5	13.1	366.5	150010007
2014	132,228	1.5	0.0	0.0	0.2	0.8	1.7	3.3	5.0	8.5	13.1	317.2	150010007
2015	129,789	1.3	0.0	0.0	0.2	0.7	1.4	2.7	4.0	7.4	12.0	341.6	150010007
2013–2015	396,722	1.5	0.0	0.0	0.2	0.8	1.6	3.2	4.8	8.2	12.7	366.5	150010007

AQS = Air Quality System; avg = average; ID = identification; mean = arithmetic average; max = maximum; N = population number; Obs = observations.

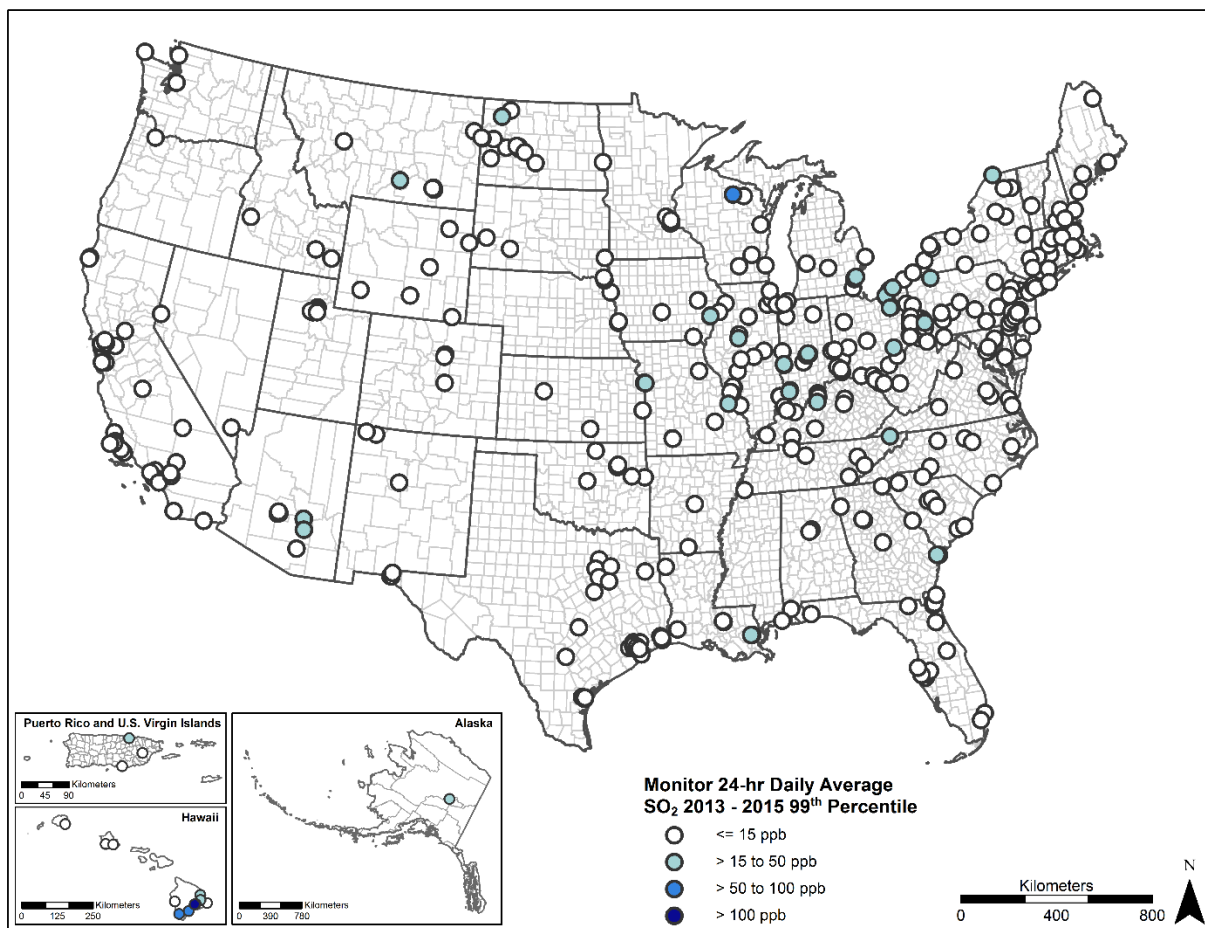
^aAQS site ID number reporting the highest 3-yr concentration across the U.S.

AQS accepts as valid any reported concentration that is \pm LDL. Data analyzed in this table include negative values in this range. Note: Not all sites collect 5-min measurements. Site ID 150010007 does not collect 5-min measurements (i.e., the location of the peak 1-hr max, 1-hr daily average and 24-hr average corresponding to a monitor adjacent to the Hawaiian volcanoes).



Max = maximum; SO₂ = sulfur dioxide.

Figure 2-11 Map of 99th percentile of 1-hour daily max sulfur dioxide concentration reported at Air Quality System monitoring sites, 2013–2015.



Note: The 24-h avg concentration is a metric often used in epidemiologic studies.
 SO₂ = sulfur dioxide.

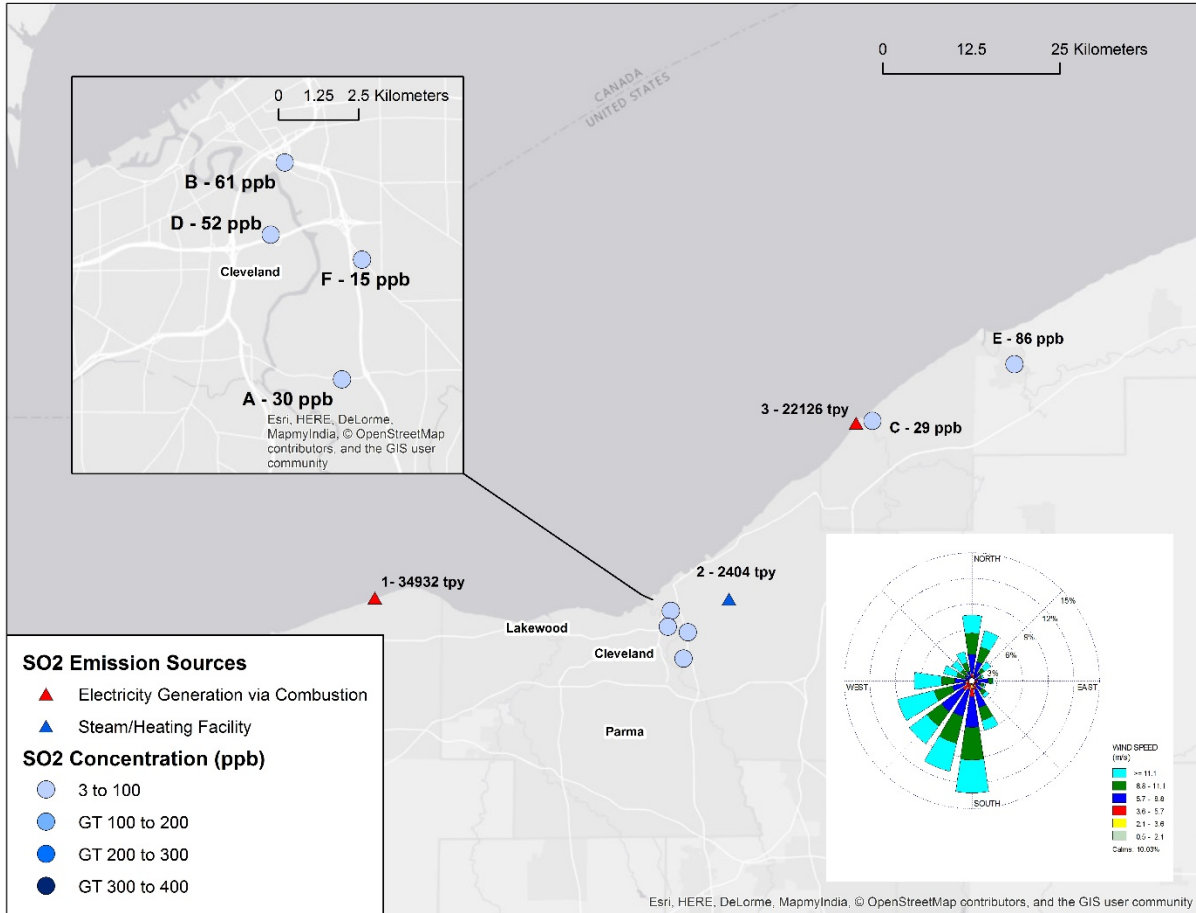
Figure 2-12 Map of 99th percentile of 24-hour avg sulfur dioxide concentration reported at Air Quality System monitoring sites, 2013–2015.

2.5.2.2 Urban Spatial Variability

Air quality measurements from centrally located, urban monitoring sites are often used to represent community-scale exposure in epidemiologic analyses. However, central site exposure estimates may not fully capture variations in pollutant concentrations over urban scales. Urban areas differ in topography, source types, and source locations. To illustrate the effects of these differences on urban scale concentration gradients, SO₂ spatial variability was characterized in six focus areas: Cleveland, OH; Pittsburgh, PA; New York City, NY; St. Louis, MO; Houston, TX; and Gila County, AZ. These focus

areas were selected based on (1) their relevance to current health studies (i.e., areas with peer-reviewed, epidemiologic analysis), (2) the existence of four or more monitoring sites located within the area boundaries, and (3) the presence of several diverse SO₂ sources within a given focus area boundary.

Maps of individual focus areas indicating 99th percentile 5-minute hourly max concentrations at monitoring sites and emissions from large point sources and their locations are presented in [Figures 2-13](#) through [2-18](#). As shown by the maps, up to 12 SO₂ monitoring sites are located in individual focus areas. Monitoring sites in each focus area are located at various distances from SO₂ sources. Due to the relatively short atmospheric lifetime of SO₂, monitoring sites adjacent to large point sources (e.g., electric generating units, industrial sources, copper smelting facilities, integrated iron and steel mills, shipping ports) are expected to detect higher SO₂ concentrations than those farther downwind. However, other variables, particularly stack height and wind speed and direction, influence concentrations observed near sources. For example, Sites C and E in Cleveland are both adjacent to large sources, but Site C has a much lower concentration than Site E despite the source near Site C emitting much more SO₂ than the source near Site E.



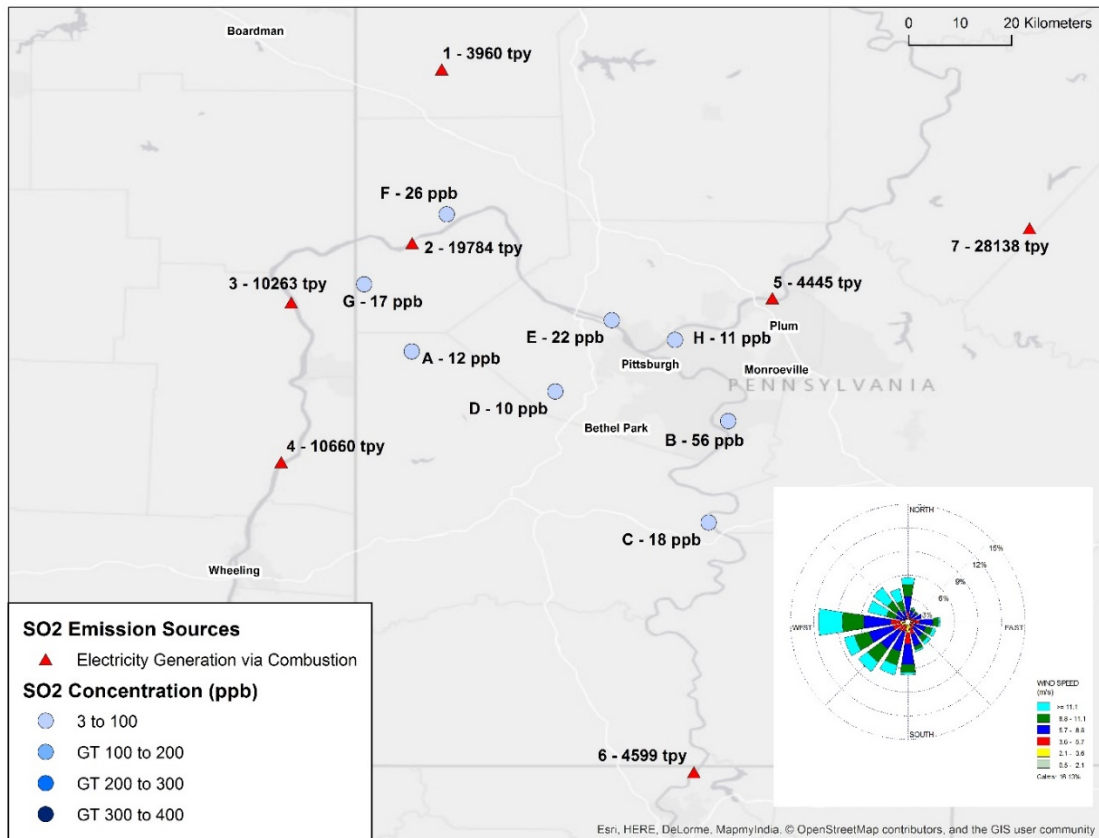
Note: Blue circles denote monitoring sites included in the U.S. Air Quality Monitoring System. Triangles denote sources emitting 2,000 tpy or more according to the 2014 U.S. National Emissions Inventory. The inset, lower right, displays a wind rose of average wind speed and direction for data acquired at Cleveland Hopkins International Airport over the 3-yr period 2013–2015.

Figure 2-13 Map of the Cleveland, OH focus area showing emissions from large sources and the 99th percentile 5-minute hourly max concentration at ambient monitors during 2013–2015.

Table 2-7 Largest SO₂ emissions sources, Cleveland, OH (as noted in Figure 2-13).

Map Code	Facility Name	Facility Type	SO ₂ Emissions (tpy)
1	Avon Lake Power Plant (0247030013)	EGU—Combustion	34,932
2	The Medical Center Company (1318003059)	Steam/Heating Facility	2,404
3	Cleveland Electric Illuminating Company, Eastlake Plant (0243160009)	EGU—Combustion	22,126

EGU = electric power generating uni; SO₂ = sulfur dioxide.



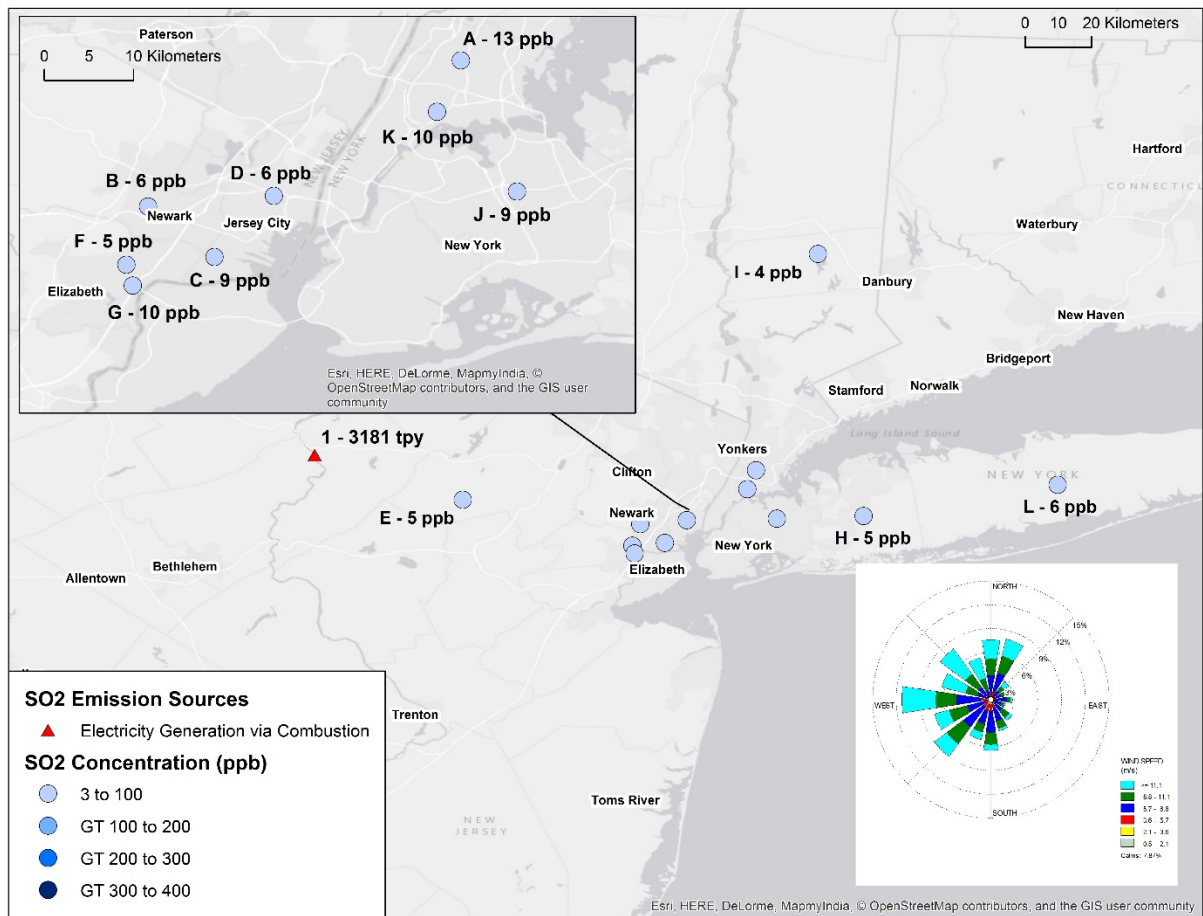
Note: Blue circles denote monitoring sites included in the U.S. Air Quality Monitoring System. Triangles denote sources emitting 2,000 tpy or more according to the 2014 U.S. National Emissions Inventory. The inset, lower right, displays a wind rose of average wind speed and direction for data acquired at Pittsburgh International Airport over the 3-yr period 2013–2015.

Figure 2-14 Map of the Pittsburgh, PA focus area showing emissions from large sources and the 99th percentile 5-minute hourly max concentration at ambient monitors during 2013–2015.

Table 2-8 Largest SO₂ emissions sources, Pittsburgh, PA (as noted in [Figure 2-14](#)).

Map Code	Facility Name	Facility Type	SO ₂ Emissions (tpy)
1	NRG Power Midwest LP/New Castle Power Plant	EGU—Combustion	3,960
2	FirstEnergy Gen LLC/Bruce Mansfield Plant	EGU—Combustion	19,784
3	W. H. Sammis Plant	EGU—Combustion	10,263
4	Cardinal Power Plant (Cardinal Operating Company)	EGU—Combustion	10,660
5	NRG Midwest LP/Cheswick	EGU—Combustion	4,445
6	Monongahela Power Company—Fort Martin Power	EGU—Combustion	4,599
7	Genon NE Mgmt Company/Keystone Station	EGU—Combustion	28,138

EGU = electric generating units; SO₂ = sulfur dioxide.



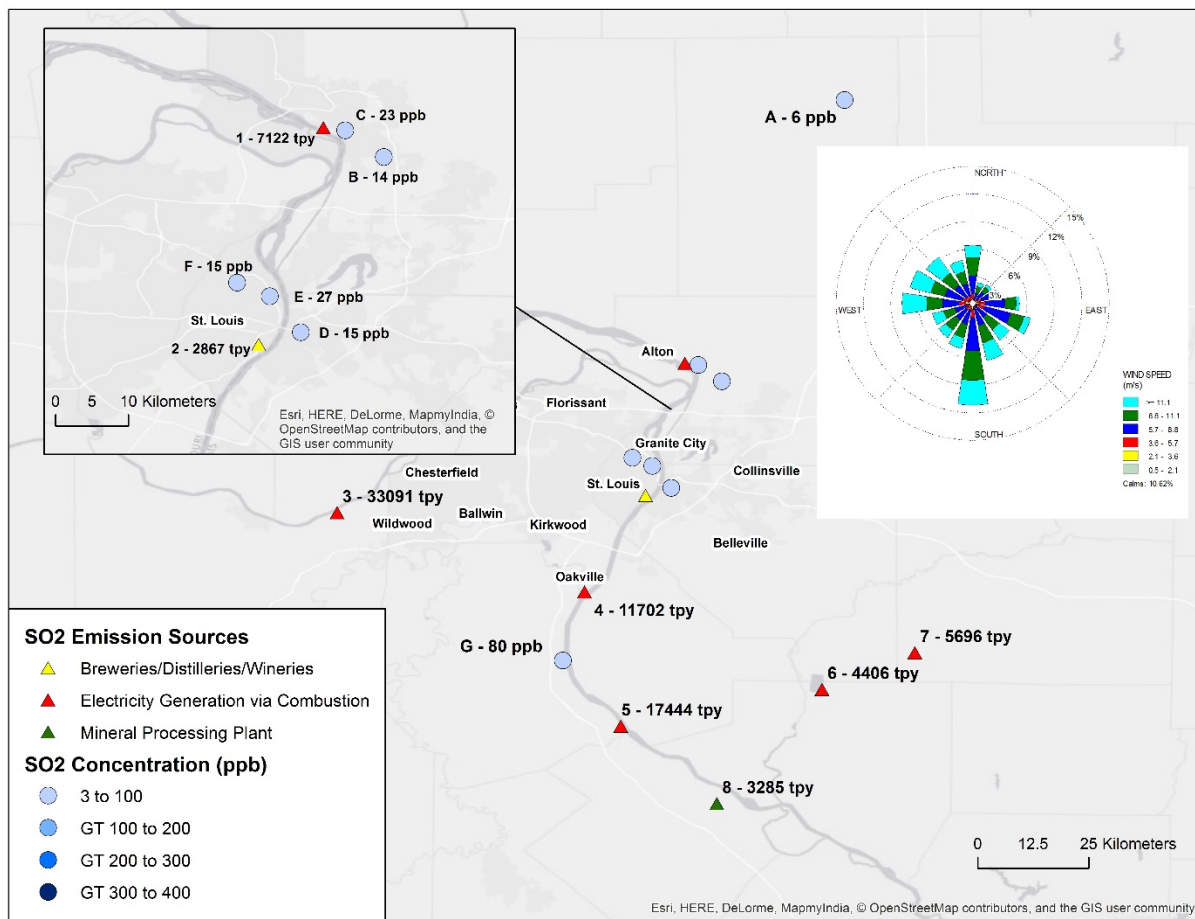
Note: Blue circles denote monitoring sites included in the U.S. Air Quality Monitoring System. Triangles denote sources emitting 2,000 tpy or more according to the 2014 U.S. National Emissions Inventory. The inset, lower right, displays a wind rose of average wind speed and direction for data acquired at Newark International Airport over the 3-yr period 2013–2015.

Figure 2-15 Map of the New York City, NY focus area showing emissions from large sources and the 99th percentile 5-minute hourly max concentration at ambient monitors during 2013–2015.

Table 2-9 Largest SO₂ emissions source, New York, NY (as noted in Figure 2-15).

Map Code	Facility Name	Facility Type	SO ₂ Emissions (tpy)
1	NRG REMA LLC/Portland Generating Station	EGU—Combustion	3,181

EGU = electric power generating unit; SO₂ = sulfur dioxide.



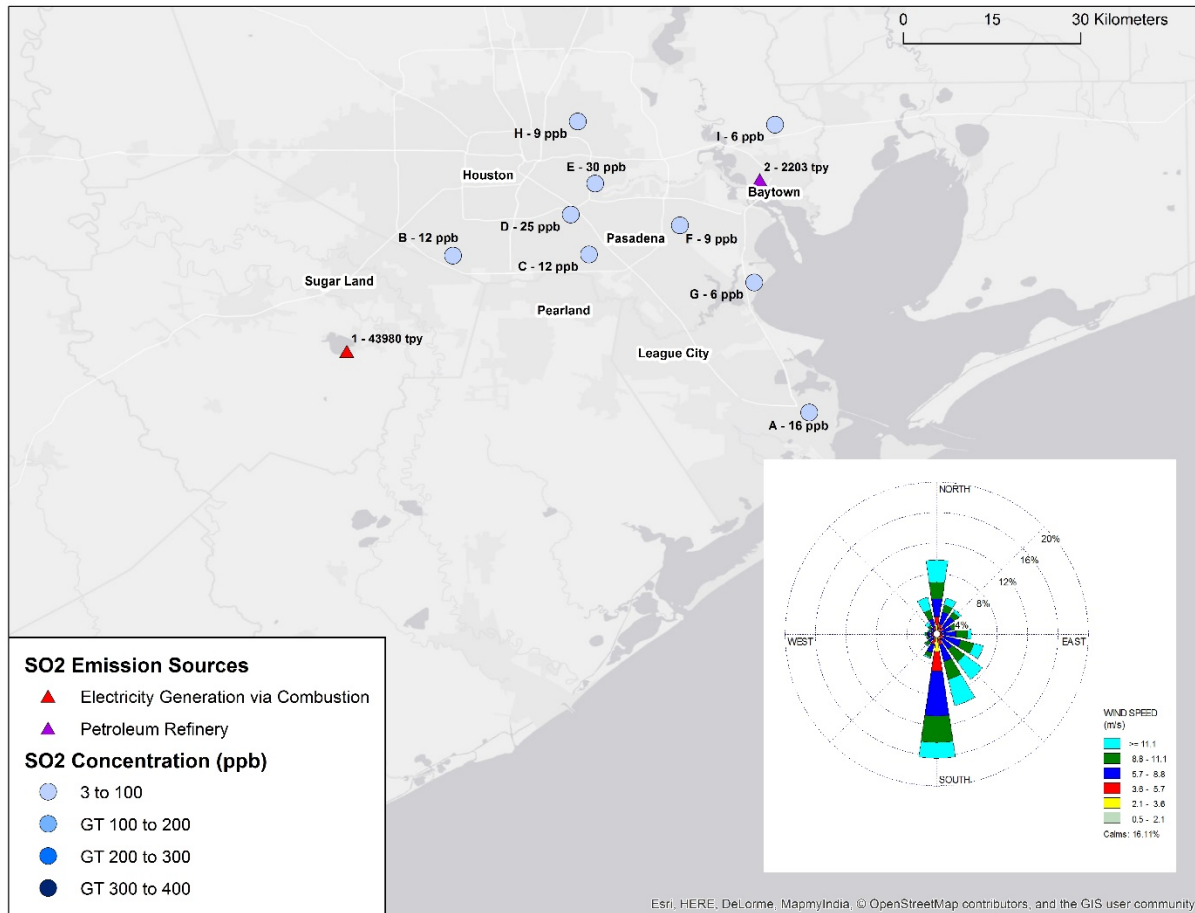
Note: Blue circles denote monitoring sites included in the U.S. Air Quality Monitoring System. Triangles denote sources emitting 2,000 tpy or more according to the 2014 U.S. National Emissions Inventory. The inset, upper right, displays a wind rose of average wind speed and direction for data acquired at Lambert—St. Louis International Airport over the 3-yr period 2013–2015.

Figure 2-16 Map of the St. Louis, MO-IL focus area showing emissions from large sources and the 99th percentile 5-minute hourly max concentration at ambient monitors during 2013–2015.

Table 2-10 Largest SO₂ emissions sources, St. Louis, MO-IL (as noted in [Figure 2-16](#)).

Map Code	Facility Name	Facility Type	SO ₂ Emissions (tpy)
1	Dynergy Midwest Generation LLC	EGU—Combustion	7,122
2	Anheuser-Busch Inc-St. Louis	Breweries/Distilleries/Wineries	2,867
3	Ameren Missouri-Labadie Plant	EGU—Combustion	33,091
4	Ameren Missouri-Meramec Plant	EGU—Combustion	11,702
5	Ameren Missouri-Rush Island Plant	EGU—Combustion	17,444
6	Dynergy Midwest Generation	EGU—Combustion	4,406
7	Prairie State Generating Station	EGU—Combustion	5,696
8	Mississippi Lime Company—Saint Genevieve	Mineral Processing Plant	3,285

EGU = electric generating units; SO₂ = sulfur dioxide.



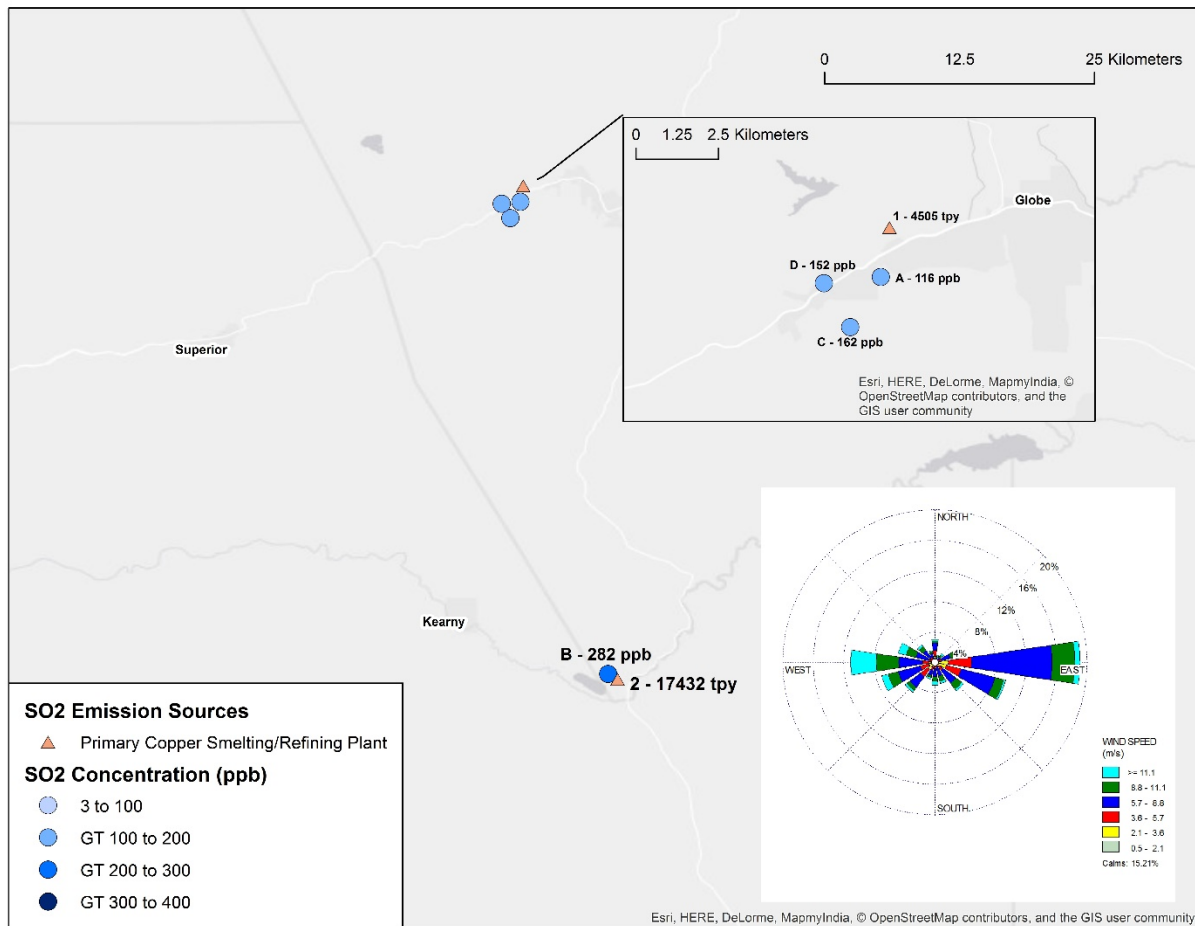
Note: Blue circles denote monitoring sites included in the U.S. Air Quality Monitoring System. Triangles denote sources emitting 2,000 tpy or more according to the 2014 U.S. National Emissions Inventory. The inset, lower right, displays a wind rose of average wind speed and direction for data acquired at George Bush Intercontinental Airport over the 3-yr period 2013–2015.

Figure 2-17 Map of the Houston, TX focus area showing emissions from large sources and the 99th percentile 5-minute hourly max concentration at ambient monitors during 2013–2015.

Table 2-11 Largest SO₂ emissions source, Houston, TX (as noted in [Figure 2-17](#)).

Map Code	Facility Name	Facility Type	SO ₂ Emissions (tpy)
1	WA Parish Electric Generating Station	EGU—Combustion	43,980
2	Baytown Refinery	Petroleum Refinery	2,203

EGU = electric generating units; SO₂ = sulfur dioxide.



Note: Blue circles denote monitoring sites included in the U.S. Air Quality Monitoring System. Yellow triangle denote sources emitting 2,000 tpy or more according to the 2011 U.S. National Emissions Inventory. The inset, lower right, displays a wind rose of average wind speed and direction for data acquired at the Phoenix Sky Harbor Intercontinental Airport over the 3-yr period 2013–2015.

Figure 2-18 Map of the Gila County, AZ focus area showing emissions from large sources and the 99th percentile 5-minute hourly max concentration at ambient monitors during 2013–2015.

Table 2-12 Largest SO₂ emissions sources, Gila County, AZ (as noted in [Figure 2-18](#)).

Map Code	Facility Name	Facility Type	SO ₂ Emissions (tpy)
1	Freeport McMoran Miami Smelter	Primary Copper Smelting/Refining Plant	4,505
2	Asarco, LLC—Hayden Smelter	Primary Copper Smelting/Refining Plant	17,432

SO₂ = sulfur dioxide.

[Table 2-13](#) provides the distribution of 1-hour daily max SO₂ concentrations and monitor type (standard vs. trace level monitor) reported at individual AQS sites in the six focus areas. Concentrations reported at these sites were similar to nationwide SO₂ concentrations discussed earlier in this section ([Section 2.5.2.1](#)). For all but one individual monitoring site, median concentrations were below 15 ppb. The one exception was the monitoring site in the Gila County, AZ focus area, for which the median concentration was 39 ppb. This particular monitoring site (Site B) is located within 1 km of a copper smelting plant with markedly high annual SO₂ emissions [greater than 17,000 tpy SO₂; ([U.S. EPA, 2016b](#))].

Table 2-13 1-h daily max sulfur dioxide concentration distribution by Air Quality System monitoring site in six focus areas, 2013–2015.

Site Label	AQS Monitoring Site ID	N of Obs	Mean	Min	10%	25%	50%	75%	90%	99%	Max	Monitor Type
Cleveland-Elyria-Mentor, OH												
A	390350065	709	6.4	0.0	0.0	1.0	3.0	7.0	13.2	55.9	125.0	Standard
B	390350060	887	11.5	0.0	0.0	2.0	6.0	16.0	32.0	62.1	92.0	Standard
C	390850003	758	7.6	0.0	2.0	3.0	6.0	10.0	15.0	37.4	95.0	Standard
D	390350038	786	14.0	0.0	1.0	4.0	10.0	20.0	32.5	61.3	105.0	Standard
E	390850007	901	11.2	0.0	2.0	3.0	6.0	11.0	22.0	117.0	201.0	Standard
F	390350045	630	3.9	0.0	0.0	0.0	2.0	5.0	9.0	30.0	51.0	Standard

Table 2-13 (Continued): 1 h daily max sulfur dioxide concentration distribution by Air Quality System monitoring site in six focus areas, 2013–2015.^a

Site Label	AQS Monitoring Site ID	N of Obs	Mean	Min	10%	25%	50%	75%	90%	99%	Max	Monitor Type
Pittsburgh, PA												
A	421255001	1,020	3.6	0.0	0.0	0.0	3.0	5.0	9.0	17.0	53.0	Standard
B	420030064	1,076	16.6	0.0	2.0	4.0	11.0	21.0	39.5	90.8	244.0	Standard
C	421250005	1,044	6.1	0.0	2.0	3.0	4.0	7.0	11.0	33.6	61.0	Standard
D	420030067	1,069	3.4	0.0	0.0	1.0	2.0	4.0	7.0	19.0	55.0	Standard
E	420030002	1,090	5.9	0.0	1.0	2.0	4.0	7.0	12.0	41.0	75.0	Standard
F	420070005	1,014	7.0	0.0	0.0	1.0	4.0	10.0	17.0	40.0	80.0	Standard
G	420070002	1,028	5.6	0.0	1.0	2.0	4.0	8.0	12.0	24.7	45.0	Standard
H	420030008	706	4.0	0.0	0.9	1.7	2.8	4.5	7.7	20.2	100.3	Trace
New York-Northern New Jersey-Long Island, NY-NJ-PA												
A	360050133	1,089	4.0	0.2	0.9	1.5	2.8	5.3	8.9	16.5	26.5	Standard
B	340130003	1,089	1.8	0.0	0.3	0.6	1.3	2.4	3.9	7.8	13.0	Trace
C	340170006	725	1.4	0.0	0.0	0.0	1.0	2.0	4.0	9.0	11.0	Standard
D	340171002	1,090	1.4	0.0	0.0	0.0	1.0	2.0	4.0	8.0	11.0	Standard
E	340273001	1,065	1.4	0.0	0.0	0.0	1.0	2.0	3.0	9.0	20.0	Standard
F	340390003	1,089	1.3	0.0	0.0	0.0	1.0	2.0	3.0	6.0	12.0	Standard
G	340390004	1,081	2.3	0.0	0.0	1.0	1.0	3.0	5.0	13.2	109.0	Standard
H	360590005	1,001	2.0	0.2	0.8	1.1	1.5	2.3	3.6	8.3	14.6	Standard
I	360790005	1,083	1.2	0.1	0.4	0.6	0.8	1.3	2.2	5.8	10.3	Standard
J	360810124	1,086	2.5	0.0	0.5	0.9	1.7	3.3	5.4	11.0	18.5	Trace
K	360050110	1,077	3.1	-0.2	0.8	1.2	2.2	4.1	6.7	14.2	32.1	Standard
L	361030009	938	1.6	-0.6	0.1	0.4	1.0	2.3	4.0	8.7	15.8	Standard
St. Louis, MO-IL												
A	171170002	646	2.2	0.0	0.8	1.0	2.0	3.0	4.0	8.5	21.0	Standard
B	171191010	1,023	4.1	0.0	0.9	1.3	3.0	5.0	9.0	18.0	40.0	Standard

Table 2-13 (Continued): 1 h daily max sulfur dioxide concentration distribution by Air Quality System monitoring site in six focus areas, 2013–2015.^a

Site Label	AQS Monitoring Site ID	N of Obs	Mean	Min	10%	25%	50%	75%	90%	99%	Max	Monitor Type
C	171193007	1,041	5.6	0.0	1.0	2.0	4.0	7.0	11.6	24.4	42.0	Standard
D	171630010	1,018	4.7	0.0	1.0	2.0	3.6	6.0	10.0	20.8	30.0	Standard
E	295100085	921	7.2	0.0	1.3	2.4	4.2	9.1	16.5	40.2	51.4	Trace
F	295100086	1,077	4.5	0.5	1.2	1.8	3.3	5.6	9.5	19.6	31.8	Standard
G	290990027	1,089	11.6	0.3	1.1	2.2	4.2	8.8	36.3	94.5	252.7	Standard
Houston-Sugar Land-Baytown, TX												
A	481670005	736	3.6	0.3	1.0	1.5	2.4	3.8	6.8	26.5	50.6	Standard
B	482010051	214	3.1	0.0	0.7	1.0	1.9	3.4	6.1	22.2	44.4	Standard
C	482010062	160	3.7	0.4	1.0	1.7	2.4	4.4	7.9	18.0	19.3	Standard
D	482010416	313	5.5	0.3	0.9	1.6	3.4	6.9	12.1	33.6	54.0	Standard
E	482011035	71	4.9	0.3	0.5	1.5	2.4	5.4	13.1	25.9	29.8	Standard
F	482011039	590	2.2	0.0	0.2	0.7	1.6	2.9	5.2	11.0	16.0	Trace
G	482011050	885	1.9	0.2	0.5	0.7	1.4	2.4	3.8	9.0	16.4	Standard
H	482010046	15	3.5	1.8	1.9	2.3	2.8	3.2	4.7	12.0	13.1	Standard
I	482011017	415	1.5	0.0	0.4	0.6	1.0	1.9	3.3	8.3	10.6	Standard
Gila County, AZ												
A	40070009	1,080	24.9	0.0	2.0	3.0	12.0	34.3	64.0	153.2	259.0	Standard
B	40071001	889	50.8	0.0	1.0	13.0	39.0	71.0	114.2	247.2	368.0	Trace
C	40070011	739	28.5	0.0	1.0	2.0	9.0	36.0	84.0	204.9	380.0	Trace
D	40070012	630	31.3	0.0	1.0	2.0	8.0	39.8	95.0	230.7	324.0	Trace

AQS = Air Quality System; ID = identification; max = maximum; mean = arithmetic average; min = minimum; N = population number; Obs = observations.

AQS accepts, as valid, reported concentrations that are \pm LDL. Data analyzed in this table include negative values in this range.

More substantial site-to-site differences were observed in the 99th percentile of SO₂ concentrations. Across these monitoring sites, 1-hour daily max 99th percentile concentrations ranged from 5.8 to 247.2 ppb, with the majority of sites exhibiting 99th percentile concentrations at or below 40 ppb. Relatively high 99th percentile concentrations were reported at monitoring sites within 5 km of a large SO₂ point source, particularly in Gila County, AZ near copper smelters. Relatively high 99th percentile concentrations were also observed in the Cleveland, OH and Pittsburgh, PA focus areas. These data were in agreement with previous studies, which generally observed higher urban SO₂ concentrations near local industrial/combustion sources related to oil-burning units, smelters, and EGUs ([Brand et al., 2016](#); [Clougherty et al., 2013](#); [Wheeler et al., 2008](#)).

Over the past decade, the number of AQS monitoring sites reporting 5-minute SO₂ concentrations has substantially increased. At the time of the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), a total of 98 monitoring sites periodically reported 5-minute hourly max concentrations. To date, approximately 380 sites report 5-minute data, including urban sites within focus areas, sites near city centers, and sites near SO₂ sources ([Figure 2-10](#) in [Section 2.4.3](#)).

Similar analyses of 5-minute hourly max concentrations were performed on more recent data reported at individual monitoring sites in the six focus areas. [Table 2-14](#) shows the range in 5-minute hourly max SO₂ concentrations reported at individual monitors, within the six focus areas in the 2013–2015 time frame. Median 5-minute hourly max concentrations are below 5 ppb, while maximum concentrations range from 15 to 1,241 ppb.

Table 2-14 5-minute hourly max sulfur dioxide concentrations by Air Quality System monitoring sites in select focus areas, 2013–2015.

Site Label	AQS Monitoring Site ID	N of Obs	Mean	Min	10%	25%	50%	75%	90%	99%	Max	Monitor Type
Cleveland-Elyria-Mentor, OH												
A	390350065	16,201	3.7	0.0	0.0	0.0	2.0	5.0	8.0	27.0	397.0	Standard
B	390350060	18,585	4.9	0.0	0.0	0.0	1.0	4.0	13.0	53.0	159.0	Standard
C	390850003	15,966	3.6	0.0	0.0	1.0	2.0	5.0	8.0	26.0	241.0	Standard
D	390350038	17,321	6.0	0.0	0.0	0.0	2.0	7.0	16.0	49.0	180.0	Standard

Table 2-14 (Continued) 5 minute hourly max sulfur dioxide concentrations by Air Quality System monitoring sites in select focus areas, 2013–2015

Site Label	AQS Monitoring Site ID	N of Obs	Mean	Min	10%	25%	50%	75%	90%	99%	Max	Monitor Type
E	390850007	19,297	5.6	0.0	0.0	1.0	3.0	5.0	9.0	69.0	428.0	Standard
F	390350045	13,720	1.5	0.0	0.0	0.0	0.0	2.0	4.0	15.0	131.0	Standard
Pittsburgh, PA												
A	421255001	24,367	1.5	0.0	0.0	0.0	0.0	2.0	4.0	12.0	73.0	Standard
B	420030064	25,602	6.1	0.0	0.0	1.0	2.0	7.0	16.0	56.0	493.0	Standard
C	421250005	24,930	3.3	0.0	1.0	1.0	2.0	4.0	6.0	18.0	137.0	Standard
D	420030067	25,480	1.4	0.0	0.0	0.0	1.0	2.0	4.0	10.0	89.0	Standard
E	420030002	26,001	2.4	0.0	0.0	0.0	1.0	3.0	6.0	22.0	112.0	Standard
F	420070005	24,264	3.1	0.0	0.0	0.0	1.0	3.0	8.0	26.0	155.0	Standard
G	420070002	24,572	2.2	0.0	0.0	0.0	1.0	3.0	6.0	17.0	64.0	Standard
H	420030008	16,095	1.7	-0.2	0.1	0.4	1.0	2.2	3.8	10.7	158.3	Trace
New York-Northern New Jersey-Long Island, NY-NJ-PA												
A	360050133	25,699	2.5	0.0	0.4	0.8	1.5	3.2	5.8	13.0	32.3	Standard
B	340130003	25,928	0.9	0.0	0.1	0.2	0.5	1.2	2.3	5.7	23.1	Trace
C	340170006	17,200	0.8	0.0	0.0	0.0	0.0	1.0	3.0	9.0	29.0	Standard
D	340171002	25,826	1.0	0.0	0.0	0.0	1.0	1.0	2.0	6.0	34.0	Standard
E	340273001	24,451	1.2	0.0	0.0	1.0	1.0	1.0	2.0	5.0	58.0	Standard
F	340390003	25,887	1.2	0.0	0.0	0.0	1.0	2.0	3.0	5.0	47.0	Standard
G	340390004	25,748	1.4	0.0	0.0	0.0	1.0	2.0	3.0	10.0	317.0	Standard
H	360590005	23,683	1.4	0.1	0.6	0.8	1.1	1.6	2.3	5.3	21.5	Standard
I	360790005	25,630	0.9	0.0	0.4	0.5	0.7	1.0	1.4	3.6	16.1	Standard
J	360810124	25,557	1.5	-0.1	0.1	0.3	0.8	1.9	3.8	9.0	26.8	Trace
K	360050110	25,333	2.0	-1.2	0.3	0.7	1.4	2.6	4.3	10.0	46.6	Standard
L	361030009	22,128	1.2	-0.7	0.1	0.4	0.9	1.7	2.9	6.4	30.5	Standard
St. Louis, MO-IL												
A	171170002	14,260	1.5	0.0	0.5	1.0	1.2	2.0	2.7	6.0	56.0	Standard

Table 2-14 (Continued) 5 minute hourly max sulfur dioxide concentrations by Air Quality System monitoring sites in select focus areas, 2013–2015

Site Label	AQS Monitoring Site ID	N of Obs	Mean	Min	10%	25%	50%	75%	90%	99%	Max	Monitor Type
B	171191010	22,801	1.7	0.0	0.0	0.0	0.9	2.0	4.0	15.0	240.0	Standard
C	171193007	23,684	2.7	0.0	0.0	0.8	1.3	3.0	6.0	24.0	94.0	Standard
D	171630010	22,691	1.9	0.0	0.0	0.0	1.0	2.0	4.2	15.0	87.4	Standard
E	295100085	20,653	3.3	0.0	0.6	1.2	2.0	3.3	6.3	26.5	93.7	Trace
F	295100086	25,720	2.4	0.2	0.8	1.1	1.5	2.5	4.5	15.2	53.0	Standard
G	290990027	26,002	5.7	0.2	0.5	0.9	2.1	3.6	8.0	80.4	657.1	Standard
Houston-Sugar Land-Baytown, TX												
A	481670005	16,307	1.9	0.0	0.4	0.6	1.1	2.1	3.6	15.8	84.9	Standard
B	482010051	4,523	1.1	0.0	0.2	0.3	0.6	1.2	2.3	10.3	65.9	Standard
C	482010062	3,399	1.6	0.0	0.3	0.5	1.0	1.8	3.1	12.5	33.4	Standard
D	482010416	6,982	2.4	0.0	0.3	0.6	1.0	2.3	5.2	24.1	90.9	Standard
E	482011035	1,482	2.4	0.0	0.3	0.5	1.0	2.3	4.4	26.3	75.8	Standard
F	482011039	12,547	0.9	0.0	0.0	0.0	0.5	1.1	2.2	6.8	25.7	Trace
G	482011050	19,894	1.0	0.0	0.3	0.4	0.6	1.1	2.1	5.7	21.3	Standard
H	482010046	313	1.8	0.0	0.3	0.5	1.5	2.6	3.3	7.2	15.2	Standard
I	482011017	8,728	0.7	0.0	0.0	0.2	0.4	0.8	1.5	5.0	25.3	Standard
Gila County, AZ												
A	40070009	25,732	9.2	0.0	1.0	1.0	3.1	4.5	21.6	115.5	461.0	Standard
B	40071001	20,222	19.6	0.0	0.0	1.0	2.0	10.6	55.0	252.2	1,241.2	Trace
C	40070011	16,630	9.1	0.0	0.0	0.0	1.0	3.0	22.0	142.1	694.0	Trace
D	40070012	14,156	7.6	0.0	0.0	1.0	1.0	2.0	11.0	148.0	993.0	Trace

AQS = Air Quality System; ID = identification; max = maximum; mean = arithmetic average; min = minimum; N = population number; Obs = observations.

AQS accepts as valid reported concentrations that are \pm LDL. Data analyzed in this table include negative values in this range.

To evaluate the extent of SO₂ spatial variability over urban geographical scales, concentration correlations between monitoring site pairs were calculated in each of the six focus areas. To estimate the degree to which concentrations at two different monitoring sites followed similar temporal trends, pairwise comparisons were evaluated using Pearson correlations. Across the six focus areas, Pearson correlations ranged from 0 to 1.0 for 24-hour avg data. Correlations close to 1 represent strong correspondence over time between pairwise monitoring site concentrations, while values close to 0 represent poor correspondence between concentrations. [Figures 2-19](#) and [2-20](#) show scatterplots of pairwise correlations of 24-hour avg and 5-minute hourly max SO₂ concentrations, respectively, versus distance between monitoring site pairs. The 24-hour avg concentrations are presented due to their frequent use in epidemiologic studies, while 5-minute hourly max concentrations are a metric of interest for short-duration exposures. Given the meandering nature of SO₂ plumes and potential for plume touchdown several kilometers from the stack ([Turner, 1970](#)), low correlation among monitoring sites would be expected in most cases for the 5-minute hourly max data.

Intersite pairwise comparisons in [Figure 2-19](#) suggest high spatial variability of the 24-hour avg SO₂ concentration time series, consistent with long-standing observations of the movement of emissions plumes. In every focus area except for New York (discussed below), low to moderate intersite pairwise correlations of 24-hour avg SO₂ concentration data were observed, with the majority of Pearson correlations below 0.6. Intersite pairwise correlations tended to decrease with distance. Even within relatively short distances (up to 15 km), most intersite pairwise correlations were low, reflecting the variable nature of ambient SO₂ across urban spatial scales, possibly due to short atmospheric residence time, variable meteorology, and the episodic nature of the emissions as discussed in [Section 2.2](#).

In comparison, 5-minute hourly max SO₂ concentrations had somewhat higher spatial variability across urban spatial scales ([Figure 2-20](#)). In most cases, intersite pairwise correlations of 5-minute hourly max concentrations are lower (less than 0.4) and decline more dramatically with distance than intersite pairwise correlations of 24-hour avg concentrations.

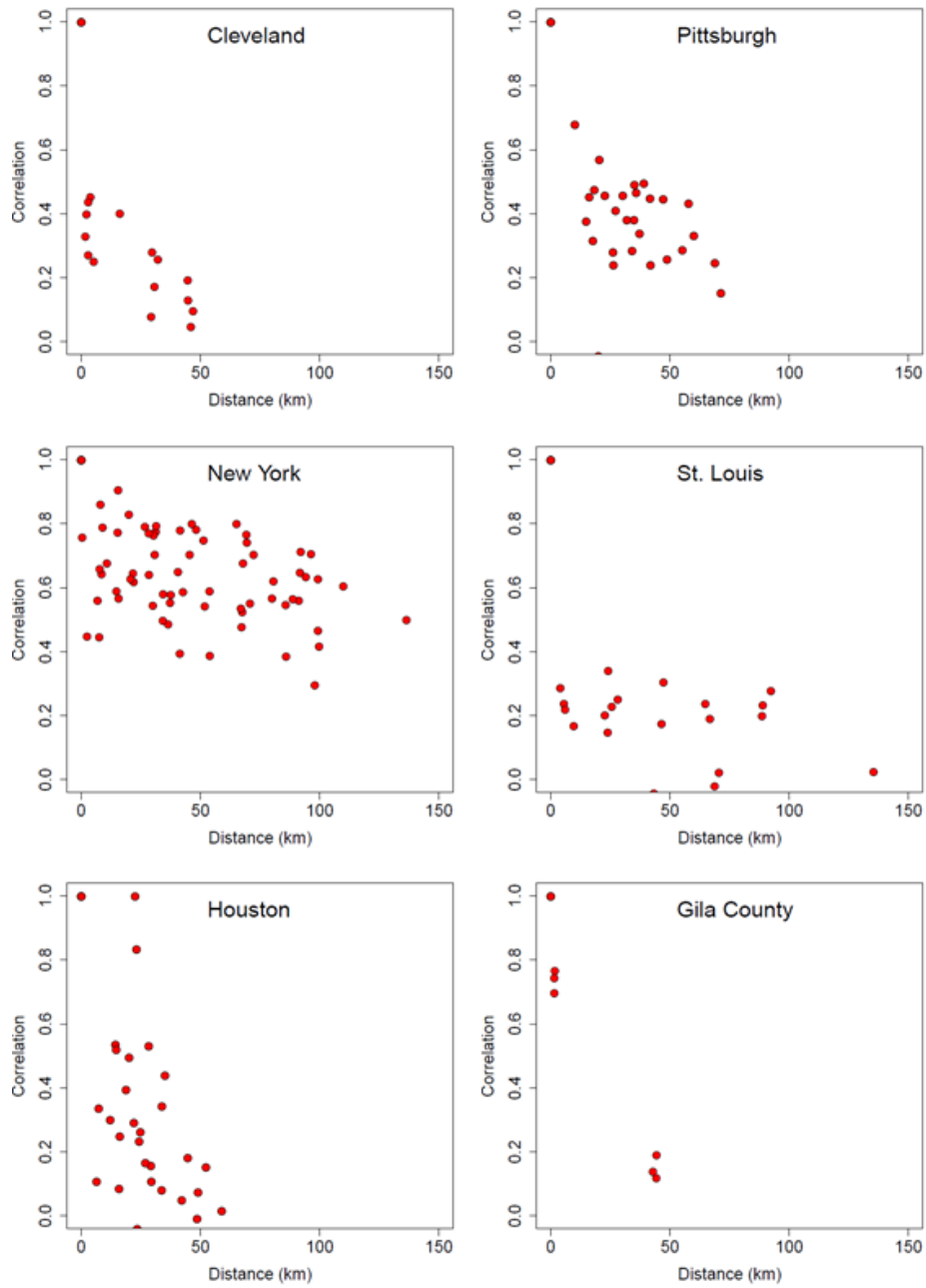


Figure 2-19 Pairwise correlations of 24-hour avg sulfur dioxide versus distance between monitoring site pairs in six focus areas, 2013–2015.

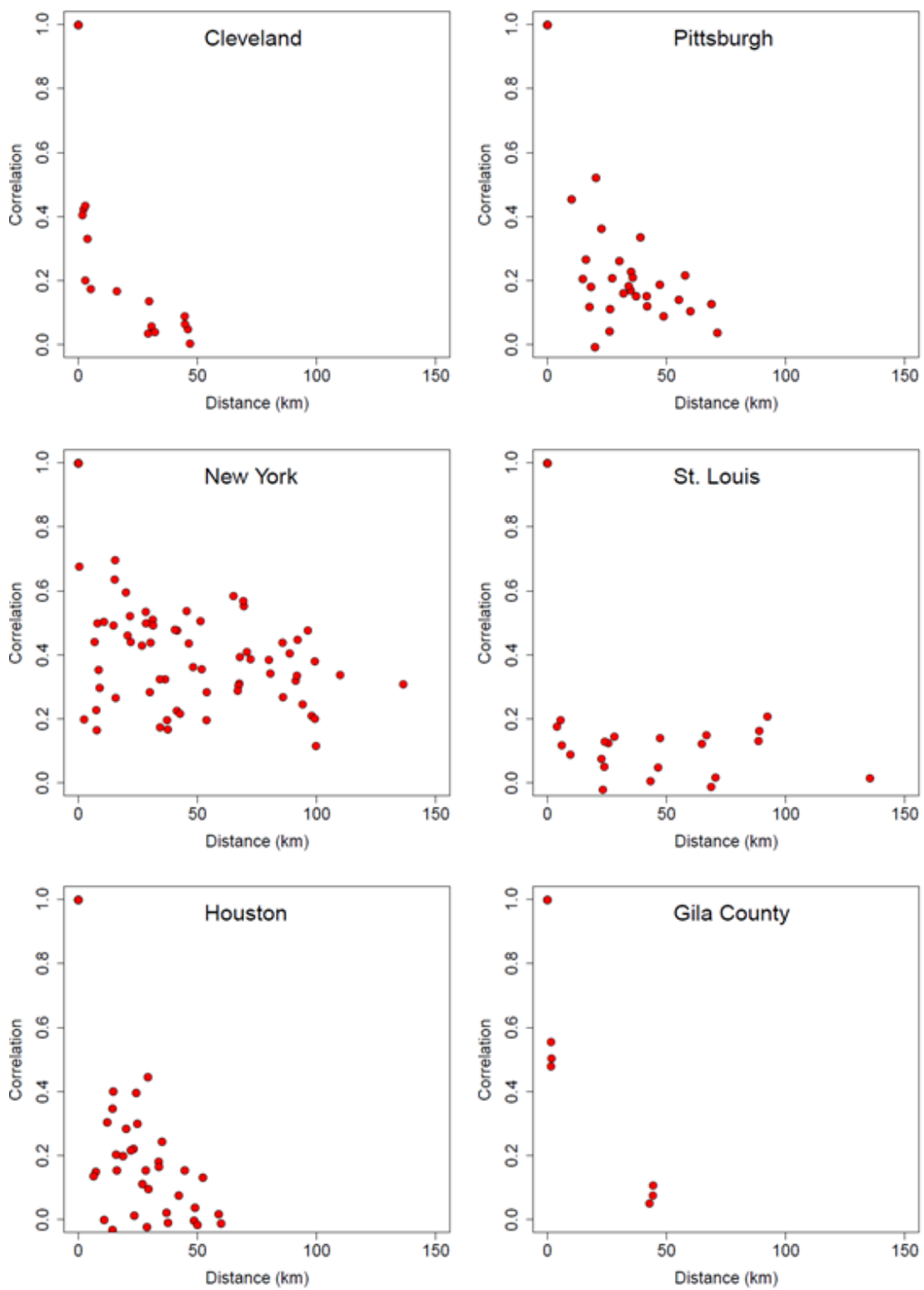


Figure 2-20 Pairwise correlations of 5-minute hourly max sulfur dioxide versus distance between monitoring sites in six focus areas, 2013–2015.

While spatial variability is evident to some degree in all urban areas, the extent of this variability is location dependent. For example, pairwise correlations in Cleveland, OH and St Louis, MO indicate strong SO₂ spatial heterogeneity. In comparison, pairwise correlations in New York City, NY are generally high and uniform across more than 100 km despite sometimes large distances between monitoring sites. Stronger pairwise correlations in New York City, NY may be related to similar temporal source patterns, given that the focus area's smaller power plants (<2,000 tpy SO₂ emissions), including gas-coal cogeneration facilities in Brooklyn, NY and Sayreville, NJ; an oil-burning facility in Queens, NY; a coal-fired power plant in Jersey City, NJ; and numerous homes using oil-burning heat likely have similar periods of high operation across the metropolitan area. This is analogous to observations about similarities in traffic patterns across large distances that promote higher correlation despite distance between the sources ([Sarnat et al., 2010](#)). Conversely, high spatial variations in Cleveland, OH and St. Louis, MO may be explained by the presence of a limited number of sources (>2,000 tpy) located at unevenly distributed sites across the metropolitan area.

In summary, SO₂ concentrations vary substantially across urban spatial scales as evidenced by poor to moderate intersite pairwise correlations observed in SO₂ data in six focus areas. Spatial heterogeneity in urban-scale SO₂ concentrations and their temporal patterns may be explained by the presence of multiple, unevenly distributed SO₂ sources, meteorological factors that lead to varying degrees of SO₂ dilution, or removal through cloud/fog chemistry and deposition. Additionally, in this analysis, metrics representing maximum SO₂ concentrations generally exhibited more spatial heterogeneity than 24-hour avg metrics.

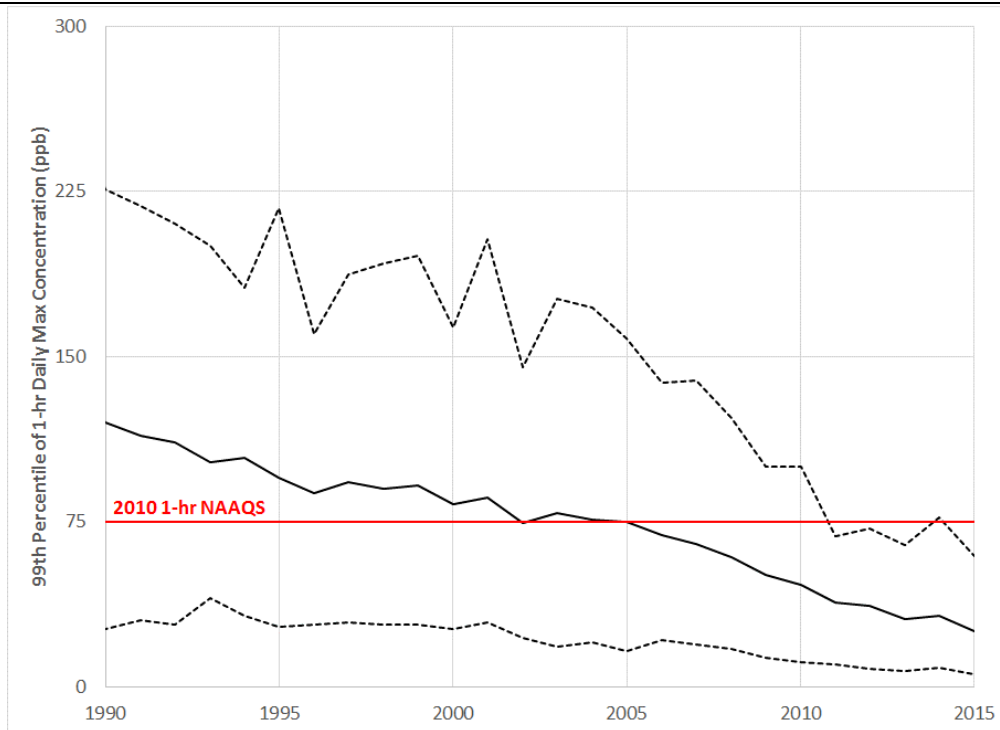
2.5.3 Temporal Variability

Temporal variations in outdoor SO₂ concentrations affect the magnitude, duration, and frequency in which humans are exposed to SO₂. In this section, different types of temporal trends are discussed, spanning long-term temporal trends on an annual basis to short-term trends on a subhourly basis.

2.5.3.1 Long-Term Trends

Trends in SO₂ concentrations reported at AQS monitoring sites across the U.S. from 1980 to 2015 are shown in [Figure 2-21](#) for the annual 99th percentile of the 1-hour daily max SO₂ concentration. Information on SO₂ concentration trends at individual, local air

monitoring sites can be found at <https://www.epa.gov/air-trends/sulfur-dioxide-trends> (U.S. EPA, 2012b).



Note: The solid line shows the mean concentrations and the upper and lower dashed lines represent the 10th and 90th percentile concentrations, respectively. The red line indicates the current NAAQS for sulfur oxides. NAAQS = National Ambient Air Quality Standards.

Source: <https://www.epa.gov/air-trends/sulfur-dioxide-trends>.

Figure 2-21 National sulfur dioxide air quality trend, based on the 99th percentile of the 1-hour daily max concentration for 163 sites, 1980–2015. A 76% decrease in the national average was observed from 1990–2015.

The steady decline in SO₂ concentrations over the past 25 years is largely attributed to emissions reductions at EGUs due to the Acid Rain and NO_x Budget Trading programs, CSAPR and other national interstate transport rules that have been implemented under the Clean Air Act Amendments of 1990 (USC Title 42 Chapter 85). Reductions in SO₂ emissions commenced in 1996 and continued into the 2000s, resulting in dramatic decreases in total, nationwide SO₂ emissions and concentrations (Figure 2-5). The NO_x Budget Trading Program and national interstate transport rules led to further reductions in

SO₂ emissions. From 1990–2014, the annual 99th percentile average of 1-hour daily max SO₂ concentration has decreased by 76% nationally.

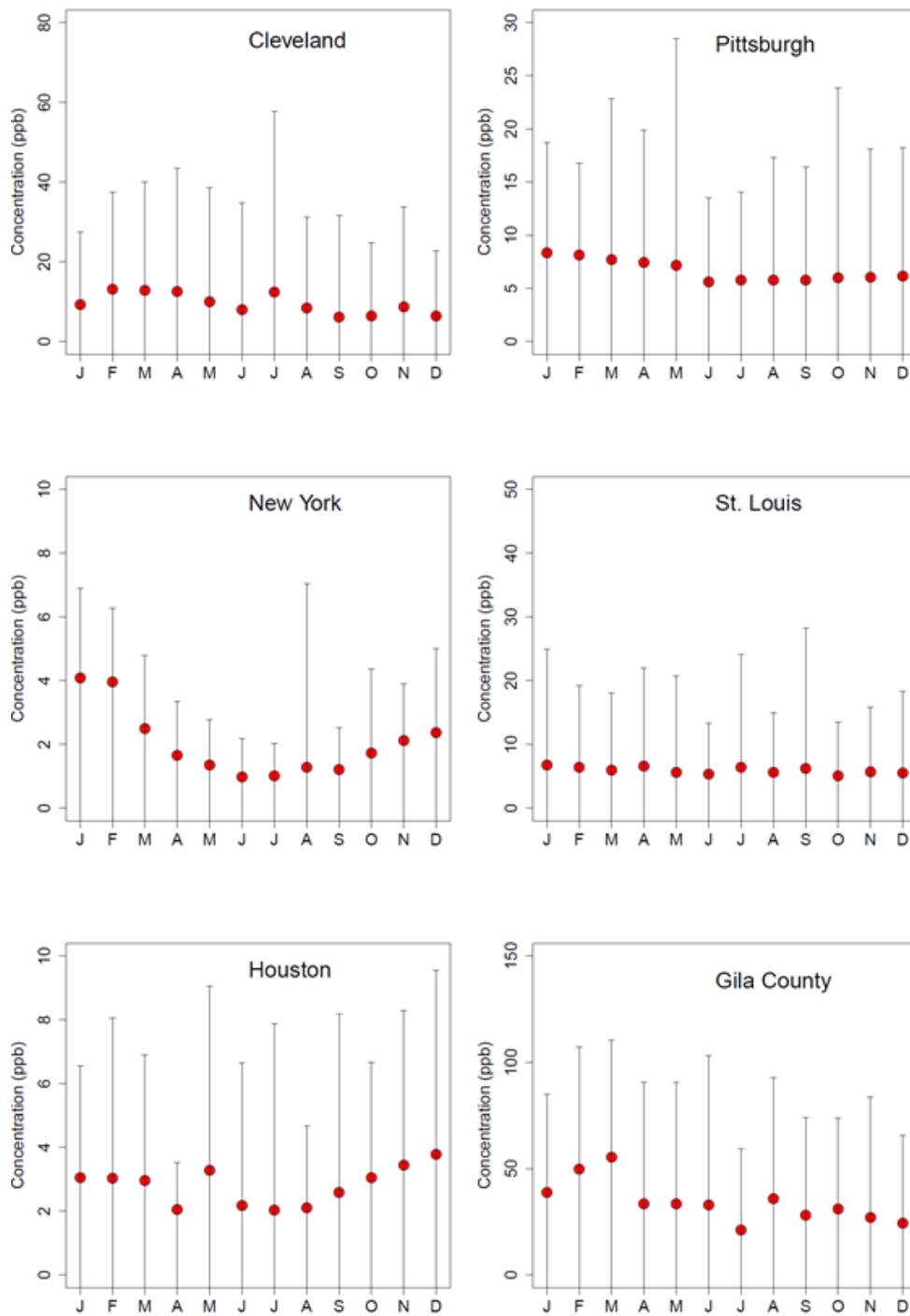
Substantial declines in SO₂ concentration over the past decades have also been observed on regional scales. [Blanchard et al. \(2013\)](#) reported an average decline of 7.6% per year ($\pm 1.6\%$) in SO₂ emissions from 1999–2010 across four southeastern U.S. states (Alabama, Florida, Georgia, Mississippi), primarily due to reductions in power plant emissions, which account for approximately 75% of total SO₂ emissions in the southeastern U.S. region. This decline corresponded to large reductions in annual SO₂ concentrations (between 5.1 and 9.7% per year) reported at monitoring sites across these four states.

2.5.3.2 Seasonal Trends

In the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), month-to-month trends for SO₂ were observed across a number of metropolitan areas, and these seasonal profiles varied by location. Some cities, such as Steubenville, OH and Phoenix, AZ showed clear wintertime maxima, while other urban areas (Philadelphia, PA; Los Angeles, CA; Riverside, CA) exhibited higher SO₂ concentrations during summer months. Differences in seasonal profiles were attributed to variations in source emissions, topography, and meteorological conditions among different areas.

Month-to-month variability based on more recent 1-hour daily max concentrations (2013–2015) is shown for the six focus areas introduced earlier in this chapter ([Section 2.5.2.2](#)). [Figure 2-22](#) displays the range of SO₂ concentrations reported at all monitoring sites within each focus area.

The data indicate that 1-hour daily max SO₂ concentrations vary across seasons, especially in the higher concentrations within monthly SO₂ concentration distributions. Among the five urban focus areas, mean concentrations (red circle) varied by no more than 10 ppb throughout the year, while the median concentration in the Gila County, AZ focus area varied by 30 ppb. Large variations across all focus areas are observed in the upper end (greater than 75th percentile) of SO₂ concentrations. Notably, mean monthly SO₂ concentrations were higher and more variable than median values, indicating that the distribution is skewed by high, infrequent observations.



Note: For every month, arithmetic mean concentrations are displayed as red circles. The whiskers represent the 95% confidence interval.

Figure 2-22 Sulfur dioxide month-to-month variability based on 1-hour daily max concentrations at Air Quality System sites in each focus area, 2013–2015.

Recent data further demonstrate that seasonal profiles vary by location. While each focus area exhibits some degree of seasonal variation, no consistent seasonal profile was observed across the focus areas. For example, springtime maxima in 1-hour daily max SO₂ are evident in Cleveland, OH and Gila County, AZ, corresponding to focus areas with the highest SO₂ concentrations. Alternatively, New York City, NY; Houston, TX; and Pittsburgh, PA show clear wintertime maxima.

Month-to-month variations in SO₂ concentrations are consistent with month-to-month emissions patterns ([Lee et al., 2011a](#)) and the atmospheric chemistry of SO₂. Summertime minima, observed in the New York City, NYC and Houston, TX focus areas, may correspond to enhanced oxidation of SO₂ to SO₄²⁻ by photochemically derived atmospheric oxidants that are more prevalent during the humid summer ([Khoder, 2002](#)). The difference in seasonality among these cities suggest that SO₂ can be substantially variable across local and regional scales.

2.5.3.3 Diel Variability

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) explored nationwide patterns in diel variability of SO₂ concentrations (i.e., variability of SO₂ concentrations across a 24-hour period), and found clear daytime maxima and nighttime minima, with larger day-night differences with increasing SO₂ concentrations. Daytime maxima were attributed to entrainment of SO₂ from elevated point sources (e.g., power plants and industrial sources) into the mixed boundary layer, which expands due to rising surface temperatures.

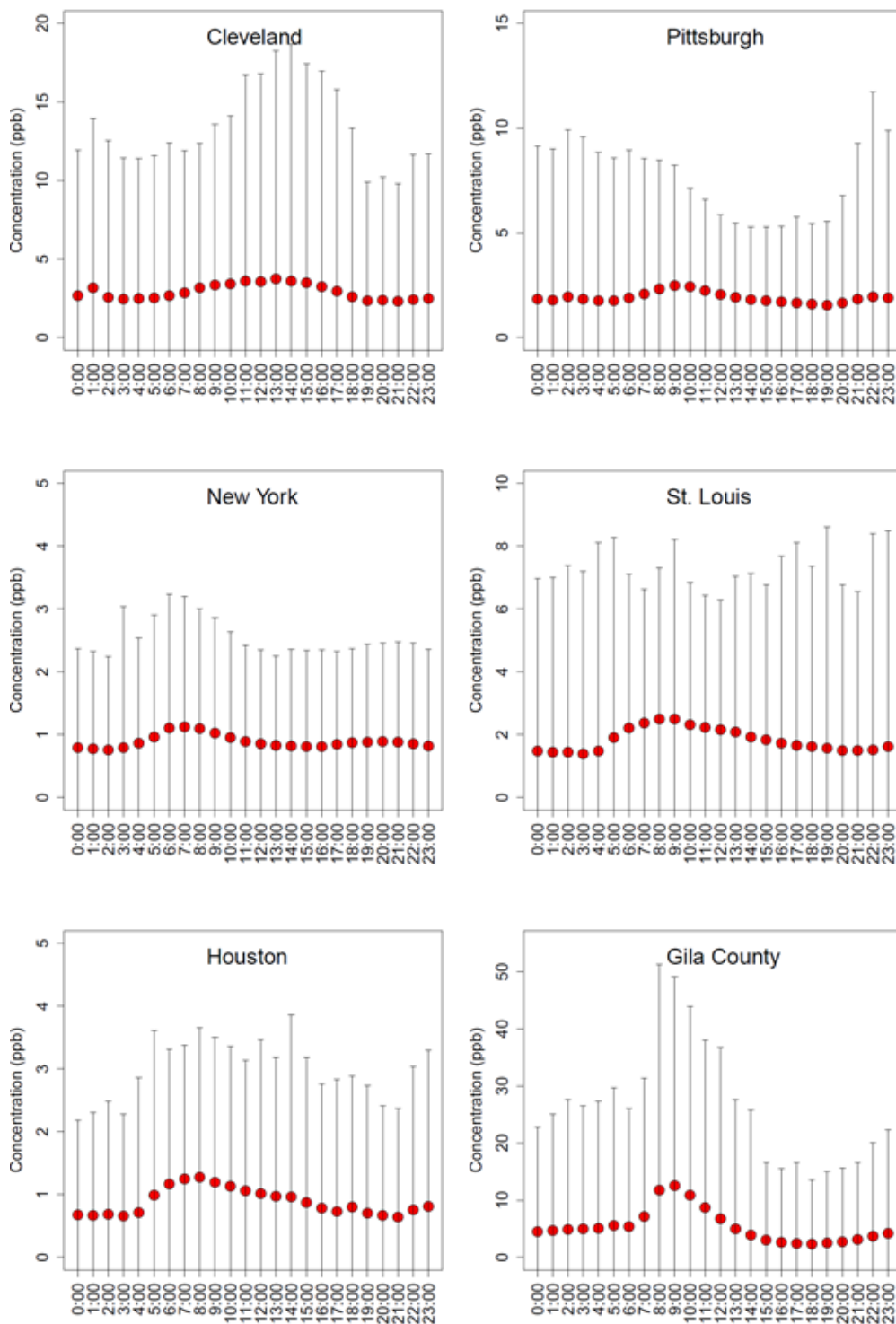
Diel patterns (i.e., patterns of change in ambient SO₂ concentrations over a 24-hour period) were investigated in the focus areas using 1-hour avg and 5-minute hourly max SO₂ data for the 2013–2015 time frame. [Figures 2-23](#) and [2-24](#) show variations in 1-hour avg and 5-minute hourly max SO₂ concentrations, respectively, in the six focus areas.

Consistent with the nationwide diel patterns reported in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), SO₂ concentrations in the six focus areas were generally low during nighttime and approach maxima values during daytime hours ([Figures 2-23](#) and [2-24](#)). In Pittsburgh, PA; New York City, NY; St. Louis, MO; Houston, TX; and Gila County, AZ, daytime maxima occurred during early morning hours (6:00 to 9:00 a.m. local standard time). In Cleveland, OH, SO₂ tended to peak later in the morning or in some cases early to mid-afternoon.

The timing and duration of daytime SO₂ peaks vary across the focus areas. These variations were likely a result of a combination of source emissions and meteorological parameters. The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) concluded that higher daytime SO₂

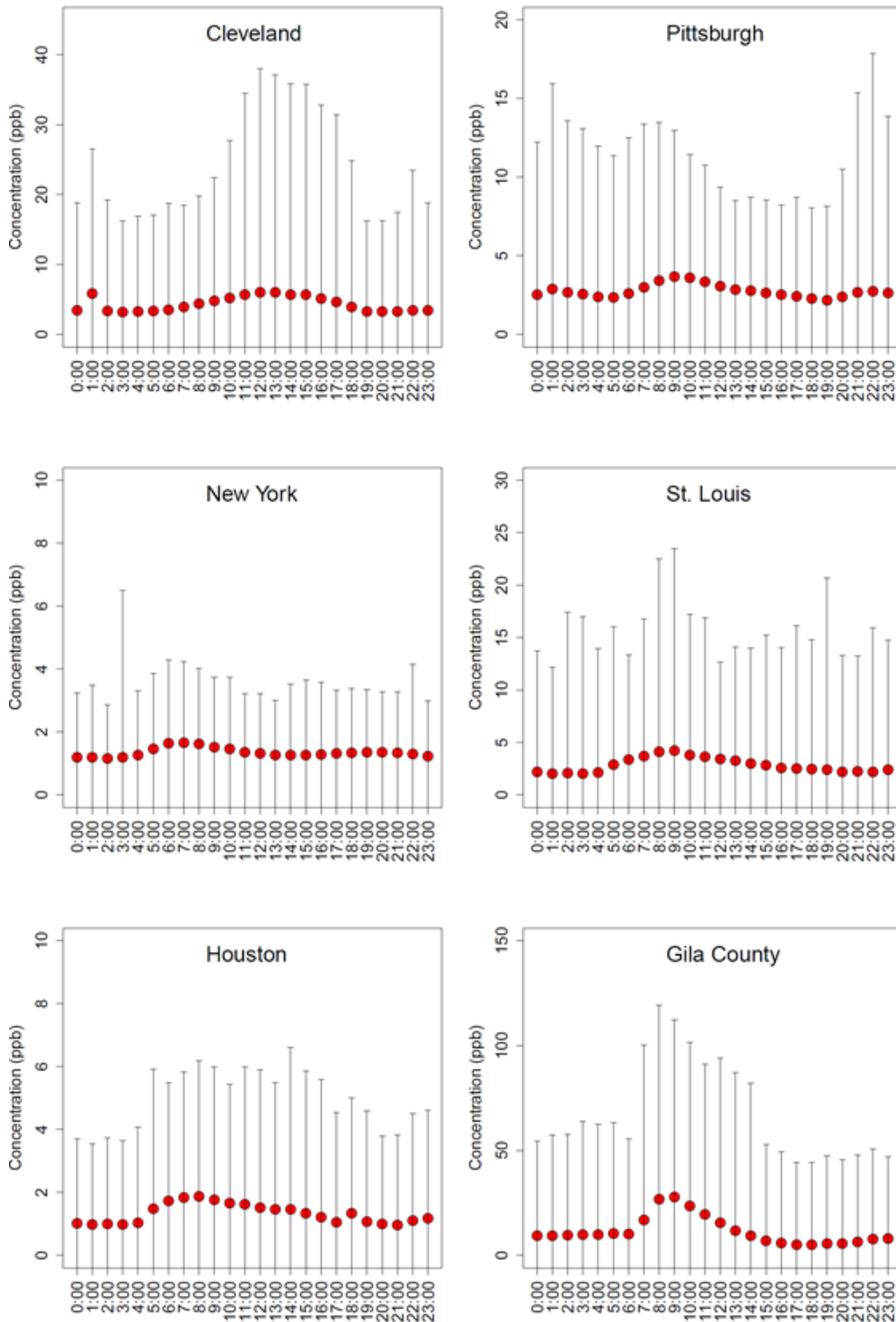
likely reflected an increase in power plant emissions coupled with an increase in entrainment of these elevated emissions into the lower atmosphere as the mixed layer expands throughout the day. Distinct morning peaks may have been related to stable atmospheric conditions, which tend to trap atmospheric pollution near the ground, resulting in an overall increase in ground-level pollution.

Notably, SO₂ concentrations were all well below the primary NAAQS level during all hours of the day in every focus area except Gila County, AZ. In the five urban focus areas, mean 5-minute hourly max and 1-hour avg concentrations were almost all less than 5 ppb. All measured SO₂ concentrations were for the most part below 15 ppb for all but Gila County, AZ, even when examining the upper end of the distribution of 5-minute hourly max concentrations. For Gila County, AZ, the upper bound of the 95% confidence interval for 5-minute hourly max and 1-hour avg SO₂ concentrations exceeded 100 ppb and 50 ppb, respectively. At this location the large copper smelter sources nearby contributed to a strong morning peak in the diel pattern.



Note: For every hour, arithmetic mean concentrations are displayed as red circles. The whiskers represent the 95% confidence interval. Hours are shown in local standard time.

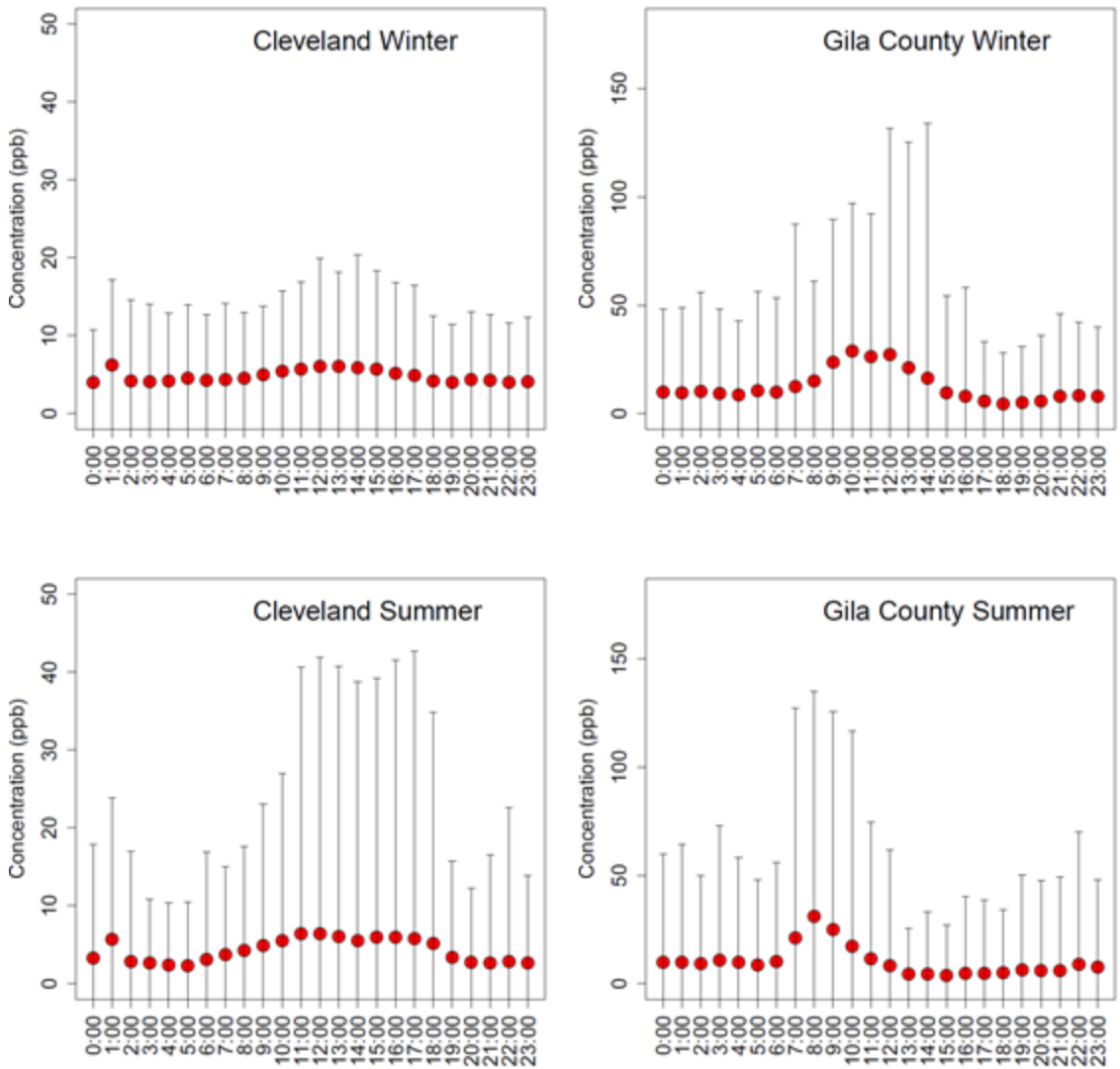
Figure 2-23 Diel variability based on 1-hour avg sulfur dioxide concentrations in the six focus areas, 2013–2015.



Note: For every hour, arithmetic mean concentrations are displayed as red circles. The whiskers represent the 95% confidence interval. Hours are shown in local standard time.

Figure 2-24 Diel trend based on 5-minute hourly max data in the six focus areas, 2013–2015.

Diel SO₂ concentration patterns may be influenced by seasonal factors. Diel plots of 5-minute hourly max for winter and summer are presented for Cleveland, OH and Gila County, AZ in [Figure 2-25](#). Concentration patterns are clearly different between the two locations. While the peak of the mode increased in the summer months indicating higher concentrations within the diel pattern, Cleveland, OH exhibited very little change in the location of the mode for the diel patterns between the cold and warm seasons. In contrast, the mode of the diel pattern occurred earlier in summer compared with winter for Gila County, AZ. Factors that may influence the mode of the diel pattern include peak smelter operation times and atmospheric mixing. For example, seasonal differences in solar radiation prolong nighttime inversion periods during the winter. Transport to downwind monitoring sites may be impeded by stable conditions. Moreover, increased solar radiation during the summer enhances mixing, increasing the probability of plume touchdown ([Slade, 1968b](#)). The average 5-minute hourly max SO₂ concentrations were also somewhat lower during the summer compared with winter in Gila County, AZ. O₃ production in the summer may have promoted oxidation of SO₂ ([Khoder, 2002](#)) to produce the observed losses.



Note: For every hour, arithmetic mean concentrations are displayed as red circles. The whiskers represent the 95% confidence interval. Hours are shown in local standard time.

Figure 2-25 Diel trend based on 5-minute hourly max data in the Cleveland, OH and Gila County, AZ focus areas during winter and summer, 2013–2015.

2.5.4 Relationships between Hourly Mean and Peak Concentrations

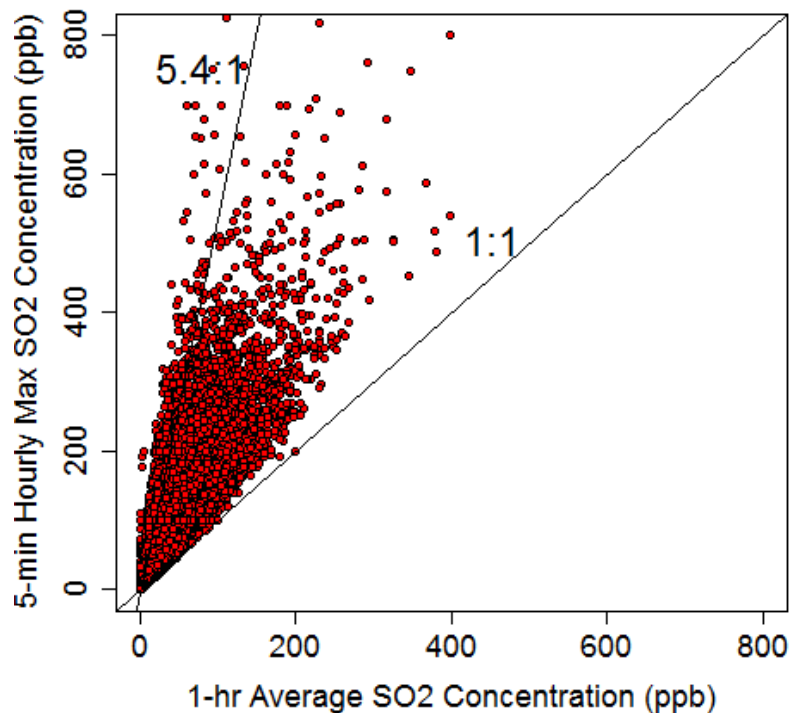
Peak concentrations within an SO₂ plume can greatly exceed the mean concentration at the plume centerline, so that exposure to the peak may be much greater than an hourly or daily SO₂ measurement. Plume dispersion is a Gaussian process, but the plume meanders so that the peak at any instant in time exceeds the mean of the plume centerline found by averaging over some longer time period, such as 1 hour or 1 day ([Slade, 1968a](#); [Gifford, 1960](#)). Several studies ([Dourado et al., 2012](#); [Schauberger et al., 2012](#); [Venkatram, 2002](#); [Turner, 1970](#)) have characterized the peak-to-mean ratio (PMR), showing that the ratio increases with longer concentration averaging time (where average concentration is found in the denominator) compared with the time over which the peak concentration is measured (where peak concentration is found in the numerator). [Venkatram \(2002\)](#) used dispersion modeling to illustrate the stochasticity of the dispersion process, where the mean over a longer time period is determined by an ensemble average across simulations. At a fixed location, the results of [Venkatram \(2002\)](#) imply that exposure to the plume peak occurs with varying probabilities based on the time scale used to represent the instantaneous plume, the time scale over which the average is computed, the intermittency of atmospheric turbulence, and atmospheric stability, where higher atmospheric turbulence would lead to lower PMR.

The PMR has been computed in the literature as a function of the ratio of the mean-to-peak concentration integration times raised to some power in the range of 0.2 ([Venkatram, 2002](#); [Turner, 1970](#)) to 0.5 ([Venkatram, 2002](#)) or 0 to 0.68 ([Schauberger et al., 2012](#)), with the increasing exponent corresponding to increased atmospheric instability. When 5-minute hourly max data are compared with 1-hour avg data, the mean-to-peak integration time ratio is 60 minutes-to-5 minutes = 12. This inverted relationship implies that a larger averaging time generally produces a larger PMR. A peak-to-mean ratio of 1 to 5.4 would be expected using the wider range of exponents (i.e., 12⁰ to 12^{0.68}). PMR varies over space and time due to differences in distance from sources, source characteristics (e.g., stack height), wind speed, and changes in atmospheric stability during the day.

Scatterplots of collocated 5-minute hourly max and 1-hour avg measurements are displayed for all monitors in [Figure 2-26](#) and by focus area in [Figure 2-27](#). Data for the PMR analyses were subject to the same completeness criteria outlined in [Table 2-5 \(Section 2.5.1\)](#).

PMRs were used in the previous SO₂ NAAQS review to evaluate the distribution of 5-minute hourly max concentrations corresponding to a given 1-hour avg SO₂ concentration ([U.S. EPA, 2009c](#)). PMRs are determined by dividing the 5-minute hourly max concentration by the 1-hour avg concentration. Using this approach, a PMR of 1

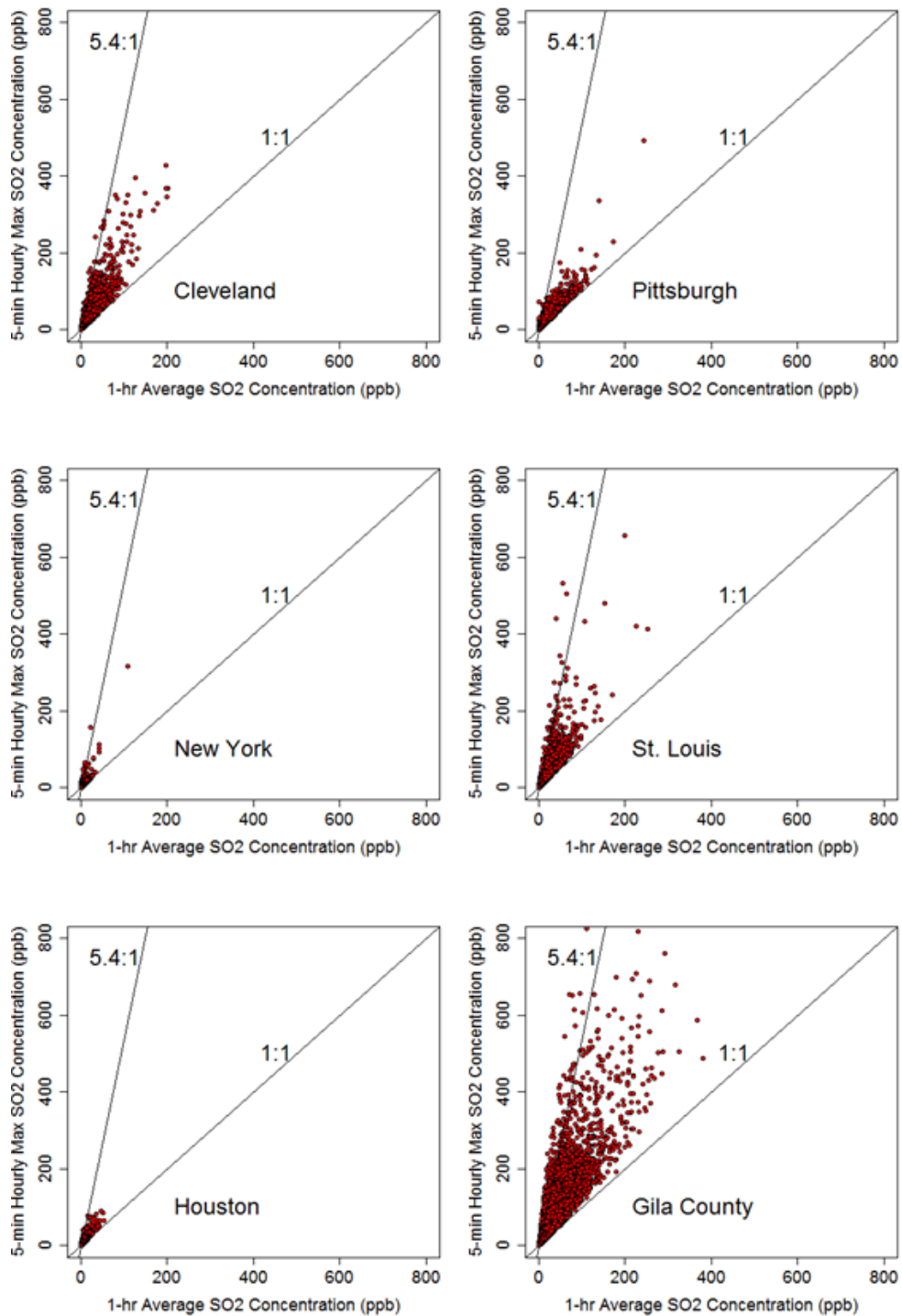
demonstrates that 5-minute hourly max and 1-hour avg concentrations are equivalent. A high PMR value (up to a maximum value of 12 in this case) indicates that the 5-minute hourly max concentration is higher than the 1-hour avg concentration. For example, a PMR of 2 indicates that 5-minute hourly max concentration is 2 times higher than the 1-hour avg concentration. PMR values of 1 and 5.4, the upper value from the literature (Schaubberger et al., 2012), are displayed as lines in Figures 2-26 and 2-27. Median PMRs obtained from comparing the 5-minute hourly max with the 1-hour avg AQS data at sites where both measures were available simultaneously had a range of 1 to 5.5 with a median of 1.3 (99th percentile = 4.7), in reasonable agreement with the predicted range of 1 to 5.4 for the PMR. Concentrations at the 99th percentile 1-hour daily max of 63.5 ppb correspond to a 5-minute hourly max of 200 ppb with a PMR of 3.2.



SO₂ = sulfur dioxide.

Note: Peak-to-mean ratios are displayed on each scatter plot as 1:1 (5-min hourly max = 1-h avg) and 5.4:1 (5-min hourly max is 5.4 times higher than 1-h avg).

Figure 2-26 Scatterplot of 5-minute hourly max versus 1-hour avg sulfur dioxide concentrations, 2013–2015.



SO₂ = sulfur dioxide.

Note: Peak-to-mean ratios are displayed on each scatter plot as 1:1 (5-min hourly max = 1-h avg) and 5.4:1 (5-min hourly max is 5.4 times higher than 1-h avg).

Figure 2-27 Scatterplot of 5-minute hourly max versus 1-hour avg sulfur dioxide concentrations by focus area, 2013–2015.

[Table 2-15](#) displays the range of temporal correlations between corresponding 5-minute hourly max and 1-hour avg concentrations and the range of PMRs computed from SO₂ measurements reported at these monitoring sites within the six focus areas shown in [Figure 2-27](#). Similar to results in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), 5-minute hourly max concentrations tend to correlate well with 1-hour avg metrics, suggesting that 1-hour avg metrics, in most cases, adequately represent changes in 5-minute hourly max data over time. However, 5-minute hourly max concentrations tend to be higher than 1-hour avg concentrations. PMRs were skewed higher for the Gila County focus area and slightly higher for the New York City focus area. However, overall 1-hour daily max concentrations in New York were relatively low (highest 99th percentile 1-hour daily max was 16.5 ppb), so a PMR of 2 or 3 would lead to a 5-minute hourly max of 33 or 49.5 ppb. In contrast, the 1-hour daily max concentrations in Gila County were much higher (highest 99th percentile 1-hour daily max was 247 ppb), which would lead to 5-minute hourly max concentrations of 494 ppb if the PMR were 2 and of 741 ppb if the PMR were 3.

Table 2-15 Pearson correlation coefficient comparing 1-hour avg with 5-minute hourly max and peak-to-mean ratio for maximum sulfur dioxide concentrations in the six focus areas, 2013–2015.

Focus Area	N Monitoring Sites	Correlation Coefficient	Median PMR ^a
Cleveland, OH	7	0.89–0.93	1.00–1.85
Pittsburgh, PA	9	0.91–0.97	1.00–1.40
New York City, NY	12	0.66–0.98	1.28–2.33
St Louis, MO	7	0.88–0.94	1.17–1.38
Houston, TX	9	0.91–0.95	1.33–1.69
Gila County, AZ	4	0.84–0.93	3.24–6.15

N = population number; PMR = peak-to-mean ratio.

^aMedian PMR = 5 min max/1-h avg. The range of data represents median PMR across each site within the focus area.

2.5.5 Background Concentrations

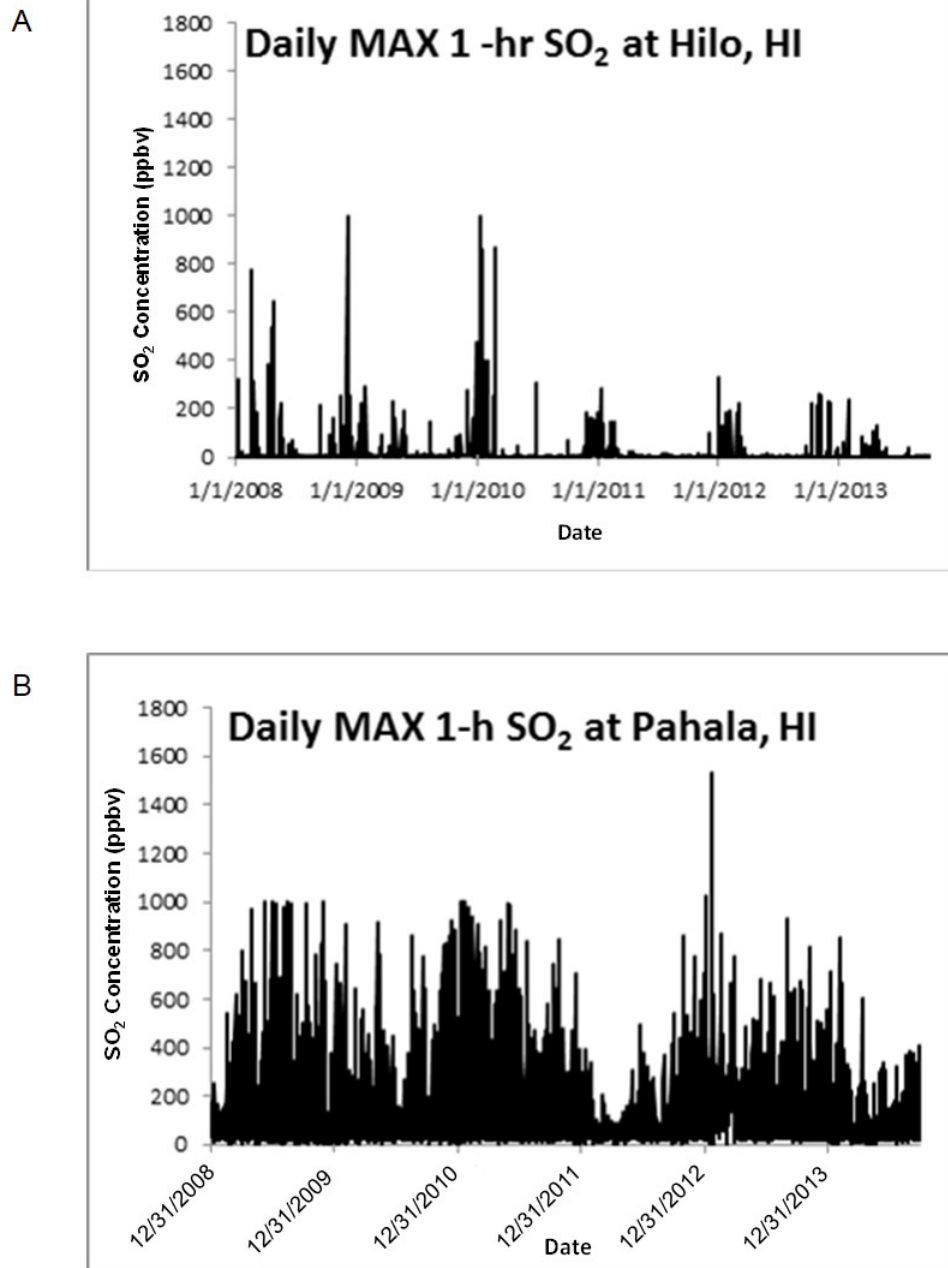
With the exception of periodic volcanic eruptions in Hawaii, natural and international transboundary sources of SO₂ make only minor contributions to the total atmospheric burden of SO₂ in the U.S. [Sections 2.2.4](#) and [2.2.5](#) describe those sources contributing to background SO₂.

No new studies have appeared that attempt to estimate background SO₂ concentrations since the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). The 2008 SO_x ISA discussed a global scale three-dimensional modeling study that estimated annual mean SO₂ concentrations in surface air including both anthropogenic and natural sources, using the MOZART-2 (Model of Ozone and Related Chemical Tracers) ([Horowitz et al., 2003](#)). Sources included in the study were emissions from fossil and biofuel combustion, biomass burning, biogenic and soil emissions, and oceanic emissions. Background SO₂ concentration estimates were below 0.01 ppb over much of the U.S. Aside from a few areas influenced by near-border sources, maximum background concentrations of SO₂ are 0.03 ppb. In the U.S. Northwest, geothermal sources of SO₂ are responsible for 70 to 80% of the background SO₂ concentration; even so, total SO₂ concentrations are still on the order of ~2 ppb or less. In these simulations, background contributed less than 1% to SO₂ concentrations in surface air in 2001 throughout much of the contiguous U.S.

Satellite-borne instruments have mapped large SO₂ sources globally and have obtained data showing intercontinental transport. [Fioletov et al. \(2013\)](#) identified a number of “hotspots” for continuous SO₂ emissions, both anthropogenic and volcanic (e.g., industrial sources in China, Russia, the U.S., the Gulf of Mexico, and Saudi Arabia; volcanic sources in Kīlauea, HI and Anahatan in the Marianas). [Clarisse et al. \(2011\)](#) showed evidence for transport of SO₂ from Asia to Alaska and Canada. In one such episode in November 2010, there was a clearly defined plume crossing the Pacific.

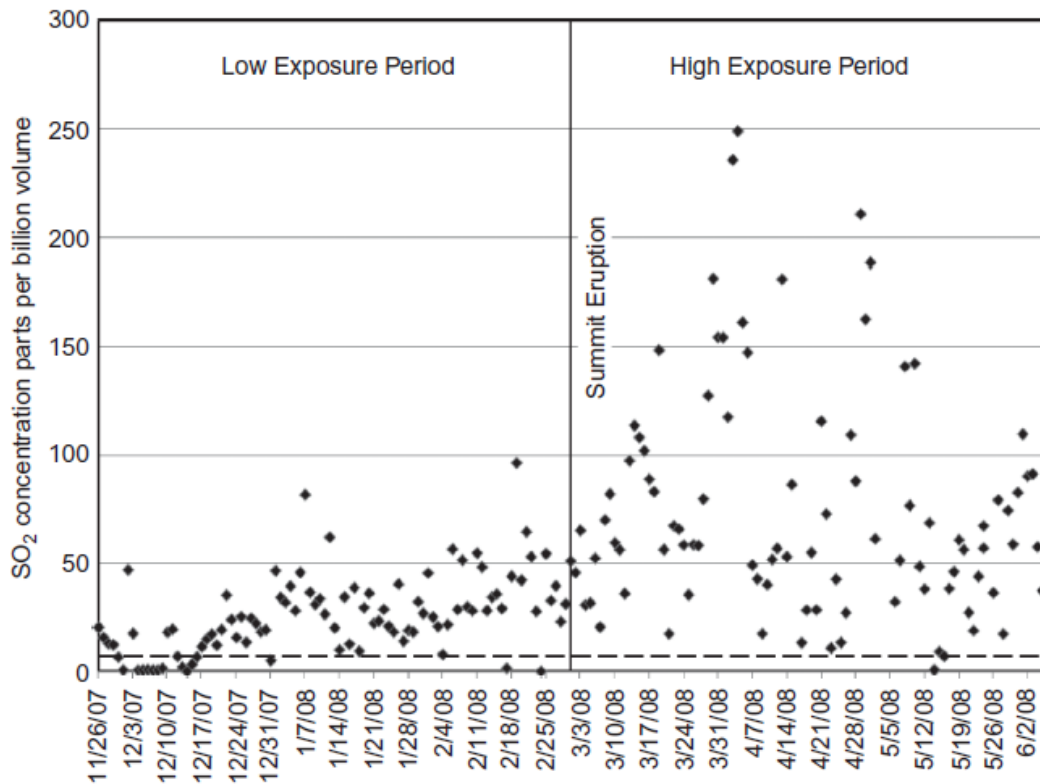
As described in [Section 2.2.4.2](#), volcanic sources of SO₂ in the U.S. are found in the Pacific Northwest, Alaska, and Hawaii. The most important domestic effects from volcanic SO₂ occur on the Hawaiian Islands. Nearly continuous venting of SO₂ from Mauna Loa and Kīlauea produces SO₂ in high concentrations that can affect populated areas on the Big Island of Hawaii (as well as others in the chain, depending on wind conditions). [Figure 2-28A](#) shows the 2008–2013 time series for 1-hour daily max SO₂ concentrations at Hilo, HI, (population of approximately 40,000), which is located about 50 km northeast of Kīlauea. [Figure 2-28B](#) shows the same time series at Pahala (population ~1,300), which is located about 30 km southeast of Kīlauea ([Longo et al., 2010](#)). As demonstrated by these figures, 1-hour daily max SO₂ concentrations can reach levels greater than 1,000 ppb. [Figure 2-29](#) shows a 6-month concentration time series for

the Ka'u District, one of the other communities scattered throughout the southern half of the island that are also exposed to high SO₂ concentrations ([Longo et al., 2010](#)).



SO₂ = sulfur dioxide.

Figure 2-28 1-hour daily max sulfur dioxide concentrations measured at (A) Hilo, HI and (B) Pahala, HI.



SO₂ = sulfur dioxide.

Note. The dashed line represents the World Health Organization 24-h avg SO₂ guideline = 7.5 ppbv ([WHO, 2006](#)).

Data source: SO₂ measured continuously by a TECO pulsed-fluorescence monitor, State of Hawaii Air Quality Division.

Source [Longo et al. \(2010\)](#). Reprinted with permission of Taylor and Francis.

Figure 2-29 Average 24-hour ambient sulfur dioxide concentrations during low and high (volcanic gas) concentration study periods (November 26, 2007 to June 6, 2008) for Ka’u District, located downwind of Kīlauea Volcano.

2.6 Atmospheric Modeling

This section discusses various modeling techniques to estimate ambient concentrations of SO₂. Different types of models are discussed in terms of their capabilities, strengths, and limitations. The section focuses on recent models that have been widely used in U.S. applications. [Section 2.6.1](#) focuses on dispersion models, which are the most widely used and the most relevant for modeling the influence of large point sources on local-scale SO₂ concentrations in the urban and other near-field environments. [Section 2.6.2](#) briefly

discusses chemical transport models (CTMs) that can be used to model SO₂ concentrations at regional and national scales.

2.6.1 Dispersion Modeling

Atmospheric transport and dispersion (ATD) models are important mathematical tools for simulating the fate of air pollutants in support of a wide variety of environmental assessments. ATD models can be used to estimate SO₂ concentration for regulatory purposes if monitoring data are not available or sufficient (75 CFR 35520). Using equations that represent the physical and chemical atmospheric processes that govern dispersal and fate, ATD models provide an estimate of the concentration distribution, both temporally and spatially, of pollutants emitted from sources such as industrial facilities, roadways, and urban areas. The processes that are most important vary depending on the particular model application. The models must specifically account for the characteristics of the source or sources of the pollutant (e.g., buoyant releases), the meteorological conditions, the surrounding surfaces and complexities (e.g., buildings, terrain, and trees), the background concentrations from sources not considered directly in the modeling and the chemical transformations of the pollutant in the atmosphere. The use of ATD models in health studies for SO₂ is described in [Section 3.3.2.4](#).

Dispersion models are particularly important to pollutant studies where monitoring is not practical or sufficient. For pollutants such as SO₂ where spatial distributions of 1-hour avg concentrations associated with large sources often contain extreme gradients, the siting of individual monitors to capture high ground-level concentrations over a wide variety of sources and meteorological conditions would be challenging at best. Extensive arrays of monitors are impractical. Thus, the implementation program for the 2010 primary SO₂ NAAQS allows for air quality modeling to be used in place of monitoring to characterize air quality, and for such air quality information to be used in the process for informing final designation decisions (75 FR 35520). The SO₂ NAAQS is currently the only criteria pollutant standard for which modeling may be used to characterize air quality for the purpose of the area designation process. In addition, modeling is critical to the assessment of the impact of future sources or proposed modifications where monitoring cannot inform. Also, modeling is helpful in the design and implementation of mitigation techniques for addressing existing pollution problems and for compliance evaluations.

ATD models take many forms. They include steady-state (emissions and meteorology), Gaussian-based formulations [e.g., AERMOD, ([Cimorelli et al., 2005](#))]; Lagrangian models [e.g., SCIPUFF, ([Sykes et al., 2007](#)); HYSPLIT, ([Draxler, 1999](#)); ([NOAA](#),

2014)], which are particularly useful when emissions and meteorological conditions are variable over the modeling increment, and Eulerian photochemical grid-based models [e.g., Community Multiscale Air Quality (CMAQ), (Byun and Schere, 2006)], which explicitly model chemical processes and have modeling resolution ranges from about one to tens of kilometers. Additionally, there are stochastic or statistical approaches using, for example, Monte Carlo techniques (Hanna et al., 1982) or those using simple regression approaches (Banerjee et al., 2011). For very complex flows such as a release within an urban canopy of a city, computational fluid dynamics models are considered. Hanna et al. (2006) demonstrated that these models are capable of reproducing the general flow and measured tracer dispersion patterns when very detailed source and three-dimensional building information are available.

In the U.S., steady-state Gaussian models are the most common dispersion models used for primary pollutants like SO₂ (U.S. EPA, 2010a). These models may be used to determine compliance with standards and primary pollutant impacts from new or proposed sources. The same is true for these types of analyses in other countries. For example, the Advanced Dispersion Modeling System (ADMS) (Carruthers et al., 1995), the Hybrid Plume Dispersion Model (HPDM) (Hanna and Chang, 1993), the Danish model, Operationelle Meteorologiske Luftkvalitetsmodeller (OML) (Olesen et al., 1992), and several other steady-state Gaussian-based models have been recommended by the European Environment Agency (van Aalst et al., 1998) for applications involving SO₂ from smoke stacks. Other examples in which Gaussian-type models are found to be applicable for near-field applications are by the U.K. Department of Environment, Food, and Rural Affairs (Williams et al., 2011) and by the New Zealand Ministry of the Environment (Bluett et al., 2004). The primary concerns for many of these compliance-type applications are the magnitude, location, and frequency of high concentrations and the strong gradients of concentrations found near sources. Often the highest concentrations are found within a few kilometers and sometimes within tens of meters of the source. Near-field or near-to-the-source dispersion is the real strength of steady-state modeling.

The American Meteorological Society (AMS)/EPA Regulatory Model (AERMOD) is the preferred model of U.S. EPA for the vast majority of near-field applications with the Off-shore and Coastal Dispersion model (OCD) being used for offshore emissions and alternative models used for unique situations [e.g. the California Puff model (CALPUFF) for Class I area screening application) where justified. AERMOD represents a modernization of applied Gaussian models with advances in areas such as: boundary layer scaling formulations; dispersion rates for both surface and elevated releases; plume interactions with buildings and complex terrain; and characteristics of point, area, and volume source types. In convective conditions, where dispersion produces a distinctly

non-Gaussian vertical pollutant distribution, AERMOD provides a three-part formulation (each Gaussian) that when combined yield distributions representative of those observed ([Weil et al., 1997](#); [Briggs, 1993](#)). The challenges faced by Gaussian models in very light wind conditions are addressed in AERMOD by simulating a meandering plume, and providing turbulence-based lower limits on the transport wind speed and an empirically based correction for the surface friction velocity. In recent years, U.S. EPA has been working to improve AERMOD predictions under light wind conditions, including an adjustment of surface friction velocity under stable light wind conditions (80 FR 45340). For modeling applications where light and variable winds are dominant and reliable wind field estimates are available, models such as SCIPUFF or HYSPLIT provide estimates of plume trajectories and more temporally resolved concentration distributions [e.g., [Wannberg et al. \(2010\)](#)].

AERMOD and models like it are designed to simulate concentrations on an hourly increment, and model evaluations are focused on averaging times of 1 hour or greater ([Perry et al., 2005](#)), although any averaging times may be used, including 5-minute averages, provided that dispersion parameters used in the model have been estimated from available data for the selected averaging time ([Pasquill and Smith, 1983](#)). Longer term concentrations are obtained by averaging the 1-hour concentrations. Spatial resolution is simply determined by the density of receptors included in the analysis (i.e., very high resolution possible). For each hour, emissions and other source characteristics, land surface characteristics, and meteorological conditions are provided to the model. Additionally, the model requires a description of buildings and complex terrain within the modeling domain that are expected to influence pollutant dispersion. The model can simulate hundreds of sources and receptors, providing for analyses in urbanized and industrialized areas.

One limitation of the Gaussian approach is the assumption of steady conditions over a 1-hour modeling period and over the plume transport distance to the receptors. The model is recommended for receptors up to 50 km from a source when steady conditions are appropriate ([U.S. EPA, 2005b](#)). However, this can be challenging, especially for light winds. Under low wind conditions, there are concerns that AERMOD can overestimate measured SO₂ concentrations without adjustment for empirical relationships between wind and concentration ([Paine et al., 2015](#)). Recent updates to AERMOD have been made by the U.S. EPA to address those concerns (80 FR 45340). AERMOD is also limited in its treatment of SO₂ chemistry, using a method much simpler than the more rigorous simulation of atmospheric transformation of SO₂ found in models such as CMAQ or SCICHEM ([Chowdhury et al., 2012](#)). AERMOD uses a simple 4-hour half-life assumption for reducing SO₂ concentration in the plume with travel time ([Turner, 1964](#)). This approach yields results consistent with the SO₂ residence time estimates by [Hidy](#)

(1994) and [Seinfeld and Pandis \(2006\)](#). Therefore, for conditions and sources where the highest hourly concentrations are expected to be relatively close to the source, chemistry is not expected to play a major role in determining compliance with primary standards.

Lagrangian puff dispersion models, such as CALPUFF, have been developed as an alternative to Gaussian dispersion models, such as AERMOD. CALPUFF models SO₂ as particles and then uses a Lagrangian step algorithm to model nonsteady-state dynamics, using time-varying winds specified by meteorological models, such as MM5 [e.g., [Atabi et al. \(2016\)](#), [Abdul-Wahab et al. \(2011\)](#), [Souto et al. \(2014\)](#), [Lee et al. \(2014\)](#), [Zhang et al. \(2015a\)](#)]. The nonsteady-state approach offered by Lagrangian puff dispersion models may be considered an alternative to Gaussian dispersion models that do not account for time dependence. Comparisons have been conducted between Lagrangian models such as CALPUFF and Gaussian plume models such as AERMOD. CALPUFF predictions of 24-hour SO₂ concentrations at an oil refinery in Sohar, Oman compared within 36% of measurements ([Abdul-Wahab et al., 2011](#)). Comparison of CALPUFF and AERMOD to SO₂ measurements at a gas refinery in South Pars, Qatar showed that, while CALPUFF and AERMOD both typically underestimated SO₂ measurements, CALPUFF predictions were usually closer to measured SO₂ concentrations compared with AERMOD ([Atabi et al., 2016](#)). However, [Rood \(2014\)](#) observed that Lagrangian puff models and Gaussian dispersion models both underpredicted 1-hour and 9-hour avg concentrations, but the magnitude of bias was larger in the Lagrangian puff models applied at a field site in Colorado with variable winds and natural topography. [Holnicki et al. \(2016\)](#) noted that the model performance improved with longer averaging times and that the 1-hour avg concentration predicted by CALPUFF was less accurate than predictions for annual average concentrations, when compared to SO₂ measurements. However, recent dispersion modeling results were compared between CALPUFF and AERMOD for the Section 126 Petition from New Jersey for the Portland Generating Station (76 FR 69052) where CALPUFF overestimated 1-hour daily max SO₂ observations taken in Columbia, NJ by 226%, while AERMOD overestimated the same observations by 14%.

Uncertainty in the model predictions is influenced by the uncertainty in model input data (in particular emission or source characterization and meteorological conditions) as well as by inadequacies in model formulations. Uncertainty related to model input variables is generally estimated by propagating the expected errors in the individual input variables (e.g., wind speed, emission rate) through the model using Monte Carlo techniques ([Dabberdt and Miller, 2000](#)). In addition, there is uncertainty related to the fundamental difference between modeled and measured concentrations. Monitored data (within sampling error) represents actual realizations of events, while modeling estimates represent ensemble mean concentrations ([Rao, 2005](#)). Based on a study comparing a variety of models (including Gaussian) to a number of tracer field study results, [Hanna et](#)

[al. \(1993\)](#) found that for continuous point releases and receptors within a kilometer of the source, uncertainty in model inputs in combination with the stochastic nature of the atmosphere result in typical mean biases on the order of 20 to 40% and normalized mean square errors up to 70%. The author points out that these levels of difference between model and monitor results would likely exist even for more sophisticated models. [Hanna \(2007\)](#) provided a comprehensive review of methods for determining sensitivity and uncertainty in ATD models.

Focusing on the uncertainties in model inputs, it is easy to see that an individual model estimate paired in time and space with a monitored concentration will likely differ, sometimes substantially. [Weil \(1992\)](#) pointed out that wind direction uncertainties alone can cause disappointing results in space and time pairings from otherwise well-performing dispersion models. With wind direction errors, the plume footprints from the model and that from the observations may not overlap. However, a model that is based on appropriate characterizations of the important physical processes should be able to reproduce the distribution of observed concentrations assuming that the distributions of model inputs is similar to that of the observed conditions ([Venkatram et al., 2001](#)). Meteorological inputs coupled with AERMOD can impact the results, and the output may depend on the use of recorded meteorological observations or meteorological models (e.g., Weather Research and Forecasting (WRF) model). Meteorological models may add error to the dispersion simulation, and that error is impacted by model selection and resolution ([Isakov et al., 2007](#)). Therefore, in evaluating a model's ability to predict concentrations within the modeling domain, it is important to include an analysis of modeled and monitored concentration distributions for any location studied, and possibly to adjust the results based on any observed bias. As part of the proposed update to the Guideline on Air Quality Models, U.S. EPA proposed to allow the use of prognostic meteorological data for regulatory applications of AERMOD (80 FR 45340). U.S. EPA conducted several assessments comparing observed meteorological data to prognostic meteorological data and found that the prognostic data performed adequately ([U.S. EPA, 2015a](#)).

[Chang and Hanna \(2004\)](#) provided a comprehensive discussion of methods for evaluating the performance of air quality models. They discuss a series of performance measures that included statistical metrics such as fractional bias (FB), geometric mean bias, normalized mean squared error and the fraction of estimates within a factor of two observations. These and other measures are included in the commonly used BOOT software ([Chang and Hanna, 2005](#)), which also allows for estimation of confidence limits on the concentrations computed and provides insight about the sources of bias in the model ([Irwin, 2014](#)). [Chang and Hanna \(2004\)](#) also discussed exploratory analysis methods of plotting and analyzing the modeled and measured concentrations. They

pointed out that the most useful model evaluation studies are those that examine a number of models and compare them with a number of field studies.

For models intended for application to compliance assessments (e.g., related to the 1-hour daily max SO₂ standard), the model's ability to capture the high end of the concentration distribution is important. Measures such as robust highest concentration (RHC) ([Cox and Tikvart, 1990](#)), and exploratory examinations of quantile-quantile plots ([Chambers et al., 1983](#)) are useful. The RHC represents a smoothed estimate of the top values in the distribution of hourly concentrations. In contrast, for dispersion modeling in support of health studies where the model must capture concentrations at specified locations and time periods, additional measures of bias and scatter are important.

The intended use of a model and the objective of a model evaluation guide the selection of evaluation criteria. [Frost \(2014\)](#) evaluated the model performance for AERMOD when it was applied to the study of 1 year of SO₂ emissions from three coal-fired EGUs. The study authors found agreement within a factor of two of the 99th percentile SO₂ design value for the majority of the data but noted performance outside a factor of two for the top 5% of measured 1-hour avg concentrations. However, [Rehbein et al. \(2014\)](#) found that the model fell within a factor of two of the monitoring data even at high concentrations for a model validation outside a nickel smelting facility in Sudbury, Ontario, Canada. U.S. EPA also conducted evaluations of prognostic meteorological data in AERMOD ([U.S. EPA, 2015a](#)), including the facility modeled by [Frost \(2014\)](#). These evaluations included data analysis adhering to the U.S. EPA Protocol for Best Performing Models, which includes a scientific and operational component of model performance ([U.S. EPA, 1992](#)). SO₂ concentrations modeled by AERMOD were within a factor of two observations in all but one simulation when using the metrics of the protocol. Meteorological parameters were modeled with FB within 20% of observations ([U.S. EPA, 2015a](#)).

At the time of its inclusion into the U.S. EPA Guideline on Air Quality Models ([U.S. EPA, 2005b](#)), the performance of AERMOD was evaluated against seventeen field-study databases over averaging times from 1 hour to 1 year ([Perry et al., 2005](#)). In each case, the emissions characteristics and background concentrations were well known, meteorological data were available on site, and tracer concentrations were measured at multiple locations where high plume impacts were expected. Four of the studies involved very dense sampler arrays. For the four intensive studies, [Perry et al. \(2005\)](#) found the ratio of modeled 1-hour avg RHC to monitored RHC ranged from 0.77 to 1.18 [i.e., relatively unbiased in estimating extreme (high) values]. For studies involving tall buoyant stacks with more limited monitoring locations, 1-hour ratios were not reported, but the 3-hour avg ratios ranged from 1.0 to 1.35 (i.e., a slight tendency to overpredict the

high concentrations). Examination of quantile-quantile plots supported the findings that the model was capturing the upper end of the 1- and 3-hour avg concentration distribution. [Hanna et al. \(2001\)](#) evaluated the AERMOD and ADMS Gaussian dispersion models with five field study databases including area sources, low releases, and tall power plant stacks in rural, flat, and complex terrain. Among the median performance measures they reported, the ratio of maximum modeled to maximum observed concentrations was 0.77 for AERMOD and 0.80 for ADMS, each a small underprediction. The median value over the five databases of the geometric mean (MG, a measure of the ratio of averaged modeled to monitored concentration) was 1.7 for AERMOD and 1.22 for ADMS. With 1.0 as the ideal value, both models were found to overpredict (with ADMS less biased). Unlike the ratio of maximum values, MG is a measure of performance over the entire distribution of concentrations. [Hurley \(2006\)](#) also evaluated AERMOD and two Australian models against seven field studies and found no database against which AERMOD performed poorly.

With the adoption of the 2010 1-hour daily max SO₂ standard, there is renewed interest in AERMOD's abilities to simulate near-field maximum short-term concentrations. A number of specific areas for model improvement were discussed at the 10th and 11th Modeling Conference on Air Quality in 2012 ([U.S. EPA, 2012a](#)) and 2015 ([U.S. EPA, 2016a](#)). Among them were concerns about simulations in stable conditions with light and meandering winds, use of prognostic meteorological data, modeling of emissions from haul roads, plume chemistry, and building downwash. Proposed improvements include an adjusted friction velocity model for stable/low wind conditions in AERMET, a new model for dispersion options in AERMOD, and an option for buoyant line sources in AERMOD ([U.S. EPA, 2016a](#)). Routine reporting of 5-minute average concentrations by air agencies may also facilitate model evaluation. Research in many of these areas is underway, and improvements to AERMOD have been made based on the outcomes of those conferences, largely as part of U.S. EPA rulemaking to revise the *Guideline*. While the stochastic nature of the atmosphere will always preclude the development of a perfect model, improvements to the model formulations will continue with the goal of estimating hourly average concentrations while reducing model uncertainty and expanding applicability.

2.6.2 Chemical Transport Models

Chemical transport models are an important tool for characterizing regional- and national-scale air quality. The scales at which they typically operate are too large to satisfactorily capture meteorological and chemical processes involving SO₂ at the local or

near-source scale. The dispersion models discussed previously are thus preferable for characterizing SO₂ concentrations at these scales.

Chemical transport models such as the CMAQ model, are deterministic models of chemical transport that account for physical and chemical processes, including advection, turbulence, diffusion, deposition, gas-phase and heterogeneous chemistry, and convective cloud transport, while following the constraint of mass conservation ([Byun and Schere, 2006](#)). CTMs provide regional concentration estimates and are typically run with horizontal grid resolutions of 4, 12, or 36 km. Temporal resolutions are typically 1 hour, although larger temporal aggregation often occurs for the purpose of maintaining reasonable data file size. CTMs are used to compute interactions among primary atmospheric pollutants and their transformation products, the production of secondary aerosols, the evolution of particle size distribution, and transport and deposition of pollutants. CTMs are driven by emissions inventories for primary species such as SO₂, NO₂, NH₃, VOCs, and primary PM, and by meteorological fields produced by other numerical weather prediction models. Values for meteorological variables such as winds and temperatures are taken from a meteorological model that is nudged by operational analyses, reanalyses, or general circulation models. In most cases, these are off-line meteorological predictions; thus, they are not modified by radiatively active species generated by the air quality model. Work to integrate meteorology and chemistry was initiated in the mid-1990s [by [Lu et al. \(1997a\)](#) and [Lu et al. \(1997b\)](#) and references therein], although limits to computing power prevented widespread application. More recently, new integrated models of meteorology and chemistry are available; see, for example, the Weather Research and Forecast model with chemistry (WRF-Chem; <http://ruc.noaa.gov/wrf/wrf-chem/>) and WRF-CMAQ ([Wong et al., 2012](#)).

Biases in SO₂ concentrations predicted by CTMs can occur as a result of error in model representation of atmospheric processes converting SO₂ to H₂SO₄ and in removal processes. For example, overestimates of cloud-based reactions converting SO₂ to H₂SO₄ have been shown to negatively bias SO₂ concentration estimates in CMAQ v4.6 ([Mueller et al., 2011](#)). Improvements to modeling these processes, such as capturing metal catalysis of the SO₂ → H₂SO₄ conversion process, have been included in CMAQ v5.0.2 to improve model estimates of SO₂ and SO₄²⁻ ([Alexander et al., 2009](#)). Therefore, when using CMAQ to estimate exposure to SO₂, attention must be given to the version of the model so that any inherent biases are understood.

The Air Quality Model Evaluation International Initiative (AQMEII) was developed by scientists in Europe and North America to evaluate several CTMs against each other using common input data sets ([Rao et al., 2011](#)). [Pouliot et al. \(2015\)](#) assembled emissions input data for European and North American simulations performed over two

phases of the AQMEII study and found a 12% reduction in SO₂ emission estimates for 2006 in both Europe and North America. These differences were attributed to differences in methodologies used to estimate emissions and to differences in input data that influence the CTM output. In a comparison of CTM models of SO₂ with surface measurements in Europe, the Modeling Atmospheric Composition and Climate (MACC) model reanalysis overestimated surface SO₂ concentrations by 40% in winter and underestimated surface SO₂ levels by 63% in summer ([Giordano et al., 2015](#)). In North America, MACC underestimated SO₂ in summer by 81%. MACC results were higher than regional CTMs in the winter for North America, and seasonal variability was not well captured ($r = 0.16$ in summer and $r = 0.19$ in winter). These errors were thought to relate to the differences in the lifetime of SO₂ transported from the domain borders to the domain center being shorter than the timescale of the model bias.

2.7 Summary

Of the sulfur oxides, SO₂ is the most abundant in the atmosphere, the most important in atmospheric chemistry, and the one most clearly linked to human health effects ([Chapter 5](#)). Thus, the NAAQS are currently set using SO₂ as the indicator species. As a consequence of several U.S. air quality regulatory programs, emissions of SO₂ have declined by approximately 79% for all NEI source categories during the time period 1990 to 2014 ([Section 2.2](#)). Coal-fired EGUs remain the dominant anthropogenic source, emitting 3.2×10^6 tons SO₂ annually according to the 2014 NEI. Natural sources include volcanoes, wildfires, and biogenic sulfides that are intermittent and of limited spatial extent.

Beyond the size of the emissions source, the important variables that determine the concentration of SO₂ downwind of a source are the photochemical removal processes occurring in the emissions plume ([Section 2.3](#)) and local meteorology. The gas-phase oxidation of SO₂ by hydroxyl radical is slow in comparison to aqueous-phase oxidation in cloud and fog droplets. Clouds and fog can reduce local SO₂ concentrations by converting it to H₂SO₄ in the droplet phase. Another gas-phase oxidation mechanism involves a Criegee intermediate biradical that participates in converting SO₂ to SO₃. The Criegee-based SO₂ oxidation mechanism may amplify the rate of SO₂ removal in areas with high concentrations of Criegee precursors (i.e., low molecular weight organic gases, such as biogenic compounds, and unsaturated hydrocarbons) present downwind of industrial sites and refineries. The atmospheric SO₂ oxidation processes, coupled with variable meteorological conditions, including wind, atmospheric stability, humidity, and cloud/fog cover, influence the observed SO₂ concentrations at urban monitoring sites.

Changes were undertaken to the existing U.S. EPA monitoring network as a result of the new 1-hour daily max primary NAAQS standard promulgated in 2010 ([Section 2.4](#)). First, the automated pulsed UVF method, the method most commonly used by state and local monitoring agencies for NAAQS compliance, was designated as an FRM. Second, new SO₂ monitoring guidelines require states to report 5-minute data in light of health effects evidence on lung function decrements among exercising individuals with asthma following a 5–10 minute exposure of SO₂ above 200 ppb ([Section 5.2.1.2](#)). There are 380 monitoring sites across the U.S. reporting 5-minute data. Analysis of environmental concentrations of SO₂ data reported in this chapter reflect the monitoring network changes, particularly the analysis of the recent 5-minute data.

On a nationwide basis, the average 1-hour daily max SO₂ concentration reported during 2013–2015 was 5.3 ppb ([Section 2.5.2.1](#)). The national 99th percentile 5-minute hourly max concentration was 23.8 ppb, suggesting that the occurrence of very high 5-minute peak values is rare. However, peak concentrations (99th percentile) of the 1-hour daily max SO₂ concentrations were greater than 75 ppb at some monitoring sites located near large anthropogenic sources (e.g., power plants). Volcanoes produce large amounts of SO₂, and hourly concentrations in their vicinity can be greater than 2,000 ppb. SO₂ concentration is highly variable across urban spatial scales ([Section 2.5.2.2](#)), exhibiting moderate to poor correlations between SO₂ concentrations measured at different monitoring sites across a metropolitan area. This high degree of urban spatial variability may not be fully captured by central site monitoring estimates.

Long-term concentration trends show a steady decline in the mean, 10th, and 90th percentile of the site-specific 99th percentile of the 1-hour daily max SO₂ concentrations ([Section 2.5.3](#)). The data show a 76% decline in 99th percentile 1-hour daily max SO₂ concentration over the period 1990–2015. Seasonal trends were examined for six focus areas, and only New York and, to a lesser extent, Houston exhibited strong intra-annual trend in which cool season 1-hour daily max SO₂ concentrations were higher than warm season 1-hour daily max SO₂ concentrations. Diel patterns in 1-hour avg SO₂ concentration mostly shows daytime concentrations peak in the morning or midday, and the time of the peak can vary by location and may be influenced by seasonal conditions.

Peak concentrations within an SO₂ plume can greatly exceed the mean concentration at the plume centerline, so that exposure to the peak may greatly exceed an hourly or daily SO₂ measurement ([Section 2.5.4](#)). PMRs obtained from comparing the 5-minute hourly max with the 1-hour avg AQS data at sites where both measures were available simultaneously had a range of 1 to 5.5 with a median of 1.3. In a city with low SO₂ concentrations, a high PMR may still be related to elevated 5-minute hourly max SO₂ concentration. For example, overall 1-hour daily max concentrations in the New York

focus area were relatively low (highest 99th percentile 1-hour daily max was 16.5 ppb), so a PMR of 2 or 3 would lead to a 5-minute hourly max of 33 or 49.5 ppb. In contrast, the 1-hour daily max concentrations in Gila County were much higher (highest 99th percentile 1-hour daily max was 247 ppb), which would suggest 5-minute hourly max concentrations of 504 ppb if the PMR were 2 and of 741 ppb if the PMR were 3.

Contributions to background concentrations include natural emissions of SO₂ and photochemical reactions involving reduced sulfur compounds of natural origin, as well as the transport of sulfur compounds from outside of the U.S. ([Section 2.5.5](#)). In the U.S. Northwest, geothermal sources of SO₂ are responsible for 70 to 80% of the background SO₂ concentration; even so, total SO₂ concentrations are still on the order of ~2 ppb or less. In model simulations, background contributed less than 1% to SO₂ concentrations in surface air in 2001 throughout much of the contiguous U.S. Even with ambient concentrations for 2013–2015 that were roughly half the magnitude of those measured around 2001, the estimated background SO₂ would contribute only 2% to ambient SO₂ concentrations in most of the contiguous U.S.

Atmospheric modeling includes dispersion and chemical transport models to estimate SO₂ concentrations in locations where monitoring is not practical or sufficient ([Section 2.6](#)). Because existing ambient SO₂ monitors may not be sited in locations to capture peak 1-h daily max concentrations, the implementation program for the 2010 primary SO₂ NAAQS allows for air quality modeling to be used to characterize air quality for informing designation decisions (75 FR 35520). Modeling is critical to assessing the impact of future sources or proposed modifications when monitoring cannot be informative, and for designing and implementing mitigation techniques.

CHAPTER 3 EXPOSURE TO AMBIENT SULFUR DIOXIDE

3.1 Introduction

The 2008 SO_x Integrated Science Assessment (ISA) ([U.S. EPA, 2008d](#)) evaluated ambient sulfur dioxide (SO₂) concentrations and exposure assessment in multiple microenvironments, presented methods for estimating personal and population exposure via monitoring and modeling, analyzed relationships between personal SO₂ exposure and ambient SO₂ concentrations, and discussed the implications of using ambient SO₂ concentrations to estimate exposure in epidemiologic studies. This chapter summarizes that information and presents new information regarding exposure to ambient SO₂. The chapter will focus on the inhalation exposure route for SO₂ from the key sources described in [Chapter 2](#) because the presence of other SO_x species in the atmosphere has not been demonstrated, as discussed previously. Exposure to particulate sulfate formed by oxidation of SO₂ is addressed in the particulate matter (PM) ISA ([U.S. EPA, 2009a](#)). Sections within the chapter are organized to first present broad exposure concepts applicable to air pollution in general, followed by SO₂-specific material. Topics addressed in the chapter include methodological considerations for use of exposure data, and exposure assessment and epidemiologic inference. Many new studies are included in this chapter to better characterize exposure and understand exposure error. This material provides context for interpreting the epidemiologic studies described in [Chapter 5](#).

3.2 Conceptual Overview of Human Exposure

3.2.1 Exposure Terminology

A variety of metrics and terms are used to characterize air pollution exposure. They are described here at the beginning of the chapter to provide clarity for the subsequent discussion.

The *concentration* of an air pollutant is defined as the mass or volume of the pollutant in a given volume of air (e.g., µg/m³ or ppb). Concentrations observed in outdoor locations are referred to as ambient concentrations. The term *exposure* refers to contact with a specific pollutant concentration over a certain period of time ([Zartarian et al., 2005](#)), in single or multiple locations. For example, contact with a concentration of 10 ppb SO₂ for

1 hour would be referred to as a 1-hour exposure to 10 ppb SO₂, and 10 ppb is referred to as the *exposure concentration*. As discussed in [Chapter 4](#), dose incorporates the concept of intake into the body (via inhalation). Exposure concentrations are particularly relevant for interpreting controlled human exposure studies, where participants are exposed to a well-defined pollutant concentration, or panel epidemiologic studies that use personal exposure monitors. Ambient concentrations are more relevant to epidemiologic studies using measured or modeled concentrations.

A location where exposure occurs is referred to as a *microenvironment*, and an individual's daily exposure consists of the time-integrated concentrations in each of the microenvironments visited during the day. Ambient air pollution may penetrate indoors (see [Section 3.4.1.1](#) on infiltration), where it combines with air pollution from indoor sources (*nonambient air pollution*) to produce the total measured indoor concentration. Exposure to the ambient fraction of this indoor concentration, together with exposure to ambient concentrations in outdoor microenvironments, is referred to as *ambient exposure* ([Wilson et al., 2000](#)).

Because personal exposures are not routinely measured, the term *surrogate* is used in this chapter to describe a quantity meant to estimate or represent exposure, such as an SO₂ concentration measured at a fixed-site monitor ([Sarnat et al., 2000](#)). When surrogates are used for exposure assignment in epidemiologic studies, exposure misclassification or exposure error can result. *Exposure misclassification* refers to exposure error for categorical variables, such as diseased and nondiseased individuals. Exposure measurement error due to exposure assignment methods and spatial and temporal variability in pollutant concentrations may be either differential or nondifferential. Differential error occurs when the measured exposure and associated error differ across groups such that the mismeasured exposure contains information about the health outcome other than that associated with the true exposure ([Armstrong, 2008](#)). An example of differential exposure error is the use of geocoding to estimate air pollution exposure by proximity to roadways, because concentrations are different upwind and downwind of a major roadway ([Lane et al., 2013](#); [Singer et al., 2004](#)). Nondifferential exposure error refers to the situation where exposure characterization is similarly accurate across all groups ([Armstrong et al., 1992](#)).

Exposure misclassification and exposure error can result in bias and reduced precision of the effect estimate in epidemiologic studies. *Bias* refers to the difference between the observed and true association, while *precision* is typically represented by the width of the confidence interval around the effect estimate. Bias toward the null, or attenuation of the effect estimate, indicates an underestimate of the magnitude of the effect and is characteristic of nondifferential measurement error, particularly classical measurement

error. Bias away from the null can occur through differential exposure measurement error or under certain exposure scenarios ([Armstrong et al., 1992](#)).

Exposure error refers to the bias and uncertainty associated with using measured or estimated concentrations to represent the actual exposure of an individual or population ([Lipfert and Wyzga, 1996](#)). Exposure error has two components: (1) exposure measurement error derived from uncertainty in the quantity being used to represent exposure, whether measured or estimated concentration or exposure, and (2) use of a surrogate in the epidemiologic study in lieu of the true exposure, which may be unobservable. [Section 3.4.4](#) provides additional definitions for specific types of exposure error and discusses the potential impact of such errors on epidemiologic study results.

3.2.2 Conceptual Model of Personal Exposure

A theoretical model of personal exposure is presented in this section to highlight measurable quantities and uncertainties. This model has been developed and presented in previous ISAs, most recently in the 2016 ISA for Oxides of Nitrogen ([U.S. EPA, 2016d](#)), and it is reproduced here to provide context for the current document.

An individual's time-integrated total exposure to SO₂ can be described based on a compartmentalization of the person's activities throughout a given time period:

$$E_T = \int C_t dt$$

Equation 3-1

$$E_T = \sum_{j=1}^n C_j t_j$$

where E_T = total exposure over a time period of interest, C_t = airborne SO₂ concentration during time t spent in a given microenvironment, and dt = portion of the time period spent in a given microenvironment, j . Total exposure can be decomposed into a model that accounts for exposure to SO₂ of ambient (E_a) and nonambient (E_{na}) origin of the form:

$$E_T = E_a + E_{na}$$

Equation 3-2

Although indoor combustion of sulfur-containing fuels, particularly kerosene, is a nonambient source of SO₂ (see [Section 3.4.1](#)), these sources are specific to individuals

and may not be important sources of population exposure. This ISA focuses on the ambient component of exposure because this is more relevant to the National Ambient Air Quality Standards (NAAQS) review. Ambient sources of SO₂, such as electric power generating units (EGUs) and industrial fuel combustion, are described in [Section 2.2.1](#). Assuming steady-state outdoor conditions, E_a can be expressed in terms of the fraction of time spent in various outdoor and indoor (including enclosed microenvironments such as vehicles) microenvironments ([U.S. EPA, 2006](#); [Wilson et al., 2000](#)):

$$E_a = \sum f_o C_o + \sum f_i F_{inf,i} C_o$$

Equation 3-3

where f = fraction of the relevant time period (equivalent to dt in [Equation 3-1](#)); subscript o denotes outdoor microenvironments; subscript i denotes indoor microenvironments; and $F_{inf,i}$ = infiltration factor for indoor microenvironment i . [Equation 3-3](#) is subject to the constraint $\sum f_o + \sum f_i = 1$ to reflect the total exposure over a specified time period, and each term on the right-hand side of the equation has a summation because it reflects various microenvironmental exposures. Here, “indoors” refers to being inside any aspect of the built environment, [e.g., home, office buildings, enclosed vehicles (automobiles, trains, buses), or recreational facilities (movie theaters, restaurants, bars)]. “Outdoor” exposure can occur in parks or yards, on sidewalks, and on bicycles or motorcycles. Assuming steady-state ventilation conditions, the infiltration factor (F_{inf}) is a function of the penetration (P) of SO₂ into the microenvironment, the air exchange rate (AER) (a) of the microenvironment, and the rate of SO₂ loss (k) in the microenvironment:

$$F_{inf} = \frac{Pa}{(a + k)}$$

Equation 3-4

In epidemiologic studies, it is often assumed that the fraction of time spent outdoors can be expressed cumulatively as f_o ; the indoor terms still retain a summation because infiltration differs for different microenvironments. If an epidemiologic study employs only one measure for C_o , then the assumed model of an individual’s exposure to ambient SO₂, given in [Equation 3-3](#), is re-expressed solely as a function of C_o :

$$E_a = \left(f_o + \sum f_i F_{inf,i} \right) C_o$$

Equation 3-5

The spatial variability of outdoor SO₂ concentrations and epidemiologic study design determine whether [Equation 3-5](#) is a reasonable approximation for [Equation 3-3](#). Spatial variability of outdoor SO₂ is influenced by proximity to sources, source characteristics, meteorology, built and natural topography, and oxidation rates. These equations also

assume steady-state microenvironmental concentrations. Errors and uncertainties inherent in using [Equation 3-5](#) in lieu of [Equation 3-3](#) are described in [Section 3.4.4](#) with respect to implications for interpreting epidemiologic studies. Epidemiologic studies may use concentration measured at a fixed-site monitor to represent ambient concentration; thus α , the ratio between personal exposure to ambient SO₂ and the ambient concentration of SO₂, is defined as:

$$\alpha = \frac{E_a}{C_a}$$

Equation 3-6

Combining [Equations 3-5](#) and [3-6](#) yields:

$$\alpha = f_o + \sum f_i F_{inf,i}$$

Equation 3-7

where α varies between 0 and 1. Estimates of α for SO₂ are provided in [Section 3.4.1.3](#). If a person's exposure occurs in a single microenvironment, the ambient component of a microenvironmental SO₂ concentration can be represented as the product of the ambient concentration and F_{inf} . Time-activity data and corresponding estimates of F_{inf} for each microenvironmental exposure are needed to compute an individual's α with accuracy ([U.S. EPA, 2006](#)). In epidemiologic studies, α is assumed to be constant in lieu of time-activity data and estimates of F_{inf} , which varies with building- and meteorology-related air exchange characteristics ([Section 3.4.1.1](#)). If important local outdoor sources and sinks exist that are not captured by fixed-site monitors, then the ambient component of the local outdoor concentration may be estimated using dispersion models, land use regression (LUR) models, receptor models, fine-scale chemical transport models (CTMs), or some combination of these techniques, which are described in [Section 3.3.2](#).

3.2.3 Exposure Considerations Specific to Sulfur Dioxide

The inhalation exposure pathway relevant for SO₂ is influenced by sources, chemistry, meteorology, and ambient concentrations, described in detail in [Chapter 2](#) and summarized briefly here. The majority of SO₂ is emitted by coal-fired EGUs ([Section 2.2](#)); the point source nature of these emissions contributes to the relatively high spatial variability of SO₂ concentrations (both ambient and exposure) compared with pollutants such as PM and ozone (O₃) ([Sections 2.5](#) and [3.4.2.2](#)). Another contributing factor to spatial variability is the dispersion and oxidation of SO₂ in the atmosphere ([Section 2.3](#)), resulting in decreasing ambient SO₂ concentrations with increasing

distance from the source. SO₂ from point sources travels as a plume, which may or may not impact portions of an urban area depending on meteorological conditions. Ambient SO₂ concentrations do not exhibit consistently strong temporal variability over daily or seasonal time scales ([Section 2.5](#)); however, in some areas, concentrations are low during nighttime and show a daytime maximum, affecting temporal exposure patterns. Due to the relative lack of indoor SO₂ sources, personal SO₂ exposure is expected to be dominated by ambient exposure ([Section 3.4.1.3](#)).

3.3 Methodological Considerations for Use of Exposure Data

This section describes techniques that have been used to measure or model concentrations of SO₂ that serve as surrogates for personal SO₂ exposures in epidemiologic studies. Common methods for assigning an exposure surrogate from monitoring data include using ambient SO₂ concentration measured at a single fixed-site monitor to represent population exposure ([Section 3.3.1.1](#)), averaging ambient SO₂ concentrations from multiple fixed-site monitors, or using personal monitoring data ([Section 3.3.1.2](#)). Fixed-site monitoring data are often used as exposure concentration surrogates in time-series epidemiologic studies to examine how changes in SO₂ exposure over time are associated with changes in a health outcome ([Section 3.4.4.1](#)). Panel epidemiologic studies may use personal monitors to estimate personal exposure ([Section 3.4.4.3](#)). Modeling methods vary in complexity from source proximity models (SPM) ([Section 3.3.2.1](#)) to monitoring data-based methods [LUR ([Section 3.3.2.2](#)) and inverse distance weighting (IDW) ([Section 3.3.2.3](#))] to physics-based models [dispersion models ([Section 3.3.2.4](#)) and CTMs ([Section 3.3.2.5](#))] to microenvironmental exposure models ([Section 3.3.2.6](#)) and are often used to produce exposure concentration surrogates for long-term epidemiologic studies ([Section 3.4.4.2](#)). Previous studies from the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) are described along with newer studies.

3.3.1 Measurements

3.3.1.1 Fixed-Site Monitoring

Fixed-site monitors are sited for the purpose of determining whether attainment goals are met under the Clean Air Act. However, fixed-site monitoring ambient SO₂ concentration data are also often used in epidemiologic studies as a surrogate for exposure to SO₂, as discussed in [Section 3.4.4](#). Methods, errors, and uncertainties regarding measurements made by fixed-site monitors are described in [Section 2.4](#). The effect of errors and

uncertainties due to instrumentation issues depends on epidemiologic study design, as described further in [Section 3.4.4](#). Various uses of these data are possible depending on the design of the epidemiologic study. Short-term (e.g., daily, hourly) data can be used for time-series studies and long-term (e.g., annual average) data for longer term studies. For a given core-based statistical area (CBSA), fixed-site monitors are sited at a fixed location based on the number of people living in the CBSA and the sources of SO₂ emissions (40 CFR 58, Appendix D). Even in CBSAs with multiple monitors, the monitors do not fully capture spatial variability in SO₂ concentration across the study area. Moreover, fixed-site monitors that only log hourly average SO₂ concentration may not fully capture temporal variability; monitors that log 5-minute avg or 5-minute hourly max SO₂ concentration data better characterize temporal variability.

3.3.1.2 Personal and Microenvironmental Monitoring Techniques

Personal and microenvironmental SO₂ monitors have been used in studies characterizing relationships between indoor and outdoor SO₂ concentrations ([Section 3.4.1.2](#)) and relationships between personal exposure to SO₂ and ambient SO₂ concentrations ([Section 3.4.1.3](#)). Additionally, personal monitoring is used infrequently in the epidemiologic studies described in [Chapter 5](#). As described in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), both active and passive samplers have been used to measure personal and microenvironmental SO₂ exposures. The Harvard-Environmental Protection Agency (EPA) annular denuder system is a stationary active sampler initially developed to measure microenvironmental concentrations of particles and acidic gases simultaneously ([Koutrakis et al., 1988](#)); [Brauer et al. \(1989\)](#) modified it to serve as a personal exposure monitor. The system draws air past an impactor to remove particles and then through an annular denuder coated with sodium carbonate to trap SO₂ and other acidic gases. Gases collected within the denuder are extracted with ultrapure water and analyzed by ion chromatography. The detection limit depends on the sensitivity of the ion chromatography analysis as well as the volume of air sampled, and is typically below 1 ppb ([Brauer et al., 1989](#)), with a collection efficiency of 99.3% for a 24-hour measurement ([Koutrakis et al., 1988](#)). Similar denuder-type systems have been used in other microenvironmental monitoring studies [e.g., ([Patterson and Eatough, 2000](#))]. Another active sampler, developed for a scripted personal exposure study in Baltimore, MD, used a hollow glass denuder coated with triethanolamine, with SO₂ detection by ion chromatography ([Chang et al., 2000](#)). At a sampling rate of 100 mL/minute for 1 hour, the detection limit was 62 ppb, resulting in many of the 1-hour SO₂ samples being below the detection limit; see [Section 2.5](#) for a summary of typical ambient SO₂ concentrations. For microenvironmental sampling, Federal Reference Method (FRM) and Federal

Equivalent Methods (FEMs), described in [Section 2.4](#), have also been deployed [e.g., ([Maggos et al., 2016](#); [Bozkurt et al., 2015](#); [Halios et al., 2014](#))].

Passive badge-type samplers have also been developed to eliminate the need for a powered sampling pump. A common version is manufactured by Ogawa USA, Inc. and consists of a cellulose fiber filter coated with triethanolamine ([Ogawa & Co, 2007](#)). SO₂ is detected via ion chromatography with a reported detection limit for a 24-hour sample of 2–6 ppb ([Sarnat et al., 2006](#); [Sarnat et al., 2005](#); [Sarnat et al., 2000](#)). Passive badge samplers can also be combined with active particle samplers to create a multipollutant sampler [e.g., ([Demokritou et al., 2001](#))]. Another type of passive sampler, useful for microenvironmental measurements of SO₂, involves a sampling cassette with a sodium carbonate-soaked glass fiber filter, again analyzed for SO₂ with ion chromatography ([Triche et al., 2005](#)). Passive samplers for measuring SO₂ concentrations are not very sensitive to potential interferants, such as temperature or relative humidity ([Swaans et al., 2007](#)). The cumulative sampling approach and the relatively high detection limit of the passive badges makes them mainly suitable for monitoring periods of 24 hours or greater. This limits their ability to measure short-term daily fluctuations in personal SO₂ exposures. If the passive sampling data below detection limits are censored, then the estimated mean and distribution of the SO₂ concentrations may be biased. Maximum likelihood estimation and bootstrap methods can be employed to estimate unbiased means and data distributions for censored data-sets ([Zhao and Frey, 2006](#); [Frey and Zhao, 2004](#); [Zhao and Frey, 2004](#)).

3.3.2 Modeling

Many models can be used to predict the outdoor concentration of SO₂ across geographic regions or at specific locations of interest where people spend time (e.g., outdoors at homes, schools, workplaces, roadways). The modeled concentration can be used as a surrogate for human exposure to SO₂. Most models do not estimate exposures to ambient SO₂ directly, because those models are not designed to include time-activity patterns and indoor concentrations of ambient SO₂ in various microenvironments. Approaches to modeling exposure concentration described below include source proximity models (SPM), LUR, IDW models, dispersion models, and CTM. These models can be employed at urban, regional, or national scales to estimate daily, or longer, average ambient SO₂ concentrations as an exposure surrogate. Short-term (e.g., daily) ambient SO₂ concentration estimates are needed for ambient SO₂ exposure surrogates in acute exposure assessments, whereas long-term (e.g., annual) ambient SO₂ concentration estimates can be used for ambient SO₂ exposure surrogates in chronic exposure assessments. Additionally, microenvironmental models, which incorporate time-activity

data to estimate exposures directly, are included in this discussion. [Table 3-1](#) provides an overview of the modeling approaches discussed in this section.

Table 3-1 Comparison of models used for estimating exposure concentration or exposure. Factors available in each model are checked.

Factors	Type of Model					
	SPM	IDW	LUR	Dispersion	CTM	Microenvironmental
Distance from source	X	X	X	X	X	X
Emission rate		X	X	X	X	X
Terrain or land use			X	X	X	X
Dispersion				X	X	X
Chemistry				X	X	X
Human activity						X
Inhalation						X

CTM = chemical transport model, IDW = inverse distance weighting, LUR = land use regression, SPM = source proximity models.

3.3.2.1 Source Proximity Models

SPMs provide a simple method to estimate ambient SO₂ concentration as a surrogate for ambient SO₂ exposure concentration. These models calculate the distance from receptors (e.g., homes, schools) to a source of SO₂ emissions (e.g., industrial facilities). It is assumed that ambient SO₂ concentration is some function of distance from the source. SO₂ emitted from a point source is thought to disperse as a meandering plume, such that average ambient SO₂ concentration decreases with distance from the source ([Section 2.6.1](#)). Source proximity models do not account for emissions, stack characteristics, plume dispersion, meteorology, or atmospheric chemistry.

These models do not necessarily account for the effect of stack height to limit ambient SO₂ concentrations in the immediate vicinity of the point source, which produces zero SO₂ concentration close to the source and then a peak ground-level concentration at the location of plume touchdown. [Burstyn et al. \(2008\)](#) avoided the stack height issue by

modeling ambient SO₂ concentration as a function of the inverse distance within 2- and 50-km buffers of each gas plant and oil well. In another study, proximity to source was treated as a Boolean variable as a surrogate for high and moderate ambient SO₂ exposure ([Cambra et al., 2011](#)). Likewise, [Liu et al. \(2012b\)](#) computed relative risk (RR) of respiratory disease using ZIP codes with fuel-fired power plants compared with the reference of ZIP codes without fuel-fired power plants.

SPMs are widely applied for exposure assessments because few input data are required. The main limitation of an SPM is the potential for error in the exposure surrogate because none of the factors other than distance from the source affecting emission rates, dispersion, and photochemical activity of pollutants (e.g., emission rates, atmospheric physics, chemistry, meteorology) are included as model variables [e.g., [Zou et al. \(2009a\)](#)].

To improve the accuracy of SPMs in providing a surrogate for exposure, an emission-weighted proximity model (EWPM) was developed that considers the emission rate and duration of each ambient SO₂ point source, in addition to the distance from the source. The model still does not account for stack characteristics, plume dispersion, meteorology, or atmospheric chemistry. [Zou et al. \(2009b\)](#) evaluated the SPM and EWPM to estimate ambient SO₂ concentrations in Dallas and Ellis counties, TX. Normalized ambient SO₂ concentration estimates based on SPM and EWPM were compared to normalized ambient SO₂ concentration measurements at three monitoring sites. EWPM-based ambient SO₂ concentration estimates agreed more closely to the observed ambient SO₂ concentrations than SPM-based ambient SO₂ concentration estimates. Epidemiologic estimates of risk also were in closer agreement between EWPM and American Meteorological Society/Environmental Protection Agency Regulatory Model (AERMOD) compared with the comparison of results using SPM and AERMOD ([Zou et al., 2011](#)). In addition, surface maps of EWPM- and SPM-predicted ambient SO₂ concentrations across two counties showed that with SPM risk of exposure is usually overestimated in the region of dense emission sources and underestimated where emission sources were sparse ([Zou et al., 2009b](#)).

SPMs provide a simple estimation of SO₂ concentrations based on location information. However, they are fundamentally limited because they are not designed to account for emissions, stack characteristics, plume dispersion, meteorology, or atmospheric chemistry. EWPMs provide an alternative to SPMs that account for emissions for improved prediction but still suffer from most of the limitations of SPMs.

3.3.2.2 Land Use Regression Models

LUR models are used to estimate ambient SO₂ concentration as a surrogate for exposure in some large health studies of long-term SO₂ exposure, because they provide spatial variability in estimates of ambient SO₂ concentration across the geographic area of the study population. A detailed description of LUR models is provided in Chapter 3 of the 2016 ISA for Oxides of Nitrogen ([U.S. EPA, 2016d](#)). Briefly, LUR fits a multiple linear regression model of concentration based on local data (e.g., proximity to SO₂ emissions sources, road length, land use, population density) and then applies that model to locations without monitors to predict the ambient SO₂ concentration field. LUR allows the ability to characterize more completely the spatial variation by predicting at arbitrary locations, compared with methods where data are available at a limited number of points, such as fixed-site monitoring ([Marshall et al., 2008](#)). Metrics used for comparing modeling approaches include spatial scale, averaging time, and out-of-sample coefficient of variation, and root mean squared error (RMSE) as measures of prediction error to cross-validate the model. In-sample coefficient of variation and RMSE are sometimes reported to illustrate the model training error. However, studies in the literature of LUR model results do not consistently report all of these parameters. The discussion of LUR models below includes the metrics provided in specific papers.

Models are typically calibrated using ambient SO₂ concentration data from passive sampler measurements and several local predictor variables. LUR models are typically based on several days, weeks, or years of data and thus do not account well for short-term temporal variability in the ambient SO₂ concentration estimates. Hence, LUR is typically used to estimate air pollution exposure in long-term epidemiologic studies. Although LUR is usually employed for nitrogen dioxide (NO₂), it has also been used to study spatial variability in ambient SO₂ concentration in a small number of studies [e.g., [Atari et al. \(2008\)](#)]. Important methodological issues for interpreting LUR model results include number of measurement sites used to fit the statistical model, predictor variable selection, and comparison of LUR performance among model formulations and with other models. These issues affect how well the spatial variability of ambient SO₂ concentration in a city is represented by the LUR. For example, [Gulliver et al. \(2011\)](#) observed a small fractional bias in the models of -4% to -6%, depending on year, for a retrospective LUR model of annual average SO₂ concentration over the U.K. for years 1962–1991. The authors note that the model tended to overpredict the SO₂ concentrations in rural areas and underpredict in urban areas without a high monitor density. Leave one out cross-validation R^2 declined from 0.71 in 1962 to 0.31 in 1991, which the authors attribute to lower spatial variation in SO₂ concentration in later years. Therefore, LUR is not appropriate for representing average conditions over time periods long enough for source conditions to change.

LUR models have been applied to estimate ambient SO₂ concentrations in close proximity to industrial SO₂ sources. [Atari et al. \(2008\)](#) developed an LUR model to predict ambient SO₂ concentrations in Sarnia, Ontario, Canada, an area known as “Chemical Valley” for its high density of chemical industries. Ambient SO₂ concentrations measured over a 2-week period by passive badge monitors were used to “train” the model, and the explanatory variables for the LUR model were distance to an industrial zone, location within 1,200 m of industrial areas, and location within 100 m of major roads. Measurements of ambient SO₂ concentration for model training were collected with passive samplers at 37 locations across the city for 2 weeks in the fall of 2005, with an average concentration of 3.4 ppb. Although the sampling time was relatively short, the authors suggested that it was sufficient to capture intra-urban variability given the spatial distribution of SO₂ sources. The in-sample coefficient of determination was $R^2 = 0.66$. An out-of-sample coefficient of determination was calculated to cross-validate the model for three separate tests: removing 5, 10, and 50% of the samplers and then predicting the SO₂ concentration at those sites. The out-of-sample coefficient ranged from $R^2 = 0.62$ to $R^2 = 0.73$, and the RMSEs of the out-of-sample predictions were 0.3 to 1 ppb. [Atari et al. \(2008\)](#) attributed this moderate validation to a skewed ambient SO₂ concentration distribution, although skewness metrics were not provided. These findings suggest that LUR simulates the spatial variability of SO₂ from a point source with reasonable accuracy but may not fully capture the distribution of SO₂ data.

LUR has also been applied to predict ambient SO₂ concentrations in the vicinity of urban sources. [Clougherty et al. \(2013\)](#) modeled concentrations of ambient SO₂, NO₂, particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm (PM_{2.5}), and black carbon (BC), measured using passive samplers during 2-week periods at different locations across New York City, NY. Ambient SO₂ concentration was predicted by the reference site mean (partial $R^2 = 0.35$), number of oil-burning units (partial $R^2 = 0.36$), and nighttime population within 1 km (partial $R^2 = 0.06$) to give an overall out-of-sample model fit of $R^2 = 0.77$, where R^2 was based on the comparison between raw ambient SO₂ concentrations and model predictions. Traffic covariates were not included in the model. The study authors thought these findings reflected the presence of large combustion boilers in Manhattan and western Bronx, where ambient SO₂ concentrations were predicted to be highest because sulfur content in residential heating fuel is high. Ambient SO₂ concentration was not influenced by vehicle traffic, unlike the other air pollutants studied. [Beelen et al. \(2007\)](#) modeled ambient SO₂, NO₂, nitric oxide, and black smoke (BS) concentrations as the sum of regional, urban, and local components. For the urban-scale model, land use variables, such as location in a nonrural, urban, or industrial area, were included. For the local-scale model, traffic intensity variables were included. The model produced an in-sample $R^2 = 0.56$. The analysis used

passive sampling data from four 2-week periods in 1999–2000, when diesel fuel contained higher concentrations of sulfur before the fuel standards (66 FR 5002) took effect to reduce sulfur concentrations in diesel fuel for highway vehicles and heavy-duty vehicles. The out-of-sample RMSE was 1.6 ppb for the regional model and 1.2 ppb for the urban model; RMSE was not reported for the local model. Ambient SO₂ concentrations modeled in the [Beelen et al. \(2007\)](#) study were used as exposure estimates in a longitudinal cohort study of vascular damage among young adults [see [Section 5.3.2.5](#) and [Lenters et al. \(2010\)](#)]. [Wheeler et al. \(2008\)](#) applied LUR for a study of ambient SO₂ concentration to estimate exposure in Windsor, Ontario and found that distance to the Ambassador Bridge, housing density, and SO₂ emission sources from Detroit within 3 km were all significant predictors of ambient SO₂ concentration with in-sample $R^2 = 0.69$ and out-of-sample $R^2 = 0.65$. [Wheeler et al. \(2008\)](#) also evaluated LUR performance for predicting ambient SO₂ concentration, measured using passive samplers over 2-week periods for each season, by comparing the LUR results with measurements to estimate air pollutant exposure in Windsor, Ontario. They found that correlation of summer predictions of ambient SO₂ concentrations with those from other seasons was lower, suggesting that photochemistry might not be well represented in the LUR model. These studies illustrate several factors that may affect the accuracy of the LUR for SO₂, including source proximity, source type, and season. Therefore, the individual study characteristics and model validation results must be evaluated to determine if LUR provides an accurate exposure prediction.

Spatial variability in ambient SO₂ concentrations predicted by LUR has been used to estimate inter-individual variability in exposure by assuming the ambient SO₂ concentration modeled at the study participants' homes matched their exposure. Ambient SO₂ concentrations computed using LUR by [Atari et al. \(2008\)](#) were used by [Atari et al. \(2009\)](#) to correlate modeled ambient SO₂ concentrations with individual and community perceptions of odor, by [Oiamo and Luginaah \(2013\)](#) to study whether males and females are affected differently by ambient SO₂ exposure, and by [Oiamo et al. \(2011\)](#) to investigate the relationship between estimated ambient SO₂ exposure and access to a general practitioner. [Kanaroglou et al. \(2013\)](#) used a spatial autocorrelation LUR model to estimate ambient SO₂ concentrations, in which the spatial autocorrelation component of the model's residuals was removed. [Kanaroglou et al. \(2013\)](#) applied the spatial autocorrelation LUR model in the vicinity of an industrial area in Hamilton, Ontario, Canada based on monitoring data obtained on a mobile platform during nonpeak traffic times (10:00 a.m.–4:00 p.m.) for 62 days over a 5-year period (2005–2010) and observed that location and difference between wind direction and direction of the industrial area to the receptor each predicted ambient SO₂ concentration (RMSE = 1.24). Inclusion of autocorrelation did not improve the model substantially. These findings suggest that LUR

captures inter-individual variability of the study population exposed to SO₂ from a point source with reasonable accuracy in these studies.

The LUR studies evaluated indicate that LUR can do a reasonable job of estimating SO₂ exposure concentrations, and LUR is capable of capturing spatial variability of SO₂ concentration well. However, some modeling decisions can detract from LUR modeling accuracy. Use of very long averaging periods (in the example cited above, 30 years) may lead to inaccurate predictions. Likewise, model variables may affect accuracy of predictions. These include source proximity, source type, and season. Therefore, the individual study characteristics and model validation results must be evaluated for each study to determine if LUR provides an accurate prediction of exposure.

3.3.2.3 Inverse Distance Weighting

IDW, in which ambient SO₂ concentration at a receptor point is calculated as the weighted average of ambient SO₂ concentration measured at monitoring locations, has been used to estimate exposure based on ambient SO₂ concentration surfaces. Several recent studies using IDW have been published. The weighting factor is an inverse function of distance between the receptor and the monitor. For example, [Brauer et al. \(2008\)](#) and [MacIntyre et al. \(2011\)](#) estimated exposure to ambient SO₂ and other industrial pollutants within 10 km of point sources using an IDW sum of ambient SO₂ concentration and the three closest monitors within 50 km for application in epidemiologic models ([Clark et al., 2010](#)). Often, the weighting factor is the inverse distance raised to some power, and a higher power is applied to increase the weight on monitors that are closer to the receptor. [Rivera-González et al. \(2015\)](#) applied an inverse-distance-squared weighting and compared the results with a citywide average, use of the nearest monitor, or kriging to develop an ambient SO₂ concentration surface. The results from IDW were correlated with the other citywide averaging, nearest monitor, and ordinary kriging ($r = 0.88-0.97$), and the mean ambient SO₂ concentration estimated with IDW was within 10% of the mean computed with the other methods. [Neupane et al. \(2010\)](#) estimated the ambient SO₂ concentration surface using both bicubic spline interpolation and IDW for a study of long-term exposure to air pollutants and risk of hospitalization for pneumonia in Hamilton, Ontario, Canada in a case-control study design. Bicubic spline interpolation produced a lower mean ambient SO₂ concentration and larger interquartile range (IQR) compared with IDW. Because there is no reference value in these studies, it is difficult to conclude that IDW presents any substantial improvement in prediction accuracy compared with other methods. These findings do indicate that the results of IDW are comparable to averaging and smoothing methods when estimating SO₂ concentration.

3.3.2.4 Dispersion Models

Gaussian dispersion models have been applied to estimate ambient SO₂ concentration as a surrogate for human exposure to SO₂ [e.g., ([Smargiassi et al., 2009](#))]. A detailed description of Gaussian dispersion modeling, along with its strengths and limitations for modeling ambient SO₂ concentrations, can be found in [Section 2.6](#). When used in health studies, evaluation of the model outputs in comparison with monitoring data can help determine the applicability of the model. Adjustment of the model surface using monitor observations to account for biases is another technique to improve population exposure concentration estimates.

[Zou et al. \(2009c\)](#) developed a hybrid modeling system to estimate source-specific ambient SO₂ concentration across space as a surrogate for population exposure to ambient SO₂ in Dallas County, TX. First, an AERMOD dispersion model was run for three source scenarios (vehicle only, industrial only, and combined vehicle and industrial), and kriging interpolation was applied to the modeling results to produce a monthly average ambient SO₂ concentration grid map (100 × 100 m). The population exposure was next estimated by multiplying the ambient SO₂ concentration value modeled by AERMOD and the corresponding population density value for each grid cell (100 × 100 m) and for the three source classifications. The results showed that monthly simulated population SO₂ exposure concentrations were moderately correlated with simulated ambient SO₂ concentrations from vehicle sources ($r = 0.440$) and weakly correlated with ambient SO₂ concentrations from industrial sources ($r = 0.069$). This study used emissions data from the year 2000 before the ultra-low sulfur diesel fuel regulations were enacted.

Lagrangian particle modeling has been used to estimate ambient SO_x concentration as a surrogate for ambient SO_x exposure from specific sources ([Ancona et al., 2015](#)) to study the relationship of long-term exposure to SO_x with mortality for all-causes ([Section 5.5.2.2](#)), cardiovascular disease ([Section 5.3.2.2](#)), and cancer ([Section 5.6.1](#)). The Lagrangian particle model tracks the movement of SO_x as nonreactive parcels (i.e., massless particles), considering SO_x to be a marker of the emission source representing some combination of directly emitted SO₂ and sulfate formed in the atmosphere ([Section 2.3](#)). [Gariazzo et al. \(2004\)](#) compared this type of Lagrangian particle model against ambient SO₂ concentration measurements and observed reasonable agreement, although the observations seemed to lag the modeled ambient SO₂ concentration at times. The results suggest that the model would have provided a reasonable estimate of exposure in the [Ancona et al. \(2015\)](#) study, especially given the long-term nature of the study.

3.3.2.5 Chemical Transport Models

Ambient SO₂ concentrations calculated with CTMs, such as the Community Multiscale Air Quality (CMAQ) model, are sometimes used to estimate human exposure to ambient SO₂ ([Section 2.6](#)). CTMs can be applied in epidemiologic studies of either short- or long-term exposure to ambient SO₂ but are more commonly used in long-term ambient SO₂ exposure studies. Given observed biases in the CTMs [e.g., ([U.S. EPA, 2008c](#))], much attention has been given to bias correction of these models for application in exposure assessment. [Chen et al. \(2014a\)](#) evaluated CMAQ v4.7.1 results for several pollutants and found that ambient SO₂ concentration was underpredicted by roughly a factor of two, but this problem was largely ameliorated through bias correction techniques. Improvements to modeling ambient SO₂-related reactions have been corrected in CMAQ v 5.0.2, so that ambient SO₂ concentrations used for exposure surrogates from this or later versions would have smaller exposure errors.

One major limitation of CTMs for estimating ambient SO₂ concentrations as exposure surrogates is that the grid resolution, typically between 4 and 36 km, can be much larger than the length scale of the meandering plume upon touchdown. This limitation presents the possibility that ambient SO₂ concentrations can be underestimated along the plume path when localized peaks are averaged over space. [Baldasano et al. \(2014\)](#) recognized this limitation and merged HYSPLIT with a CTM simulation of ambient SO₂ and particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm (PM₁₀) transport in the vicinity of a refinery. HYSPLIT models dispersion of pollutants, such as ambient SO₂, as particle trajectories; the Weather Research and Forecast meteorological model is coupled with the particle trajectory model to account for wind speed, wind direction, and atmospheric turbulence. [Ching et al. \(2006\)](#) nested smaller grids (1, 4, 12 km) within larger grids (36 km) to improve spatial variability of the simulation. Similarly, [Karamchandani et al. \(2010\)](#) coupled a plume-in-grid model with CTM that treats dispersion as a Gaussian process with parameters that are set using micrometeorological conditions. Inclusion of subgrid-scale modeling enables calculation of the ambient SO₂ plume at finer spatial scales so that maximum ambient SO₂ concentration, and potentially maximum exposures, can be estimated by the model suite ([Baldasano et al., 2014](#)).

3.3.2.6 Microenvironmental Exposure Models

Microenvironmental exposure models are designed to account for variations in the amount of time people spend in different locations by using time-weighted SO₂ concentrations in each microenvironment (e.g., outdoors; indoors at home, school,

workplace; in-vehicle) for the exposure surrogate. Models such as the Stochastic Human Exposure and Dose Simulation (SHEDS) and Air Pollutants Exposure Model (APEX) have been used occasionally for exposure assessment in epidemiologic studies ([Dionisio et al., 2014](#); [Mannshardt et al., 2013](#); [Chang et al., 2012a](#)), and they are also used for the risk assessment performed as part of the NAAQS review process, as was done for the risk and exposure assessment during the last review of the SO₂ NAAQS ([U.S. EPA, 2009c](#)).

The fundamental principles of stochastic population-based exposure models are described in detail in the 2008 NO_x ISA Annex 3.6 ([U.S. EPA, 2008a](#)). Briefly, the models combine ambient concentration data with information on infiltration into enclosed microenvironments, such as buildings and vehicles (see [Section 3.4.1.1](#)), to estimate microenvironmental concentrations. The models then use demographic variables such as age and sex to select appropriate activity patterns from a database. For the risk assessment done during the last review of the SO₂ NAAQS, the U.S. EPA used the Consolidated Human Activity Database (CHAD), which is described in [Section 3.4.2.1](#) and in the 2016 NO_x ISA ([U.S. EPA, 2016d](#)). Inhalation rates are determined from the level of effort associated with each activity (e.g., sitting, walking, or running). Inhalation rates and microenvironmental concentrations are combined to estimate dose. Depending on the availability of controlled human exposure data, response functions based either on microenvironmental exposure concentrations or inhaled dose are used to characterize expected health effects. For population-level exposure assessments, stochastic exposure models such as SHEDS and APEX estimate the distribution of exposures across the population of interest ([U.S. EPA, 2012c](#); [Burke et al., 2001](#)).

3.3.3 Choice of Exposure Surrogates in Epidemiologic Studies

Epidemiologic studies use a variety of methods to assign a surrogate for ambient SO₂ exposure. Study design, data availability, and research objectives are all important factors when selecting an exposure assessment method. Epidemiologic study design may influence the choice of exposure surrogate. The influence of exposure error on effect estimates from epidemiologic studies of different designs is discussed in detail in [Section 3.4.4](#), [Table 3-2](#) and the following text summarizes various metrics used in epidemiologic studies of ambient SO₂ exposure, appropriate applications for the metrics, and errors and uncertainties that may be associated with the metrics. Elements included on the table are based on the studies described in [Section 3.3.2](#). Elements presented are exposure concentration assignment method (which can include monitor deployment or making predictions using different types of models), description of the method, description of the type of epidemiologic applications where this method has been used in [Chapter 5](#), strengths of the method, limitations of the method, and exposure errors associated with the exposure concentration data collection or model prediction method.

Table 3-2 Summary of exposure assignment methods, their typical uses in sulfur dioxide epidemiologic studies, strengths, limitations, related errors, and uncertainties.

Exposure Concentration Assignment Method	Description	Epidemiologic Application	Strengths	Limitations	Exposure Errors
Measurement Methods					
Fixed-site monitors (Section 3.3.1.1).	An FRM or FEM monitor located at a fixed location to measure ambient SO ₂ concentration.	Short-term community time-series studies: surrogate for ambient SO ₂ exposure concentration of a population within a city.	Ambient SO ₂ concentration measurements undergo rigorous quality assurance.	Measurements of ambient SO ₂ concentration made at a fixed location may differ from an exposed individual's true exposure, and no spatial variation is assumed.	Correlation between outdoor SO ₂ concentrations proximal to the receptors and ambient SO ₂ concentration measurements typically decreases with increasing distance from the monitor, potentially leading simultaneously to decreased precision and to bias towards the null, as increased noise drives the slope towards zero.
		Long-term epidemiologic studies: surrogate for ambient SO ₂ exposure concentration to compare populations among multiple cities.			Potential for bias if ambient SO ₂ concentration at a receptor location is higher or lower than the ambient SO ₂ concentration measured at the monitor; potential for imprecision from assumption of constant SO ₂ concentration within some radius of the monitor.

Table 3-2 (Continued): Summary of exposure assignment methods, their typical use in sulfur dioxide epidemiologic studies, strengths, limitations, and related errors and uncertainties.

Exposure Concentration Assignment Method	Description	Epidemiologic Application	Strengths	Limitations	Exposure Errors
Microenvironmental monitors (Section 3.3.1.2).	Typically an FRM or FEM monitor located in an outdoor or indoor microenvironment to measure ambient SO ₂ concentration.	Panel epidemiologic studies: SO ₂ exposure (e.g., personal or residential samples) within a geographic area.	Ambient SO ₂ concentration measurements undergo rigorous quality assurance.	Instrument expense may make it difficult to perform sampling simultaneously in multiple environments.	High detection limit may lead to bias if appropriate statistical methods are not used for handling biased data; nonambient SO ₂ exposure sampling may lead to bias.
Active personal exposure monitors (Section 3.3.1.2).	Air is pulled through a pump and sampled for ambient SO ₂ concentration using ion chromatography to measure personal SO ₂ exposure.	Panel epidemiologic studies: SO ₂ exposure (e.g., personal or residential samples) within a geographic area.	SO ₂ concentrations are obtained at the site of the exposed person.	High detection limit.	High detection limit and potential for nonambient SO ₂ exposure sampling may lead to bias if appropriate statistical methods are not used for handling biased data.
Passive personal exposure monitors (Section 3.3.1.2).	SO ₂ is captured on a coated filter via passive exposure for a time period to measure a personal or area sample.	Panel studies: ambient SO ₂ exposure within a city or among multiple cities.	SO ₂ concentrations are obtained at the site of the exposed person.	Integrated sample does not allow for time-series analysis; high detection limit.	High detection limit may lead to bias if appropriate statistical methods are not used for handling biased data; nonambient SO ₂ exposure sampling may lead to bias.

Table 3-2 (Continued): Summary of exposure assignment methods, their typical use in sulfur dioxide epidemiologic studies, strengths, limitations, and related errors and uncertainties.

Exposure Concentration Assignment Method	Description	Epidemiologic Application	Strengths	Limitations	Exposure Errors
Modeling Methods					
Source proximity model (Section 3.3.2.1).	Ambient SO ₂ concentrations are estimated from distance of receptor from source.	Long-term epidemiologic studies: surrogate for ambient SO ₂ exposure concentration within a city or among multiple cities or regions.	Few input data required.	Does not account for emission rate and duration, stack parameters, plume dispersion, meteorology, or atmospheric chemistry; over-smoothing based on assumption that ambient SO ₂ concentration is constant for a given distance from the source or based on smoothing function between monitors.	Potential for bias if ambient SO ₂ concentration at a receptor location is higher or lower than the average ambient SO ₂ concentration between the source and receptor; potential for imprecision from overly smoothed SO ₂ concentration.
Emission weighted proximity model (Section 3.3.2.1).	Ambient SO ₂ concentrations are estimated from distance of receptor to pollution source, emission rate, and duration.	Long-term epidemiologic studies: surrogate for ambient SO ₂ exposure concentration within a city or among multiple cities or regions.	Considers emission rate and duration.	Does not account for stack parameters, plume dispersion, meteorology, or atmospheric chemistry; over-smoothing based on assumption that ambient SO ₂ concentration is constant for a given distance from the source or based on smoothing function between monitors.	Potential for bias if ambient SO ₂ concentration at a receptor location is higher or lower than the average ambient SO ₂ concentration between the source and receptor; potential for imprecision from overly smoothed SO ₂ concentration.

Table 3-2 (Continued): Summary of exposure assignment methods, their typical use in sulfur dioxide epidemiologic studies, strengths, limitations, and related errors and uncertainties.

Exposure Concentration Assignment Method	Description	Epidemiologic Application	Strengths	Limitations	Exposure Errors
Land use regression model (Section 3.3.2.2).	Measured ambient SO ₂ concentrations are regressed on local variables (e.g., land use factors), and the resulting model is used to estimate ambient SO ₂ concentrations at specific locations.	Long-term epidemiologic studies: surrogate for ambient SO ₂ exposure concentration, usually across a city but sometimes among multiple cities.	High spatial resolution.	Does not account for emission rates, stack parameters, plume dispersion, or atmospheric chemistry and may account for meteorology only in terms of wind speed and wind direction, depending on model formulation; potential for model misspecification; has limited generalizability, and moderate resources are needed.	Potential for bias if grid is not finely resolved or if the model is misspecified or applied to a location different from where the model was fit.
Inverse distance weighting (Section 3.3.2.3).	Measured ambient SO ₂ concentrations are interpolated to estimate ambient SO ₂ concentration surfaces across regions. IDW uses an inverse function of distance to monitors.	Long-term epidemiologic studies: surrogate for ambient SO ₂ exposure concentration, usually within a city or geographic region.	High spatial resolution.	Does not fully capture spatial variability of ambient SO ₂ concentration among monitors, and does not account for emissions, stack parameters, and dispersion; only accounts for meteorology and chemistry to the extent that it is calibrated to data with similar meteorology and chemistry; over-smoothing based on assumption that ambient SO ₂ concentration is constant for a given distance from the source or based on smoothing function between monitors.	Potential for negative bias if ambient SO ₂ sources are not captured or SO ₂ concentration is overly smoothed; potential for positive bias if SO ₂ deposition or other loss processes; potential for imprecision from overly smoothed SO ₂ concentration.

Table 3-2 (Continued): Summary of exposure assignment methods, their typical use in sulfur dioxide epidemiologic studies, strengths, limitations, and related errors and uncertainties.

Exposure Concentration Assignment Method	Description	Epidemiologic Application	Strengths	Limitations	Exposure Errors
Dispersion modeling (Section 3.3.2.4).	Ambient SO ₂ concentrations at specific locations are estimated from emissions, meteorology, and atmospheric physics.	Long-term epidemiologic studies: surrogate for ambient SO ₂ exposure concentration within a city or geographic region.	High spatial and temporal resolution, accounts for atmospheric physics from local emission sources	Resource intensive, very limited representation of atmospheric chemistry or background SO ₂ concentrations.	Potential for bias where the dispersion model does not capture boundary conditions and resulting fluid dynamics well (e.g., in large cities with urban topography affecting dispersion).
Chemical transport model (Section 3.3.2.5).	Grid-based ambient SO ₂ concentrations are estimated from emissions, meteorology, and atmospheric chemistry and physics.	Long-term epidemiologic studies: surrogate for ambient SO ₂ exposure concentration, sometimes within a city but more typically across a larger region.	Strengths include accounting for stack parameters, emission rates, mixing height, atmospheric stability, meteorology, atmospheric chemistry, and complex terrain.	Limited grid cell resolution (i.e., grid cell length scale is typically 4–36 km and much larger than plume width), resource-intensive, spatial smoothing of local SO ₂ emissions sources.	Potential for bias if grid cells are too large to capture spatial variability of ambient SO ₂ exposures.
Microenvironmental model (e.g., APEX, SHEDS) (Section 3.3.2.6).	Estimates distributions of micro-environmental SO ₂ concentrations, exposures, and doses for populations (e.g., census tracts) based on air quality data, demographic variables, and activity patterns.	Panel epidemiologic studies; no epidemiologic studies cited here use micro-environmental models.	Accounts for variability of SO ₂ exposures across large populations, accounts for different concentrations in different microenvironments, accounts for location-activity information.	Models simulate individuals and their exposures, but they do not represent an actual population.	Potential for bias when the modeled distributions of ambient SO ₂ concentration, indoor:outdoor pollutant ratios, and time-activity patterns differ from the true distributions.

APEX = Air Pollutants Exposure model; FEM = Federal Equivalent Method; FRM = Federal Reference Method; IDW = inverse distance weighting; SHEDS = Stochastic Human Exposure and Dose Simulation; SO₂ = sulfur dioxide.

Time-series epidemiologic studies examine how changes in SO₂ exposure (or a surrogate for SO₂ exposure concentration) are associated with changes in a health outcome over time. Fixed site monitors that measure SO₂ concentration or personal SO₂ monitors that provide an estimate of exposure are often used for that purpose ([Section 3.3.1.1](#)). The advantage of fixed-site monitors is that the monitors undergo rigorous quality assurance, while the limitation is that the ambient SO₂ concentration at the monitor location may differ from an exposed individual's true exposure. Low correlation between SO₂ concentrations at the fixed-site monitor and exposure concentrations at the location of a receptor can lead to decreased precision of the effect estimate and to bias of the effect estimate towards the null.

Personal SO₂ monitors are useful because obtaining SO₂ concentrations at study participants' changing locations allows for an estimate of exposure, but these monitors are limited by a high detection limit ([Section 3.3.1.2](#)). High detection limit and potential for nonambient SO₂ exposure sampling may both lead to bias in effect estimates where the exposure was estimated by personal samples. Panel studies often use personal monitors as well and can suffer from bias for the same reasons.

Long-term average epidemiologic studies examine the influence of SO₂ exposure to a population over a time period of at least a month and often of years. Long-term studies are often used to compare health effect estimates in populations for different cities or to examine the impact of spatial variability of SO₂ concentrations within a city or region. Fixed-site monitors are often used for long-term studies that compare health effect estimates among cities ([Section 3.3.1.1](#)). As for time-series studies, bias and reduced precision can result from differences between SO₂ exposure concentrations and SO₂ concentrations measured at the monitor. Passive monitors have been employed for long-term average epidemiologic studies ([Section 3.3.1.2](#)). The integrated design of passive samplers does not permit time-series analysis and also has a high detection limit, which could lead to bias if appropriate statistical techniques are not applied to the data. Most of the modeling approaches described in [Section 3.3.2](#) are commonly used for studies of the health effects of SO₂ exposure. Data requirements of the models increase with model complexity, from SPM to CTM. The strength of SPM ([Section 3.3.2.1](#)), EWPM ([Section 3.3.2.1](#)), and IDW ([Section 3.3.2.3](#)) is their simplicity. However, the limitation of these models is that they do not account for emissions, dispersion, meteorology, or atmospheric chemistry. This can lead to either positive or negative biases. LUR also does not account for emissions, dispersion, or atmospheric chemistry, but its strength is that it has high spatial resolution, which reduces but does not eliminate the potential for bias in health effect estimates ([Section 3.3.2.2](#)). Model misspecification may also bias long-term health effect estimates using LUR to estimate the exposure concentration. Dispersion models ([Section 3.3.2.4](#)) and CTMs

([Section 3.3.2.5](#)) rely on atmospheric chemistry and physics to develop estimates of SO₂ exposure concentration. Limited representation of the fluid dynamics is a limitation of dispersion models, and spatial smoothing due to grid resolution can be a limitation of CTMs. Both factors can lead to bias in the health effect estimates for long-term studies. Microenvironmental models ([Section 3.3.2.6](#)) incorporate time-activity data to overcome some limitations of spatial smoothing in grid-based models, but they are rarely used in epidemiologic models.

3.4 Exposure Assessment, Error, and Implications for Epidemiologic Inference

This section describes exposure assessment issues related to the use of surrogates for ambient SO₂ exposure in epidemiologic studies that may influence or introduce error into the observed health effect estimate. The section reviews several factors that may influence exposure to SO₂ ([Section 3.4.1](#)) and errors in its estimation ([Section 3.4.2](#)), copollutant relationships that may confound the relationship between SO₂ exposure and health effects ([Section 3.4.3](#)), and how epidemiologic study results may be influenced by these factors and relationships ([Section 3.4.4](#)).

[Section 3.4.4](#) focuses on three types of epidemiologic studies that are discussed frequently in [Chapter 5](#): community time-series studies, long-term cohort studies, and panel studies. *Community time-series studies* assess the daily health status of a population of thousands or millions of people over the course of multiple years by estimating population exposure concentrations across an area using a short monitoring interval (hours to days). In these studies, the community-averaged concentration of an air pollutant measured at fixed-site monitors is typically used as a surrogate for individual or population ambient exposure. *Long-term cohort studies*, such as the American Cancer Society (ACS) cohort study, usually involve hundreds or thousands of subjects followed over several years or decades [e.g., ([Jerrett et al., 2009](#))]. Concentrations are generally aggregated over time and by community to estimate exposures. *Panel studies*, which consist of a relatively small sample (typically tens) of study participants followed over a period of days to months, have been used to examine the health effects associated with short-term exposure to ambient concentrations of air pollutants [e.g., ([Delfino et al., 1996](#))]. Panel studies may also apply a microenvironmental model to represent exposure concentrations for an air pollutant.

Some information presented in this section, including parameters influencing infiltration factors ([Section 3.4.1.1](#)) and activity patterns ([Section 3.4.2.1](#)), can be used in health risk assessment. They are discussed here, based on their inclusion in a small number of

epidemiologic studies ([Dionisio et al., 2014](#); [Mannshardt et al., 2013](#); [Chang et al., 2012a](#)). The more detailed information provides a more comprehensive exposure characterization that can help explain population-level variability in the exposure and health effect estimates ([Mannshardt et al., 2013](#)).

3.4.1 Relationships between Personal Exposure and Ambient Concentration

Several factors influence the relationship between personal SO₂ exposure and ambient SO₂ concentration. Indoor SO₂ concentrations are highly dependent on AER due to the lack of indoor SO₂ sources and the rapid deposition of ambient SO₂ after it penetrates into enclosed microenvironments ([Section 3.4.1.1](#)). Generally, indoor SO₂ concentrations are lower than ambient SO₂ concentrations measured outdoors. Because people spend the bulk of their time indoors ([Section 3.4.2.1](#)), personal SO₂ exposures are often much lower than ambient SO₂ concentrations. For example, [Brown et al. \(2009\)](#) reported the mean winter personal SO₂ exposure concentrations in Boston to be 1.8 ppb, while the ambient SO₂ concentration was 11.3 ppb. Both personal SO₂ exposure concentration and ambient SO₂ concentration were even lower in summer, with mean values of near zero and 3.6 ppb, respectively. The following sections describe studies evaluating AER, relationships between indoor and outdoor SO₂ concentrations, and personal-ambient relationships for SO₂.

3.4.1.1 Parameters Influencing Infiltration Factors

Air Exchange Rate

AER, which is the airflow into and out of a building and is represented by a in the conceptual model presented in [Section 3.2.2](#), influences the rate of entry of ambient SO₂ and hence personal exposure to SO₂, because people spend an average of more than 80% of their time indoors ([Spalt et al., 2015](#); [Klepeis et al., 2001](#)). Several factors affect the AER, including the physical driving forces of the airflows (e.g., pressure differences across the building envelope from wind, indoor-outdoor temperature differences, and mechanical ventilation), building characteristics (e.g., local wind sheltering, tightness of the building envelope), and occupant behavior (e.g., opening windows, operating outdoor-vented fans, adjusting thermostat temperature during heating and cooling seasons). Therefore, substantial spatial and temporal AER variations can occur due to temporal and geographical differences in weather conditions, building characteristics, and occupant behavior. The resulting spatial-temporal variations in ambient SO₂ exposure may help explain possible differences in epidemiologic associations between ambient

SO₂ concentrations and health effects in different U.S. communities ([Baxter and Sacks, 2014](#)).

Field studies indicate that the AER of U.S. residences varies by season and region, with substantial variability among different residences within a region, and variability across regions due to differences in occupant behavior, building age, air conditioning prevalence, and building type. [Yamamoto et al. \(2010\)](#) reported AER measured at residences in Los Angeles, CA, Elizabeth, NJ, and Houston, TX as part of the Relationship Among Indoor, Outdoor, and Personal Air (RIOPA) Study conducted between 1999 and 2001. Among the three cities and across seasons, median AER was 0.71/hour. Regional differences can be seen when breaking the data down by season and location. Median AERs in Los Angeles, Elizabeth, and Houston were 0.87/hour, 0.88/hour, and 0.47/hour. Differences between AER for Houston and AER for Los Angeles and Elizabeth may in part be related to larger home sizes (average home volume was 304 m³ for Houston, compared with 163 m³ in Los Angeles and 252 m³ in Elizabeth). Seasonally, median AER was higher in summer compared to winter in Los Angeles (summer: 1.14/hour; winter: 0.61/hour). However, the opposite pattern occurred in Houston (summer: 0.37/hour; winter: 0.63/hour) and to some degree in Elizabeth (summer: 0.88/hour; winter: 1.07/hour). A similar pattern was reported by [Jiao et al. \(2012\)](#), with lower mean AER in Houston during summer compared to winter, but higher mean summer AER in New York city and central North Carolina. More prevalent use of open windows in Los Angeles, where summertime tends to be less humid than in Elizabeth or Houston, may promote greater air exchange. This difference may grow smaller with the increased prevalence of air conditioning, because air conditioning usage is an important factor in infiltration ([Allen et al., 2012](#)). Low AER values in the mild temperatures of autumn may be due to a diminished “stack effect,” which refers to airflow through the building due to indoor-outdoor temperature differential ([Breen et al., 2014b](#)).

Intra- and inter-home variability in AER was also tested in the RIOPA study ([Yamamoto et al., 2010](#)). Intra-home variability in AER indicated that individual homes’ AER changed considerably between seasons (32, 37, and 37% for Los Angeles, Elizabeth, and Houston, respectively). Inter-home variability also differed substantially for all three cities, with the interquartile range of AER exceeding the median AER consistently across seasons and cities.

Vehicle AERs can be substantially higher than residential AERs, leading to rapid infiltration of on-road pollutants. While on-road SO₂ emissions have declined due to reductions in fuel sulfur content ([Section 2.2.3](#)), high vehicle AER would increase exposure in areas with high ambient SO₂ concentrations. Window position has been

shown to be the most important determinant of vehicle AER, even when windows are only partially open ([Ott et al., 2008](#)). When windows are closed, several other factors can affect vehicle AER, including vehicle make and model, vehicle age, driving speed, and fan/recirculation setting on the vehicle ventilation system. The combined effect of these factors result in AERs that vary by more than two orders of magnitude, from less than 1/hour (approximately equivalent to a typical residential AER) to more than 100/hour ([Hudda et al., 2011](#)). In a model fit to AER measurements on 59 vehicles driven at three different speeds under recirculation conditions, the most important variables were vehicle age, mileage (which is related to age), and speed, plus an adjustment for manufacturer ([Fruin et al., 2011](#)). Fan speed and vehicle shape were not influential variables.

Penetration Factor

Limited information was identified regarding the penetration factor P ([Equation 3-4](#)). [López-Aparicio et al. \(2011\)](#) measured SO₂ concentrations indoors and outdoors at the National Library in Prague, Czech Republic from July 2009 to March 2010 and observed SO₂ penetration values ranging from $P = 0.25$ to 0.74. The authors described the library as naturally ventilated, and it is unclear whether windows were open or closed during the measurement periods. Measured outdoor SO₂ concentrations were higher for the cold months of January, February, and March compared with the remainder of the sampling campaign, and penetration was lower during that period ($P = 0.25$ to 0.48). The literature search only produced this one recent study of SO₂ infiltration.

Deposition Factor

Indoor sources of SO₂ are relatively scarce and SO₂ rapidly reacts with indoor surfaces [see [Grontoft and Raychaudhuri \(2004\)](#) and references cited therein] or oxidizes rapidly via indoor Criegee intermediates [see [Section 2.3.1](#) for a description of Criegee chemistry or [Shallcross et al. \(2014\)](#) for the role of indoor Criegee intermediates in SO₂ losses]. The main indoor source of SO₂ is combustion of sulfur-containing fuels, such as kerosene, which is generally used for an emergency or supplemental source of heat in the U.S. Unlike fireplaces, woodstoves, or gas space heaters, kerosene heaters caused elevated SO₂ concentrations indoors in a study conducted in Connecticut and Virginia ([Triche et al., 2005](#)). The median indoor SO₂ concentration measured by passive sampler over 2 weeks in homes using kerosene heat was 6.4 ppb, compared with 0.22 ppb for homes that did not use kerosene heat in the 2-week period. This relatively low concentration when the kerosene heater was not in use is consistent with the rapid removal rate of infiltrated ambient SO₂. As discussed in [Section 2.3](#), SO₂ is removed from the atmosphere by both dry and wet deposition to surfaces, represented by k in the conceptual

model presented in [Section 3.2.2](#). The deposition rate of SO₂ in apartments in Athens, Greece was found to range from 0.76–4.3/hour, similar to the rate observed for O₃, but an order of magnitude higher than the deposition rate measured for NO₂ ([Halios et al., 2009](#)).

Understanding air exchange rate, penetration, and deposition inform characterization of indoor exposure to ambient SO₂. Given that people generally spend the majority of their time indoors ([Section 3.4.2.1](#)), failure to account for these processes could result in biasing SO₂ exposure estimates and subsequent addition of bias and/or variability in the health effect estimates obtained from epidemiologic studies.

3.4.1.2 Indoor-Outdoor Relationships

A number of studies from the U.S., Canada, Europe, and Asia summarized in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), as well as a few new studies conducted outside the U.S., have characterized the relationship between outdoor and indoor SO₂ concentrations. Ratios and slopes of the indoor SO₂ concentration versus the SO₂ concentration immediately outside the indoor microenvironment had an extremely wide range in the studies described in the 2008 SO_x ISA, from near zero to near unity. One of the most detailed older studies of SO₂ in a school was able to detect an indoor-outdoor slope of 0.02–0.03, with near-zero intercept and a correlation of 0.79–0.91, while measuring indoor concentrations <1 ppb, obtained over 10-hour periods when school was in session and 14-hour periods when the school was vacant ([Patterson and Eatough, 2000](#)). Studies conducted since the 2008 SO_x ISA have focused on public buildings and show generally similar results to older studies. A historic library in Prague without heating or air conditioning had indoor:outdoor ratios of 0.25–0.74 (mean = 0.49) for monthly average outdoor SO₂ concentrations of 1–7 ppb obtained with passive samplers ([López-Aparicio et al., 2011](#)). In Brazil, ratios of average indoor and outdoor SO₂ concentrations from 2-week passive samples were 0.7 and 1.0 for urban and suburban schools, respectively ([Godoi et al., 2013](#)).

Several factors could contribute to the differences observed among studies, including building characteristics (e.g., forced ventilation, building age, and building type such as residences or public buildings), behaviors affecting air exchange rates such as opening windows, indoor deposition of SO₂, and analytical capabilities. When reported, correlations between indoor and outdoor ambient SO₂ concentrations were relatively high (>0.75), suggesting that variations in outdoor ambient SO₂ concentration are driving indoor SO₂ concentrations. These high correlations were observed across seasons and geographic locations. This is consistent with the relative lack of indoor sources of SO₂

([Section 3.4.1.1](#)). For other criteria pollutants, nonambient sources can be an important contributor to total personal exposure, but personal SO₂ exposure is expected to be dominated by ambient SO₂ in outdoor microenvironments and in enclosed microenvironments with high air exchange rates (e.g., buildings with open windows and vehicles).

3.4.1.3 Personal-Ambient Relationships

As discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), personal monitoring studies for SO₂ exposure assessment have frequently found that most SO₂ exposure concentrations are below the detection limit of the personal samplers used in the study. Several studies using passive samplers ([Section 3.4.1.2](#)) found that 95% or more of the personal SO₂ exposure concentrations were less than the field detection limit of 2–6 ppb for 24-h avg samples ([Sarnat et al., 2006](#); [Sarnat et al., 2005](#); [Sarnat et al., 2001](#); [Sarnat et al., 2000](#)). Thus, these data are not suitable for evaluating the relationship between personal exposure and ambient concentration for SO₂.

A study in Boston using a different type of sampler, a personal annular denuder ([Section 3.3.1.2](#)) with a detection limit of 0.19 ppb, found that the slope between 24-hour personal SO₂ exposure concentration and ambient SO₂ concentration was 0.13, with a standard error of 0.02 and zero intercept ([Brauer et al., 1989](#)). The 2008 SO_x ISA reported slopes of 0.03–0.13. Assuming that there are no nonambient sources of SO₂ ([Section 3.2.2](#)), the slope serves as an estimate for α . The R^2 value was 0.43 ($r = 0.66$) in this analysis, which excluded values below the detection limit, indicating that personal SO₂ exposure concentration was moderately correlated with ambient SO₂ concentration.

3.4.2 Factors Contributing to Error in Estimating Exposure to Ambient Sulfur Dioxide

Ambient SO₂ concentrations measured at fixed monitoring sites are commonly used for exposure surrogates in community time-series ([Sections 5.2.1, 5.3.1](#)) and long-term cohort ([Sections 5.2.2, 5.3.2, 5.5.2](#)) epidemiologic studies. As noted in [Section 3.3.1.1](#), use of a fixed-site SO₂ monitor to capture a surrogate for true, likely unobserved ambient SO₂ exposure may lead to exposure error. Factors that may influence this type of error include human activity patterns, spatial and temporal variation in ambient SO₂ concentration, and indoor exposure to ambient SO₂ including variation in infiltration parameters ([Brown et al., 2009](#); [Zeger et al., 2000](#)). Additionally, uncertainty in the

metric used to represent exposure may be influenced by method detection limit, accuracy, and precision of the instrument. These factors are discussed in the following section.

3.4.2.1 Activity Patterns

The activity pattern of individuals is an important determinant of their exposure. Variation in SO₂ exposure concentrations among microenvironments means that the amount of time spent in each location will influence an individual's exposure to ambient SO₂. The effect of activity pattern on exposure is explicitly accounted for in [Equation 3-3](#) by the fraction of time spent in different microenvironments. As discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), although activity patterns vary both among and within individuals, resulting in corresponding variations in exposure across a population and over time, people generally spend more than 80% of their time indoors ([Spalt et al., 2015](#); [Klepeis et al., 2001](#)).

Time spent in different locations has been found to vary by age. [Table 3-3](#) summarizes National Human Activity Pattern Survey (NHAPS) data reported for four age groups, termed very young (0–4 years), school age (5–17 years), working (18–64 years), and retired (65+ years) ([Klepeis et al., 1996](#)). Although this survey represents older data, an uncertainty analysis conducted for the 2014 O₃ REA suggests that historical activity patterns can generally be used to represent current activity patterns ([U.S. EPA, 2014a](#)). The working population spent the least time outdoors, while the school age population spent the most time outdoors. NHAPS respondents aged 65 years and over spent somewhat more time outdoors than adults aged 18–64 years, with a greater fraction of time spent outdoors at a residence. Children aged 0–4 years also spent most of their outdoor time in a residential outdoor location. On average, the fraction of time spent outdoors by school age respondents was 2.62 percentage points higher than working respondents, corresponding to approximately 38 minutes more time outdoors per day. Moreover, in a survey comparing children (mostly less than age 8 years), their parents who were mostly under age 55 years, and adults older than age 55 years, a larger proportion of children reported spending over 30 minutes performing vigorous outdoor physical activity ([Wu et al., 2011b](#)).

Table 3-3 Mean fraction of time spent in outdoor locations by various age groups in the National Human Activity Pattern Survey study.

Age Group (yr)	Residential-Outdoor (%)	Other Outdoor (%)	Total Outdoors (%)
0-4	5.38	0.96	6.34
5-17	5.05	2.83	7.88
18-64	2.93	2.33	5.26
65+	4.48	1.27	5.75

Source: Data from ([Klepeis et al., 1996](#)).

Longitudinal activity pattern information is also an important determinant of exposure, as different people may exhibit different patterns of time spent outdoors over time due to race/ethnicity, age, sex, employment, health status, and lifestyle-dependent factors. [Spalt et al. \(2015\)](#) analyzed the relationship between time-activity patterns and demographic patterns for the Multi-Ethnic Study of Atherosclerosis Air cohort. They found that time spent outdoors was best predicted by employment status, and white, black, or Hispanic study participants were more likely to spend time outdoors compared with participants of Chinese ethnicity. These differences may manifest as higher mean SO₂ exposures or more frequent high-exposure episodes for some individuals. The extent to which longitudinal variability in individuals contributes to the population variability in activity and location can be quantified by the ratio of between-person variance to total variance in time spent in different locations and activities [the intra-class correlation coefficient (ICC)]. [Nethery et al. \(2009\)](#) studied time-activity patterns in a cohort of 62 pregnant women in Vancouver, Canada observed that activity reduced over the course of pregnancy. [Xue et al. \(2004\)](#) quantified ICC values in time-activity data collected by Harvard University for 160 children aged 7–12 years in Southern California ([Geyh et al., 2000](#)). For time spent outdoors, the ICC was approximately 0.15, indicating that 15% of the variance in outdoor time was due to between-person differences. The ICC value might be different for other population groups. [Schwab et al. \(1992\)](#) studied time-activity patterns among fourth- through sixth-grade children in a cohort of 50 children with asthma or persistent wheeze and 50 children not reporting respiratory symptoms. Those with asthma or wheeze reported more time outside on both school and nonschool days.

Several methods are available for sampling diary information, and the method chosen can affect estimated personal SO₂ exposures and related exposure errors. [Che et al. \(2014\)](#) evaluated how diary sampling methods influenced estimates of children's exposure (in this case, to ambient PM_{2.5}). Random resampling, diversity and autocorrelation, and

Markov-chain cluster methods of diary sampling were tested. The three sampling methods provided similar results for total ambient exposure, outdoor ambient exposure, and ambient exposure at homes and indoor locations not including home, school, or vehicles.

The U.S. EPA's National Exposure Research Laboratory has consolidated many of the most important human activity databases into one comprehensive database called CHAD. The current version of CHAD contains data from 22 human activity pattern studies (including NHAPS), which were conducted between 1982 and 2010 and evaluated to obtain over 54,000 person-days of 24-hour human activities in CHAD ([Isaacs, 2014](#); [McCurdy et al., 2000](#)). Five studies conducted between 1997 and 2010 comprising over 30,000 person-days have been added to CHAD since the previous SO_x ISA ([University of Michigan, 2016](#); [Isaacs et al., 2013](#); [Wu et al., 2012](#); [Hertz-Picciotto et al., 2010](#); [Knowledge Networks, 2009](#); [Williams et al., 2009](#)). The surveys include probability-based recall studies conducted by U.S. EPA and the California Air Resources Board, as well as real-time diary studies, telephone interviews, and internet-based surveys conducted nationally and in individual U.S. metropolitan areas using both probability-based and volunteer subject panels. All ages of both sexes are represented in CHAD. The data for each subject typically consist of 1 to 3 days of activity diaries, in which each activity is defined by start time, duration, activity type, and microenvironmental classification (i.e., location). Activities vary from 1 minute to 1 hour in duration, with longer activities being subdivided into clock-hour durations to facilitate exposure modeling. CHAD also provides information on the level of exertion associated with each activity, which can be used by exposure models, including the APEX model, to estimate ventilation rate and pollutant dose ([Section 3.3.2.6](#)).

Time-activity patterns vary both across and within different populations and lifestyles. CHAD has a large set of activity diaries, but it cannot represent all of the variation observed in study area(s) chosen for risk assessment. Previous uncertainty characterizations conducted for the 2014 O₃ REA ([U.S. EPA, 2014a](#)) suggested that activity data for individuals with asthma is comparable to activity data for healthy individuals and similar activity patterns are observed across different regions of the U.S. Other factors that are not accounted for, but could be important influences on exposure, include SES and intra-urban differences in outdoor and other activities ([U.S. EPA, 2014a](#)).

Algorithms for generating longitudinal activity patterns reduce uncertainty that would result from repeatedly sampling CHAD for activities on successive days, given that individuals often have similar day-to-day activity patterns. The method used in the 2014 O₃ REA ([U.S. EPA, 2014a](#)) involves a within-person autocorrelation statistic and a

population diversity statistic, which together can help represent the repeated nature of individual activities while allowing for variability across the population.

To improve the characterization of activity patterns, mobile electronic devices, such as smartphones with embedded global positioning system (GPS) receivers and dedicated GPS data loggers, are increasingly used to collect time-location information. GPS technology has the potential to provide increased resolution in recording activity patterns. For example, [Glasgow et al. \(2014\)](#) analyzed the frequency of positional data collection by Android-based smartphones among a panel of study participants and found that on average 74% of the data were collected over intervals shorter than 5 minutes, which is a marked improvement over many time-activity studies using diaries. However, manual processing of GPS data to determine time spent in different microenvironments is limited due to large (potentially thousands of samples per person per day) and multidimensional (location, speed, time, signal quality) data sets, missing data due to loss of GPS signal reception while inside certain buildings, and difficulty discriminating among certain microenvironments (e.g., steel structures have substantial indoor/outdoor differences in satellite signal strength). To address these limitations, automated microenvironmental classification models have been developed ([Breen et al., 2014a](#); [Kim et al., 2012](#); [Wu et al., 2011a](#); [Adams et al., 2009](#); [Elgethun et al., 2007](#)). For example, [Breen et al. \(2014a\)](#) developed a classification model to estimate time of day and duration spent in eight microenvironments (indoors and outdoors at home, work, school; inside vehicles; other locations) from GPS data and geocoded building boundaries. The classification model estimates were compared with diary data and correctly classified the microenvironment for 99.5% of the daily time spent by the participants. In conjunction with accelerometers, air pollutant monitors, and health monitors, GPS-based time-activity data and related monitors have the potential to reduce error in exposure assessment ([NRC, 2012](#)).

Positional errors are a concern for geographic information systems (GIS) and GPS-based technologies, although post-processing algorithms can compensate for loss of signal (e.g., when inside a steel-frame building) to some degree. [Lane et al. \(2013\)](#) compared three geocoding techniques with aerial photography and observed median positional errors of 7–23 m. [Glasgow et al. \(2014\)](#) also compared smartphone positions with geocoded diary-based locations to test the positional accuracy of the phones. For all data combined, the smartphones had a median positional accuracy of 342.3 m, although this includes any error due to the participant imprecisely recording location information in the diary. When broken down by network, the median positional accuracy varied from 98.0 to 1,168.8 m. [Wu et al. \(2010\)](#) compared several portable GPS devices to aerial photography. Median positional errors were 7.3–20.8 m for indoor measurements taken 3 m from a door or window. For outdoor measurements taken 6.1 m from a window or door, median positional errors were 4.1–16.3 m, and for on-road measurements, median

positional errors were 3.5–5.5 m. [Ganguly et al. \(2015\)](#) compared two automated (GIS-based) geocoding techniques with GPS positional data in Detroit, MI. Median positional errors for two GIS methods were 26 m for both methods in comparison with GPS.

3.4.2.2 Spatial Variability

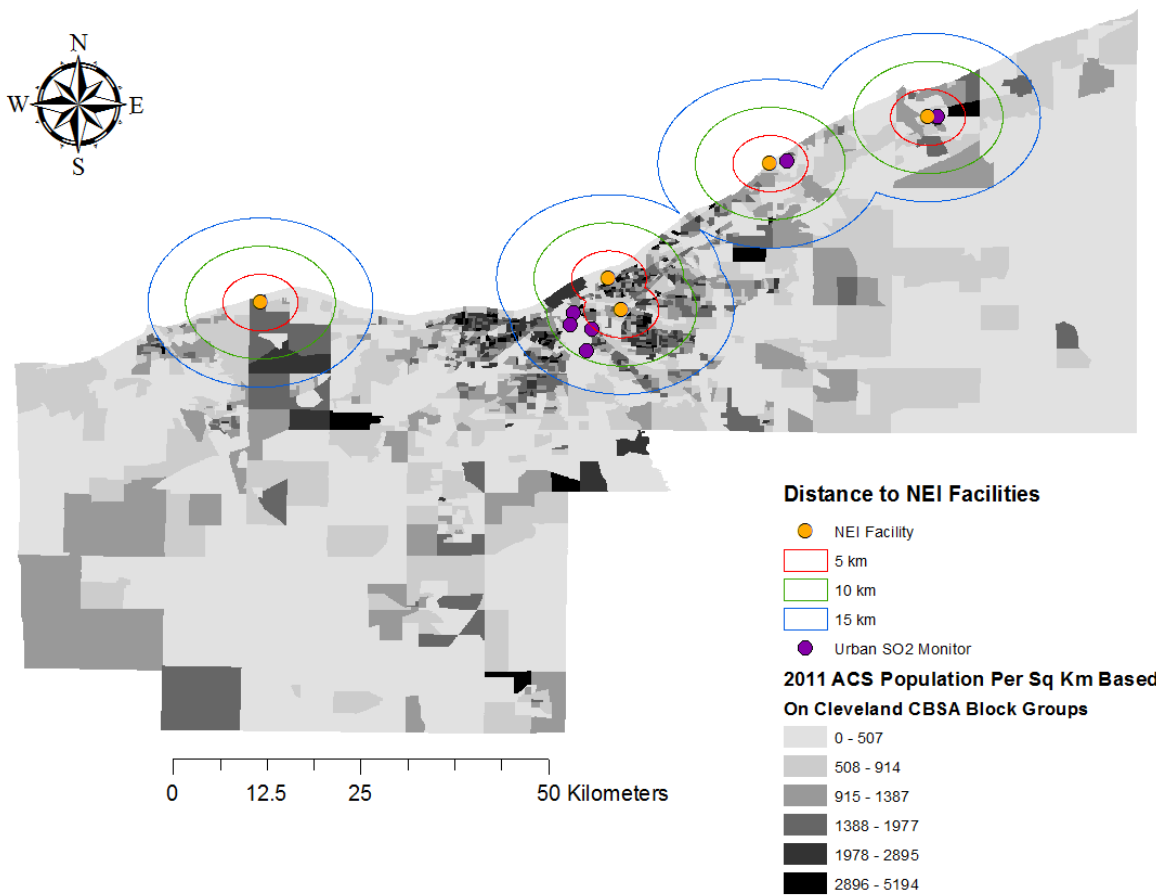
Failure to account for spatial variability in ambient SO₂ concentrations can contribute to exposure error and error in the health effects estimates produced by epidemiologic studies, whether the studies rely on fixed-site monitor data or model output as a surrogate for exposure concentration. Low correlations between the monitor used to measure concentration as an exposure surrogate and the true exposure concentrations at the locations of the study population contribute to exposure error in time-series studies ([Goldman et al., 2010](#)).

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) discussed spatial variability in ambient SO₂ concentrations and the impact of this variability on effect estimates from community time-series epidemiologic studies. Inter-monitor correlations within urban areas ranged from very low to very high values, suggesting that ambient SO₂ concentrations at some monitors may not be highly correlated with the community average SO₂ exposure concentration. Of particular concern for SO₂ is the predominance of point sources, resulting in an uneven distribution of ambient SO₂ concentrations across an urban area. Factors contributing to differences among monitors include the presence of point sources, proximity to point sources, terrain features, and uncertainty regarding the measurement of low ambient SO₂ concentrations. The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) concluded that low correlation between a specific monitor and the community average ambient SO₂ exposure concentration will tend to bias an effect estimate toward the null.

Because ambient SO₂ concentrations can have high spatial variability, average SO₂ exposure concentration estimates may have less error for populations living close to a monitor in a community time-series epidemiologic model. [Figures 3-1](#) and [3-2](#) illustrate proximity of populations and SO₂ monitors to multiple ambient SO₂ sources in the Cleveland and Pittsburgh CBSAs, respectively (discussed in [Chapter 2](#)). These CBSAs were chosen for further discussion here, because they have both high population density and numerous sources above 2,000 tpy. [Figure 3-1](#) shows the location of fixed-site SO₂ monitors and sources with respect to population density for the Cleveland, OH CBSA. Four of the monitors are centrally located in the urban area, and are also within 10 km of SO₂ sources, while two other monitors are located much closer to point sources (<5 km). While some densely populated areas are near fixed-site SO₂ monitors, some of the

highest density census block groups are located more than 10–15 km from fixed-site monitors despite proximity to the sources. [Table 3-4](#) indicates that approximately two-thirds of the population in various age groups lives within 15 km from a fixed-site SO₂ monitor. For the Pittsburgh CBSA ([Figure 3-2](#)), only two of the monitors are located near sources, with the other monitors distributed among population centers and less densely populated areas. Here, approximately 40% of the population lives more than 15 km from a fixed-site SO₂ monitor ([Table 3-5](#)). Such variability in the proximity of populations to fixed-site monitors suggests that some portions of an urban area may be subject to increased exposure error. While only minor differences were noted among age groups in the portion of the population living at specific distances from monitors, the potential exists for exposure error to differ among other potentially at-risk groups due to monitor proximity.

Several recent studies have evaluated the impact of spatial variability in ambient SO₂ concentration on community time-series health effect estimates. [Strickland et al. \(2011\)](#) reported a 0–6% lower RR association [fixed-site monitor: RR = 1.009 (0.992, 1.027); unweighted average: RR = 1.023 (1.006, 1.042); population-weighted average: RR = 1.020 (1.001, 1.039)] per IQR increases in ambient 1-hour SO₂ exposure concentration (from a fixed-site monitor, unweighted average across monitors, and population-weighted average) compared with other criteria pollutants in Atlanta, GA. The authors attributed lower RR to spatial heterogeneity in ambient SO₂ exposure concentrations used as exposure surrogates and the inability of a fixed-site monitor to capture ambient SO₂ plume touchdowns in other parts of the city. The chi-squared statistic moderately increased when average ambient SO₂ exposure concentrations (both population-weighted and unweighted) from monitors across the city were used. Effect estimates were higher for the monitor average metrics than for the fixed-site monitor, and this difference was magnified when effect estimates were based on a standardized increment rather than the IQR. Because the IQR of the data covered the range of values observed across the monitors in Atlanta for the [Strickland et al. \(2011\)](#) study, the IQR partially accounted for spatial variability. The different exposure assignment approaches only altered the magnitude, not direction, of observed associations.



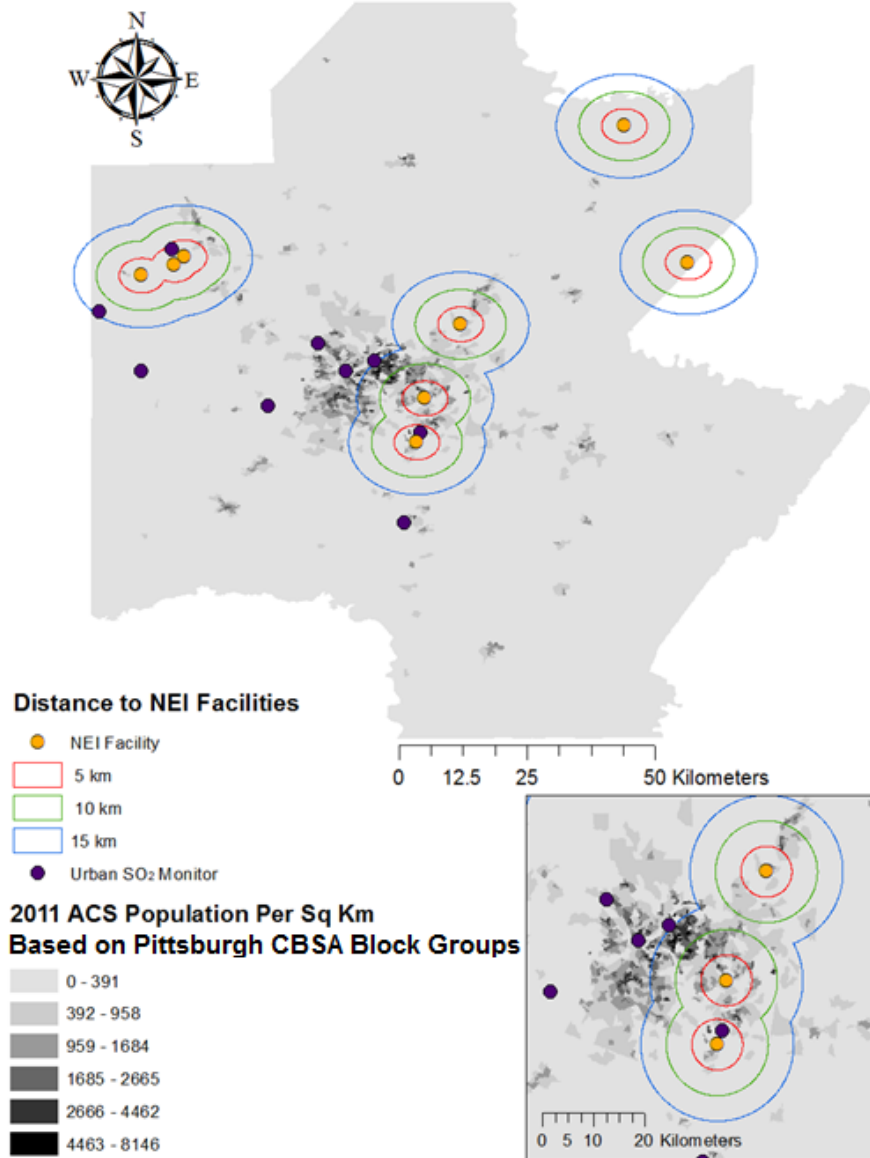
ACS = American Cancer Society; CBSA = core-based statistical area; NEI = National Emissions Inventory; SO₂ = sulfur dioxide. Note that the current map projection (GCS-WGS-1984) creates buffers that take on an elliptical shape instead of a circle. The map projection was chosen to preserve the projection integrity across the data files and reduce error associated with merging data projections.

Figure 3-1 Map of the Cleveland, OH core-based statistical area including National Emissions Inventory facility locations, urban sulfur dioxide monitor locations, and distance to each facility with respect to core-based statistical area block group population density estimates for 2011. National Emissions Inventory facility emissions ranged from 1,942 tons/year to 48,300 tons/year.

Table 3-4 2011 American Community Survey population estimates of people living within a specified distance of an urban sulfur dioxide monitor in the Cleveland, OH core-based statistical area. Population estimates are based on census block group estimates.

Age Group (yr)	Total Population	Within 1 km	Within 5 km	Within 10 km	Within 15 km
Total	2,080,318	11,816	266,777	759,078	1,310,309
≤4	121,820	781	17,608	46,551	75,947
5–17	364,740	1,872	44,719	129,432	222,401
18–64	1,280,478	7,793	178,439	482,808	822,787
≥65	313,280	1,370	26,011	100,287	189,174

Source: Data from the 2011 American Community Survey ([U.S. Census Bureau, 2011](https://www.census.gov)).



ACS = American Cancer Society; CBSA = core-based statistical area; NEI = National Emissions Inventory; SO₂ = sulfur dioxide. Note that the current map projection (GCS-WGS-1984) creates buffers that take on an elliptical shape instead of a circle. The map projection was chosen to preserve the projection integrity across the data files and reduce error associated with merging data projections.

The inset map shows National Emissions Inventory facilities located to the southeast of the highly urbanized areas.

Figure 3-2 Map of the Pittsburgh, PA core-based statistical area including National Emissions Inventory facility locations, urban sulfur dioxide monitor locations, and distance to each facility with respect to core-based statistical area block group population density estimates for 2011. National Emissions Inventory facility emissions ranged from 1,279 tons/year to 46,467 tons/year.

Table 3-5 2011 American Community Survey population estimates of people living within a specified distance of an urban sulfur dioxide monitor in the Pittsburgh, PA core-based statistical area. Population estimates are based on census block group estimates.

	Total Population	Within 1 km	Within 5 km	Within 10 km	Within 15 km
Population (yr)	2,357,769	64,224	494,382	1,076,465	1,428,871
≤4	121,101	2,646	24,748	56,178	73,853
5–17	358,500	8,641	65,882	152,858	211,204
18–64	1,471,310	41,989	325,041	683,445	897,459
≥65	406,858	10,948	78,711	183,984	246,355

Source: Data from the 2011 American Community Survey ([U.S. Census Bureau, 2011](https://www.census.gov/data/tables/2011/acs/2011-01-01)).

High spatial and temporal variability in ambient SO₂ concentration leading to a null-biased effect estimate was also observed in Atlanta by [Goldman et al. \(2010\)](#) when using 1-h daily max SO₂ concentration as an exposure surrogate in a community time-series epidemiologic study. In this study, the authors used a semivariance analysis incorporating both spatial and temporal variability to show that secondary pollutants such as PM_{2.5} and O₃ have lower exposure error (where ambient concentration is a surrogate for exposure) than primary pollutants such as carbon monoxide (CO) and SO₂, for which concentrations tend to have higher spatial variability than those of secondary pollutants. [Goldman et al. \(2010\)](#) simulated exposure error as the difference between concentration measured at the fixed-site monitor and the concentration estimated at a receptor's location. The study authors computed a semivariance term over distance to the fixed-site monitor to concentration at a distance from the monitor. The estimated error for SO₂ was then added to a base case scenario, in which the authors assumed that the fixed-site monitor would produce an accurate exposure. Both the fixed-site monitor estimate and the estimate at the receptor location were used in epidemiologic models to estimate the risk ratio for cardiovascular emergency department (ED) visits. The authors estimated that the risk ratio was biased towards the null by approximately 60% when estimating exposure using the fixed-site monitor in lieu of estimating exposure at the receptors' locations. In a related study, [Goldman et al. \(2012\)](#) used different methods to obtain the surrogate for exposure: fixed-site monitor, unweighted average across monitors, population-weighted average across monitors, and area-weighted average across monitors. The bias decreased for 1-h daily max SO₂ when using unweighted, population-weighted, and area-weighted averages of concentrations from multiple

monitors for the exposure estimate compared with using concentration from a fixed-site monitor for the exposure estimate. Similarly, epidemiologic studies in the U.S. ([Kumar, 2012](#); [Morello-Frosch et al., 2010](#)) and Australia ([Jalaludin et al., 2007](#)) found higher associations between ambient SO₂ concentrations (used as exposure surrogates) and birth outcomes when the analysis was restricted to mothers matched with an ambient SO₂ monitor within 3–5 km of their residence, suggesting bias towards the null remained in the spatial averages used in the base case ([Section 5.4](#)).

3.4.2.3 Temporal Variability

The influence of plume dynamics on human exposures is important for considering results of time-series studies of ambient SO₂ exposure. As described in [Section 2.5.4](#), peak concentrations within the ambient SO₂ plume can exceed concentrations averaged over an hour by up to a factor of five; for the observations made in this assessment, the peak was observed to exceed the mean by up to a factor of 5.5. Hence, SO₂ fixed-site monitoring with averaging times of 1 hour or 1 day, commonly used in time-series epidemiologic studies as an exposure metric ([Chapter 5](#)), may fail to characterize the variability and peak SO₂ exposure concentrations associated with a meandering plume, resulting in exposure error. Moreover, controlled human exposure studies have demonstrated health effects at 5-minute time scales ([Chapter 5](#)). The longer averaging times used in epidemiologic studies may be misaligned with the critical time window of the health effect corresponding to peak SO₂ exposure.

Most of the community time-series epidemiologic studies on the health effects of ambient SO₂ exposure described in [Chapter 5](#) use 24-h avg concentration as a surrogate for exposure. Correlations among different temporal aggregations (1-h avg vs. 5-minute hourly max, 24-h avg vs. 1-h daily max, and 24-h avg vs. 5-minute daily max) were computed from the Air Quality System (AQS) data presented in [Section 2.5.4](#) to provide an indication of how well the 24-h avg represents the 1-h daily max and 5-minute daily max measures that correspond to peak SO₂ plume exposure ([Figure 3-3](#)). Approximately 75% of correlations between 1-h avg and 5-minute hourly max were above 0.9. Correlations between 24-h avg and 1-h daily max were slightly lower, with roughly 75% of the data having correlations above 0.75. A larger range of data was observed for the correlations between 24-h avg and 5-minute daily max, with 75% of the data having correlations above 0.60 and more than 50% of the data having correlations above 0.70. These moderate-high correlations suggest that 24-h avg data used in many time-series epidemiologic studies capture the peak exposure reasonably well, but exceptions may be found for specific sites, as suggested by the lower outliers ($r < 0.35$) and lower whisker ($r < 0.6$) of the correlation between 24-h avg and 5-minute daily max data.

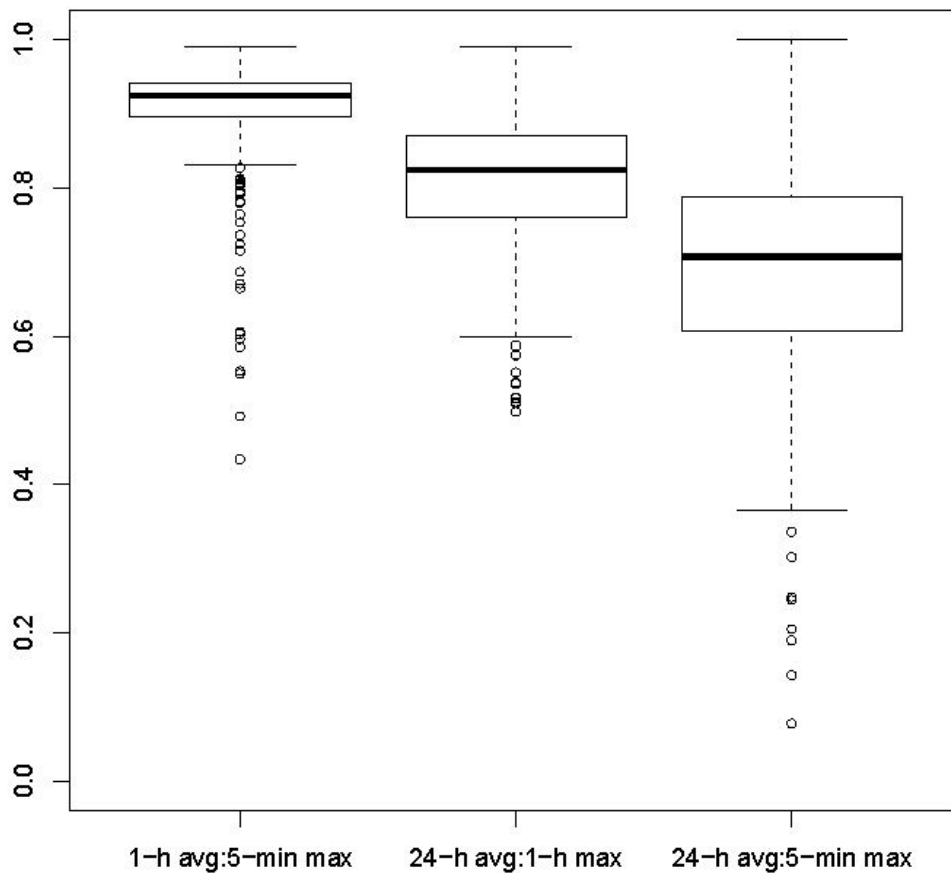


Figure 3-3 Pearson correlations between 1-h avg and 5-minute hourly max, 24-h avg and 1-h daily max, and 24-h avg and 5-minute daily max sulfur dioxide concentrations.

A study in Canada suggests that ambient SO₂ concentration measured over a single year can represent ambient SO₂ exposure concentration over a multidecade period. The authors compared measurement methods used to represent long-term SO₂ exposure concentration and found that the annual average ambient SO₂ exposure concentration in the census tract of a subject's residence during 1980 and 1994 was well correlated (Pearson $R = 0.83$ and 0.85 for all subjects, respectively) with an ambient SO₂ exposure concentration metric accounting for movement among census subdivisions during 1980–2002, despite likely decreases in emissions and concentrations over this period ([Guay et al., 2011](#)). This result may have been due in part to a relatively low rate of movement, with subjects residing on average for 71% of the 22-year period in the same census subdivision they were in during 1980. The percentage of study participants staying in the same place was thought to be similar for those living in urban and rural

areas. [Guay et al. \(2011\)](#) also found that coverage of the study population reduced from 40% for the fixed-time exposure assignments, to 31% when averaging fixed-time exposure assignments with exposure assignments based on census subdivision, to 29% when assigning exposures based only on census subdivision, suggesting that improved spatial and temporal resolution in long-term studies may come at the expense of data completeness. It should also be recognized that there may be differences in population mobility between the U.S. and Canada that could influence these findings.

3.4.2.4 Method Detection Limit, Instrument Accuracy, and Instrument Precision

Personal SO₂ exposure measurements with ambient SO₂ concentration typically have correlations of $0.4 < r < 0.9$ when personal SO₂ exposure measurements are above the method detection limit (MDL). However, the magnitude of personal SO₂ exposure measurements is often much lower than the magnitude of ambient SO₂ concentrations [[Section 3.4.1.3](#); ([U.S. EPA, 2008d](#))]. Moderate to high correlation indicates that using ambient concentration as a surrogate for personal exposure captures the variability needed for epidemiologic studies, particularly for time-series and panel studies. Low personal-ambient correlations reported in the literature are strongly influenced by low personal exposures relative to the detection limits of personal samplers. When this happens, personal samplers are unable to provide a signal to correlate with variations in ambient concentration. Low correlations ($r < 0.4$) in situations with a high proportion of samples below the detection limit should not be interpreted as evidence for the lack of a relationship between personal exposure and ambient SO₂ concentrations. Instead, a low personal sample value likely represents a true low exposure and thus appropriately leads to a low personal:ambient ratio. Low personal:ambient ratios may be due to low penetration and high deposition of SO₂ in indoor microenvironments where people spend most of their time. In a study of personal:ambient exposure ratios by [Brown et al. \(2009\)](#), the authors cited personal SO₂ samples below MDL and extremely low SO₂ levels to rationalize not pursuing further analysis.

Instrument error occurs when the measured SO₂ concentrations are subject to interferences that cause biases or noise leading to error in estimating exposure. Ambient SO₂ concentrations measured by FRM or FEM are subject to positive bias from the detection of interfering compounds. See [Section 2.4.1.2](#) for details on errors that affect FRMs and FEMs used for fixed-site monitoring. Inter-monitor comparison is often used to estimate instrument precision. [Goldman et al. \(2010\)](#) used a simulation to investigate the influence of instrument precision error at locations where ambient SO₂ fixed-site monitors were collocated. Instrument precision error increased with increasing ambient concentration for the fixed-site monitors. When instrument error and ambient SO₂

concentration were correlated, error was larger in locations with more prevalent or stronger sources or at times when SO₂ emissions were higher for a given location. For example, the magnitude of the instrument error was expected to be largest when SO₂ emissions were highest, such as during peak energy usage times. Instrument error was also observed to exhibit some autocorrelation at 1- and 2-day lags in the [Goldman et al. \(2010\)](#) simulation. Hence, the diurnal variability in relative SO₂ instrument error does not change substantially from day to day. For epidemiologic studies of short-term SO₂ exposure that use fixed-site-monitored ambient SO₂ concentration as a surrogate for exposure, instrument error would not be expected to influence the exposure surrogate on a daily basis. When comparing health effect estimates among cities for an epidemiologic study of long-term SO₂ exposure, differences in instrument error among cities could lead to biased exposure surrogates, given the reliance on differences in magnitude of the exposure surrogate to study spatial contrasts. [Section 3.4.4](#) describes the influence of instrument error and high MDL on exposure error and health effect estimates for community time-series ([Section 3.4.4.1](#)), long-term average ([Section 3.4.4.2](#)), and panel ([Section 3.4.4.3](#)) epidemiologic studies.

3.4.3 Copollutant Relationships

Confounding is described in the Preamble to the ISAs ([U.S. EPA, 2015b](#)). Briefly, confounding occurs when the copollutant exposure concentrations are correlated with those of the pollutant of interest and the health effect. Confounding can cause misleading results for estimating the health effect of SO₂ if the study does not account for the copollutant ([Rothman and Greenland, 1998](#)). This differs from effect modification, where the health effect estimate for SO₂ is conditional upon the copollutant exposure concentration via interaction of the SO₂ and copollutant exposures.

Health effect estimates derived from both time-series and long-term average epidemiologic studies are subject to confounding if the health effect model does not account for copollutant correlation. Simulations by [Zeger et al. \(2000\)](#) indicate that unaccounted correlation among exposure concentrations or exposure errors for copollutants may lead to bias and uncertainty in the health effect estimates obtained in community time-series epidemiologic studies using fixed-site monitoring data. Positive correlation among copollutant exposure concentrations was shown to amplify the health effect estimates in [Zeger et al. \(2000\)](#), while negative correlation attenuated the health effect estimates. Correlation of the errors in measuring copollutant concentrations may add bias to the health effect estimate, especially when one is measured with more error than the other ([Zeger et al., 2000](#)). Similarly, simulations of spatial confounding among correlated copollutants for a long-term average epidemiologic study showed that the

health effect estimate was attenuated when copollutants were omitted from the model ([Paciorek, 2010](#)). In some cases, this could promote a false conclusion about the strength of an association between a health effect and the copollutant exposure concentrations.

To assess the independent health effects of ambient SO₂ exposure in an epidemiologic study, it is necessary to identify (1) measurement error for copollutants; (2) which copollutants [e.g., NO₂, PM_{2.5}, ultrafine particulate matter (UFP), BC] are potential confounders of the health effect-SO₂ relationship so that their correlation and collinearity with SO₂ can be tested and, if needed, accounted for in the epidemiologic model; (3) appropriate time lags for SO₂ and copollutants; and (4) the spatial correlation structure across multiple pollutants, if the epidemiologic study design is for long-term exposure ([Paciorek, 2010](#); [Bateson et al., 2007](#)). Additionally, confounding can also vary by the health endpoint studied.

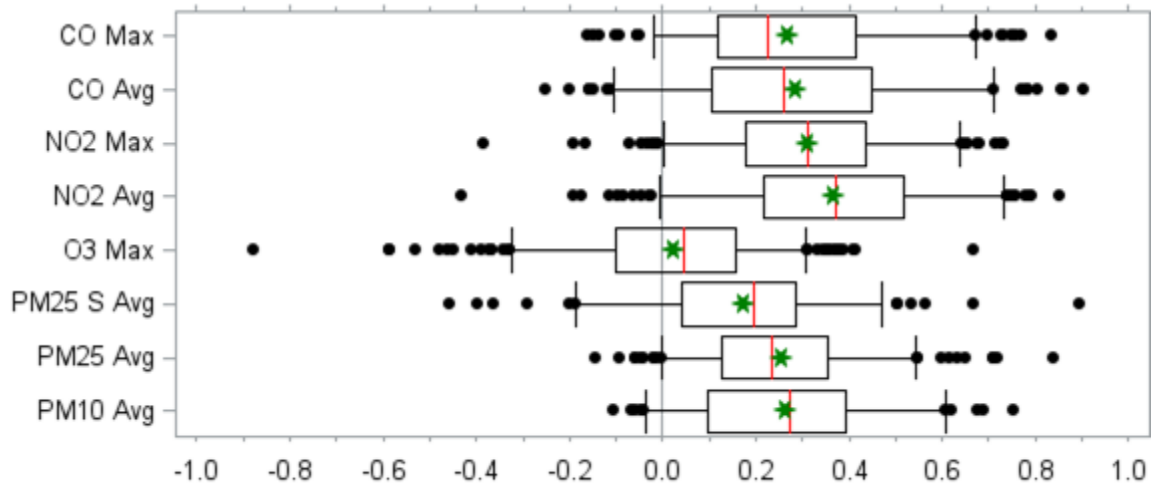
In many cases, correlation of SO₂ with copollutants, such as PM_{2.5} and O₃, is sufficiently low to assume that any health effects identified in SO₂ models are independent of other air pollutants. However, when SO₂ and a copollutant are correlated (for example, SO₂ and NO₂ are often moderately correlated, [Section 5.2](#) and [Figure 3-8](#)), copollutant epidemiologic models may be used to adjust the SO₂ effect estimate for potential confounding by the copollutant ([Tolbert et al., 2007](#)). Two-pollutant models can help identify which is the better predictor of the effect, particularly if the etiologically linked pollutant is measured with more error than the other pollutant ([Zeger et al., 2000](#)). However, collinearity potentially affects the epidemiologic model's effect estimate when highly correlated pollutants are modeled simultaneously, and differences in the spatial distribution of ambient SO₂ concentration and the copollutants' ambient concentrations may also complicate model interpretation [[Section 5.1.2.1](#) and [Gryparis et al. \(2007\)](#)]. Because ambient SO₂ exhibits a relatively high degree of exposure error compared with other criteria pollutants [e.g., [Section 3.4.4.1](#); [Goldman et al. \(2010\)](#)], time-series studies using two-pollutant models in which the SO₂ effect estimate remains robust may provide additional support for a health effect to be associated with SO₂ exposure [e.g., [Ito et al. \(2007\)](#)].

This section considers temporal copollutant correlations and how relationships among copollutants may change in space using AQS data and data reported in the epidemiologic literature ([Chapter 5](#)). Temporal copollutant correlations are computed from the time series of ambient concentrations for two copollutants measured with collocated AQS monitors. Spatial relationships are computed from the set of ambient concentrations across space for two copollutants at a point in time. The following sections review coexposures that can potentially confound the relationship between a health effect and ambient SO₂ exposure over different temporal and spatial resolutions.

3.4.3.1 Temporal Relationships among Ambient Sulfur Dioxide and Copollutant Exposures

Short-Term Temporal Correlations

Short-term copollutant correlations were studied using collocated air quality data reported within the U.S. EPA AQS repository system during 2013–2015. Of the sites reporting SO₂ data to AQS, 438 sites met the 75% data completeness criteria presented in [Section 2.5.1](#). Daily air quality metrics representing either 1-h daily max or 24-h avg ambient SO₂ concentration values were used. Pearson correlations were used to evaluate temporal correlations among ambient SO₂ concentrations and NAAQS copollutant concentrations at the same location. In addition, correlations between ambient SO₂ and PM_{2.5}-sulfur were examined because PM_{2.5}-sulfur serves as a surrogate for SO₂ oxidation products (i.e., sulfate) and may have confounding effects on health outcomes associated with ambient SO₂ exposure. [Figures 3-4](#) and [3-5](#) display the distribution of correlations between NAAQS copollutants and SO₂ daily metrics (24-h avg, 1-h daily max) for all data combined, and [Figures 3-6](#) and [3-7](#) display those copollutant correlations broken down by season. Because epidemiologic studies may use either daily average or daily maximum metrics, correlations are presented for both metrics, when available. 1-h daily maximum concentrations are used for CO and NO₂, while 8-h daily max concentrations are used for O₃.

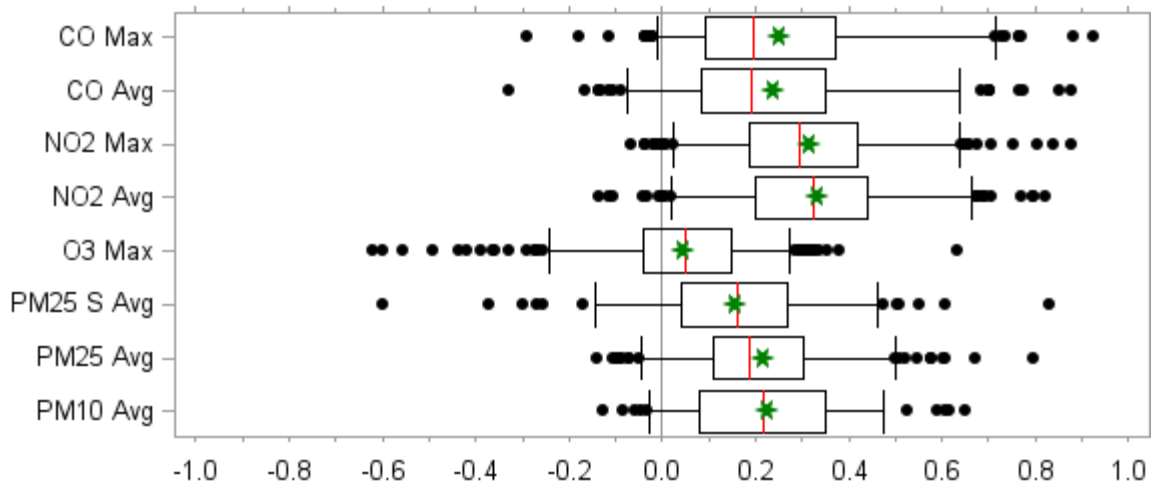


CO = carbon monoxide; NO₂ = nitrogen dioxide; O₃ = ozone; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; S = sulfur.

Notes: Shown are the median (red line), mean (green star), and interquartile range (box), 5th and 95th percentile (whiskers) and extremes (black circles).

Number of monitoring sites by pollutant: SO₂: 438, CO: 171, NO₂: 206, O₃: 310, PM₁₀: 110, PM_{2.5}: 214, PM_{2.5} S: 137.

Figure 3-4 Distribution of Pearson correlation coefficients for comparison of 24-h avg sulfur dioxide concentration from the year-round data set with collocated National Ambient Air Quality Standards pollutants (and sulfur in PM_{2.5}) from Air Quality System during 2013–2015.

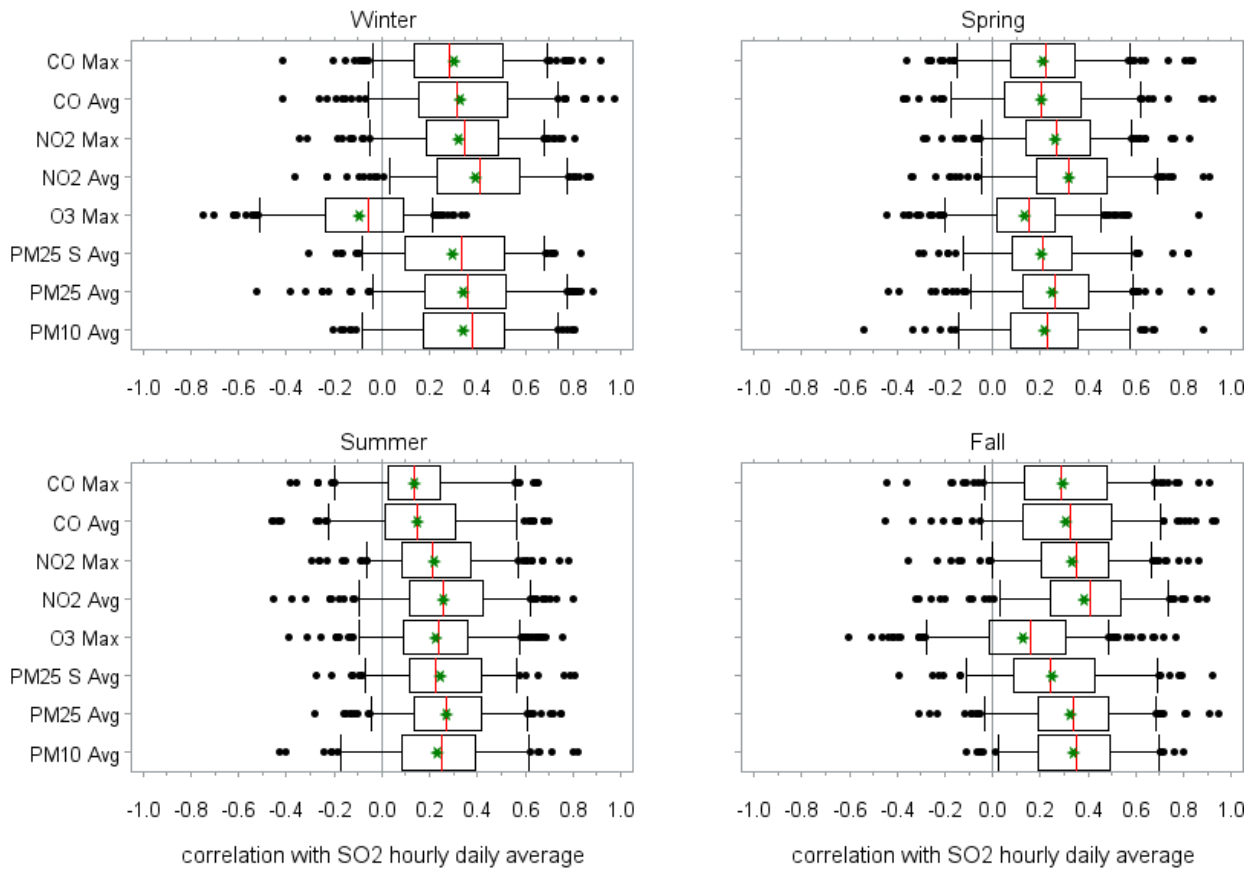


CO = carbon monoxide; NO₂ = nitrogen dioxide; O₃ = ozone; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; S = sulfur.

Notes: Shown are the median (red line), mean (green star), and interquartile range (box), 5th and 95th percentile (whiskers) and extremes (black circles).

Number of monitoring sites by pollutant: SO₂: 438, CO: 171, NO₂: 206, O₃: 310, PM₁₀: 110, PM_{2.5}: 214, PM_{2.5} S: 137.

Figure 3-5 Distribution of Pearson correlation coefficients for comparison of daily 1-h max sulfur dioxide concentration from the year-round data set with collocated National Ambient Air Quality Standards pollutants (and sulfur in PM_{2.5}) from Air Quality System during 2013–2015.

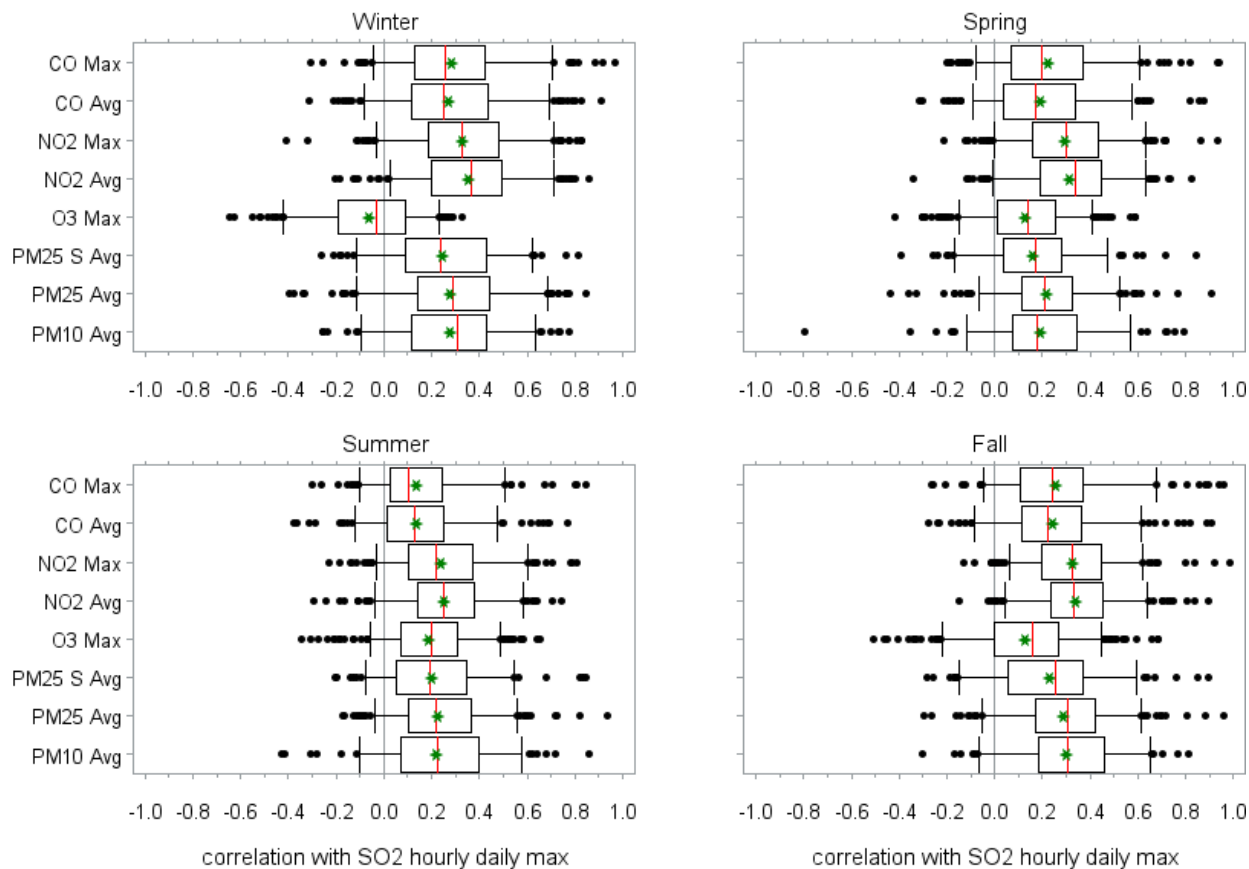


CO = carbon monoxide; NO₂ = nitrogen dioxide; O₃ = ozone; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; S = sulfur; SO₂ = sulfur dioxide.

Notes: Shown are the median (red line), mean (green star), and interquartile range (box), 5th and 95th percentile (whiskers) and extremes (black circles).

Number of monitoring sites by pollutant: SO₂: 438, CO: 171, NO₂: 206, O₃: 310, PM₁₀: 110, PM_{2.5}: 214, PM_{2.5} S: 137.

Figure 3-6 Distribution of Pearson correlation coefficients for comparison of daily 24-h avg sulfur dioxide ambient concentration stratified by season with collocated National Ambient Air Quality Standards pollutants (and PM_{2.5}) from Air Quality System during 2013–2015.



CO = carbon monoxide; NO₂ = nitrogen dioxide; O₃ = ozone; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; S = sulfur; SO₂ = sulfur dioxide.

Notes: Shown are the median (red line), mean (green star), and interquartile range (box), 5th and 95th percentile (whiskers) and extremes (black circles).

Number of monitoring sites by pollutant: SO₂: 438, CO: 171, NO₂: 206, O₃: 310, PM₁₀: 110, PM_{2.5}: 214, PM_{2.5} S: 137.

Figure 3-7 Distribution of Pearson correlation coefficients for comparison of daily 1-h max sulfur dioxide ambient concentration stratified by season with collocated National Ambient Air Quality Standards pollutants (and PM_{2.5}) from Air Quality System during 2013–2015.

While 24-h avg ambient SO₂ concentration exhibits a wide range of correlations with NAAQS copollutants, median correlations are all below 0.4 (Figure 3-4). The lowest correlations are observed between ambient SO₂ concentration and ambient O₃ concentration, with median correlations below 0.1. Slightly higher correlations are observed between ambient SO₂ concentration and other primary NAAQS pollutant concentrations (NO₂ and CO), with median correlations between 0.3 and 0.4. Common fuel combustion sources may be responsible for these correlations (Section 2.2). Lower correlations with PM_{2.5} sulfur than PM_{2.5} mass may reflect the secondary formation of

sulfate by oxidation of SO₂, while PM_{2.5} mass also has a primary component. Correlations close to 1 or below 0 are sometimes observed but only occur at a few outlier monitoring sites. Comparatively, copollutant correlations of daily 1-h max ambient SO₂ in [Figure 3-5](#) are also slightly lower than the copollutant correlations based on ambient SO₂ 24-h avg values in [Figure 3-4](#). The medians of correlations between daily 1-h max ambient SO₂ concentrations and other NAAQS pollutants are below 0.3, with the exception of NO₂, which exhibits median correlations slightly above 0.3. Less than 5% of the data had correlations above $R = 0.7$. Higher correlations may introduce a greater degree of confounding into results of short-term epidemiologic studies. It is notable that the nature of correlations between SO₂ and copollutants is changing given rulemaking on use of ultra-low sulfur diesel fuel that went into effect in 2006 (66 FR 5002). Some of the epidemiologic studies cited in [Chapter 5](#) included data obtained before 2006 and 2007, when the new sulfur standards took effect for highway vehicles and heavy-duty vehicles, respectively. Between 2004 and 2014, SO₂ from highway emissions decreased by 86% ([Table 2-1](#)). Because on-road vehicles are the largest source for ambient NO_x ([U.S. EPA, 2016d](#)) and CO ([U.S. EPA, 2010b](#)) and contribute to ambient organic carbon (OC) and elemental carbon (EC) ([U.S. EPA, 2009a](#)), the new sulfur standards may have contributed to the wider variation observed in correlation between ambient SO₂ and copollutant concentrations. Note that potential for confounding also varies by health endpoint.

Correlations between ambient SO₂ and NAAQS copollutant concentrations demonstrate very little variability across seasons ([Figures 3-6](#) and [3-7](#)). All median and average copollutant correlations are below 0.4 across every season. The only substantial seasonal difference in correlations between ambient SO₂ and copollutant concentrations occurs during the winter, when ambient SO₂ concentration exhibits lower negative correlations with ambient O₃ concentration (median winter correlations = -0.1). SO₂-O₃ correlations are generally low year-round, potentially because the regional nature of O₃ formation contrasts with the local nature of SO₂ plumes from point sources. In winter, the low correlations could be directly linked to relatively low ambient O₃ concentrations during this time of year due to less photochemical O₃ production and SO₂ oxidation.

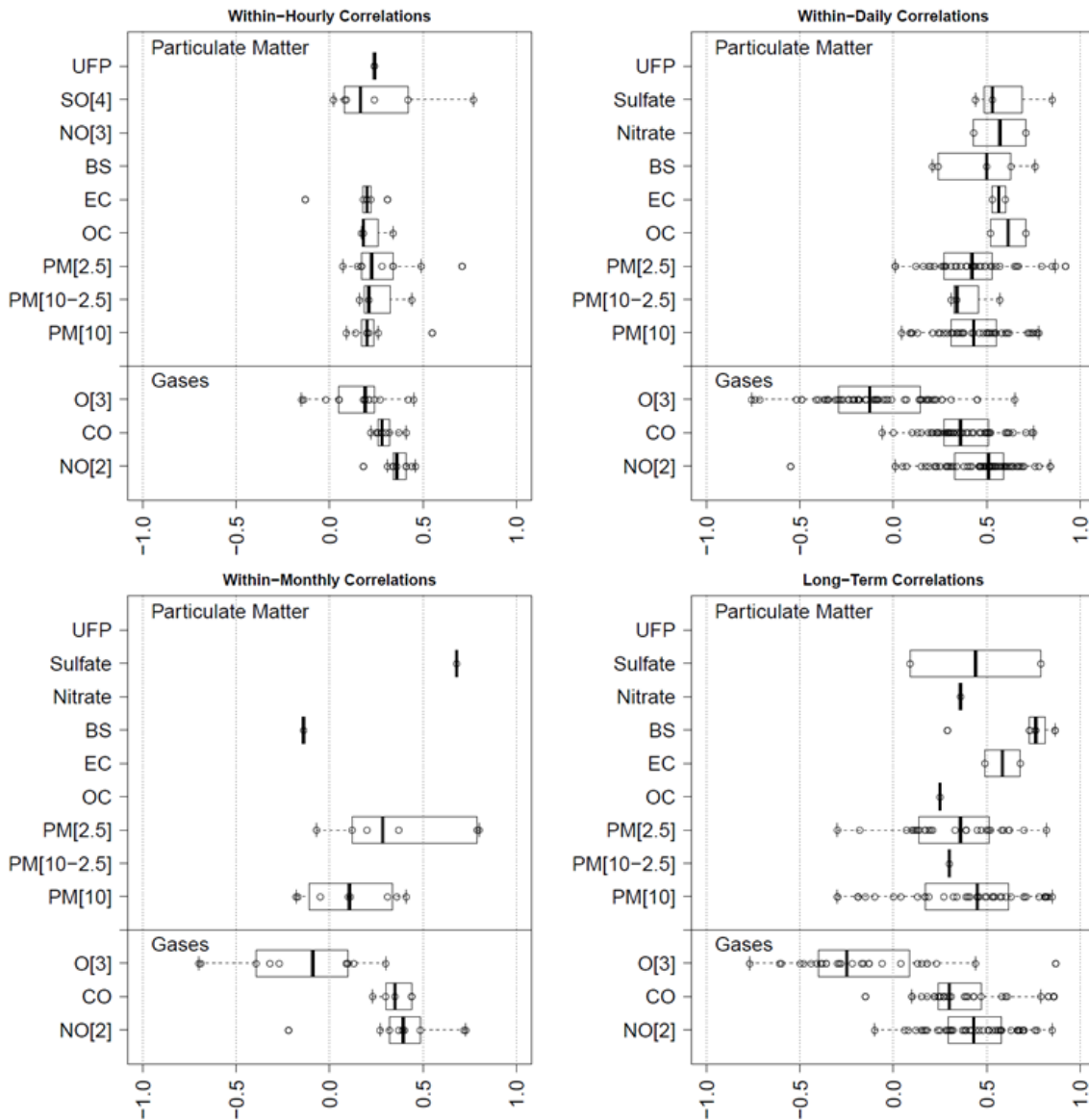
Daily and hourly ambient SO₂ concentrations generally exhibit median correlations around 0.2–0.4 with respect to other collocated NAAQS copollutants at AQS monitoring sites. However, a small subset of sites report relatively higher copollutant correlations. Tests to identify potential confounding in epidemiologic models may need to be performed if high copollutant correlations are reported in the individual studies. High copollutant correlations in the national distribution could be due either to consistently similar concentrations for both SO₂ and the copollutant or to consistent fluctuations in concentrations of both pollutants due to source behavior and meteorology.

Exposure studies have also examined correlations between ambient SO₂ concentration and ambient or personal copollutant exposure concentrations, generally reporting low or moderate correlations. For SO₂, within-hourly concentrations have median correlations around 0.2 for most PM of different cutpoints and species. For gases, median correlations of within-hourly data were lower for O₃ than for CO and NO₂, respectively, but median correlations did not surpass 0.4. Correlations were mostly positive for all but O₃, which exhibits both negative and positive correlations. See [Figure 3-8](#) and references cited therein for copollutant correlation data reported in the literature cited in ([U.S. EPA, 2017b](#)).

More data were available for within-daily correlations of SO₂ and copollutant exposure concentrations. Median correlation around 0.5 were observed for sulfate, nitrate, BS, and OC PM_{2.5} species, PM₁₀, and NO₂ for the within daily time-scale. Median correlation was around 0.3 for particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm and greater than a nominal 2.5 µm (PM_{10-2.5}), around 0.4 for CO and PM_{2.5}, and around -0.2 for O₃. Both data availability and inter-site variability were much greater for the gases, PM_{2.5}, and PM₁₀ compared with the individual PM_{2.5} species or PM_{10-2.5}. Where data were available, a large degree of scatter was evident in the data. In studies where within-daily correlations of SO₂ exposure concentrations with NO₂ and CO exposure concentrations were observed to be high, it is possible the data were collected before the rulemaking to reduce sulfur content in diesel fuel went into effect in 2006 (66 FR 5002) or when coal was in greater use in energy generation ([Section 2.2](#)). The minority of sites with stronger correlations have the potential to reflect a greater degree of confounding into the epidemiologic results if the copollutant correlations at those sites are similar to the copollutant correlations experienced at the locations of exposure. It is possible that the observed correlation at a single site may not reflect copollutant correlations at the sites of exposure, particularly in areas with a large amount of spatial heterogeneity of SO₂.

Data for correlations between ambient SO₂ concentrations and personal copollutant exposures were reported in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), and no studies have been produced to substantiate or revise the observations reported at that time.

Between-subject correlations of daily ambient SO₂ concentration with personal PM_{2.5} exposures were found to vary widely with positive and negative correlations in the [Sarnat et al. \(2005\)](#) and [Sarnat et al. \(2001\)](#) studies. In the [Sarnat et al. \(2005\)](#) study, 95–97% of the SO₂ data were below the MDL, indicating high uncertainty. This evidence suggests that correlations between personal copollutant exposures and ambient SO₂ concentration vary among individuals, and thus the potential for copollutant confounding cannot be ruled in or out.



BS = black smoke; CO = carbon monoxide; EC = elemental carbon; LUR = land use regression; NO₂ = nitrogen dioxide; NO₃ = nitrate ion; O₃ = ozone; OC = organic carbon; PM_{2.5} = in general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm, a measure of fine particles; PM₁₀ = in general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm, a measure of thoracic particles (i.e., that subset of inhalable particles thought small enough to penetrate beyond the larynx into the thoracic region of the respiratory tract); PM_{10-2.5} = in general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm and greater than 2.5 μm, a measure of thoracic coarse particulate matter or the coarse fraction of PM₁₀; SO₂ = sulfur dioxide; SO₄ = sulfate; UFP = ultrafine particulate matter.

Notes: Each data point in each boxplot represents the correlation between SO₂ and copollutants. Boxes represent the interquartile range of the data with the median line plotted, and 90th and 10th percentile of the data are plotted as the whiskers. Correlation data computed from LUR studies are not included here. Correlations shown by open black circles either come from urban-regional scale studies or do not specify the study's spatial scale. Within-monthly correlations include correlations obtained over 5 weeks or less for SO₂. Long-term correlations refers to correlations obtained over a period longer than 5 weeks.

Sources of data for this figure are listed in ([U.S. EPA, 2017b](#)).

Figure 3-8 Summary of temporal sulfur dioxide-copollutant correlation coefficients from measurements reported in the literature, sorted by temporal averaging period.

Long-Term Correlations

Long-term epidemiologic studies that have reported copollutant correlations are also displayed in [Figure 3-8](#) and references cited therein for correlations developed from data averaged over periods longer than 5 weeks. Data were limited for many of the PM_{2.5} components. For exposure concentrations of PM_{2.5}, PM₁₀, O₃, CO, and NO₂, a wide range of correlations has been estimated. Median correlation was lower for PM_{2.5} exposure concentration ($r = 0.2$) compared with that of PM₁₀ ($r = 0.4$), CO ($r = 0.3$), and NO₂ ($r = 0.3$). Median correlation was negative ($r = -0.3$) for O₃ exposure concentration. For correlations between exposure concentrations of SO₂ and PM_{2.5}, most of the data were clustered around the median, while variability in the correlations was larger for the other copollutants. As for short-term copollutant relationships, no clear conclusion can be drawn regarding the potential for confounding of long-term SO₂ epidemiologic estimates by copollutants. Wide variability in copollutant correlations was observed, with the highest correlations around 0.7–0.8 for PM_{2.5}, PM₁₀, CO, and NO₂. For the studies with high copollutant correlations, health effect estimates may be inflated by copollutant confounding ([Zeger et al., 2000](#)). Copollutant correlations are reported for tabulated epidemiologic studies throughout [Chapter 5](#) to illustrate where copollutant confounding may influence health effect estimates, and two-pollutant models designed to test for copollutant confounding are presented where available ([Section 5.2.2](#)).

3.4.3.2 Spatial Relationships among Ambient Sulfur Dioxide and Copollutants

Spatial confounding can potentially bias health effect estimates in epidemiologic studies of long-term SO₂ exposure. [Paciorek \(2010\)](#) performed simulations to test the effect of spatial confounding on health effect estimates in long-term exposure epidemiologic studies. He identified unmeasured spatial confounding as a key driver in biasing health effect estimates in a spatial regression towards the null. [Paciorek \(2010\)](#) maintained that bias in the health effect estimate is lower when variation in the exposure metric occurs at a smaller spatial scale than that of the unmeasured confounder compared with bias in the health effect estimate when the spatial scale of the exposure metric is larger than that of the unmeasured confounder.

3.4.4 Implications for Epidemiologic Studies of Different Designs

Exposure error is defined in [Section 3.2.1](#). To summarize, exposure error refers to the bias and uncertainty associated with using concentration to represent the actual exposure of an individual or population. Exposure error has two components: (1) uncertainty in the

metric used to represent exposure concentration and (2) the difference between the exposure surrogate in the epidemiologic study and the true exposure (which may not be observable) ([Zeger et al., 2000](#)). Classical error is the component of exposure measurement error derived from uncertainty in the metric being used to represent exposure concentration. Classical error is defined as error scattered around the true personal exposure and independent of the measured exposure concentration. Classical error results in bias of the epidemiologic health effect estimate that is typically towards the null (no effect of the exposure). Classical error can also cause inflation or reduction of the standard error of the health effect estimate. Berkson error is the component of exposure error related to the use of an exposure surrogate in the epidemiologic study in lieu of the true exposure. Berkson error is defined as the unobserved portion of the true exposure, and it is independent of the observed portion of the true exposure ([Goldman et al., 2011](#); [Armstrong, 2008](#); [Reeves et al., 1998](#)). Pure Berkson error generally does not bias the health effect estimate.

When investigators use statistical models to predict exposure concentrations, the exposure error is no longer purely classical or purely Berkson and may have characteristics of each error type. Measurement error for modeled exposure concentrations has been decomposed into Berkson-like and classical-like components, sharing some characteristics with pure Berkson and classical errors, respectively, but with key differences ([Szpiro et al., 2011](#)). Berkson-like errors tend to occur when the modeled exposure concentration does not capture all of the variability in the true exposure. Under ideal conditions, Berkson-like error increases the variability around the health effect estimate in a manner similar to pure Berkson error and does not induce bias, but Berkson-like error is spatially correlated and not independent of predicted exposure concentrations, so it results in underestimation of standard errors ([Goldman et al., 2011](#); [Szpiro et al., 2011](#)). [Szpiro and Paciorek \(2013\)](#) analyzed the impact of Berkson-like error under more general conditions and found that it can bias health effect estimates either toward the null or away from the null. For example, in one simulation study in which the spatial distributions of monitor and subject locations were dramatically different, the health effect estimates were biased away from the null. In another example, where spatially structured covariates were included in the health model but not in the exposure model, the health effect estimates were biased toward the null. Hence, Berkson-like error can lead to bias of the health effect estimate in either direction and should not be ignored. Classical-like errors result from uncertainty in estimating exposure model parameters. It can add variability to predicted exposure concentrations and can bias health effect estimates in a manner similar to pure classical error, but it differs from pure classical error in that the additional variability in estimated exposure concentrations is also not independent across space. Exposure error can bias epidemiologic associations between ambient pollutant concentrations and health outcomes, compared with the effect

estimate obtained using the true exposure, and it tends to widen confidence intervals around those estimates so that nominal coverage of the confidence intervals is below 95% for exposure effect estimates conditional on mismeasured covariates ([Sheppard et al., 2005](#); [Zeger et al., 2000](#)).

Exposure error can be an important contributor to uncertainty and variability in epidemiologic study results. The importance of exposure error varies with study design and is dependent on the spatial and temporal aspects of the design. Factors that could influence estimates of exposure to ambient SO₂ include time-activity patterns of the study population, topography of the natural and built environment, meteorology, instrument errors, use of ambient SO₂ concentration as a surrogate for exposure to ambient SO₂, and the presence of SO₂ in a mixture of pollutants. The following sections will consider various sources of error and how they affect the interpretation of results from epidemiologic studies of different designs.

3.4.4.1 Community Time-Series Studies

In most short-term exposure epidemiologic studies of the health effects of SO₂, the health effect endpoint is modeled as a function of ambient exposure, E_a , which is defined as the product of C_a and α , a term encompassing time-weighted averaging of exposure concentration and infiltration of SO₂ ([Section 3.2.2](#)). Community time-series epidemiologic studies capturing the exposures and health outcomes of a large cohort frequently use the ambient concentration at a fixed-site monitor ($C_{a,f}$) as a surrogate for E_a in an epidemiologic model ([Wilson et al., 2000](#)). At times, an average of fixed-site-monitored concentrations is used for the E_a surrogate. For studies involving thousands of participants, it is not feasible to measure personal exposure concentrations or time-activity patterns. Moreover, for community time-series epidemiology studies of short-term exposure, the temporal variability in ambient SO₂ concentration is of primary importance to relate to variability in the health effect estimate ([Zeger et al., 2000](#)). $C_{a,f}$ can be an acceptable surrogate if the fixed-site monitor captures the temporal variability of the actual SO₂ exposure concentration. Spatial variability in ambient SO₂ concentrations across the study area could attenuate an epidemiologic health effect estimate if the exposures are not correlated in time with $C_{a,f}$ when fixed-site monitoring is used to represent exposure in the epidemiologic model. If exposure assessment methods that more accurately capture spatial variability in the concentration distribution over a study area are employed, then the confidence intervals around the health effect estimate may decrease. $C_{a,f}$ may be an acceptable surrogate for E_a if the concentration time series at the fixed-site monitor is correlated in time with the exposures.

In a time-series study of ED visits for cardiovascular disease, [Goldman et al. \(2011\)](#) simulated the effect of classical and Berkson errors due to spatiotemporal variability among ambient air pollutant concentrations over a large urban area. For 1-h daily max SO₂, the RR per ppm was negatively biased in the case of classical error (-1.3%) and negligibly positively biased in the case of Berkson error (0.0042%). The 95% confidence interval range for RR per ppm was wider for Berkson error (0.028) compared with classical error (0.0025).

Recent studies have explored the effect of spatial exposure error on health effect estimates to test the appropriateness of using fixed-site monitoring for time-series studies. [Goldman et al. \(2010\)](#) simulated spatial exposure error based on a semivariogram function across monitor sites with and without temporal autocorrelation at 1- and 2-day lags to analyze the influence of spatiotemporal variability among ambient concentrations over a large urban area on a time-series study of ED visits for cardiovascular disease. A semivariogram estimated the change in spatial variability in exposure concentration with distance from the monitoring site. The average of the calculated random term was added to an ambient fixed-site monitoring SO₂ concentration time series (designated in this study to be the base case) to estimate SO₂ population exposure concentration subject to spatial error. For the analysis with temporal autocorrelation accounted for, RR per ppm for 1-h daily max SO₂ dropped slightly to 1.0045 (95% CI: 1.0023, 1.0065) when it was compared with the fixed-site monitor RR per ppm = 1.0139 (for all air pollutants).¹ When the model did not account for temporal autocorrelation, RR per ppm dropped very slightly to 1.0042 for 1-h daily max SO₂. The results of [Goldman et al. \(2010\)](#) suggest that spatial exposure error from the use of ambient fixed-site SO₂ concentration monitoring data results in biasing the health effect estimate towards the null, but the magnitude of the change in effect was small.

In another simulation study analyzing the influence of spatiotemporal variability among ambient concentrations over a large urban area on health effect estimates, [Goldman et al. \(2012\)](#) evaluated the effect of different types of spatial averaging on bias in the health effect risk ratio and the effect of correlation between measured and reference ambient concentrations of SO₂ and other air pollutants. Ambient concentrations were simulated at alternate monitoring locations using the geostatistical approach described above ([Goldman et al., 2010](#)) for the 20-county Atlanta metropolitan area for comparison with ambient concentration measurements obtained directly from monitors at those sites. Geostatistical-simulated ambient exposure concentrations were designated as the reference in this study, and other exposure assessment methods were assumed to have

¹ Note that 95% CIs were not reported for the fixed-site monitor RR or for the cases where temporal autocorrelation was not modeled.

some error. Five different exposure assessment approaches were tested: (1) using a single fixed-site monitor, (2) averaging the simulated exposures across all monitoring sites, (3) performing a population-weighted average across all monitoring sites, (4) performing an area-weighted average across all monitoring sites, and (5) performing population-weighted averaging of the geostatistical simulation. [Goldman et al. \(2012\)](#) observed that the exposure error was somewhat correlated with both the measured exposure concentration and the reference ambient concentrations, reflecting both Berkson and classical error components. For the fixed-site monitor, the exposure errors were somewhat inversely correlated with the exposure concentration reference value but had relatively higher positive correlation with the measured ambient concentration. For the other exposure estimation methods, the exposure errors were inversely correlated with the reference exposure concentration, while having positive but lower magnitude correlation with the measured ambient concentration. Additionally, the exposure bias, given by the ratio of the exposure error to the measured value, was much higher in magnitude at the fixed-site monitor than for the spatial averaging techniques for SO₂. Hence, compared with other exposure assessment methods, the health effect estimate would likely have greater bias towards the null with reduced precision when a fixed-site monitor is used to measure ambient SO₂ concentration as a surrogate for exposure. However, exposure error is likely to cause some bias and imprecision for other exposure surrogate methods as well.

In addition to the effect of the correlations and ratios themselves, spatial variation across urban areas also impacts time-series epidemiologic results. The [Goldman et al. \(2010\)](#) and [Goldman et al. \(2012\)](#) findings suggest more Berkson error in the spatially resolved exposure concentration metrics compared with the fixed-site monitor ambient concentration and more classical error for the fixed-site monitor ambient concentration estimate compared with the other exposure concentration measurement techniques. Hence, more bias would be expected for the health effect estimate calculated from the fixed-site monitor ambient concentration, and more variability would be expected for the health effect estimate calculated from exposure concentrations estimated by the more spatially resolved methods. Differences in the magnitude of exposure concentration estimates are not likely to cause substantial bias, but they tend more to widen confidence intervals and thus reduce the precision of the effect estimate beyond the nominal coverage of the confidence intervals that would be obtained if using the true, albeit unknown, exposure concentration ([Zeger et al., 2000](#)). The more spatially variable air pollutants studied in [Goldman et al. \(2012\)](#) also had more bias in the health effect estimates. This occurred across exposure assignment methods but was more pronounced for the fixed-site measurement ambient concentration data. Note that the [Goldman et al. \(2010\)](#), [Goldman et al. \(2011\)](#), and [Goldman et al. \(2012\)](#) studies were performed only in

Atlanta, GA. These simulation studies are informative, but similar simulation studies in additional cities would aid generalization of these study results.

[Section 3.4.2.4](#) describes the influence of high MDL on the relationship between measured ambient SO₂ concentrations and personal SO₂ exposures. When measurements are above MDL, then the amount of correlation between personal SO₂ exposure and ambient SO₂ concentrations may influence the extent of bias in a time-series study. If the reported values of personal exposure measurements are below MDL, correlation between personal SO₂ exposure measurements and ambient SO₂ concentrations will likely be low due to random noise in the signal. To the extent that true correlations are less than one, epidemiologic effect estimates based on ambient concentration will be biased toward the null, based on simulations by [Zeger et al. \(2000\)](#). Time-series epidemiologic studies employing data below MDL may demonstrate attenuated effect, but this scenario cannot be used to reject the hypothesis of a health effect.

[Section 3.4.2.4](#) also describes the influence of instrument accuracy and precision on the relationship between ambient SO₂ concentrations and personal SO₂ exposures. Exposure measurement error related to instrument precision has a smaller influence on health effect estimates in time-series studies compared with error related to spatial gradients in the ambient SO₂ concentration because instrument precision would not be expected to modify the ability of the instruments to respond to changes in ambient concentration over time. [Goldman et al. \(2010\)](#) investigated the influence of instrument error on health effect estimates in a time-series epidemiology study by studying differences in exposure concentration estimates and health effect estimates obtained using collocated monitors. Very small changes in the risk ratios were observed for 1-h daily max SO₂ ambient concentrations. For 1-h daily max SO₂ ambient concentration, the RR per ppm of SO₂ ambient concentration with simulated instrument precision error was 1.0132 compared with RR per ppm = 1.0139 for the fixed-site monitor. The amount of bias in the health effect estimate related to instrument precision was very small.

As described in [Section 3.4.1](#), nonambient sources of SO₂ are rare. Even in microenvironments where nonambient SO₂ exposure is substantial, such as in a room with a kerosene heater, such nonambient exposure concentrations are unlikely to be temporally correlated with ambient SO₂ exposure concentrations ([Wilson and Suh, 1997](#)), and therefore, would not affect epidemiologic associations between ambient SO₂ exposure concentrations and a health effect in a time-series study. [Sheppard et al. \(2005\)](#) concluded that nonambient exposure does not influence the health outcome effect estimate if ambient and nonambient exposure concentrations are independent. Personal exposure to ambient SO₂ is some fraction of the ambient concentration. Therefore, effect estimates based on personal SO₂ exposure rather than ambient SO₂ concentration will be

positively biased in proportion to the ratio of ambient SO₂ concentration to ambient SO₂ exposure concentration. Daily fluctuations in this ratio can widen the confidence intervals in the ambient SO₂ concentration effect estimate beyond the nominal coverage of the confidence intervals obtained using the true exposure. Uncorrelated nonambient exposure concentration will not bias the effect estimate but may also widen the confidence intervals ([Sheppard et al., 2005](#); [Wilson and Suh, 1997](#)).

3.4.4.2 Long-Term Cohort Studies

For long-term epidemiologic studies, the lack of personal exposure data means that investigators must rely on fixed-site ambient SO₂ concentration data or model estimates of SO₂ exposure concentration. Ambient SO₂ concentration data may be used directly, averaged across counties or other geographic areas, or used to construct geospatial or regression models to estimate exposure concentrations at unmonitored locations ([Sections 3.3.2.1 through 3.3.2.3](#)). For cohort epidemiologic studies of long-term human exposure to SO₂ where the spatial difference in the magnitude of the ambient SO₂ exposure is often of most interest and if $C_{a,f}$ is used as a surrogate for E_a , then α can represent the exposure measurement error related to uncertainties in the time-activity data and air exchange rate. Spatial variability in ambient SO₂ exposure concentrations across the study area could lead to bias in the health effect estimate if $C_{a,f}$ is not representative of E_a . This could occur, for example, if the study participants were clustered in a location where their SO₂ exposure concentration is higher or lower than the exposure concentration estimated at a model or measurement site. $C_{a,f}$ may be an acceptable surrogate for E_a if the fixed-site monitor is located close to the study participants and the ambient SO₂ source (e.g., near the plume touchdown of a power plant) and spatial variability of the ambient SO₂ concentration across the study area where the study participants are located is minimal in the vicinity of each sample group.

Model surfaces can also be constructed from physics-based models ([Sections 3.3.2.4 and 3.3.2.5](#)), but the amount of exposure error depends on model specifications. For example, [Lipfert et al. \(2009\)](#) estimated ambient SO₂ concentration based on the CMAQ model for use as an exposure surrogate. Annual average ambient SO₂ concentrations were estimated with a 36 by 36-km grid across the contiguous U.S. The modeled ambient SO₂ concentrations were used as exposure surrogates to determine their association with county-level mortality data for the Washington University-Electric Power Research Institute Veterans Cohort Mortality Study. To assign exposures at the county level, the CMAQ grid that included the largest city within each county was determined, and the associated CMAQ ambient SO₂ concentration was used as a surrogate for exposure concentration for the entire county. [Lipfert et al. \(2009\)](#) did not provide data to validate

the model results with measurements, but they found that the magnitude of the mortality risk coefficients was lower for modeling results compared with that obtained by measurements (risk coefficient for measurement: -0.047 , risk coefficient for model: -0.012), suggesting that averaging over the 36-km grid may have caused attenuation of the effect estimate.

The number of recent long-term studies of SO₂ exposure that permit evaluation of the relationship between long-term average ambient SO₂ concentrations and personal or population exposures is limited, and the value of short-term exposure concentration data for evaluating long-term exposure concentration relationships is uncertain. If the longer averaging time (annual vs. daily or hourly) smoothes out short-term fluctuations, long-term ambient SO₂ concentrations employed in long-term epidemiologic studies may be well correlated with true long-term exposures. For example, [Guay et al. \(2011\)](#) observed high correlation between single-year/single-location ambient SO₂ concentrations used for an exposure surrogate with ambient SO₂ concentrations averaged over a 22-year period when the annual SO₂ concentrations were assigned based on the study participants' census subdivision. However, lower correlation between long-term exposure and ambient SO₂ concentration could occur if important exposure determinants change over a period of several years, including activity pattern and residential air exchange rate.

Minimization of error in the exposure concentration prediction does not always minimize error in the health effect estimate. [Szpiro et al. \(2011\)](#) used simulation studies to evaluate the bias and uncertainty of the health effect estimate obtained when using correctly specified and misspecified long-term exposure concentration models. The correct exposure concentration model was designated by the authors to be an LUR with three covariates while the misspecified model included only two of these three covariates. The study authors estimated the exposure concentration model parameters using monitor data and predicted exposure concentrations at subject locations. They studied two conditions: where the variation in the third covariate was identical in the monitor and subject data versus where it was much smaller in the monitor data than in the subject data. [Szpiro et al. \(2011\)](#) showed that prediction accuracy of the exposure concentration estimate was always higher for the correctly specified model compared with the misspecified model. The health effect estimate had lower RMSE for the correct model when the third covariate had identical variability in the monitor and subject data. However, when the third covariate was much less variable in the monitor data, then the health effect estimate had lower RMSE for the misspecified model. The results of the [Szpiro et al. \(2011\)](#) simulations demonstrate one situation where use of a more accurately defined exposure concentration metric does not improve the health effect estimate.

Error correction is a relatively new approach to estimate the correct standard error and to potentially correct for bias in air pollution cohort studies. [Szpiro and Paciorek \(2013\)](#) established that two conditions must hold for the health effect estimate to be predicted correctly: (1) the exposure concentration estimates from monitors must come from the same underlying distribution as the true exposure concentrations and (2) the health effect model includes all covariates relevant to the population. [Szpiro and Paciorek \(2013\)](#) and [Bergen and Szpiro \(2015\)](#) developed methods to correct for bias from classical-like measurement error. Valid standard error estimates are obtained by means of the nonparametric bootstrap. Methods have also been proposed to correct for bias from Berkson-like error, but these require stronger conditions, including compatibility between subject and monitor locations and inclusion of spatially structured health model covariates in the exposure concentration model.

In the [Szpiro and Paciorek \(2013\)](#) study, when the assigned exposure concentration measurements were set to be uniform across space, the health effect estimate was biased away from the null with different standard error compared with the case when the exposure subjects were collocated with the study participants. When an additional spatial covariate was omitted, the health effect estimate was biased towards the null with different standard errors compared with the correctly specified model. Bias correction and bootstrap calculation of the standard errors reduced bias in the model prediction. Furthermore, bias correction with bootstrapped simulation of standard error improved the confidence interval obtained from the simulation. With no correction or bootstrapping, the standard errors and confidence interval were underestimated. Bias correction with bootstrapped simulation of standard errors produced standard errors and confidence intervals close to the true standard errors and confidence intervals. These findings imply that without bias correction, effect estimates would be biased with standard errors that underestimate the true standard error. None of the epidemiologic studies cited in [Chapter 5](#) applied bias correction or standard error bootstrapping. The findings of [Szpiro and Paciorek \(2013\)](#) suggest that bias away from the null may occur when spatial resolution of the exposure model is insufficient, while bias towards the null may occur when the model is misspecified. Without bootstrapped simulation of the standard error, it is possible that the confidence intervals around the effect estimates seem more precise than they actually are. [Spiegelman \(2013\)](#) noted that the new measurement error correction methods developed by [Szpiro and Paciorek \(2013\)](#) are a version of regression calibration. This study illustrated the influence of classical-like and Berkson-like errors on long-term exposure cohort study health effect estimates through these simulations.

Instrumentation bias could be expected to influence health effect estimates from epidemiologic studies of long-term SO₂ exposures in some situations. [Section 2.4.1](#) describes how the presence of copollutants can cause ambient SO₂ concentrations

measured using fixed-site monitors to be overestimated and how high relative humidity can cause ambient SO₂ concentration measurements to be underestimated. Relative humidity would not be expected to vary greatly within a city. However, local ambient copollutant concentrations may be spatially variable such that failure to account for differences in measurement errors could lead to some differential bias in health effect estimates across a city related to instrument error. Because climate and ambient sources are more likely to differ among cities, instrumentation error could have a larger influence on the comparison of health effect estimates among cities when fixed-site monitors are used to estimate exposure concentrations.

3.4.4.3 Panel Studies

Panel or small-scale cohort studies involving dozens of individuals (including some studies cited in [Sections 5.2.2.2](#) and [5.2.2.3](#)) may use more individualized exposure concentration measurements, including personal exposures, residential indoor or outdoor concentration measurements, or concentration data from local study-specific monitors. Modeled concentrations are typically not used as exposure surrogates in panel epidemiologic studies. A main disadvantage of the modeling approach is that the results of modeling exposure concentration must be compared to an independent set of measured exposure concentration levels ([Klepeis, 1999](#)). In addition, a modeling approach requires resource-intensive development of validated and representative model inputs, such as human activity patterns, distributions of AER, and deposition rate. Therefore, modeled exposure concentrations are used infrequently in panel epidemiologic studies.

[Section 3.4.2.4](#) describes the influence of high MDL on the relationship between measured ambient SO₂ concentrations and personal exposures for ambient SO₂. Personal exposure measurements below MDL will likely cause the correlation between personal exposure measurements and ambient SO₂ concentrations to be low due to random noise in the signal. Noise in the exposure signal would add noise to the health effect estimate in a panel epidemiologic study as well. Below MDL measurements would be unlikely to bias the effect estimate, however, because the magnitude of exposure would be low whether measured with a high-precision or low-precision device.

It is also possible that the ratio of personal SO₂ exposure to ambient SO₂ concentration in panel studies is low due to the compound's low penetration and high reactivity. This results in attenuation of the magnitude of the exposure concentration-based effect estimate relative to the ambient concentration-based effect estimate (see [Equation 3-6](#)). However, if the ratio is approximately constant over time, the strength of the statistical

association would be similar for ambient concentration- and exposure concentration-based effect estimates ([Sheppard, 2005](#); [Sheppard et al., 2005](#)).

3.5 Summary and Conclusions

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) evaluated studies of ambient SO₂ concentrations and exposures in multiple microenvironments, discussed methods for estimating personal and population exposure concentrations via monitoring and modeling, analyzed relationships between personal exposure and ambient concentrations, and discussed the implications of using ambient SO₂ concentrations as estimates of exposure concentration in epidemiologic studies. Key findings were that indoor SO₂ concentrations and personal SO₂ exposure concentrations tended to be below the detection limit of personal SO₂ samplers for averaging times of 24 hours or less, making it difficult to evaluate the relationship between ambient SO₂ concentrations and indoor or personal SO₂ exposure concentrations. However, in studies with the bulk of personal samples above the detection limit, personal measurements of SO₂ exposure were moderately correlated with ambient SO₂ concentrations. Regarding the influence of exposure concentration estimates on epidemiologic study results, high spatial variability of ambient SO₂ concentrations across an urban area results in highly variable correlations among urban SO₂ monitors. Low correlations between individual monitored ambient SO₂ concentrations and the community average ambient SO₂ concentration tend to bias effect estimates toward the null, while variations in individual personal-ambient relationships across a community will tend to widen confidence intervals around the effect estimates compared with what would be obtained if the true exposure were used in the epidemiologic model. All of these findings are supported by the recent evidence available since the previous ISA.

In the current ISA, increased focus has been placed on the use of exposure surrogates in epidemiologic studies. Multiple techniques can be used to assign SO₂ exposure concentrations for epidemiologic studies, including the use of fixed-site monitor ambient SO₂ concentrations, personal SO₂ monitors, and various types of models. Each has strengths and limitations, as summarized in [Table 3-2](#). Fixed-site monitors provide a continuous record of ambient SO₂ concentrations at their locations over many days or years, but they do not capture the relatively high spatial variability in ambient SO₂ concentration across an urban area and may not capture temporal variability of the plume when the monitor is not in the plume's path. These features tend to attenuate health effect estimates in time-series epidemiologic studies. For long-term studies, bias may occur in either direction depending on whether the monitor is over- or underestimating ambient SO₂ exposure concentration for the population of interest. In all study types, use of fixed-site monitor ambient SO₂ concentrations in lieu of the true SO₂ exposure

concentrations is expected to widen confidence intervals compared with confidence intervals that would be obtained if the true exposure were used. Personal SO₂ monitors directly measure exposure, but low ambient SO₂ concentrations often result in a substantial fraction of the samples falling below the MDL for averaging times of 24 hours or less. Personal monitors also provide a relatively limited data set, making them more suitable for panel epidemiologic studies. For microenvironmental sampling, FRM and FEMs, described in [Section 2.4](#), have also been deployed for panel studies as well.

Computational models can be used to develop exposure concentration surrogates for individuals and large populations when personal exposure measurements are unavailable. Modeling approaches may include LUR models, IDW models, dispersion models, and CTMs. Strengths and limitations of each method are discussed in [Table 3-2](#). Briefly, LUR and IDW do not take into account atmospheric chemistry and physics. IDW is a weighted average of ambient SO₂ concentrations measured at several monitors. Other spatial interpolation techniques, such as kriging, also require ambient SO₂ concentrations from several monitors and apply more complex mathematical functions to interpolate among monitors. LUR regresses measured ambient SO₂ concentrations on local variables and then uses the resulting model to predict ambient SO₂ concentrations across a study area or at the locations of specific receptors. As such, LUR enables higher spatial resolution of predicted ambient SO₂ concentrations and requires more detailed input data compared with IDW. Mechanistic models, such as dispersion models and CTMs, simulate the transport and dispersion of ambient SO₂, and in the case of CTMs, the atmospheric chemistry. The strength of mechanistic models is increased accuracy of the ambient SO₂ concentration field over time and space. However, they are much more computationally intensive. Microenvironmental models require personal sensor data for input and are resource intensive. The strength of these models is that they account for time the exposed population spend in different microenvironments. Microenvironmental models ([Section 3.3.2.6](#)) incorporate time-activity data to overcome some limitations of spatial smoothing in grid-based models, but they are rarely used in epidemiologic models. With the exception of microenvironmental models, these methods tend to be used in epidemiologic studies of long-term ambient SO₂ exposure. Depending on the modeling approach, there is the potential for bias and reduced precision due to model misspecification, missing sources, smoothing of concentration gradients, and complex topography. Evaluation of model results helps demonstrate the suitability of that approach for particular applications.

The current ISA also reviews the newly available literature regarding indoor and personal exposures to SO₂. New studies of the relationship between indoor and outdoor SO₂ concentrations have focused on public buildings and are consistent with previous studies showing that indoor:outdoor ratios and slopes cover an extremely wide range, from near

zero to near one. Differences in results among studies are due to the lack of indoor sources of SO₂, indoor deposition of ambient SO₂, building characteristics (e.g., forced ventilation, building age, and building type such as residences or public buildings), personal activities, and analytical approaches. When reported, correlations between indoor and outdoor SO₂ concentrations were relatively high (>0.75), suggesting that variations in outdoor SO₂ concentrations are driving indoor SO₂ concentrations. Several studies of personal-ambient SO₂ relationships available at the time of the previous ISA showed a large fraction of samples below the MDL, making them unsuitable for determining personal-ambient correlations. In a study with all personal samples above the MDL, personal exposure was moderately correlated with ambient concentration.

Additional factors that could contribute to error in estimating exposure to ambient SO₂ include time-location-activity patterns, spatial and temporal variability in SO₂ concentrations, and proximity of populations to sources and monitor or model receptor sites. Activity patterns vary both among and within individuals, resulting in corresponding variations in exposure across a population and over time. Ambient SO₂ concentrations among different microenvironments and the amount of time spent in each location will jointly influence an individual's exposure to ambient SO₂ (see [Equation 3-3](#)). Time spent in different locations has also been found to vary by age, with younger and older age groups spending a greater percentage of time outdoors than adults of typical working age (18–64 years). These variations in activity pattern contribute to differences in exposure and introduce error into population-averaged SO₂ exposure estimates.

Failure to account for spatial and temporal variability in ambient SO₂ concentrations can contribute to exposure error in epidemiologic studies, whether the study relies on fixed-site monitor data or concentration modeling for exposure assessment. Ambient SO₂ concentrations have low to moderate spatial correlations between ambient monitors across urban geographic scales. This implies that a finer geographic scale is needed to measure exposure concentration. Thus, using ambient SO₂ concentration data measured at fixed-site monitors as exposure surrogates in epidemiologic studies may introduce exposure error into the resulting health effect estimate. Spatial variability in the magnitude of ambient SO₂ concentrations can affect cross-sectional and large-scale cohort studies by undermining the assumption that intra-urban ambient SO₂ exposure differences across space are less important than inter-urban differences. This issue may be less important for time-series studies, which rely on day-to-day temporal variability in ambient SO₂ exposure concentrations to evaluate health effects. Modeling techniques to capture spatial variability can reduce exposure error in long-term average epidemiologic models.

Proximity of populations to ambient SO₂ monitors may influence how well human exposure to ambient SO₂ is represented by measurements at the monitors, although factors other than distance also play an important role. Many ambient SO₂ monitoring sites are located near dense population centers, but other near-source sites may not be near population centers. Use of monitoring sites in epidemiologic studies introduces exposure error into health effect estimates. The literature has shown that exposure error and related bias in the health effect estimate is reduced by using averaging schemes in lieu of a single fixed-site monitor ([Section 3.4.2.2](#)).

Exposure to copollutants may result in confounding of health effect estimates. For ambient SO₂, daily concentrations generally exhibit low to moderate correlations with daily NAAQS copollutant concentrations at collocated monitors ([Figure 3-4](#)). However, a wide range of copollutant correlations is observed at different monitoring sites, from moderately negative to moderately positive. In studies where daily correlations of ambient SO₂ concentrations with ambient NO₂ and CO concentrations were observed to be high, it is possible the data were collected before rulemaking to reduce sulfur content in diesel fuel went into effect in 2006 (66 FR 5002). The minority of sites with stronger correlations have the potential to reflect a greater degree of confounding into the epidemiologic results if the copollutant correlations at those sites are similar to the copollutant correlations experienced at the locations of exposure. It is possible that the observed correlation at a single site may not reflect copollutant correlations at the sites of exposure, particularly in areas with a large amount of spatial heterogeneity of SO₂. A similar impact is expected for epidemiologic studies of long-term ambient SO₂ exposure, because a wide range of copollutant correlations have also been reported over time periods of months to years.

Exposure error can contribute to variability in epidemiologic study results by biasing effect estimates toward or away from the null and changing the size of the confidence intervals compared with the confidence intervals around the effect estimate if the true ambient SO₂ exposure could have been used. The importance of exposure error varies according to the study design, especially regarding the study's spatial and temporal aspects. For example, in time-series and panel studies, low personal-ambient correlations tend to bias the effect estimate toward the null, while spatial variation in personal-ambient correlations across an urban area contributes to widening of the confidence interval around the effect estimate compared with the confidence interval produced using the true ambient SO₂ exposure. For long-term studies, bias of the health effect estimate may occur in either direction depending on whether the monitor is over- or underestimating true ambient SO₂ exposure concentration for the population of interest. In all study types, use of fixed-site monitors in lieu of the true ambient SO₂ exposure concentration is expected to decrease precision of the health effect estimate

because spatial variation in personal-ambient correlations across an urban area can alter the confidence interval around the effect estimate compared with the confidence interval that would be obtained if the true ambient SO₂ exposure concentration were used. Choice of exposure concentration estimation method also influences the impact of exposure error on epidemiologic study results. Fixed-site monitors offer a convenient source of time-series data. However, because they are in a fixed location, ambient SO₂ concentration measurements obtained from a fixed-site monitor do not account for the effects of spatial variation in ambient SO₂ concentration, ambient and nonambient concentration differences, and varying activity patterns on personal exposure to ambient SO₂. Personal exposure measurements, such as those made in panel epidemiologic studies, provide specific exposure estimates that may more accurately reflect spatial and temporal variability, but sample size is often small and only a limited set of health outcomes can be studied. Modeled ambient SO₂ concentration or exposure concentration estimates offer alternatives or supplementation to measurements, with the advantage of estimating ambient SO₂ exposure concentrations over a wide range of scales, populations, and scenarios, particularly for locations lacking monitoring data. Model estimates are most useful when compared to an independent set of measured ambient SO₂ concentrations or exposure concentrations. The various sources of exposure error and their potential impact on epidemiologic study results are evaluated in this ISA.

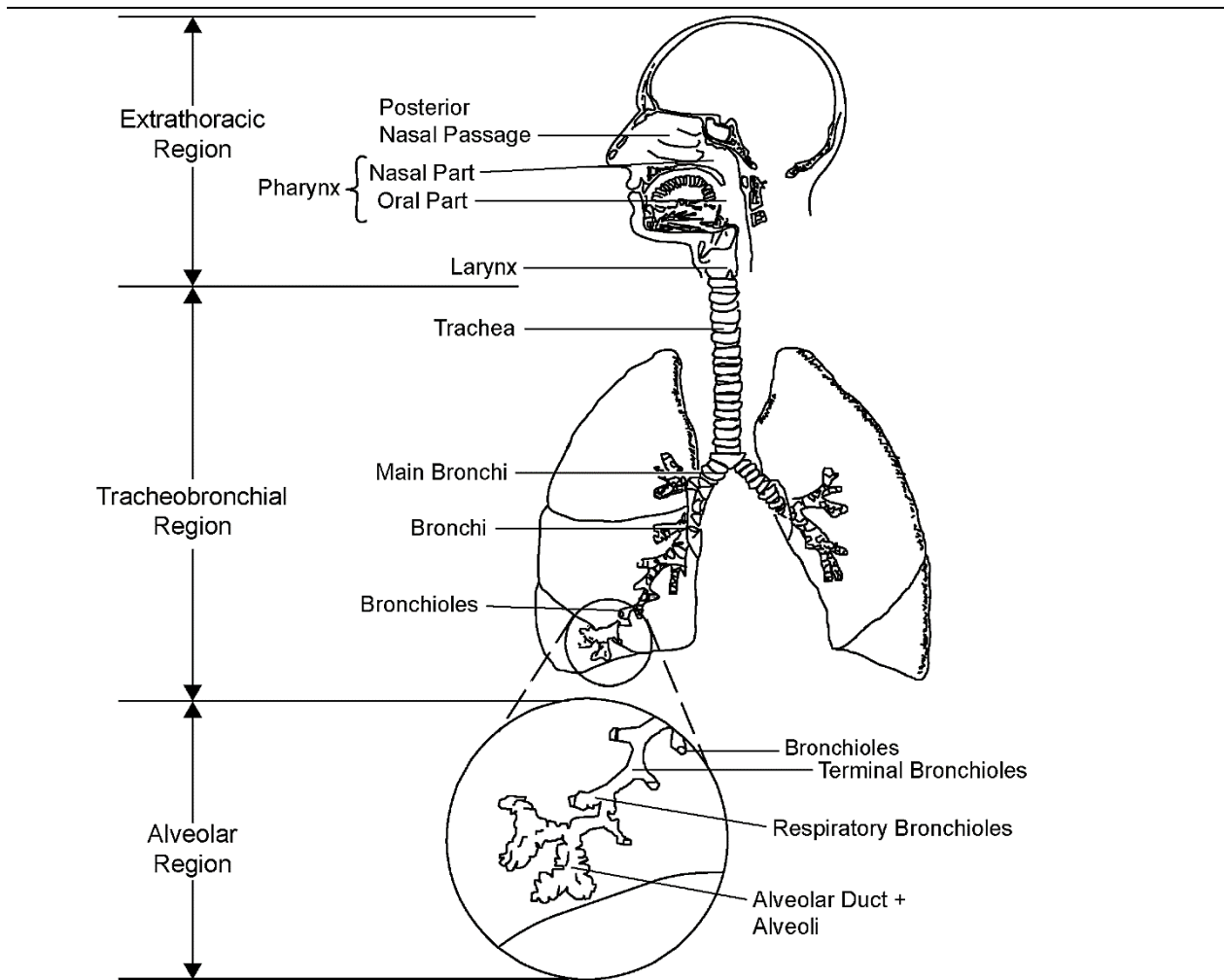
CHAPTER 4 DOSIMETRY AND MODE OF ACTION

4.1 Introduction

[Chapter 4](#) begins by providing background information on the structure and function of the respiratory tract ([Section 4.1.1](#)) and breathing rates and habits ([Section 4.1.2](#)). The subsequent discussion of dosimetry of inhaled sulfur dioxide (SO₂) ([Section 4.2](#)) considers the chemical properties of SO₂ and the processes of absorption, distribution, metabolism, and elimination, as well as sources and levels of exogenous and endogenous sulfite. The biological pathways that potentially underlie health effects are described in “Modes of Action of Inhaled Sulfur Dioxide” ([Section 4.3](#)). This section includes a description of processes by which inhaled SO₂ initiates a cascade of molecular and cellular responses and the organ-level responses that follow. Together, these sections provide the foundation for understanding how exposure to inhaled SO₂ may lead to health effects. This understanding may provide biological plausibility for effects observed in the epidemiologic studies.

4.1.1 Structure and Function of the Respiratory Tract

The basic structure of the human respiratory tract is illustrated in [Figure 4-1](#). In the literature, the terms extrathoracic (ET) region and upper airways or upper respiratory tract are used synonymously. The terms lower airways and lower respiratory tract refer to the intra-thoracic airways [i.e., the tracheobronchial (TB) and alveolar regions of the lung].



Source: Based on [ICRP \(1994\)](#).

Figure 4-1 Diagrammatic representation of respiratory tract regions in humans.

4.1.2 Breathing Rates and Breathing Habit

4.1.2.1 Breathing Rates

Breathing rates vary across the day and are generally a function of an individual's age, sex, and activity level. [Table 4-1](#) provides median ventilation rates extracted from Tables 6–17 (males) and 6–19 (females) of the *Exposure Factors Handbook* ([U.S. EPA, 2011](#)). The original source of these ventilation rates is Table C4 of [U.S. EPA \(2009b\)](#) which describes their derivation. The median ventilation rates in [Table 4-1](#) represent central tendency estimates across a distribution of body weights for each age group and a distribution of metabolic equivalents of work for each age group. Additional information for other ages and percentiles of the ventilation rate distribution are available from those tables. Except for the oldest age range, ventilation rates (volume/time) increase with activity level and age and are greater in men than women.

Table 4-1 Ventilation rates in humans as a function of activity.

Median Ventilation Rate (L/min)					
Sex	Age (Yr)	Sleep	Light Activity	Moderate Activity	Strenuous Activity
Male	3 to <6	4.29	11.1	20.6	37.8
	6 to <11	4.46	11.3	21.6	41.9
	21 to <61	5.17	13.6	29.7	52.9
	≥81	5.90	13.8	28.2	50.9
Female	3 to <6	4.1	10.7	19.8	33.3
	6 to <11	4.24	10.8	20.4	38.0
	21 to <61	4.06	11.1	23.0	44.2
	≥81	4.39	10.3	20.6	41.4

Source: [U.S. EPA \(2011\)](#) and [U.S. EPA \(2009b\)](#).

Ventilation rates are also higher in overweight individuals compared with those of normal weight ([Brochu et al., 2014](#)). For example, median daily ventilation rates (m³/day) are about 1.2 times greater in overweight [>85th percentile body mass index (BMI)] than normal-weight children (5–10 years of age). In 35–45-year-old adult males and females,

ventilation rates are 1.4 times greater in overweight ($\text{BMI} \geq 25 \text{ kg/m}^2$) than normal-weight ($18.5 < \text{BMI} < 25 \text{ kg/m}^2$) individuals. Across all ages, overweight/obese individuals respire greater amounts of air and associated pollutants than age-matched normal-weight individuals.

One way to consider differences in ventilation rates between adults and children is to normalize to body weight. This metric is relevant especially for SO_2 absorbed in the nasal airways and the fraction of absorbed SO_2 that distributes systemically ([Section 4.2.3](#)). Normalized to body mass, median daily ventilation rates ($\text{m}^3/\text{kg}\text{-day}$) decrease over the course of life ([Brochu et al., 2011](#)). This decrease in ventilation relative to body mass is rapid and nearly linear from infancy through early adulthood. Relative to normal-weight male and female adults (25–45 years of age; $0.271 \text{ m}^3/\text{kg}\text{-day}$), ventilation rates normalized to body mass are increased 1.5 times in normal-weight children (7–10 years of age; $0.402 \text{ m}^3/\text{kg}\text{-day}$) and doubled in normal-weight infants (0.22–0.5 years of age; $0.538 \text{ m}^3/\text{kg}\text{-day}$). Although adults have greater absolute ventilation rates than children in terms of inhaled volume per unit time, normalized to body size, children intake greater volumes of air and associated pollutants than adults.

A potentially useful metric for assessing SO_2 dose to the bronchi and differences between children and adults in bronchial effects of SO_2 is SO_2 absorbed dose per bronchial surface area ([Section 4.2.2](#)). Ventilation per tracheobronchial surface area is also used to approximate absorbed dose per bronchial surface area for inter-species extrapolation [[Appendix A of U.S. EPA \(2009d\)](#)].

4.1.2.2 Breathing Habit

As humans, we breathe oronasally (i.e., through both our nose and mouth). In general, we breathe through our nose when at rest and increasingly through the mouth with increasing activity level. Few people breathe solely through their mouth. In contrast to the oronasal breathing of humans, rodents are obligate nasal breathers. Described in [Section 4.2.2](#), the nasal passages more efficiently remove SO_2 from inhaled air than the oral passage. As the fraction of inhaled air passing through the mouth increases so too does the amount of inhaled SO_2 reaching the tracheobronchial airways where SO_2 may cause bronchoconstriction. Thus, route of breathing (namely, the fraction of inhaled air passing through the mouth) is a critical determinate of dose to the lower airways and the potential respiratory effects of SO_2 . This section describes how route of breathing, also referred to as “respiratory mode” or “breathing habit” in the literature, is affected by age, sex, obesity, activity level, and upper respiratory tract anomalies.

One of the more commonly referenced studies in dosimetric papers is [Niinimaa et al. \(1981\)](#). These investigators found that most people, 87% (26 of 30) in the study, breathed through their nose until an activity level was reached where they switched to oronasal breathing. Thirteen percent (4 of 30) of the subjects, however, were oronasal breathers even at rest. These two subject groups are commonly referred to in the literature (e.g., [ICRP, 1994](#)) as “normal augmenters” and “mouth breathers,” respectively. [Bennett et al. \(2003\)](#) reported a more gradual increase in oronasal breathing, with males (n = 11; 22 ± 4 years) tending to have a greater oral contribution than females (n = 11; 22 ± 2 years) at rest (87 vs. 100% nasal, respectively) and during exercise (45 vs. 63% nasal at 60% max workload, respectively).

Consistent with this trend for women to have a greater nasal contribution ([Bennett et al., 2003](#)); in a large study of children (63 M, 57 F; 4–19 years), [Leiberman et al. \(1990\)](#) reported a statistically greater nasal fraction during inspiration in girls relative to boys (77 and 62%, respectively; $p = 0.03$) and a marginally significant difference during expiration (78 and 66%, respectively; $p = 0.052$). Another large study (88 M, 109 F; 5–73 years) also reported females as having a significantly greater fraction of nasal breathing than males ([Vig and Zajac, 1993](#)). This effect was largest in children (5–12 years) with an inspiratory nasal fraction, under resting conditions, of 66% in males and 86% in females. This study also reported that the partitioning between the nose and mouth was almost identical between inspiration and expiration. In children and adults, sex explains some inter-individual variability in route of breathing, with females breathing more through the nose than males.

A few studies have attempted to measure oronasal breathing in children as compared to adults ([Bennett et al., 2008](#); [Becquemin et al., 1999](#); [James et al., 1997](#); [Vig and Zajac, 1993](#)). [James et al. \(1997\)](#) found that children (n = 10; 7–16 years) displayed more variability than older age groups (n = 27; 17–72 years) with respect to their oronasal pattern of breathing with exercise. [Becquemin et al. \(1999\)](#) found that children (n = 10; 8–16 years) tended to display more oral breathing both at rest and during exercise than adults (n = 10; 27–56 years). The highest oral fractions were also found in the youngest children. Similarly, [Bennett et al. \(2008\)](#) reported children (n = 12; 6–10 years) tended to have a greater oral contribution than adults (n = 11; 18–27 years) at rest (68 vs. 88% nasal, respectively) and during exercise (47 vs. 59% nasal at 40% max workload, respectively). [Vig and Zajac \(1993\)](#) reported a statistically significant effect of age on route of breathing which was most apparent in males with the fraction of nasal breathing increasing from 67% in children (5–12 year olds) to 82% in teens (13–19 year olds), and 86% in adults (20–73 years). Females had a nasal fraction of 86% in children and teens and 93% in adults. Based on these studies, the nasal fraction increases with age until adulthood.

Several large studies have reported an inverse correlation ($r = -0.3$ to -0.6) between nasal resistance and nasal breathing fraction ([Vig and Zajac, 1993](#); [Leiberman et al., 1990](#); [Leiter and Baker, 1989](#)). However, neither pharmaceutical constriction nor dilation of the nasal passages affected the nasal fraction ([Leiberman et al., 1990](#); [Leiter and Baker, 1989](#)). Nasal resistance decreases with age and is lower in females and may account for larger nasal fractions in adults and females ([Vig and Zajac, 1993](#)). Smaller studies ($n = 37$) have not found a significant correlation between nasal resistance and nasal fraction, but have noted that those having high resistance breathe less through the nose ([James et al., 1997](#)). [Bennett et al. \(2003\)](#) reported a tendency of lower nasal resistance in African-American blacks (5 M, 6 F; 22 ± 4 years) relative to Caucasians (6 M, 5 F; 22 ± 3 years). The nasal fraction in blacks tended to be greater at rest and 40% max workload and achieved statistical significance relative to Caucasians at 20 and 60% max workload. [Leiter and Baker \(1989\)](#) reported that of the 15 mouth-breathing children as identified by a dentist, pediatrician, or otolaryngologist in their study, the 3 having greatest nasal resistance breathed 100% through the mouth. These investigators also reported that the nasal fraction was negatively correlated ($p \leq 0.004$) with nasal resistance during both inspiration and expiration; however, the correlation appears driven by the three individuals with 100% mouth breathing. Overall, breathing habit is related to nasal resistance, which may explain some of the effects of age and sex on breathing habit.

Diseases affecting nasal resistance may also affect breathing route. [Chadha et al. \(1987\)](#) found that the majority (11 of 12) of patients with asthma or allergic rhinitis breathe oronasally (i.e., they breathe partially through the mouth) even at rest. [James et al. \(1997\)](#) also reported the subjects ($n = 37$; 7–72 years) having hay fever, sinus disease, or recent upper respiratory tract symptoms tended to have a greater oral contribution relative to those absent upper respiratory tract symptoms. [James et al. \(1997\)](#) additionally observed that two subjects (5.4%) breathed purely through the mouth, but provided no other characteristics of these individuals. Greater oral breathing may occur due to upper respiratory tract infection and inflammation.

Some studies of children suggest obesity also affects breathing habit. Using MRI, [Schwab et al. \(2015\)](#) examined anatomic risk factors of obstructive sleep apnea in children ($n = 49$ obese with sleep apnea, 38 obese control, 50 lean controls; 11–16 years of age). In obese children with sleep apnea, adenoid size was increased relative to both obese and lean controls not having sleep apnea. The size of the adenoid was also increased in male obese controls ($n = 24$) relative to male lean controls ($n = 35$), whereas adenoid size was similar between female obese controls ($n = 14$) and female lean controls ($n = 15$). Both nasopharyngeal cross-sectional area and minimum area were similar between lean and obese controls, but decreased in obese children with obstructive sleep apnea. In a longitudinal study of children ($n = 47$ F, 35 M) assessed annually from 9 to

13 years of age, [Crouse et al. \(1999\)](#) found nasal cross section was minimal at 10 years of age. The authors speculated this may be due to prepubertal enlargement of the adenoids. In a 5-year longitudinal study of children (n = 17 M, 9 F) following adenoidectomy, [Kerr et al. \(1989\)](#) reported a change in mode of breathing from oral to nasal. These studies suggest that obese children, especially boys, may have increased oral breathing relative to normal-weight children.

In summary, breathing habit is affected by age, sex, nasal resistance, and perhaps by obesity. Numerous studies show children to inhale a larger fraction of air through their mouth than adults. Across all ages, males also inhale a larger fraction of air through their mouth than females. Other factors that increase nasal resistance such as allergies or acute upper respiratory infections can also increase the fraction of oral breathing. Obesity, especially in boys, may also contribute to increased nasal resistance and an increased oral fraction of breathing relative to normal-weight children.

4.2 Dosimetry of Inhaled Sulfur Dioxide

This section provides a brief overview of SO₂ dosimetry and updates information provided in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). Dosimetry of SO₂ refers to the measurement or estimation of the amount of SO₂ and its reaction products reaching and/or persisting at specific sites within the respiratory tract or systemically after exposure. One principal effect of inhaled SO₂ is to stimulate bronchial epithelial irritant receptors and initiate a reflexive contraction of smooth muscles in the bronchial airways. Health effects may be due to the inhaled SO₂ or its chemical reaction products. Complete identification of the causative agents and their integration into SO₂ dosimetry is a complex issue that has not been thoroughly evaluated. The major factors affecting the transport and fate of gases and aerosols in the respiratory tract are the morphology of the respiratory tract; the physicochemical properties of the epithelial lining fluid (ELF); respiratory functional parameters, such as tidal volume, flow rate, and route of breathing; the physicochemical properties of the gas; and the physical processes that govern gas transport. Few studies have investigated SO₂ dosimetry since the 1982 AQCD for Particulate Matter and Sulfur Oxides ([U.S. EPA, 1982a](#)) and the 1986 Second Addendum ([U.S. EPA, 1986](#)).

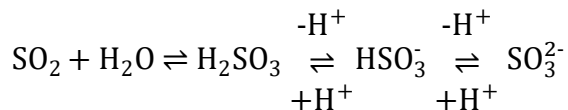
The following sections address SO₂ chemistry and the processes of absorption, distribution, metabolism, and elimination that pertain to the dosimetry of inhaled SO₂. Studies investigating the dosimetry of SO₂ generally are for concentrations of SO₂ that are higher than those present in ambient air. However, these studies are included here because they provide the foundation for understanding SO₂ toxicokinetics and

toxicodynamics. The discussion of dosimetry concludes with a consideration of other sources of SO₂-derived products in the body.

4.2.1 Chemistry

Physicochemical properties of SO₂ most relevant to respiratory tract uptake include its solubility in the ELF and its chemical transformations and reactions that occur there. Henry's law relates the gas-phase and liquid-phase interfacial concentrations at equilibrium and is a function of temperature and pressure. The Henry's law constant, defined as the ratio of partial pressure or concentration of SO₂ in the gas phase to SO₂ dissolved in the liquid phase, is an inverse measure of solubility. Although the solubility of SO₂ in the ELF is not known, the effective Henry's law constant or solubility factor is known for SO₂ in water and is 0.047 (mol/L)_{air} per (mol/L)_{water} at 37°C and 1 atmosphere ([Hales and Sutter, 1973](#)). For comparison, Henry's law constant for ozone (O₃) is 6.4 (mol/L)_{air} per (mol/L)_{water} under the same conditions ([Kimbell and Miller, 1999](#)). Thus, SO₂ is nearly 140 times more soluble than O₃ in water. In general, the more soluble a gas is in biological fluids, the more rapid and proximal its absorption will be in the respiratory tract. In addition to the Henry's law constant, it is also necessary to consider the transport of SO₂ from the lumen to the ELF of the tracheobronchial airways ([Section 4.2.2](#)). When the partial pressure of SO₂ on mucosal surfaces exceeds that of the gas phase, such as during expiration, some desorption of SO₂ from the ELF may be expected ([Section 4.2.5](#)).

Once SO₂ contacts the fluids lining the airways, it dissolves into the aqueous compartment and rapidly hydrates to form sulfurous acid (H₂SO₃), which forms hydrogen (H⁺) ions, bisulfite (HSO₃⁻) anions, and sulfite (SO₃²⁻) anions ([Gunnison et al., 1987a](#); [Gunnison, 1981](#)).



Equation 4-1

The prevalence of these sulfur species in solution is determined primarily by pH, and to a lesser extent, by temperature and ionic strength. In the human respiratory tract (pH of 7.4 and 37°C), dissolved SO₂ exists exclusively as a mixture of bisulfite and sulfite, with the latter predominating ([Gunnison, 1981](#)). Subsequent reactions of bisulfite and sulfite such as sulfitolysis, enzymatic detoxification, and auto-oxidation are described below. Hydrogen ions may impact airway physiology via acidification reactions.

4.2.2 Absorption

Because SO₂ is highly soluble in water, it is expected to be almost completely absorbed in the nasal passages of both humans and laboratory animals under resting conditions. The dosimetry of SO₂ can be contrasted with the lower solubility gas, O₃, for which the predicted tissue doses (O₃ flux to the liquid-tissue interface) are very low in the trachea and increase to a maximum in the terminal bronchioles or first airway generation in the pulmonary region [Chapter 5 of [U.S. EPA \(2013b\)](#)].

[Melville \(1970\)](#) measured the absorption of SO₂ (1.5 to 3.4 ppm) during nasal and oral breathing in 12 healthy volunteers. Total respiratory tract absorption of SO₂ (expressed as a percentage of the amount inhaled) was significantly greater ($p < 0.01$) during nasal than oral breathing (85 vs. 70%, respectively) and was independent of the inspired concentration. Respired flows were not reported. [Andersen et al. \(1974\)](#) measured the nasal absorption of SO₂ (25.5 ppm) in seven volunteers at an average inspired flow of 23 L/minute [i.e., eucapnic hyperpnea (presumably to simulate light exertion)]. These investigators reported that the oropharyngeal SO₂ concentration was below their limit of detection (0.25 ppm), implying that at least 99% of SO₂ had been absorbed in the nose of subjects during inspiration. [Speizer and Frank \(1966\)](#) also measured the absorption of SO₂ (16.1 ppm) in seven healthy subjects at an average ventilation of 8.5 L/minute (i.e., at rest). They reported that 14% of the inhaled SO₂ was absorbed within the first 2 cm into the nose. The concentration of SO₂ reaching the pharynx was below the limit of detection, suggesting that at least 99% was absorbed in the nose during inspiration.

[Frank et al. \(1969\)](#) and [Brain \(1970\)](#) investigated the oral and nasal absorption of SO₂ in the surgically isolated upper respiratory tract of anesthetized dogs. Radiolabeled SO₂ (³⁵SO₂) at concentrations of 1, 10, 25, or 50 ppm was passed separately through the nose and mouth at steady flows of 3.5 and 35 L/minute for 5 minutes by [Brain \(1970\)](#). The nasal absorption of SO₂ (1 ppm) was effectively 100% at 3.5 L/minute and 96.8% at 35 L/minute. The effect of SO₂ concentration on nasal absorption was negligible, with nasal absorption increasing from 99.9% at 1 ppm to 99.99% at 10 ppm and 99.999% at 50 ppm at 3.5 L/minute. The oral absorption of SO₂ (1 ppm) was 99.56% at 3.5 L/minute, but only 34% at 35 L/minute. There was a slight decrease in oral SO₂ absorption from 99.56 to 96.3% when the concentration was increased from 1 to 10 ppm at 3.5 L/minute, whereas nasal absorption was unaffected by changes in concentration (1–50 ppm). In an earlier experiment, [Frank et al. \(1967\)](#) showed that nasal absorption of 2.2 ppm ³⁵SO₂ at 3.5 L/minute was 100% throughout the first 20 minutes of exposure. On average, there was a small reduction in ³⁵SO₂ absorption to 94% approaching 30 minutes of exposure. [Frank et al. \(1969\)](#) noted that the aperture of the mouth may vary considerably, and that this variation may affect SO₂ uptake in the mouth. Although there was a minor effect of

inhaled concentration on SO₂ absorption, the route of breathing and rate of flow were the main factors affecting the magnitude of SO₂ absorption in the upper airways of dogs.

Modeling shows that virtually all SO₂ reaching the lower airways in young adults, as well as in dogs and rats, is absorbed in the bronchi and does not penetrate into the bronchioles or alveolar region ([Tsuji no et al., 2005](#)). Thus, ventilation per bronchial surface area can serve as a surrogate for inhaled SO₂ dose per unit bronchial surface area. Considering the effect of age on SO₂ dose to the airways of humans, dose as ventilation per bronchial surface area can be estimated using bronchial morphology from [Phalen et al. \(1985\)](#) and 50th-percentile ventilation rates from [Brochu et al. \(2011\)](#). This approximation shows a gradual reduction in bronchial surface dose with decreasing age from young adults to infants. Using this approximation, an infant (4-months old) would have approximately 80% of the bronchial surface dose of a young adult (18-years old). However, as described in [Section 4.1.2.2](#), children breathe more through the mouth than adults, which is associated with greater SO₂ penetration to the lower respiratory tract. SO₂ penetrating through the upper airways is rapidly removed in the trachea and the first several generations of bronchi, possibly resulting in somewhat greater airway surface doses of SO₂ of children than adults in proximal bronchi due to the greater oral breathing contribution of children than adults.

In summary, inhaled SO₂ is readily absorbed in the upper airways of both humans and laboratory animals. During nasal breathing, the majority of available data suggests 95% or greater SO₂ absorption occurs in the nasal passages, even under ventilation levels comparable to those during exercise. Somewhat less SO₂ is absorbed in the oral passage than in the nasal passages. The difference in SO₂ absorption between the mouth and the nose is highly dependent on respired flow rates. With an increase in flow from 3.5 to 35 L/minute, nasal absorption is relatively unaffected, whereas oral absorption is reduced from 100 to 34%. Inhaled SO₂ concentration has a negligible effect on nasal absorption, whereas oral absorption may decrease slightly with increasing concentration from 1 ppm to 10 ppm SO₂. Thus, the rate of breathing (namely, for oral breathing) and the route of breathing (i.e., the contribution through the nose vs. mouth) have a great effect on the magnitude of SO₂ absorption in the upper airways and on the penetration of SO₂ to the lower airways. Overall, the available data clearly show a pattern of SO₂ absorption that shifts from the upper airways to the tracheobronchial airways in conjunction with a shift from nasal to oronasal breathing and associated increased ventilatory rates in exercising humans. Due to their increased amount of oral breathing, children (particularly boys and perhaps the obese) and individuals with allergies or upper airway infections may be expected to have greater SO₂ penetration into the lower respiratory tract than healthy adults ([Section 4.1.2](#)).

4.2.3 Distribution

Once inhaled, SO₂ is absorbed in the respiratory tract, and SO₂-derived products are widely distributed throughout the body, as demonstrated in studies using radiolabeled ³⁵SO₂. Although rapid extra-pulmonary distribution of SO₂-derived products occurs, the highest tissue concentrations of the ³⁵S retained in the body at any given time are found primarily in the respiratory tract (upper and lower) and may be detected there for up to a week following inhalation ([Balchum et al., 1960, 1959](#)). [Frank et al. \(1967\)](#) observed ³⁵S in the blood and urine of dogs within 5 minutes, the first time point, after starting 22 ppm ³⁵SO₂ exposures of the surgically isolated nasal airways. At the end of 30–60-minute exposures, the authors estimated that 5–18% of the administered ³⁵S was in the blood. [Balchum et al. \(1959\)](#) investigated the tissue distribution of ³⁵S in dogs exposed for 20–40 minutes to ³⁵SO₂ ranging in concentration from 1.1 to 141 ppm via tracheostomy or by nose/mouth breathing. At approximately 1-hour post-exposure, regardless of the exposure route or the ³⁵SO₂ exposure concentration, about 6% of the retained ³⁵S was found in the liver, with lesser amounts found in the heart, spleen, kidney, brain, and other tissues. However, the percent of retained ³⁵S was, on average, 13 times greater in the trachea and lungs of the tracheostomized group than in the nose/mouth breathing group, demonstrating the protection of the lower respiratory tract provided by SO₂ removal in the upper airways. Comparison of dogs retaining similar total amounts of ³⁵S (i.e., controlling for retained dose), showed that the blood concentrations of ³⁵S were higher in the tracheostomized dogs than in the nose/mouth breathing dogs. Given very high ³⁵S concentrations in the tongues of the nose/mouth breathing dogs and that blood concentrations had not decreased in two-thirds of these dogs by 1-hour post-exposure, the authors postulated that a substantial portion of the ³⁵SO₂ products may have been retained within the upper airways with only slow absorption into the blood. Studies in rabbits and rats also show that there can be an accumulation and retention of SO₂-derived products within proximal regions of the respiratory tract (discussed below).

The distribution and clearance of inhaled SO₂ from the respiratory tract may involve several intermediate chemical reactions and transformations. In particular, hydrated SO₂ transforms to sulfite/bisulfite at physiologic pH. Sulfite can diffuse across cell membranes, and bisulfite can react with disulfide bonds (R₁-S-S-R₂) to form thiols (R₁-SH) and S-sulfonates (R₂-S-SO₃⁻) by a process termed sulfitolysis ([Gunnison and Benton, 1971](#)). Disulfide bonds are important determinants of protein structure and function in biological systems. Secreted airway mucins contain many disulfide bonds, and breaking these bonds might alter their function and thereby alter mucociliary clearance.

Studies in rabbits and rats found measurable levels of sulfite and S-sulfonates in the upper respiratory tract following inhalation of 10–30 ppm SO₂. Levels of sulfite and S-sulfonates were increased in tracheal washings of rabbits exposed to 10 ppm SO₂ for up to 72 hours ([Gunnison et al., 1981](#)), implying reaction of sulfite with disulfide groups in mucus proteins in the ELF. In addition, tracheal tissue contained elevated levels of S-sulfonates, implicating reaction of sulfite with disulfide groups in tissue proteins. Bronchial tissue from rats had increased levels of sulfites and S-sulfonates when higher concentrations (30 ppm) of SO₂ were employed ([Gunnison et al., 1987b](#)). Under these conditions, no S-sulfonates were found in lung parenchyma, and neither sulfites nor S-sulfonates were found in the plasma. The lack of sulfites and S-sulfonates in the plasma of rats may have been due to these animals' high levels of sulfite oxidase and rapid metabolism of sulfite ([Section 4.2.4](#)). Consistent with ³⁵S rapidly appearing in the blood of ³⁵SO₂-exposed dogs, S-sulfonates were found in plasma of rabbits following 10 ppm SO₂ exposure, providing evidence for absorption of sulfite into the blood of rabbits ([Gunnison et al., 1981](#); [Gunnison and Palmes, 1973](#)). Experiments examining in vivo and ex vivo plasma have shown that sulfite reacts with disulfide bonds in albumin and fibronectin to produce S-sulfonates ([Gregory and Gunnison, 1984](#)).

Exposure of humans to SO₂ also resulted in measurable S-sulfonates in plasma ([Gunnison and Palmes, 1974](#)). In this study, humans were exposed continuously to 0.3–6 ppm SO₂ for up to 120 hours, and plasma levels of S-sulfonates were positively correlated with concentrations of SO₂ inhaled. The regression line for this relationship had a correlation coefficient of 0.61 and the slope was 1.1 nmol/mL of plasma S-sulfonate for each 1-ppm increment in SO₂ concentration. A subacute study measured sulfite plus S-sulfonate content of the lung, liver, and brain of mice exposed to 5, 10, or 20 ppm SO₂, 4 hours/day for 7 days ([Meng et al., 2005a](#)). A concentration-dependent increase in sulfite and S-sulfonate levels was observed. Thus, in humans and mice, the amount of SO₂-derived species in blood and other tissues increases with the concentration in inhaled air. It should also be noted that measurable amounts of sulfite/S-sulfonate were found in tissues of humans and mice inhaling filtered air instead of SO₂ ([Meng et al., 2005a](#); [Gunnison and Palmes, 1974](#)). Besides inhaled SO₂, sulfite is derived from other exogenous, as well as endogenous sources ([Section 4.2.6](#)).

Inhaled SO₂ need not reach the lower airways for SO₂-derived species to be found in the blood. During the 5 full days of SO₂ exposure in the [Gunnison and Palmes \(1974\)](#) study, volunteers were likely at rest or sleeping for much of their exposures. Given that ventilation rates would be relatively low and breathing would be largely nasal ([Section 4.1.2](#)), most inhaled SO₂ would likely be absorbed in the extrathoracic airways ([Section 4.2.2](#)). A number of studies also exposed the surgically isolated upper airways of dogs to ³⁵SO₂ and observed ³⁵S to rapidly appear in the blood and for the concentration in

blood to continuously increase during exposure (e.g., [Yokoyama et al., 1971](#); [Frank et al., 1967](#)). [Frank et al. \(1969\)](#) proposed the majority of SO₂-derived products found in the blood originated from SO₂ absorbed in the upper airways.

In summary, inhaled SO₂ is readily dissolved in the ELF where it exists as a mixture of bisulfite and sulfite with the latter predominating. Bisulfite reacts with disulfide groups forming S-sulfonates; sulfite can diffuse across cell membranes and reach the circulation. Following absorption in the respiratory tract, SO₂-derived products (e.g., sulfite and/or S-sulfonates) are widely distributed throughout the body and have been observed in the blood and urine within 5 minutes of starting an SO₂ exposure of surgically isolated nasal airways. Measurable levels of S-sulfonates have been observed in plasma following inhalation of SO₂ in humans, dogs, mice, and rabbits. Perhaps due to higher levels of hepatic sulfite oxidase relative to other species, sulfites, and S-sulfonates are not found in the plasma of rats. Although the majority of SO₂-derived products remain in the respiratory tract following exposure, extrapulmonary SO₂-derived products are found in the liver, with lesser amounts found in the heart, spleen, kidney, brain, and other tissues. The amount of SO₂-derived species in blood and other tissues increases with the concentration of SO₂ in inhaled air, while the distribution within the body is generally unaffected. A substantial portion of SO₂-derived products appear to be retained within the upper airways, particularly during nasal breathing, with only slow absorption into the blood.

4.2.4 Metabolism

The primary route of sulfite metabolism is by sulfite oxidase-catalyzed enzymatic oxidation to sulfate ([Gunnison, 1981](#)). Because of this pathway, intra-cellular steady-state concentrations of sulfite are low in normal individuals ([Gunnison et al., 1987a](#)). Sulfite oxidase is a molybdenum-containing enzyme that is found in mitochondria. Its distribution varies widely across tissues. While lung tissue has very low sulfite oxidase activity, liver has high sulfite oxidase activity and plays a major role in detoxification of circulating sulfite. [Maier et al. \(1999\)](#) examined the distribution of sulfite oxidase activity in the respiratory tract and liver of four beagle dogs. Sulfite oxidase activity was highest in the liver. The median sulfite oxidase activity in the nose was about 30% of the liver. Median activity levels in the trachea and bronchi were about 20% of the liver and the median activity levels in the lung parenchyma were only 10% of those in the liver. The 1982 AQCD ([U.S. EPA, 1982a](#)) noted that depleting the activity of sulfite oxidase in an animal model through a low-molybdenum diet supplemented with the competitive inhibitor tungsten resulted in a substantial lowering of the lethal dose for intra-peritoneally injected bisulfite. A deficiency in sulfite oxidase activity may lead to

toxicity even in the absence of exogenous sulfite or bisulfite exposures. For example, humans and mice with homozygous genetic defects in the sulfite oxidase protein or in the enzymes required to synthesize the essential molybdenum cofactor develop ultimately lethal neurologic disease attributable to accumulation of endogenous sulfite postnatally (i.e., following loss of maternal protection in utero) ([Johnson-Winters et al., 2010](#); [Reiss et al., 2005](#)).

Sulfite oxidase activity is highly variable among species. Liver sulfite oxidase activity in the rat is 10–20 times that in humans. Rapid metabolism of circulating sulfite to sulfate may explain the lack of sulfite/S-sulfonates found in blood of rats exposed by inhalation to 30 ppm SO₂, whereas these products were found in other species ([Gunnison et al., 1987a](#)). In sulfite oxidase-deficient rats, plasma sulfite levels increase with the severity of the deficiency ([Gunnison et al., 1987b](#)).

[Gunnison and Benton \(1971\)](#) also identified S-sulfonate in blood as a reaction product of inhaled SO₂. S-sulfonates, which are produced by the reaction of bisulfite with disulfide bonds, may be metabolized back to disulfides. Although the enzymatic pathways and cofactors are not clearly established for this repair process, it requires reducing equivalents, and thus, has a metabolic cost.

In summary, the primary route of sulfite metabolism is by sulfite oxidase-catalyzed oxidation into sulfate. The sulfite oxidase levels vary widely among tissues with very low levels found in the lung and high levels found in the liver, which plays a major role in the detoxification of circulating sulfite. Sulfite oxidase activity is also highly variable among species with liver sulfite oxidase activity in rats being 10–20 times greater than in humans.

4.2.5 Elimination

Mechanisms involved in elimination include both desorption of SO₂ from the respiratory tract and the clearance of reaction products from the body.

When the partial pressure of SO₂ on mucosal surfaces exceeds that of the gas phase, such as during expiration, some desorption of SO₂ from respiratory tract lining fluids may be expected. [Speizer and Frank \(1966\)](#) found that on expiration, 12% of the SO₂ absorbed during inspiration was desorbed into the expired air. During the first 15 minutes after the 25- to 30-minute SO₂ exposure, another 3% was desorbed. In total, 15% of the amount of originally inspired and absorbed SO₂ was desorbed from the nasal mucosa. [Frank et al. \(1969\)](#) reported that up to 18% of the SO₂ was desorbed within ~10 minutes after exposure.

SO₂ that does not desorb is transformed to bisulfite/sulfite ([Section 4.2.1](#)). Because the lung tissue has a low activity of sulfite oxidase, diffusion into the circulation may be a more important route of sulfite clearance from the lung than enzyme-catalyzed transformation to sulfates. Within a period of minutes after starting ³⁵SO₂ inhalation exposures, ³⁵S was observed in the blood and urine of dogs and distributed about the body ([Frank et al., 1967](#); [Balchum et al., 1959](#)). At the end of 30–60-minute exposures, 5–18% of the administered ³⁵S was in the blood, and 1–6% had been excreted in the urine by 3 hours post-exposure ([Yokoyama et al., 1971](#); [Frank et al., 1967](#)). The rate of urinary excretion was proportional to the blood concentration, and 92% of the urinary ³⁵S was in the form of sulfate ([Yokoyama et al., 1971](#)). In contrast, S-sulfonates formed in the circulation were reported to have a clearance half-time of 3.2 days ([Gunnison and Palmes, 1973](#)).

In summary, when the partial pressure of SO₂ on mucosal surfaces exceeds that of the gas phase, such as during expiration or following exposure, some desorption of SO₂ from the respiratory tract lining fluids may be expected. SO₂ that does not desorb is transformed to bisulfite/sulfite. Given the low activity of sulfite oxidase in the respiratory tract, sulfite is more likely to diffuse into the circulation or react with tissue constituents than be metabolized to sulfate. Circulating sulfite may subsequently react with constituents of the blood to form S-sulfonates or other species. Sulfite reaching the liver ([Section 4.2.3](#)) is efficiently metabolized to sulfate ([Section 4.2.4](#)). Urinary excretion of sulfate is rapid and proportional to the concentration of SO₂ products in the blood. S-sulfonates are cleared more slowly from the circulation with a clearance half-time of days. The portion of SO₂-derived products that are retained within the respiratory tract are only slowly absorbed into the blood ([Section 4.2.3](#)).

4.2.6 Sources and Levels of Exogenous and Endogenous Sulfite

The primary endogenous contribution of sulfite is from the catabolism of sulfur-containing amino acids (namely, cysteine and methionine). Sulfite may subsequently be metabolized to sulfate in a reaction catalyzed by sulfite oxidase in most tissues, but especially in the liver ([Section 4.2.4](#)). Mean daily sulfate produced following ingestion of cysteine and methionine in the U.S. increases from 70 mg/kg-day in infants (ages 2–6 months) to 100 mg/kg-day in young children (1–3 years old) and then decreases to 30 and 40 mg/kg-day in adult (19–50 years old) females and males, respectively ([IOM, 2005](#)). To facilitate comparison with exogenous sources, a mole of SO₂ can produce a mole of sulfate, but the SO₂ mass is only two-thirds of the sulfate mass.

Sulfite is also added to foods because it has antioxidant and antimicrobial properties ([Vandevijvere et al., 2010](#); [Gunnison, 1981](#)). In a study considering actual food consumption of Belgian adults and measured sulfite levels in food, [Vandevijvere et al. \(2010\)](#) observed a wide distribution in exogenous sulfite from ingestion. Expressed in terms of SO₂ equivalents, rates of exogenous sulfite ingestion may be described by a log-normal distribution with a median intake of 0.14 SO₂ mg/kg-day and a geometric standard deviation of 2.15. Individuals at the 5th and 95th percentiles of this distribution are estimated to consume 0.04 and 0.49 SO₂ mg/kg-day. In a comparison of theoretical food-consumption data with maximum permissible SO₂/sulfites to foods, the Belgian adults in the [Vandevijvere et al. \(2010\)](#) study had a similar potential sulfite intake to U.S. adults. The estimated intake for children could be in the range of that for adults or less due to the likely minimal consumption of sulfite sources such as wine. Endogenous sulfite from catabolism of ingested sulfur-containing amino acids far exceeds exogenous sulfite from ingestion of food additives [by 140 and 180 times in adult (19–50 years old) females and males, respectively, and by 500 times or more in young children (1–3 years old)].

Exogenous sulfite may also be derived from SO₂ inhalation. For the purposes of comparisons herein, all inhaled SO₂ is assumed to contribute to systemic sulfite levels. In reality, as discussed in [Section 4.2.3](#), the majority of SO₂-derived products from SO₂ inhalation are retained in the respiratory tract and may be detected there for up to a week following inhalation. The potential contribution of inhaled SO₂ to systemic sulfite levels varies with age, activity level, and SO₂ concentration. Using median and 97.5th percentile daily ventilation rates from [Brochu et al. \(2011\)](#), normal-weight adults (25–45 years of age) are estimated to receive 0.004 and 0.006 mg SO₂ per kg body mass, respectively, from a full day exposure to 5 ppb SO₂. As an upper-bound estimate for ambient exposure in most locations, a full-day exposure to 75 ppb SO₂ (the level of the current National Ambient Air Quality Standard for SO_x) would result in 0.053 SO₂ mg/kg-day and 0.085 SO₂ mg/kg-day for adults having median and 97.5th percentile ventilation rates, respectively. The estimated daily SO₂ intake (mg/kg-day) would be roughly 1.5 times greater in children (7–10 years of age) and doubled in infants (0.22–0.5 years of age) due to the greater ventilation rate per body mass of children compared to adults (25–45 years of age). Even upper-bound sulfite levels from inhalation (75 ppb SO₂, 24 hours, 97.5th percentile ventilation) are far less than those derived from catabolism of sulfur-containing amino acids, by 230 to 300 times in adults (25–45 years) and nearly 500 times in young children (1–3 years).

Comparison of sulfite derived from SO₂ inhalation with that from ingestion of food additives is more complicated. In adults (25–45 years), sulfite intake (mg/kg-day) from inhalation (75 ppb SO₂, 24 hours, 97.5th percentile ventilation) is 1.6 times lower than

median sulfite intake from ingestion of food additives. In children (<10 years), assuming similar levels of sulfite intake as adults, sulfite intake from inhalation (75 ppb SO₂, 24 hours, 97.5th percentile ventilation) is approximately the same as median sulfite intake from ingestion of food additives. However, ingested sulfite absorbed into the blood goes directly to the liver where much of it is metabolized into sulfate. The majority of sulfite derived from inhalation that enters the blood is rapidly distributed [as either sulfite or S-sulfonate ([Yokoyama et al., 1971](#); [Balchum et al., 1959](#))] about the body with around a quarter of total blood flow going to the liver ([ICRP, 2002](#)) where there is a high activity of sulfite oxidase compared to other tissues. For lower exposure concentrations and durations than considered above, sulfite (and/or S-sulfonate) levels in the blood following SO₂ inhalation could exceed those from ingestion of food additives, particularly in children.

In summary, exogenous sources contribute hundreds of times lower amounts of sulfite than the catabolism of sulfur-containing amino acids, when averaged across the entire body. Sulfite and sulfate derived from the catabolism of sulfur-containing amino acids are distributed broadly and do not accumulate in respiratory tract tissues. Following ingestion of sulfite-containing food additives, sulfite enters the circulation and is subject to first pass clearance in the liver where it is metabolized to sulfate. Following inhalation, a substantial portion of SO₂-derived products accumulate and are retained within the respiratory tract. SO₂-derived products that enter the circulation are rapidly distributed throughout the body, appear primarily in the liver, and are excreted via the urine ([Section 4.2.5](#)).

4.3 Mode of Action of Inhaled Sulfur Dioxide

This section describes the biological pathways that potentially underlie health effects resulting from short-term and long-term exposure to SO₂. Extensive research carried out over several decades in humans and in laboratory animals has yielded much information about these pathways. This section is not intended to be a comprehensive overview, but rather, it updates the basic concepts derived from the SO₂ literature presented in the 1982 AQCD ([U.S. EPA, 1982a](#)) and the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) and introduces the recent relevant literature. While this section highlights findings of studies published since the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), earlier studies that represent the current state of the science are also discussed. Studies conducted at more environmentally relevant concentrations of SO₂ (i.e., ≤2 ppm, see [Section 1.2](#)) are of greater interest because biological pathways responsible for effects at higher concentrations may not be identical to those occurring at lower concentrations. Some studies at higher concentrations are included if they were early demonstrations of key biological pathways

or if they are recent demonstrations of potentially important new pathways. This information is used to develop a mode of action framework for inhaled SO₂ that serves as a guide to interpreting health effects evidence presented in [Chapter 5](#).

Mode of action refers to a sequence of key events, endpoints, and outcomes that result in a given toxic effect ([U.S. EPA, 2005a](#)). Elucidation of mechanism of action provides a more detailed understanding of key events, usually at the molecular level ([U.S. EPA, 2005a](#)). The framework developed in this chapter will include some mechanistic information on initiating events at the molecular level, but will mainly focus on the effects of SO₂ at the cellular, tissue, and organism level.

SO₂ is a highly reactive gas. At physiologic pH, its hydrated forms include sulfurous acid, bisulfite, and sulfite, with the latter species predominating. Sulfite is a strong nucleophilic anion that readily reacts with nucleic acids, proteins, lipids, and other classes of biomolecules. It participates in many important types of reactions including sulfonation (sulfitolysis) and auto-oxidation with the generation of free radicals. This latter reaction may be responsible for the induction of oxidative stress that occurs as a result of exposure to SO₂.

As described in [Section 4.2](#), SO₂ is a water-soluble gas that is absorbed almost entirely in the upper respiratory tract. However, under conditions of mouth breathing and exercise, some SO₂ may penetrate to the tracheobronchial region. The main effects of SO₂ inhalation are seen at the sites of absorption (i.e., the respiratory tract) and include (1) activation of sensory nerves in the respiratory tract resulting in neural reflex responses, (2) injury to airway mucosa, and (3) increased airway responsiveness and allergic inflammation. Effects outside the respiratory tract may occur at very high concentrations of inhaled SO₂. Biologic pathways involved in mediating these responses to inhaled SO₂ are discussed below. In addition, a brief synopsis of pathways involved in mediating the effects of endogenous SO₂/sulfite is presented.

4.3.1 Activation of Sensory Nerves in the Respiratory Tract

SO₂ is classified as a sensory (or nasal) irritant in mice, guinea pigs, rats, and humans ([Alarie, 1973](#)). As such, it may stimulate trigeminal nerve endings when inhaled by the nose, which results in an inhibition of respiration. It may also stimulate trigeminal nerves in the larynx, which results in coughing, and in the cornea, which induces tearing. Other reflexes stimulated by trigeminal nerve endings include decreased heart rate, peripheral vasoconstriction, closure of the glottis, closure of the nares, and increased nasal flow resistance. These responses are variable among species. Increased nasal flow resistance has been demonstrated in humans breathing SO₂ gas through the nose. Furthermore,

desensitization of the respiratory rate response occurs with repeated exposure. Most sensory (or nasal) irritants, including SO₂, also cause bronchoconstriction, but at concentrations higher than those stimulating nerve endings in the nose.

SO₂ is also classified as a pulmonary (or bronchial) irritant that evokes reflex reactions through effects on pulmonary nerve endings ([Alarie, 1973](#)). These reactions usually include an increase in respiratory rate accompanied by a decrease in tidal volume, sometimes preceded by coughing and brief apnea, and sometimes accompanied by bronchoconstriction. These responses have been observed in guinea pigs and cats breathing via a tracheal cannula, which bypasses the nose. In the cat, SO₂ exposure increased the activity of vagal afferent fibers by either stimulating or sensitizing tracheobronchial receptors on the nerve endings. SO₂ also increased airway resistance in guinea pigs and humans breathing through the nose, mouth, and/or tracheal cannula. Increased airway resistance may occur via a variety of mechanisms including accumulation of secretions, inflammatory changes of the airway walls, collapsing airways, and constrictions of the central and peripheral airways. Constriction may be due to direct action on the smooth muscle, axonal reflexes, vagal nerve stimulation, and release of mediators such as histamine.

Continuous or repeated exposure to inhaled SO₂ has a different pattern of responses in different species ([Alarie, 1973](#)). In guinea pigs, the increase in airway resistance rose to a plateau upon exposure and decreased to baseline with cessation of exposure. In humans and dogs, resistance increased with exposure but decreased after 10 minutes (humans) or 3 minutes (dogs) despite the continuous presence of the gas. Studies in adults with asthma demonstrated a different pattern. When exposure to SO₂ occurred during a 30-minute period with continuous exercise, the response to SO₂ developed rapidly and was maintained throughout the 30-minute exposure ([Kehrl et al., 1987](#); [Linn et al., 1987](#); [Linn et al., 1984c](#)). Sequential exposures in nonasthmatic human subjects and in cats resulted in a decreased response to SO₂ in the second exposure compared with the first, indicative of desensitization.

Early experiments demonstrated that SO₂-induced reflexes were mediated by cholinergic parasympathetic pathways involving the vagus nerve and inhibited by atropine ([Grunstein et al., 1977](#); [Nadel et al., 1965a, b](#)). Bronchoconstriction was found to involve smooth muscle contraction because β -adrenergic agonists such as isoproterenol reversed the effects. Rapid shallow breathing was observed in SO₂-exposed tracheotomized cats (bypassing the nose). Histamine was proposed to play a role in SO₂-induced bronchoconstriction ([U.S. EPA, 1982a](#)), but this hypothesis remains unsupported. Hydrogen ions, sulfurous acid, sulfite, and bisulfite are all putative mediators of the reflex responses ([Gunnison et al., 1987a](#)). In particular, sulfite-mediated sulfitolysis of

disulfides present in receptor proteins on sensory nerve fibers has been postulated because S-sulfonate formation may potentially disrupt protein structure or function ([Alarie, 1973](#)).

More recent experiments in animal models conducted since 1982 have demonstrated that both cholinergic and noncholinergic mechanisms may be involved in SO₂-induced effects. In two studies using bilateral vagotomy, vagal afferents were found to mediate the immediate ventilatory responses to SO₂ ([Wang et al., 1996](#)), but not the prolonged bronchoconstrictor response ([Barthelemy et al., 1988](#)). Other studies showed that atropine failed to block SO₂-induced bronchoconstriction, and that a local axon reflex resulting in C-fiber secretion of neuropeptides (i.e., neurogenic inflammation) was responsible for the effect ([Hajj et al., 1996](#); [Atzori et al., 1992](#)). Neurogenic inflammation has been shown to play a key role in animal models of airway inflammatory disease ([Groneberg et al., 2004](#)). Furthermore, in isolated perfused and ventilated guinea pig lungs, bronchoconstriction to SO₂ was biphasic. The initial phase was mediated by a local axon reflex involving the release of the neuropeptide calcitonin gene-related peptide from sensory nerves, while the later phase involved other mechanisms ([Bannenberget al., 1994](#)).

In humans, the mechanisms responsible for SO₂-induced bronchoconstriction are not entirely understood. In nonasthmatic subjects, near complete attenuation of bronchoconstriction has been demonstrated using the anticholinergic agents atropine and ipratropium bromide ([Yildirim et al., 2005](#); [Snashall and Baldwin, 1982](#); [Tan et al., 1982](#)). However, in asthmatic subjects, these same anticholinergic agents ([Field et al., 1996](#); [Myers et al., 1986a](#)), as well as short- and long-acting β_2 -adrenergic agonists ([Gong et al., 1996](#); [Linn et al., 1988](#)), theophylline ([Koenig et al., 1992](#)), cromolyn sodium ([Myers et al., 1986a](#)), nedocromil sodium ([Bigby and Boushey, 1993](#)), and leukotriene receptor antagonists ([Gong et al., 2001](#); [Lazarus et al., 1997](#)) only partially blocked SO₂-induced bronchoconstriction. That none of these therapies have been shown to completely attenuate the effects of SO₂ implies the involvement of both parasympathetic pathways and inflammatory mediators in asthmatic individuals. Strong evidence of this was borne out in a study by [Myers et al. \(1986a\)](#) in which asthmatic adults were exposed to SO₂ following pretreatment with cromolyn sodium (a mast cell stabilizer), atropine (a muscarinic receptor antagonist), and the two medications together. While both treatments individually provided some protection against the bronchoconstrictive effects of SO₂, there was a much stronger and statistically significant effect following concurrent administration of the two medications. Besides mast cell stabilization, cromolyn sodium may also reduce the activity of lung irritant receptors ([Harries et al., 1981](#)), providing an alternative mechanism for the reduction in SO₂-induced bronchoconstriction observed.

It has been proposed that inflammation contributes to the enhanced sensitivity to SO₂ seen in asthmatic human subjects by altering autonomic responses ([Tunnicliffe et al., 2001](#)), enhancing mediator release ([Tan et al., 1982](#)), and/or sensitizing C-fibers and rapidly adapting receptors ([Lee and Widdicombe, 2001](#)). Whether local axon reflexes also play a role in SO₂-induced bronchoconstriction in asthmatic individuals is not known ([Groneberg et al., 2004](#); [Widdicombe, 2003](#); [Lee and Widdicombe, 2001](#)). However, differences in respiratory tract innervation between rodents and humans suggest that C-fiber-mediated neurogenic inflammation may be unimportant in humans ([Groneberg et al., 2004](#); [Widdicombe, 2003](#); [Widdicombe and Lee, 2001](#)). Furthermore, enhanced sensitivity to SO₂ in asthmatic individuals may be related to genetic polymorphisms of inflammatory mediators, such as TNF- α ([Winterton et al., 2001](#)).

Studies conducted in vitro provide support for SO₂ exposure-mediated effects that involve inflammatory cells. It is known that sulfite exposure of cultured rat basophil leukemia cells, a mast cell analog, causes immunoglobulin E (IgE)-independent degranulation, release of histamine, serotonin and other mediators, and intra-cellular production of reactive oxygen species ([Collaco et al., 2006](#)). In addition, peroxidases, such as neutrophil myeloperoxidase, oxidize bisulfite anion to several radical species that in turn attack proteins ([Ranguelova et al., 2013](#); [Ranguelova et al., 2012](#)). This represents a potentially important new toxicological pathway for sulfite, especially in the presence of neutrophilic and/or eosinophilic inflammation.

Irritant responses are indicative of a chemical's ability to damage the respiratory tract ([Alarie and Luo, 1986](#); [Alarie, 1981](#)). In the case of sensory irritation, there is a characteristic decrease in respiratory rate, which is often used to set standards for occupational exposures. Chemicals that are pulmonary irritants often lead to rapid shallow breathing. They typically induce pulmonary edema or congestion if inhaled for a long enough period of time. Some chemicals are both sensory and pulmonary irritants and pulmonary irritation may occur at concentrations below which sensory irritation occurs. In the case of SO₂, a concentration-dependent hierarchy of effects has been noted in humans ([Kane et al., 1979](#)). Lethal or extremely severe injury to the respiratory tract has been reported at and above 190 ppm. Intolerable sensory irritation and respiratory tract injury that may occur with extended exposure has been associated with 10–15-minute exposures to 30–100 ppm SO₂, and tolerable sensory irritation has been associated with 10-minute exposures to 5–11.5 ppm SO₂. Minimal sensory irritation has been associated with exposures at and below 1 ppm. Increased airway resistance, likely due to pulmonary irritation and reflex bronchoconstriction, has been observed in adults without asthma at 5 ppm while at rest and at 1 ppm SO₂ while exercising ([Arts et al., 2006](#)). However, lung function changes have been observed at concentrations of SO₂ lower than 1 ppm in exercising adults with asthma. Thus, pulmonary irritation may occur

at levels of SO₂ below those that cause sensory irritation, especially in exercising adults with asthma.

In summary, SO₂ acts as both a sensory and a pulmonary irritant through activation of sensory nerves in the respiratory tract resulting in neural reflex responses. This occurs in a variety of species, including humans. Pulmonary irritant responses due to SO₂ exposure result in reflex bronchoconstriction, especially in adults with asthma. Both cholinergic parasympathetic pathways involving the vagus nerve and inflammation contribute to reflex bronchoconstriction in asthmatic individuals.

4.3.2 Injury to Airway Mucosa

A common feature of irritant gases, including SO₂, is the capacity to injure airway mucosa, resulting in decreased epithelial barrier function, inflammation, and compromised ciliary function ([Carson et al., 2013](#)). Despite being the initial site of SO₂ absorption and having low activity of sulfite oxidase, the respiratory tract of healthy humans is thought to be capable of detoxifying 5 ppm inhaled SO₂ ([Gunnison et al., 1987a](#)). In fact, exposure to 0.5–2 ppm SO₂ for 4 hours did not result in any measurable changes in biomarkers of oxidative stress or inflammation in exhaled breath condensate (EBC) or nasal lavage fluid (NALF) from healthy adults subjected to two periods of moderate exercise ([Raulf-Heimsoth et al., 2010](#)). In addition, no changes in nasal lining fluid ascorbic acid or uric acid levels were observed following 1-hour exposure of adults with asthma to 0.2 ppm SO₂ ([Tunncliffe et al., 2003](#)).

However, respiratory tract injury has been observed in humans exposed for extended periods to SO₂ concentrations of 30 ppm and greater. In animal models, airway injury and histopathological changes, such as mucous cell metaplasia and intra-mural fibrosis, have generally been observed following chronic exposure to SO₂ concentrations of 10 ppm and higher ([U.S. EPA, 2008d](#)). Rats exposed to 20 ppm SO₂ for several weeks exhibit fibrotic remodeling of airway epithelium and mucus hypersecretion, key features of chronic obstructive pulmonary disease and chronic asthma in humans ([Wagner et al., 2006](#)). Inflammatory changes have been noted in some animal models following subacute exposure to 5–100 ppm SO₂ ([U.S. EPA, 2008d](#)). Impaired mucociliary clearance has also been demonstrated at high concentrations of SO₂. In humans, nasal mucus flow was decreased during a 5-hour exposure to 5 and 25 ppm SO₂ ([Gunnison et al., 1981](#)). Impaired mucus flow in the trachea has been observed in rats exposed subacutely to 11.4 ppm SO₂ and in dogs exposed chronically to 1 ppm SO₂ ([Gunnison et al., 1981](#); [Hirsch et al., 1975](#)). Whether these effects were due to compromised ciliary function or

altered properties of the mucus due to sulfite-mediated sulfitolysis of disulfide bonds in mucus was not investigated.

Recent studies provide additional insight. An ultrastructural examination of nasal biopsy tissue by freeze fracture microscopy was conducted in healthy humans exposed to 0.75 ppm SO₂ for 2 h ([Carson et al., 2013](#)). Evidence of fragmentation of the tight junctional complex and polymorphonuclear infiltrate was reported although no effects on ciliary membranes were observed. These subtle responses suggest a slight decrease in barrier function due to acute SO₂ exposure at this level. Furthermore, a subacute exposure of rats to 2.67 ppm SO₂ (6 hours/day, 7 days) resulted in altered lung mRNA levels for inducible nitric oxide synthase (involved in inflammation) and for bax (or B-cell lymphoma 2-like protein 4; involved in regulating apoptosis) ([Sang et al., 2010](#)). In this study, gene expression changes were also found in the heart and they were more pronounced than in the lung. These results suggest that, despite low sulfite oxidase activity, the respiratory tract may be more resistant than the heart to the effects of inhaled SO₂.

In summary, exposure to SO₂ results in injury to airway mucosa, especially at higher concentrations and following extended periods of exposure. There is little evidence of injury or inflammation in response to acute exposures to concentrations of 2 ppm SO₂ or less in healthy human subjects. However, one new study found subtle histopathological changes at the ultrastructural level following a 2-hour exposure to 0.75 ppm SO₂. New evidence also suggests subtle changes in the lung related to inflammation and apoptosis in rats exposed over several days to 2.67 ppm SO₂.

4.3.3 Modulation of Airway Responsiveness and Allergic Inflammation

Respiratory irritants, including SO₂, are thought to be a major cause of occupational asthma ([Baur et al., 2012](#); [Andersson et al., 2006](#)). Both peak high-level exposures and low-level persistent exposures to respiratory irritants have been associated with the development of asthma. The propensity for airways to narrow following inhalation of some stimuli is termed airway responsiveness. The term airway hyperresponsiveness (AHR) is generally used in cases where airway responsiveness to methacholine or histamine is assessed and the provocative concentration is determined to be sufficiently low to classify the subjects as having AHR based on criteria such as [ATS \(2000a\)](#). Along with symptoms, variable airway obstruction, and airway inflammation, AHR is a primary feature in the clinical definition and characterization of asthma severity ([Reddel et al., 2009](#)).

Studies in several different animal species have shown that a single exposure to SO₂ at a concentration of 10 ppm or less failed to increase airway responsiveness following a challenge agent ([U.S. EPA, 2008d](#)). However, in an animal model of allergic airway disease, SO₂ exposure enhanced airway responsiveness. In this study, sheep previously sensitized and challenged with *Ascaris suum* extract were exposed to 5 ppm SO₂ for 4 hours ([Abraham et al., 1981](#)). Airway responsiveness to carbachol was increased at 24 hours, but not immediately, after SO₂ exposure. This response was not observed in sheep that had not been sensitized and challenged with *Ascaris suum* extract. The mechanism underlying the SO₂-mediated increases in airway responsiveness was not investigated in this study. However, this response could have resulted from sensitization of vagal irritant receptors, greater sensitivity of smooth muscle to bronchoconstriction agents, or enhanced concentrations of bronchoconstriction agents reaching the receptors or bronchial smooth muscle. The delayed nature of the response points to a possible role of inflammation in increasing airway responsiveness.

Two controlled human exposure studies in adults with asthma provide further evidence of increased responsiveness to an allergen when exposure to SO₂ was in combination with nitrogen dioxide (NO₂). In one of these studies, exposure to 0.2 ppm SO₂ or 0.4 ppm NO₂ alone did not affect airway responsiveness to house dust mite allergen immediately after a 6-hour exposure at rest ([Devalia et al., 1994](#)). However, following exposure to the two pollutants in combination, subjects demonstrated an increased response to the inhaled allergen. [Rusznak et al. \(1996\)](#) confirmed these results in a similar study and found that increased responsiveness to dust mites persisted up to 48 hours post-exposure. These results provide evidence that exposure to SO₂ and NO₂ in combination elicits an increase in airway responsiveness to an allergen. This effect is longer in duration than other effects typically associated with exposure to SO₂.

Several other studies have examined the effects of SO₂ exposure on allergic inflammation. One of these was a controlled human exposure study of adults with asthma. Subjects were exposed for 10 minutes to 0.75 ppm SO₂ while exercising at a moderate level ([Gong et al., 2001](#)). In addition to changes in lung function and symptoms, there was a statistically significant increase in eosinophil count in induced sputum 2 hours post-exposure. Pretreatment with a leukotriene receptor antagonist dampened these responses, implicating a role for leukotrienes in mediating SO₂ exposure-induced effects.

The other studies investigated the effects of repeated exposure to SO₂ on inflammatory and immune responses in an animal model of allergic airways disease. [Li et al. \(2007\)](#) demonstrated that exposure of ovalbumin-sensitized rats to 2 ppm SO₂ for 1 hour followed by challenge with ovalbumin each day for 7 days resulted in an increased

number of inflammatory cells in bronchoalveolar lavage fluid (BALF) and an enhanced histopathological response compared with rats treated with SO₂ or ovalbumin alone. Similarly, inter-cellular adhesion molecule 1 (ICAM-1), a protein involved in regulating inflammation, and mucin 5AC glycoprotein (MUC5AC), a mucin protein, were upregulated in lungs and trachea to a greater extent in rats treated both with SO₂ and ovalbumin. A follow-up study involving the same exposure regimen (2 ppm SO₂ for 1 hour) in the same allergic animal model (rats sensitized and challenged with ovalbumin) also found that repeated SO₂ exposure enhanced inflammatory and allergic responses to ovalbumin ([Li et al., 2014](#)). Numbers of eosinophils, lymphocytes, and macrophages were greater in the BALF of SO₂-exposed and ovalbumin-treated animals than in animals treated only with ovalbumin. In addition, SO₂ exposure enhanced upregulation and activation of nuclear factor kappa-light-chain-enhancer of activated B cells (NFκB), a transcription factor involved in inflammation, and upregulation of the cytokines interleukin-6 (IL-6) and interleukin-4 (IL-4) in lung tissue. Furthermore, BALF levels of IL-6 and IL-4 were increased to a greater extent in SO₂-exposed and ovalbumin-treated animals compared with ovalbumin treatment alone. These results indicate that repeated SO₂ exposure enhanced activation of the NFκB inflammatory pathway and upregulation of inflammatory cytokines in ovalbumin-treated animals. Furthermore, SO₂ exposure enhanced the effects of ovalbumin on levels of inter-feron gamma (IFN-γ) (decreased) and IL-4 (increased) in BALF and on IgE levels in serum (increased). Because levels of IL-4 are often indicative of T helper 2 (Th2) status and levels of IFN-γ are indicative of a T helper 1 (Th1) status, these results suggest a shift in Th1/Th2 balance towards Th2 in rats made allergic to ovalbumin, an effect that was exacerbated by SO₂ exposure. These Th2-related changes are consistent with the observed increases in serum IgE and BALF eosinophils in ovalbumin-treated animals, effects that were also enhanced by SO₂ exposure. Taken together, these results indicate that repeated exposure to SO₂ exacerbated inflammatory and allergic responses in this animal model. It should be noted, however, that Group 2 innate lymphoid cells can mediate Type 2 immunity, as has been described for O₃-mediated responses in mice ([Ong et al., 2016](#)). Whether Group 2 innate lymphoid cells mediate effects of inhalation of SO₂, which like O₃ is an irritant gas, is unexplored.

Two other follow-up studies by the same laboratory examined the effects of inhaled SO₂ on the asthma-related genes encoding epidermal growth factor (EGF), epidermal growth factor receptor (EGFR), and cyclooxygenase-2 (COX-2), and on apoptosis-related genes and proteins in this same model based on sensitization with ovalbumin ([Xie et al., 2009](#); [Li et al., 2008](#)). While EGF and EGFR are related to mucus production and airway remodeling, COX-2 is related to apoptosis and may play a role in regulating airway inflammation. SO₂ exposure enhanced the effects of ovalbumin in this model, resulting in greater increases in mRNA and protein levels of EGF, EGFR, and COX-2 in the trachea

compared with ovalbumin treatment alone. SO₂ exposure enhanced other effects of ovalbumin in this model, resulting in a greater decline in mRNA and protein levels of tumor protein p53 (p53) and bax and a greater increase in mRNA and protein levels of B-cell lymphoma 2 (bcl-2) in the lungs compared with ovalbumin challenge alone. The increased ratio of bcl-2:bax, an indicator of susceptibility to apoptosis, observed following ovalbumin challenge, was similarly enhanced by SO₂. Thus, repeated exposure to SO₂ may impact numerous processes involved in inflammation and/or airway remodeling in allergic airway disease.

The effects of repeated SO₂ exposure on the development of an allergic phenotype and altered physiologic responses in naive animals was examined in two studies in which SO₂ exposure preceded allergen sensitization. Repeated exposure of guinea pigs to SO₂ promoted allergic sensitization and subsequently enhanced allergen-induced bronchial obstruction, as reported by [U.S. EPA \(2008d\)](#). [Riedel et al. \(1988\)](#) examined the effect of SO₂ exposure on local bronchial sensitization to inhaled antigen. Guinea pigs were exposed by inhalation to 0.1, 4.3, and 16.6 ppm SO₂ for 8 hours/day for 5 days. During the last 3 days, SO₂ exposure was followed by exposure to nebulized ovalbumin for 45 minutes. Following bronchial provocation with inhaled ovalbumin (0.1%) 1 week later, bronchial obstruction was measured by examining the respiratory loop obtained by whole-body plethysmography. In addition, specific antibodies against ovalbumin were measured in serum and BALF. Results showed significantly higher bronchial obstruction in animals exposed to SO₂ (at all concentration levels) and ovalbumin, compared with animals exposed only to ovalbumin. In addition, significant increases in antiovalbumin immunoglobulin G (IgG) antibodies were detected in BALF lavage fluid of animals exposed to 0.1, 4.3, and 16.6 ppm SO₂ and in serum from animals exposed to 4.3 and 16.6 ppm SO₂ compared with controls exposed only to ovalbumin. These results demonstrate that repeated exposure to SO₂ enhanced allergic sensitization in the guinea pig at a concentration as low as 0.1 ppm. In a second study, guinea pigs were exposed to 0.1 ppm SO₂ for 5 hours/day for 5 days and sensitized with 0.1% ovalbumin aerosols for 45 minutes on Days 4 to 5 ([Park et al., 2001](#)). One week later, animals were subjected to bronchial challenge with 0.1% ovalbumin and lung function was evaluated 24 hours later by whole-body plethysmography. Results demonstrated a significant increase in enhanced pause, a measure of airway obstruction, in animals exposed to SO₂ and ovalbumin but not in animals treated with ovalbumin or SO₂ alone. Results also demonstrated increased numbers of eosinophils in lavage fluid and an infiltration of inflammatory cells, bronchiolar epithelial cell damage, and plugging of the airway lumen with mucus and cells in the bronchial tissues of animals treated with both SO₂ and ovalbumin, but not in animals treated with ovalbumin or SO₂ alone. These experiments indicate that repeated exposure to near ambient levels of SO₂ plays a role in allergic sensitization and also exacerbates allergic inflammatory responses in the guinea pig.

Furthermore, increases in bronchial obstruction observed in both studies suggest that repeated SO₂ exposure increased airway responsiveness.

Longer term exposure of naive newborn rats to SO₂ (2 ppm, 4 hours/day for 28 days) resulted in altered cytokine levels that suggest a shift in Th1/Th2 balance towards Th2 ([Song et al., 2012](#)). Th2 polarization is one of the steps involved in allergic sensitization. It should be noted, however, that Group 2 innate lymphoid cells can mediate Type 2 immunity, as has been described for O₃-mediated responses in mice ([Ong et al., 2016](#)). Whether Group 2 innate lymphoid cells mediate effects of inhalation of SO₂, which like O₃ is an irritant gas, is unexplored. In naive animals exposed to SO₂, levels of IL-4, which is indicative of a Th2 response, were increased and levels of IFN- γ , indicative of a Th1 response, were decreased in BALF. In ovalbumin-sensitized newborn rats, SO₂ exposure resulted in a greater enhancement of lavage fluid IL-4 and an increase in serum IL-4 levels compared with ovalbumin-sensitization alone. In addition, SO₂ exposure led to increased bronchial obstruction and airway remodeling, as indicated by increased content of airway smooth muscle, in the ovalbumin-sensitized animals. Stiffness and contractility of airway smooth muscle was assessed in vitro using cells from experimentally treated animals. In allergic rats, both stiffness and contractility were increased as a result of SO₂ exposure, suggesting an effect on the biomechanics of airway smooth muscle. This study provides evidence for allergic sensitization by SO₂ in naive newborn rats and for enhanced allergic inflammation, increased airway responsiveness, and airway remodeling in SO₂-exposed allergic newborn rats.

Supportive evidence that SO₂ may promote allergic sensitization is provided by a study in mice that were first treated with sodium sulfite and then sensitized and challenged with house dust mite allergen ([Lin et al., 2011a](#)). Sulfite is formed in ELF following inhalation of SO₂ ([Section 4.2.1](#)). Repeated intra-nasal treatment with 10 μ L of a 5-mM solution of sodium sulfite aggravated inflammation (measured by histopathology) and allergic sensitization in this model. Specific IgE levels were higher in sulfite-treated and allergen-challenged animals compared with either sulfite treatment or allergen challenge alone. Specific IgG2 α levels, indicative of a Th1 response, were decreased as a result of sulfite treatment in house dust mite-challenged mice. In addition, inter-leukin-5 (IL-5) levels, indicative of a Th2 response, and the ratio of IL-5:IFN- γ , a marker of Th2 polarization, were higher in lung tissue from sulfite-treated and allergen-challenged mice compared with either sulfite treatment or allergen challenge alone.

Mixtures of SO₂ and other criteria pollutants have also been shown to modulate airway responsiveness and/or allergic inflammation. As discussed above, increased airway responsiveness to house dust mite allergen occurred in human subjects with mild allergy and asthma immediately following 6 hours of concurrent exposure to 0.2 ppm SO₂ and

0.4 ppm NO₂, but not to either pollutant alone ([Rusznak et al., 1996](#); [Devalia et al., 1994](#)). This effect persisted for 48 hours. Recently, the effects of simulated downwind coal combustion emissions (SDCCE), which contains SO₂, on allergic airway responses was investigated in mice ([Barrett et al., 2011](#)). Mice were sensitized and challenged with ovalbumin and exposed for 6 hours/day for 3 days to several concentrations of SDCCE with and without a particle filter. Concentration of SO₂ in the highest exposure was 0.2 ppm. Other gases present in this exposure were NO₂ (0.29 ppm), NO (0.59 ppm), and carbon monoxide (0.02 ppm). SDCCE exposure was followed by another challenge with ovalbumin in some animals. Results demonstrated that both the particulate and the gaseous phases of SDCCE exacerbated allergic airways responses. Airway responsiveness to methacholine (measured by the forced oscillation technique) was enhanced by the gaseous phase of SDCCE in mice that were challenged with ovalbumin after SDCCE exposure. While results of this study are consistent with SO₂ playing a role in enhancing allergic responses and increases in airway responsiveness, a role for other components in the mixture cannot be ruled out.

In summary, a growing body of evidence supports a role for SO₂ in increasing airway responsiveness and/or allergic inflammation in animal models of allergic airway disease, as well as in asthmatic individuals. Some responses in asthmatic individuals were observed only when exposure to SO₂ occurred in combination with NO₂. In animal studies, repeated or prolonged exposure to SO₂ promoted allergic sensitization. One study in newborn allergic rats suggested that airway remodeling may contribute to increases in airway responsiveness following prolonged exposure to SO₂.

4.3.4 Induction of Systemic Effects

As described in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), two controlled human exposure studies reported that acute exposure to 0.2 ppm SO₂ resulted in changes in heart rate variability in healthy adults and in asthmatic adults ([Routledge et al., 2006](#); [Tunnicliffe et al., 2001](#)). More recently, altered parasympathetic regulation of heart rate was reported in rats exposed to 5 ppm SO₂ during the peri-natal and post-natal period ([Woerman and Mendelowitz, 2013a, b](#)). Whether these responses were due to activation of sensory nerves in the respiratory tract resulting in a neural reflex response and altered autonomic function or some other mechanism is not known.

Numerous studies over several decades have reported other extrapulmonary effects of inhaled SO₂ ([U.S. EPA, 2008d](#)). Most of these occur at concentrations far higher than those measured in ambient air. As discussed in [Section 4.2.3](#), studies in mice and humans demonstrating the presence of sulfite and S-sulfonates in blood and tissues outside of the

respiratory tract point to the likely role of circulating sulfite in mediating these responses. A subacute study measured sulfite plus S-sulfonate content of the lung, liver, and brain of mice exposed to 5, 10, or 20 ppm SO₂ for 4 hours/day for 7 days ([Meng et al., 2005a](#)) and found a concentration-dependent increase. Similarly, exposure of human subjects to 0.3–6 ppm SO₂ for up to 120 hours resulted in the appearance in the plasma of sulfite plus S-sulfonates ([Gunnison and Palmes, 1974](#)). The relationship between sulfite/sulfonate concentration and chamber SO₂ concentration was linear (regression coefficient of 0.61) with a slope of 1.1 nmol/mL of plasma S-sulfonate for each 1-ppm increment in SO₂ concentration. These results indicate that prolonged (i.e., hours to days) exposure to as low as 0.3 ppm SO₂ results in measurable amounts of circulating sulfite in humans. The relationship between circulating sulfite/S-sulfonate and extrapulmonary effects of inhaled SO₂ has not yet been explored in human subjects.

Because the activity of sulfite oxidase is variable among species, the degree of sensitivity to SO₂-mediated effects is likely to be variable among species. For example, sulfite oxidase in rats is 10–20 times greater than in humans and 3–5 times greater than in rabbits or rhesus monkeys ([Gunnison et al., 1987a](#); [Gunnison, 1981](#)). Thus, the toxicity of SO₂ may be less in rats due to more rapid metabolism of sulfite to sulfate.

Systemic effects are likely due to oxidative stress, possibly from sulfite auto-oxidation. Alternatively, sulfite-mediated S-sulfonate formation may disrupt protein function, and metabolic reduction of S-sulfonates may alter reduction-oxidation (redox) status. Moreover, sulfite may serve as a substrate for peroxidases, such as myeloperoxidase and eosinophil peroxidase, to produce free radicals, as has been demonstrated in neutrophils and eosinophils ([Rangelova et al., 2013](#); [Rangelova et al., 2012](#); [Rangelova et al., 2010](#)). These sulfur-based free radical species may then initiate protein or lipid oxidation.

[Baskurt \(1988\)](#) found that exposure of rats to 0.87 ppm SO₂ for 24 hours resulted in increased hematocrit, sulfhemoglobin, and osmotic fragility, as well as decreased whole blood and packed cell viscosities. These results indicate a systemic effect of inhaled SO₂ and are consistent with an oxidative injury to red blood cells. Other studies have reported lipid peroxidation in erythrocytes and tissues of animals exposed to SO₂ ([Qin et al., 2012](#); [Ziemann et al., 2010](#); [Haider et al., 1982](#)). Supplementation with ascorbate and α -tocopherol decreased SO₂-induced lipid peroxidation in erythrocytes ([Etlik et al., 1995](#)). Additionally, recent studies report mitochondrial changes in the hearts and brains of rats exposed to 1.34 ppm (4 hours/day) SO₂ for several weeks ([Qin et al., 2016](#); [Qin et al., 2012](#)). Demonstration of mitochondrial biogenesis in rat brain suggests that SO₂ exposure induces an adaptive response to oxidative stress ([Qin et al., 2012](#)). Changes in cardiac function were observed at higher concentrations (2.7 ppm SO₂); pretreatment with antioxidants blocked this effect ([Qin et al., 2016](#)). Other recent studies report altered

markers of brain inflammation and synaptic plasticity following several weeks to months of exposure to 1.34 ppm (4 hours/day) SO₂ ([Yao et al., 2015](#); [Yao et al., 2014](#)). Further studies are required to confirm that inhalation exposures of SO₂ at or near ambient levels increase blood sulfite levels sufficiently for oxidative injury to occur in blood cells or other tissues.

In summary, exposure to SO₂ may result in effects outside the respiratory tract via activation of sensory nerves in the respiratory tract resulting in a neural reflex response or mediated by circulating sulfite. A few studies employing concentrations of 2 ppm SO₂ or less have demonstrated effects that are consistent with sulfite-mediated redox stress, such as increased sulfhemoglobin in red blood cells and lipid peroxidation in the brain. Recent studies also suggest possible inflammation and other effects in tissues distal to the absorption site following several weeks to months of exposure to 1.34 ppm SO₂.

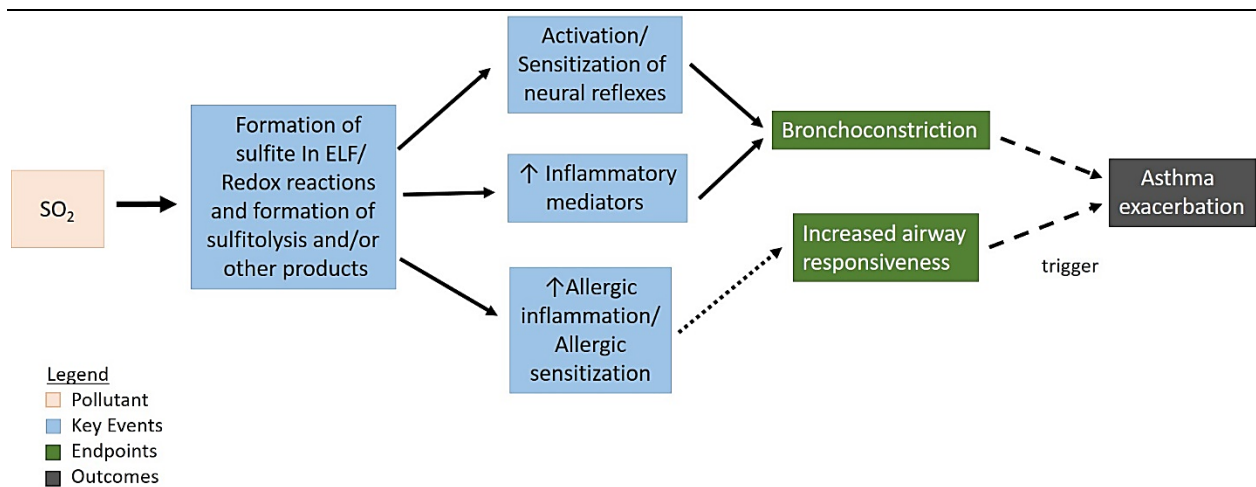
4.3.5 Role of Endogenous Sulfur Dioxide/Sulfite

Endogenous SO₂/sulfite is a product of normal metabolism of sulfur-containing amino acids (e.g., cysteine and methionine) ([Liu et al., 2010](#)). While SO₂ gas is measured in the head space gas of preparations of various tissues or bodily fluids ([Balazy et al., 2003](#)), sulfite/bisulfite is measured in soluble fractions. The distribution of SO₂ and enzymes responsible for SO₂ generation has been reported in tissues of the rat ([Luo et al., 2011](#)). Chemical transformations between bisulfite/sulfite/SO₂ and the gasotransmitter H₂S also occur. H₂S is similarly derived from sulfur-containing amino acids. Evidence has accumulated that endogenous H₂S acts as a biological signaling molecule ([Filipovic et al., 2012](#)) and plays important roles in the cardiovascular ([Coletta et al., 2012](#)) and other systems. Recent studies suggest that endogenous SO₂ may also be a gasotransmitter ([Liu et al., 2010](#)). Like the other gasotransmitters NO and CO, SO₂ at physiologic levels may activate guanylyl cyclase to generate cyclic guanosine monophosphate (cGMP), which mediates effects through cGMP-dependent kinases ([Li et al., 2009](#)). However, SO₂ may also act through non-cGMP-dependent pathways. Experimental studies in animal models and in vitro systems demonstrate myriad effects of exogenous SO₂ on the cardiovascular system, including vasorelaxation, negative inotropic effects on cardiac function, anti-inflammatory and antioxidant effects, and decreased blood pressure (BP) and vascular remodeling in hypertensive animals ([Liu et al., 2010](#)). Effects were concentration dependent in many cases. In vivo studies generally were conducted using 5 ppm and higher concentrations of SO₂ (or sulfite/bisulfite) ([Liu et al., 2010](#)). In summary, endogenous SO₂ is a newly recognized gasotransmitter that may play important roles in cardiovascular and other systems.

4.3.6 Mode of Action Framework

This section describes the key events, endpoints, and outcomes that comprise the modes of action of inhaled SO₂. Here, key events are subclinical effects, endpoints are effects that are generally measured in the clinic, and outcomes are health effects at the organism level. Biological pathways discussed above that may contribute to health effects resulting from short-term and long-term exposures to SO₂ (Chapter 5) are summarized as a part of this analysis. These proposed modes of action are based on the available evidence and may not reflect all of the pathophysiology underlying health effects.

Figure 4-2 depicts the proposed mode of action for respiratory effects due to short-term exposure to SO₂.



ELF = epithelial lining fluid; redox = reduction-oxidation; SO₂ = sulfur dioxide.

Note: Pathways indicated by a dotted line are those for which evidence is limited to findings from experimental animal studies, while evidence from controlled human exposure studies is available for pathways indicated by a solid line. Dashed lines indicate proposed links to the outcomes of asthma exacerbation. Key events are subclinical effects, endpoints are effects that are generally measured in the clinic, and outcomes are health effects at the organism level.

Source: National Center for Environmental Assessment.

Figure 4-2 Summary of evidence for the proposed mode of action linking short-term exposure to sulfur dioxide and respiratory effects.

The propensity for airways to narrow following inhalation of some stimuli is termed airway responsiveness. A characteristic feature of individuals with asthma is an increased propensity of their airways to narrow in response to bronchoconstrictive stimuli relative to nonatopic individuals without asthma. Different kinds of stimuli can elicit

bronchoconstriction, but in general they act on airway smooth muscle receptors (direct stimuli, e.g., methacholine) or act via the release of inflammatory mediators (indirect stimuli, e.g., allergens) (O'Byrne et al., 2009). SO₂ is a nonspecific bronchoconstrictive stimuli that is not easily classified as a direct or indirect stimuli, as was discussed in [Section 4.3.1](#).

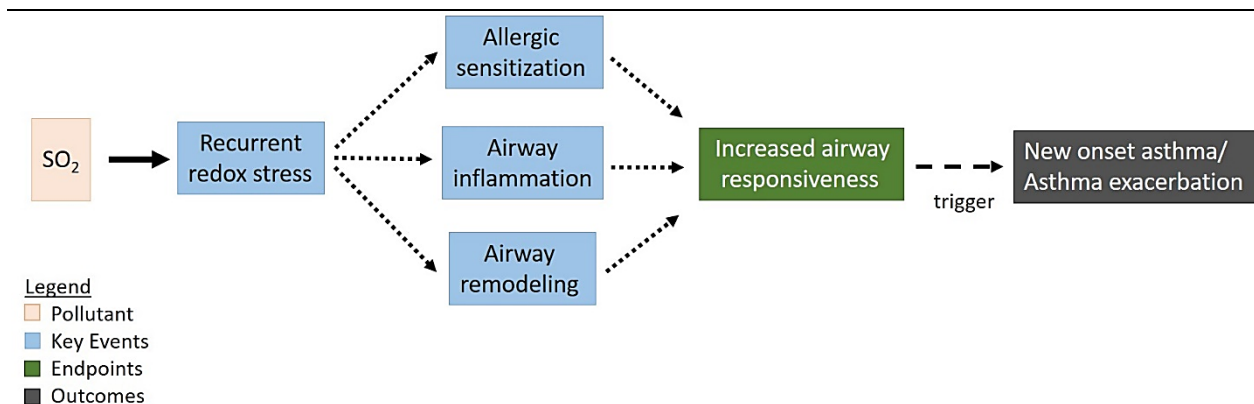
Because inhalation of SO₂ results in chemical reactions in the ELF, the initiating event in the development of respiratory effects is the formation of sulfite, sulfiteolysis products, hydrogen ion, and/or other products. Both sulfite and S-sulfonates have been measured in tracheal and bronchial tissue as well as in tracheal washings of experimental animals exposed to SO₂. Reactive products formed as a result of SO₂ inhalation are responsible for a variety of downstream key events, which may include activation or sensitization of sensory nerves in the respiratory tract resulting in neural reflex responses, release of inflammatory mediators, and modulation of allergic inflammation or sensitization. These key events may collectively lead to several endpoints, including bronchoconstriction and increased airway responsiveness. Bronchoconstriction is characteristic of an asthma attack. However, individuals who are not asthmatic may also experience bronchoconstriction in response to SO₂ inhalation; generally, this occurs at higher concentrations than in an individual who is asthmatic (>1 ppm). Additionally, SO₂ exposure may increase airway responsiveness to subsequent exposures of other stimuli such as allergens or methacholine. These pathways may be linked to the epidemiologic outcome of asthma exacerbation.

The strongest evidence for this mode of action comes from controlled human exposure studies. SO₂ exposure resulted in increased airway resistance due to bronchoconstriction in healthy adults and in adults with asthma. In adults without asthma, this response occurred primarily as a result of activation of sensory nerves in the respiratory tract resulting in neural reflex responses mediated by cholinergic parasympathetic pathways involving the vagus nerve. However, in adults with asthma, evidence indicates that the response is only partially due to vagal pathways and that inflammatory mediators such as histamine and leukotrienes also play an important role. Activation of sensory nerves in the respiratory tract, which result in neural reflex responses, has been studied in humans exposed to occupationally relevant concentrations of SO₂ (up to 2 ppm). Responses measured in these studies include increased respiratory rate and decreased tidal volume, which involve the vagus nerve, and increased nasal air-flow resistance, which involves the trigeminal nerve. These responses are not a part of the mode of action described here, but are mentioned because they are known irritant effects of SO₂. Studies in experimental animals demonstrate that SO₂ exposure activates reflexes that are mediated by cholinergic parasympathetic pathways involving the vagus nerve. However, noncholinergic mechanisms may also play a role because some studies demonstrate that a local axon

reflex resulting in C-fiber secretion of neuropeptides (i.e., neurogenic inflammation) is responsible for the effects of SO₂.

Evidence demonstrates that SO₂ exposure modulates allergic inflammatory responses. Enhancement of allergic inflammation was observed in adults with asthma who were exposed for 10 minutes to 0.75 ppm SO₂ (i.e., leukotriene-mediated increases in numbers of sputum eosinophils). In an animal model of allergic airway disease, repeated exposure to 2 ppm SO₂ led to an enhanced inflammatory response, as measured by numbers of BALF inflammatory cells, levels of BALF cytokines, histopathology, activation of the NFκB pathway, and upregulation of intra-cellular adhesion molecules, mucin, and cytokines, in lung tissue. Furthermore, repeated exposure to SO₂ enhanced Th2 polarization (or Group 2 innate lymphoid cell-mediated Type 2 immunity), numbers of BALF eosinophils, and serum IgE levels in this same model. Other studies demonstrated that repeated exposure of naive animals to SO₂ (as low as 0.1 ppm) over several days promoted allergic sensitization (allergen-specific IgG levels) and enhanced allergen-induced bronchial obstruction (an indicator of increased airway responsiveness) and inflammation (airway fluid eosinophils and histopathology) when animals were subsequently sensitized and challenged with an allergen. Similarly, intra-nasal treatment with sulfite both aggravated allergic sensitization (Th2 cytokines and allergen specific IgE levels) and exacerbated allergic inflammatory responses (histopathology) in animals subsequently sensitized and challenged with allergen. These changes in allergic inflammation may enhance airway responsiveness and promote bronchoconstriction in response to a trigger. Thus, allergic inflammation and increased airway responsiveness may link short-term SO₂ exposure to asthma exacerbation.

[Figure 4-3](#) depicts the proposed mode of action for respiratory effects due to long-term exposure to SO₂.



Redox = reduction-oxidation; SO₂ = sulfur dioxide.

Note: Pathways indicated by a dotted line are those for which evidence is limited to findings from experimental animal studies, while evidence from controlled human exposure studies is available for pathways indicated by a solid line. Dashed lines indicate proposed links to the outcomes of new onset asthma/asthma exacerbation. Key events are subclinical effects, endpoints are effects that are generally measured in the clinic, and outcomes are health effects at the organism level.

Source: National Center for Environmental Assessment.

Figure 4-3 Summary of evidence for the proposed mode of action linking long-term exposure to sulfur dioxide and respiratory effects.

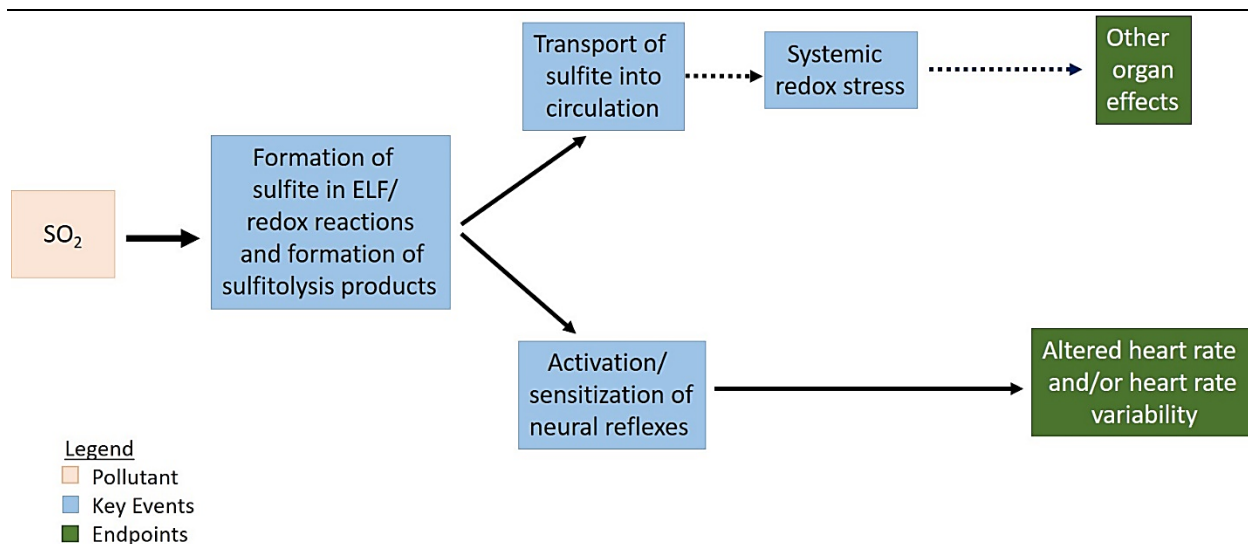
The initiating event in the development of respiratory effects due to long-term SO₂ exposure is the recurrent or prolonged redox stress due to the formation of reactive products in the ELF. This is the driving factor for the potential downstream key events, airway inflammation, allergic sensitization, and airway remodeling that may lead to the endpoint increased airway responsiveness. The term AHR is generally used in cases where airway responsiveness to methacholine or histamine is assessed and the provocative concentration is determined to be sufficiently low to classify the subjects as having AHR based on criteria such as [ATS \(2000a\)](#). Airway inflammation, airway remodeling, and AHR are characteristic of asthma. The resulting outcome may be new asthma onset, which presents as an asthma exacerbation that leads to physician-diagnosed asthma.

Evidence for this mode of action comes from studies in both naive and allergic experimental animals. Exposure of naive newborn animals to SO₂ (2 ppm) for several weeks resulted in hyperemia in lung parenchyma, inflammation in the airways, and Th2 polarization (or Group 2 innate lymphoid cell-mediated Type 2 immunity), the latter of which is a key step involved in allergic sensitization. Support is also provided by short-term studies in naive animals in which repeated exposure to SO₂ (2 ppm) over several days led to pathologic changes, including inflammatory cell influx. Th2 polarization (or other Type 2 immune responses) and airway inflammation may set the

stage for increases in airway responsiveness. In addition, short-term SO₂ exposure (0.1 ppm) promoted allergic sensitization, enhanced other allergic inflammatory responses, and increased airway responsiveness when animals were subsequently sensitized with an allergen. Further, repeated exposure of allergic newborn animals to SO₂ (2 ppm) over several weeks enhanced allergic responses and resulted in morphologic responses indicative of airway remodeling and increased airway responsiveness. Thus, repeated exposure to SO₂ in naive animals may lead to the development of allergic airway disease, which shares many features with asthma. Furthermore, repeated exposure of allergic animals to SO₂ may promote airway remodeling and increased airway responsiveness. Increased airway responsiveness in animal models and the development of AHR in humans may link long-term exposure to SO₂ to the epidemiologic outcome of new onset asthma.

[Figure 4-4](#) depicts the proposed mode of action for extrapulmonary effects due to short-term or long-term exposure to SO₂.

Although experimental studies have shown extrapulmonary effects resulting from SO₂ inhalation (see [Section 4.3.4](#)), there is uncertainty regarding the mode of action underlying these responses. Evidence from controlled human exposure studies (0.2 ppm, 1 hour) points to SO₂ exposure-induced activation/sensitization of neural reflex responses as a key event leading to the endpoint of altered heart rate or heart rate variability. Evidence also points to transport of sulfite into the circulation. Controlled human exposure and experimental animal studies have demonstrated the presence of sulfite and S-sulfonates in plasma, liver, or brain following SO₂ exposure. This occurred at a concentration as low as 0.3 ppm SO₂ in humans exposed for up to 120 hours. Sulfite is highly reactive and may be responsible for redox stress (possibly through auto-oxidation or peroxidase-mediated reactions to produce free radicals) in the circulation and extrapulmonary tissues. However, this is likely to occur only at very high concentrations or during prolonged exposures because circulating sulfite is efficiently metabolized to sulfate in a reaction catalyzed by hepatic sulfite oxidase.



ELF = epithelial lining fluid; redox = reduction-oxidation; SO₂ = sulfur dioxide.

Note: Pathways indicated by a dotted line are those for which evidence is limited to findings from experimental animal studies, while evidence from controlled human exposure studies is available for pathways indicated by a solid line. No links to outcomes are proposed. Key events are subclinical effects and endpoints are effects that are generally measured in the clinic.

Source: National Center for Environmental Assessment.

Figure 4-4 Summary of evidence for the proposed mode of action linking exposure to sulfur dioxide and extrapulmonary effects.

Besides inhalation of SO₂, the ingestion of food additives and the catabolism of sulfur-containing amino acids also contribute to levels of sulfite in the body (Section 4.3.5). In humans, the amount of sulfite derived from inhaled SO₂ (assuming 100% absorption, 75 ppb and 24-hour exposure) is comparable to that derived from the expected daily consumption of food additives. The amount of sulfite derived from the breakdown of endogenous sulfur-containing amino acids is far greater. Sulfite derived from inhaled SO₂, unlike that derived from food additives, enters the circulation without first passing through the liver, which efficiently metabolizes sulfite to sulfate. Thus, the potential exists for inhaled SO₂ to have a greater impact on circulating sulfite levels than sulfite derived from food additives. While the amount of sulfite derived from the breakdown of endogenous sulfur-containing amino acids is far greater, its metabolic pathways and impact on circulating sulfite levels are not clear. Thus, the potential exists for prolonged exposure to high concentrations of inhaled ambient SO₂ to result in extrapulmonary effects due to circulating sulfite.

In summary, this section provides a foundation for understanding how exposure to the gaseous air pollutant SO₂ may lead to health effects. This encompasses the many steps between uptake into the respiratory tract and biological responses that ensue. The

reaction of inhaled SO₂ with components of the ELF initiates a cascade of events occurring at the cellular, organ, and organism level. Biological responses discussed in this section were organized in a mode of action framework that serves as a guide to interpreting health effects evidence presented in [Chapter 5](#).

CHAPTER 5 INTEGRATED HEALTH EFFECTS OF EXPOSURE TO SULFUR OXIDES

5.1 Introduction

5.1.1 Scope of the Chapter

While the term “sulfur oxides” (SO_x) refers to multiple gaseous oxidized sulfur compounds [e.g., sulfur dioxide (SO₂), sulfur trioxide], this chapter focuses on evaluating the health effects associated with exposure to SO₂. As discussed in [Section 2.1](#), SO₂ is the most abundant SO_x species in the atmosphere, and the available health evidence examines SO₂. The health effects of particulate sulfur-containing compounds (e.g., sulfate) are considered in the current review of the National Ambient Air Quality Standards (NAAQS) for particulate matter (PM) and were evaluated in the 2009 Integrated Science Assessment (ISA) for PM ([U.S. EPA, 2009a](#)) (see [Section 1.1](#)).

This chapter evaluates the epidemiologic, controlled human exposure, and animal toxicological evidence of SO₂-related respiratory ([Section 5.2](#)), cardiovascular ([Section 5.3](#)), reproductive and developmental ([Section 5.4](#)), total mortality ([Section 5.5](#)), and cancer ([Section 5.6](#)) effects. Evidence from epidemiologic and animal toxicological studies of other SO₂-related effects are included in Supplemental Tables 5S-1 ([U.S. EPA, 2017c](#)) and 5S-2 ([U.S. EPA, 2017c](#)). Sections for respiratory, cardiovascular, and mortality effects are divided into subsections describing the evidence for short-term (i.e., 1 month or less) and long-term (i.e., more than 1 month) exposures. The evidence for reproductive and developmental and cancer effects is considered within one long-term exposure section, with time windows of exposure addressed as appropriate. Causal conclusions are determined for both short- and long-term exposures by evaluating the evidence for each health effect and exposure category independently, using the causal framework [described in the Preamble to the ISAs ([U.S. EPA, 2015b](#))].

Each chapter section begins with a summary of the conclusions from the 2008 ISA for Sulfur Oxides, followed by an evaluation of recent health studies (i.e., those published since the completion of the 2008 ISA for Sulfur Oxides) that build upon evidence from previous reviews. The collective body of evidence, including recent studies and studies included in previous assessments, is integrated across scientific disciplines to develop conclusions and causality determinations. Within each of the sections focusing on morbidity outcomes (e.g., respiratory morbidity, cardiovascular morbidity), the evidence

is organized into more refined outcome groupings [e.g., asthma exacerbation, myocardial infarction (MI)] that comprise a continuum of subclinical to clinical effects. The discussion of specific health outcomes is then organized by scientific discipline (i.e., epidemiology, controlled human exposure, toxicology). This structure helps in evaluating coherence and biological plausibility of the effects observed in association with exposure to SO₂ and promotes the transparent characterization of the weight of evidence in drawing the causal conclusions found at the end of each section (e.g., see [Section 5.2.1.9](#)). Causal determinations for total mortality are based on the evidence for nonaccidental causes of mortality and informed by the extent to which evidence for the spectrum of cardiovascular and respiratory effects provides biological plausibility for SO₂-related total mortality. Findings for cause-specific mortality inform multiple causal determinations. For example, studies of respiratory and cardiovascular mortality are used to assess the continuum of effects and inform the causal determinations for respiratory and cardiovascular morbidity. As described in [Section 1.2](#), judgments regarding causality are made by evaluating the evidence over the full range of exposures in animal toxicological, controlled human exposure, and epidemiologic studies defined in this ISA to be relevant to ambient air concentrations (i.e., ≤2,000 ppb).

5.1.2 Evidence Evaluation and Integration to Form Causal Determinations

5.1.2.1 Evaluation of Individual Studies

As described in the Preamble to the ISAs ([U.S. EPA, 2015b](#)) (Section 5.a), causal determinations were informed by integrating evidence across scientific disciplines (e.g., exposure, animal toxicology, controlled human exposure studies, epidemiology) and related outcomes, as well as by judgments on the strength of inference from individual studies. These judgments were based on evaluating strengths, as well as various sources of bias and uncertainty related to study design, study population characterization, exposure assessment, outcome assessment, consideration of confounding, statistical methodology, and other factors. This evaluation was applied to controlled human exposure, animal toxicological, and epidemiologic studies included in this ISA, comprising studies from previous assessments as well as those studies published since the 2008 ISA for Sulfur Oxides. The major considerations in evaluating individual studies are described in the [Preamble](#) and are consistent with current best practices employed in other approaches for reporting or evaluating health science data.¹

¹ For example, National Toxicology Program Office of Health Assessment and Translation approach ([Rooney et al., 2014](#)), Integrated Risk Information System Preamble ([U.S. EPA, 2013d](#)), ToxRTool ([Klimisch et al., 1997](#)),

Additionally, these considerations are compatible with published U.S. EPA (Environmental Protection Agency) guidelines related to cancer, neurotoxicity, reproductive toxicity, and developmental toxicity ([U.S. EPA, 2005a](#), [1998](#), [1996a](#), [1991](#)).

The evaluation factors described in the [Preamble](#) were used as a guideline rather than a checklist or criteria to define the quality of a study. The presence or absence of a particular feature did not necessarily define a less informative study or preclude a study from consideration in the ISA. Further, these aspects were not criteria for a particular determination of causality in the five-level hierarchy. As described in the Preamble to the ISAs ([U.S. EPA, 2015b](#)), causal determinations were based on judgments of the overall strengths and limitations of the collective body of available studies and the coherence of evidence across scientific disciplines and related outcomes. Where possible, considerations such as exposure assessment and confounding (i.e., bias due to a relationship with the outcome and correlation with exposures to SO₂), were framed to be specific to sulfur oxides. Thus, judgments of the strength of inference from a study can vary depending on the specific pollutant being assessed.

Evaluation of the extent to which the available scientific information informs the understanding of uncertainties related to the independent effect of sulfur oxides is of particular relevance in the review process. Because examination of copollutant confounding is based largely on copollutant models, the inherent limitations of such models are considered in drawing inferences about independent associations for SO₂. For example, collinearity potentially affects model performance when highly correlated pollutants are modeled simultaneously, and inference can be limited if there are differences in the spatial distributions of SO₂ and the copollutant such that model assumptions of equal measurement error or constant correlations for SO₂ and the copollutant are not satisfied ([Section 3.4.3](#)). Correlations of short-term SO₂ concentrations with other NAAQS pollutants are generally low to moderate, but may vary by location ([Section 3.4.3](#)). Thus, the interpretation of copollutant model results reported in epidemiologic studies depends on a variety of factors, which are discussed throughout the chapter, generally in the context of a specific study and/or health endpoint.

5.1.2.2 Integration of Scientific Evidence

Causal determinations are made by considering the strength of inference from individual studies and on integrating multiple lines of evidence. As detailed in the Preamble to the

STROBE guidelines ([von Elm et al., 2007](#)), Animals in Research: Reporting In Vivo Experiments guidelines ([Kilkenny et al., 2010](#)).

ISAs ([U.S. EPA, 2015b](#)), evidence integration involved evaluating the consistency and coherence of findings within and across disciplines, as well as within and across related outcomes. Cross-disciplinary integration often addresses uncertainties within a particular discipline. Controlled human exposure and animal toxicological studies can provide direct evidence for health effects related to SO₂ exposures. Coherence of experimental evidence with epidemiologic findings can advance our understanding about whether epidemiologic associations with health outcomes plausibly reflect an independent effect of exposure to SO₂ in ambient air. For example, the coherence of effects observed in epidemiologic studies with human clinical studies demonstrating direct effects of SO₂ on lung function ([Section 5.2.1.2](#)), is drawn upon to reduce uncertainties in epidemiologic studies. Thus, evidence across a spectrum of related outcomes and across scientific disciplines (e.g., epidemiologic and controlled human exposure studies) was integrated and used to clarify the understanding of uncertainties for a particular outcome or discipline due to chance, publication bias, selection bias, and confounding by copollutant exposures or other factors.

The integration of the scientific evidence is facilitated by presenting data from multiple studies within and across disciplines. To increase comparability of results across epidemiologic studies, the ISA presents effect estimates for associations with health outcomes scaled to the same increment of SO₂ concentration.¹ The increments for standardization vary by averaging time. For 24-h avg, effect estimates were scaled to a 10-ppb increase for SO₂. For 1-h daily maximum (max), effect estimates were scaled to a 40-ppb increase for SO₂. Effect estimates for long-term exposures to SO₂ (i.e., annual or multiyear averages) were scaled to a 5-ppb increase. These increments were derived by calculating the U.S.-wide percentile distributions for a given averaging time and then calculating the approximate difference between the median (i.e., 50th percentile, a typical pollution day) and the 98th percentile (a more polluted day) for a given averaging time. Units of dose in toxicological studies are typically presented in ppm; however, when toxicological data are summarized in the context of epidemiologic findings, units are converted to ppb for comparability.

5.1.3 Summary

The subsequent sections review and synthesize the evidence of SO₂-related health effects from multiple disciplines (e.g., exposure, animal toxicology, and epidemiology). Information on dosimetry and modes of action ([Chapter 4](#)) provides the foundation for understanding how exposure to inhaled SO₂ may lead to health effects, providing

¹ This is as opposed to reporting effect estimates that are scaled to variable changes in concentration such as interquartile range (IQR) for the study period or an arbitrary unit.

biological plausibility for effects observed in the health studies. The science related to sources, emissions, and atmospheric concentrations ([Chapter 2](#)), as well as the potential for human exposure to ambient air sulfur oxides ([Chapter 3](#)), also informs the interpretation of the health effects evidence. Integrative “Summary and Causal Determination” sections for short- and long-term exposures follow the discussion of the evidence for each health outcome category. These integrative summary sections include assessments of the strength of inference from studies comprising the evidence base and integrate multiple lines of evidence to characterize relationships between sulfur oxides and various health effects.

5.2 Respiratory Effects

5.2.1 Short-Term Exposure

5.2.1.1 Introduction

The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) concluded that there is a causal relationship between respiratory effects and short-term exposure to SO₂. The rationale for this causal determination was heavily based on evidence from multiple, high-quality controlled human exposure studies demonstrating decreased lung function and increased respiratory symptoms following SO₂ exposures of 5–10 minutes in exercising adults with asthma.

Epidemiologic evidence also indicated associations between short-term increases in ambient SO₂ concentration and respiratory effects in populations living in locations with ambient concentrations below the previous 24-h avg NAAQS level of 140 ppb. Evidence was strongest for increased respiratory symptoms and respiratory-related hospital admissions and emergency department (ED) visits, especially in children. Due to inadequate examination, a key uncertainty was potential confounding by copollutants, particularly PM. However, controlled human exposure studies of individuals with asthma clearly show that respiratory effects are caused by 5–10 minute SO₂ exposures.

In contrast with asthma exacerbation, there was little information to assess whether short-term SO₂ exposure exacerbated allergy or chronic obstructive pulmonary disease (COPD) or increased risk of respiratory infection. There was some experimental evidence, however, for respiratory effects in healthy humans (>1,000 ppb) and animal models (100 ppb) exposed to SO₂. Epidemiologic evidence in healthy populations was limited and inconsistent.

As described in the following sections, evidence from recent studies is generally consistent with that in the 2008 ISA and 1982 Air Quality Criteria Document (AQCD) for Sulfur Oxides ([U.S. EPA, 2008d, 1982a](#)). To clearly characterize differences in the weight of evidence and the extent of coherence among disciplines and related outcomes, the sections are organized by respiratory outcome group [asthma exacerbation ([Section 5.2.1.2](#)), allergy exacerbation ([Section 5.2.1.3](#)), COPD exacerbation ([Section 5.2.1.4](#)), respiratory infection ([Section 5.2.1.5](#)), aggregated respiratory conditions ([Section 5.2.1.6](#)), respiratory effects in the general population and healthy individuals ([Section 5.2.1.7](#)), and respiratory mortality ([Section 5.2.1.8](#))]. Epidemiologic studies comprise most of the recent evidence base, and previous controlled human exposure and animal toxicological studies form the basis for characterizing and integrating evidence across disciplines. Recent epidemiologic evidence supports associations between ambient SO₂ concentrations and asthma-related symptoms, hospital admissions, and ED visits, but uncertainties related to exposure measurement error and copollutant confounding remain. Recent epidemiologic studies add information on allergy and COPD exacerbation, respiratory infection, and respiratory effects in healthy populations, but relationships of these outcomes with short-term SO₂ exposure remain unclear because of inconsistent evidence or limited coherence among disciplines.

5.2.1.2 Asthma Exacerbation

Asthma is a chronic inflammatory lung disease with a broad range of characteristics and disease severity. SO₂ exposure has been demonstrated to induce clinical features of asthma exacerbation, including decreased lung function [e.g., decreased forced expiratory volume in 1 sec (FEV₁) or increased specific airway resistance (sRaw)], and increased symptoms (e.g., wheezing, cough, shortness of breath), as well as some subclinical effects such as inflammation. This section describes evidence for SO₂-associated lung function changes and respiratory symptoms in people with asthma, hospital admissions and emergency department visits for asthma and related respiratory conditions, and subclinical effects underlying asthma such as pulmonary inflammation and oxidative stress.

As detailed in the previous 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), controlled human exposure studies reported increased respiratory symptoms and decreased lung function after short-term exposures of 5–10 minutes to 0.2–0.6 ppm SO₂ during exercise or eucapnic hyperpnea (a rapid and deep breathing technique through a mouthpiece that prevents an imbalance of carbon dioxide due to hyperventilation) in adults and adolescents (12–18 years old) with asthma. In contrast, healthy adults demonstrated increased airway resistance and decreased FEV₁ following exposure to higher

concentrations (>1.0–5.0 ppm) in the majority of controlled human exposure studies evaluating the respiratory effects of SO₂ ([Section 5.2.1.7](#)). While children may be especially susceptible to the respiratory effects of SO₂ for dosimetric reasons ([Section 4.2.2](#)), there are no available controlled human exposure studies in children under 12, partly due to ethical concerns.

Coherent with controlled human exposure findings, epidemiologic evidence indicated that short-term increases in ambient SO₂ concentration were associated with asthma-related hospital admissions, ED visits, and symptoms. The strongest evidence was for children, which is consistent with their greater oral breathing and higher ventilation rates relative to their size than adults and the consequent potential for them receiving a higher SO₂ dose to the tracheobronchial airways of the lower respiratory tract ([Sections 4.1.2, 4.2.2](#)). Epidemiologic evidence for SO₂-related lung function decrements was inconsistent among both children and adults with asthma. A key uncertainty in the epidemiologic evidence was whether the findings reflected an independent association for SO₂ because the studies assigned exposure from a limited set of fixed-site monitors that may not adequately reflect the spatial and temporal heterogeneity of SO₂ concentrations ([Section 3.3.1.1](#)). Also, few of the studies examined potential confounding by particulate matter with an aerodynamic diameter less than or equal to 2.5 μm (PM_{2.5}) or other copollutants.

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) also provided limited evidence for a relationship between SO₂ concentrations and allergic responses and inflammation in individuals with asthma. Children and adults with atopy plus asthma were found to be at greater risk of SO₂-associated respiratory effects such as respiratory symptoms and lung function decrements. In addition, animal toxicological studies demonstrated that repeated exposure to SO₂ enhanced inflammation and allergic responses in animal models of allergic airway disease.

Both recent studies and the evidence presented in the 2008 SO_x ISA link short-term SO₂ exposure to asthma exacerbation. Most recent studies are epidemiologic, and they continue to show ambient SO₂-associated increases in asthma symptoms, hospital admissions, and ED visits among children. However, uncertainty regarding exposure measurement error and copollutant confounding remains in the epidemiologic evidence. A few recent animal toxicological studies add support for SO₂-induced allergic inflammation. While there are no recent controlled human exposure studies in individuals with asthma (see [Section 5.2.1.7](#) for recent studies in healthy individuals), previous evidence from controlled human exposure studies provides support for an independent effect of SO₂ exposure on asthma exacerbation.

Lung Function Changes in Populations with Asthma

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) reported strong evidence for the effects of SO₂ exposure on decrements in lung function in controlled human exposure studies in adults with asthma under increased ventilation conditions. Controlled human exposure studies, none of which are new since the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), also demonstrated a subset of individuals within this population who are particularly sensitive to the effects of SO₂ exposure. This finding is most evident in the recent analysis of several published studies by [Johns et al. \(2010\)](#). Some additional data from the previous studies has also become available since the 2008 SO_x ISA and is summarized in [Tables 5-2, 5-3, and 5-4](#). Recent epidemiologic findings are inconsistent overall. A few recent epidemiologic studies add evidence for SO₂ measured at a children's school or in copollutant models with PM, nitrogen dioxide (NO₂), or ozone (O₃), although their reliance on fixed site monitors that may not capture the spatial and temporal variation of SO₂ represents a limitation. There is a paucity of evidence from animal toxicological studies. While some animal toxicological studies of short-term exposure to SO₂ have examined changes in lung function, these experiments were conducted in naive animals rather than in models of allergic airway disease, which share many phenotypic features with asthma in humans.

Controlled Human Exposure Studies

Bronchoconstriction in individuals with asthma is the most sensitive indicator of SO₂-induced lung function effects. A characteristic feature of individuals with asthma is an increased propensity of their airways to narrow in response to bronchoconstrictive stimuli relative to nonatopic individuals without asthma. Different kinds of stimuli can elicit bronchoconstriction, but in general, stimuli act on airway smooth muscle receptors (direct stimuli, e.g., methacholine) or act via the release of inflammatory mediators (indirect stimuli, e.g., allergens) ([O'Byrne et al., 2009](#)). SO₂ is a nonspecific bronchoconstrictive stimulus that is not easily classified as direct or indirect, as discussed in [Section 4.3.1](#).

The propensity for airways to narrow following inhalation of some stimuli is termed bronchial or airway responsiveness. The term airway hyperresponsiveness (AHR) is generally used in cases where airway responsiveness to methacholine or histamine is assessed and the provocative concentration (PC) is determined to be sufficiently low to classify the subjects as having AHR based on criteria such as [ATS \(2000a\)](#). Along with symptoms, variable airway obstruction, and airway inflammation, AHR is a primary feature in the clinical definition and characterization of asthma severity ([Reddel et al., 2009](#)).

Bronchoconstriction, evidenced by decrements in lung function, is observed in controlled human exposure studies after approximately 5–10-minute exposures and can occur at SO₂ concentrations as low as 0.2 ppm in exercising individuals with asthma; more consistent decrements are seen at concentrations of 0.4 ppm and greater ([U.S. EPA, 2008d](#)). In contrast, healthy adults are relatively insensitive to the respiratory effects of SO₂ below 1 ppm ([Section 5.2.1.7](#)). In all individuals, bronchoconstriction is mainly seen during conditions of increased ventilation rates, such as exercise or eucapnic hyperpnea, and is likely due to a shift from nasal breathing to oronasal breathing, which increases the concentration of SO₂ reaching the bronchial airways ([Section 4.2.2](#)). The majority of controlled human exposures to SO₂ were conducted with adult volunteers, although a limited number were also conducted with adolescents (12–18 years). Characteristics of controlled exposure studies in individuals with asthma are summarized in [Table 5-1](#). Controlled exposure studies in individuals without asthma are discussed in [Section 5.2.1.7](#).

Table 5-1 Study-specific details from controlled human exposure studies of individuals with asthma.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Endpoints Examined
Balmes et al. (1987)	Asthma; n = 8; 6 M, 2 F (23–39 yr)	0, 0.5, or 1 ppm SO ₂ for 1, 3, or 5 min during eucapnic hyperpnea (60 L/min) via mouthpiece	sRaw, symptoms
Bethel et al. (1983)	Asthma; n = 10; 8 M, 2 F (22–36 yr)	0 or 0.5 ppm SO ₂ for 5 min with exercise 750 kg m/min (125 watts)	sRaw
Bethel et al. (1984)	Asthma; n = 7; 5 M, 2 F (24–36 yr)	0 or 0.5 ppm SO ₂ for 3 min with humidified room-temperature or cold dry air via mouthpiece	sRaw
Bethel et al. (1985)	Asthma; n = 19; 16 M, 3 F (22–46 yr)	0 or 0.25 ppm SO ₂ for 5 min during heavy exercise [bicycle, 750 (n = 19) or 1,000 (n = 9) kg m/min; 125 or 167 watts, respectively]	sRaw
Gong et al. (1995)	Asthma; n = 14; 12 M, 2 F (18–50 yr)	0 or 0.5, 1.0 ppm SO ₂ with light, medium, and heavy exercise (avg ventilation 30, 36, and 43 L/min, respectively) for 10 min	sRaw, FEV ₁ , symptoms, psychophysical (stamina) changes
Gong et al. (1996)	Asthma; n = 10; 2 M, 8 F (19–49 yr)	0 or 0.75 ppm SO ₂ for 10 min with exercise (29 L/min) at 1, 12, 18, and 24 h after pretreatment with placebo or salmeterol (long-acting β ₂ -agonist)	FEV ₁ , symptoms

Table 5-1 (Continued): Study specific details from controlled human exposure studies of individuals with asthma.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Endpoints Examined
Gong et al. (2001)	Asthma; n = 12; 2 M, 10 F (20–48 yr)	0 or 0.75 ppm SO ₂ for 10 min with exercise (35 L/min) with or without pretreatment to montelukast sodium (10 mg/day for 3 days)	sRaw, FEV ₁ , symptoms, eosinophil counts in induced sputum
Horstman et al. (1986)	(1) Asthma; n = 27; 27 M with asthma and sensitive to inhaled methacholine (19–33 yr) (2) n = 4 from study population above	(1) 0, 0.25, 0.5, or 1.00 ppm SO ₂ for 10 min with exercise (treadmill, 21 L/min per m ² body surface area) (2) 2 ppm SO ₂ for 10 min with exercise (treadmill, 21 L/min per m ² body surface area)	sRaw
Horstman et al. (1988)	Asthma; n = 12; 12 M (22–37 yr)	0 or 1.0 ppm SO ₂ for 0, 0.5, 1.0, 2.0, and 5.0 min with exercise (treadmill 40 L/min)	sRaw, symptoms
Jörres and Magnussen (1990)	Asthma; n = 14; 10 M, 4 F (21–55 yr, 34 ± 14 yr)	0 or 0.25 ppm NO ₂ , or 0.5 ppm SO ₂ at rest for 30 min via mouthpiece followed by challenge with 0.75 ppm SO ₂ during voluntary eucapnic hyperpnea via mouthpiece. Ventilation increased in 15 L/min steps, each lasting 3 min	sRaw
Kehrl et al. (1987)	Asthma; n = 10; 10 M (20–30 yr)	0 or 1 ppm SO ₂ for 1 h with exercise (3 × 10 min at 41 L/min on a treadmill)	sRaw
Koenig et al. (1980)	Asthma; n = 9; 7 M, 2 F (14–18 yr)	0 or 1 ppm SO ₂ with 1 mg/m ³ of NaCl droplet aerosol and filtered air (no SO ₂ or NaCl) exposures for 60 min via facemask with mouth breathing at rest, no exposure to SO ₂ alone	FEV ₁ , RT, FRC, V _{max50} , V _{max75} , symptoms
Koenig et al. (1981)	Asthma; n = 8; 6 M, 2 F (14–18 yr)	0 or 1 ppm SO ₂ with 1 mg/m ³ of NaCl droplet aerosol and filtered air (no SO ₂ or NaCl) exposures for 30 min via mouthpiece at rest followed by 10 min exercise on a treadmill (sixfold increase in minute ventilation), no exposure to SO ₂ alone	FEV ₁ , RT, FRC, V _{max50} , V _{max75} , symptoms
Koenig et al. (1983)	(1) Asthma with EIB; n = 9; 6 M, 3 F (12–16 yr) (2) Asthma with EIB; n = 7 from study population above	(1) 1 g/m ³ of NaCl droplet aerosol, 1 ppm SO ₂ + 1 mg/m ³ NaCl, 0.5 ppm SO ₂ + 1 mg/m ³ NaCl for 30 min exposure via mouthpiece at rest followed by 10 min exercise on treadmill (five- to sixfold increase in V _E), no exposure to SO ₂ alone or filtered air (2) 0.5 ppm SO ₂ + 1 mg/m ³ NaCl via a face mask without nose clip with exercise conditions the same as above, no exposure to SO ₂ alone or filtered air	FEV ₁ , RT, FRC, V _{max50} , V _{max75} , symptoms
Koenig et al. (1987)	Allergic with EIB; n = 10; 3 M 7 F (13–17 yr)	0 or 0.75 ppm SO ₂ (mouthpiece) with exercise (33.7 L/min) for 10 and 20 min prior pretreatment (placebo or 180 µg albuterol)	FEV ₁ , RT, FRC, symptoms

Table 5-1 (Continued): Study specific details from controlled human exposure studies of individuals with asthma.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Endpoints Examined
Koenig et al. (1988)	Asthma with EIB; n = 8; 2 M, 6 F (13–17 yr)	1.0 ppm SO ₂ 10 min (mouthpiece, treadmill, 35 L/min) with pretreatment (placebo 20, 40, 60 mg cromolyn) 20 min prior, no air control exposure	FEV ₁ , RT
Koenig et al. (1990)	Asthma with EIB; n = 13; 8 M, 5 F (12–18 yr)	0.1 ppm SO ₂ for 15 min preceded by air or 0.12 ppm O ₃ for 45 min during intermittent exercise (2 × 15 min at 30 L/min on a treadmill), all exposures were via mouthpiece, no air control exposure	FEV ₁ , RT, FRC, V _{max50} , symptoms
Koenig et al. (1992)	Asthma; n = 8; 2 M, 6 F (18–46 yr; 27.5 ± 9.6 yr)	0 or 1 ppm SO ₂ for 10 min via mouthpiece with exercise (13.4–31.3 L/min) with or without pretreatment to theophylline	FEV ₁ , RT
Lazarus et al. (1997)	Asthma; n = 12; 7 M, 5 F (24–43 yr)	0, 0.25, 0.5, 1.0, 2.0, 4.0, or 8.0 ppm SO ₂ with eucapnic hyperpnea (20 L/min) via mouthpiece for 4 min sequential exposures with pretreatment with Zafirlukast (placebo or 20 mg) 2 or 10 h earlier	sRaw
Linn et al. (1983b)	Asthma; n = 23; 13 M, 10 F (19–31 yr)	(1) 0, 0.2, 0.4, or 0.6 ppm SO ₂ with low humidity or high humidity for 10 min with exercise (bicycle, 5 min 50 L/min) (2) 0 or 0.6 ppm SO ₂ with warm air or cold air with exercise (bicycle, 50 L/min, ~5 min)	sRaw, sGaw, FVC, FEV ₁ , symptoms
Linn et al. (1983a)	Asthma; n = 23; 15 M, 8 F (18–30 yr, 23 ± 4 yr)	0 or 0.75 ppm SO ₂ with unencumbered breathing and mouth only breathing (with exercise 40 L/m, 10 min bicycle)	sRaw, thoracic gas volume, symptoms, FVC, FEV ₁ , PEFR, V _{max50} , V _{max25}
Linn et al. (1984c)	Asthma; n = 24; 13 M, 11 F (19–31 yr)	0, 0.3, or 0.6 ppm SO ₂ at 21°C, 7°C, and –6°C, RH 80% (bicycle 50 L/min, ~5 min)	sRaw, sGaw, symptoms
Linn et al. (1984a)	Asthma; n = 14; 12 M, 2 F (18–33 yr)	0 or 0.6 ppm SO ₂ for 6 h with exercise on Day 1 and 2 (2 × 5-min exercise, bicycle, 50 L/min per exposure)	sRaw, sGaw, symptoms
Linn et al. (1984b)	(1) Asthma; n = 8; 4 M, 4 F (19–29 yr) (2) Asthma; n = 24; 17 M 7 F (18–30 yr)	(1) 0, 0.2, 0.4, or 0.6 ppm SO ₂ at 5°C, 50 and 85%RH with exercise (5 min, 50 L/min) (2) 0 or 0.6 ppm SO ₂ at 5°C and 22°C, 85% RH with exercise (5 min, 50 L/min)	sRaw, sGaw, FEV ₁ , symptoms
Linn et al. (1985b)	Asthma; n = 22; 13 M, 9 F (18–33 yr)	0 or 0.6 ppm SO ₂ at 21°C and 38°C and 20 and 80% RH with exercise (~5 min, 50 L/min)	sRaw, sGaw, symptoms
Linn et al. (1985a)	COPD; n = 24; 15 M, 9 F (49–68 yr)	0, 0.4, or 0.8 ppm SO ₂ for 1 h with exercise (2 × 15 min, bicycle, 18 L/min)	sRaw, FVC, FEV ₁ , MMFR, symptoms

Table 5-1 (Continued): Study specific details from controlled human exposure studies of individuals with asthma.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Endpoints Examined
Linn et al. (1987)	Healthy; n = 24; 15 M, 9 F (18–37 yr) Atopic; n = 21; 12 M, 9 F (18–32 yr) Minimal or mild asthma; n = 16; 10 M, 6 F (20–33 yr) Moderate or severe asthma; n = 24; 10 M, 14 F (18–35 yr) Moderate or severe asthma; n = 24	0, 0.2, 0.4, or 0.6 ppm SO ₂ 1 h exposures 3 × 10-min exercise (bicycle) periods ~40 L/min Two rounds of exposures were conducted	Lung function measure pre-exposure, ~15 min and ~55 min into exposure sRaw, FVC, FEV ₁ , peak expiratory flow rate, maximal midexpiratory flow rate Continuously—EKG Midway—HR Before, during, 1-day after, and 1-wk after-symptom score, self-rated activity Immediately after exposure—bronchial reactivity percentage change in FEV ₁ induced by 3 min normocapnic hyperpnea with cold, dry air
Linn et al. (1988)	Asthma; n = 20; 13 M, 7 F (19–36 yr)	Three pretreatment groups (1) metaproterenol sulfate (2) placebo (3) no treatment 0, 0.3, or 0.6 ppm SO ₂ 10 min with exercise (bike 50 L/min)	Lung function—pre, post 60 min, 90 min 120 min, Symptoms—pre, post, 20 min post, 60 min post, 120 min post, 24 h post, 1 wk post
Linn et al. (1990)	Asthma; n = 21; 6 M, 15 F (19–48 yr)	0, 0.3, or 0.6 ppm SO ₂ 10 min with exercise 50 L/min (1) low medication use; (2) normal; (3) high (usual medication supplemented by inhaled metaproterenol before exposure)	Lung function and symptoms measured before and after exposure
Magnussen et al. (1990)	Asthma; n = 46; 24 M, 22 F (28 ± 14 yr) Healthy; n = 12 (24 ± 5 yr)	0 or 0.5 ppm SO ₂ 10 min tidal breathing followed by 10 min of isocapnic hyperventilation (30 L/min) via mouthpiece Histamine challenge—(8 mg/mL)	sRaw
Myers et al. (1986a)	Asthma; n = 10; 7 M, 3 F (19–40 yr)	0, 0.25, 0.5, 1, 2, 4, or 8 ppm SO ₂ 3 min sequential exposures (mouthpiece, 40 L/min) with pretreatment 30 min prior with cromolyn (placebo, 20, or 200 mg)	sRaw
Myers et al. (1986b)	(1) Asthma; n = 9; 7 M, 2 F (19–40 yr) (2) Asthma; n = 7; 7 M (19–40 yr)	0, 0.25, 0.5, 1, 2, 4, or 8 ppm SO ₂ 3 min sequential exposures (mouthpiece, eucapnic hyperpnea 40 L/min) with pretreatment 30 min prior (1) atropine (2 mg) and cromolyn (200 mg); (2) placebo and cromolyn (200 mg); (3) atropine (2 mg) and placebo; (4) placebo	sRaw

Table 5-1 (Continued): Study specific details from controlled human exposure studies of individuals with asthma.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Endpoints Examined
Roger et al. (1985)	Asthma; n = 28; 28 M (19–33 yr)	0, 0.25, 0.5, or 1.0 ppm SO ₂ with three 10 min periods of exercise (42.4 L/min) each followed by 15 min of exposure at rest for a total exposure duration of 75 min	Raw; sRaw; FVC, FEV ₁ , FEF _{25–75} , FEF _{max} , FEF ₅₀ , FEF ₇₅ ,
Rubinstein et al. (1990)	Asthma; n = 9; 5 M, 4 F (23–34 yr)	0 or 0.3 ppm NO ₂ during exercise in a chamber followed by challenge with 0.25 to 4.0 ppm SO ₂ , in doubling dose increments, for 4 min each via mouthpiece during light exercise (20 L/min) until sRaw increased by 8 SRaw units above baseline	sRaw, FVC, FEV ₁ , single-breath nitrogen test
Sheppard et al. (1983)	Asthma; n = 8; 4 M, 4 F (22–36 yr)	0.5 ppm SO ₂ for 3 min eucapnic hyperpnea via mouthpiece	sRaw, symptoms
Trenga et al. (1999)	Asthma; n = 47; 14 M, 33 F (18–39 yr)	0.5 ppm SO ₂ for 10 min via mouthpiece during moderate exercise	FEV ₁ , FVC, FEV ₁ /FVC, PEF, FEF _{25–75} , symptoms ratings
Trenga et al. (2001)	Asthma; n = 17; 5 M, 12 F (19–38 yr)	0.1 or 0.25 ppm SO ₂ for 10 min via mouthpiece with moderate exercise (treadmill) following exposure to air or 0.12 O ₃ ppm for 45 min via mouthpiece with intermittent moderate exercise.	FVC, FEV ₁ , FEF _{25–75} , PEF, symptoms
Tunnicliffe et al. (2003)	Asthma; n = 12 (adults, 35.7 yr) Healthy; n = 12 (adults, 34.5 yr)	0 or 0.2 ppm SO ₂ via head dome at rest	Symptoms, FEV ₁ , FVC, MMEF, exhaled NO, ascorbic and uric acid in nasal lavage fluid

COPD = chronic obstructive pulmonary disease; EIB = exercise-induced bronchospasm; EKG = electrocardiogram; F = female; FEV₁ = forced expiratory volume in 1 sec; FEF_{25–75%} = forced expiratory flow at 25–75% of forced vital capacity; FEF_{50%} = forced expiratory flow at 50% of forced vital capacity; FEF_{75%} = forced expiratory flow at 75% of forced vital capacity; FEF_{max} = maximum forced expiratory flow; FRC = functional residual capacity; FVC = forced vital capacity; HR = heart rate; M = male; MMEF = maximum midexpiratory flow; MMFR = maximal midexpiratory flow rate; n = sample size; NaCl = sodium chloride; NO = nitric oxide; NO₂ = nitrogen dioxide; O₃ = ozone; PEF = peak expiratory flow; PEFR = peak expiratory flow rates; ppm = parts per million; Raw = airway resistance; RH = relative humidity; RT = total respiratory resistance; SD = standard deviation; sGAW = specific airway conductance; sRaw = specific airway resistance; SO₂ = sulfur dioxide; V_E = minute volume; V_{max} = maximal flow of expired vital capacity; V_{max75} = flow rate with 75% of FVC remaining to be expired; V_{max50} = flow rate with 50% of FVC remaining to be expired; V_{max25} = flow rate with 25% of FVC remaining to be expired.

^aRange or Mean ± SD.

Several investigators ([Linn et al., 1990](#); [Linn et al., 1988](#); [Linn et al., 1987](#); [Bethel et al., 1985](#); [Linn et al., 1984a](#); [Linn et al., 1983b](#)) demonstrated ≥100% increase in sRaw or ≥15% decrease in FEV₁ after 5–10-minute exposures to low concentrations (0.2–0.3 ppm) of SO₂ in exercising adults with asthma, with effects being more pronounced following 5–10-minute exposures to 0.4–0.6 ppm SO₂ ([Linn et al., 1990](#); [Magnussen et al., 1990](#); [Linn et al., 1988](#); [Linn et al., 1987](#); [Roger et al., 1985](#); [Linn et al., 1983b](#)).

SO₂-induced bronchoconstriction in individuals with asthma occurs rapidly when exposed while at increased ventilation and is transient with recovery following cessation of such conditions. Bronchoconstriction occurs in as little as 2 minutes from the start of exposure in adults with asthma who have increased ventilation rates due to exercise or eucapnic hyperpnea ([Horstman et al., 1988](#); [Balmes et al., 1987](#); [Sheppard et al., 1983](#)). During exposure to SO₂ over a 30-minute period with continuous exercise, the response to SO₂ develops rapidly and is maintained throughout the 30-minute exposure ([Kehrl et al., 1987](#)). [Linn et al. \(1984a\)](#) reported decrements in lung function in adults with asthma immediately after each of two 5-min exercise periods (one after entering the chamber and the second 5 hours later) in two 6-hour exposures to 0.6 ppm SO₂ on successive days. The decrements in lung function observed in the early and late exercise periods were not statistically significantly different from each other, but tended to be diminished in the late exercise period relative to the first. The responses observed after the second day of SO₂ exposure were also slightly (minimal biologically, but statistically less based on sGaw data) less than the response observed after the first day of SO₂ exposure. These effects are generally observed to return to baseline levels within 1 hour after cessation of exercise, even with continued exposure ([Linn et al., 1984a](#)).

Other factors that affect responses to SO₂ include temperature and humidity. The majority of controlled human exposure studies were conducted at 20–25°C and at relative humidities ranging from ~25–90%. Some evidence indicates that the respiratory effects of SO₂ are exacerbated by colder and dryer conditions ([Linn et al., 1985b](#); [Bethel et al., 1984](#); [Linn et al., 1984b](#)).

Responders versus nonresponders to SO₂. At the time of the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), it was well documented that some individuals have a greater response to SO₂ than others with similar disease status ([Table 5-2](#)) ([Linn et al., 1990](#); [Magnussen et al., 1990](#); [Linn et al., 1988](#); [Linn et al., 1987](#); [Horstman et al., 1986](#); [Bethel et al., 1985](#); [Roger et al., 1985](#); [Linn et al., 1984b](#); [Linn et al., 1983b](#)).

[Horstman et al. \(1986\)](#) reported that individuals required different concentrations of SO₂ to produce a doubling of sRaw ($\geq 100\%$) compared to clean air exposure [provocative concentration of SO₂, PC(SO₂)] ([Figure 5-1](#)). This study described the distribution of individual bronchial sensitivity to SO₂, measured by sRaw, in 27 subjects with asthma that were sensitive to methacholine; insensitive volunteers were excluded from further participation in the study. Individuals were exposed to concentrations of SO₂ between 0 and 2 ppm for 10 minutes under exercising conditions [minute ventilation (V_E) = 42 L/minute]. While six of the subjects (22%) reached a PC(SO₂) below 0.5 ppm SO₂, two subjects (7.4%) experienced a moderate decrease ≤ 0.3 ppm ([Figure 5-1](#)). On the other end of the spectrum, four subjects (14.8%) did not demonstrate $\geq 100\%$ increase in

sRaw even when exposed to 2.0 ppm SO₂ and eight (29.6%) subjects required an SO₂ concentration between 1.0 and 2.0 ppm to elicit a response. The authors noted that the effects of SO₂ on sRaw are similar to a variety of nonspecific bronchoconstrictive stimuli. However, they observed only a weak correlation between airway responsiveness to SO₂ and methacholine ($r = 0.31, p = 0.12$). This study demonstrates substantial interindividual variability in sensitivity to the bronchoconstrictive effects of SO₂ in exercising adults with asthma.

Table 5-2 Percentage of adults with asthma in controlled human exposure studies experiencing sulfur dioxide-induced decrements in lung function and respiratory symptoms.

SO ₂ Conc (ppm)	Exposure Duration (min)	n	Ventil-ation (L/min)	Cumulative Percentage of Subjects (Number of Subjects) ^a				Study	Respiratory Symptoms: Supporting Studies
				↑sRaw	≥100%	≥200%	≥300%		
				↓ FEV ₁	≥15%	≥20%	≥30%		
0.2	5	23	~48	sRaw	9% (2) ^b	0	0	Linn et al. (1983b)	Limited evidence of SO ₂ -induced increases in respiratory symptoms in some people with asthma: (Linn et al. (1990) ; Linn et al. (1988) ; Linn et al. (1987) ; Schachter et al. (1984) ; Linn et al. (1983b))
	10	40	~40	sRaw	7.5% (3) ^c	2.5% (1) ^c	0 ^c	Linn et al. (1987)^c	
	10	40	~40	FEV ₁	9% (3.5) ^c	2.5% (1) ^c	1% (0.5) ^c	Linn et al. (1987)^c	
0.25	5	19	~50–60	sRaw	32% (6)	16% (3)	0	Bethel et al. (1985)	
	5	9	~80–90	sRaw	22% (2)	0	0	Bethel et al. (1985)	
	10	28	~40	sRaw	4% (1)	0	0	Roger et al. (1985)	
0.3	10	20	~50	sRaw	10% (2)	5% (1)	5% (1)	Linn et al. (1988)^d	
	10	21	~50	sRaw	33% (7)	10% (2)	0	Linn et al. (1990)^d	
	10	20	~50	FEV ₁	15% (3)	0	0	Linn et al. (1988)	
	10	21	~50	FEV ₁	24% (5)	14% (3)	10% (2)	Linn et al. (1990)	

Table 5-2 (Continued): Percentage of adults with asthma in controlled human exposure studies experiencing sulfur dioxide induced decrements in lung function and respiratory symptoms.

SO ₂ Conc (ppm)	Exposure Duration (min)	n	Ventil-ation (L/min)	Cumulative Percentage of Subjects (Number of Subjects) ^a				Study	Respiratory Symptoms: Supporting Studies
				↑sRaw	≥100%	≥200%	≥300%		
				↓ FEV ₁	≥15%	≥20%	≥30%		
0.4	5	23	~48	sRaw	13% (3)	4% (1)	0	Linn et al. (1983b)	Stronger evidence with some statistically significant increases in respiratory symptoms:
	10	40	~40	sRaw	24% (9.5) ^c	9% (3.5) ^c	4% (1.5) ^c	Linn et al. (1987)^c	
	10	40	~40	FEV ₁	27.5% (11) ^c	17.5% (7) ^c	10% (4) ^c	Linn et al. (1987)^c	
0.5	5	10	~50-60	sRaw	60% (6)	40% (4)	20% (2)	Bethel et al. (1983)	Balmes et al. (1987)^f ; Gong et al. (1995) (Linn et al. (1987) ; Linn et al. (1983b)) Roger et al. (1985)
	10	28	~40	sRaw	18% (5)	4% (1)	4% (1)	Roger et al. (1985)	
	10	45	~30	sRaw	36% (16)	16% (7)	13% (6)	Magnussen et al. (1990)^f	
0.6	5	23	~48	sRaw	39% (9)	26% (6)	17% (4)	Linn et al. (1983b)	
	10	40	~40	sRaw	34% (13.5) ^c	24% (9.5) ^c	19% (7.5) ^c	Linn et al. (1987)^c	
	10	20	~50	sRaw	60% (12)	35% (7)	10% (2)	Linn et al. (1988)	
	10	21	~50	sRaw	62% (13)	29% (6)	14% (3)	Linn et al. (1990)	
	10	40	~40	FEV ₁	47.5% (19) ^c	39% (15.5) ^c	17.5% (7) ^c	Linn et al. (1987)^c	
	10	20	~50	FEV ₁	55% (11)	55% (11)	5% (1)	Linn et al. (1988)	
	10	21	~50	FEV ₁	43% (9)	38% (8)	14% (3)	Linn et al. (1990)	

Table 5-2 (Continued): Percentage of adults with asthma in controlled human exposure studies experiencing sulfur dioxide induced decrements in lung function and respiratory symptoms.

SO ₂ Conc (ppm)	Exposure Duration (min)	n	Ventilation (L/min)	Cumulative Percentage of Subjects (Number of Subjects) ^a				Study	Respiratory Symptoms: Supporting Studies
				↑sRaw	≥100%	≥200%	≥300%		
				↓ FEV ₁	≥15%	≥20%	≥30%		
1.0	10	28	~40	sRaw	50% (14)	25% (7)	14% (4)	Roger et al. (1985)^e	Clear and consistent increases in SO ₂ -induced respiratory symptoms: Linn et al. (1990) ; Linn et al. (1988) ; Linn et al. (1987) ; Linn et al. (1983b)), Gong et al. (1995) , Horstman et al. (1988)
10	10	~40	sRaw	60% (6)	20% (2)	0	Kehrl et al. (1987)		

Conc = concentration; FEV₁ = forced expiratory volume in 1 sec; n = sample size; sRaw = specific airway resistance; SO₂ = sulfur dioxide.

^aData presented from all references from which individual data were available. Percentage of individuals who experienced greater than or equal to a 100, 200, or 300% increase in specific airway resistance, or a 15, 20, or 30% decrease in FEV₁. Lung function decrements are adjusted for the effects of exercise in clean air (calculated as the difference between the percent change relative to baseline with exercise/SO₂ and the percent change relative to baseline with exercise/clean air).

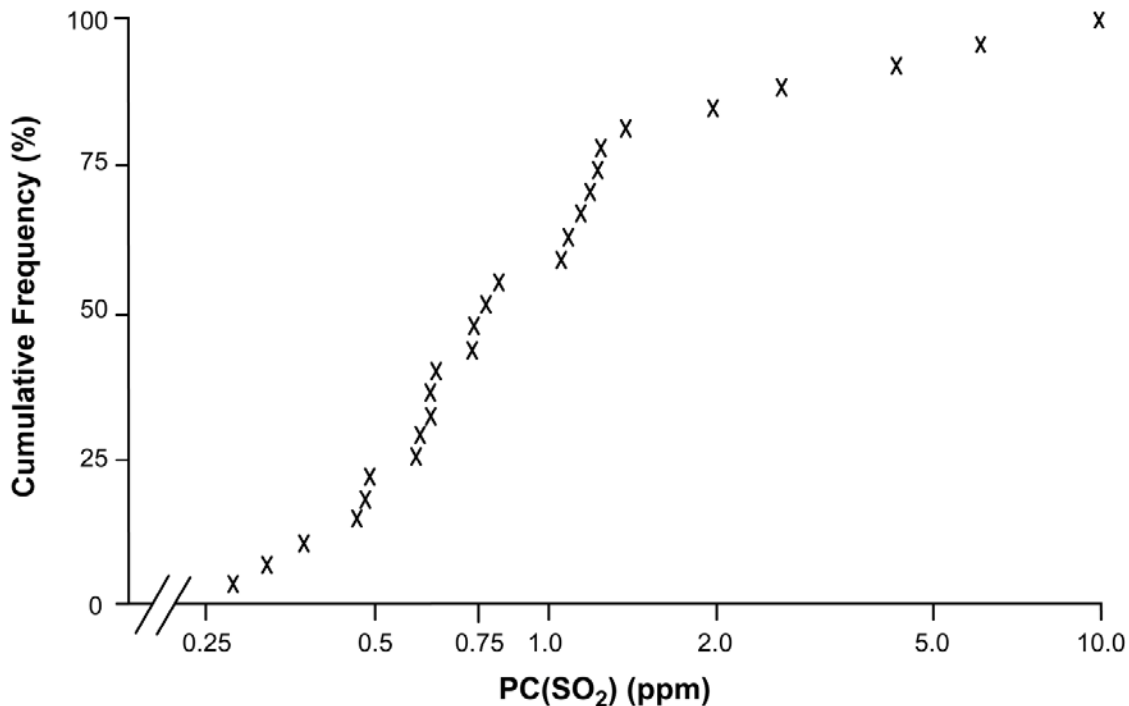
^bNumbers in parenthesis represent the number of subjects experiencing the indicated effect.

^cResponses of people with mild and moderate asthma reported in [Linn et al. \(1987\)](#) have been combined. Data are the average of the first- and second-round exposure responses following the first 10 min period of exercise. In some cases, the average had a first decimal place value of 5, which is reported in the table to avoid a high bias in values due to rounding. In all other cases, the calculated percentages were rounded to the nearest integer.

^dAnalysis includes data from only people with mild ([Linn et al., 1988](#)) and moderate ([Linn et al., 1990](#)) asthma who were not receiving supplemental medication.

^eOne subject was not exposed to 1 ppm due to excessive wheezing and chest tightness experienced at 0.5 ppm. For this subject, the values used for 0.5 ppm were also used for 1.0 ppm under the assumption that the response at 1.0 ppm would be equal to or greater than the response at 0.5 ppm.

^fIndicates studies in which exposures were conducted using a mouthpiece rather than a chamber.



PC = provocative concentration; SO₂ = sulfur dioxide.
 Note: Each data point represents the PC(SO₂) for an individual subject.
 Source: [Horstman et al. \(1986\)](#). Reprinted with permission of Sage Publications.

Figure 5-1 Distribution of individual airway sensitivity to sulfur dioxide. The cumulative percentage of subjects is plotted as a function of provocative concentration, which is the concentration of sulfur dioxide that provoked a 100% increase in specific airway resistance compared to clean air.

Completed after the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), an analysis by [Johns et al. \(2010\)](#) of publicly available primary data from published studies clearly demonstrates disparate responses among 177 adults with asthma. Data from five studies of individuals with asthma exposed to multiple concentrations of SO₂ for 5–10 minutes with elevated ventilation rates ([Linn et al., 1990](#); [Linn et al., 1988](#); [Linn et al., 1987](#); [Roger et al., 1985](#); [Linn et al., 1983b](#)) were analyzed after classifying individuals by responder status. Classification of responders versus nonresponders was based on the magnitude of sRaw and FEV₁ changes in response to the highest SO₂ concentration to which subjects were exposed (0.6 or 1.0 ppm). Responders were defined as subjects experiencing ≥100% increase in sRaw or ≥15% decrease in FEV₁ after exposure. Response status was assigned

separately for sRaw and FEV₁. Among responders, significant decreases in FEV₁ were observed for concentrations as low as 0.3 ppm SO₂ ($p = 0.005$) (Table 5-3). In addition, marginally significant increases in sRaw were demonstrated at 0.3 ppm SO₂ ($p = 0.009$), with statistically significant increases observed at 0.4 and 0.5 ppm ($p < 0.001$) (Table 5-4). [Due to multiple comparisons, Johns et al. (2010) designated a critical p -value of 0.005 as significant, using the Bonferroni multiple comparison correction.] Overall, these data demonstrate a bimodal distribution of airway responsiveness to SO₂ in individuals with asthma, with one subpopulation that is insensitive to the bronchoconstrictive effects of SO₂ even at concentrations as high as 1.0 ppm, and another subpopulation that has an increased risk for bronchoconstriction at low concentrations of SO₂. The Winterton et al. (2001) study suggests that a tumor necrosis factor alpha (TNF- α) promoter polymorphism in some individuals with asthma may be associated with increased airway responsiveness to SO₂.

Table 5-3 Percent change in post- versus pre-exposure measures of forced expiratory volume in 1 second relative to clean air control after 5–10-minute exposures to sulfur dioxide during exercise.

	SO ₂ Concentration ppm	Number of Exposures	FEV ₁			
			% Decrease	95% Confidence Limits		p -Value
				Lower	Upper	
Responders	0.2	37	-5.0	-8.9	-1.1	0.012
	0.3	20	-7.6	-13.0	-2.3	0.005 ^{a,b}
	0.4	37	-17.4	-21.3	-13.6	<0.001 ^{a,b}
Nonresponders	0.2	43	0.4	-4.3	5.2	0.854
	0.3	21	-3.6	-9.6	2.5	0.252
	0.4	43	-4.3	-9.2	0.6	0.086

FEV₁ = forced expiratory volume in 1 sec; ppm = parts per million; SO₂ = sulfur dioxide.

^aIndicates significance at 0.05 level using the Bonferroni multiple comparison correction.

^bIndicates significance at 0.05 level using Dunnett's test.

A generalized linear latent and mixed models (GLLAMM) procedure was used that included study as a fixed effect, concentration dummy variables as a covariate, and subject and the times a subject was exposed to a sequence of exposures as random variables. Data were included from Linn et al. (1987), Linn et al. (1988), and Linn et al. (1990).

Source: Table 2 of Johns et al. (2010).

Table 5-4 Percent change in post- versus pre-exposure measures of specific airway resistance relative to clean air control after 5–10-minute exposures to sulfur dioxide during exercise.

	SO ₂ Concentration ppm	Number of Exposures	% Increase	sRaw		p-Value
				95% Confidence Limits		
				Lower	Upper	
Responders	0.2	36	10.2	-3.6	24.0	0.147
	0.25	14	19.5	-4.0	43.1	0.104
	0.3	25	25.4	6.5	44.3	0.009
	0.4	36	75.7	53.4	98.0	<0.001 ^{a,b}
	0.5	14	68.0	33.2	102.8	<0.001 ^{a,b}
Nonresponders	0.2	67	7.9	-4.9	20.7	0.227
	0.25	14	12.6	-10.5	35.7	0.286
	0.3	16	16.4	-5.2	38.1	0.137
	0.4	67	16.2	1.8	30.6	0.028
	0.5	14	14.7	-12.3	41.7	0.285

sRaw = specific airway resistance; ppm = parts per million; SO₂ = sulfur dioxide.

^aIndicates significance at 0.05 p level, using the Bonferroni multiple comparison correction.

^bIndicates significance at 0.05 level using Dunnett's test.

A generalized linear latent and mixed models (GLLMM) procedure was used that included study as a fixed effect, concentration dummy variables as a covariate, and subject and the times a subject was exposed to a sequence of exposures as random variables. Data were included from [Linn et al. \(1983b\)](#), [Linn et al. \(1987\)](#), [Linn et al. \(1988\)](#), [Linn et al. \(1990\)](#), and [Roger et al. \(1985\)](#).

Source: Table 1 of [Johns et al. \(2010\)](#).

A recent analysis of four previously published studies ([Horstman et al., 1988](#); [Horstman et al., 1986](#); [Schachter et al., 1984](#); [Sheppard et al., 1980](#)) in which individuals with asthma were exposed to multiple SO₂ concentrations or had their response recorded over multiple durations of SO₂ exposure was provided by [Goodman et al. \(2015\)](#). Eight of 56 individuals were identified as sensitive to the effects of SO₂ by [Goodman et al. \(2015\)](#). However, the analysis conducted by [Goodman et al. \(2015\)](#) did not consider the log-normal distribution of airway responsiveness data and instead used an arithmetic mean and standard deviation in their analysis.

Effects of asthma severity on SO₂-induced response. The influence of asthma severity on the degree of responsiveness to SO₂ exposure has been examined ([Trenga et al., 1999](#); [Linn et al., 1987](#)). One study involved exposure to SO₂ under conditions of increased ventilation (i.e., exercise) ([Linn et al., 1987](#)). Adults with asthma were divided into two groups, minimal/mild and moderate/severe, mainly based on the individual's use of medication to control asthma. Individuals that did not regularly use asthma medication were classified as minimal/mild; however, even the moderate/severe group consisted of adults who had well-controlled asthma, were generally able to withhold medication, were not dependent on corticosteroids, and were able to engage in moderate to heavy levels of exercise. Thus, this moderate/severe group would likely be classified as moderate by today's classification standards ([Johns et al., 2010](#); [Reddel, 2009](#)). [Linn et al. \(1987\)](#) found similar relative decrements in lung function in response to SO₂ exposure between the groups. Nevertheless, the moderate/severe group demonstrated larger absolute changes in lung function compared with the mild group ([Linn et al., 1987](#)). This greater decrement in lung function was attributable to a larger response to the exercise component of the exposure protocol in the moderate/severe group compared with the mild group. [Trenga et al. \(1999\)](#) found a correlation between asthma severity and response to SO₂. Adults with asthma were divided into four groups based on medication usage as an indicator of asthma severity. The role of exercise was not determined in this study, so it was unclear whether individuals with more severe asthma had a greater response to exercise compared to individuals with less severe asthma. However, both studies suggest that adults with moderate/severe asthma may have more limited reserve to deal with an insult compared with individuals with mild asthma.

Asthma with medication. Asthma medications have been shown to mitigate SO₂-induced bronchoconstriction ([U.S. EPA, 2008d](#)). Medications evaluated include short-acting and long-acting beta-adrenergic bronchodilators ([Gong et al., 1996](#); [Linn et al., 1990](#); [Linn et al., 1988](#); [Koenig et al., 1987](#)), cromolyn sodium ([Koenig et al., 1988](#); [Myers et al., 1986b](#)), theophylline ([Koenig et al., 1992](#)), and leukotriene receptor antagonists ([Gong et al., 2001](#); [Lazarus et al., 1997](#)). While these therapies have been shown to mitigate the respiratory effects of SO₂, they did not completely eliminate these effects in all studies.

Children and adolescents. Several studies have examined the responsiveness to SO₂ of adolescents (ages 12–18 years) with asthma or allergic with exercise-induced bronchospasm (EIB) ([Koenig et al., 1990](#); [Koenig et al., 1988](#); [Koenig et al., 1987](#)). Of these studies, only [Koenig et al. \(1987\)](#) included a control air exposure, so that the bronchoconstrictive effects of SO₂ itself (rather than, for example, due to EIB), can be assessed. On average, based on the data provided in Table 1 of this paper, adolescents experienced a pre- to post-exposure reduction in FEV₁ of 15.4% following exposure to

0.75 ppm SO₂ and a reduction in FEV₁ of 3.46% following air exposure. Although the adolescents in this study were allergic with EIB, they did not have extrinsic asthma. Nevertheless, they are discussed here because allergies affect airway responsiveness ([Burrows et al., 1995](#)) and because their response to SO₂ is similar to that observed in other studies of individuals with asthma. The pre- to post-reduction in FEV₁ of 15.4% following 0.75 ppm SO₂ observed by [Koenig et al. \(1987\)](#) is similar to the pre- to post-reduction in FEV₁ of 13.9% found in adolescents with asthma following exposure to 1.0 ppm SO₂ observed by [Koenig et al. \(1988\)](#). For potential comparison to the results of adolescents, three studies of adults with asthma were conducted at 0.75 ppm ([Gong et al., 2001](#); [Gong et al., 1996](#); [Linn et al., 1983a](#)). Of these, only [Gong et al. \(2001\)](#) provided pre-to-post data for both exposures to air and SO₂. Similar to the [Koenig et al. \(1987\)](#) results, [Gong et al. \(2001\)](#) observed a pre- to post-reduction of 15.8% in FEV₁ following SO₂ exposure in adults based on Table 2 of their paper. Adjusted for the responses occurring with air exposure, [Koenig et al. \(1987\)](#) observed an 11.8% reduction in FEV₁ in adolescents, similar to the 12.7% reduction observed in adults by [Gong et al. \(2001\)](#). These two studies differ in that the adolescents were exposed via a mouthpiece, whereas the adults were exposed in a chamber without a mouthpiece. Breathing on a mouthpiece is expected to produce a somewhat larger FEV₁ decrement than unencumbered breathing ([Linn et al., 1983a](#)). Although generally similar effects of SO₂ on adolescents and adults have been observed, exact comparisons of SO₂ effects between adolescents and adults are not possible given the available data.

There is also evidence that adolescents (ages 12–18 years) with asthma or atopy are responsive to coexposures of SO₂ and sodium chloride (NaCl) droplet aerosol ([Koenig et al., 1983, 1981](#); [Koenig et al., 1980](#)). Exposure concentrations in these studies ranged from 0.1 to 1.0 ppm SO₂. [Koenig et al. \(1983\)](#) observed average FEV₁ decrements of 15 and 23% in exercising adolescents (12–16 years old) with asthma after a 10-minute exposure to 0.5 ppm SO₂ plus 1 mg/m³ NaCl droplet aerosols or 1.0 ppm SO₂ plus 1 mg/m³ NaCl droplet aerosols, respectively. No significant changes were observed following exposure to the NaCl droplet aerosol alone. However, the observed effect may be the result of the presence of hygroscopic particles that carry SO₂ deeper into the lung.

There are no controlled human exposure studies for children less than 12 years of age who were exposed to SO₂. However, the responsiveness of children to SO₂ relative to adolescents and adults may be inferred by the responses to other nonspecific bronchoconstrictive stimuli. [Horstman et al. \(1986\)](#) noted that the effects of SO₂ on sRaw are similar to that of a variety of nonspecific bronchoconstrictive stimuli. Indeed, SO₂ is a nonspecific bronchial challenge agent that has been used to assess changes in airway responsiveness of individuals with asthma following NO₂ and O₃ exposures ([Trenka et al., 2001](#); [Jörres and Magnussen, 1990](#); [Rubinstein et al., 1990](#)). Airway responsiveness

to methacholine, a history of respiratory symptoms, and atopy were significant predictors of airway responsiveness to SO₂ in healthy adults ([Nowak et al., 1997](#)). Thus, potential differences in airway responsiveness of children to SO₂ relative to adolescents and adults may be gleaned from the literature on airway responsiveness to other nonspecific stimuli such as methacholine.

A number of cross-sectional studies have assessed airway responsiveness of children with and without asthma to methacholine [e.g., ([Mochizuki et al., 1995](#); [Morikawa et al., 1994](#); [Avital et al., 1991](#); [Hopp et al., 1986](#); [Hopp et al., 1985](#))]. Studies show a clear decrease in airway responsiveness of healthy children with increasing age beyond 5–7 years of age through adolescence ([Mochizuki et al., 1995](#); [Hopp et al., 1986](#); [Hopp et al., 1985](#)). In studies of children with asthma, some have reported airway responsiveness increased with asthma severity but was not affected by age ([Avital et al., 1991](#); [Hopp et al., 1986](#)), whereas others have found airway responsiveness to increase with asthma severity and decrease with age beyond 6–7 years of age ([Mochizuki et al., 1995](#); [Morikawa et al., 1994](#)). The study by [Mochizuki et al. \(1995\)](#) suggested that airway responsiveness in both healthy children and those affected by asthma increases from ages 2–3 years up to 6–7 years, after which airway responsiveness begins decreasing.

More confidence in the effect of age on airway responsiveness may be placed on data from longitudinal studies than from the cross-sectional studies discussed above. In a longitudinal study of methacholine responsiveness conducted at 9, 11, 13, and 15 years of age, [Le Souëf et al. \(1995\)](#) found that responsiveness (1) decreases with age; (2) is greater in boys (n = 389) than girls (n = 429); and (3) is greater in those reporting wheeze, although responsiveness decreased with age in these individuals as well. Asthma prevalence and symptoms such as wheeze are greater in boys than girls during childhood and become similar or reversed around the time of puberty ([Almqvist et al., 2008](#)). In a subset of the cohort as used by [Le Souëf et al. \(1995\)](#), [Burrows et al. \(1995\)](#) investigated the effects of age (n = 573, 49% female), atopy (n = 558), and serum immunoglobulin E (IgE) (n = 473) on airway responsiveness. At 9 years of age, a larger fraction of boys experienced AHR compared with girls. By the age of 15 years, there was little to no difference in the fraction with AHR between the sexes. Relative to atopic children, those without atopy or with only minimal atopy had a lower fraction with AHR and showed a more evident decrease in the fraction having AHR with increasing age. In the most atopic children (41 of 558), about 20-30% experienced severe AHR, which did not decrease with age. Across all ranges of serum IgE, there was a decrease in the fraction having AHR from age 9 to age 15 years. By 15 years of age, only a small fraction of the children with low serum IgE levels had AHR. At both 9 and 15 years of age, the fraction having AHR increased with increasing serum IgE levels ($p < 0.0001$). In biennial assessments of childhood responsiveness, [Burrows et al. \(1995\)](#) observed considerable intra-individual

variability in bronchial reactivity, but they observed a statistically significant trend for the more allergic children to experience persistent AHR among their biennial assessments.

A number of factors may influence bronchial responsiveness to SO₂ including innate responsiveness of the airways, route of breathing, disease status, and age. To the extent that variability in bronchial responsiveness to SO₂ may be inferred from studies evaluating responsiveness to methacholine, these studies suggest that greater airway responsiveness to SO₂ may occur in school-aged children (~5–11 years of age), particularly boys, than in adolescents. Additionally, the methacholine data also suggest that greater airway responsiveness to SO₂ in school-aged children and adolescents who are allergic or experience wheeze is expected to occur than in those without these conditions. Children, particularly boys, breathe more through the mouth than adults, and ventilation rates relative to body mass are greater in children than adults (see [Section 4.1.2](#)). Allergic rhinitis can lead to increased nasal resistance, which also results in less nasal and more oral breathing. Obese children also tend to have increased nasal resistance, increased oral breathing, and increased ventilation rates relative to normal-weight children (see [Section 4.1.2](#)). Oral breathing allows greater SO₂ penetration into the lower airways, where it may cause bronchoconstriction, than does nasal breathing (see [Section 4.2.2](#)). Overall, normal-weight and obese school-aged children (~5–11 years of age) having asthma-like symptoms might be expected to experience greater responsiveness (i.e., larger decrements in pulmonary function) following exposure to SO₂ than normal-weight adolescents and adults. Boys may be particularly affected due the combined effects of increased bronchial responsiveness and a greater degree of mouth breathing.

Mixtures effects. The health effects of SO₂ can be potentially modified by the interaction with other pollutants during or prior to exposure. A few studies involving mixtures with NaCl droplet aerosol are discussed above. A few controlled human exposure studies have examined the interactive effects of O₃ and SO₂ both sequentially and in combination. Exercising adolescents with asthma exposed to 0.1 ppm SO₂ for 15 minutes after a 45-minute exposure to 0.12 ppm O₃ had a significant decrease (8%) in FEV₁ (8%) ($p < 0.05$), a significant increase in total respiratory resistance (R_T) (19%) ($p < 0.05$), and a significant decrease in maximal flow at 50% of expired vital capacity (V_{max50}) (15%) ($p < 0.05$); while air followed by SO₂, and O₃ followed by O₃ exposures did not cause significant changes in lung function ([Koenig et al., 1990](#)). In a more recent study in exercising adults with asthma, [Trenga et al. \(2001\)](#) observed greater decrements in lung function after 45 minutes of exposure to 0.12 ppm O₃ followed by 15 minutes of 0.25 ppm SO₂ compared to air followed by SO₂.

[Jörres and Magnussen \(1990\)](#) and [Rubinstein et al. \(1990\)](#) investigated the effects of prior NO₂ exposure on SO₂-induced bronchoconstriction in adults with asthma. [Jörres and Magnussen \(1990\)](#) observed that tidal breathing of NO₂ at rest increased airway responsiveness to subsequent hyperventilation of SO₂. [Rubinstein et al. \(1990\)](#) noted NO₂ exposure during exercise induced greater airway responsiveness to inhaled SO₂ in only one subject of nine. The effect of exercising versus resting exposures to NO₂ on airway responsiveness is discussed elsewhere ([Brown, 2015](#)).

While SO₂ acts as a nonspecific bronchial challenge agent that causes reductions in lung function in individuals with asthma after brief exposure, it can also affect airway responsiveness to subsequent exposures involving other stimuli such as allergens or methacholine. Two studies of adults with asthma exposed at rest to SO₂ in combination with NO₂ demonstrated increases in airway responsiveness to subsequent allergen challenge ([Rusznak et al., 1996](#); [Devalia et al., 1994](#)). In the first of these studies, exposure to 0.2 ppm SO₂ or 0.4 ppm NO₂ did not affect airway responsiveness to house dust mite allergen immediately after a 6-hour exposure. In considering the effect of SO₂ alone, it is unlikely that enough SO₂ reached the bronchial airways to cause an effect because volunteers were exposed at rest. Following exposure to the two pollutants in combination, volunteers demonstrated an increased response to inhaled allergen ([Devalia et al., 1994](#)). [Rusznak et al. \(1996\)](#) confirmed these results in a similar study and found that increased airway responsiveness to dust mites persisted up to 48 hours post-exposure. These results provide evidence that exposure to SO₂ in combination with NO₂ increases airway responsiveness to a subsequent allergen challenge. This effect is longer in duration than other effects typically associated with exposure to SO₂.

Epidemiologic Studies

Unlike controlled human exposure studies, epidemiologic studies inconsistently indicate SO₂-related lung function decrements in populations with asthma. This inconsistency applies to previous ([U.S. EPA, 2008d](#)) and recent ([Tables 5-5](#) and [5-6](#)) studies, as well as those involving adults and children with asthma. Epidemiologic studies examined longer SO₂ averaging times and lags and had uncertainty in exposures because the exposures were estimated from fixed-site monitors. For the few findings of SO₂-associated lung function decrements, confounding by moderately to highly correlated PM and NO₂ ($r = 0.54-0.9$) was not examined. A few recent studies address some of these uncertainties, but they persist in the evidence overall.

Table 5-5 Epidemiologic studies of lung function in adults with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Qian et al. (2009b) Boston, MA; New York City, NY; Philadelphia, PA; Madison, WI; Denver, CO; San Francisco, CA; 1997–1999 N = 154, ages 12–65 yr. 100% persistent asthma. 1/3 ICS use, 1/3 beta-agonist use, 1/3 placebo use. Daily measures for 16 wk. Home PEF. Recruited from clinics as part of an asthma medication trial. Multiple comparisons—many pollutants, lags, medication use analyzed.</p>	<p>Monitors avg within 32 km of subject ZIP code centroid. Mean (SD): 4.8 (3.9) 75th percentile: 6.2 Max: 32</p>	24-h avg 0	<p>Change in PEF (L/min) All subjects: -0.12 (-3.0, 2.7) ICS: -8.4 (-13, -3.4) Beta-agonist: 4.4 (-0.49, 9.3) Placebo: 3.3 (-1.4, 8.0)</p>	<p>Persist with: PM₁₀, NO₂, or O₃ (ICS users) PM_{2.5} not examined. r = 0.58 NO₂, NR for PM₁₀.</p>
		0–2 avg	<p>All subjects: -1.9 (-5.6, 1.7) ICS: -13 (-18, -6.4) Beta-agonist: 6.4 (0.14, 13) Placebo: 0.85 (-5.2, 6.9)</p>	
<p>†Maestrelli et al. (2011) Padua, Italy, 2004–2005 N = 32, mean (SD) age 40 (7.5) yr. 81% persistent asthma. 69% ICS use. 90% atopy. 6 measures over 2 yr. Supervised spirometry. Recruited from database of beta-agonist users (>6 times per yr for 3 yr).</p>	<p>Two monitors in city Medians across seasons: 0.87–2.7 75th percentiles across seasons: 1.3–4.1</p>	24-h avg 0	<p>Change in % predicted FEV₁ All subjects: -2.1 (-6.6, 2.3) Nonsmokers: -11 (-40, 18)</p>	<p>No copollutant model Copollutant correlations NR.</p>
<p>†Canova et al. (2010) Padua, Italy, 2004–2005 N = 19, ages 15–44 yr. 79% moderate/severe asthma. 58% ICS use. Daily measures for five 30-day periods over 2 yr. Home PEF/FEV₁. Part of same cohort as Maestrelli et al. (2011) above.</p>	<p>Two monitors in city Mean (SD): 1.4 (1.1) Max: 4.9</p>	24-h avg 0, 1, 2, 3, 0–1 avg, 0–3 avg	<p>Quantitative effect estimates NR. Figure shows negative but imprecise associations for PEF and FEV₁ with wide 95% CIs.</p>	<p>No copollutant model PM_{2.5} not examined. Spearman r = 0.50 CO, 0.51 PM₁₀, 0.54 NO₂.</p>

Table 5-5 (Continued): Epidemiologic studies of lung function in adults with asthma published since the 2008 Integrated Science Assessment for Sulfur Oxides.

<p>†Wiwatanadate and Liwsrisakun (2011) Chiang Mai, Thailand, 2005–2006 N = 121, ages 13–78 yr. 48% moderate/severe persistent asthma. Daily measures for 10 mo. Home PEF. Recruited from allergy clinics.</p>	<p>Monitor within 10 km of home Mean (SD): 1.7 (0.62) 90th percentile: 2.4 Max: 3.9</p>	<p>24-h avg 4</p>	<p>NR</p>	<p>Only multipollutant models analyzed $r = 0.23 \text{ NO}_2, -0.07 \text{ PM}_{2.5}$.</p>
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Avg = average; CI = confidence interval; CO = carbon monoxide; FEV₁ = forced expiratory volume in 1 sec; ICS = inhaled corticosteroid; ISA = Integrated Science Assessment; max = maximum; n = sample size; N = population number; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; PEF = peak expiratory flow; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; r = correlation coefficient; SD = standard deviation; SO₂ = sulfur dioxide.

^aEffect estimates are standardized to a 10-ppb increase in 24-h avg SO₂.

†Studies published since the 2008 ISA for Sulfur Oxides.

Adults. Previous studies were limited to Europe and Asia. A recent study shows an SO₂-associated decrease in lung function in adults with asthma in the U.S. ([Qian et al., 2009b](#)); however, recent studies in Europe and Asia do not show this decrease ([Maestrelli et al., 2011](#); [Wiwatanadate and Liwsrisakun, 2011](#); [Canova et al., 2010](#)) (Table 5-5). Mean and upper percentile SO₂ concentrations tended to be lower in recent studies than in previous studies [e.g., means for 24-h avg 0.87–4.8 ppb vs. 1.6–90 ppb in ([Boezen et al., 2005](#); [Neukirch et al., 1998](#); [Peters et al., 1996a](#))]. However, lower concentrations do not appear to account for the weak recent evidence in adults with asthma because previous studies with mean SO₂ concentrations of 5.2 to 90 ppb did not observe SO₂-associated lung function decrements ([Park et al., 2005](#); [Peters et al., 1996a](#)). Recent studies did not differ in temporal variability (e.g., ratio of the mean concentration to standard deviation) in SO₂ concentrations, which is the basis of analysis in these repeated measures studies.

The U.S. multicity study provides supporting evidence but has the same uncertainty in the exposure estimate as do other studies in adults with asthma. All studies estimated SO₂ exposure from fixed-site monitors, either a single monitor or average of many monitors. Ambient SO₂ concentrations tend to show high spatiotemporal variability within a city, and correlations with personal exposure are low to moderate ([Section 3.4.1.3](#)). The studies did not discuss whether measurements at the monitors adequately represented the spatiotemporal variability in ambient SO₂ concentrations in the study area. Uncertainty is high in the U.S. study, which averages SO₂ concentrations across monitors within 32 km of subjects' ZIP code centroid ([Qian et al., 2009b](#)). Ambient SO₂ concentrations show large, transient peaks ([Section 2.5.3](#)), which may be important based on results from controlled human exposure studies showing that 5- to 10-minute exposures to 200–600 ppb SO₂ induce rapid and short-lived lung function decrements. Epidemiologic studies examined same-day (lag 0) SO₂ concentrations. Daily average SO₂ concentrations may not represent peak exposures or capture the transient effects of peak exposures implicated in controlled human exposure studies.

Some recent studies that did not observe SO₂-related lung function decrements had small sample sizes (N = 19 or 32) ([Maestrelli et al., 2011](#); [Canova et al., 2010](#)). It is unclear, however, whether sample size explains the inconsistency among adults with asthma or AHR overall. Similarly sized studies ([Boezen et al., 2005](#); [Neukirch et al., 1998](#)) observed associations, and larger studies do not show evidence for association ([Wiwatanadate and Liwsrisakun, 2011](#); [Park et al., 2005](#); [Peters et al., 1996a](#)). In panel studies, the number of repeated measurements is also important, and [Canova et al. \(2010\)](#) measured lung function for five 30-day periods. Many studies that had a large number of repeated measurements examined lung function measured by subjects at home not supervised by a trained technician. Results were inconsistent for both methodologies.

A few epidemiologic studies add information on response modification by asthma phenotype but produce no clear finding. Previous results found an association between SO₂ exposure and decreased lung function in adults with AHR and elevated IgE ([Boezen et al., 2005](#)), and an association between SO₂ exposure and increased AHR in adults with physician-diagnosed asthma and AHR ([Taggart et al., 1996](#)). A recent study did not find an association between SO₂ exposure and lung function decrements in adults with physician-diagnosed asthma ([Maestrelli et al., 2011](#)). Most of the subjects in [Taggart et al. \(1996\)](#) and [Maestrelli et al. \(2011\)](#) were atopic. A 10-ppb increase in 24-h avg SO₂ was associated with a -2.1 point change [95% confidence interval (CI): -6.6, 2.3] in percent predicted FEV₁. Like the controlled human exposure studies, the epidemiologic studies do not clearly show that SO₂-associated lung function decrements depend on asthma severity. An association was observed in adults with mild to moderate asthma ([Neukirch et al., 1998](#)), and the results varied among populations with more severe asthma ([Maestrelli et al., 2011](#); [Canova et al., 2010](#); [Qian et al., 2009b](#)). In contrast with the controlled human exposure studies, the U.S. asthma medication trial observed an SO₂-related decrease in lung function in adults randomized to daily inhaled corticosteroid (ICS) use [-8.4 L/minute (95% CI: -13, -3.4) change in peak expiratory flow (PEF) per 10-ppb increase in 24-h avg SO₂] ([Qian et al., 2009b](#)). Decrements were not observed in the beta-agonist or placebo groups ([Table 5-5](#)). These two groups had more frequent asthma exacerbation during the study than the corticosteroid group but similar PEF and mean age ([Lazarus et al., 2001](#)). All three groups had persistent asthma. Thus, a clear explanation for the pattern of SO₂ associations is not apparent. There is no clear rationale for attributing null findings to the lack of analysis stratified by corticosteroid use, particularly for results that were adjusted for such use ([Maestrelli et al., 2011](#); [Canova et al., 2010](#)).

Across studies, the potential influence of copollutants is largely unaddressed. No study in adults with asthma examined PM_{2.5} total mass, and previous studies observed lung function decrements in association with larger sized PM metrics that were highly correlated with SO₂ concentrations ($r = 0.8-0.9$) and sulfate ([Neukirch et al., 1998](#); [Peters et al., 1996a](#)). That some cities had a coal-fired power plant or used coal for heating may explain some of the high correlations with PM and moderate correlations with NO₂ ($r = 0.54$) ([Neukirch et al., 1998](#); [Taggart et al., 1996](#)). Copollutant interactions were not assessed. Only the recent U.S. study analyzed confounding. SO₂ was moderately correlated with NO₂ [$r = 0.58$, correlations with other pollutants NR)] ([Qian et al., 2009b](#)). SO₂ was negatively associated with PEF in the corticosteroid group, and effect estimates persisted with adjustment for PM₁₀, NO₂, or O₃ ([Table 5-5](#)). However, inference from the results is weak due to numerous comparisons across pollutants, lags, and medication groups and questionable reliability in the exposures estimated from monitors up to 32 km away.

Children. As with adults, evidence from neither the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) nor recent studies ([Table 5-6](#)) consistently links increases in ambient SO₂ concentration with lung function decrements in children with asthma, including recent U.S. multicity studies ([Ierodiakonou et al., 2015](#); [O'Connor et al., 2008](#)). The inconsistency does not appear to be explained by lung function measured under supervised conditions or by subjects at home, asthma severity, or prevalence of asthma medication use. In contrast to adults with asthma, SO₂-associated lung function decrements were not observed in children with asthma who took inhaled corticosteroids ([Ierodiakonou et al., 2015](#); [Liu et al., 2009b](#)). Among children with asthma in Windsor, ON, the association was limited to nonusers ([Liu et al., 2009b](#)). For some recent studies, including a U.S. multicity study, inference about an SO₂ effect is weak because the association was isolated to one lung function parameter or exposure lag among numerous lung function parameters, lags, pollutants, and/or asthma medication groups examined ([Ierodiakonou et al., 2015](#); [Wiwatanadate and Trakultivakorn, 2010](#)).

Table 5-6 Epidemiologic studies of lung function in children with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Greenwald et al. (2013) El Paso, TX, Mar–Jun 2010 N = 38, mean age 10 yr. 47% daily asthma medication use. Weekly measures for 13 wk. Supervised spirometry. Recruited from schools.</p>	<p>Monitor at school A: residential area B: 91 m from major road Mean (SD): 1.2 (0.44) and 0.84 (0.54) Upper percentiles NR.</p>	<p>24-h avg 0–3 avg</p>	<p>Percent change in FEV₁ A: 15 (–60, 210) B: –31 (–52, –2.0)</p>	<p>No copollutant model PM_{2.5} not associated Pearson <i>r</i> = –0.14 BC, –0.22 NO₂, –0.07 BTEX, 0.14 cleaning product VOCs.</p>
<p>†Dales et al. (2009) Windsor, ON, Oct–Dec 2005 N = 182, ages 9–14 yr. 37% ICS use, 35% beta-agonist use. Daily measures for 4 wk. Home FEV₁. Recruited from schools. Mean 1.6 and 2.2 h/day spent outdoors for two study groups.</p>	<p>Two monitors avg 99% homes within 10 km of sites. Median: 4.5 95th percentile: 16</p>	<p>12-h avg 8 a.m.–8 p.m. 8 p.m.–8 a.m. 8-h avg 12 a.m.–8 a.m. 24-h avg</p>	<p>Percent change in FEV₁ Bedtime: 0 (–0.92, 0.93) Diurnal: –1.41 (–2.73, –0.08) Bedtime: –0.17 (–0.98, 0.65) Morning: 0.63 (–0.28, 1.55) Bedtime: –0.14 (–1.03, 0.76)</p>	<p>Persists with: PM_{2.5}, NO₂, or O₃. Pearson <i>r</i> = 0.43 PM_{2.5}, 0.31 NO₂.</p>
<p>†Liu et al. (2009b), Liu (2013) Windsor, ON, Oct–Dec 2005 N = 182, ages 9–14 yr. 37% ICS use, 35% beta-agonist use. Weekly measures for 4 wk. Supervised spirometry. Same cohort as Dales et al. (2009) above.</p>	<p>Two monitors avg 99% homes within 10 km of sites. Median: 4.5 95th percentile: 16</p>	<p>24-h avg 0 0–2 avg</p>	<p>Percent change FEV₁: –0.46 (–2.0, 1.1) FEF_{25–75%}: –1.5 (–4.7, 2.0) Change in percent predicted FEV₁: –2.0 (–4.6, 0.74) FEF_{25–75%}: –5.7 (–11, –2.2)</p>	<p>FEF_{25–75%}, lag 0–2 avg Persists with: O₃ Does not persist with: PM_{2.5} or NO₂ Spearman <i>r</i> = 0.56 PM_{2.5}, 0.18 NO₂, –0.02 O₃.</p>

Table 5-6 (Continued): Epidemiologic studies of lung function in children with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†O'Connor et al. (2008) Inner-City Asthma Study cohort: Boston, MA; Bronx, NY; New York City, NY; Chicago, IL; Dallas, TX; Tucson, AZ; Seattle, WA; 1998–2001 N = 861, ages 5–12 yr. 100% persistent asthma. 100% atopy. Daily measures for four 2-wk periods. Home FEV₁/PEF. Recruited from intervention study.</p>	<p>Monitors avg close to home and not near industry. Median 2.3 km to site. Quantitative SO₂ data NR.</p>	<p>24-h avg 1–5 avg</p>	<p>Change in percent predicted FEV₁: –1.29 (–2.04, –0.54) PEF: –1.73 (–2.49, –0.96) No association for lag 1.</p>	<p>No copollutant model r = 0.37 PM_{2.5}, 0.59 NO₂.</p>
<p>†Amadeo et al. (2015) Pointe-à-Pitre, Guadeloupe, 2008–2009 N = 71, ages 8–13 yr. Cross-sectional. Supervised spirometry. Recruited from schools.</p>	<p>Monitors in city Number and distance NR Mean (SD): 1.8 (1.4) Max: 4.9</p>	<p>24-h avg 0–13 avg</p>	<p>Change in prerun PEF (L/min) 93 (–28, 214) Percent change post 6-min run –1.6 (–36, 33)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>
<p>†Ierodiakonou et al. (2015) Childhood Asthma Management Program cohort: Boston, MA; Baltimore, MD; St. Louis, MO; Denver, CO; Albuquerque, NM; San Diego, CA; Toronto, ON, 1993–1999 N = 1,003, ages, 5–12 yr. 100% mild/moderate asthma. 30% ICS use. 30% mast cell inhibitor use. 14 measures over 4 yr. Supervised spirometry. Recruited from clinics. Multiple comparisons—many pollutants, lags, exposure durations, medication use analyzed.</p>	<p>Nearest monitor within 50 km of ZIP code centroid. Medians across cities: 2–6 90th percentiles across cities: 5–24</p>	<p>24-h avg 0</p>	<p>Change in percent predicted Prebronchodilator FEV₁ All subjects 0.25 (–0.13, 0.63) ICS: 0.38 (–0.30, 1.1) Post-bronchodilator FEV₁ ICS: 0 (–0.73, 0.75) Change in methacholine that induces a 20% drop in FEV₁ Mast cell inhibitor: –13% (–25, 1.3)</p>	<p>No copollutant model PM_{2.5} not examined. Spearman r across cities = 0.19–0.34 CO, –0.41 to –0.05 O₃, 0.15–0.54 NO₂.</p>

Table 5-6 (Continued): Epidemiologic studies of lung function in children with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
†Wiwatanadate and Trakultivakorn (2010) Chiang Mai, Thailand, 2005–2006 N = 31, ages 4–11 yr. 100% with symptoms in previous yr. 52% mild intermittent asthma Daily measures for 1 yr. Home PEF. Recruited from allergy clinic. Multiple comparisons—many pollutants, lags, lung function parameters analyzed.	Monitor within 25 km of home Mean (SD): 1.7 (0.62) 90th percentile: 2.4 Max: 3.9 ppb	24-h avg	Change in PEF (L/min)	Daily avg PEF lag 4
		0	Evening PEF	Persists with: O ₃
		4	–8.1 (–25, 9.2)	PM _{2.5} not associated
		4	–21 (–38, –4.1)	$r = -0.04$ O ₃ , -0.07 PM _{2.5} , 0.38 CO, 0.23 NO ₂ .
		0	Daily avg PEF	
		4	–0.3 (–15, 15)	
			–18 (–32, –2.8)	

Avg = average; BC = black carbon; BTEX = benzene, toluene, ethylbenzene, xylene; CI = confidence interval; CO = carbon monoxide; FEF_{25–75%} = forced expiratory flow at 25–75% of forced vital capacity; FEV₁ = forced expiratory volume in 1 sec; ICS = inhaled corticosteroid; ISA = Integrated Science Assessment; max = maximum; n = sample size; N = population number; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; PEF = peak expiratory flow; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; r = correlation coefficient; SD = standard deviation; SO₂ = sulfur dioxide; VOC = volatile organic compound.

^aEffect estimates are standardized to a 10-ppb increase in 8-h to 24-h avg SO₂.

†Studies published since the 2008 ISA for Sulfur Oxides.

Considerations in Interpreting Epidemiologic Evidence for Lung Function. A few recent studies aimed to address uncertainty in the exposure estimates, lag structure of associations, or copollutant confounding ([Greenwald et al., 2013](#); [Dales et al., 2009](#); [Liu et al., 2009b](#)) and provide limited indication of SO₂-associated lung function decrements. For children in El Paso, TX, [Greenwald et al. \(2013\)](#) measured SO₂ concentrations inside and outside of two schools. For children attending the school near a major road, a 10-ppb increase in lag 0–3 avg SO₂ was associated with a –31% change (95% CI: –52, –2.0) in FEV₁. This is the largest effect estimate among children or adults with asthma, but a 10-ppb increase in 4-day avg SO₂ is unlikely in the area [school mean 0.84 (SD: 0.54) ppb]. Results are inconsistent for 24-h avg SO₂ assigned from monitors 2.3 to 50 km from children’s homes or schools ([Amadeo et al., 2015](#); [Ierodiakonou et al., 2015](#); [Dales et al., 2009](#); [Liu et al., 2009b](#); [O’Connor et al., 2008](#)). Lung function decreased with increases in SO₂ concentrations at a monitor located a median distance of 2.3 km from children’s homes ([O’Connor et al., 2008](#)) but not a monitor within 50 km of children’s ZIP code centroid ([Ierodiakonou et al., 2015](#)) ([Table 5-6](#)). The studies did not discuss the adequacy of using monitors at these distances to represent temporal variation in SO₂ exposure. No association was observed with the change in PEF after a 6-minute exercise period ([Amadeo et al., 2015](#)), but this protocol does not mimic controlled human exposure studies because PEF was examined in relation to 13-day avg SO₂.

The lag structure of associations between ambient SO₂ concentration and observed reductions in lung function varies among epidemiologic studies. Previous studies reported associations with same day (i.e., lag 0) SO₂ concentrations ([Delfino et al., 2003b](#); [Peters et al., 1996a](#)). Recent studies point to a potential prolonged response to SO₂ exposure through evidence of associations with 3- to 5-day avg SO₂ concentrations ([Greenwald et al., 2013](#); [Liu et al., 2009b](#); [O’Connor et al., 2008](#)) that are larger in magnitude than those for lag 0 or 1 day ([Table 5-6](#)).

Correlations of copollutants [PM_{2.5}, PM₁₀, sulfate, black carbon (BC), organic carbon (OC), total suspended solids (TSP), NO₂ or volatile organic compound (VOCs)] with SO₂ were moderate ($r = 0.56–0.59$) in some recent studies ([Liu et al., 2009b](#); [O’Connor et al., 2008](#)) and high in previous studies ($r = 0.8–0.9$) ([Delfino et al., 2003b](#); [Peters et al., 1996a](#)). SO₂ averaging times varied across studies, making it difficult to assess whether higher correlations were due to higher air pollution levels in the past. Correlations were weak for school measurements but were not reported specifically for a school near a major road ([Greenwald et al., 2013](#)) where confounding by BC and VOCs could be more likely. Copollutant confounding and interactions are poorly studied, and unstudied for children living near a coal-fired power plant ([Peters et al., 1996a](#)). SO₂ and O₃ measurements at fixed-site monitors were not correlated ($r = -0.02$), and SO₂ associations persisted with adjustment for O₃ ([Dales et al., 2009](#); [Liu et al., 2009b](#)). A

recent study adds information on SO₂ results adjusted for correlated copollutants. Among children with asthma in Windsor, ON, the SO₂ association persisted with adjustment for PM_{2.5} or NO₂ for 12-h avg SO₂ ([Dales et al., 2009](#)) but not 24-h avg SO₂ ([Liu, 2013](#); [Liu et al., 2009b](#)) ([Table 5-6](#)). Associations for PM_{2.5} were robust to SO₂ adjustment, but inference about confounding is weak due to the moderate SO₂-PM_{2.5} correlation ($r = 0.56$) and the potential differences in exposure error for SO₂ and PM_{2.5}, which were made up to 10 km from subjects' homes. Weak inference also applies to results in a Los Angeles, CA cohort not supporting an association for SO₂ after adjustment for benzene [-34 L/minute change in PEF (95% CI: -120, 52) per 40-ppb increase in 1-h max SO₂] ([Delfino et al., 2003b](#)). SO₂ was highly correlated with benzene ($r = 0.70$), and pollutants were measured up to 4.8 km from home or school.

Summary of Lung Function Changes in Populations with Asthma

Controlled human exposure studies provide strong evidence for SO₂-induced lung function decrements in adults with asthma under increased ventilation conditions. Short-term exposures for 5–10 minutes to 0.2–0.3 ppm SO₂ resulted in 5–30% of exercising individuals with asthma experiencing moderate or greater decrements (defined in terms of a $\geq 15\%$ decrease in FEV₁ or $\geq 100\%$ increase in sRaw; [Table 5-2](#)). Exposures for 5–10-minutes to SO₂ at concentrations ≥ 0.4 ppm results in moderate or greater decrements in lung function in 20–60% of exercising individuals with asthma. A group of responders (defined as having $\geq 15\%$ decrease in FEV₁ after exposure to 0.6 or 1.0 ppm SO₂) showed statistically significant decrements in FEV₁ following exposure for 5–10 minutes to 0.3 ppm SO₂ ([Table 5-3](#)). Less evidence is available from controlled human exposure studies to assess SO₂-induced lung function decrements in children with asthma. However, school-aged children (~5–11 years of age), particularly boys and perhaps obese children, might be expected to experience greater responsiveness (i.e., larger decrements in lung function) following exposure to SO₂ than normal-weight adolescents and adults.

For both adults and children with asthma, epidemiologic evidence is inconsistent for lung function decrements associated with ambient SO₂ concentrations ([Tables 5-5](#) and [5-6](#)). However, one study found an association between ambient SO₂ concentration and lung function decrements in a population with AHR and elevated IgE, and another found an association between ambient SO₂ concentration and AHR in a population with asthma and a high prevalence of atopy. Evidence from animal toxicological studies, provides coherence for the timing of effects observed in recent epidemiologic studies. Specifically, when examining associations between ambient SO₂ concentration and observed reductions in lung function, the pattern of associations was found to vary among epidemiologic studies, with some studies reporting associations with same day SO₂

concentrations and recent studies pointing to a potential prolonged response ([Table 5-6](#)). The potential prolonged effect of SO₂ is supported by rodent studies demonstrating enhanced allergic inflammation after repeated SO₂ exposures. Allergic inflammation may mediate lung function decrements and provide biological plausibility for the epidemiologic associations observed due to multiday SO₂ concentrations, particularly in populations with elevated IgE or atopy. Findings of increased airway responsiveness could not be attributed to exposure to SO₂ alone as epidemiologic studies did not examine copollutant confounding, and controlled human exposure studies only examined SO₂ and NO₂ coexposures. A limitation across epidemiologic studies is the uncertainty in the SO₂ exposure estimates. A recent study observed an association with SO₂ measured at children's schools, but others used monitors located 2.3–50 km from subjects' homes or schools. It is unclear whether the SO₂ concentrations at the available fixed-site monitors adequately represent the variation in personal exposure, especially if peak exposures are as important as indicated by controlled human exposure studies. The influence of copollutants on epidemiologic results remains largely uncharacterized, including associations in populations with AHR and elevated IgE or asthma and a high prevalence of atopy, and populations living near SO₂ sources. SO₂-related lung function decrements in adults and children with asthma are inconsistently observed after adjustment for PM_{2.5}, PM₁₀, or NO₂, but the implications of these results are unclear because of the uncertainty in the SO₂ exposure estimates and potential differences in exposure error for PM_{2.5} ([Table A-1](#)).

Respiratory Symptoms in Populations with Asthma

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) reported strong evidence for the effects of SO₂ exposure on respiratory symptoms in controlled human exposure studies in individuals with asthma under increased ventilation conditions. The same studies also observed SO₂-induced decrements in lung function, although respiratory symptoms occurred consistently at higher SO₂ concentrations ([Table 5-2](#)). No new controlled human exposure studies have been reported since the previous ISA. The available epidemiologic studies do not provide insight into the concurrence between lung function and symptom changes. In contrast to evidence from controlled human exposure studies, previous and recent epidemiologic evidence for SO₂-associated increases in respiratory symptoms is weak in adults with asthma. However, epidemiologic evidence supports associations in children with asthma, and recent studies add evidence for estimates of SO₂ exposure at school and/or home. Overall, the influence of copollutants remains largely unexamined.

Controlled Human Exposure Studies

As reviewed in the 2008 ISA for Sulfur Oxides and the 1986 Supplement to the Second Addendum ([U.S. EPA, 2008d, 1994](#)), controlled human exposure studies demonstrate increases in incidence or severity of respiratory symptoms (i.e., cough, chest tightness, throat irritation) in individuals with asthma exposed to SO₂ concentrations between 0.2 and 0.6 ppm for 5–10 minutes during exercise ([Tables 5-2](#) and [5-7](#)). Statistically significant increases are observed at SO₂ concentrations ≥0.4 ppm [e.g., [Linn et al. \(1983b\)](#)].

Table 5-7 Study-specific details from controlled human exposure studies of respiratory symptoms.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Time of Symptom Assessment
Balmes et al. (1987)	Asthma; n = 8; 6 M, 2 F (23–39 yr)	0, 0.5, or 1 ppm SO ₂ for 1, 3, or 5 min during eucapnic hyperpnea (60 L/min)	After exposure
Gong et al. (1995)	Asthma; n = 14; 12 M, 2 F; (27 ± 11 yr)	0, 0.5, or 1.0 ppm SO ₂ with light, medium, and heavy exercise (avg ventilation 30, 36, and 43 L/min) for 10 min	Before, during, and immediately after exposure
Gong et al. (1996)	Asthma; n = 10; 2 M, 8 F; (30.3 ± 9.2 yr)	0 or 0.75 ppm SO ₂ with exercise (29 L/min) for up to 24 h with or without pretreatment with salmeterol (long-acting β ₂ -agonist)	Before and immediately after exposure
Gong et al. (2001)	Asthma; n = 11; 2 M, 9 F; (30.8 ± 11.3 yr)	0 or 0.75 ppm SO ₂ for 10 min with exercise (35 L/min) with or without pretreatment to montelukast sodium (10 mg/day for 3 days)	Before, immediately after, and 1 and 2 h after exposure
Horstman et al. (1988)	Asthma; n = 12 M; (28.6 ± 5.5 yr)	0 or 1.0 ppm SO ₂ for 0, 0.5, 1.0, 2.0, and 5.0 min with exercise (treadmill, 40 L/min)	Before and immediately after exposure
Kehrl et al. (1987)	Asthma; n = 10 M; (26.8 ± 4.4 yr)	0 or 1 ppm SO ₂ for 1 h with exercise (3 × 10 min, 41 L/min, treadmill)	Before and during exposure/exercise
Koenig et al. (1980)	Asthma; n = 9; 7 M, 2 F; (15.7 ± 1.1 yr)	0 or 1 ppm SO ₂ with 1 mg/m ³ of NaCl droplet aerosol, 1 mg/m ³ NaCl droplet aerosol for 60-min exposure with mouthpiece at rest	Before, during, and immediately after exposure
Koenig et al. (1981)	Asthma; n = 8; 6 M, 2 F; (14–18 yr)	0 or 1 ppm SO ₂ with 1 mg/m ³ of NaCl droplet aerosol, 1 mg/m ³ NaCl droplet aerosol for 30-min exposure via mouthpiece at rest followed by 10 min exercise on a treadmill (sixfold increase in V _E)	Before, during, and immediately after exposure

Table 5-7 (Continued): Study specific details from controlled human exposure studies of respiratory symptoms.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Time of Symptom Assessment
Koenig et al. (1983)	Phase 1: Asthma with EIB; n = 9; 6 M, 3 F; (12–16 yr) Phase 2: Asthma with EIB; n = 7 (sex NR); (12–16 yr)	Phase 1: 1 g/m ³ of NaCl droplet aerosol, 1 ppm SO ₂ , 1 mg/m ³ NaCl, 0.5 ppm SO ₂ + 1 mg/m ³ NaCl for 30-min exposure via mouthpiece at rest followed by 10 min exercise on treadmill (five- to sixfold increase in V _E) Phase 2: 0.5 ppm SO ₂ + 1 mg/m ³ NaCl via a face mask with no nose clip with exercise conditions the same as Koenig et al. (1981)	Before and immediately after exposure
Koenig et al. (1987)	Allergy with EIB; n = 10; 3 M, 7 F; (13–17 yr)	0 or 0.75 ppm SO ₂ (mouthpiece) with exercise (33.7 L/min) for 10 min and 20 min prior pretreatment (0 or 180 µg albuterol)	Before and immediately after pretreatment and exposure
Koenig et al. (1990)	Asthma with EIB; n = 13; 8 M, 5 F (14.3 ± 1.8 yr)	0.1 ppm SO ₂ for 15 min preceded by air or 0.12 ppm O ₃ for 45 min during intermittent exercise (2 × 15 min, 30 L/min, treadmill), no control, air exposure	Before and immediately after exposure
Koenig et al. (1992)	Asthma; n = 8; 2 M, 6 F; (27.5 ± 9.6 yr)	1 ppm SO ₂ for 10 min with exercise (13.4–31.3 L/min) with or without pretreatment to theophylline	Before and immediately after exposure
Linn et al. (1983b)	Asthma; n = 23; 13 M, 10 F; (23.3 ± 4.4 yr)	0, 0.2, 0.4, or 0.6 ppm SO ₂ with low humidity or high humidity for 10 min with exercise (bicycle, 5 min 50 L/min) 0 or 0.6 ppm SO ₂ with warm air or cold air with exercise (bicycle, 50 L/min, ~5 min)	Before and immediately after exposure
Linn et al. (1983a)	Asthma; n = 23; 15 M, 8 F (23 ± 4 yr)	0 or 0.75 ppm SO ₂ with unencumbered breathing and mouth only breathing with exercise (40 L/m, 10 min, bicycle)	Before and immediately after exposure
Linn et al. (1984a)	Asthma; n = 14; 12 M, 2 F (24.1 ± 4.7 yr)	0, 0.3, or 0.6 ppm SO ₂ at 21°C, 7°C, and –6°C, RH 80% with exercise (bicycle, 50 L/min, ~5 min)	Before, during, immediately after, and a week after exposure
Linn et al. (1984c)	Asthma; n = 24; 13 M, 11 F; (24.0 ± 4.3 yr)	0, 0.3, or 0.6 ppm SO ₂ at 21°C, 7°C, and –6°C and 80% RH with exercise (5 min, 50 L/min)	Before, immediately after, and 24 h after exposure
Linn et al. (1984b)	Asthma; Phase 1 (Pilot) n = 8; 4 M, 4 F; (24.5 ± 3.9 yr) Phase 2 n = 24; 19 M, 5 F; (24.0 ± 4.3 yr)	Phase 1: 0, 0.2, 0.4, or 0.6 ppm SO ₂ at 5°C, 50, and 85% RH with exercise (5 min, 50 L/min) Phase 2: 0 and 0.6 ppm SO ₂ at 5°C and 22°C, 85% RH with exercise (5 min, 50 L/min)	Phase 1: before and immediately after exposure Phase 2: before, immediately after, 1 day after, and 1 wk after exposure

Table 5-7 (Continued): Study specific details from controlled human exposure studies of respiratory symptoms.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Time of Symptom Assessment
Linn et al. (1985b)	Asthma; n = 22; 13 M, 9 F; (23.5 ± 4.0 yr)	0 or 0.6 ppm SO ₂ at 21 and 38°C, 20 and 80% RH with exercise (~5 min, 50 L/min)	Before, immediately after, and 24 h after exposure
Linn et al. (1985a)	Asthma with COPD; n = 24; 15 M, 9 F; (60 yr; Range: 49–68 yr)	0, 0.4, or 0.8 ppm SO ₂ for 1 h with exercise (2 × 15 min, bicycle, 18 L/min)	Before, during, immediately after, 24 h after, and 7 days after exposure
Linn et al. (1987)	Healthy; n = 24; 15 M, 9 F; (18–37 yr) Atopic (sensitive to common airborne allergens but no asthma); n = 21; 12 M, 9 F; (18–35 yr) Minimal or mild asthma; n = 16; 10 M, 6 F; (20–33 yr) Moderate or severe asthma; n = 24; 10 M, 14 F; (18–35 yr)	0, 0.2, 0.4, or 0.6 ppm SO ₂ for 1 h with exercise (3 × 10-min, bicycle, ~40 L/min)	Before and during exposure (after first exercise and after last exercise)
Linn et al. (1988)	Asthma; n = 20; 13 M, 7 F; (28 ± 5 yr)	Three pretreatment groups (1) metaproterenol sulfate, (2) placebo, (3) no treatment 0, 0.3, and 0.6 ppm SO ₂ for 10 min with exercise (bike, 50 L/min)	Before, immediately after, 10 min, 30 min, 60 min, 120 min, 24 h, and 1 wk after exposure
Linn et al. (1990)	Asthma; n = 21; 6 M, 15 F; (34.8 ± 8.9 yr)	0, 0.3, or 0.6 ppm SO ₂ 10 min with exercise (50 L/min) (1) low medication use, (2) normal, (3) high usual medication supplemented by inhaled metaproterenol before exposure	Before exposure, after pretreatment, immediately after, 30 min after, and 60 min after exposure
Magnussen et al. (1990)	Asthma; n = 46; 21 M, 25 F; (28 ± 14 yr)	0 or 0.5 ppm SO ₂ for 20 min. 10-min rest followed by 10 min isocapnic hyperventilation (30 L/min)	Before exposure and immediately after hyperventilation
Myers et al. (1986a)	Asthma; n = 10; 7 M, 3 F; (27.6 ± 5.5 yr)	Three pretreatment groups (1) 200 mg cromolyn, (2) 20 mg cromolyn, (3) placebo Doubling concentrations of SO ₂ during sequential 3 min exposures, from 0.25 to 8 ppm	Before and after each 3-min exposure to an increasing SO ₂ concentration

Table 5-7 (Continued): Study specific details from controlled human exposure studies of respiratory symptoms.

Study	Disease Status; n; Sex; (Age ^a)	Exposure Details (Concentration; Duration)	Time of Symptom Assessment
Sheppard et al. (1983)	Asthma; n = 8; 4 M, 4 F; (26.6 ± 4.3 yr)	0.5 ppm SO ₂ for 3 min eucapnic hyperpnea	Before and immediately after exposure
Trenga et al. (1999)	Asthma; n = 47; 14 M, 33 F; (21.1 yr; Range: 18–39 yr)	0.5 ppm SO ₂ for 10 min with moderate exercise	Before and immediately after exposure
Trenga et al. (2001)	Asthma; n = 17; 5 M, 12 F; (27.4 ± 6.3 yr)	0.5 ppm SO ₂ for 10 min with moderate exercise (treadmill)	Before and immediately after exposure

COPD = chronic obstructive pulmonary disease; EIB = exercise-induced bronchospasm; F = female; M = male; n = sample size; NaCl = sodium chloride; NR = not reported; O₃ = ozone; ppm = parts per million; RH = relative humidity; SD = standard deviation; SO₂ = sulfur dioxide; V_E = minute volume.

^aRange or Mean ± SD.

[Linn et al. \(1983b\)](#) reported the severity of respiratory symptoms following 5-minute exposures to 0, 0.2, 0.4, and 0.6 ppm SO₂ in heavily exercising individuals with mild to moderate asthma. Total symptom score changes were significant ($0.01 < p < 0.05$, one-tailed Dunnett comparison) after 0.2 ppm SO₂ exposure and there was a statistically significant effect of increasing SO₂ concentration on symptoms. Considering individual symptoms, there was also a significant effect of increasing SO₂ concentration on the symptoms of cough, substernal irritation, wheeze, and chest tightness. Subsequently, a similar study with a slightly lower level of exercise demonstrated that 43% of subjects with asthma experienced increases in respiratory symptoms after three 10-minute exposures ([Linn et al., 1987](#)) to 0.6 ppm SO₂, with a smaller percentage experiencing respiratory symptoms at 0.4 ppm SO₂ ([Smith, 1993](#); [Linn et al., 1987](#)).

Additional studies examining concentrations of ≥ 0.5 ppm SO₂ demonstrated SO₂-induced increases in respiratory symptoms. Total and lower respiratory symptom scores were significantly increased with increasing SO₂ concentrations (0, 0.5, and 1.0 ppm SO₂) following 10-minute exposures with varying levels of exercise ([Gong et al., 1995](#)). [Trenga et al. \(1999\)](#) confirmed these results, observing a significant correlation between FEV₁ decrements and increases in respiratory symptoms following 10-minute exposures to 0.5 ppm SO₂ via mouthpiece. Respiratory symptoms have also been observed following exposure durations as low as 3 minutes to 0.5 ppm SO₂ via mouthpiece during eucapnic hyperpnea (V_E = 60 L/minute), in which seven out of eight individuals with asthma developed respiratory symptoms ([Balmes et al., 1987](#)).

As with lung function, increased respiratory symptoms in response to short-term exposure to SO₂ in individuals with asthma is dependent on exercise. [Linn et al. \(1983b\)](#) reported significant changes in total symptom scores after 0.2 ppm SO₂ exposure in heavily exercising individuals with asthma. In contrast, in adults with asthma at rest, no association was found between respiratory symptoms (i.e., throat irritation, cough, wheeze) and 1-hour exposures to 0.2 ppm SO₂ ([Tunncliffe et al., 2003](#)).

Epidemiologic Studies

Compared with controlled human exposure studies, epidemiologic evidence for SO₂-associated increases in symptoms is variable, being supportive in children with asthma but weak in adults with asthma. A recent study of children and adults combined does not support an association with asthma medication use. The analysis, which only reported the lack of statistically significant associations, was limited by analysis of beta-agonist levels in wastewater rather than medication use ascertained for individual subjects ([Fattore et al., 2016](#)) ([Table 5-8](#)). The evidence base specifically in children with asthma is larger and more informative, providing results for home and/or school SO₂ exposure estimates and temporally resolved SO₂ metrics. Also, while they do not settle questions, studies in children with asthma aim to assess copollutant confounding and interactions. Although the evidence overall is less consistent in recent than previous studies, the aforementioned strengths are features of many recent studies of children with asthma.

Adults. Ambient air SO₂ concentrations were lower in recent than previous studies (0.87–2.7 ppb vs. 1.6–90 ppb for means), but this reduction does not appear to explain the weak evidence because previous results are also inconsistent [[Supplemental Figure 5S-1 and Table 5S-3 \(U.S. EPA, 2017c\)](#)]. All studies have uncertainty in the SO₂ exposure estimates assigned from a single monitor or averaged across multiple monitors. No study indicated whether measurements at the monitors adequately represented the spatiotemporal variability in ambient SO₂ concentrations in the study area or the temporal variation in people's exposures.

Table 5-8 Epidemiologic studies of respiratory symptoms in populations with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
Adults with asthma				
†Maestrelli et al. (2011) Padua, Italy, 2004–2005 N = 32, mean (SD) age 40 (7.5 yr). 81% persistent asthma. 69% ICS use. 90% atopy. Six measures over 2 yr. Symptoms assessed in clinic. Recruited from database of beta-agonist users (>6 times per year for 3 yr).	Two monitors in city Medians across seasons: 0.87–2.7 75th percentiles across seasons: 1.3–4.1	24-h avg 0	Asthma control score Increase = better control All subjects: 0.77 (–1.1, 2.6) Nonsmokers: 0.10 (–2.2, 2.4) n = 22	No copollutant model No association with personal or fixed-site PM _{2.5} Copollutant correlations NR
		2	Daytime symptoms OR: 0.90 (0.81, 0.99)	Not statistically significant with: NO ₂ Quantitative results NR. PM _{2.5} not examined. <i>r</i> = 0.23 NO ₂ and PM ₁₀
		5	Nighttime symptoms OR: 1.16 (1.04, 1.29)	
†Wiwatanadate and Liwsrisakun (2011) Chiang Mai, Thailand, 2005–2006 N = 121, ages 13–78 yr. 48% moderate/severe persistent asthma. Daily diary for 10 mo. Recruited from allergy clinics. Multiple comparisons—many pollutants, lags, health endpoints analyzed.	Monitor within 10 km of home Mean (SD): 1.7 (0.62) 90th percentile: 2.4 Max: 3.9	24-h avg	SO ₂ increment NR. Results reported only for statistically significant lags	
		2	Daytime symptoms OR: 0.90 (0.81, 0.99)	
		5	Nighttime symptoms OR: 1.16 (1.04, 1.29)	
		0–2 avg	2.53 (1.05, 6.08)	
†Anyenda et al. (2016) Kanazawa, Japan, Jan–June 2011 N = 83, ages 23–84 yr. 54% atopy. Daily diary for mean 153 days. Recruited from hospital outpatients.	One monitor in city Mean (SD): 1.6 (1.3) Max: 7.3	24-h avg	Cough	Persists with: PAH or NO ₂ (lag 2) Spearman <i>r</i> = 0.60 PAH, 0.56 NO ₂
		0	0.67 (0.34, 1.31)	
		2	2.19 (1.34, 3.54)	
		0–2 avg	2.53 (1.05, 6.08)	

Table 5-8 (Continued): Epidemiologic studies of respiratory symptoms in populations with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
Children with asthma				
<p>†Spira-Cohen et al. (2011), Spira-Cohen (2013) Bronx, NY, 2002–2005 N = 40, ages 10–12 yr. 44% with asthma ED visit or hospital admission in previous 12 mo. Daily diaries for 1 mo. Recruited from schools by referrals from school nurses.</p>	<p>Monitor at school Concentrations NR Most children walk to school</p>	<p>1-h max (a.m.) 0</p>	<p>Cough RR: 1.60 (1.20, 2.12) Wheeze RR: 1.81 (1.15, 2.84) Shortness of breath RR: 1.45 (0.90, 2.84)</p>	<p>Cough Does not persist with: school EC PM_{2.5} not associated <i>r</i> = 0.45 EC</p>
<p>†Velická et al. (2015) Ostrava, Czech Republic, Nov 2013–Feb 2014 N = 147, ages 6–18 yr. 67% mild persistent asthma. 33% moderate persistent asthma. 79% atopy. 97% regular asthma medication use. Daily diaries for 4 mo. Recruited from clinics.</p>	<p>Five monitors and dispersion model 0.5 × 0.5 km resolution Weighted avg by time at home and school Median: 4.0 75th percentile: 12</p>	<p>24-h avg 0</p>	<p>Cough OR: 0.92 (0.74, 1.17) Breathing difficulty-wheeze OR: 2.29 (1.55, 3.39) Reliever inhaler use OR: 1.84 (1.32, 2.56) Restricted activities OR: 1.25 (1.00, 1.62)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR</p>
<p>†Dales et al. (2009) Windsor, ON, Oct–Dec 2005 N = 182, ages 9–14 yr. 37% ICS use. 35% beta-agonist use. Daily diaries for 4 wk. Recruited from schools. Mean 1.6 and 2.2 h/day spent outdoors.</p>	<p>Two monitors avg 99% homes within 10 km of sites Median: 4.5 95th percentile: 16</p>	<p>24-h avg</p>	<p>OR for SO₂ ≥8.8 vs. <2.3 ppb Chest tightness 1.30 (1.06, 1.58) ORs for difficulty breathing, cough, and wheeze reported not statistically significant</p>	<p>No copollutant model Quantitative results NR</p>
<p>†O'Connor et al. (2008) Inner-City Asthma Study cohort: Boston, MA; Bronx, NY; New York City, NY; Chicago, IL; Dallas, TX; Tucson, AZ; Seattle, WA; 1998–2001 N = 861, ages 5–12 yr. 100% persistent asthma. 100% atopy. Daily diaries for four 2-wk periods. Recruited from intervention study.</p>	<p>Monitors avg close to home and not near industry Median 2.3 km to site Quantitative SO₂ data NR</p>	<p>24-h avg 1–19 avg</p>	<p>Wheeze-cough RR: 1.05 (0.89, 1.23) Nighttime asthma RR: 1.11 (0.91, 1.36) Slow play RR: 1.06 (0.88, 1.27) Missed school RR: 1.10 (0.82, 1.49)</p>	<p>No copollutant model <i>r</i> = 0.59 NO₂, 0.32 CO, 0.37 PM_{2.5}</p>

Table 5-8 (Continued): Epidemiologic studies of respiratory symptoms in populations with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Gent et al. (2009) New Haven county, CT, 2000–2004 N = 149, ages 4–12 yr. 45% intermittent asthma. Daily diaries reported monthly for 1 yr. Recruited from larger cohort, clinic, and school.</p>	<p>Monitor 0.9–30 km of home Mean 10 km to site Concentrations NR</p>	<p>24-h avg 0</p>	<p>NR</p>	<p>Only multipollutant model analyzed with six PM_{2.5} component factors r = 0.45 motor vehicle factor</p>
Children and adults with asthma				
<p>†Fattore et al. (2016) Milan, Italy, Sep–Dec 2013 N = 84 days Daily wastewater samples for 84 days analyzed for levels of the beta-agonist salbutamol.</p>	<p>Three monitors avg Mean (SD): 2.2 (1.3) Max: 5.9</p>	<p>24-h avg 0 to 10 (Single-day)</p>	<p>Beta-agonist levels in wastewater No quantitative results. RRs reported not statistically significant</p>	<p>No copollutant model Pearson r = 0.66 PM_{2.5}, 0.65 PM₁₀</p>

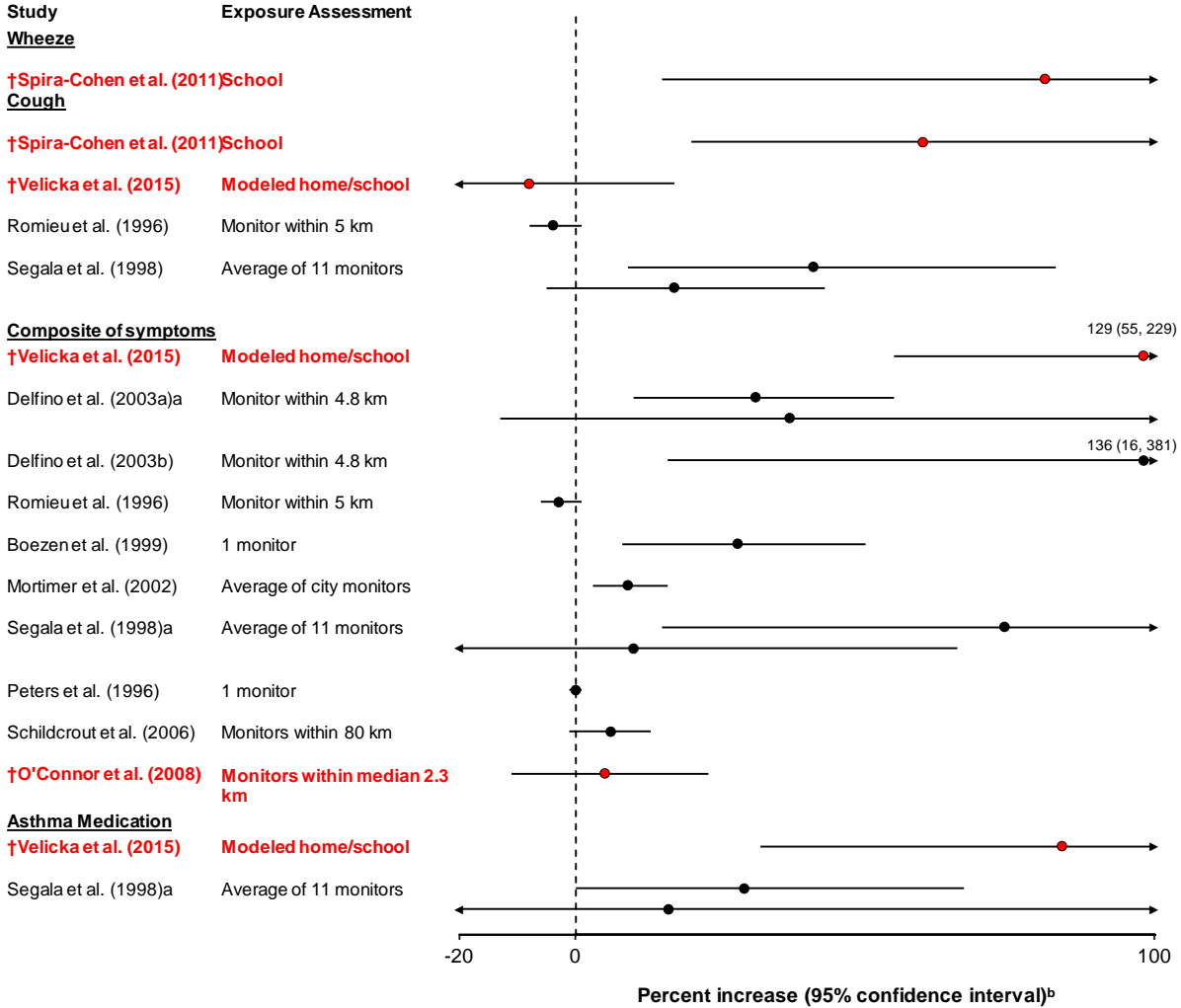
Avg = average; CI = confidence interval; CO = carbon monoxide; EC = elemental carbon; ED = emergency department; ICS = inhaled corticosteroids; ISA = Integrated Science Assessment; max = maximum; n = sample size; N = population number; NO₂ = nitrogen dioxide; NR = not reported; OR = odds ratio; PAH = polycyclic aromatic hydrocarbon; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; r = correlation coefficient; RR = relative risk; SD = standard deviation; SO₂ = sulfur dioxide.

^aEffect estimates are standardized to a 10-ppb increase in 24-h avg SO₂ and 40-ppb increase in 1-h max SO₂.

†Studies published since the 2008 ISA for Sulfur Oxides.

All epidemiologic studies of adults examined 24-h avg SO₂ concentrations (Table 5-8), longer than the 5–10-minute exposures used in controlled human exposure studies (Table 5-2). As in previous studies, recent epidemiologic evidence does not indicate associations for respiratory symptoms with same-day (lag 0) SO₂ concentrations (Anyenda et al., 2016; Maestrelli et al., 2011). Atopy was prevalent in Maestrelli et al. (2011) (90%); previous findings supported an association in adults with AHR and elevated IgE (Boezen et al., 2005). A recent study linked an increase in SO₂ concentration to an increase in nighttime asthma symptoms with a 5-day lag (Wiwatanadate and Liwsrisakun, 2011), but inference is weak because results were inconsistent among the many lags, pollutants, and health effects examined. Also, SO₂ exposures were assessed from a monitor up to 10 km from subjects' homes. There is some consistency for SO₂ concentrations lagged 2 or 5 days, or averaged over 3 or 5 days, [Supplemental Figure 5S-1 and Table 5S3 (U.S. EPA, 2017c) and Anyenda et al. (2016)]. PM metrics also were associated with symptoms and moderately to highly correlated with SO₂ ($r = 0.60\text{--}0.9$) (Boezen et al., 2005; Neukirch et al., 1998; Peters et al., 1996a). Whether the magnitude of copollutant correlations influences the consistency of association for SO₂ with respiratory symptoms in adults with asthma cannot be determined in this small evidence base. As examined only in a recent study, SO₂ associations persisted with adjustment for polycyclic aromatic hydrocarbon (PAH) or NO₂ (Anyenda et al., 2016). However, uncertainty in the exposures estimated from a single monitor and a different site for PAH limits the inferences that can be drawn about an independent association for SO₂. Controlled human exposure studies show that symptoms resolve once exposure ends, but SO₂-induced allergic inflammation could be a pathway by which SO₂ exposure induces symptoms after several days or over multiple days.

Children. Overall, the epidemiologic evidence indicates associations between higher SO₂ concentrations and increased respiratory symptoms in children with asthma, particularly when effects are examined as a composite index of multiple symptoms (Figure 5-2). Associations also are observed for asthma medication use or activity restriction but not consistently for wheeze or cough. Results vary in magnitude and precision (Figure 5-2). In some study areas, the SO₂ concentrations were much lower (Spira-Cohen et al., 2011; Delfino et al., 2003a; Delfino et al., 2003b) or higher (Mortimer et al., 2002) than the 10-ppb increment used to standardize the effect estimates. Although recent studies give inconsistent results (Table 5-8), associations are observed with SO₂ exposure estimates that are measured or modeled for the school or home. Recent studies reported lower SO₂ concentrations than many previous studies [for 24-h avg, median ~4 ppb vs. means 8.3 and 90 ppb in (Segala et al., 1998; Romieu et al., 1996)]. It is unclear whether the inconsistency is due to lower concentrations; previous studies observed associations in locations with similar SO₂ concentrations [median 24-h avg 2.2–7.4 ppb in Schildcrout et al. (2006), mean 8-h max 4.6 ppb in Delfino et al. (2003a), Delfino et al. (2003b)].



^aThe two results for [Delfino et al. \(2003a\)](#) refer to symptoms not interfering with activity and symptoms interfering with activity. The two results for [Segala et al. \(1998\)](#) refer to children with mild asthma and children with moderate asthma.

^bEffect estimates are standardized to a 10-ppb increase in 24-h avg sulfur dioxide concentration and a 40-ppb increase in 1-h max concentrations. Study details are presented in [Table 5-8](#). Results from [Gent et al. \(2009\)](#) are not presented in the figure because they are based on a multipollutant model. Corresponding quantitative results are reported in Supplemental Table 5S-4 ([U.S. EPA, 2017c](#)).

Note: † and red = recent studies published since the 2008 ISA for Sulfur Oxides, black = studies from the 2008 ISA for Sulfur Oxides.

Figure 5-2 Associations between short-term average ambient sulfur dioxide concentrations and respiratory symptoms and asthma medication use in children with asthma.

[Spira-Cohen et al. \(2011\)](#) is notable not only for monitoring SO₂ at schools but also for examining 1-h max concentrations. In the population of children in Bronx, NY, increases in SO₂ were linked to increased odds of cough and wheeze but not shortness of breath ([Table 5-8](#)). Previous U.S. studies also associated symptoms with temporally resolved SO₂ metrics [i.e., 1-h max, 8-h max, 3-h avg (8–11 a.m.)] but had more uncertainty in

exposures estimated from monitors up to 4.8 km from children's homes/schools ([Delfino et al., 2003a](#); [Delfino et al., 2003b](#)) or monitors averaged across the city ([Mortimer et al., 2002](#)). [Spira-Cohen et al. \(2011\)](#) did not report SO₂ concentrations to compare to previous studies but reported that most children walked to school. [Velická et al. \(2015\)](#) also aimed to improve exposure assessment for children in Ostrava, Czech Republic. A dispersion model and five monitors were used to estimate SO₂ concentrations at 0.5-km resolution and calculate a time-weighted 24-h avg for each child based on the school and home location. SO₂ was associated with breathing difficulty-wheeze, reliever inhaler use, and restricted activities, but not cough ([Table 5-8](#)). The study population had a high prevalence of atopy (79%); thus, results agree with [Segala et al. \(1998\)](#). However, compared to the previous studies, [Velická et al. \(2015\)](#) may have less uncertainty in exposure estimates ([Section 3.5](#)).

Other recent studies largely do not provide evidence for SO₂-associated increases in respiratory symptoms in children with asthma ([Dales et al., 2009](#); [Gent et al., 2009](#); [O'Connor et al., 2008](#)). But, they are limited because of (1) the large distance between the SO₂ monitor and children's homes (e.g., up to 10 km, median 2.3 km, mean 10 km); (2) a lack of quantitative results ([Dales et al., 2009](#)); (3) use of 19-day avg SO₂ concentrations, which are more subject to residual temporal confounding ([O'Connor et al., 2008](#)); or (4) use of SO₂ only as part of a multipollutant model with six PM_{2.5} component source factors ([Gent et al., 2009](#)).

For the associations observed between SO₂ and respiratory symptoms in children with asthma, including those with atopy, the influence of copollutants is poorly addressed. Symptoms were not associated with personal or school PM_{2.5} but with other PM metrics, including PM₁₀, elemental carbon (EC), OC, black smoke (BS), and TSP. Associations were also observed with NO₂, VOCs such as benzene and xylene, and O₃ ([Table 5-8](#)). Except for O₃, these copollutants were moderately to highly correlated with SO₂ ($r = 0.45-0.9$). Correlations were highest in previous studies, but recent studies did not report SO₂ concentrations ([Spira-Cohen et al., 2011](#)) or copollutant correlations ([Velická et al., 2015](#)) to assess whether the magnitude of correlation varied by SO₂ levels. Copollutant models were analyzed in few studies and for few copollutants. For a Los Angeles, CA cohort, no SO₂-VOC interaction was indicated, and SO₂ associations persisted with adjustment for benzene, xylene, or toluene for some but not all symptoms ([Delfino et al., 2003a](#); [Delfino et al., 2003b](#)). Associations for VOCs were attenuated as well, and copollutant model results were uncertain because of the moderate to high correlations with SO₂ ($r = 0.58-0.78$) and because exposures were assessed from monitors 4.8 km from children's homes or schools. Potential exposure error also limits inference from results showing associations for joint increases in SO₂ with PM₁₀, NO₂, or carbon monoxide (CO) that were similar to each single-pollutant association ([Schildcrout](#)

[et al., 2006](#)). The recent Bronx, NY study analyzed copollutant models for school SO₂ and EC, which may have more comparable exposure error. SO₂ and EC were moderately correlated ($r = 0.45$), consistent with the location in a high diesel truck traffic area (data were obtained for 2002–2005, before the Diesel Fuel Standard went into effect in 2006, see [Section 2.2.3](#)) ([Spira-Cohen et al., 2011](#)). In the copollutant model, the odds ratio (OR) for cough was robust for EC but decreased in magnitude and precision for SO₂ from 1.60 (95% CI: 1.20, 2.12) to 1.32 (95% CI: 0.93, 1.87) per 40-ppb increase in 1-h max SO₂.

Summary of Respiratory Symptoms in Populations with Asthma

Controlled human exposure studies provide strong evidence for the effects of SO₂ exposure on respiratory symptoms in adults with asthma under increased ventilation conditions. Exposures for 5–10 minutes to 0.2–0.6 ppm SO₂ induced respiratory symptoms in exercising individuals with asthma, with the most consistent evidence from exposures to 0.4–0.6 ppm SO₂ ([Table 5-2](#)). Epidemiologic evidence in adults with asthma is weak, but increases in ambient SO₂ concentration are generally associated with increased risk of asthma symptoms in children ([Figure 5-2](#); [Table 5-8](#)). Assessing coherence specifically with controlled human exposure studies of adolescents with asthma is difficult because those studies lacked an appropriate control exposure. Limited findings support associations in children and adults with AHR and elevated IgE.

Epidemiologic results in children are less consistent in recent than previous studies but support associations for 1-h max SO₂ measured at schools or 24-h avg SO₂ modeled for school and home. School or home SO₂ measures may better represent exposures than the concentrations at fixed-site monitors examined in most studies, particularly for 1-h max. These SO₂ metrics are longer than the 5–10 minute SO₂ exposures in controlled human exposure studies, which show transient responses. And, the extent to which confounding or an interaction with copollutants such as PM_{2.5}, EC, NO₂, and VOCs contributed to epidemiologic associations, including those for populations with asthma and a high prevalence of atopy or AHR and elevated IgE, and for residents near a coal-fired power plant, is not fully characterized in the epidemiologic studies. However, an independent effect of SO₂ exposure is indicated by experimental evidence in rodents of allergic inflammation enhanced by repeated 1-hour exposures to 2 ppm SO₂.

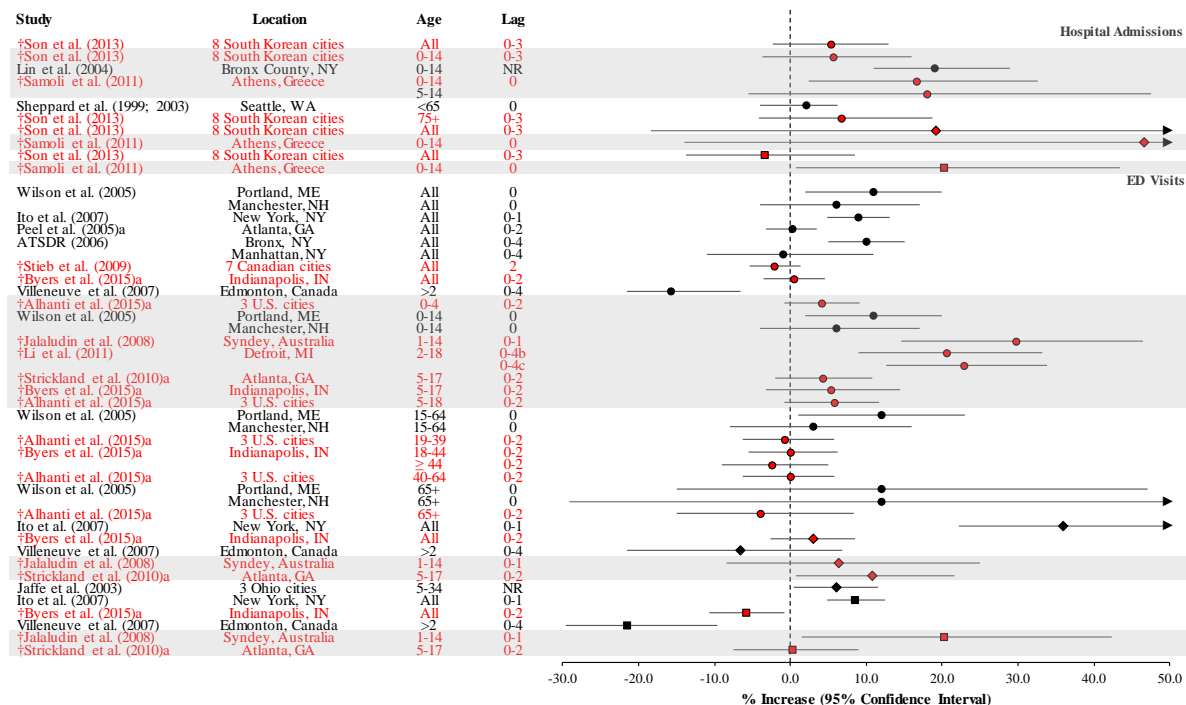
Hospital Admission and Emergency Department Visits for Asthma

Since the completion of the 2008 SO_x ISA, epidemiologic studies have continued to examine the association between short-term exposure to ambient SO₂ concentrations and respiratory-related hospital admissions and ED visits, but are primarily limited to

single-city studies. The sections within this chapter detailing the respiratory-related hospital admissions and ED visits studies characterize recent studies in the context of the collective body of evidence evaluated in the 2008 SO_x ISA. The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) included the first thorough evaluation of respiratory morbidity in the form of respiratory-related hospital admissions and ED visits, including asthma. These studies reported generally positive associations with short-term SO₂ exposures, with associations that are often larger in magnitude for children ([Figure 5-3](#)). Additionally, SO₂ associations with asthma hospital admissions and ED visits were often attenuated, but remained positive in copollutant models with PM, NO₂, or O₃.

Within this section focusing on asthma, as well as the rest of the chapter, respiratory-related hospital admissions and ED visit studies are evaluated separately because only a small percentage of respiratory-related ED visits result in hospital admission. Additionally, when evaluating asthma ED visit and hospital admission studies that focus on children (i.e., defined age ranges <18 years of age), it is important to note the difficulty of reliably diagnosing asthma in children <5 years of age. Thus, including children under the age of 5 years in a study population may result in the overestimation of the number of asthma ED visit and hospital admissions ([NAEPP, 2007](#)).

For each of the studies evaluated in this section, [Table 5-9](#) presents the air quality characteristics of each city, or across all cities, the exposure assignment approach used, and information on copollutants examined in each asthma hospital admission and ED visit study. Other recent studies of asthma hospital admissions and ED visits are not the focus of this evaluation because these studies were conducted in small single-cities, encompassed a short study duration, had insufficient sample size, or did not examine potential copollutant confounding. The full list of these studies, as well as study-specific details, can be found in Supplemental Table 5S-6 ([U.S. EPA, 2017c](#)).



ED = emergency department; ISA = Integrated Science Assessment.

^a = studies that used a 1-h max exposure metric.

^b = time-series results.

^c = case-crossover results.

Note: † and red text/symbols = recent studies published since the 2008 ISA for Sulfur Oxides. Black text/symbols = U.S. and Canadian studies evaluated in the 2008 ISA for Sulfur Oxides; Circle = all-year; diamond = warm/summer months; square = cold/winter months. Gray shading depicts studies that present results for children (i.e., <18 yr of age). Corresponding quantitative results are reported in Supplemental Table 5S-5 ([U.S. EPA, 2017c](#)).

Figure 5-3 Percent increase in asthma hospital admissions and emergency department visits from U.S. and Canadian studies evaluated in the 2008 SO_x ISA and recent studies in all-year and seasonal analyses for a 10-ppb increase in 24-h avg or 40-ppb increase in 1-h max sulfur dioxide concentrations.

Table 5-9 Study-specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
Hospital admissions						
Lin et al. (2004)	Bronx County, NY (1991–1993)	Avg of SO ₂ concentrations from two monitoring sites	24-h avg	Cases: 16.8 Controls: 15.6	NR	NR
Sheppard et al. (1999) , Sheppard (2003)	Seattle, WA (1987–1994)	Avg of SO ₂ concentrations from multiple monitors	24-h avg	8.0	75th: 10.0 90th: 13.0	Correlation (<i>r</i>): PM ₁₀ : 0.31 PM _{2.5} : 0.22 PM _{10–2.5} : 0.34 O ₃ : 0.07 CO: 0.24 Copollutant models: none
†Son et al. (2013)	Eight South Korean cities (2003–2008)	Avg of hourly ambient SO ₂ concentrations from monitors in each city	24-h avg	3.2–7.3	NR	Correlation (<i>r</i>): PM ₁₀ : 0.5 O ₃ : –0.1 NO ₂ : 0.6 CO: 0.6 Copollutant models: none
†Zheng et al. (2015)	Meta-analysis (1988–2014)	NR	24-h avg	3.1–45.5 ^a	NR	Correlations (<i>r</i>): NR Copollutant models: none
†Samoli et al. (2011)	Athens, Greece (2001–2004)	Avg of SO ₂ concentrations across multiple monitors	24-h avg	6.4	75th: 8.4	Correlation (<i>r</i>): O ₃ : –0.19 NO ₂ : 0.55 Copollutant models: PM ₁₀ , SO ₂ , NO ₂ , O ₃

Table 5-9 (Continued): Study specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
ED visits						
Jaffe et al. (2003)	Cincinnati, Cleveland, and Columbus, OH (1991–1996)	When more than one monitoring station operating in a day, monitor reporting highest 24-h avg SO ₂ concentration used	24-h avg	Cincinnati: 13.7 Cleveland: 15.0 Columbus: 4.2	Max: Cincinnati: 50 Cleveland: 64 Columbus: 22	Correlations (<i>r</i>) (range across cities) NO ₂ : 0.07–0.28 O ₃ : 0.14–0.26 PM ₁₀ : 0.29–0.42 Copollutant models: none
Ito et al. (2007)	New York, NY (1999–2002)	Average SO ₂ concentrations across 19 monitors	24-h avg	7.8	75th: 10 95th: 17	Correlations (<i>r</i>): NR Copollutant models: PM _{2.5} , NO ₂ , O ₃ , CO
ATSDR (2006)	Bronx and Manhattan, NY (1999–2000)	SO ₂ concentrations from one monitor in Bronx and one in Manhattan Note: monitors used in series not simultaneously	24-h avg	Manhattan: 12 Bronx: 11	NR	Correlations (<i>r</i>): Bronx: O ₃ : -0.49 NO ₂ : 0.50 PM _{2.5} : 0.39 Max PM ₁₀ : 0.0.34 Manhattan: O ₃ : -0.40 NO ₂ : 0.47 PM _{2.5} : 0.26 PM ₁₀ : 0.24 Copollutant models: O ₃ , FRM and Max PM _{2.5} , NO ₂

Table 5-9 (Continued): Study specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
Peel et al. (2005)	Atlanta, GA (1993–2000)	Average of SO ₂ concentrations from monitors for several monitoring networks	1-h max	16.5	90th: 39.0	Correlations (<i>r</i>): PM _{2.5} : 0.17 PM ₁₀ : 0.20 PM _{10-2.5} : 0.21 UFP: 0.24 PM _{2.5} water soluble metals: 0.00 PM _{2.5} sulfate: 0.08 PM _{2.5} acidity: -0.03 PM _{2.5} OC: 0.18 PM _{2.5} EC: 0.20 Oxygenated HCs: 0.14 O ₃ : 0.19 CO: 0.26 NO ₂ : 0.34 Copollutant models: none
Wilson et al. (2005)	Portland, ME, and Manchester, NH (1996–2000)	SO ₂ concentrations from one monitor in each city	24-h avg	Portland: 11.1 Manchester: 16.5	NR	Correlation (<i>r</i>) (Range across cities): O ₃ : 0.05–0.24 Copollutant models: none
†Stieb et al. (2009)	Seven Canadian cities (1992–2003)	Average SO ₂ concentrations across all monitors in each city. Number of SO ₂ monitors in each city ranged from 1–11.	24-h avg	2.6–10.0	75th: 3.3–13.4	Correlations (<i>r</i>) only reported by city and season Copollutant models: none

Table 5-9 (Continued): Study specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
†Orazio et al. (2009)	Six Italian cities (1996–2002)	Average of SO ₂ concentrations across all monitors in each city	24-h avg	All-year: 2.1–8.1 Warm (Apr–Sep): 1.3–9.0 Cold (Oct–Mar): 2.6–7.3	NR	Correlations (<i>r</i>): NR Copollutant models: none
†Alhanti et al. (2016)	Three U.S. cities Atlanta, GA (1993–2009) Dallas, TX (2006–2009) St. Louis, MO (2001–2007)	Population-weighted average using data available from all monitors measuring SO ₂	1-h max	Atlanta: 10.7 Dallas: 2.7 St. Louis: 10.7	NR	Correlations (<i>r</i>): NR Copollutant models: none
†Zheng et al. (2015)	Meta-analysis (1988–2014)	NR	24-h avg	4.6–39.1 ^a	NR	Correlations (<i>r</i>): NR Copollutant models: none
†Strickland et al. (2010)	Atlanta, GA (1993–2004)	Population-weighted average using data available from all monitors measuring SO ₂	1-h max	All-year: 10.8 Warm (May–Oct): 9.6 Cold (Nov–Apr): 12.0	NR	Correlations (<i>r</i>): NR Copollutant models: none
†Li et al. (2011)	Detroit, MI (2004–2006)	Average of SO ₂ concentrations across two monitors in Detroit metropolitan area that measure SO ₂	24-h avg	3.8	75th: 5.1 Max: 27.3	Correlations (<i>r</i>), range across monitors: CO: 0.17–0.31 PM _{2.5} : 0.40–0.53 NO ₂ : 0.42–0.55 Copollutant models: none

Table 5-9 (Continued): Study specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
†Byers et al. (2015)	Indianapolis, IN (2007–2011)	Double-weighted average (distance from monitor to ZIP code centroid and age-specific census population) of two SO ₂ monitors	1-h max	All-year: 10.1 Warm: 10.5 Cold: 9.8	NR	Correlations (r): All-year: PM _{2.5} : 0.34 Warm: 1-h max O ₃ : 0.45 8-h max O ₃ : 0.42 PM _{2.5} : 0.38 Cold: PM _{2.5} : 0.29
†Villeneuve et al. (2007)	Edmonton, AB (1992–2002)	Average of SO ₂ concentrations across three monitoring stations	24-h avg	Summer (Apr–Sep) 50th: 2.0 Winter (Oct–Mar) 50th: 3.0	Summer 75th: 3.0 Winter 75th: 4.0	Correlations (r): NR Copollutant models: NR
†Jalaludin et al. (2008)	Sydney, Australia (1997–2001)	Average of SO ₂ concentrations across 14 monitoring stations	24-h avg	All-year: 1.07 Warm: 1.03 Cold: 1.1	Max All-year: 4.1 Warm: 4.1 Cold: 3.9	Correlations (r): (warm, cold) PM ₁₀ : 0.37, 0.46 PM _{2.5} : 0.27, 0.46 O ₃ : 0.45, -0.04 CO: 0.46, 0.51 NO ₂ : 0.52, 0.56 Copollutant models: PM ₁₀ , PM _{2.5} , O ₃ , CO, NO ₂

Table 5-9 (Continued): Study specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
†Smargiassi et al. (2009)	Montreal, QC (1996–2004)	SO ₂ concentrations measured at two monitoring sites east and southwest of the refinery At-home estimates of daily exposure by estimating SO ₂ concentrations at centroid of residential postal codes using AERMOD	24-h avg	Regional: 4.3 East: 6.9 Southwest: 4.4 AERMOD: East + Southwest: 3.0 East: 3.7 Southwest: 2.4	75th: Regional: 5.3 East: 9.2 Southwest: 5.9 AERMOD: East + Southwest: 4.3 East: 5.5 Southwest: 3.0	NR
†Brand et al. (2016)	Quebec and British Columbia, Canada	Average of SO ₂ concentrations from all monitors within 7.5 km from a major facility (i.e., refinery, smelter, pulp mill)	24-h avg 1-h max	24-h avg Quebec: Cases: 2.35 Controls: 2.40 British Columbia Cases: 2.04 Controls: 2.23 Total: Cases: 2.04 Controls: 2.23	NR	Correlations (<i>r</i>): NR Copollutant models: none

Table 5-9 (Continued): Study specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
†Winguist et al. (2014)	Atlanta, GA, U.S. (1998–2004)	Population-weighted average using data available from all monitors measuring SO ₂	1-h max	Warm (May–Oct): 8.3 Cold (Nov–April): 10.8	75th: Warm: 11.4 Cold: 14.6	Correlations (<i>r</i>): Warm: O ₃ : 0.27 CO: 0.32 NO ₂ : 0.44 PM _{2.5} : 0.28 EC: 0.31 Sulfate: 0.24 Secondary PM _{2.5} : 0.24 Cold: O ₃ : 0.05 CO: 0.22 NO ₂ : 0.41 PM _{2.5} : 0.07 EC: 0.18 Sulfate: 0.02 Secondary PM _{2.5} : 0.08 Copollutant models: none
†Pearce et al. (2015)	Atlanta, GA	SO ₂ concentrations from one monitor	1-h max	14.6	NR	Correlations (<i>r</i>): NR Copollutant models: none
Outpatient and physician visits						
†Burra et al. (2009)	Toronto, ON (1992–2001)	Average of SO ₂ concentrations across six monitors	1-h max	9.7	75th: 12.0 95th: 35.0 Max: 62.0	Correlations (<i>r</i>): NR Copollutant models: none

Table 5-9 (Continued): Study specific details and mean and upper percentile concentrations from asthma hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examination
† Sinclair et al. (2010)	Atlanta, GA, U.S. (1998–2002)	SO ₂ concentrations collected as part of AIRES at SEARCH Jefferson street site	1-h max	1998–2000: 19.3 2000–2002: 17.6 1998–2002: 18.3	NR	Correlations (<i>r</i>): NR Copollutant models: none

AERMOD = American Meteorological Society/U.S. EPA Regulatory Model; AIRES = Aerosol Research Inhalation Epidemiology Study; avg = average; CO = carbon monoxide; EC = elemental carbon; FRM = Federal Reference Method; HCs = hydrocarbons; ISA = Integrated Science Assessment; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OC = organic carbon; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; PM_{10-2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm and greater than 2.5 µm; *r* = Pearson correlation coefficient; SEARCH = Southeast Aerosol Research Characterization; SO₂ = sulfur dioxide; UFP = ultrafine particle.

^aRange of mean concentrations across all studies included in the meta-analysis.

† = studies published since the 2008 ISA for Sulfur Oxides.

Hospital Admissions

The 2008 SO_x ISA identified only two U.S.-based studies and no Canadian studies that examined the association between short-term SO₂ exposures and asthma hospital admissions. These studies reported positive associations; however, they were limited to studies of individual cities ([Figure 5-3](#)). The asthma hospital admission studies averaged SO₂ concentrations over multiple monitors and only examined 24-h avg exposure metrics, which may not adequately capture the spatial and temporal variability in SO₂ concentrations ([Sections 3.4.2.2](#) and [3.4.2.3](#)). While correlations between 24-h avg and 1-h max SO₂ concentrations are high ($r > 0.75$) at most monitors, lower correlations may occur at some monitors and in individual studies, adding uncertainty to the ability of 24-h avg metrics to capture peak SO₂ concentrations. Additionally, relatively few studies have examined the potential confounding effects of other pollutants on the SO₂-asthma hospital admissions relationship.

To date, a limited number of studies have been published since the 2008 SO_x ISA that focus on the relationship between short-term SO₂ exposures and asthma hospital admissions. In a time-series study conducted in Athens, Greece, [Samoli et al. \(2011\)](#) evaluated the association between multiple ambient air pollutants and pediatric asthma hospital admissions for ages 0–14 years. In an all-year analysis, the authors reported a positive association with SO₂ [16.5 % (95% CI: 2.3, 32.6); lag 0 increase for a 10-ppb

increase in 24-h avg SO₂ concentrations]. In copollutant analyses, the authors found SO₂ risk estimates to be robust in models with PM₁₀ [13.0% (95% CI: -1.5, 29.7)] and O₃ [16.5% (95% CI: 2.3, 32.6)]. However, models with NO₂ showed an increase in the SO₂ risk estimate [21.3% (95% CI: 1.1, 45.5)]. SO₂ was low ($r < 0.4$) to moderately (r ranging from 0.4–0.7) correlated with other pollutants examined in the study, with the highest correlation with NO₂ ($r = 0.55$).

The association between short-term SO₂ exposures and asthma hospital admissions was also examined by [Son et al. \(2013\)](#) in a study of eight South Korean cities. In addition to focusing on asthma, the authors examined allergic disease hospital admissions, which encompass asthma. For all ages, the authors reported a 5.3% increase (95% CI: -2.4, 13.0) in asthma hospital admissions for a 10-ppb increase in 24-h avg SO₂ concentrations and a 3.1% increase (95% CI: -3.7, 10.7) in allergic diseases hospital admissions. In analyses focusing on children (ages 0–14 years) and older adults (≥ 75 years of age), the authors reported associations that were larger in magnitude, compared to all ages for both asthma and allergic diseases hospital admissions ([Figure 5-3](#)).

The evidence from studies evaluated in the 2008 SO_x ISA, as well as recent studies indicating a positive association between short-term SO₂ exposure and asthma hospital admissions, is supported by a meta-analysis conducted by [Zheng et al. \(2015\)](#) that focused on all studies examining air pollution and asthma hospital admissions and ED visits published between 1988 and 2014. For SO₂, the authors reported a 2.1% increase (95% CI: 0.5, 3.70) in asthma hospital admissions for a 10-ppb increase in 24-h avg SO₂ concentrations based on estimates from 31 studies. The results from [Zheng et al. \(2015\)](#) are smaller in magnitude compared to the other asthma hospital admission studies summarized in [Figure 5-3](#), perhaps due to the meta-analysis only including single-day lag estimates from each of the studies. The results of the meta-analysis were found to be robust in sensitivity analyses examining publication bias; however, the publication bias analysis was not conducted separately for asthma hospital admissions and ED visits results.

Emergency Department Visits

The majority of studies, examining respiratory-related hospital admissions and ED visits, have focused on asthma ED visits. Studies evaluated in the 2008 SO_x ISA were primarily limited to single-city studies that provided generally positive associations between SO₂ and asthma ED visits, with positive associations being reported in some study locations and evidence of no association in other locations ([Figure 5-3](#)). Additionally, there was limited evidence for potential seasonal differences in SO₂ associations with asthma ED visits. As with the hospital admission studies, there has been limited analyses examining the potential confounding effects of copollutants on the SO₂-asthma ED visit relationship.

Recent studies that examined the association between short-term SO₂ exposures and asthma ED visits have primarily focused on either children or the entire population, with a few studies examining whether effects differ by lifestage. Additionally, unlike the hospital admission studies, the ED visit studies examined both 24-h avg and 1-h max exposure metrics, which can provide some additional insight, on a population level, into the short-term exposures that result in respiratory effects in controlled human exposure and animal toxicological studies (see previous subsections of [Section 5.2.1.2](#)).

[Strickland et al. \(2010\)](#) examined the association between SO₂ exposure and pediatric asthma ED visits (ages 5–17 years) in Atlanta, GA, using air quality data over the same years as [Tolbert et al. \(2007\)](#), who examined all respiratory ED visits. However, unlike [Tolbert et al. \(2007\)](#), who used a single-site monitor, [Strickland et al. \(2010\)](#) used population-weighting, a more refined exposure assignment approach, to combine daily pollutant concentrations across monitors. As discussed in [Section 3.4.4.2](#), a study by [Goldman et al. \(2012\)](#) shows that the bias in health effect estimates decreases when using population-weighted averages instead of the values from a fixed-site monitor for assigning exposure. In [Strickland et al. \(2010\)](#), the authors developed a statistical model using hospital-specific, time-series data that is essentially equivalent to a time-stratified, case-crossover analysis (i.e., using interaction terms between year, month, and day-of-week to mimic the approach of selecting referent days within the same month and year as the case day). [Strickland et al. \(2010\)](#) observed a 4.2% (95% CI: –2.1, 10.8) increase in ED visits for a 40-ppb increase in 1-h max SO₂ concentrations at lag 0–2 days in an all-year analysis. The potential confounding effects of other pollutants on the SO₂–asthma ED visit relationship was not assessed in this study, and correlations between pollutants were not presented. However, when evaluating the correlation of pollutants examined over the same study years in [Tolbert et al. \(2007\)](#), SO₂ had a low correlation with all pollutants ($r \leq 0.36$).

Positive associations between short-term SO₂ exposures and pediatric asthma ED visits were also observed in a study conducted by [Li et al. \(2011\)](#) in Detroit, MI that focused on whether there was evidence of a threshold in the air pollution–asthma ED visit relationship. In the main nonthreshold analysis, the authors conducted both time-series and time-stratified case-crossover analyses. [Li et al. \(2011\)](#) observed similar results in both analyses, which indicated an association between SO₂ and asthma ED visits, [time series: 20.5% (95% CI: 8.9, 33.2); lag 0–4 for a 10-ppb increase in 24-h avg SO₂ concentrations; case-crossover: 22.8% (95% CI: 12.6, 33.7); lag 0–4]. The results of the U.S.-based studies focusing on children conducted by [Strickland et al. \(2010\)](#) and [Li et al. \(2011\)](#) are consistent with those of [Jalaludin et al. \(2008\)](#) in a study of children 1–14 years of age conducted in Sydney, Australia. In addition to conducting the analysis focusing on ages 1–14 years, the authors also examined whether risks varied among age

ranges within this study population ([Chapter 6](#)). [Jalaludin et al. \(2008\)](#) examined single-day lags ranging from 0 to 3 days as well as the average of 0–1 days. In the 1–14 years of age analysis, the authors observed slightly larger associations at lag 0–1 days [29.7% (95% CI: 14.7, 46.5)] compared to lag 0 [22.0% (95% CI: 9.1, 34.5)] for a 10-ppb increase in 24-h avg SO₂ concentrations. An examination of the potential confounding effects of other pollutants was assessed in copollutant models with PM₁₀, PM_{2.5}, O₃, CO, or NO₂ at lag 0. SO₂ was found to be weakly to moderately correlated with these pollutants, $r = 0.27–0.52$. [Jalaludin et al. \(2008\)](#) reported that the SO₂-asthma ED visit association was slightly attenuated, but remained positive in all copollutant models, with the magnitude of the association ranging from a 13.2–16.1% increase in asthma ED visits.

[Byers et al. \(2015\)](#), in a study conducted in Indianapolis, IN, examined asthma ED visits across all ages as well as various lifestages (i.e., 5–17, 18–44, and ≥45 years of age). The authors used a double-weighted approach to assign exposure by first weighting air pollution concentrations by distance from a monitor to the ZIP code centroid and then weighting the concentrations by the age-specific census population. In an all-year analysis for all ages, the authors reported a 0.4% increase in asthma ED visits (95% CI: –3.6, 4.5) at lag 0–2 for a 40-ppb increase in 1-h max SO₂ concentrations, with evidence of a larger association when focusing on pediatric asthma ED visits [5.4% (95% CI: –3.2, 14.5); lag 0–2], which is consistent with [Strickland et al. \(2010\)](#), [Li et al. \(2011\)](#), and [Jalaludin et al. \(2008\)](#). Although copollutant analyses were not conducted, SO₂ was found to have a low correlation with PM_{2.5} ($r < 0.4$) in all-year and seasonal analyses, and moderate correlation with 1-h max and 8-h max O₃ in warm season analyses ($r = 0.42–0.45$). Additionally, when examining SO₂ concentrations across the entire study period, the authors noted that only 36 days (i.e., 2.1% of days) had 1-h max SO₂ concentrations that exceeded the NAAQS.

[Alhanti et al. \(2016\)](#) also used the approach of assigning exposure using population-weighting similar to [Strickland et al. \(2010\)](#), but expanded the study area from Atlanta, GA to include two additional cities: Dallas, TX and St. Louis, MO. The analysis focused on examining whether there was evidence of differential risk across lifestages (i.e., 0–4, 5–18, 19–39, 40–64, and 65+ years of age) for asthma ED visits across a number of air pollutants, including SO₂. Analyses were conducted for each individual city, and an overall estimate across all three cities was calculated by taking the inverse-variance weighted average of the city-specific risk estimate. Across the individual cities, there was evidence of positive and negative associations for all age categories examined except ages 5–18 years for which positive associations were observed across all cities, which is consistent with the single-city studies detailed above. In the combined analysis across the three cities, [Alhanti et al. \(2016\)](#) reported positive

associations for ages 0–4 years [4.1% (95% CI: –0.8, 9.2); lag 0–2 for 40-ppb increase in 1-h max SO₂ concentrations] and 5–18 years [5.7% (95% CI: –0.8, 11.8); lag 0–2] ([Sarnat, 2016](#)). In sensitivity analyses, the results were found to be robust to alternative model specifications for both control for temporal trends and weather covariates.

As detailed in the asthma hospital admissions section, [Zheng et al. \(2015\)](#) conducted a meta-analysis of asthma hospital admission and ED visit studies. In the analysis focusing on ED visit studies, the authors reported a 3.5% increase (95% CI: 1.9, 5.1) in asthma ED visits for a 10-ppb increase in 24-h avg SO₂ concentrations based on single-day lag estimates from 34 studies. This result is in the range of risk estimates reported in studies that observed positive associations between short-term SO₂ exposure and asthma ED visits ([Figure 5-3](#)).

Although a number of recent studies add to the evidence from the 2008 SO_x ISA indicating a positive association between asthma ED visits and short-term SO₂ exposures, not all studies have reported positive associations. Both [Stieb et al. \(2009\)](#) and [Villeneuve et al. \(2007\)](#), in studies conducted in seven Canadian cities and Edmonton, AB, respectively, did not observe evidence of a positive association between short-term SO₂ exposures and asthma ED visits ([Figure 5-3](#) and [Table 5-9](#)). The evidence of no association was observed over multiple lag structures (i.e., both single and multiday lags) ([Stieb et al., 2009](#); [Villeneuve et al., 2007](#)) as well as subdaily exposure metrics (i.e., 3-h avg pollutant concentrations) ([Stieb et al., 2009](#)).

Hospital Admissions and Emergency Department Visits for Respiratory Conditions Associated with Asthma

As stated previously, asthma is difficult to diagnose in children less than 5 years of age ([NAEPP, 2007](#)); however, asthma-like symptoms in children within this age range are often presented in the form of transient wheeze. Although studies that examine ED visits for wheeze do not directly inform upon the relationship between short-term SO₂ exposures and asthma, they can add supporting evidence. [Orazzo et al. \(2009\)](#) examined the association between short-term SO₂ exposures and wheeze ED visits, in children (ages 0–2 years) in six Italian cities. In a time-stratified case-crossover analysis, [Orazzo et al. \(2009\)](#) examined associations for multiday lags ranging from 0–1 to 0–6 days. The authors reported the strongest evidence for an association between short-term SO₂ exposures and wheeze ED visits at lags of 0–3 to 0–6 days with estimates ranging from 2.1 to 4.3%, respectively, for a 10-ppb increase in 24-h avg SO₂ concentrations. Within this study, copollutant analyses or correlations with other pollutants were not presented.

[Smargiassi et al. \(2009\)](#) and [Brand et al. \(2016\)](#) also provided additional information on whether there is an association between short-term SO₂ exposures and health effects that may be closely related to asthma. The distinction between asthma and asthma-related

outcomes is made in this case because of the focus on children 2–4 years of age in [Smargiassi et al. \(2009\)](#), which examined asthma hospital admissions and ED visits, and [Brand et al. \(2016\)](#), which examined hospital admissions for the combination of asthma and bronchiolitis. An asthma exacerbation for children 2–4 years of age may not necessarily represent the same health outcome as those studies discussed earlier in this section which include older individuals in whom asthma is more easily diagnosed.

In both [Smargiassi et al. \(2009\)](#) and [Brand et al. \(2016\)](#), the authors aimed to examine whether industrial sources of air pollution result in higher exposures to air pollutants, including SO₂, and subsequently an increase in asthma-related hospital admissions and ED visits. Within [Smargiassi et al. \(2009\)](#), the authors examined the influence of a point source of SO₂ (i.e., stack emissions from a refinery) in Montreal on asthma hospital admissions and ED visits using data from two fixed-site monitors as well as estimates of SO₂ concentrations from a dispersion model [American Meteorological Society/Environmental Protection Agency Regulatory Model (AERMOD)]. The authors examined both daily mean and daily peak SO₂ concentrations. When comparing SO₂ concentrations at one monitoring site east of the refinery with those obtained via AERMOD, the authors observed a modest correlation (daily mean SO₂, $r = 0.43$; daily peak SO₂, $r = 0.36$). An examination of hospital admissions and ED visits for locations east and southwest of the refinery found that associations with SO₂ estimates from AERMOD were slightly larger in magnitude for the same-day daily peak [hospital admissions: 1.46 (95% CI: 1.10, 1.93); ED visits: 1.18 (95% CI: 1.05, 1.33) for a 40-ppb increase in 1-h max SO₂ concentrations] compared to daily mean concentrations [hospital admissions: 1.36 (95% CI: 1.05, 1.81); ED visits: 1.15 (95% CI: 1.02, 1.27) for a 10-ppb increase in 24-h avg SO₂ concentrations] in an unadjusted model. When examining associations using SO₂ concentrations from ambient monitors, [Smargiassi et al. \(2009\)](#) did not find consistent evidence of an increase in asthma hospital admissions or ED visits. The authors also examined an adjusted model to control for daily weather variables and all other regional pollutants (i.e., PM_{2.5}, SO₂, NO₂, and O₃), but these results are not presented because, as discussed within this ISA, the evaluation of potential copollutant confounding is limited to two-pollutant models because the results from multipollutant models are difficult to interpret due to multicollinearity between pollutants. However, the results from the unadjusted (i.e., single-pollutant model) and adjusted models were generally similar.

Whereas [Smargiassi et al. \(2009\)](#) focused on a population residing near two refineries in Montreal, [Brand et al. \(2016\)](#) examined the association between air pollutant emissions and concentrations and asthma-related hospital admissions from a number of industrial facilities (i.e., metal smelters, pulp mills, and oil refineries) in both Montreal and British Columbia. To capture the potential influence of air pollutants, including SO₂, from

industrial facilities on asthma-related hospital admissions, the authors limited the analysis to air quality monitors and the population residing within 7.5 km of a facility. In a time-stratified, case-crossover analysis, the authors reported no evidence of an association when examining the relationship between 24-h avg and 1-h max SO₂ concentrations at lag 0 from “any industry” in individual city analyses as well as a pooled analysis across both cities with ORs ranging from 0.79–0.88 for a 10 ppb increase in 24-h avg SO₂ concentrations and 0.79–0.86 for a 40 ppb increase in 1-h max SO₂ concentrations. Only the analysis in Quebec examined associations between SO₂ from smelters and asthma-related hospital admissions, and as a result was excluded from the “any industry” pooled analysis. [Brand et al. \(2016\)](#) reported no evidence of an association in analyses using 24-h avg [OR = 0.88 (95% CI: 0.50, 1.50)] and 1-h max [OR = 0.79 (95% CI: 0.48, 1.25)] SO₂ concentrations that were identified as being from smelters. Collectively the results from [Smargiassi et al. \(2009\)](#) and [Brand et al. \(2016\)](#) provide initial evidence that SO₂ emitted from industrial facilities, as captured by measurements from ambient monitors, is not associated with an increase in asthma-related hospital admissions and ED visits. However, these results may reflect the fact that a monitor may not adequately capture spatial and temporal variability in SO₂ concentrations, including peak exposures of residents (see [Section 3.4.2](#)).

Outpatient and Physician Visits Studies of Asthma

Several recent studies examined the association between ambient SO₂ concentrations and physician or outpatient (nonhospital, non-ED) visits for asthma. In Toronto, [Burra et al. \(2009\)](#) examined asthma physician visits among patients aged 1–17 and 18–64 years in a study focusing on differences by sex and income within each age category. For children, the authors reported evidence of consistent positive associations between short-term increases in SO₂ concentrations and asthma physician visits for most of the single and multiday lags examined (i.e., 0, 0–1, 0–2, 0–3), with no evidence of an association for a 0–4-day lag. In the analysis of adults, a similar pattern of associations was observed; however, there was no evidence of an association at the two longest lags examined, 0–3 and 0–4 days.

In a study conducted in Atlanta, GA, [Sinclair et al. \(2010\)](#) examined the association between multiple respiratory outcomes, including asthma and outpatient visits, from a managed care organization. The authors separated the analysis into two time periods (the first 25 months of the study period and the second 28 months of the study period) to compare the air pollutant concentrations and relationships between air pollutants and acute respiratory visits for the 25-month time period examined in [Sinclair and Tolsma \(2004\)](#) (i.e., August 1998–August 2000), and an additional 28-month time period of available data from the Atlanta Aerosol Research and Inhalation Epidemiology Study

(ARIES) (i.e., September 2000–December 2002). As detailed in [Table 5-9](#), SO₂ concentrations were relatively similar between periods, differing by less than 2 ppb. A comparison of the two time periods indicated that risk estimates across outcomes tended to be larger in the earlier 25-month period compared to the later 28-month period, with evidence of consistent positive associations across the lags examined for asthma (both child and adult), but confidence intervals were relatively large.

Examination of Seasonal Differences

In addition to examining the association between short-term SO₂ exposures and asthma hospital admissions and ED visits in all-year analyses, some studies also conducted seasonal analyses. When evaluating these studies, it is important to note that the difference in the geographic locations examined across studies complicates the ability to draw overall conclusions regarding the seasonal patterns of associations.

In the study of eight South Korean cities, [Son et al. \(2013\)](#) examined potential seasonal differences across respiratory hospital admission outcomes. For asthma and allergic disease hospital admissions, the association with SO₂ was largest in magnitude during the summer, although confidence intervals were quite large [asthma: 19.1% (95% CI: –18.3, 73.9), lag 0–3; allergic disease: 21.9% (95% CI: –6.7, 58.6), lag 0–3 for a 10-ppb increase in 24-h avg SO₂ concentrations]. Across the eight cities, mean 24-h avg SO₂ concentrations were lowest during the summer season (4.4 ppb compared to a range of 4.8 to 7.0 in the other seasons), which was also observed for NO₂, PM₁₀, and CO. The seasonal asthma hospital admission results of [Son et al. \(2013\)](#) are similar to those reported in [Samoli et al. \(2011\)](#) in a study conducted in Athens, Greece. [Samoli et al. \(2011\)](#) observed the largest magnitude of an association during the summer months [46.6% (95% CI: –13.8, 149.3); lag 0 for a 10-ppb increase in 24-h avg SO₂ concentrations], but also reported a similar association in the autumn months [42.6% (95% CI: –0.5, 104.4); lag 0]. Although positive, associations for the winter and spring months were smaller in magnitude, 20.2 and 31.8%, respectively.

The initial indication of larger associations during the summer for asthma hospital admissions is further supported by the analysis of [Strickland et al. \(2010\)](#) examining short-term SO₂ exposures and pediatric asthma ED visits in Atlanta. The authors reported evidence of asthma ED visit associations larger in magnitude during the summer [10.8% (95% CI: 0.7, 21.7); lag 0–2 for a 40-ppb increase in 1-h max SO₂ concentrations], with no evidence of an association during the winter [0.4% (95% CI: –7.5, 9.0)]. These results are consistent with [Byers et al. \(2015\)](#), who reported associations larger in magnitude in the summer for all ages [3.1% (95% CI: –2.6, 8.6); lag 0–2 for a 40-ppb increase in 1-h max SO₂ concentrations], and particularly children 5–17 years of age [13.0% (95% CI: 0.8, 26.8); lag 0–2], and no evidence of an association in the cold season across all ages

examined. However, in another study focusing on asthma physician visits in Atlanta, [Sinclair et al. \(2010\)](#) reported inconsistent evidence of seasonal differences in risk estimates, with the pattern of associations being different in each of the time periods examined in the study. It is important to note that the results of [Sinclair et al. \(2010\)](#) may be a reflection of the severity of asthma exacerbations requiring medical attention and people proceeding directly to a hospital for treatment instead of first visiting a physician. Therefore, the study may not be able to adequately capture associations, and specifically, any potential seasonal differences.

The meta-analysis conducted by [Zheng et al. \(2015\)](#) provides some additional supporting evidence for potential seasonal differences in SO₂-asthma hospital admission and ED visit associations. In a combined analysis including both asthma hospital admission and ED visit studies that reported seasonal results, [Zheng et al. \(2015\)](#) reported slightly larger associations in the warm [4.8% (95% CI: 2.7, 7.0) for a 10-ppb increase in 24-h avg SO₂ concentrations] compared to the cold season [3.2% (95% CI: 0.5, 5.9)], but confidence intervals did overlap.

Although there is some evidence for larger associations during the summer, studies conducted by [Villeneuve et al. \(2007\)](#) in Edmonton, AB and [Jalaludin et al. \(2008\)](#) in Sydney, Australia present conflicting results. As stated above, [Villeneuve et al. \(2007\)](#) did not find evidence of an association between short-term SO₂ exposures and asthma ED visits, including in the seasonal analysis, while [Jalaludin et al. \(2008\)](#) reported evidence of larger associations during the cold months (May–October) compared to the warm months (November–April) ([Figure 5-3](#)).

Overall, the results of [Samoli et al. \(2011\)](#), [Son et al. \(2013\)](#), [Strickland et al. \(2010\)](#), and [Byers et al. \(2015\)](#) suggest that associations are larger in magnitude during the summer season, but this conclusion should be viewed with caution because the results of each study are highly imprecise, as reflected by the wide confidence intervals for each seasonal result. Additionally, the interpretation of results from these studies is complicated by the lack of copollutant analyses and by the results from [Villeneuve et al. \(2007\)](#) and [Jalaludin et al. \(2008\)](#) which do not show evidence of larger associations during the summer or warm season.

Lag Structure of Associations

When examining associations between air pollution and a specific health outcome, such as respiratory-related hospital admissions, it is informative to assess whether exposure to an air pollutant results in an immediate, delayed, or prolonged effect on health. Recent studies that examine both multiple single- and multiday lags can help provide information

on whether there is a specific exposure window(s) that contribute to SO₂-related asthma hospital admissions and ED visits.

[Son et al. \(2013\)](#) examined the lag structure of associations for multiple respiratory-related hospital admissions, including asthma and allergic disease, by analyzing both single- and multiday lags. Across single-day lags of 0 to 3 days, positive associations were observed across each lag, but the magnitude of the association varied across single-day lags for each outcome. For both asthma and allergic disease hospital admissions, the largest association, in terms of magnitude, for SO₂ was observed for each of the multiday lags examined, with the largest occurring at lag 0–3 days [asthma: 5.3% (95% CI: –2.4, 13.0); allergic disease: 3.1% (95% CI: –3.7, 10.7) for a 10-ppb increase in 24-h avg SO₂ concentrations].

Studies conducted by [Samoli et al. \(2011\)](#) and [Jalaludin et al. \(2008\)](#) report evidence for the strongest SO₂-asthma hospital admission and ED visit associations occurring rather immediately (lag 0) as well as over the first few days after exposure, average of lags from 0 up to 2 days. [Samoli et al. \(2011\)](#) in the examination of single- and multiday lags for associations between SO₂ and asthma hospital admissions in Athens, Greece found associations of similar magnitude at lag 0 and a 0–2 day distributed lag, but the confidence interval around the association from the distributed lag model was wide (quantitative results not presented). The associations reported for single-day lags of 1 and 2 days were small and close to null. [Jalaludin et al. \(2008\)](#), in a study in Sydney, Australia, found when examining single-day lags of 0 to 3 days that asthma ED visit associations were largest for lag 0 [22.0% (95% CI: 9.1, 34.5) for a 10-ppb increase in 24-h avg SO₂ concentrations] and 1 day [16.1% (95% CI: 5.1, 26.5)]. This result is further reflected in the largest SO₂ association being observed for the multiday lag of 0–1 days [29.7% (95% CI: 14.7, 46.5)].

Only a limited number of studies have examined the lag structure of associations, and the results across studies are not fully supported by the rest of the literature base. [Villeneuve et al. \(2007\)](#), when studying asthma ED visits in seven Canadian cities, examined single-day lags (0 and 1 day) as well as multiday lags of 0–2 and 0–4 days. The authors reported no evidence of an association between short-term SO₂ exposures and asthma ED visits at any lag. Additionally, [Orazio et al. \(2009\)](#) in the study of wheeze ED visits in six Italian cities, examined multiday lags ranging from 0–1 to 0–6 days. Across the lags examined, the authors reported evidence of increasing magnitude of the association as the length of the multiday lag increased, with lag 0–6 days showing the largest association.

Exposure Assignment

Questions often arise in air pollution epidemiologic studies about the method used to assign exposure (see [Section 3.3.3](#)). [Strickland et al. \(2011\)](#), using ED visit data from Atlanta, GA, assessed the effect of various exposure assignment approaches on the relationship between short-term air pollution exposures and asthma ED visits. The authors used warm season data from [Strickland et al. \(2010\)](#) to examine the relative influence of different exposure assignment approaches (i.e., central monitor, unweighted average across available monitors, and population-weighted average) on the magnitude and direction of associations between SO₂ and pediatric asthma ED visits. SO₂ exhibited a relatively low chi-square goodness-of-fit statistic compared with other pollutants, which the authors attributed to spatial heterogeneity in SO₂ concentrations ([Section 3.4.2.2](#)). [Strickland et al. \(2011\)](#) reported that effect estimates per interquartile range (IQR) increase in SO₂ were similar across the metrics; however, based on a standardized increment (i.e., 20 ppb in the study), the magnitude of the association between SO₂ and pediatric asthma ED visits varied [central monitor 3.0% (95% CI: -0.4, 8.4); unweighted average 12.8% (95% CI: 2.8, 23.4); population-weighted average 10.9% (95% CI: 0.8, 21.9) for a 40-ppb increase in 1-h max SO₂ concentrations at lag 0–2 days]. The difference in associations observed across the various exposure assignment approaches when using the standardized increment can be attributed to the value (i.e., a 1-h max SO₂ concentration of 20 ppb) not reflecting an increase in SO₂ concentrations that is reflective of the SO₂ distribution in Atlanta (e.g., in the study, the standardized increment for 1-h max SO₂ is 20 ppb, but the IQR, which is often used to calculate the relative risk (RR), differs across the exposure assignment approaches, varying from 9.6 to 13.9 ppb).

Concentration-Response Relationship

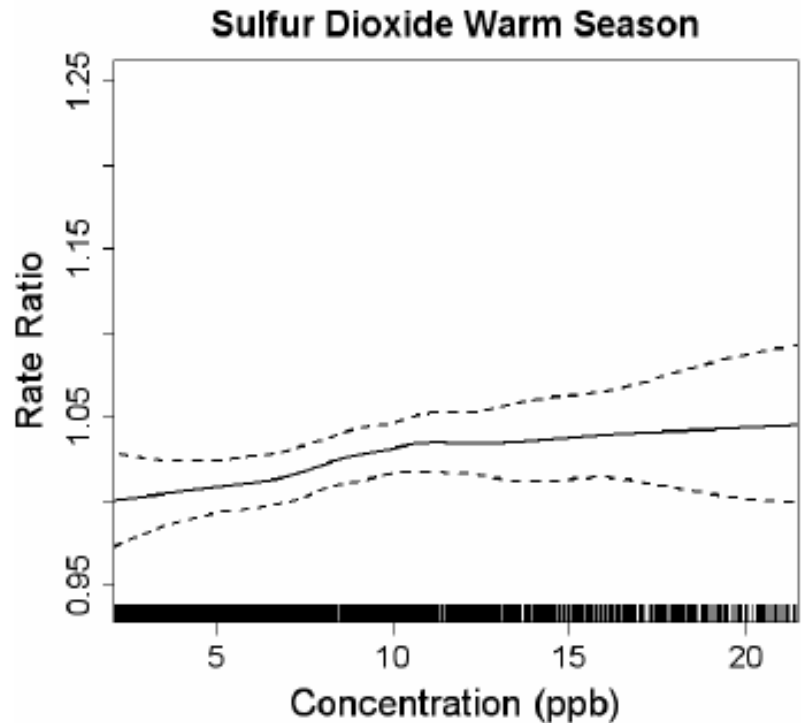
To date, few studies have examined the concentration-response (C-R) relationship between SO₂ concentration in ambient air and respiratory morbidity. In recent epidemiologic studies, [Strickland et al. \(2010\)](#) and [Li et al. \(2011\)](#) examined the shape of the SO₂-pediatric asthma ED visit relationship using different analytical approaches.

[Strickland et al. \(2010\)](#) examined the C-R relationship by conducting quintile and locally weighted scatterplot smoothing (LOESS) C-R analyses. In the quintile analysis, SO₂ associations were examined in both the warm and cold seasons; however, no associations were observed for the cold season for any quintile. Focusing on the warm season, the authors found evidence of an increase in the magnitude of the association for 1-h max concentrations within the range of 7 to <24.2 ppb, relative to the first quintile (i.e., SO₂ concentrations <3.1 ppb). The smallest associations were observed for the 5th quintile, which represented 1-h max concentrations ranging from 24.2 to ≤149 ppb; however, this quintile represented the extreme end of the distribution of SO₂ concentrations where data

density was low. Additionally, the LOESS C-R relationship analysis provides evidence of a linear relationship between short-term SO₂ exposures and asthma ED visits along the distribution of 1-h max concentrations from the 5th (2.1 ppb) to 95th (21.5 ppb) percentile ([Sacks, 2015](#)) ([Figure 5-4](#)). Collectively, these analyses do not provide evidence of a threshold.

In a study conducted in Detroit, MI, [Li et al. \(2011\)](#) examined whether there is evidence of a nonlinear C-R relationship for air pollutants and pediatric asthma ED visits. Associations with SO₂ were examined in both a time-series and time-stratified, case-crossover study design assuming (1) a linear relationship and (2) a nonlinear relationship starting at a 24-h avg concentration of 8 ppb [i.e., the maximum likelihood estimate within the 10th to 95th percentile concentration where a change in linearity may occur (~91st percentile)]. It is important to note the analysis that assumed a nonlinear relationship did not assume zero risk below the inflection point. The focus of the analysis was on identifying whether risk increased above that observed in the linear models at 24-h avg SO₂ concentrations above 8 ppb. In the analyses assuming linearity, the authors examined single-day lags of 3 and 5 days and multiday lags of 0–2 and 0–4 days. Positive associations were observed for all lags examined and were relatively consistent across models, with the strongest association for a 0–4-day lag [time series: 20.5% (95% CI: 8.9, 33.2); case-crossover: 22.8% (95% CI: 12.6, 33.7) for a 10-ppb increase in 24-h avg SO₂ concentrations]. In the models that assumed a nonlinear relationship, the authors did not observe evidence of increased risk above ~8 ppb. However, it is important to note that the data density is low at 24-h avg concentrations greater than 8 ppb, as reflected by this value representing the ~91st percentile of SO₂ concentrations.

Collectively, [Strickland et al. \(2010\)](#) and [Li et al. \(2011\)](#) provide initial evidence of a log-linear, no-threshold relationship between short-term SO₂ concentrations and asthma ED visits. However, it is important to note that these studies have not fully explored potential alternatives to linearity when examining the shape of the C-R relationship, which in combination with the potential measurement error due to lack of characterization of the spatial and temporal variability in SO₂ exposure concentrations, complicates the interpretation of the SO₂-asthma ED visit C-R relationship ([Sections 3.4.2.2](#) and [3.4.2.3](#)).



Note: solid line = smoothed concentration-response estimate. Dashed line = twice-standard error estimates.
 Source: [Strickland et al. \(2010\)](#). Reprinted with permission of the American Thoracic Society.

Figure 5-4 Concentration-response for associations between 3-day average (lag 0–2) sulfur dioxide concentrations and emergency department visits for pediatric asthma at the 5th to 95th percentile of sulfur dioxide concentrations in the Atlanta, GA area.

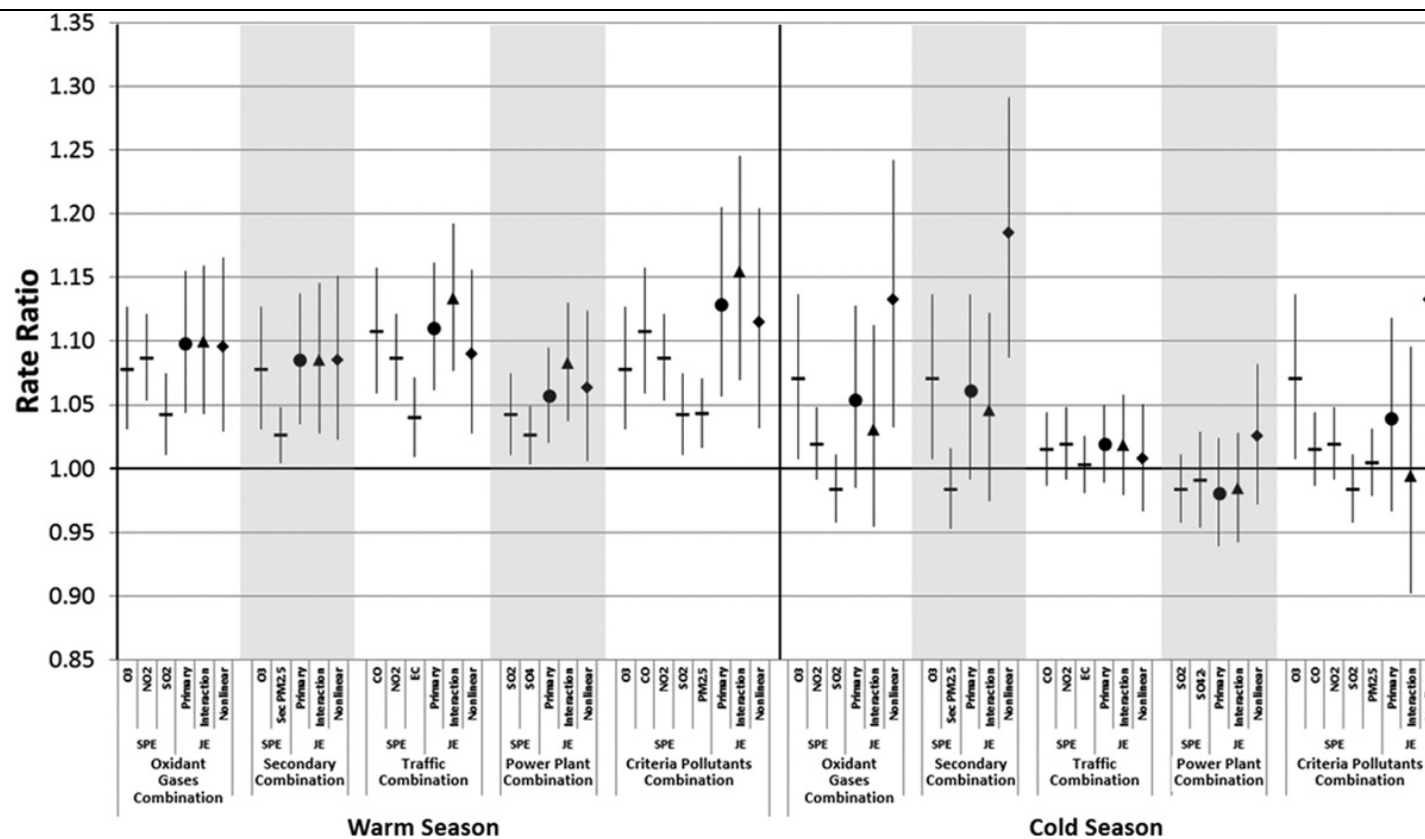
Sulfur Dioxide within the Multipollutant Mixture

An important question during the review of any criteria air pollutant, is whether the pollutant has an independent effect on human health. Ambient exposures to criteria air pollutants are in the form of mixtures, however, making this question difficult to answer in epidemiologic studies, especially when the pollutant of interest is highly correlated with other pollutants in the mixture. Epidemiologic studies traditionally try to identify the independent effect of a criteria air pollutant through the use of copollutant models, but these methods do not consider the broader air pollution mixture. Recent studies conducted by [Winqvist et al. \(2014\)](#) and [Pearce et al. \(2015\)](#) using pediatric asthma ED visits data from Atlanta assessed whether specific mixtures are more strongly associated with health effects compared to others. Although the primary objective of these types of studies is not to directly assess the independent effects of a pollutant, they can inform the

understanding of the role of SO₂ in the air pollution mixture (e.g., contributing to an additive or synergistic effect).

[Winguist et al. \(2014\)](#) examined multipollutant mixtures by focusing on the joint effect (i.e., the combined effect of multiple pollutants) of pollutants often associated with specific air pollution sources. Associations between short-term SO₂ exposures and pediatric asthma ED visits (i.e., ages 5–17 years) were examined in single-pollutant models and also in a multipollutant context in joint models for pollutant combinations representative of irritant gases (i.e., O₃, NO₂, and SO₂), power plants (i.e., SO₂ and SO₄²⁻), and NAAQS pollutants (i.e., O₃, CO, NO₂, SO₂, and PM_{2.5}). It is important to note that the pollutant combination analyses attempt to address a different question (i.e., what is the risk associated with exposure to a combination of pollutants?) than a traditional copollutant analysis, which focuses on identifying the independent effect of a pollutant. Using the model detailed in [Strickland et al. \(2010\)](#), the authors examined the relationship between each combination and pediatric asthma ED visits using a Poisson model in the context of a time-referent case-crossover analysis. The authors reported results for an IQR increase for lag 0–2 days in single-pollutant analyses as well as three types of joint effect models [i.e., no interaction terms (primary), first-order multiplicative interactions between pollutants (interactions), and nonlinear pollutant terms (nonlinear)] ([Figure 5-5](#)).

In single-pollutant analyses, SO₂ associations were smaller in magnitude compared to the other pollutants that comprised each pollutant combination, but the confidence interval surrounding each SO₂ estimate was relatively narrow. Across pollutant combinations that contained SO₂, joint effect models reported consistent positive associations with pediatric asthma ED visits in the warm season. Additionally, for each pollutant combination the association observed was larger in magnitude than any single-pollutant association, including SO₂, but not equivalent to the sum of each individual pollutant association for a specific combination. In the warm season analyses, associations across the different joint effect models were relatively similar. Overall, the results during the cold season were more variable.



CO = carbon monoxide; EC = elemental carbon; JE = joint model estimate; NO₂ = nitrogen dioxide; O₃ = ozone; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; SO₂ = sulfur dioxide; SO₄ = sulfate; SPE = single-pollutant model estimate.

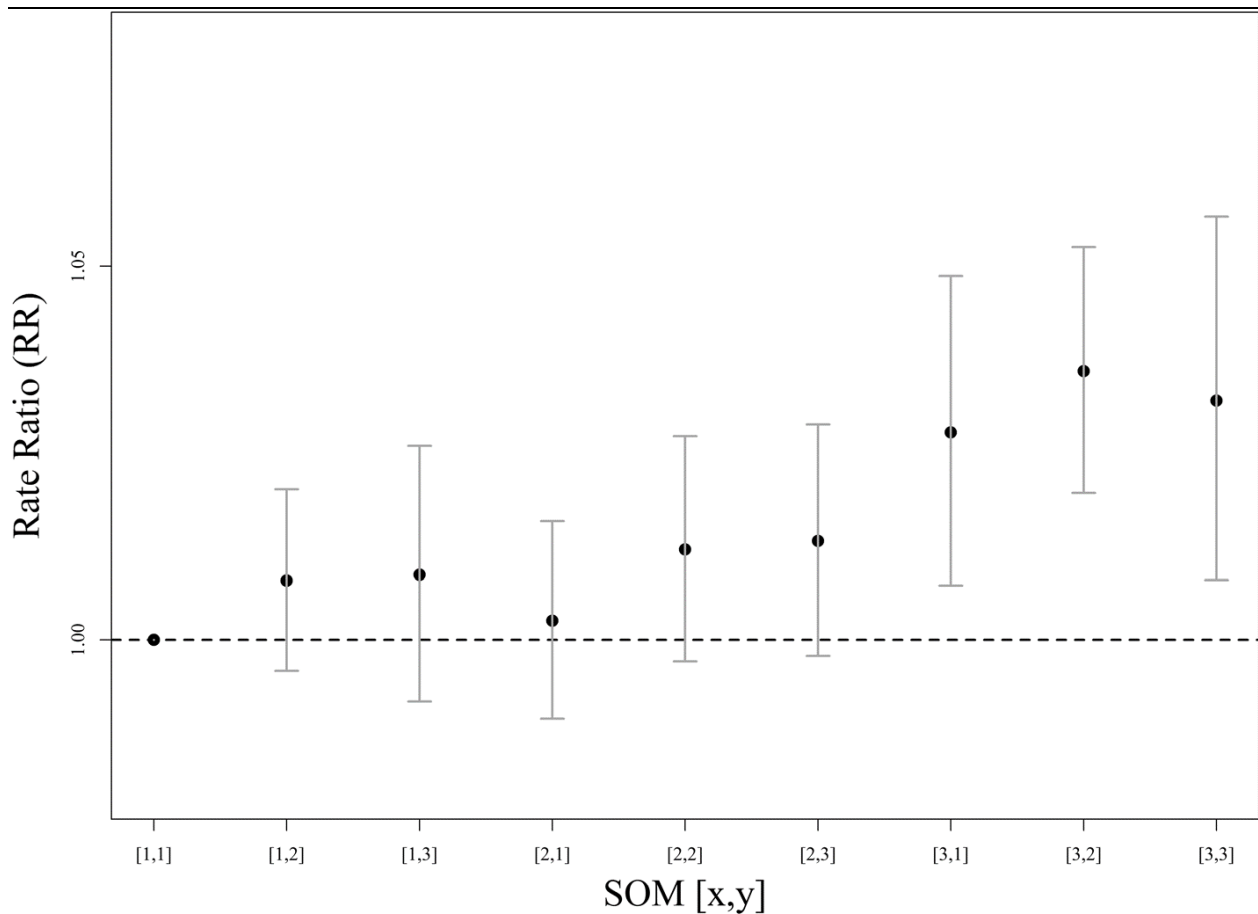
Note: Interquartile range for 1-h max SO₂ concentrations = 10.51 ppb.

Source: (Winquist et al., 2014). Reprinted with permission of Wolters Kluwer Health.

Figure 5-5 Rate ratio and 95% confidence intervals for single-pollutant and joint effect models for each pollutant combination in warm and cold season analyses for an interquartile range increase in each pollutant at lag 0–2 days.

[Pearce et al. \(2015\)](#) took a different approach to examining multipollutant mixtures by using an unsupervised learning tool, the self-organizing map (SOM), which is similar to cluster analysis. Using air pollution concentrations for 10 pollutants from a single monitor, the authors identified nine distinct day types representative of air quality in Atlanta during the study period. These unique days were then used as indicator variables to examine associations with pediatric asthma ED visits using the same statistical approach as [Strickland et al. \(2010\)](#) and [Winqvist et al. \(2014\)](#). Across the nine SOMs, some pollutant combinations represented days consisting of high single pollutant extremes, which included a day with high 1-h max SO₂ concentrations (i.e., mean concentration of 48.8 ppb and concentrations ranging from 8.5–23.7 ppb for all other SOMs). In analyses of all SOMs focusing on lag 1, the strongest associations were observed for days representing above average concentrations for all pollutants, and for days representing a collection of primary (i.e., CO, NO₂, NO_x, EC, and OC) or secondary pollutants (i.e., O₃, ammonium ion (NH₄⁺), and SO₄) ([Figure 5-6](#)). Additional evidence of associations with pediatric asthma ED visits was observed for days with single pollutant extremes, including days with high SO₂ concentrations and generally lower concentrations for all other pollutants ([Figure 5-6](#)). Interestingly, when comparing SOMs results with single-pollutant results in sensitivity analyses, the authors reported a null association with SO₂ at lag 1. This result differs from that observed in [Strickland et al. \(2010\)](#) and [Winqvist et al. \(2014\)](#), but the difference could be due to the fact that [Pearce et al. \(2015\)](#) focused only on lag 1 because they were examining distinct pollution profiles that often do not occur on multiple days in a row. In contrast, [Strickland et al. \(2010\)](#) and [Winqvist et al. \(2014\)](#) examined associations over a multiday average of 0–2 days. Additionally, the difference between the SOM and single-pollutant SO₂ result could be because the SOM with high SO₂ concentrations was better able to capture the immediate respiratory response due to higher peak concentrations, which would be consistent with the effects observed in controlled human exposure and animal toxicological studies.

Although the single-pollutant results of [Winqvist et al. \(2014\)](#) and [Pearce et al. \(2015\)](#) differ due to the lags examined, the studies contribute to evidence that SO₂ alone and in combination with other pollutants is associated with asthma ED visits. The studies also highlight the difficulty in separating out the independent effect of a pollutant that is part of a mixture because multiple pollutants are often highly correlated.



SOM = self-organizing map.

Note: [2,2] = days with high sulfur dioxide concentrations. [3,3] and [3,1] = days with primary and secondary pollutants, respectively. [3,2] = days with above average concentrations for all pollutants.

Source: (Pearce et al., 2015). Reprinted from Environmental Health.

Figure 5-6 Rate ratio and 95% confidence interval for association between self-organizing map-based multipollutant day type and pediatric asthma emergency department visits at lag 1.

Summary of Asthma Hospital Admission and Emergency Department Visits

Recent studies that examined the association between short-term SO₂ exposure and asthma hospital admissions and ED visits generally report positive associations in studies examining all ages, children (i.e., <18 years of age), and older adults (i.e., 65 years of age and older) (Figure 5-3). The pattern of associations observed across studies focusing on all ages as well as age-stratified analyses is consistent with those studies evaluated in the 2008 SO_x ISA. Across asthma hospital admission and ED visit studies that evaluated the lag structure of associations, the most consistent evidence indicated that associations

were largest in magnitude for multiday lags that encompassed the first few days after exposure (i.e., average of 0–2 and 0–3 day lags). This evidence generally supports the timing of SO₂ effects observed in the controlled human exposure and animal toxicological studies ([Section 5.2.1.2](#)). The examination of potential copollutant confounding was rather limited in the body of studies that focused on asthma hospital admissions and ED visits ([Samoli et al., 2011](#); [Jalaludin et al., 2008](#)). Across studies, SO₂ was found to be low ($r < 0.4$) to moderately ($r = 0.4–0.7$) correlated with other pollutants examined. Evidence from these studies is consistent with those studies evaluated in the 2008 SO_x ISA and adds to the body of evidence indicating that SO₂-asthma hospital admission and ED visit associations remain positive, although in some instances attenuated in magnitude, in copollutant models.

A number of recent studies also examined whether there was evidence that the association between short-term SO₂ exposures and asthma hospital admissions and ED visits was modified by season or some other individual- or population-level factor ([Chapter 6](#)). An examination of seasonal differences in SO₂-asthma hospital admission and ED visit associations provide some evidence of SO₂ effects being larger in magnitude in the summer or warm season, but the lack of this pattern across all studies that conducted seasonal analyses suggests that seasonal associations may vary by geographic location. Studies of individual- and population-level factors provide some evidence of differences in associations by lifestage, for children and older adults, and more limited evidence for differences by sex ([Chapter 6](#)).

Additionally, some recent studies examined various study design issues, including model specification and exposure assignment. An examination of model specification, as detailed in [Section 5.2.1.6](#), indicates that the relationship between short-term SO₂ exposures and respiratory-related hospital admissions, including those for asthma and allergic disease, are sensitive to using less than 7 degrees of freedom (df) per year to account for temporal trends, but robust to alternative lags and df ranging from 3 to 6 for weather covariates ([Son et al., 2013](#)). The results of [Son et al. \(2013\)](#) are supported by the sensitivity analyses examining model specification conducted by [Alhanti et al. \(2016\)](#) for asthma ED visits where the results were relatively consistent when the number of df for temporal trends was increased and alternative covariates for weather used. An examination of various exposure assignment approaches, including single fixed-site, average of multiple monitors, and population-weighted average, suggests that each approach may influence the magnitude, but not direction, of the SO₂-asthma ED visit risk estimate ([Strickland et al., 2011](#)).

Finally, recent studies examined whether the shape of the SO₂-asthma ED visits C-R relationship is linear or provides evidence of a threshold. These studies provide initial

evidence of a log-linear, no-threshold relationship between short-term SO₂ exposures and asthma ED visits ([Li et al., 2011](#); [Strickland et al., 2010](#)), but a thorough empirical exploration of alternatives to linearity has not been conducted. An examination of seasonal differences in SO₂-asthma hospital admission and ED visit associations provide some evidence of SO₂ effects being larger in magnitude in the summer or warm season, but the lack of this pattern across all studies that conducted seasonal analyses suggests that seasonal associations may vary by geographic location.

Subclinical Effects Underlying Asthma Exacerbation: Pulmonary Inflammation and Oxidative Stress

Pulmonary inflammation is a key subclinical effect in the pathogenesis of asthma. It consists of both acute and chronic responses and involves the orchestrated interplay of the respiratory epithelium and both the innate and adaptive immune system. The immunohistopathologic features of chronic inflammation involve infiltration of airway epithelium by inflammatory cells such as eosinophils, lymphocytes, mast cells, and macrophages and the release of inflammatory mediators such as cytokines and leukotrienes. Oxidative stress is also relevant to asthma exacerbation. Many transcription factors regulating the expression of pro-inflammatory cytokines are redox sensitive, and inflammatory cells that respond to cytokine signaling can generate reactive oxygen species leading to oxidative stress.

This section characterizes the evidence on SO₂ exposure effects on pulmonary inflammation and oxidative stress in humans with asthma and in animal models of allergic airway disease (see [Section 5.2.1.7](#) for healthy humans and animal models). The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) concluded that evidence from the limited number of controlled human exposure, epidemiologic, and animal toxicological studies was insufficient to determine that exposure to SO₂ at current ambient concentrations was associated with inflammation in the airway. However, several studies provided evidence for subclinical effects related to allergic inflammation. There are no recent controlled human exposure studies, but there is additional investigation in epidemiologic and animal toxicological studies. Epidemiologic results are inconsistent for pulmonary inflammation and oxidative stress, including those for SO₂ measured at or near children's schools. However, recent findings in rats link short-term SO₂ exposure to allergic inflammation.

Controlled Human Exposure Studies

Pulmonary inflammation following 5–10-minute exposure to SO₂ was discussed in the previous ISA; no new studies were available for review. Briefly, [Tunnicliffe et al. \(2003\)](#) measured levels of exhaled nitric oxide (eNO), an indirect marker for pulmonary inflammation, in individuals with asthma before and after a 1-hour exposure to 0.2 ppm

SO₂ under resting conditions. Nasal lavage fluid levels of the antioxidants, ascorbic and uric acid, were also measured pre- and post-exposure. No statistically significant differences were observed between pre- and post-exposure for any of these indicators. Because subjects were exposed at rest and exposed to low concentrations, it is unlikely that enough SO₂ reached the airways to cause an effect. [Gong et al. \(2001\)](#) evaluated the response of individuals with asthma to 0.75 ppm SO₂ during exercise. In addition to changes in lung function and symptoms, there was a statistically significant increase in eosinophil count in induced sputum 2 hours after a 10-minute exposure. This response was significantly dampened by pretreatment with a leukotriene receptor antagonist. These results provided some evidence that SO₂ elicits an inflammatory response in the airways of individuals with asthma that extends beyond the immediate bronchoconstriction response typically associated with SO₂ exposure. Additionally, this study provides further evidence that the bronchoconstriction response is only partially due to neural reflexes and that inflammatory mediators play an important role ([Section 4.3.1](#)).

Epidemiologic Studies

Recent epidemiologic evidence is inconsistent for associations of short-term increases in ambient SO₂ concentration with pulmonary inflammation and oxidative stress in adults and children with asthma ([Table 5-10](#)). Outcomes were assessed at varying frequency: daily, weekly, or seasonally. All studies examined eNO. Higher eNO has been linked to higher eosinophil counts ([Brody et al., 2013](#)) as well as prevalence and exacerbation of asthma ([Soto-Ramos et al., 2013](#); [Carraro et al., 2007](#); [Jones et al., 2001](#); [Kharitonov and Barnes, 2000](#)). An SO₂-associated increase in eNO was observed in a population of adults with asthma with high prevalence of atopy (90%) ([Maestrelli et al., 2011](#)) ([Table 5-10](#)), although associations with lung function or asthma control score were not observed. The results of [Maestrelli et al. \(2011\)](#) are also supported by findings that allergic inflammation in rats persists 24 hours after SO₂ exposures repeated over many days. The multicity U.S. asthma medication trial observed imprecise associations for eNO with wide 95% CIs in the ICS, beta-agonist, and placebo groups ([Qian et al., 2009a](#)). Both studies of adults with asthma estimated SO₂ exposure from fixed-site monitors. Neither indicated whether the measurements adequately represented the spatiotemporal variability in SO₂ concentrations in the study area, and the U.S. study averaged concentrations from monitors within 32 km of each subject's ZIP code centroid.

Two recent studies measured SO₂ at or 0.65 km from children's schools ([Greenwald et al., 2013](#); [Lin et al., 2011b](#)), which may better represent some component of subjects' exposure. Results are inconsistent. Percent changes in eNO were 31 (95% CI: -24, 119) per 10-ppb increase in SO₂ measured at a school in El Paso, TX ([Greenwald et al., 2013](#)) and 5.5 (95% CI: 2.7, 8.3) per 10-ppb increase in SO₂ measured near a school in Beijing,

China before and after the 2008 Olympics ([Lin et al., 2011b](#)). Among children with asthma not using ICS in Windsor, ON, SO₂ concentrations at a monitor within 10 km of homes were not associated with eNO but were associated with markers of oxidative stress in exhaled breath condensate (EBC) ([Liu et al., 2009b](#)). The school-based studies differed in lags examined, and an association was observed with lag 0 SO₂ ([Lin et al., 2011b](#)) but not lag 0–3 avg SO₂ ([Greenwald et al., 2013](#)). For SO₂ measured at fixed-site monitors, associations were observed with both lag 0 and lag 0–2 avg concentrations ([Liu et al., 2009b](#)). Prevalence of atopy was not reported for the study populations of children.

Findings are inconsistent for SO₂ associations with pulmonary inflammation and oxidative stress in children and adults with asthma. Copollutant confounding is an additional uncertainty. Associations were observed with PM_{2.5}, BC, CO, O₃, and NO₂ ([Lin et al., 2011b](#); [Maestrelli et al., 2011](#); [Liu et al., 2009b](#)). Only [Liu et al. \(2009b\)](#) reported SO₂-copollutant correlations, indicating the potential for confounding with PM_{2.5} ($r = 0.56$), less so with NO₂ ($r = 0.18$), and likely not with O₃ ($r = -0.02$). [Maestrelli et al. \(2011\)](#) did not examine copollutant models, and results in children with asthma are conflicting. For pollutants measured 0.65 km from school, SO₂ associations with eNO persisted but decreased with adjustment for PM_{2.5} or BC ([Lin et al., 2011b](#)). Based on pollutants measured up to 10 km from home, the SO₂ association with oxidative stress decreased with adjustment for NO₂ and did not persist with adjustment for PM_{2.5} ([Liu et al., 2009b](#)) ([Table 5-10](#)).

Animal Toxicological Studies

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) discussed several studies that investigated the effects of exposure to SO₂ on inflammatory responses. While one study failed to demonstrate inflammation following a single subacute exposure to 1 ppm SO₂ ([U.S. EPA, 2008d](#)), other studies found that repeated SO₂ exposure enhanced the development of an allergic phenotype and altered physiologic responses in animal models of allergic airway disease. These studies demonstrating effects of repeated SO₂ exposures in models of allergic airway disease are listed in [Table 5-11](#) and described here. In addition, other studies involving repeated SO₂ exposures in naive rats, including studies that demonstrate increased sensitivity to allergens, have been conducted and are described below in [Section 5.2.1.7](#)

Table 5-10 Epidemiologic studies of pulmonary inflammation and oxidative stress in populations with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
Adults with asthma				
<p>†Qian et al. (2009b) Boston, MA; New York City, NY; Philadelphia, PA; Madison, WI; Denver, CO; San Francisco, CA; 1997–1999</p> <p>N = 119, ages 12–65 yr. 100% persistent asthma. 1/3 ICS use, 1/3 beta-agonist use, 1/3 placebo use. Examined every 2–4 wk for 16 wk. Recruited from clinics as part of an asthma medication trial. Multiple comparisons—many pollutants, lags, medication use analyzed.</p>	<p>Monitors averaged within 32 km of subject ZIP code centroid.</p> <p>Mean (SD): 5.3 (4.4)</p> <p>75th percentile: 7.6</p> <p>Max: 27</p>	24-h avg 0	<p>Change in eNO (ppb)</p> <p>All subjects: 0.09 (–0.07, 0.25)</p> <p>ICS: 0.17 (–0.11, 0.44)</p> <p>Beta-agonist: 0.04 (–0.18, 0.27)</p>	<p>Remains null with: PM₁₀, NO₂, or O₃ (all subjects, lag 0)</p> <p><i>r</i> = 0.58 NO₂, NR PM₁₀.</p>
		0–3 avg	<p>All subjects: 0.07 (–0.12, 0.26)</p> <p>ICS: 0.15 (–0.13, 0.43)</p> <p>Beta-agonist: 0.10 (–0.19, 0.38)</p>	
<p>†Maestrelli et al. (2011) Padua, Italy, 2004–2005</p> <p>N = 32, mean (SD) age 40 (7.5) yr. 81% persistent asthma. 69% ICS use. 90% atopy.</p> <p>Six measures over 2 yr. Recruited from database of beta-agonist users (>6 times per yr for 3 yr).</p>	<p>Two monitors in city</p> <p>Medians across seasons: 0.87–2.7</p> <p>75th percentiles across seasons: 1.3–4.1</p>	24-h avg 0	<p>Change in eNO (ppb)</p> <p>All subjects: 55 (–2.3, 113)</p> <p>Nonsmokers: 82 (3.1, 161)</p>	<p>No copollutant model</p> <p>No association with personal or fixed-site PM_{2.5} or PM₁₀.</p> <p>Copollutant correlations NR.</p>
			<p>Change in EBC pH</p> <p>Decrease = more inflammation</p> <p>All subjects: 0.46 (–0.20, 1.1)</p> <p>Nonsmokers: 0.18 (–0.34, 0.69)</p> <p>n = 22</p>	

Table 5-10 (Continued): Epidemiologic studies of pulmonary inflammation and oxidative stress in populations with asthma published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
Children with asthma				
† Greenwald et al. (2013) El Paso, TX, Mar–Jun 2010 N = 38, mean age 10 yr. 47% daily asthma medication use. Weekly measures for 13 wk. Recruited from schools.	Monitor at school A: residential area B: 91 m from major road Mean (SD): 1.2 (0.44) and 0.84 (0.54) Upper percentiles NR.	24-h avg 0–3 avg	Percent change in eNO A: –59 (–89, 36) B: 31 (–24, 119)	No copollutant model No association with PM _{2.5} . Pearson <i>r</i> = –0.14 BC, –0.22 NO ₂ , –0.07 BTEX, 0.14 cleaning product VOCs.
† Lin et al. (2011b) Beijing, China N = 8, ages 9–12 yr Daily measures for five 2-wk periods before and after Olympics. Recruitment from school.	Monitor 0.65 km from school Means across five periods before and after Olympics: 3.7–45	24-h avg 0 1	Percent change in eNO 5.5 (2.7, 8.3) 3.4 (1.4, 5.4)	Persist with: BC or PM _{2.5} Copollutant correlations NR.
† Liu et al. (2009b) , Liu (2013) Windsor, ON Oct–Dec 2005 N = 182, ages 9–14 yr. 37% ICS use. 35% beta-agonist use. Weekly measures for 4 wk. Recruited from schools. Mean 1.6 and 2.2 h/day spent outdoors for two study groups.	Two monitors averaged 99% homes within 10 km of sites Median: 4.5 95th percentile: 16	24-h avg 0 0–2 avg	Percent change eNO: 9.0 (–7.6, 29) TBARS: 28 (0.46, 63) 8-Isoprostane: 23 (3.9, 44) eNO: –5.6 (–28, 24) TBARS: 77 (31, 131) 8-Isoprostane: –0.55 (–28, 38)	TBARS, lag 0–2 avg Persists with: NO ₂ or O ₃ Does not persist with: PM _{2.5} Spearman <i>r</i> = 0.56 PM _{2.5} , 0.18 NO ₂ , –0.02 O ₃ .

Avg = average; BC = black carbon; BTEX = benzene, toluene, ethylbenzene, xylene; CI = confidence interval; EBC = exhaled breath condensate; eNO = exhaled nitric oxide; ICS = inhaled corticosteroid; ISA = Integrated Science Assessment; max = maximum; n = sample size; N = population number; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; PM_{2.5} = particulate matter with nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with nominal aerodynamic diameter less than or equal to 10 μm; *r* = correlation coefficient; SD = standard deviation; SO₂ = sulfur dioxide; TBARS = thiobarbituric acid reactive substances; VOCs = volatile organic compound.

^aEffect estimates are standardized to a 10-ppb increase in 24-h avg SO₂.

†Studies published since the 2008 ISA for Sulfur Oxides.

Table 5-11 Study-specific details from animal toxicological studies of subclinical effects underlying asthma.

Study	Species (Strain); n; Sex; Lifestage/Age or Weight	Exposure Details (Concentration; Duration)	Endpoints Examined
Li et al. (2007)	Rats (Wistar); n = 6/group; M; age NR	Sensitization by i.p. injection of 100 mg ovalbumin followed by booster injection of 10 mg ovalbumin after 7 days followed by: (1) Challenge with 1% ovalbumin aerosol for 30 min for 7 days beginning at 15 days, (2) Exposure to 2 ppm SO ₂ for 1 h/day for 7 days, or (3) SO ₂ exposure followed by ovalbumin aerosol challenge for 7 days	Endpoints examined 24 h following the last challenge BALF— <i>inflammatory cell counts</i> Lung— <i>histopathology, immunohistochemistry</i> Lung and tracheal tissue— <i>mRNA and protein levels of MUC5AC and ICAM-1</i>
Li et al. (2008)	Rats (Wistar); n = 6/group; M; age NR; 180–200 g	Sensitization by i.p. injection of 100 mg ovalbumin followed by booster injection of 10 mg ovalbumin after 7 days followed by: (1) Challenge with 1% ovalbumin aerosol for 30 min for 7 days beginning at 15 days, (2) Exposure to 2 ppm SO ₂ for 1 h/day for 7 days, or (3) SO ₂ exposure followed by ovalbumin aerosol challenge for 7 days	Endpoints examined 24 h following the last challenge BALF— <i>total and differential cell counts, EGF</i> Lung tissue— <i>histopathology</i> Lung and tracheal tissue— <i>mRNA levels of EGF, EGFR, COX-2</i> Lung tissue— <i>protein levels of EGFR, COX-2</i>
Xie et al. (2009)	Rats (Wistar); n = 6/group; M; age NR	Sensitization by i.p. injection of 100 mg ovalbumin followed by booster injection of 10 mg ovalbumin after 7 days followed by: (1) Challenge with 1% ovalbumin aerosol for 30 min for 7 days beginning at 15 days, (2) Exposure to 2 ppm SO ₂ for 1 h/day for 7 days, or (3) SO ₂ exposure followed by ovalbumin aerosol challenge for 7 days	Endpoints examined 24 h following the last challenge Lung tissue— <i>mRNA levels of p53, bax, bcl-2</i> Lung— <i>protein levels of p53, bax, bcl-2</i>

Table 5-11 (Continued): Study-specific details from animal toxicological studies of subclinical effects underlying asthma.

Study	Species (Strain); n; Sex; Lifestage/Age or Weight	Exposure Details (Concentration; Duration)	Endpoints Examined
Li et al. (2014)	Rats (Wistar); n = 6/group; M; age NR; 180–200 g	Sensitization by i.p. injection of 100 mg ovalbumin followed by booster injection of 10 mg ovalbumin after 7 days followed by: (1) Challenge with 1% ovalbumin aerosol for 30 min for 7 days beginning at 15 days, (2) Exposure to 2 ppm SO ₂ for 1 h/day for 7 days, or (3) SO ₂ exposure followed by ovalbumin aerosol challenge for 7 days	Endpoints examined BALF—inflammatory cell counts and cytokines IL-4, IFN-γ, TNF-α, IL-6 Serum—IgE Lung—histopathology Lung and tracheal tissue—mRNA and protein levels of NFκB, IκBα, IKKβ, IL-6, IL-4, TNF-α, FOXP3 EMSA NFκB binding activity

BALF = bronchoalveolar lavage fluid; bax = B-cell lymphoma 2-like protein 4; bcl-2 = B-cell lymphoma 2; COX-2 = cyclooxygenase-2; EGF = epidermal growth factor; EGFR = epidermal growth factor receptor; EMSA = electrophoretic mobility shift assay; FOXP3 = forkhead box p3; ICAM-1 = intercellular adhesion molecule 1; IFN-γ = interferon gamma; IgE = immunoglobulin E; IKKβ = inhibitor of nuclear factor kappa-B kinase subunit beta; IL-4 = interleukin-4; IL-6 = interleukin-6; IκBα = nuclear factor of kappa light polypeptide gene enhancer in B-cells inhibitor, alpha; i.p. = intraperitoneal; M = male; MUC5AC = mucin 5AC glycoprotein; n = sample size; NFκB = nuclear factor kappa-light-chain-enhancer of activated B cells; NR = not reported; p53 = tumor protein p53; SO₂ = sulfur dioxide; TNF-α = tumor necrosis factor alpha.

Repeated exposure to SO₂ promoted an allergic phenotype when ovalbumin sensitization and challenge preceded SO₂ exposure. As described in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), [Li et al. \(2007\)](#) demonstrated that rats, which were first sensitized and challenged with ovalbumin and subsequently exposed to 2 ppm SO₂ for 1 hour/day for 7 days, had an increased number of inflammatory cells in BALF and an enhanced histopathological response compared with those treated with ovalbumin or SO₂ alone. Similarly, intercellular adhesion molecule 1 (ICAM-1), a protein involved in regulating inflammation, and MUC5AC, a mucin protein, were upregulated in lungs and trachea to a greater extent in rats treated with ovalbumin and SO₂ than in those treated with ovalbumin or SO₂ alone. A follow-up study involving the same exposure regimen (2 ppm SO₂ for 1 hour) in the same allergic animal model (rats sensitized and challenge with ovalbumin) also found that repeated SO₂ exposure enhanced inflammatory and allergic responses to ovalbumin ([Li et al., 2014](#)). Numbers of eosinophils, lymphocytes, and macrophages were greater in BALF of SO₂-exposed and ovalbumin-treated animals than in animals treated only with ovalbumin. In addition, SO₂ exposure enhanced upregulation and activation of NFκB, a transcription factor involved in inflammation and upregulation of the cytokines interleukin-6 (IL-6) and interleukin-4 (IL-4) in lung tissue in this model of allergic airway disease. Furthermore, BALF levels of IL-6 and IL-4 were increased to a greater extent in SO₂-exposed and ovalbumin-treated animals compared with ovalbumin treatment alone. These results indicate that repeated SO₂ exposure enhanced activation of the nuclear factor kappa-light-chain-enhancer of activated B cells (NFκB)

inflammatory pathway and upregulation of inflammatory cytokines in ovalbumin-treated animals. Furthermore, SO₂ exposure enhanced the effects of ovalbumin on levels of interferon gamma (IFN- γ , decreased) and IL-4 (increased) in BALF and on IgE levels in serum (increased). Because levels of IL-4 are indicative of T-derived lymphocyte helper 2 (Th2) status and levels of IFN- γ are indicative of T helper 1 (Th1) status, these results suggest a shift in Th1/Th2 balance towards Th2 in rats made allergic to ovalbumin, an effect exacerbated by SO₂ exposure. These Th2-related changes are consistent with the observed increases in serum IgE and BALF eosinophils in ovalbumin-treated animals, effects which were also enhanced by SO₂ exposure. Alternatively, Th2-related changes may reflect a Type 2 immune response mediated by Group 2 innate lymphoid cells. Taken together, these results indicate that repeated exposure to SO₂ exacerbated inflammatory and allergic responses in this animal model.

Two other follow-up studies by the same laboratory examined the effects of inhaled SO₂ on the asthma-related genes encoding epidermal growth factor (EGF), epidermal growth factor receptor (EGFR), and cyclooxygenase-2 (COX-2) and on apoptosis-related genes and proteins in this same model based on sensitization with ovalbumin ([Xie et al., 2009](#); [Li et al., 2008](#)). While EGF and EGFR are related to mucus production and airway remodeling, COX-2 is related to inflammation and apoptosis and may play a role in regulating airway inflammation. SO₂ exposure enhanced the effects of ovalbumin challenge in this model, resulting in greater increases in mRNA and protein levels of EGF, EGFR, and COX-2 in the trachea compared with ovalbumin challenge alone. SO₂ exposure enhanced other effects of ovalbumin in this model, resulting in a greater decline in mRNA and protein levels of p53 and B cell lymphoma 2 like protein 4 (bax) and a greater increase in mRNA and protein levels of B cell lymphoma 2 (bcl-2) in the lungs compared with ovalbumin challenge alone. The increased ratio of bcl-2:bax, an indicator of susceptibility to apoptosis, observed following ovalbumin challenge, was similarly enhanced by SO₂. Thus, repeated exposure to SO₂ may impact numerous processes that may be involved in inflammation and/or airway remodeling in allergic airway disease.

Summary of Subclinical Effects Underlying Asthma Exacerbation

Whereas previous evidence was limited and inconsistent, recent evidence from experimental studies supports a relationship between short-term exposure to SO₂ and allergic responses related to asthma. This includes a finding of eosinophilic inflammation in individuals with asthma exposed acutely to SO₂. In addition, enhanced inflammation and allergic responses were demonstrated in animals made allergic to ovalbumin and exposed repeatedly to SO₂. Epidemiologic findings are inconsistent overall, including recent results based on SO₂ measured at or near children's schools. However, coherent with experimental studies, an SO₂-associated increase in pulmonary inflammation was

observed in a population of adults with asthma and a high prevalence of atopy. Copollutant confounding is not addressed in these results, but the evidence from animal toxicological studies provides some biological plausibility for an effect of SO₂ exposure, particularly because effects in rats were shown to occur with repeated exposures and 24 hours after exposure ended. The evidence for SO₂-related allergic inflammation also supports evidence across disciplines for SO₂ effects on asthma symptoms, hospital admissions, and ED visits, as well as lung function decrements in people with asthma.

Summary of Asthma Exacerbation

The 2008 ISA for Sulfur Oxides did not explicitly draw a conclusion about a relationship between short-term SO₂ exposure and asthma exacerbation but described strong support from controlled human exposure studies for SO₂-induced lung function decrements and increases in respiratory symptoms in adults with asthma when ventilation rates were increased. Such effects in adolescents with asthma are less clear due to a paucity of data, but effects appear similar to adults. There are no laboratory studies of children exposed to SO₂; however, a number of studies have assessed airway responsiveness of children and adults exposed to the bronchoconstrictive stimuli methacholine. Based largely on those studies, school-aged children (~5–11 years of age), particularly boys and perhaps obese children, might be expected to have greater responses (i.e., larger decrements in lung function) following exposure to SO₂ than adolescents and adults.

In adults with asthma, short-term exposures for 5–10 minutes to 0.2–0.3 ppm SO₂ resulted in 5–30% of exercising individuals experiencing moderate or greater decrements (i.e., ≥15% decrease in FEV₁ or ≥100% increase in sRaw; [Table 5-2](#)). Decrements in FEV₁ at 0.3 ppm SO₂ were statistically significant in responsive individuals (defined as those having an FEV₁ decrease of ≥15% after exposure to 0.6 or 1.0 ppm SO₂; [Table 5-3](#)). At concentrations greater than or equal to 0.4 ppm, 20–60% of people with asthma experienced SO₂-induced decrements in lung function, which were frequently accompanied by respiratory symptoms. There is a clear concentration-response relationship for exposures to SO₂ between 0.2 and 1.0 ppm, both in terms of increasing severity of effect and percentage of asthmatics affected. These concentrations are in the range of the highest 5-minute ambient SO₂ concentrations in some U.S. cities during 2013–2015 ([Section 2.5.2.2](#)).

Epidemiologic evidence generally supports SO₂-associated increases in asthma hospital admissions and ED visits, particularly in children ([Figure 5-3](#)), and respiratory symptoms in children with asthma ([Figure 5-2](#); [Table 5-8](#)). Epidemiologic evidence is inconsistent for SO₂ associations with lung function decrements in adults and children with asthma ([Tables 5-6](#) and [5-7](#)). For the limited results from previous epidemiologic and controlled

human exposure studies on airway responsiveness (i.e., response to methacholine), an independent effect of SO₂ is unclear. Two controlled human exposure studies demonstrated increased airway responsiveness to subsequent allergen challenge for at least 48 hours following SO₂ exposure in combination with a copollutant (i.e., NO₂). Most epidemiologic studies estimated SO₂ exposure from fixed-site monitors without fully characterizing the spatial and temporal variation in SO₂ exposure. A few recent studies aimed to address the uncertainty in exposure estimates and observed asthma-related effects in association with SO₂ measured or modeled at or near schools or homes. Studies did not statistically correct for exposure measurement error. As in the 2008 ISA for Sulfur Oxides, copollutant confounding is unresolved in the epidemiologic evidence. Many recent studies continue to indicate that SO₂ associations with asthma hospital admissions and ED visits remain relatively unchanged in magnitude in copollutant models, but SO₂ associations with asthma symptoms and pulmonary inflammation often did not persist after adjustment for PM_{2.5}, EC/BC, or NO₂. The role of SO₂ in ambient multipollutant mixtures is not clearly elucidated. Controlled human exposure studies show asthma-related effects when SO₂ exposure occurs with O₃ or NO₂, and limited epidemiologic examination shows associations for multipollutant mixtures that contain SO₂. However, associations for mixtures containing SO₂ are similar to or less than the sum of single-pollutant effect estimates for SO₂, CO, NO₂, PM₁₀, or PM_{2.5}, indicating an overlap in associations for copollutants.

Expanded evidence for SO₂-induced allergic inflammation supports an effect of SO₂ exposure on asthma exacerbation. Epidemiologic findings of SO₂-associated increases in pulmonary inflammation are inconsistent, but enhanced allergic inflammation and allergic responses are demonstrated in a previous controlled human exposure study of adults with asthma and in multiple recent studies from a single laboratory in rats made allergic to ovalbumin and exposed repeatedly to 2 ppm SO₂. These findings provide some support for the epidemiologic associations for SO₂ with decreased lung function as well as increased airway responsiveness, respiratory symptoms, and pulmonary inflammation observed in most studies in populations of children and adults with asthma and a high prevalence of atopy, or with AHR and elevated IgE. However, it is not clear whether risk of SO₂-related respiratory effects differs in individuals with atopic asthma. Differential effects by asthma severity or other asthma phenotypes are not well characterized.

Much of the epidemiologic evidence for SO₂-associated asthma exacerbation is for 24-h avg SO₂ concentrations. Although 24-h avg and 1-h max SO₂ concentrations are correlated at the same monitor, it is not clear whether this correlation applies across a community. Some recent studies add evidence of association for asthma symptoms and ED visits with increases in 1-h max SO₂ concentrations, including SO₂ measured at schools. For lung function decrements, pulmonary inflammation, and asthma hospital

admission and ED visit studies, several results indicate associations for 3- or 4-day avg SO₂ concentrations. The evidence for enhanced allergic inflammation, which is seen after repeated 2 ppm SO₂ exposures and 24 hours after exposure ended, somewhat supports the biological plausibility of epidemiologic associations with asthma-related outcomes. Moreover, controlled human exposure studies clearly demonstrate that SO₂ exposures of 0.2–0.6 ppm can induce effects related to asthma exacerbation.

5.2.1.3 Allergy Exacerbation

The evidence described in the preceding section for SO₂ and allergen coexposure enhancing inflammation in rodent models of allergic airway disease indicates that SO₂ exposure may increase the sensitivity of people with allergic asthma to an allergen. This evidence also suggests the potential for SO₂ exposure to affect respiratory responses in people with allergy but not asthma. The 2008 ISA for Sulfur Oxides did not make distinct statements about a relationship with SO₂ exposure, but relevant epidemiologic studies had inconsistent findings. Recent epidemiologic evidence is also uncertain, including that for school SO₂ measurements.

Lung Function in Populations with Allergy

Previous epidemiologic studies examining children or adults with high serum IgE levels, but without AHR, did not find associations between short-term increases in ambient SO₂ concentration and decreases in lung function ([Boezen et al., 2005](#); [Boezen et al., 1999](#)). The same studies observed associations for groups with high serum IgE and AHR. Previous findings were based on 24-h avg SO₂ measured at a single site in each city. The recent study improves on previous studies, measuring SO₂ at children's schools ([Correia-Deur et al., 2012](#)). Also, the 2-h avg metric used in this study is more comparable to the exposure durations examined in experimental studies. In this group of children with allergy in São Paulo, Brazil, SO₂ was not associated with PEF [−0.82% (95% CI: −1.9, 0.31) per 10-ppb increase in 2-h avg SO₂]. Results were similar for allergy defined by high serum IgE levels alone like previous studies and by multiple criteria (i.e., high IgE levels, positive skin prick test, and high blood eosinophil levels). There was evidence for an association among all children (with and without allergy), but that was attenuated in copollutant models with PM₁₀, NO₂, or CO. Correlations with SO₂ were not reported.

Respiratory Symptoms and Physician Visits in Populations with Allergy

Limited to epidemiologic studies, evidence for an association between short-term SO₂ exposure and allergy symptoms is inconsistent. Nonspecific upper and lower respiratory symptoms were examined in children and adults with high IgE levels but without AHR, and there were no associations with SO₂ exposure ([Boezen et al., 2005](#); [Boezen et al., 1999](#)). For symptoms specific to allergy, [Villeneuve et al. \(2006b\)](#) observed an SO₂-associated increase in physician visits for allergic rhinitis in older adults. Recent findings for allergic rhinitis or eczema in children are mixed. However, inference about an SO₂ effect is weak both for results indicating an association ([Kim et al., 2016a](#)) and results not indicating an association ([Annesi-Maesano et al., 2012](#); [Linares et al., 2010](#)). Limitations include cross-sectional design ([Annesi-Maesano et al., 2012](#); [Linares et al., 2010](#)), analysis of a multipollutant model with NO₂, O₃, PM₁₀, and pollen ([Kim et al., 2016a](#); [Annesi-Maesano et al., 2012](#)), lack of consideration of confounding by meteorological factors ([Kim et al., 2016a](#)), or inclusion of children with and without allergy in analysis of eczema ([Linares et al., 2010](#)). For results supporting a relationship with allergy symptoms, associations were observed with same-day (lag 0) 24-h avg SO₂ concentrations. These concentrations were from a single monitor in the city, and information was not reported on the extent to which the measurements represented the spatiotemporal variability in SO₂ concentrations in the study area. Associations were observed with copollutants such as NO₂, PM₁₀, and BS, although these results were inconsistent as well ([Villeneuve et al., 2006b](#); [Boezen et al., 2005](#); [Boezen et al., 1999](#)). Correlations with SO₂ concentrations were not reported, and copollutant models were not analyzed. Thus, the extent to which the supporting findings may indicate an independent association for SO₂ is unclear.

Subclinical Effects Underlying Allergy Exacerbation

In addition to the animal toxicological evidence for SO₂-enhanced allergic inflammation, a previous epidemiologic study of a population of children in which 8% had asthma and 18% had atopy found an SO₂-associated decrease in blood eosinophil number, which was presumed to reflect increased recruitment to the airways ([Soyseth et al., 1995](#)). Exposure assessment from a monitor 2 km from most subjects' homes is an uncertainty, as is confounding by PM. The study was conducted in a European city with an aluminum smelter that emitted SO₂ and PM, and PM was not examined for association with eosinophils.

Summary of Allergy Exacerbation

Epidemiology studies found little evidence of a relationship between short-term exposure to SO₂ and lung function, respiratory symptoms, or physician visits in populations with allergy. Animal toxicological studies reported that SO₂ exposure enhanced allergic inflammation ([Section 5.2.1.2](#)).

5.2.1.4 Chronic Obstructive Pulmonary Disease Exacerbation

COPD is a lung disease characterized by destruction of alveolar tissue, airway remodeling, and airflow limitation. Reduced airflow can decrease lung function, and clinical symptoms demonstrating exacerbation of COPD include cough, dyspnea, sputum production, and shortness of breath. Severe exacerbation can lead to hospital admissions or ED visits. This spectrum of outcomes has been evaluated in relation to short-term SO₂ exposure, and evidence across outcomes and disciplines is inconsistent. This applies to the small body of studies available for the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) as well as the few available recent studies. Recent findings come from epidemiologic studies, and most are for hospital admissions and ED visits.

Lung Function and Respiratory Symptoms

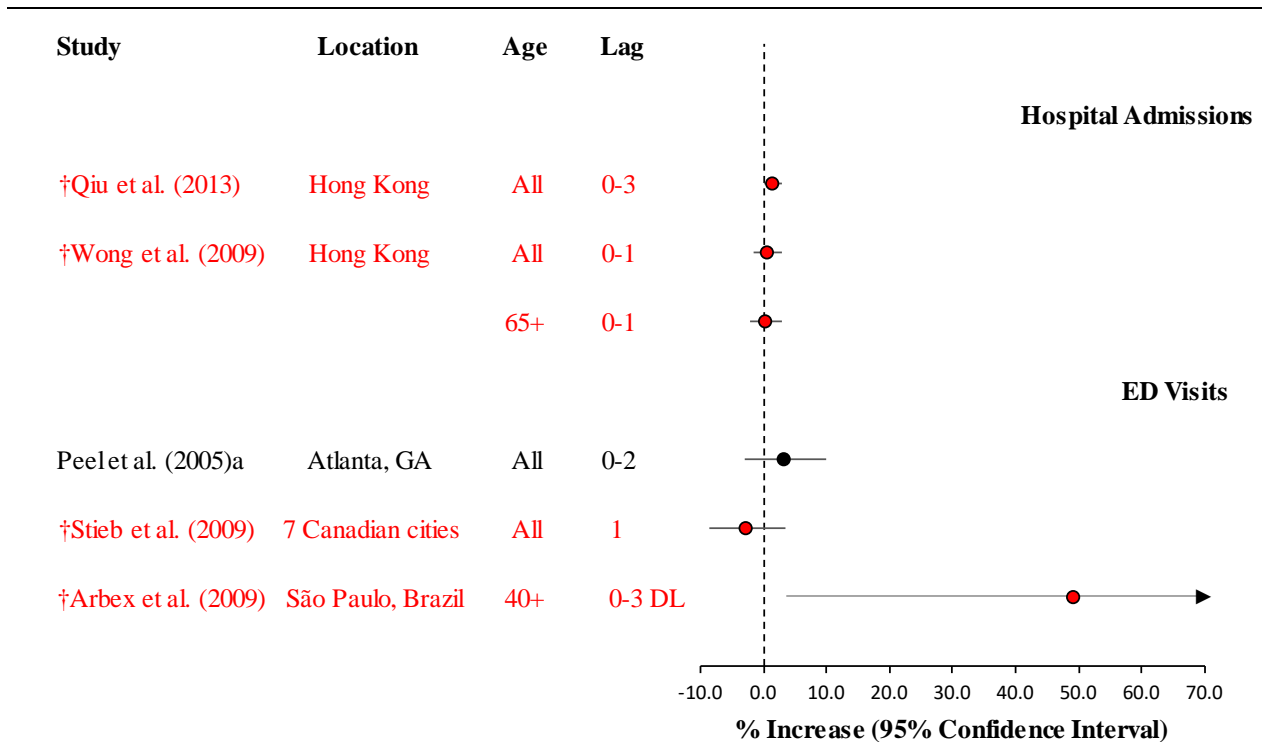
Evidence from a controlled human exposure study and epidemiologic studies does not support an effect of SO₂ exposure on lung function in adults with COPD. Recent epidemiologic studies add information on respiratory symptoms and mostly do not indicate an association with ambient SO₂ concentrations.

[Linn et al. \(1985a\)](#) reported that a 15-minute exposure to 0.4 and 0.8 ppm SO₂ had no effect on lung function in older adults with physician-diagnosed COPD. These adults were much older than the adults with asthma ([Table 5-2](#)) or healthy adults ([Table 5-15](#)) examined in controlled human exposure studies. Also, the level of exercise in adults with COPD ($V_E = 18$ L/minute) was lower than that of individuals with asthma, which effectively lowers the SO₂ dose delivered to the lungs ([Section 4.2.2](#)). Neither the previous nor recent epidemiologic study observed SO₂-associated decrements in lung function in adults with COPD ([Peacock et al., 2011](#); [Harre et al., 1997](#)). Both studies estimated SO₂ exposure from a fixed-site monitor(s), and examined 24-h avg concentrations lagged 1 day. Whereas previous results were based on a multipollutant model (with PM₁₀, NO₂, O₃), results of the more recent study were based on a single-pollutant model. Associations were imprecise with wide 95% CIs [e.g., 0.31 L/minute (95% CI: -0.10, 0.72) change in PEF per 10-ppb increase in SO₂ and OR 1.01 (95% CI: 0.89, 1.15) for

PEF decrement greater than 20%] ([Peacock et al., 2011](#)). Mean and 75th percentile SO₂ concentrations were 7.5 and 9.3 ppb, respectively. SO₂ mostly was not associated with dyspnea, sputum changes, wheeze/tight chest, or other respiratory symptoms ([Wu et al., 2016](#); [Peacock et al., 2011](#)). [Wu et al. \(2016\)](#) examined a period of higher SO₂ concentration (median 17 ppb and 75th percentile 27 ppb) and observed dyspnea to increase with an increase in 3- to 6-day avg SO₂ (OR: 1.88 [95% CI: 1.06, 3.34] per 10-ppb increase in 3-day avg SO₂). However, there was a wide range of distance from subjects to the monitor (1.6–8.8 km), and associations also were observed with moderately correlated ($r = 0.51$ – 0.68) PM_{2.5}, PM₁₀, and NO₂.

Hospital Admissions and Emergency Department Visits

Of the studies evaluated in the 2008 SO_x ISA, only one U.S. or Canadian-based study examined the association between short-term SO₂ exposure and COPD hospital admissions or ED visits ([Figure 5-7](#)). Recent studies add to the initial evidence, which generally indicates no association between short-term SO₂ exposures and COPD hospital admissions and ED visits. Additionally, most studies averaged SO₂ concentrations over multiple monitors and examined 24-h avg exposure metrics, which, may not adequately capture the spatial and temporal variability in SO₂ concentrations ([Section 3.4.2](#)). For each of the studies evaluated in this section, [Table 5-12](#) presents the air quality characteristics of each city or across all cities, the exposure assignment approach used, and information on copollutants examined in each COPD hospital admission and ED visit study. Other recent studies of COPD hospital admissions and ED visits are not the focus of this evaluation because of various study design issues, as initially detailed in [Section 5.2.1.2](#), but the full list of these studies, as well as study-specific details, can be found in Supplemental Table 5S-6 ([U.S. EPA, 2017c](#)).



ED = emergency department; ISA = Integrated Science Assessment.

a = studies that used a 1-h max exposure metric.

Note: † and red text/symbols = recent studies published since the 2008 ISA for Sulfur Oxides; black text/symbols = U.S. and Canadian studies evaluated in the 2008 ISA for Sulfur Oxides. Corresponding quantitative results are reported in Supplemental Table 5S-7 ([U.S. EPA, 2017c](#)).

Figure 5-7 Percent increase in chronic obstructive pulmonary disease hospital admissions and emergency department visits from U.S. and Canadian studies evaluated in the 2008 SO_x ISA and recent studies in all-year analyses for a 10-ppb increase in 24-h avg or 40-ppb increase in 1-h max sulfur dioxide concentrations.

Table 5-12 Study-specific details and mean and upper percentile concentrations from chronic obstructive pulmonary disease hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location Years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examined
Hospital admissions						
†Qiu et al. (2013b) , Ko et al. (2007a)	Hong Kong, China (1998–2007)	Average of SO ₂ concentrations from 10 monitoring stations	24-h avg	7.4	NR	Correlations (<i>r</i>): O ₃ : 0.173 Copollutant models: PM ₁₀
†Wong et al. (2009)	Hong Kong, China (1996–2002)	Average of SO ₂ concentrations from eight monitoring stations	24-h avg	6.8	75th: 8.4 Max: 41.8	Correlations (<i>r</i>): NR Copollutant models: none
ED visits						
Peel et al. (2005)	Atlanta, GA (1993–2000)	Average of SO ₂ concentrations across monitors for several monitoring networks	1-h max	16.5	90th: 39.0	Correlations (<i>r</i>): PM _{2.5} : 0.17 PM ₁₀ : 0.20 PM _{10-2.5} : 0.21 UFP: 0.24 PM _{2.5} water soluble metals: 0.00 PM _{2.5} sulfate: 0.08 PM _{2.5} acidity: -0.03 PM _{2.5} OC: 0.18 PM _{2.5} EC: 0.20 Oxygenated HCs: 0.14 O ₃ : 0.19 CO: 0.26 NO ₂ : 0.34 Copollutant models: none

Table 5-12 (Continued): Study specific details and mean and upper percentile concentrations from chronic obstructive pulmonary disease hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location Years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile Concentrations ppb	Copollutants Examined
† Stieb et al. (2009)	Seven Canadian cities (1992–2003)	Average SO ₂ concentrations across all monitors in each city. Number of SO ₂ monitors in each city ranged from 1–11.	24-h avg	2.6–10.0	75th: 3.3–13.4	Correlations (<i>r</i>) only reported by city and season. Copollutant models: none
† Arbex et al. (2009)	São Paulo, Brazil (2001–2003)	Average of SO ₂ concentrations across 13 monitoring stations	24-h avg	5.3	75th: 6.6 Max: 16.4	Correlations (<i>r</i>): PM ₁₀ : 0.77 NO ₂ : 0.63 CO: 0.52 Copollutant models: none

Avg = average; CO = carbon monoxide; EC = elemental carbon; HC = hydrocarbon; ISA = Integrated Science Assessment; max = maximum; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OC = organic carbon; PM₁₀ = particulate matter with nominal aerodynamic diameter less than or equal to 10 µm; PM_{2.5} = particulate matter with nominal aerodynamic diameter less than or equal to 2.5 µm; PM_{10-2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm and greater than a nominal 2.5 µm; *r* = correlation coefficient; SO₂ = sulfur dioxide; UFP = ultrafine particle.

† = Studies published since the 2008 ISA for Sulfur Oxides.

Hospital Admissions

Of the studies evaluated in the 2008 SO_x ISA, relatively few examined the association between short-term SO₂ exposure and COPD hospital admissions, and evidence of an association was inconsistent across studies. Although several recent studies assessed the relationship between short-term SO₂ exposures and COPD hospital admissions, the overall body of evidence remains limited.

[Wong et al. \(2009\)](#) in a study that examined the potential modification of the relationship between air pollution and respiratory-related hospital admissions by influenza, also focused on cause-specific respiratory hospital admissions, including COPD. When focusing on the baseline effect of short-term SO₂ exposures on COPD hospital admissions, the authors found limited evidence of an association at lag 0–1 days for a 10-ppb increase in 24-h avg SO₂ concentrations in analyses of both all ages [0.8% (95% CI: –1.5, 3.1)] and individuals over the age of 65 [0.5% (95% CI: –2.0, 3.0)].

In an additional study conducted in Hong Kong, [Qiu et al. \(2013b\)](#) focused on whether there is evidence of modification of the air pollution-COPD hospital admissions relationship by season and humidity. Compared to [Wong et al. \(2009\)](#), [Qiu et al. \(2013b\)](#) included 5 additional years of recent data through the year 2007. In single-pollutant models focusing on the association between short-term SO₂ exposures and COPD hospital admissions, for a multiday lag of 0–3 days, the authors reported a 1.6% increase (95% CI: 0.1, 3.1) for a 10-ppb increase in 24-h avg SO₂ concentrations. The magnitude of the SO₂ association was found to differ between [Qiu et al. \(2013b\)](#) and [Wong et al. \(2009\)](#), but the reason for the difference remains unclear, considering that similar data sources were used in each study. It is important to note that neither study conducted copollutant analyses for the entire study duration nor provided detailed information on the correlation between the air pollutants examined to help in assessing whether SO₂ has an independent effect on COPD hospital admissions.

Emergency Department Visits

The 2008 SO_x ISA identified relatively few studies that examined the association between short-term SO₂ exposure and COPD ED visits, and across studies there was inconsistent evidence of an association. Although recent studies continued to assess the relationship between short-term SO₂ exposures and COPD ED visits, the overall body of evidence remains limited.

In the seven Canadian cities study discussed previously, and consistent with the asthma ED visits results, [Stieb et al. \(2009\)](#) did not find any evidence of associations between 24-h avg SO₂ and COPD ED visits at single-day lags of 0 to 2 days. Additionally, there was no evidence of consistent associations between any pollutant and COPD ED visits at subdaily time scales (i.e., 3-h avg of ED visits vs. 3-h avg pollutant concentrations).

[Arbex et al. \(2009\)](#) also examined the association between COPD and several ambient air pollutants, including SO₂, in a single-city study conducted in São Paulo, Brazil for individuals over the age of 40 years. The authors examined associations between short-term SO₂ exposures and COPD ED visits at single-day lags (0 to 6 days) and in a polynomial distributed lag model (0–6 days). The authors found evidence that the magnitude of the association was larger at multiday lags compared to single-day lags, with the lag of 0–3 days from the distributed lag model [49.4% (95% CI: 4.1, 113.7) for a 10-ppb increase in 24-h avg SO₂ concentrations] most representative of the pattern of associations across single-day lags. Although the 0–6-day distributed lag model had the largest risk estimate, it was not supported by the single-day lag results that showed the strongest associations at lags of 0 and 1 day. It is important to note that [Arbex et al. \(2009\)](#) did not conduct copollutant analyses, but unlike correlations with SO₂ observed in other locations, SO₂ was highly correlated with PM₁₀ ($r = 0.77$) and moderately

correlated with NO₂ ($r = 0.63$) and CO ($r = 0.52$) in this study. The results of [Arbex et al. \(2009\)](#) provide evidence of a potentially prolonged SO₂ effect on COPD ED visits; however, the results should be viewed with caution because effect estimates are not precise (i.e., wide confidence intervals), time series is short, and copollutant confounding is possible.

Seasonal Analyses

Traditionally, epidemiologic studies have examined potential seasonal differences in associations by stratifying by season. In the study of air pollution and COPD hospital admissions in Hong Kong, [Qiu et al. \(2013b\)](#) examined potential seasonal differences in associations by this traditional approach but also examined whether the combination of season and humidity modifies the air pollution-health effect association. In seasonal analyses, the authors found a stronger association at lag 0–3 for a 10-ppb increase in 24-h avg SO₂ concentrations during the cool season (November–April) [2.7% (95% CI: 0.5, 4.9)] compared to the warm season (May–October) [0.6% (95% CI: –1.1, 2.3)]. [Qiu et al. \(2013b\)](#) then examined whether the seasonal differences in associations observed were due to low humidity days (i.e., relative humidity <80%) or high humidity days (i.e., relative humidity ≥80%) by examining the interaction between the various combinations of season and humidity. When focusing on the combined effect of season and humidity, SO₂ concentrations were found to be highest on days with low humidity in both seasons. In the warm season, there was no evidence of an association regardless of whether the interaction between season and low or high humidity days were examined. In the cold season, at lag 0–3 for a 10-ppb increase in 24-h avg SO₂ concentrations, [Qiu et al. \(2013b\)](#) reported the strongest association during days with low humidity [5.3% (95% CI: 2.4, 8.3)] compared to high humidity [0.5% (95% CI: –2.6, 3.7)], suggesting that the combination of season and humidity plays a role in the relationship between air pollution and health effects. However, when examining copollutant models with PM₁₀, associations in all season and humidity combinations were attenuated, with only the association in the cool season and low humidity combination remaining positive, albeit with large uncertainty estimates [0.8% (95% CI: –2.1, 3.9); lag 0–3 for a 10-ppb increase in 24-h avg SO₂ concentrations]. The results from [Qiu et al. \(2013b\)](#) are consistent with evidence from controlled human exposure studies demonstrating that SO₂ responses are exacerbated in colder and dryer conditions ([Section 5.2.1.2](#)). However, these studies focused on lung function changes in people with asthma, and it is unclear how these results correspond to results from an epidemiologic study of COPD hospital admissions. Additionally, it is important to note the potential influence of geographic location on the results from studies that examine the seasonal patterns of associations.

Lag Structure of Associations

Only a limited number of studies examined the lag structure of associations for SO₂-related COPD hospital admissions and ED visits. In the examination of air pollution and COPD hospital admissions in Hong Kong, [Qiu et al. \(2013b\)](#) conducted analyses to evaluate associations with SO₂ at both single-day and multiday lags of 0–3 days. The authors found the strongest evidence for an SO₂-COPD hospital admission association at a multiday lag of 0–3 days, with additional evidence of positive associations at single-day lags of 1 day and 3 days.

[Arbex et al. \(2009\)](#), when examining associations between SO₂ exposure and COPD ED visits in São Paulo, Brazil, focused on both single-day lags (0 to 6 days) and a polynomial distributed lag (0–6 day) model. The authors found evidence that the magnitude of the association was larger at multiday lags compared to single-day lags, and the magnitude of the association increased as the number of lag days examined increased, specifically across lags of 0–1, 0–2, and 0–5 days. However, the 0–5-day distributed lag model results were not supported by the single-day lag results, which indicated that the effect of SO₂ on COPD ED visits was rather immediate, occurring in the range of lag 0 and 1 days. Collectively, the results of [Qiu et al. \(2013b\)](#) and [Arbex et al. \(2009\)](#) provide initial evidence suggesting a potential prolonged effect of SO₂ on COPD hospital admissions and ED visits. However, the collective evidence indicating a potential association between short-term SO₂ exposures and COPD hospital admissions and ED visits remains relatively small.

Summary of Chronic Obstructive Pulmonary Disease Exacerbation

Across disciplines and outcomes, evidence from previous and recent studies does not clearly support a relationship between short-term SO₂ exposure and COPD exacerbation. The evidence base is much smaller than that for asthma exacerbation and mostly comprises epidemiologic studies. Neither the single controlled human exposure study nor the few epidemiologic studies indicate SO₂-related lung function changes in adults with COPD, and recent epidemiologic studies mostly reported no association with an array of respiratory symptoms, including sputum changes and dyspnea, which are characteristic of COPD exacerbation. Evidence is similarly inconsistent for association between short-term increases in ambient SO₂ concentration and hospital admissions and ED visits for COPD ([Figure 5-7](#)). Hospital admissions, ED visits, lung function, and symptoms were examined in relation to 24-h avg SO₂ concentrations, but an association was not observed with 1-h max SO₂ either. The supporting evidence is limited largely to an association of COPD hospital admissions and ED visits with same-day and 4-day avg SO₂ concentrations. All epidemiologic studies estimated SO₂ exposure from fixed-site

monitors. SO₂ generally has low to moderate spatial correlations across urban geographical scales, and the potential error in the exposure estimates in adequately representing the spatiotemporal variability is uncharacterized in the evidence ([Section 3.4.2.2](#)). The uncertainty in exposure estimates especially applies to 1-h max SO₂. COPD hospital admissions were associated with PM₁₀, NO₂, and O₃. PM₁₀ was highly correlated with SO₂ ($r = 0.77$) ([Arbex et al., 2009](#)), or when analyzed in a copollutant model, attenuated the SO₂ association and produced wide 95% CIs ([Qiu et al., 2013b](#)). The copollutant model results have unclear implication as a result of uncertainty in the exposure estimates and the unreported SO₂-PM₁₀ correlation. Overall, the studies do not consistently report positive associations between short-term SO₂ exposure and COPD exacerbation. The consideration of copollutant confounding is limited so the independent effect of SO₂ is unclear.

5.2.1.5 Respiratory Infection

The respiratory tract is protected from exogenous pathogens and particles through various lung host defense mechanisms that include mucociliary clearance, phagocytosis by alveolar macrophages, and innate and adaptive immunity. There is a paucity of evidence related to host defense from animal toxicological experiments using ambient-relevant concentrations of SO₂. Several studies of short-term exposure to SO₂ were reported in the 1982 AQCD ([U.S. EPA, 1982a](#)) and discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Findings of short-term studies included some effects of 0.1–1 ppm SO₂ on the clearance of labeled particles. No new animal studies of the effects of SO₂ exposure on lung host defense have been conducted since the previous review. A small number of previous epidemiologic studies reported SO₂-associated increases in respiratory infections as self-reported or indicated by hospital admissions and ED visits. However, many results were noted as being unreliable because they were based on statistical methods prone to bias.

Recent contributions to the evidence are limited to epidemiologic studies, and the evaluation of this evidence focuses on hospital admissions and ED visits. There are recent studies of self-reported infections, and they inconsistently show associations with ambient SO₂ concentrations, [Supplemental Figure 5S-2 and Table 5S-8 ([U.S. EPA, 2017c](#))]. Results based on school or home SO₂ exposure estimates are limited by their cross-sectional design or examination of nonspecific symptoms such as fever. Other studies do not provide insight over studies of hospital admissions and ED visits on issues such as exposure measurement error, copollutant confounding, or potentially relevant exposure durations and concentrations. Recent studies of respiratory infection hospital

admissions and ED visits provide some evidence for association with ambient SO₂ concentrations. However, copollutant confounding remains an uncertainty.

Hospital Admissions and Emergency Department Visits

The 2008 SO_x ISA contained limited evidence of an association between short-term SO₂ concentrations and respiratory conditions other than asthma or COPD including respiratory tract infections and pneumonia. Of these studies, only one study was conducted in the U.S. or Canada [i.e., ([Peel et al., 2005](#))]. Recent studies have examined a variety of outcomes indicative of respiratory infection; however, none have examined the same respiratory infection outcome. Additionally, most studies averaged SO₂ concentrations over multiple monitors and examined 24-h avg exposure metrics, which may not adequately capture the spatial and temporal variability in SO₂ concentrations ([Section 3.4.2](#)). For each of the studies evaluated in this section, [Table 5-13](#) presents the air quality characteristics of each city, or across all cities, the exposure assignment approach used, and information on copollutants examined in each respiratory infection hospital admission and ED visit study. Other recent studies of respiratory infection hospital admissions and ED visits are not the focus of this evaluation because of various study design issues, as initially detailed in [Section 5.2.1.2](#), but the full list of these studies, as well as study specific details, can be found in Supplemental Table 5S-6 ([U.S. EPA, 2017c](#)).

Table 5-13 Study-specific details and mean and upper percentile concentrations from respiratory infection hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location (Years)	Type of Visit (ICD 9/10)	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb	Copollutants Examined
Hospital admissions							
†HEI (2012) Mehta et al. (2013)	Ho Chi Minh City, Vietnam (2003–2005)	Acute lower respiratory infection (J13–16, 18, 21)	Average of SO ₂ concentrations across nine monitors	24-h avg	8.2	Max: 30.5	Correlations (<i>r</i>): Dry season: PM ₁₀ : 0.32 O ₃ : 0.19 NO ₂ : 0.29 Rainy season: PM ₁₀ : 0.36 O ₃ : 0.65 NO ₂ : 0.01 Copollutant models: NO ₂ , PM ₁₀ , O ₃
†Ségala et al. (2008)	Paris, France (1997–2001)	Bronchiolitis	Average SO ₂ concentrations across 30 monitors	24-h avg	4.0	Max: 27.4	Correlations (<i>r</i>): BS: 0.76 PM ₁₀ : 0.73 NO ₂ : 0.78 Copollutant models: none

Table 5-13 (Continued): Study specific details and mean and upper percentile concentrations from respiratory infection hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location (Years)	Type of Visit (ICD 9/10)	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb	Copollutants Examined
ED visits							
Peel et al. (2005)	Atlanta, GA (1993–2000)	Pneumonia (480–486)	Average of SO ₂ concentrations from monitors for several monitoring networks	1-h max	16.5	90th: 39.0	Correlations (<i>r</i>): PM _{2.5} : 0.17 PM ₁₀ : 0.20 PM _{10-2.5} : 0.21 UFP: 0.24 PM _{2.5} water soluble metals: 0.00 PM _{2.5} sulfate: 0.08 PM _{2.5} acidity: -0.03 PM _{2.5} OC: 0.18 PM _{2.5} EC: 0.20 Oxygenated HCs: 0.14 O ₃ : 0.19 CO: 0.26 NO ₂ : 0.34 Copollutant models: none
†Stieb et al. (2009)	Seven Canadian cities (1992–2003)	Respiratory infection (464, 466, 480–487)	Average SO ₂ concentrations across all monitors in each city. Number of SO ₂ monitors in each city ranged from 1–11.	24-h avg	2.6–10.0	75th: 3.3–13.4	Correlations (<i>r</i>) only reported by city and season. Copollutant models: none

Table 5-13 (Continued): Study specific details and mean and upper percentile concentrations from respiratory infection hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location (Years)	Type of Visit (ICD 9/10)	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb	Copollutants Examined
†Ségala et al. (2008)	Paris, France (1997–2001)	Bronchiolitis	Average SO ₂ concentrations across 30 monitors	24-h avg	4.0	Max: 27.4	Correlations (<i>r</i>): BS: 0.76 PM ₁₀ : 0.73 NO ₂ : 0.78 Copollutant models: none
†Zemek et al. (2010)	Edmonton, AB (1992–2002)	Otitis media (382.9)	Average of SO ₂ concentrations across three monitors	24-h avg	All-year: 2.6 Warm (Apr–Sep): 2.1 Cold (Oct–Mar): 3.1	All-year 75th: 3.5	Correlations (<i>r</i>): NR Copollutant models: none
Outpatient and physician visits							
†Sinclair et al. (2010)	Atlanta, GA (1998–2002)	Upper respiratory infection Lower respiratory infection	SO ₂ concentrations collected as part of AIREs at SEARCH Jefferson Street site	1-h max	1998–2000: 19.3 2000–2002: 17.6 1998–2002: 18.3	NR	Correlations (<i>r</i>): NR Copollutant models: none

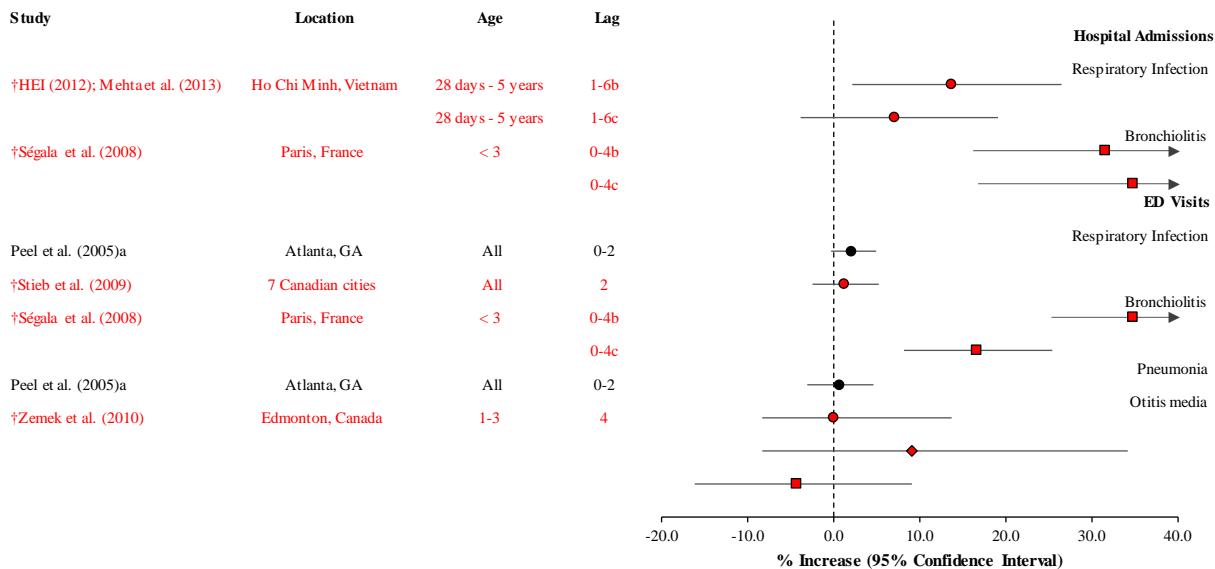
Avg = average; AIREs = Aerosol Research Inhalation Epidemiology Study; BS = black smoke; CO = carbon monoxide; EC = elemental carbon; HC = hydrocarbon; ICD = International Classification of Diseases; ISA = Integrated Science Assessment; max = maximum; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OC = organic carbon; PM_{2.5} = in general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = in general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; PM_{10-2.5} = in general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm and greater than a nominal 2.5 μm; *r* = correlation coefficient; SEARCH = Southeast Aerosol Research Characterization; SO₂ = sulfur dioxide; UFP = ultrafine particle.

† = studies published since the 2008 ISA for Sulfur Oxides.

Hospital Admissions

Although recent studies have continued to examine the association between short-term SO₂ exposures and respiratory infection hospital admissions, the overall evidence remains limited, primarily due to the variety of respiratory infection outcomes examined. In a study conducted in Ho Chi Minh City, Vietnam [Mehta et al. \(2013\)](#) and [HEI \(2012\)](#) examined the association between short-term air pollution exposures and pediatric (ages 28 days–5 years) hospital admissions for acute lower respiratory infections (ALRI, including bronchiolitis and pneumonia). In a time-stratified, case-crossover analysis focusing only on the average of a 1–6 day lag, the study authors reported a positive association, with large uncertainty estimates, between SO₂ and ALRI hospital admissions in the all-year analysis [7.0% (95% CI: –3.0, 19.1) for a 10-ppb increase in 24-h avg SO₂ concentrations]. A larger association was observed in the time-series analysis ([HEI, 2012](#)) ([Figure 5-8](#)). When examining copollutant models with PM₁₀ and O₃, SO₂ associations increased slightly, with the percent increase ranging from 7.5–8.0%, respectively. However, in models with NO₂, the SO₂ association was attenuated, but remained positive [4.9% (95% CI: –6.0, 17.0) for a 10-ppb increase in 24-h avg SO₂ concentrations].

In another study that also examined respiratory infections (i.e., bronchiolitis) in children, [Ségala et al. \(2008\)](#) focused on associations with winter (October–January) air pollution because that is the season when respiratory syncytial virus (RSV) activity peaks. [Ségala et al. \(2008\)](#) hypothesized that air pollution exposures may increase the risk of respiratory infections, including bronchiolitis due to RSV. Focusing on children <3 years of age in Paris, France, the study authors conducted a bidirectional case-crossover analysis along with a time-series analysis to examine air pollution associations with bronchiolitis hospital admissions and ED visits (see ED visits section below). Although the authors specified that the bidirectional case-crossover approach was used to “avoid time-trend bias,” it must be noted that the bidirectional approach has been shown to bias results ([Ségala et al., 2008](#); [Levy et al., 2001](#)). In the case-crossover analysis, SO₂ was associated with bronchiolitis hospital admissions at lag 0–4 days for a 10-ppb increase in 24-h avg SO₂ concentrations [34.8% (95% CI: 19.5, 47.8)] with a similar risk estimate observed for the time-series analysis [31.6% (95% CI: 13.7, 51.2)]. Although a positive association was observed, the authors did not conduct copollutant analyses. This omission complicates the interpretation of the results because SO₂ was highly correlated with the other pollutants examined, with correlations ranging from $r = 0.73$ – 0.87 .



ED = emergency department; ISA = Integrated Science Assessment.

a = studies that used a 1-h max exposure metric.

Note: † and red text/symbols = recent studies published since the 2008 ISA for Sulfur Oxides; Black text/symbols = U.S. and Canadian studies evaluated in the 2008 ISA for Sulfur Oxides; circles = all-year results, diamonds = warm season results, squares = cold season results. Corresponding quantitative results are found in Supplemental Table 5S-9 (U.S. EPA, 2017c).

Figure 5-8 Percent increase in respiratory infection hospital admissions and emergency department visits from U.S. and Canadian studies evaluated in the 2008 SO_x ISA and recent studies in all-year and seasonal analyses for a 10-ppb increase in 24-h avg or 40-ppb increase in 1-h max sulfur dioxide concentrations.

Emergency Department Visits

Similar to respiratory infection hospital admissions, recent studies have examined respiratory infection ED visits; however, these studies overall have not consistently examined the same respiratory infection outcomes (Figure 5-8). In their study of seven Canadian cities, Stieb et al. (2009) examined the association between short-term SO₂ exposure and respiratory infection ED visits. The authors reported a positive association at a 2-day lag [1.2% (95% CI: -2.5, 5.2) for a 10-ppb increase in 24-h avg SO₂ concentrations], but this result was uncertain, with no evidence of an association at single-day lags of 0 and 1 days. However, Ségala et al. (2008), in addition to examining bronchiolitis hospital admissions, also examined bronchiolitis ED visits. The authors reported evidence of an association between short-term SO₂ exposures and bronchiolitis ED visits [34.7% (95% CI: 25.5, 44.5); lag 0-4 for a 10-ppb increase in 24-h avg SO₂ concentrations]. However, as mentioned previously, the interpretation of these results is complicated by the lack of copollutant analyses and the high correlation between the

pollutants examined ($r = 0.73$ to 0.87), along with the use of a bidirectional case-crossover approach.

In an additional study conducted in Edmonton, AB, [Zemek et al. \(2010\)](#) examined a new outcome for SO₂, otitis media (i.e., ear infections) ED visits, for ages 1–3 years. Associations were examined for single-day lags of 0 to 4 days in all-year as well as seasonal analyses. The authors found no evidence of an association between short-term SO₂ exposures and increases in ED visits for otitis media at any single-day lag in the all-year analysis.

Physician/Outpatient Visits

In a study conducted in Atlanta, GA as discussed in [Section 5.2.1.2](#), [Sinclair et al. \(2010\)](#) examined the association between air pollution and respiratory infection (e.g., upper respiratory infections, lower respiratory infections) outpatient visits from a managed care organization. As detailed previously, the authors separated the analysis into two time periods—the first 25 months of the study period (i.e., August 1998–August 2000) and the second 28 months of the study period (i.e., September 2000–December 2002). A comparison of the two time periods indicated that risk estimates across outcomes tended to be larger in the earlier 25-month period compared to the later 28-month period. An examination of the respiratory infection outcomes found no evidence of an association for upper respiratory infections at any lag and a positive association for lower respiratory infections for only lag 0–2.

Multiday Lags

In the case of respiratory infection hospital admission and ED visit studies, none of the studies evaluated conducted an extensive analysis of the lag structure of associations. However, [Ségala et al. \(2008\)](#) in a study of acute bronchiolitis examined multiday lags of 0–1 and 0–4 days, which does provide some indication of the lag structure of associations. The authors found relatively similar associations for both multiday lags, but the association was slightly larger for lag 0–4 days (i.e., 31.6 vs. 34.8%). These initial results indicate a potential prolonged effect of SO₂ that could lead to a respiratory infection hospital admission or ED visit.

Seasonal Analyses

A few of the recent studies that examined respiratory infection-related hospital admissions and ED visits also examined whether there was evidence of seasonal differences in associations. Note that interpreting the results from these studies is complicated by the different geographic locations as well as the respiratory infection outcome examined in each study. [Mehta et al. \(2013\)](#) in the study of ALRI hospital

admissions in Vietnam examined potential seasonal differences in associations by dividing the year into the dry (November–April) and rainy seasons (May–October). Within these seasons, SO₂ concentrations differed drastically, with mean 24-h avg SO₂ concentrations being 10.1 ppb in the dry season and 5.7 ppb in the rainy season. In seasonal analyses, [Mehta et al. \(2013\)](#) reported that SO₂ was consistently associated with ALRI hospital admissions in the dry season [16.1% (95% CI: 1.2, 33.3) for a 10-ppb increase in 24-h avg SO₂ concentrations, lag 1–6 day avg], with no evidence of an association in the rainy season. Of the other pollutants that were found to be positively associated with ALRI hospital admissions during the dry season (i.e., PM₁₀ and NO₂), none were associated during the rainy season. In copollutant analyses for the dry season, SO₂ was robust to the inclusion of PM₁₀ and O₃ in the model, with the magnitude of the effect remaining similar, 15.0 and 15.8%, respectively. However, in models with NO₂, the SO₂-ALRI hospital admission association was attenuated, but remained positive with large uncertainty estimates [10.0% (95% CI: –4.6, 26.9) for a 10-ppb increase in 24-h avg SO₂ concentrations, lag 1–6 day avg].

Additionally, [Zemek et al. \(2010\)](#) in the study of otitis media ED visits in Alberta, reported that the magnitude of the association was larger, albeit with wide confidence intervals, in the warm months (April–September), 9.0% (95% CI: –8.4, 34.2), compared to the cold months, (October–March), –4.3% (95% CI: –16.30, 9.0) at lag 4 for a 10-ppb increase in 24-h avg SO₂ concentrations.

Summary of Respiratory Infection

Recent evidence, which comes from epidemiologic studies, expands on that presented in the 2008 ISA for Sulfur Oxides and provides some, but not entirely consistent, support for an association between ambient SO₂ concentrations and respiratory infection. Whereas cross-sectional studies do not consistently link SO₂ exposures estimated for school or home to respiratory infections self-reported by children [Supplemental Figure 5S-2 and Table 5S-8 ([U.S. EPA, 2017c](#))], some evidence points to an association with hospital admission and ED visits ([Figure 5-8](#)). Associations are observed for all respiratory infections combined and bronchiolitis but not pneumonia or otitis media. The lack of multiple studies examining the same respiratory infection outcome complicates the interpretation of the collective body of evidence, specifically because the etiologies of upper and lower respiratory infections are vastly different.

Most supporting evidence points to associations with 24-h avg SO₂ concentrations averaged over 3 to 7 days, but an association was observed with temporally resolved 1-h max as well. The relatively small number of studies does not provide a strong basis for drawing inferences about the lag structure of associations with respiratory infection or

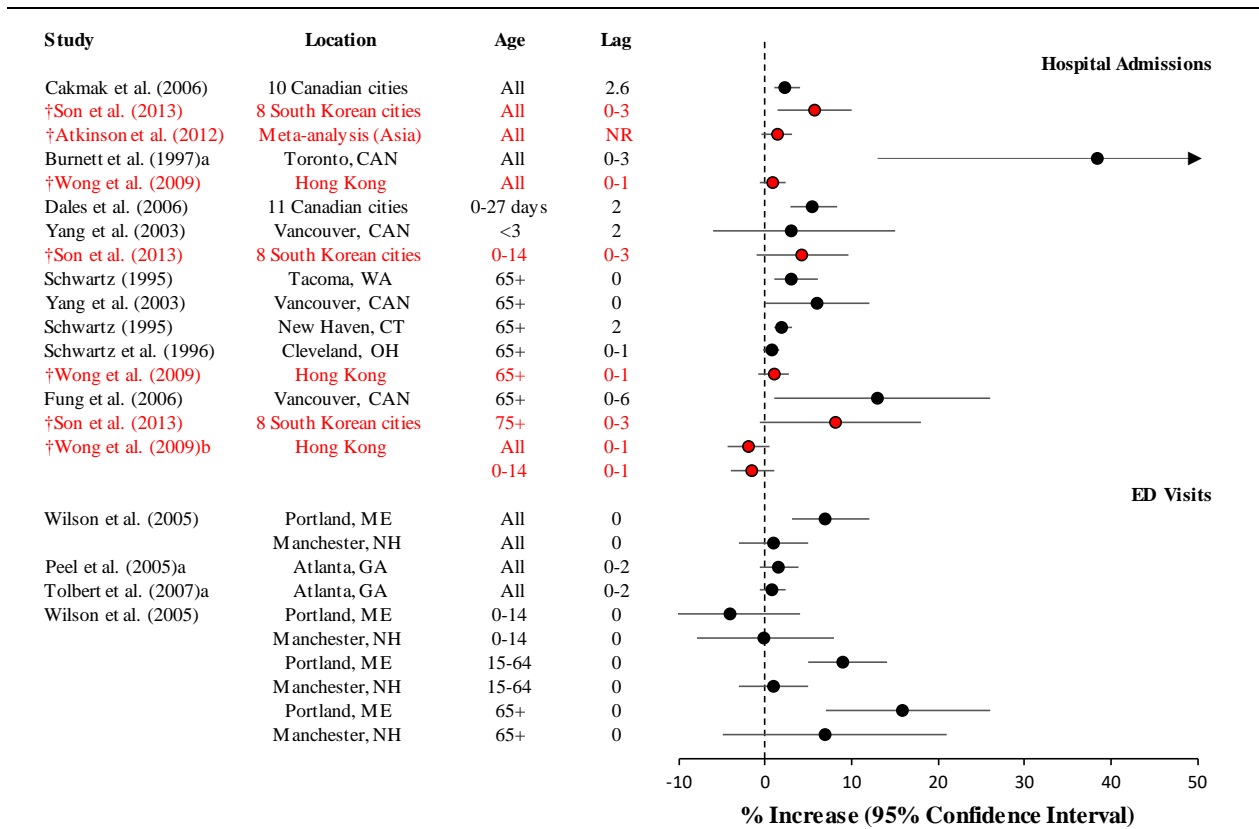
potential seasonal differences in associations. An examination of potential factors that could modify the SO₂-respiratory infection hospital admission or ED visit association finds differences by SES but inconsistent differences by sex ([Chapter 6](#)). Recent studies continued to rely on fixed-site monitors. SO₂ generally has low to moderate spatial correlations across urban geographical scales, and the uncharacterized spatial and temporal variability may contribute to some degree of exposure measurement error ([Section 3.4.2.2](#)). Another uncertainty that persists in the recent evidence is copollutant confounding. Respiratory infection hospital admissions and ED visits were associated with PM_{2.5}, PM₁₀, BS, and NO₂. High SO₂-copollutant correlations were observed ($r = 0.73-0.78$). Correlations were low in some locations ($r = 0.17-0.34$) ([Table 5-13](#)), but differences in exposure measurement error among different pollutants may influence effect estimates, particularly for copollutants with different averaging times. New information from copollutant models shows an SO₂ association that is attenuated and often imprecise (i.e., wide confidence intervals) with adjustment for NO₂, but uncertainty in the exposure estimates weakens inference about independent associations. Information to assess the biological plausibility of epidemiologic findings is limited. There is some evidence in rodents that SO₂ exposures of 0.1–1 ppm alter clearance of particles, but responses to infectious agents have not been examined in relation to ambient-relevant exposures.

5.2.1.6 Aggregated Respiratory Conditions

In addition to individual respiratory conditions, epidemiologic studies examined respiratory effects as an aggregate of multiple respiratory conditions (e.g., asthma, COPD, respiratory infections). Epidemiologic studies examining the association between short-term SO₂ exposures and respiratory-related hospital admissions or ED visits, including those discussed earlier in this chapter, were not available until after the completion of the 1986 Supplement to the Second Addendum of the 1982 SO_x AQCD ([U.S. EPA, 1994](#)). Therefore, the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) included the first thorough evaluation of respiratory morbidity in the form of respiratory-related hospital admissions and ED visits. Of the studies evaluated, the majority consisted of single-city, time-series studies that primarily examined all respiratory disease or asthma hospital admissions or ED visits, with a more limited number of studies examining other respiratory outcomes, as discussed in previous sections. Additionally, most studies averaged SO₂ concentrations over multiple monitors and examined 24-h avg exposure metrics, which may not adequately capture the spatial and temporal variability in SO₂ concentrations ([Section 3.4.2](#)). The studies that examined all respiratory disease hospital admissions and ED visits generally reported positive associations ([Figure 5-9](#)). These

associations were found to remain generally positive with some evidence of an attenuation of the association in models with gaseous pollutants (i.e., NO₂ and O₃) and particulate matter ([U.S. EPA, 2008d](#)).

Since the completion of the 2008 SO_x ISA, recent studies have examined the association between short-term exposure to ambient SO₂ and all respiratory disease hospital admissions and ED visits. For each of the studies evaluated in this section, [Table 5-14](#) presents the air quality characteristics of each city or across all cities, the exposure assignment approach used, and information on copollutants examined in each hospital admission and ED visit study that examined all respiratory diseases. Other recent studies that have examined all respiratory disease hospital admissions and ED visits are not the focus of this evaluation because of various study design issues, as initially detailed in [Section 5.2.1.2](#), but the full list of these studies, as well as study specific details, can be found in Supplemental Table 5S-6 ([U.S. EPA, 2017c](#)).



ED = emergency department; ISA = Integrated Science Assessment.

^a = studies that used a 1-h max exposure metric.

^b = [Wong et al. \(2009\)](#) also presented results for acute respiratory disease hospital admissions, which is a subset of total respiratory hospital admissions.

Note: † and red text/symbols = recent studies published since the 2008 ISA for Sulfur Oxides; Black text/symbols = U.S. and Canadian studies evaluated in the 2008 ISA for Sulfur Oxides. Corresponding quantitative results are found in Supplemental Table 5S-10 ([U.S. EPA, 2017c](#)).

Figure 5-9 Percent increase in respiratory disease hospital admissions and emergency department visits from U.S. and Canadian studies evaluated in the 2008 SO_x ISA and recent studies in all-year and seasonal analyses for a 10-ppb increase in 24-h avg or 40-ppb increase in 1-h max sulfur dioxide concentrations.

Table 5-14 Study-specific details and mean and upper percentile concentrations from respiratory disease hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location Years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb	Copollutants Examined
Hospital admissions						
Cakmak et al. (2006)	10 Canadian cities (1993–2000)	Average of SO ₂ concentrations across all monitors in each city	24-h avg	4.6	Max: 14–75	Correlations (<i>r</i>): NR Copollutant models: none
Dales et al. (2006)	11 Canadian cities (1986–2000)	Average of SO ₂ concentrations across all monitors in each city	24-h avg	4.3 ^a	95th: 3.5–23.5	Correlations (<i>r</i>): PM ₁₀ : –0.09 to 0.61 O ₃ : –0.41 to 0.13 NO ₂ : 0.20 to 0.67 CO: 0.19 to 0.66 Copollutant models: none
Burnett et al. (1997)	Toronto, ON (1992–1994)	Average of SO ₂ concentrations from 4–6 monitors during the course of the study	1-h max	7.9	75th: 11 95th: 18 Max: 26	Correlations (<i>r</i>): H ⁺ : 0.45 SO ₄ : 0.42 PM ₁₀ : 0.55 PM _{2.5} : 0.49 PM _{10-2.5} : 0.44 COH: 0.50 O ₃ : 0.18 NO ₂ : 0.46 CO: 0.37 Copollutant models: COH, PM ₁₀ , PM _{10-2.5} , PM _{2.5}

Table 5-14 (Continued): Study specific details and mean and upper percentile concentrations from respiratory disease hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location Years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb	Copollutants Examined
Fung et al. (2006)	Vancouver, BC (1995–1999)	Average of SO ₂ concentrations across all monitors within Vancouver	24-h avg	3.46	Max: 12.5	Correlations (<i>r</i>): CO: 0.61 COH: 0.65 O ₃ : -0.35 NO ₂ : 0.57 PM ₁₀ : 0.61 PM _{2.5} : 0.42 PM _{10-2.5} : 0.57 Copollutant models: none
Schwartz (1995)	New Haven, CT Tacoma, WA (1988–1990)	Average of SO ₂ concentrations across all monitors in each city	24-h avg	New Haven: 29.8 Tacoma: 11.5	New Haven: 75th: 38.2 90th: 60.7 Tacoma: 75th: 21.4 90th: 28.2	Correlations (<i>r</i>): NR Copollutant models: PM ₁₀ , O ₃
Schwartz et al. (1996)	Cleveland, OH (1988–1990)	Average of SO ₂ concentrations across all monitors	24-h avg	35.0	75th: 45.0 90th: 61.0	Correlations (<i>r</i>): NR Copollutant models: none
Yang et al. (2003)	Vancouver, BC (1986–1998)	Average of SO ₂ concentrations across four monitors	24-h avg	4.8	75th: 6.3 Max: 24.0	Correlation (<i>r</i>): O ₃ : -0.37 Copollutant models: O ₃
† Son et al. (2013)	Eight South Korean cities (2003–2008)	Average of hourly ambient SO ₂ concentrations from monitors in each city	24-h avg	3.2–7.3	NR	Correlation (<i>r</i>): PM ₁₀ : 0.5 O ₃ : -0.1 NO ₂ : 0.6 Copollutant models: none
† Atkinson et al. (2012)	Meta-analysis (Asia) (1980–2007)	NR	24-h avg	NR	NR	Correlation (<i>r</i>): NR Copollutant models: none

Table 5-14 (Continued): Study specific details and mean and upper percentile concentrations from respiratory disease hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location Years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb	Copollutants Examined
†Wong et al. (2009)	Hong Kong, China (1996–2002)	Average of SO ₂ concentrations from eight monitoring stations	24-h avg	6.8	75th: 8.4 Max: 41.8	Correlation (r): NR Copollutant models: none
ED visits						
Peel et al. (2005)	Atlanta, GA (1993–2000)	Average of SO ₂ concentrations from monitors for several monitoring networks	1-h max	16.5	90th: 39.0	Correlations (r): PM _{2.5} : 0.17 PM ₁₀ : 0.20 PM _{10-2.5} : 0.21 UFP: 0.24 PM _{2.5} water soluble metals: 0.00 PM _{2.5} sulfate: 0.08 PM _{2.5} acidity: -0.03 PM _{2.5} OC: 0.18 PM _{2.5} EC: 0.20 Oxygenated HCs: 0.14 O ₃ : 0.19 CO: 0.26 NO ₂ : 0.34 Copollutant models: none

Table 5-14 (Continued): Study specific details and mean and upper percentile concentrations from respiratory disease hospital admission and emergency department visit studies conducted in the U.S. and Canada and evaluated in the 2008 SO_x ISA and studies published since the 2008 SO_x ISA.

Study	Location Years	Exposure Assignment	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb	Copollutants Examined
Tolbert et al. (2007)	Atlanta, GA (1993–2004)	Average of SO ₂ concentrations from monitors for several monitoring networks	1-h max	14.9	75th: 20.0 90th: 35.0	Correlations (<i>r</i>): PM ₁₀ : 0.21 O ₃ : 0.21 NO ₂ : 0.36 CO: 0.28 PM _{10-2.5} : 0.16 PM _{2.5} : 0.17 PM _{2.5} SO ₄ : 0.09 PM _{2.5} EC: 0.22 PM _{2.5} OC: 0.17 PM _{2.5} TC: 0.19 PM _{2.5} water soluble metals: 0.06 Organic hydrocarbon: 0.05 Copollutant models: none
Wilson et al. (2005)	Portland, ME Manchester, NH (1996–2000)	SO ₂ concentrations from one monitor in each city	24-h avg	Portland: 11.1 Manchester: 16.5	NR	Correlation (<i>r</i>): Portland O ₃ : 0.05 Manchester O ₃ : 0.01 Copollutant models: none

Avg = average; CO = carbon monoxide; COH = coefficient of haze; EC = elemental carbon; H⁺ = hydrogen ion; HC = hydrocarbon; ISA = Integrated Science Assessment; max = maximum; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OC = organic carbon; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM_{10-2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm and greater than a nominal 2.5 μm; *r* = correlation coefficient; SO₂ = sulfur dioxide; SO₄ = sulfate; TC = total hydrocarbon; UFP = ultrafine particle.

† = studies published since the 2008 SO_x ISA.

Hospital Admissions

A recent multicity study conducted in Korea ([Son et al., 2013](#)) and a single-city study conducted in Hong Kong ([Wong et al., 2009](#)) provide additional insight into the relationship between short-term SO₂ exposures and hospital admissions for all respiratory diseases.

[Son et al. \(2013\)](#) examined the association between short-term exposures to air pollution and respiratory-related hospital admissions in eight South Korean cities. The city-specific means ranged from 3.2 to 7.3 ppb in this study. It is important to note that South Korea has unique demographic characteristics with some indicators more in line with other developed countries (e.g., life expectancy, percent of population living in urban areas), but because it represents a rapidly developing Asian country, it is likely to have different air pollution, social, and health patterns than less industrialized Asian nations or Western nations that developed earlier ([Son et al., 2013](#)). In a time-series analysis using a two-stage Bayesian hierarchical model, [Son et al. \(2013\)](#) examined both single-day lags and multiday lags up to 3 days (i.e., lag 0–3). For a lag of 0–3 days the authors reported a 5.6% increase (95% CI: 1.4, 10.0) in respiratory disease hospital admissions for a 10-ppb increase in 24-h avg SO₂ concentrations. The authors did not conduct copollutant analyses; however, SO₂ was found to be moderately correlated with PM₁₀ ($r = 0.5$), NO₂ ($r = 0.6$), and CO ($r = 0.6$). The results of [Son et al. \(2013\)](#) add additional support to the results from the multicity studies evaluated in the 2008 SO_x ISA [i.e., ([Cakmak et al. \(2006\)](#); [Dales et al. \(2006\)](#))] in terms of the lag in which the strongest associations were observed and the magnitude of the association ([Figure 5-9](#)).

A greater degree of variability in the magnitude of the association between short-term SO₂ exposures and all respiratory hospital admissions was observed when evaluating single-city studies in the 2008 SO_x ISA ([Figure 5-9](#)). [Wong et al. \(2009\)](#) in a study conducted in Hong Kong reported results consistent with these earlier single-city studies for individuals over the age of 65 [1.0% (95% CI: -0.8, 2.8) for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 0–1]. However, compared to studies that examined all ages, the magnitude of the association was much smaller [0.8% (95% CI: -0.6, 2.3) for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 0–1]. [Wong et al. \(2009\)](#) also examined acute respiratory disease, which represents a smaller subset of outcomes within all respiratory diseases. When focusing on only acute respiratory disease, [Wong et al. \(2009\)](#) reported no evidence of an association at a 0–1-day lag for all ages [-2.0% (95% CI: -4.4, 0.4) for a 10-ppb increase in 24-h avg SO₂ concentrations].

The all-respiratory-disease hospital admissions results of [Son et al. \(2013\)](#) and [Wong et al. \(2009\)](#) are supported by the results of a meta-analysis conducted by [Atkinson et al.](#)

(2012) that focused on studies conducted in Asian cities since 1980. The six estimates from studies that examined the association between SO₂ and all respiratory hospital admissions were included in a random effects model, which yielded a 1.3% increase in respiratory hospital admissions (95% CI: -0.4, 3.2) for a 10-ppb increase in 24-h avg SO₂ concentrations. However, [Atkinson et al. \(2012\)](#) found some evidence of publication bias for associations between SO₂ and respiratory hospital admissions.

Emergency Department Visits

The 2008 SO_x ISA evaluated a few studies that examined the association between short-term SO₂ exposures and all respiratory ED visits [[Figure 5-9](#), Supplemental Table 5S-10 ([U.S. EPA, 2017c](#))]. These studies reported evidence of a positive association, but the magnitude of the association varied across study locations. However, these studies were limited in that they did not examine copollutant confounding. Recent studies that examined the association between air pollution and all respiratory ED visits have not examined associations with SO₂.

Model Specification—Sensitivity Analyses

A question that often arises when evaluating studies that examine the association between air pollution and a health effect is whether the statistical model employed adequately controls for the potential confounding effects of temporal trends and meteorological conditions. [Son et al. \(2013\)](#), in the study of eight South Korean cities, conducted sensitivity analyses to identify whether risk estimates changed depending on the df used to control for temporal trends and meteorological covariates (i.e., temperature, humidity, and barometric pressure). The authors reported that the association between short-term SO₂ exposures and all of the respiratory hospital admission outcomes examined (i.e., all respiratory diseases, allergic disease, and asthma) was sensitive to using less than 7 df per year, indicating inadequate control for temporal trends, but was stable when using 7–10 df per year. These results suggest that at least 7 df per year are needed to adequately account for temporal trends when examining the relationship between short-term SO₂ exposures and respiratory disease hospital admissions. However, additional studies have not systematically examined this issue.

In an additional sensitivity analysis focusing on meteorological covariates (i.e., temperature, relative humidity, and barometric pressure), [Son et al. \(2013\)](#) examined whether risk estimates were sensitive to the degree of smoothing used and to the lag structure. The authors found that when varying the number of df for each covariate from 3 to 6 df and varying the lag structure (i.e., lag 0 and lag 0–3 days), the SO₂ association remained robust for all respiratory hospital admission outcomes.

Lag Structure of Associations

As stated previously, when examining associations between air pollution and a specific health outcome, it is informative to assess whether there is a specific exposure window for SO₂ that results in the strongest association with the health outcome of interest. In the examination of all respiratory disease hospital admissions, [Son et al. \(2013\)](#) focused on both single-day and multiday lags to address whether there is evidence of an immediate or persistent effect of SO₂. Across single-day lags of 0 to 3 days, positive associations were observed across each lag with the magnitude of the association being relatively similar across each lag (i.e., 2.4% for lag 0 and 2.1% for lags 1 to 3 days for a 10-ppb increase in 24-h avg SO₂ concentrations). When examining multiday lags of 0–1, 0–2, and 0–3 days, the authors reported an increase in the magnitude of the association as the length of the multiday lag increased with a 3.5% increase reported at lag 0–1 and a 5.6% increase reported for lag 0–3 days. Therefore, the limited evidence suggests that SO₂ effects occur within the first few days after exposure, but also that SO₂ effects on respiratory disease hospital admissions may persist over several days.

Examination of Seasonal Differences

Of the studies that examined all respiratory disease hospital admissions or ED visits, only [Son et al. \(2013\)](#) in the analysis of eight South Korean cities examined potential seasonal differences in SO₂ associations. However, it is important to note the potential influence of geographic location on the results from studies that examine potential seasonal differences in associations. For all outcomes examined, including respiratory diseases, the association with SO₂ was largest in magnitude during the summer, although confidence intervals were quite large [respiratory diseases: 21.5% (95% CI: –0.7, 48.3), lag 0–3, for a 10-ppb increase in 24-h avg SO₂ concentrations] with additional evidence of a positive association in the fall [8.9% (95% CI: –1.4, 20.7), lag 0–3, for a 10-ppb increase in 24-h avg SO₂ concentrations]. There was no evidence of an association between short-term SO₂ exposures and respiratory disease hospital admissions in either the spring or winter seasons. Across the eight cities, mean 24-h avg SO₂ concentrations were lowest during the summer season (4.4 ppb compared to a range of 4.8 to 7.0 in the other seasons) as was also the case for NO₂ and CO.

Summary of Aggregate Respiratory Conditions

Recent studies add to the evidence detailed in the 2008 SO_x ISA that indicated a generally positive association between short-term SO₂ exposures and respiratory disease hospital admissions and ED visits ([Figure 5-9](#)). These recent studies provide some insight

into previously identified limitations (i.e., model specification, lag structure of associations, and potential seasonal differences) in the SO₂-respiratory disease hospital admission and ED visits relationship. Initial evidence from a limited number of studies suggests that SO₂ associations are robust to alternative model specifications for weather covariates and that SO₂ associations are relatively stable in the range of df per year indicative of reasonable control for temporal trends (i.e., 7–10 df per year); however, more studies are needed to confirm these findings. Additionally, an examination of the lag structure of associations is in line with the results reported in studies that focused on a priori lags [i.e., associations tend to be strongest within the first few days after exposure, primarily within the range of 0 to 3 days ([Figure 5-9](#))]. The potential seasonal patterns in SO₂ associations remain unclear due to the variability in SO₂ associations observed across different geographic locations, as reflected in studies of other respiratory hospital admission and ED visit outcomes. Some studies have also examined whether there is evidence that specific factors modify the SO₂-respiratory disease hospital admission or ED visit relationship and have found some evidence for potential differences by lifestage and influenza intensity (see [Chapter 6](#)). Studies of all respiratory hospital admissions and ED visits have not conducted extensive analyses to examine potential copollutant confounding. However, studies that reported SO₂ correlations with other pollutants found low ($r < 0.4$) to moderate ($r = 0.4$ – 0.7) correlations. Overall, the results of recent studies are limited in that they do not further inform the understanding of potential confounding by copollutants on the relationship between short-term SO₂ concentrations and respiratory disease hospital admissions and ED visits.

5.2.1.7 Respiratory Effects in General Populations and Healthy Individuals

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) reported respiratory effects of SO₂ in general populations and healthy individuals but did not make specific conclusions about the relationship. Respiratory effects were demonstrated in healthy individuals following SO₂ exposures ≥ 1.0 ppm in controlled human exposure studies. Animal toxicological studies demonstrated bronchoconstriction after a single SO₂ exposure and provided evidence for increased airway responsiveness and inflammation after repeated SO₂ exposures. Epidemiologic evidence was weak. The few recent toxicological studies corroborate previous results, but recent epidemiologic and controlled human exposure studies provide inconsistent results, including new results for pulmonary inflammation.

Lung Function Changes in General Populations and Healthy Individuals

Compared with evidence for lung function changes in individuals with asthma, evidence for SO₂-induced lung function effects in healthy individuals is weak. Most of the controlled human exposure studies evaluating these effects in healthy individuals were discussed in the 1982 SO_x AQCD ([U.S. EPA, 1982a](#)). While some studies showed that transient decreases in lung function can occur at concentrations of 1.0 ppm SO₂ under exercising or forced oral breathing conditions, the evidence was more consistent for exposures >1.0 ppm ([U.S. EPA, 2008d](#)). Epidemiologic associations between ambient SO₂ concentrations and lung function continue to be inconsistent in children. While recent results indicate associations in adults, inferences about SO₂ exposure still are weak because of uncertainty in the exposure estimates and copollutant confounding.

Controlled Human Exposure Studies

Evidence from controlled human exposure studies evaluating SO₂-induced lung function changes in healthy adults was extensively discussed in the 1982 AQCD ([U.S. EPA, 1982a](#)). In general, these studies demonstrated respiratory effects such as increased airway resistance and decreased FEV₁ following exposures to concentrations >1.0–5.0 ppm, while some studies demonstrated respiratory effects at 1.0 ppm.

Lung function changes in response to SO₂ exposure in controlled human exposure studies have been investigated since the early 1950s. Respiratory effects including increased respiration rates, decrements in peak flow, bronchoconstriction, and increased airway resistance have been observed in healthy human volunteers at concentrations ≥1.0 ppm ([Lawther et al., 1975](#); [Andersen et al., 1974](#); [Snell and Luchsinger, 1969](#); [Abe, 1967](#); [Frank et al., 1962](#); [Sim and Pattle, 1957](#); [Lawther, 1955](#); [Amdur et al., 1953](#)). Although bronchoconstriction was observed in healthy subjects exposed to concentrations ≥5.0 ppm, shallow rapid respiration and increased pulse rate, decreased maximum expiratory flow from one-half vital capacity, and increased sRaw were observed following exposures as low as 1.0 ppm ([Lawther et al., 1975](#); [Snell and Luchsinger, 1969](#); [Amdur et al., 1953](#)). Overall, only these few studies have reported SO₂-induced respiratory effects in healthy individuals for 5–10-minute exposures at concentrations ≥1.0 ppm SO₂.

A limited number of studies examined lung function changes in healthy populations in response to ≥1 hour exposures to SO₂. Controlled human exposure studies examining lung function changes in healthy individuals exposed to SO₂ are summarized in [Table 5-15](#). [Andersen et al. \(1974\)](#) reported that exposures of up to 6 hours to 1.0 ppm SO₂ in resting healthy adults induced decreases in FEF_{25–75} and to a lesser extent FEV₁. Another human exposure study ([van Thriel et al., 2010](#)) reported that healthy subjects

exposed to SO₂ concentrations of 0.5, 1.0, or 2.0 ppm for 4 hours with two 15-minute exercise periods separated by about 140 minutes did not show changes in FEV₁ following the 4-hour SO₂ exposure. However, lung function measurements in this study were not performed until at least 35–85 minutes after exercise. Healthy individuals at rest or exercising exhibited no changes in several measures of lung function following a 1-hour exposure to 0.2–0.6 ppm SO₂ ([Tunnicliffe et al., 2003](#); [Linn et al., 1987](#)).

The interaction of SO₂ exposure with O₃ was reported in two studies. [Hazucha and Bates \(1975\)](#) demonstrated that a combined 2-hour exposure to low concentrations of O₃ (0.37 ppm) and SO₂ (0.37 ppm) had a greater effect on lung function than exposure to either agent alone in exercising adults. However, [Bedi et al. \(1979\)](#), using a similar study design, did not observe a greater effect of the combined exposures compared with exposure to only O₃; exposure to SO₂ alone had no effect.

Table 5-15 Study-specific details from controlled human exposure studies of lung function and respiratory symptoms in healthy adults.

Reference	Disease Status; n; Sex; Age (mean ± SD)	Exposure Details (Concentration; Duration)	Endpoints Examined
Andersen et al. (1974)	Healthy; n = 15; 15 M; 20–28 yr	0, 1, 5, or 25 ppm SO ₂ for 6 h at rest	Nasal mucociliary flow Area of the nasal airway Airway resistance (FEV ₁ , FEF _{25–75%}) Nasal removal of SO ₂ Discomfort level symptoms
Linn et al. (1987)	Healthy; n = 24; 15 M, 9 F; 18–37 yr	0, 0.2, 0.4, or 0.6 ppm SO ₂ 1 h exposures 3 × 10-min exercise (bicycle) periods ~40 L/min Exposures were repeated for a total of eight	Lung function measure pre-exposure, ~15 min, and ~55 min into exposure sRaw, FVC, FEV ₁ , peak expiratory flow rate, maximal midexpiratory flow rate Continuously EKG Midway-HR Before, during, 1-day after, and 1-wk after symptom score, self-rated activity Immediately after exposure-bronchial reactivity percent change in FEV induced by 3 min normocapnic hyperpnea with cold, dry air
Raulf-Heimsoth et al. (2010)	Healthy; n = 16; 8 M, 8 F; 19–36 yr	0, 0.5, 1.0, or 2.0 SO ₂ for 4 h with exercise for 15 min (bicycle, 75 Watts) two times during each session	Exhaled NO, biomarkers of airway inflammation in EBC and NALF
Tunncliffe et al. (2003)	Asthma; n = 12 adults, 35.7 yr Healthy; n = 12 adults, 34.5 yr	0 or 0.2 ppm SO ₂ for 1 h at rest	Symptoms, FEV ₁ , FVC, MMEF, exhaled NO, ascorbic and uric acid in nasal lavage fluid
van Thriel et al. (2010)	Healthy; n = 16; 8 M, 8 F; M: 28.4 ± 3.9 yr, F: 24.3 ± 5.2 yr	0, 0.5, 1.0, or 2.0 ppm SO ₂ for 4 h with exercise for 15 min (bicycle, 75 Watts) two times during each session	Symptoms, FEV ₁

EBC = exhaled breath condensate; EKG = electrocardiogram; F = female; FEF_{25–75%} = forced expiratory flow at 25–75% of exhaled volume; FEV = forced expiratory volume; FEV₁ = forced expiratory volume in 1 sec; FVC = forced vital capacity; HR = heart rate; M = male; MMEF = maximum midexpiratory flow; n = sample size; NALF = nasal lavage fluid; NO = nitric oxide; SD = standard deviation; SO₂ = sulfur dioxide; sRaw = specific airway resistance.

Epidemiologic Studies

Previous epidemiologic evidence was inconsistent for an association between ambient SO₂ concentrations and lung function in healthy adults or children and people recruited from the general population ([U.S. EPA, 2008d](#)). Studies mostly estimated SO₂ exposure from fixed-site monitors and did not report whether the measurements well captured the spatiotemporal variability in the study areas. Some recent studies measured SO₂ at subjects' locations and observed associations with lung function decrements in adults but not consistently in children. Most studies examined 24-h avg SO₂ concentrations, which are much longer than the 5–10 minute exposures inducing lung function decrements in experimental studies. Inconsistency also is observed among recent results for temporally resolved metrics such as 1-h max and 1- to 10-h avg SO₂ concentrations, which is similar to controlled human exposure findings for 1- to 6-hour exposures to SO₂.

Adults. An SO₂-associated decrease in lung function was observed in adults in Beijing, China where coal was used for domestic heating ([Xu et al., 1991](#)). Recent results are based on much lower SO₂ concentrations [means 7.3–8.6 ppb vs. 6.8–49 ppb in [Xu et al. \(1991\)](#)]. Associations are observed with lung function decrements in adults without respiratory disease ([Table 5-16](#)), with some based on relatively good exposure characterization ([Dales et al., 2013](#)).

The exposure characterization of [Dales et al. \(2013\)](#) is judged to be good because SO₂ was measured on site of adults' scripted exposures near (0.87 km) and away from (4.5 km at a college campus) a steel plant in Ontario. Another strength was the well-defined 8-hour exposure duration and lag between exposure and lung function testing. Higher SO₂ concentrations averaged over 10 hours (8 a.m.–6 p.m.) were associated with decreases in several lung function parameters measured just after exposure ([Table 5-16](#)). For example, a 10-ppb increase in SO₂ was linked to a –0.50% FEV₁ change (95% CI: –1.0, 0.05). [Son et al. \(2010\)](#) also examined air pollution from industry, in this case a petrochemical complex in Ulsan, South Korea. Ambient SO₂ concentrations across the study area were highly variable. Between-monitor correlation varied widely (0–0.8), even for those 5 km apart, and the mean decreased from about 0.4 to 0.2 with increasing distance up to 20 km. Investigators aimed to capture this spatiotemporal variability by combining SO₂ measurements across monitors with inverse distance weighting or kriging. These metrics and that for the nearest monitor to the subjects' home, all 24-h avg SO₂, were associated with forced vital capacity (FVC) but not FEV₁ ([Table 5-16](#)). The implications overall are unclear because many subjects lived far from a monitor, and potential confounding by meteorological factors and season were not considered. Both studies observed associations with copollutants among PM_{2.5}, PM₁₀, ultrafine particle (UFP), CO, NO₂, and O₃. Correlations between copollutants and

analyses of confounding or interactions were not reported for personal exposures near the steel plant ([Dales et al., 2013](#)). For the study near the petrochemical complex, the decrease in FEV₁ for kriged SO₂ was larger after CO adjustment ([Son et al., 2010](#)) ([Table 5-16](#)). The effect estimate for CO became null, but the range of between-monitor correlations was 0–0.8. The effect estimate for SO₂ was attenuated with adjustment for O₃, and the attenuation could be influenced by differences in exposure measurement error between the two pollutants. Between-monitor correlations were 0.4 to 0.8 for O₃.

Other studies reported SO₂-associated lung function decrements, but inference about SO₂ is weaker ([Steinvil et al., 2009](#); [Min et al., 2008a](#)). Associations were observed for SO₂ after adjustment for NO₂ or CO, but correlations with SO₂ were 0.62–0.70, and single-pollutant associations for SO₂ were in opposing directions across lags and limited to lags of 3 or more days ([Steinvil et al., 2009](#)). Associations were observed with 1-h avg SO₂ concentrations lagged 5–30 hours, but confounding by meteorological factors was not considered ([Min et al., 2008a](#)). Also, both studies had cross-sectional design and estimated SO₂ exposure from monitors up to 11 km or unspecified distance from homes.

Children. Similar to previous studies, many recent studies of children examined populations with high prevalence (8–35%) of respiratory disease, such as asthma, and populations outside the U.S. and Canada. As examined in several recent studies, SO₂ at schools was inconsistently associated with lung function ([Table 5-17](#)). Previously, 1-h max SO₂ concentrations at school were not associated with lung function. Additional results for temporally resolved SO₂ metrics, both school and fixed-site, are inconsistent.

Table 5-16 Epidemiologic studies of lung function in healthy adults and adults in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Dales et al. (2013) Sault Ste. Marie, ON, May–Aug 2010 N = 61, mean age 24 yr. 100% healthy. Cross-over, with scripted outdoor exposures near and away from steel plant. Five consecutive 8-h days at each site, with 9-day washout period in between. Supervised spirometry. Recruited from university. Required not to live in neighborhood bordering steel plant.</p>	<p>Monitor on site of outdoor exposures Mean (SD) Near steel plant 7.8 (13) College campus 1.6 (4.2)</p>	<p>10-h avg (8 a.m.–6 p.m.) Lag 0 h</p>	<p>Percent change FEV₁: -0.50 (-1.0, 0.05) FVC: -0.45 (-1.1, 0.19) FEV₁/FVC: -0.15 (-0.31, 0.01) FEF_{25–75%}: -0.44 (-0.74, -0.14) Total lung capacity -0.42 (-0.70, -0.13) Residual volume -2.1 (-4.1, -0.18)</p>	<p>No copollutant model Copollutant correlations NR.</p>

Table 5-16 (Continued): Epidemiologic studies of lung function in healthy adults and adults in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Son et al. (2010) Ulsan, South Korea, 2003–2007 N = 2,102, ages 7–97 yr. Mean age 45 yr. Mean percent predicted FEV₁ 83%. Cross-sectional. Supervised spirometry. Recruited from a meeting of residents near a petrochemical complex. Did not examine confounding by meteorological factors or season.</p>	<p>13 monitors in city Mean (SD), 75th percentile, max Kriging 8.3 (4.4), 9.6, 25 Nearest monitor 7.3 (5.9), 9.5, 34 IDW 8.4 (5.3), 11, 29 Average of 13 monitors 8.6 (4.1), 10, 24</p>	<p>24-h avg 0–2 avg</p>	<p>Change in percent predicted FVC Kriging -6.2 (-8.2, -4.2) IDW -5.3 (-7.1, -3.5) Nearest monitor -5.6 (-7.4, -3.9) Average of 13 monitors -7.0 (-9.0, -4.8) FEV₁ Kriging -0.08 (-0.76, 0.60) IDW 0.31 (-0.32, 0.95) Nearest monitor 0.35 (-0.21, 0.92) Average of 13 monitors -0.15 (-0.89, 0.58)</p>	<p>FVC, lag 0–2 avg, kriged SO₂ Persists with: CO or O₃ PM_{2.5} not examined. Copollutant correlations NR.</p>
<p>†Steinvil et al. (2009) Tel Aviv, Israel, 2002–2007 N = 2,380, mean age 43 yr. 100% healthy. Cross-sectional. Supervised spirometry. Recruited from ongoing survey of individuals attending health center.</p>	<p>Three monitors within 11 km of home Mean (SD): 2.8 (1.2) 75th percentile: 3.4 Max: 9.4</p>	<p>24-h avg 0 5 0–6 avg 0 5 0–6 avg 0 5 0–6 avg</p>	<p>Change in FEV₁ (mL) 93 (-90, 277) -300 (-487, -113) -447 (-750, -143) Change in FVC (mL) 53 (-167, 273) -373 (-600, -147) -560 (-927, -193) Percent change in FEV₁/FVC 716 (-6.5, 4,233) 237 (-79, 2,195) 220 (-217, 657)</p>	<p>FEV₁ (mL), lag 5 Persists with: O₃, NO₂, or CO r = 0.70 NO₂, 0.62 CO, -0.24 O₃.</p>

Table 5-16 (Continued): Epidemiologic studies of lung function in healthy adults and adults in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Min et al. (2008a) South Korea, 2006 N = 867, ages 20–86 yr. 100% no serious medical conditions. Cross-sectional. Supervised spirometry. Recruitment not described. Did not examine confounding by meteorological factors.</p>	<p>Monitors in city Number and distance NR Mean: 6</p>	<p>1-h avg Lag 1 h</p>	<p>Results presented only in figure. Associations observed only in smokers. FEV₁ and FVC decrease after lag of 5–6 h. No association after 30 h.</p>	<p>No copollutants examined.</p>

Avg = average; CI = confidence interval; CO = carbon monoxide; FEF_{25–75%} = forced expiratory flow at 25–75% of forced vital capacity; FEV₁ = forced expiratory volume in 1 sec; FVC = forced vital capacity; IDW = inverse distance weighting; max = maximum; N = population number; n=sample size; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; r = correlation coefficient; PM_{2.5} = particulate matter with nominal aerodynamic diameter less than or equal to 2.5 µm; PM₁₀ = particulate matter with nominal aerodynamic diameter less than or equal to 10 µm; SD = standard deviation; SO₂ = sulfur dioxide.

^aEffect estimates are standardized to a 10-ppb increase in 1-h to 24-h avg SO₂.

† = Studies published since the 2008 ISA for Sulfur Oxides.

Table 5-17 Epidemiologic studies of lung function in healthy children and children in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
†Correia-Deur et al. (2012) São Paolo, Brazil, Apr–Jul 2004 N = 31, ages 9–11 yr. 100% no allergic sensitization. Daily measures for 15 days. Supervised spirometry. Recruited from schools.	Monitor at school Mean (SD): 8.8 (3.3) 75th percentile: 11 90th percentile: 13	2-h avg	Percent change in PEF	Remains null with: PM ₁₀ , NO ₂ , or CO (analysis includes 65 children with atopy). Pearson <i>r</i> = 0.75 PM ₁₀ , 0.60 NO ₂ , 0.60 CO.
		0	–0.24 (–0.96, 0.49)	
		24-h avg	–0.20 (–1.4, 0.96)	
		0	No association for 3-, 5-, 7-, or 10-day avg	
†Altuğ et al. (2014) Eskisehir, Turkey, Feb–Mar 2007 N = 535, ages 9–13 yr Cross-sectional. Supervised spirometry. Recruited from schools from participants of a larger study.	Monitor at school Mean and max Suburban: 21, 29 Urban: 29, 44 Urban-traffic: 22, 27	24-h avg	Relative ratio for change	No copollutant model PM _{2.5} and PM ₁₀ not examined. <i>r</i> = 0.49 NO ₂ , –0.40 O ₃ .
		0–6 avg	Subjects without URS	
			FVC: 1.00 (0.97, 1.03)	
			FEV ₁ : 1.00 (0.97, 1.03)	
			PEF: 1.00 (0.97, 1.03)	
			MMEF: 1.00 (0.92, 1.08)	
			Subjects with URS	
			FVC: 1.00 (0.97, 1.03)	
			FEV ₁ : 1.00 (0.97, 1.03)	
			PEF: 1.00 (0.97, 1.03)	
			MMEF: 1.03 (0.95, 1.11)	

Table 5-17 (Continued): Epidemiologic studies of lung function in healthy children and children in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Altuğ et al. (2013) Eskisehir, Turkey, Jan 2008–Mar 2009 N = 1,880, 9–13 yr. 7% asthma. 11% hay fever Two measures: summer and winter. Supervised spirometry. Recruited from schools. Did not examine confounding by meteorological factors.</p>	<p>Monitor at school Mean and max Summer Suburban: 8.5, 16 Urban: 10, 16 Urban-traffic: 6.3, 8.9 Winter Suburban: 21, 29 Urban: 29, 44 Urban-traffic: 22, 33</p>	<p>24-h avg 0–6 avg</p>	<p>OR for impaired lung function (predicted values <85% for FEV₁ or FVC or <75% for PEF or MMEF) Summer Girls: 1.22 (0.72, 2.09) Boys: 0.83 (0.47, 1.45) Winter Girls: 1.00 (0.76, 1.32) Boys: 0.83 (0.61, 1.11)</p>	<p>Remains null with: O₃ or NO₂ PM_{2.5} and PM₁₀ not examined. <i>r</i> = 0.49 NO₂, –0.40 O₃. (winter). Summer correlations NR.</p>
<p>†Castro et al. (2009) Rio de Janeiro, Brazil, 2004 N = 118, ages 6–15 yr. 18% asthma. Daily measures for 6 wk. Supervised PEF. Recruited from schools.</p>	<p>Monitor at school Mean (SD): 7.1 (6.8) 90th percentile: 16 Max: 37</p>	<p>24-h avg 1 2 3 0–1 avg 0–2 avg</p>	<p>Change in PEF (L/min) –0.73 (–2.5, 0.99) –0.99 (–2.6, 0.61) 0.34 (–1.1, 1.8) –1.8 (–3.8, 0.17) –1.5 (–3.4, 0.46)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>
<p>†Chang et al. (2012b) Taipei, Taiwan, 1996–1997 N = 2,919, ages 12–16 yr. Cross-sectional. Supervised spirometry. Recruited from schools.</p>	<p>Five monitors averaged within 2 km of schools Means across districts 4-h avg (8 a.m.–12 p.m.): 4.6–10 10-h avg (8 a.m.–6 p.m.): 1.8–5.4 1-h max: 5.9–35</p>	<p>4-h avg 0 10-h avg 1 1-h max 0 1</p>	<p>Change in FEV₁ (mL) 0.4 (–32, 33) –117 (–193, –42) 3.6 (–21, 28) –85 (–129, –41)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>

Table 5-17 (Continued): Epidemiologic studies of lung function in healthy children and children in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Linares et al. (2010) Salamanca, Mexico, Mar 2004–Feb 2005 N = 464, ages 6–14 yr. 0.6% asthma. Daily measures for 20 days in each season. Supervised spirometry. Recruited from schools.</p>	<p>Monitors within 2 km of school Means spring–winter School 1: 12, 12, 10, 9.8 School 2: 9.1, 8.7, 10, 13</p>	<p>24-h avg 0</p>	<p>Units not reported FVC: –0.06 (–0.13, 0) FEV₁: –0.01 (–0.01, –0.00) PEF: –0.03 (–0.05, 0) FEV₁/FVC: –0.07 (–0.18, 0.03)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>
<p>†Reddy et al. (2012) Durban, South Africa, 2004–2005 N = 129, ages 9–11 yr. 37% asthma. Daily measures for 3 wk each season. Supervised spirometry. Recruited from schools. Did not examine confounding by meteorological factors except season.</p>	<p>Monitor at school Mean (SD): 5.8 (0.2) Max: 41</p>	<p>24-h avg 0–4 avg</p>	<p>Percent change FEV₁ diurnal variability (increase = poorer function) By <i>GSTM1</i> gene variant Null: –1.2 (–3.0, 0.54) Positive: 1.1 (0.45, 2.7)</p> <hr/> <p>3 By <i>GSTP1</i> gene variant AG/GG: 3.1 (1.6, 4.7) AA: –0.73 (–2.2, 0.70)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>
<p>†Makamure et al. (2016a) Durban, South Africa, 2004–2005 N = 71, ages 9–11 yr. 35% asthma. Part of the same cohort as Reddy et al. (2012) above. Daily measures for 3 wk each season. Supervised spirometry. Recruited from schools. Did not examine confounding by meteorological factors except season.</p>	<p>Monitor at school Mean (SD): 5.8 (0.2) Max: 41</p>	<p>24-h avg 1</p>	<p>Percent change FEV₁ diurnal variability (increase = poorer function) All subjects: 1.6 (–0.03, 3.3) By <i>CD14</i> gene variant CC: –1.5 (–3.4, –0.37) CT/TT: –3.6 (–7.1, –0.17)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>

Table 5-17 (Continued): Epidemiologic studies of lung function in healthy children and children in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Makamure et al. (2016b) Durban, South Africa, 2004–2005 N = 104, ages 9–11 yr. 39% asthma. Part of the same cohort as Reddy et al. (2012) above. Daily measures for 3 wk each season. Supervised spirometry. Recruited from schools. Did not examine confounding by meteorological factors except season.</p>	<p>Monitor at school Mean (SD): 5.8 (0.2) Max: 41</p>	24-h avg	Percent change FEV ₁ diurnal variability (increase = poorer function)	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>
		1	By TNF-α gene variant AA/GA: 2.3 (–0.29, 5.0) GG: 0.83 (–1.32, 3.0)	
		2	AA/GA: 2.7 (0.52, 4.8) GG: 0.24 (–1.19, 1.68)	
<p>†Amadeo et al. (2015) Pointe-à-Pitre, Guadeloupe, 2008–2009 N = 354, ages 8–13 yr. 17% asthma. Cross-sectional. Supervised spirometry. Recruited from schools.</p>	<p>Monitors in city Number and distance NR Mean (SD): 1.8 (1.4) Max: 4.9</p>	1-h max	All subjects	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>
		0	Percent change post 6-min run 43 (–3,787, 3,873)	
		24-h avg 0–13 avg	Children without asthma Change in pre-run PEF (L/min) 18 (–84, 119) Percent change post 6-min run 4.5 (–24, 33)	

Avg = average; CI = confidence interval; CO = carbon monoxide; FEV₁ = forced expiratory volume in 1 sec; FVC = forced vital capacity; ISA = Integrated Science Assessment; max = maximum; MMEF = maximum midexpiratory flow; N = population number; n = sample size; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OR = odds ratio; PEF = peak expiratory flow; PM_{2.5} = particulate matter with nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with nominal aerodynamic diameter less than or equal to 10 μm; r = correlation coefficient; SD = standard deviation; SO₂ = sulfur dioxide; TNF-α = tumor necrosis factor-alpha; URS = upper respiratory symptoms.

^aEffect estimates are standardized to a 10-ppb increase in 1-h to 24-h avg SO₂ or a 40-ppb increase in 1-h max SO₂.

† = Studies published since the 2008 ISA for Sulfur Oxides.

For SO₂ measured at schools, there is no evidence for association with lung function in groups of children without respiratory disease or symptoms in Turkey or Brazil ([Altuğ et al., 2014](#); [Correia-Deur et al., 2012](#)). [Altuğ et al. \(2014\)](#) examined only 1-wk avg SO₂, but [Correia-Deur et al. \(2012\)](#) was noteworthy for examining multiple averaging times and lags (i.e., 3- to 10-day avg). PEF, which also was measured at school, was not associated with SO₂ concentrations averaged over the preceding 2 hours [−0.24% change (95% CI: −1.4, 0.96) in PEF per 10-ppb increase in SO₂]. Another strength of this study over similar ones is its repeated-measures design and clinical assessment of children’s respiratory health status. Among the studies of school SO₂, an association with lung function was observed in another cohort of children from Brazil ([Castro et al., 2009](#)). The impact of the 18% of children with asthma on these results is unknown. The effect estimate was largest for 2-day avg SO₂ concentrations and imprecise for lag 1 and 2 ([Table 5-17](#)). Missing SO₂ concentration data for 52% of days could be one reason for the imprecision.

Some results for SO₂ measured at children’s schools have more ambiguous implication ([Makamure et al., 2016a, b](#); [Altuğ et al., 2013](#); [Reddy et al., 2012](#)) ([Table 5-17](#)). For children in Turkey, lung function was analyzed dichotomously based on a cutpoint of 85 or 75% of the predicted value ([Altuğ et al., 2013](#)). Healthy children may not experience such decrements, and the 7% of the cohort with asthma may influence results. In a South African cohort, results were in opposing directions across the many comparisons made among lung function parameters, pollutants, exposure lags, and gene variants ([Makamure et al., 2016a, b](#); [Reddy et al., 2012](#)). For example, an association of SO₂ and FEV₁ was found in children with the *GSTP1* variant but not in children with the *GSTM1* variant ([Table 5-17](#) and [Section 6.4](#)). Confounding by meteorology was not considered in either cohort.

For exposures estimated from fixed-site monitors, lung function associations were inconsistent for 1-h max SO₂ ([Amadeo et al., 2015](#); [Chang et al., 2012b](#)), which may be more variable within a community and if uncharacterized, is subject to greater exposure error. For children in Taiwan, a 40-ppb increase in 1-h max SO₂ lagged 1 day was associated with a −85 mL (95% CI: −129, −41) change in FEV₁ ([Chang et al., 2012b](#)). SO₂ concentrations were averaged from five monitors within 2 km of children’s schools. For children in Guadeloupe, West Indies, the distance to monitors was not reported. Daily 1-h max SO₂ concentrations were not associated with PEF ([Amadeo et al., 2015](#)). Although PEF was measured before and after a 6-minute exercise period, similar to procedures in controlled human exposure studies, the SO₂ metric was not likely matched temporally with PEF measurements. Lung function in populations of children with low or no prevalence of asthma was inconsistently associated with 24-h avg SO₂ measured at fixed-site monitors ([Amadeo et al., 2015](#); [Linares et al., 2010](#)), although the null findings

are for 13-day avg SO₂ ([Amadeo et al., 2015](#)). Airway responsiveness increased with increases in 24-h avg SO₂ in a population of children in which 8% had asthma and 18% had atopy ([Soyseth et al., 1995](#)). SO₂ exposures were estimated from monitors within 2 km of homes, which is similar to studies observing associations with 24-h avg and 1-h max SO₂ ([Chang et al., 2012b](#); [Linares et al., 2010](#)).

For the few associations observed for SO₂ with lung function or airway responsiveness, the potential for copollutant confounding or interactions was not addressed, including the study conducted near an aluminum smelter that also emitted PM ([Soyseth et al., 1995](#)). Associations were observed for PM₁₀, CO, NO₂, and O₃ measured at schools and fixed-site monitors, but neither correlations with SO₂ nor copollutant model results were reported ([Chang et al., 2012b](#); [Linares et al., 2010](#); [Castro et al., 2009](#)). [Altuğ et al. \(2014\)](#) reported a moderate correlation with NO₂ of 0.49 and observed no association for either NO₂ or SO₂. Copollutant models were analyzed for long-term SO₂, which was not associated with lung function decrements in single-pollutant models ([Linares et al., 2010](#)). Importantly, none of the studies examined PM_{2.5}.

Animal Toxicological Studies

Lung function was examined in numerous studies reported in the 1982 SO_x AQCD ([U.S. EPA, 1982a](#)) and the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). The majority of these were conducted in naive animals rather than in animal models of allergic airway disease. Bronchoconstriction, indicated by increased pulmonary resistance, was identified as the most sensitive indicator of lung function effects of acute SO₂ exposure, based on the observation of increased pulmonary resistance in guinea pigs acutely exposed to 0.16 ppm SO₂ ([U.S. EPA, 2008d, 1982a](#)). The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) reported a few additional studies conducted at concentrations below 2 ppm. Animal toxicological studies examining lung function changes in naive animals exposed to SO₂ are summarized in [Table 5-18](#). Increased pulmonary resistance and decreased dynamic compliance were observed in conscious guinea pigs exposed to 1 ppm SO₂ for 1 hour ([Amdur et al., 1983](#)). Effects were seen immediately after exposure and were not present 1 hour post-exposure. No changes in tidal volume, minute volume, or breathing frequency were found. These same investigators also exposed guinea pigs to 1 ppm SO₂ for 3 hours/day for 6 days ([Conner et al., 1985](#)). No changes were observed in lung function or respiratory parameters (i.e., diffusing capacity for CO, functional reserve capacity, vital capacity, total lung capacity, respiratory frequency, tidal volume, pulmonary resistance, or pulmonary compliance). In another study, [Barthelemy et al. \(1988\)](#) demonstrated a 16% increase in airway resistance following a 45-minute exposure of anesthetized rabbits to 0.5 ppm SO₂ via an endotracheal tube. This latter exposure is more relevant to oronasal than to nasal breathing.

Table 5-18 Study-specific details from animal toxicological studies of lung function.

Study	Species (Strain); n; Sex; Lifestage/Age or Weight	Exposure Details (Concentration; Duration)	Endpoints Examined
Amdur et al. (1983)	Hartley guinea pig; n = 8–23/group; M; age NR; 200–300 g	~1 ppm (2.62 mg/m ³); head only for 1 h	Endpoints examined during exposure and up to 1 h post-exposure. Lung function—pulmonary resistance, dynamic compliance, breathing frequency, tidal volume, and minute volume
Conner et al. (1985)	Hartley guinea pig; n ≤ 18/group/time point; M; age NR; 250–320 g	1 ppm (2.62 mg/m ³); nose only for 3 h/day for 6 days	Endpoints examined 1, 24, and 48 h after the sixth exposure. Lung function—residual volume, functional residual capacity, vital capacity, total lung capacity, respiratory frequency, tidal volume, pulmonary resistance, pulmonary compliance, diffusing capacity for CO, and alveolar volume
Barthelemy et al. (1988)	Rabbit; n = 5–9/group; sex NR; adult; mean 2.0 kg; rabbits were mechanically ventilated	0.5 ppm (1.3 mg/m ³) for 45 min; intratracheal	Endpoints examined 5 min before and up to 1 h post-exposure. Lung function—pulmonary resistance
Amdur et al. (1988)	Guinea pig; n = 8	1 ppm for 1 h	Endpoints examined 2 h following exposure Airway responsiveness to acetylcholine
Riedel et al. (1988)	Guinea pigs (Perlbright-White); n = 5–14; M; age NR; 300–350 g	0.1, 4.3, and 16.6 ppm whole body; 8 h/day for 5 days Animals were sensitized to ovalbumin (ovalbumin aerosol) on the last 3 days of exposure Bronchial provocation every other day with aerosolized 0.1% ovalbumin began at 1 wk after the last exposure to SO ₂ and continued for 14 days 4 groups: Control 0.1 ppm SO ₂ 4.3 ppm SO ₂ 16.6 ppm SO ₂	Bronchial obstruction determined by examination of the respiratory loop measured by whole-body plethysmography in spontaneously breathing animals after each bronchial provocation

Table 5-18 (Continued): Study specific details from animal toxicological studies of lung function.

Study	Species (Strain); n; Sex; Lifestage/Age or Weight	Exposure Details (Concentration; Duration)	Endpoints Examined
Park et al. (2001)	Guinea pigs (Dunkin-Hartley); n = 7–12/group; M; age NR; 250–350 g	0.1 ppm whole body; 5 h/day for 5 days Animals were sensitized to ovalbumin (0.1% ovalbumin aerosol) on the last 3 days of exposure Bronchial challenge with 1% ovalbumin aerosol occurred at 1 wk after the last exposure to SO ₂ 4 groups: Control Ovalbumin	Bronchial obstruction—measurement of Penh by whole-body plethysmography

CO = carbon monoxide; M = males; n = sample size; NR = not reported; Penh = enhanced pause; SO₂ = sulfur dioxide.

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) also described studies that examined airway responsiveness following SO₂ exposure. In several different animal species, a single exposure to SO₂ at a concentration up to 10 ppm failed to increase airway responsiveness to a challenge agent. These studies were mainly conducted in naive animals rather than in models of allergic airways disease. Only one was conducted at a SO₂ concentration of less than 2 ppm. This study found no change in airway responsiveness to acetylcholine measured 2 hours following a 1-hour exposure in guinea pigs to 1 ppm SO₂ ([Amdur et al., 1988](#)). However, two toxicological studies ([Park et al., 2001](#); [Riedel et al., 1988](#)) described in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), provided evidence that repeated SO₂ exposure of guinea pigs to concentrations as low as 0.1 ppm resulted in increased airway responsiveness following subsequent sensitization and challenge with ovalbumin.

Summary of Lung Function Changes in General Populations and Healthy Individuals

Across disciplines, there is limited evidence that short-term SO₂ exposure induces lung function changes in healthy people. Evidence from controlled human exposure studies of healthy individuals shows that transient decreases in lung function can occur at concentrations of 1.0 ppm SO₂ under exercising or forced oral breathing conditions, but the evidence is more consistent for exposures >1.0 ppm. Animal toxicological studies demonstrated that acute exposure of guinea pigs to 0.16–1.0 ppm SO₂ results in increased airway resistance, and provided evidence that repeated exposure of guinea pigs to concentrations of SO₂ as low as 0.1 ppm led to increased airway responsiveness following sensitization and challenge with an allergen. Epidemiologic studies do not clearly indicate SO₂-associated decreases in lung function in healthy adults or children or

groups from the general population with varying prevalence of respiratory disease. Results are mixed for SO₂ measured at subjects' locations and at fixed-site monitors. Similar to experimental studies in healthy humans and animals without allergen challenge plus 1- to 6-hour SO₂ exposures, epidemiologic findings are mixed for temporally resolved metrics such as 1-h max or 1- to 4-h avg SO₂. Associations were observed for populations living in locations with steel, aluminum, or petrochemical industry or coal heating, but SO₂ was one of many pollutants implicated.

Respiratory Symptoms in General Populations and Healthy Individuals

Respiratory symptoms in relation to short-term SO₂ exposure have been investigated in a limited number of studies of general populations or healthy individuals. The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) described some controlled human exposure and epidemiologic studies of respiratory symptoms among children or adults without asthma. Most controlled human exposure studies reported no respiratory symptoms at concentrations up to 2.0 ppm. Evidence from both previous and recent epidemiologic studies is inconsistent.

Controlled Human Exposure Studies

Controlled human exposure studies examining respiratory symptoms in healthy individuals exposed to SO₂ are summarized in [Table 5-15](#). Briefly, [Tunnicliffe et al. \(2003\)](#) found no association between respiratory symptoms (i.e., throat irritation, cough, and wheeze) and 1-hour exposures at rest to 0.2 ppm SO₂ in either healthy adults or those with asthma. Similarly, [Andersen et al. \(1974\)](#) reported no change in respiratory symptoms in resting adults exposed to 1.0 ppm SO₂ for 6 hours. A more recent study in which exercising healthy adults were exposed to SO₂ concentrations as high as 2.0 ppm for 4 hours confirms these null findings ([van Thriel et al., 2010](#)).

Epidemiologic Studies

Associations for ambient SO₂ with respiratory symptoms in populations of healthy adults and children are inconsistent. Most results are from Europe and Asia. There are more studies of children than adults, but studies of adults focus on healthy individuals. Many previous studies of children examined populations in which 5-81% of the children had chronic wheeze, asthma, or atopy ([Boezen et al., 1999](#); [Neas et al., 1995](#)). Some recent studies examine populations of children with low (0.6–4%) prevalence of respiratory disease, but like previous studies do not consistently associate increases in SO₂ concentrations with respiratory symptoms ([Table 5-19](#)). Previous results were largely based on 24-h avg SO₂ concentrations measured at fixed-site monitors. Many recent studies have improved exposure assessment, examining temporally resolved 1-hour SO₂

concentrations for adults or SO₂ concentrations at children's schools. These associations with respiratory symptoms also are inconsistent. Other uncertainties include confounding by meteorological factors and copollutants.

For adults, a study on Miyakejima Island, Japan 5 years after a volcano eruption provided information on effects related to SO₂ concentrations and durations comparable to those examined in experimental studies ([Ishigami et al., 2008](#)). Compared to the reference group (those living near monitors with 1-h avg SO₂ concentrations below 10 ppb), incidence of many symptoms increased among people living near monitors with concentrations above 100 ppb ([Table 5-19](#)). Although temporally resolved metrics were analyzed, inference about an SO₂ effect is weak. SO₂ concentrations were measured within 2 km of volunteer workers' homes and work sites, no other air pollutants or other potential confounders were examined, and 80% of concentrations were in the reference category. Results linking long-term air pollution from volcanoes to respiratory symptoms also are uncertain because they are based on ecological comparisons of areas with low and high air pollution mixtures in which SO₂ is one constituent ([Section 5.2.2.1](#)).

Table 5-19 Epidemiologic studies of respiratory symptoms in healthy adults and children and groups in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Models	SO ₂ Copollutant Model Results and Correlations
Adults				
†Ishigami et al. (2008) Miyakejima Island, Japan, 2005 N = 611, ages ≥15 yr, 100% healthy Daily diaries for 1–15 days. Recruited from volunteers working on an active volcanic island 5 yr after eruption. Did not examine potential confounding factors	Monitors within 2 km of residence/work area Means across monitors 0–3,550 Max across monitors 3,790–10,320	1-h avg	Cough crude incidence rate, males <10 ppb: 4.8, 10–20 ppb: 1.4, 20–30 ppb: 2.9, 30–100 ppb: 6.6, >100 ppb: 19.3. <i>p</i> for trend <0.01	No copollutant model No copollutants examined
		1-h max	<10 ppb: 4.7, 10–20 ppb: 4.3, 20–60 ppb: 8.1, 60–2,000 ppb: 16.4, >2,000 ppb: 58.3. <i>p</i> for trend <0.01	
Children				
†Zhao et al. (2008) Taiyuan, China, Dec 2004 N = 1,993, ages 11–15 yr. 2% asthma. 4% with furry pet or pollen allergy. Cross-sectional. Recruited from schools. Likely temporal mismatch between current SO ₂ concentrations and symptoms assessed as any occurrence in preceding 12 mo.	Monitor at school Mean (SD) and max Outdoor: 271 (72), 386 Indoor: 101 (53), 244	24-h avg	Outdoor SO ₂	No copollutant model
		0–6 avg	Wheeze OR: 1.01 (0.98, 1.04) Daytime attacks of breathlessness OR: 0.99 (0.97, 1.01) Nocturnal attacks of breathlessness OR: 1.01 (0.96, 1.06) Indoor SO ₂ Wheeze OR: 1.04 (1.01, 1.08) Daytime attacks of breathlessness OR: 1.02 (0.99, 1.04) Nocturnal attacks of breathlessness OR: 1.07 (1.01, 1.13)	PM _{2.5} not examined. <i>r</i> = 0.74 NO ₂ .

Table 5-19 (Continued): Recent epidemiologic studies of respiratory symptoms in healthy adults and children and groups in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Altuğ et al. (2014) Eskisehir, Turkey, Feb–Mar 2007 N = 605, ages 9–13 yr. 7% asthma, 44% eczema. Cross-sectional. Recruited from schools from participants of a larger study.</p>	<p>Monitor at school Mean and max Suburban: 21, 29 Urban: 29, 44 Urban-traffic: 22, 27</p>	<p>24-h avg 0–6 avg</p>	<p>Complaints of the throat in last 7 days RR: 0.83 (0.59, 1.15) Complaints of the throat at the moment RR: 1.03 (0.72, 1.47) Runny nose in last 7 days RR: 0.95 (0.74, 1.22) Runny nose at the moment RR: 0.92 (0.69, 1.23) Shortness of breath/wheeze in last 7 days RR: 1.72 (1.05, 2.81) Medication for shortness of breath/wheeze in last 7 days RR: 1.44 (0.69, 2.99) Shortness of breath/wheeze today RR: 1.79 (0.90, 3.58) Medication for shortness of breath/wheeze today RR: 0.74 (0.16, 3.33)</p>	<p>No copollutant model PM_{2.5} not examined. <i>r</i> = 0.40 O₃, 0.49 NO₂.</p>
<p>†Linares et al. (2010) Salamanca, Mexico, Mar 2004–Feb 2005 N = 464, ages 6–14 yr. 0.6% asthma. Cross-sectional. Recruited from schools.</p>	<p>Monitors within 2 km of school Means spring–winter School 1: 12, 12, 10, 9.8 School 2: 9.1, 8.7, 10, 13</p>	<p>24-h avg 0</p>	<p>Wheezing OR: 1.06 (1.00, 1.11) Rhinorrhea OR: 0.98 (0.92, 1.05) Dyspnea OR: 1.02 (0.97, 1.07)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>

Table 5-19 (Continued): Recent epidemiologic studies of respiratory symptoms in healthy adults and children and groups in the general population published since the 2008 ISA for Sulfur Oxides.

Study Population and Methodological Details	SO ₂ Exposure Estimates (ppb)	SO ₂ Averaging Time and Lag Day	Effect Estimate (95% CI) Single-Pollutant Model ^a	SO ₂ Copollutant Model Results and Correlations
<p>†Moon et al. (2009) Seoul, Incheon, Busan, Jeju, South Korea, 2003 N = 696, ages <13 yr Daily diaries for 2 mo. Recruited from schools.</p>	<p>Monitors in city Number and distance NR Means NR Max: 38</p>	<p>24-h avg 0</p>	<p>LRS OR: 1.00 (0.93, 1.08) URS OR: 1.11 (1.03, 1.20)</p>	<p>No copollutant model PM_{2.5} not examined. Copollutant correlations NR.</p>

Avg = average; CI = confidence interval; ISA = Integrated Science Assessment; LRS = lower respiratory symptoms; max = maximum; n = sample size; N = population number; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OR = odds ratio; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; r = correlation coefficient; RR = relative risk or ratio; SD = standard deviation; SO₂ = sulfur dioxide; URS = upper respiratory symptoms.

^aEffect estimates are standardized to a 10-ppb increase in 1-h avg and 24-h avg SO₂ and a 40-ppb increase in 1-h max SO₂.

†Studies published since the 2008 ISA for Sulfur Oxides.

For children, associations with SO₂ concentrations were inconsistent within studies among the array of symptoms examined ([Table 5-19](#)). Results across studies were consistent for wheeze, an asthma symptom that is less likely to be experienced by healthy children. The study in South Korea ([Moon et al., 2009](#)) has many limitations, including estimating SO₂ exposure from fixed-site monitors at an unspecified distance from children and observing only a few isolated associations among the numerous pollutants, symptoms, exposure lags, and cities examined. Other studies had cross-sectional design and measured SO₂ at school or within 2 km from school ([Altuğ et al., 2014](#); [Linares et al., 2010](#); [Zhao et al., 2008](#)). The study in China examined high SO₂ concentrations similar to those in the Japanese volcano study. Mean school SO₂ concentrations were 101 ppb indoors and 271 ppb outdoors. Indoor, but not outdoor, 1-wk avg SO₂ concentrations were associated with symptoms ([Zhao et al., 2008](#)) ([Table 5-19](#)). However, temporal mismatch is likely because measured SO₂ concentrations were compared with symptoms that may have appeared at any time in the 12 months preceding the measurement. The other study with 1-wk avg school SO₂ measures, conducted in Turkey, observed an association with any shortness of breath or wheeze in the previous 7 days but not throat symptoms, runny nose, or medication use concurrently or in the previous 7 days ([Altuğ et al., 2014](#)). It is not clear whether the single positive association applied to the entire population, the 7% with asthma, or 27% with hay fever. Among mostly healthy children (0.6% asthma) in Mexico, lag 0 SO₂ concentration was associated with wheeze, but SO₂ was measured up to 2 km from children's schools ([Linares et al., 2010](#)). SO₂ concentrations were not associated with runny nose or difficulty breathing.

For the few observations of SO₂-associated increases in respiratory symptoms in healthy adults and children, the potential for copollutant confounding was not examined. PM₁₀, CO, and formaldehyde were also associated with symptoms, but PM_{2.5} was not examined ([Table 5-19](#)). Most studies did not report copollutant correlations, and none examined copollutant models. Symptoms were not associated with outdoor NO₂ ([Altuğ et al., 2014](#); [Linares et al., 2010](#); [Zhao et al., 2008](#)), but an association was observed with indoor NO₂ ([Zhao et al., 2008](#)). Indoor school SO₂ and NO₂ were highly correlated ($r = 0.74$), and it is unclear to what extent the association with breathlessness can be attributed independently to SO₂, NO₂, or to a combined effect of those and other copollutants.

Summary of Respiratory Symptoms in General Populations and Healthy Individuals

There is little evidence for an effect of short-term SO₂ exposure on respiratory symptoms in healthy individuals. Controlled human exposure studies of healthy adults did not demonstrate effects for 1- to 6-hour SO₂ exposures up to 2 ppm, and epidemiologic findings are inconsistent for healthy adults and children. For epidemiologic studies, there is uncertain representativeness of SO₂ exposures estimated from fixed-site monitors.

However, as shown in recent studies, respiratory symptoms are also inconsistently associated with SO₂ measured at children's schools. A biological explanation for associations observed with 1-wk avg SO₂ concentrations is unclear. For associations observed with 1-h avg or max concentrations and the evidence overall, potential for confounding by PM_{2.5}, PM₁₀, NO₂, CO, and formaldehyde is not addressed.

Subclinical Respiratory Effects in Healthy Individuals

Pulmonary inflammation is a key subclinical effect in the pathogenesis of asthma and other respiratory diseases. It consists of both acute and chronic responses and involves the orchestrated interplay of the respiratory epithelium and both the innate and adaptive immune system. The immunohistopathologic features of chronic inflammation involve the infiltration of airway epithelium by inflammatory cells, such as eosinophils, lymphocytes, mast cells, and macrophages, and the release of inflammatory mediators like cytokines and leukotrienes. The 2008 ISA for Sulfur Oxides described limited evidence from animal toxicological studies for SO₂-induced pulmonary inflammation and allergic sensitization in rodents exposed to an allergen. A single new animal toxicological study also provides evidence of SO₂-induced pulmonary inflammation. Recent controlled human exposure and epidemiologic studies add to the evidence base but do not clearly support SO₂-related pulmonary inflammation in healthy populations.

Controlled Human Exposure Studies

A recent controlled human exposure study examined eNO and other biomarkers of pulmonary inflammation in the nasal lavage fluid (NALF) and EBC after exposures to 0, 0.5, 1, and 2 ppm SO₂ for 4 hours in exercising healthy adults ([Raulf-Heimsoth et al., 2010](#)). Data demonstrated no statistically significant changes in eNO; leukotriene B₄, prostaglandin E₂, and 8-iso-prostaglandin F₂ alpha in EBC; or substance P, interleukin-8, and brain derived neurotrophic factor in NALF after SO₂ exposures, compared to air.

Epidemiologic Studies

Unlike the study reviewed in the 2008 ISA for Sulfur Oxides ([Adamkiewicz et al., 2004](#)), recent studies measured SO₂ near subjects' homes, schools, or work. SO₂ concentrations at a site within 1 km of most homes were not associated with pulmonary inflammation in a population of children with high prevalence (33%) of asthma or atopy ([Chen et al., 2012a](#)). Previous results were similar for a population of older adults that included people with respiratory disease. Recent examination of healthy adults and children in Beijing, China indicates SO₂-associated increases in pulmonary inflammation or oxidative stress. These recent studies were conducted before, during, and after the 2008 Olympics ([Roy et al., 2014](#); [Lin et al., 2011b](#)). Concentrations of SO₂ and other pollutants were lower

during the Olympics than before or after (e.g., mean 24-h avg 3.0 vs. 7.5 and 6.8 ppb). During one period in winter 2007, mean 24-h avg SO₂ concentrations were 45 ppb ([Lin et al., 2011b](#)). Pollutants were measured 0.65 km from the school that the study children attended and the hospital where most of the study adults worked. A 10-ppb increase in lag 0 24-h avg SO₂ was associated with a 7.6% (95% CI: 5.9, 9.3) increase in eNO of children ([Lin et al., 2011b](#)) and, in adults, a 0.67 standard deviation (95% CI: 0.48, 0.86) increase in an index of pulmonary inflammation and oxidative stress combining eNO and EBC markers ([Roy et al., 2014](#)). Associations were also observed with PM_{2.5}, sulfate, EC/BC, CO, NO₂, and OC. Copollutant models were analyzed for children, and SO₂ effect estimates remained positive but decreased substantially with adjustment for PM_{2.5} or BC ([Lin et al., 2011b](#)). Conversely, the effect estimate for BC was robust to adjustment for SO₂. Correlations with SO₂ concentrations were not reported, and confounding by other copollutants was not examined.

Animal Toxicological Studies

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) described several animal toxicological studies that examined the effects of repeated exposure to SO₂ on inflammation. These and other animal toxicological studies examining inflammation in naive animals exposed to SO₂ are summarized in [Table 5-20](#). In two of these studies, repeated exposure to SO₂ was found to promote allergic sensitization and to enhance allergen-induced bronchial obstruction in guinea pigs. [Riedel et al. \(1988\)](#) examined the effect of SO₂ exposure on local bronchial sensitization to inhaled antigen. Guinea pigs were exposed by inhalation to 0.1, 4.3, and 16.6 ppm SO₂ for 8 hours/day for 5 days. During the last 3 days, SO₂ exposure was followed by exposure to nebulized ovalbumin for 45 minutes. Following bronchial provocation with inhaled ovalbumin (0.1%) 1 week later, bronchial obstruction was measured by examining the respiratory loop obtained by whole-body plethysmography. In addition, specific antibodies against ovalbumin were measured in serum and BALF. The results showed significantly higher bronchial obstruction in animals exposed to both SO₂, at all concentration levels, and ovalbumin compared with animals exposed only to ovalbumin. In addition, significant increases in antiovalbumin immunoglobulin G (IgG) antibodies were detected in BALF of animals exposed to 0.1, 4.3, and 16.6 ppm SO₂ and in serum from animals exposed to 4.3 and 16.6 ppm SO₂ and ovalbumin compared with controls exposed only to ovalbumin. These results demonstrated that repeated exposure to SO₂ enhanced allergic sensitization and bronchial obstruction in the guinea pig at a concentration as low as 0.1 ppm.

Table 5-20 Study-specific details from animal toxicological studies of subclinical effects.

Study	Species (strain); n; Sex; Lifestage/Age or Weight	Exposure Details (Concentration; Duration)	Endpoints Examined
Conner et al. (1989)	Guinea pigs (Hartley); n = 4; M; age NR; 250–300 g;	1 ppm nose only; 3 h/day for 1–5 days	BAL performed each day. BALF—total and differential cell counts
Riedel et al. (1988)	Guinea pigs (Perlbright-White); n = 5–14/group; M; age NR; 300–350 g;	0.1, 4.3, and 16.6 ppm whole body; 8 h/day for 5 days Animals were sensitized to ovalbumin (ovalbumin aerosol) on the last 3 days of exposure Bronchial provocation every other day with 0.1% ovalbumin aerosol began at 1 wk after the last exposure to SO ₂ and continued for 14 days Four groups: Control 0.1 ppm SO ₂ 4.3 ppm SO ₂ 16.6 ppm SO ₂	Endpoints examined 48 h after the last provocation. Serum—anti IgG levels BALF—anti IgG levels
Park et al. (2001)	Guinea pigs (Dunkin-Hartley); n = 7–12/group; M; age NR; 250–350 g;	0.1 ppm whole body; 5 h/day for 5 days Animals were sensitized to ovalbumin (0.1% ovalbumin aerosol) on the last 3 days of exposure Bronchial challenge with 1% ovalbumin aerosol occurred at 1 wk after the last exposure to SO ₂ Four groups: Control Ovalbumin SO ₂ Ovalbumin/SO ₂	Endpoints examined 24 h after the bronchial challenge. BALF—differential cell counts cells Lung and bronchial tissue—histopathology
Li et al. (2007)	Rats (Wistar); n = 6/group; M; age NR	2 ppm SO ₂ for 1 h/day for 7 days	Endpoints examined 24 h following the last exposure BALF—inflammatory cell counts Lung—histopathology and immunohistochemistry Lung and tracheal tissue—mRNA and protein levels of MUC5AC and ICAM-1

Table 5-20 (Continued): Study specific details from animal toxicological studies of subclinical effects

Study	Species (strain); n; Sex; Lifestage/Age or Weight	Exposure Details (Concentration; Duration)	Endpoints Examined
Li et al. (2014)	Rats (Wistar); n = 6/group; M; age NR; 180–200 g	2 ppm SO ₂ for 1 h/day for 7 days	Endpoints examined BALF—inflammatory cell counts and cytokines IL-4, IFN-γ, TNF-α, IL-6 Serum—IgE Lung—histopathology, Lung and tracheal tissue—mRNA and protein levels NFκB, IκBα, IKKβ, IL-6, IL-4, TNF-α, FOXP3, EMSA NFκB binding activity

BAL = bronchoalveolar lavage; BALF = bronchoalveolar lavage fluid; EMSA = electrophoretic mobility shift assay; FOXP3 = forkhead box p3; ICAM-1 = intercellular adhesion molecule 1; IFN-γ = interferon gamma; IgE = immunoglobulin E; IgG = immunoglobulin G; IKKβ = inhibitor of nuclear factor kappa-B kinase subunit beta; IL-4 = interleukin-4; IL-6 = interleukin-6; IκBα = nuclear factor of kappa light polypeptide gene enhancer in B-cells inhibitor, alpha; M = male; MUC5AC = mucin 5AC glycoprotein; n = sample size; NFκB = nuclear factor kappa-light-chain-enhancer of activated B cells; NR = not reported; SO₂ = sulfur dioxide; TNF-α = tumor necrosis factor alpha.

Similarly, guinea pigs were exposed to 0.1 ppm SO₂ for 5 hours/day for 5 days and sensitized with 0.1% ovalbumin aerosols for 45 minutes on Days 4 and 5 ([Park et al., 2001](#)). One week later, the animals were subjected to bronchial challenge with 0.1% ovalbumin, and lung function was evaluated 24 hours later by whole-body plethysmography. The results demonstrated a significant increase in enhanced pause (Penh), a measure of airway obstruction, in animals exposed to both SO₂ and ovalbumin but not in animals treated with ovalbumin or SO₂ alone. In animals treated with both SO₂ and ovalbumin, increased numbers of eosinophils were found in the lavage fluid. In addition, infiltration of inflammatory cells, bronchiolar epithelial cell damage, and plugging of the airway lumen with mucus and cells were observed in bronchial tissues. These cellular changes were not observed in animals treated with ovalbumin or SO₂ alone. Results indicate that repeated exposure to near-ambient levels of SO₂ enhanced the allergic inflammation due to subsequent sensitization and challenge with ovalbumin in the guinea pig. Furthermore, increases in bronchial obstruction suggest that SO₂ exposure increased airway responsiveness in the animals subsequently made allergic to ovalbumin.

[Park et al. \(2001\)](#) demonstrated that repeated exposure of guinea pigs to 0.1 ppm SO₂ alone did not lead to allergic inflammation or morphologic changes in the lung. [Conner et al. \(1989\)](#) found no changes in total cells and neutrophils in BALF from guinea pigs exposed repeatedly to 1 ppm SO₂. In contrast, repeated exposure of rats to 2 ppm SO₂ resulted in mild pathologic changes in the lung, including inflammatory cell influx and

smooth muscle hyperplasia ([Li et al., 2014](#); [Li et al., 2007](#)). Several other indicators of inflammation and immune response were not changed by exposure to SO₂ alone.

Summary of Subclinical Respiratory Effects in Healthy Individuals

There is limited evidence for inflammatory and other subclinical respiratory effects in healthy populations following short-term exposure to SO₂. As newly informed by recent studies, SO₂ is not clearly related to pulmonary inflammation in healthy populations in controlled human exposure or epidemiologic studies. Associations were observed in some epidemiologic studies, but confounding by PM_{2.5}, sulfate, BC, or NO₂ is not well addressed. Studies in animals demonstrated that repeated exposure of guinea pigs to 0.1 or 1 ppm SO₂ had no effect on inflammation. However, when followed by sensitization with an allergen, exposure of guinea pigs to 0.1 ppm SO₂ enhanced allergic inflammation and bronchial obstruction to an allergen challenge. These results point to the potential for SO₂ exposure to increase airway responsiveness to an allergen. In addition, repeated exposure of rats to 2 ppm SO₂ resulted in inflammation and smooth muscle hyperplasia, early indicators of airway remodeling.

Studies of Mixtures of Particles and Sulfur Oxides in Healthy Individuals

The 1982 AQCD ([U.S. EPA, 1982a](#)) addressed the question of possible effects of PM on the response to SO₂. It was noted that sorption of SO₂ onto liquid or solid particles, which may act as carriers, tended to increase its potency in animal toxicological experiments. However, the mechanism for the effect was not known. Between the 1982 AQCD and the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), additional animal studies demonstrated respiratory responses following inhalation of SO₂ that was adsorbed onto metal oxide or carbon particles. These studies are summarized in Annex Table E-4 of the 2008 ISA for Sulfur Oxides and confirm and extend earlier findings. In all of the more recent studies, the resulting particles were submicron in size and would be expected to deposit in the lower respiratory tract. Acute and subacute exposures to SO₂ and PM resulted in additive or more-than-additive effects on pulmonary resistance, diffusing capacity for CO, airway responsiveness following an acetylcholine challenge, and host defense responses. Many of these studies reported transformation of SO₂ to sulfite, sulfate, sulfur trioxide and sulfuric acid (H₂SO₄), depending on temperature and relative humidity. Respiratory responses observed in these experiments were in some cases attributed to the formation of particulate sulfur-containing species. For example, repeated exposure to 20 µg/m³ carbon black-associated sulfate resulted in impaired host defense. Some studies of laboratory-generated complex mixtures did not include a SO₂-only or a metal-only exposure group, making it difficult to determine the relative contribution of these species. In addition, the relevance of these animal toxicological studies has been called into

question because the concentrations of both PM (1 mg/m³ and higher) and SO₂ (1 ppm and higher) used in these studies are much higher than ambient levels. Furthermore, the SO₂-adsorbed PM in some of these studies is not representative of ambient PM. For example, some of the laboratory-generated aerosols contained sulfite but atmospheric chemistry studies do not indicate significant amounts of sulfite ion in ambient PM. In summary, animal toxicological studies suggest that SO₂ effects may be potentiated by coexposure to PM, but the relevance of these results to ambient exposures is not clear.

Summary of Respiratory Effects in General Populations and Healthy Individuals

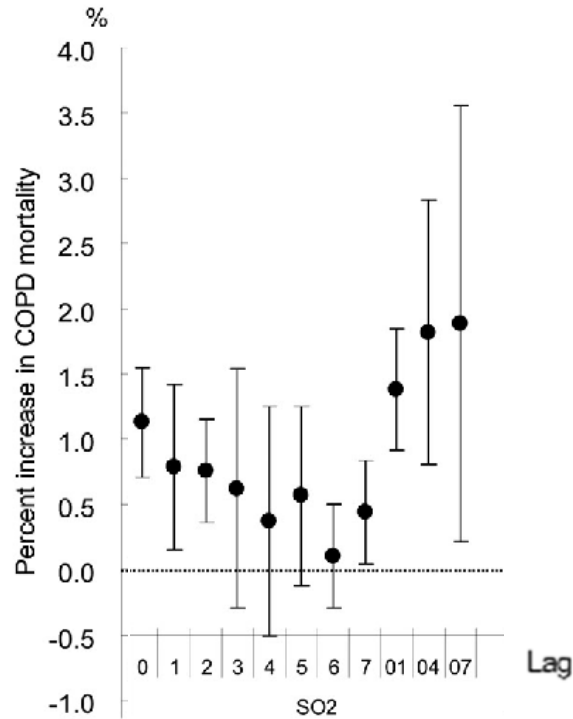
Overall, epidemiologic evidence for SO₂-induced lung function effects, symptoms, or inflammation in healthy individuals is weak. Epidemiologic studies do not clearly support associations of SO₂ with lung function in healthy children or adults. Associations of SO₂ with respiratory symptoms in children and pulmonary inflammation in healthy populations are not consistently observed. Decrements in lung function, but not increases in respiratory symptoms or inflammation were demonstrated in healthy individuals following SO₂ exposures ≥ 1.0 ppm in controlled human exposure studies. Animal toxicological studies demonstrated bronchoconstriction after a single SO₂ exposure. Repeated exposures to SO₂ were found to enhance allergic sensitization, allergic inflammation, and bronchial obstruction in response to a subsequent allergen challenge. In the absence of an allergen challenge, repeated exposure to SO₂ resulted in inflammation and smooth muscle hyperplasia, which is an early indicator of airway remodeling. Furthermore, studies of mixtures of particles and sulfur oxides indicate some enhanced effects on lung function parameters, airway responsiveness, and host defense. However, some of these studies lack appropriate controls and others involve species of sulfur oxides that may not be representative of ambient exposures.

5.2.1.8 Respiratory Mortality

Studies evaluated in the 2008 SO_x ISA that examined the association between short-term SO₂ exposure and cause-specific mortality found consistent positive associations with respiratory mortality using a 24-h avg exposure metric with some evidence indicating that the magnitude of the association was larger compared to all-cause and cardiovascular mortality. Recent multicity studies conducted in Asia ([Chen et al., 2012b](#); [Kan et al., 2010](#)) and Italy ([Bellini et al., 2007](#)), a meta-analysis of studies conducted in Asia ([Atkinson et al., 2012](#)), and a four-city study conducted in China that focused specifically on COPD mortality ([Meng et al., 2013](#)) add to the initial body of evidence indicating

increased respiratory mortality with short-term SO₂ exposure ([Section 5.5.1.3](#), [Figure 5-18](#)).

Studies evaluated in the 2008 SO_x ISA or in earlier documents [e.g., 1982 AQCD ([U.S. EPA, 1982a](#))] that examined the association between short-term SO₂ exposures and respiratory mortality focused exclusively on single-pollutant analyses. Therefore, questions arose regarding the independent effect of SO₂ on respiratory mortality, and whether associations remained robust in copollutant models. A few recent multicity studies conducted in China ([Meng et al., 2013](#); [Chen et al., 2012b](#)) and multiple Asian cities ([Kan et al., 2010](#)) examined both of these questions. [Chen et al. \(2012b\)](#) found that the SO₂-respiratory mortality association was attenuated, but remained positive in copollutant models with PM₁₀ [2.03% (95% CI: 0.89, 3.17) for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 0–1 days] and NO₂ [1.16% (95% CI: –0.03, 2.37) for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 0–1 days]. These results are similar to what the authors reported when examining the SO₂-total mortality association in models with PM₁₀ (i.e., ~40% reduction), but more attenuation was observed in models with NO₂ (i.e., ~80% reduction for total mortality and 65% reduction for respiratory mortality) ([Section 5.5.1.4](#)). [Kan et al. \(2010\)](#), as part of the Public Health and Air Pollution in Asia (PAPA) study, also examined the effect of copollutants (i.e., NO₂, PM₁₀, and O₃), but only in each city individually. The study authors found that although the SO₂-respiratory mortality association remained positive in copollutant models, there was evidence of an attenuation of the association in models with PM₁₀ and more so in models with NO₂ ([Figure 5-10](#)). [Meng et al. \(2013\)](#) in a four-city analysis of COPD mortality in China reported evidence consistent with [Chen et al. \(2012b\)](#) and [Kan et al. \(2010\)](#). The authors observed a 3.7% (95% CI: 2.4, 4.9) increase in COPD mortality for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 0–1 days. However, compared to the results for respiratory mortality from copollutant models reported in [Chen et al. \(2012b\)](#), [Meng et al. \(2013\)](#) found attenuation in models with PM₁₀, ~50% reduction [1.9% (95% CI: 0.3, 3.5)] and NO₂, ~99% reduction [0.0% (95% CI: –1.8, 1.9)] compared to the SO₂ results from the single pollutant model. The larger degree of attenuation of the SO₂-COPD mortality association in [Meng et al. \(2013\)](#), compared to respiratory mortality in [Chen et al. \(2012b\)](#) could be a reflection of the smaller sample size and smaller number of cities included in [Meng et al. \(2013\)](#). Overall, the studies that examined the potential confounding effects of copollutants on the SO₂-respiratory mortality relationship show results consistent with what has been observed for total mortality. However, the overall assessment of potential copollutant confounding remains limited, and it is unclear how the results observed in Asia are generalizable to other locations, specifically due to the unique air pollution mixture and higher concentrations observed in Asian cities.



COPD = chronic obstructive pulmonary disease; SO₂ = sulfur dioxide.
 Source: Adapted from [Meng et al. \(2013\)](#). Reprinted with permission of Elsevier.

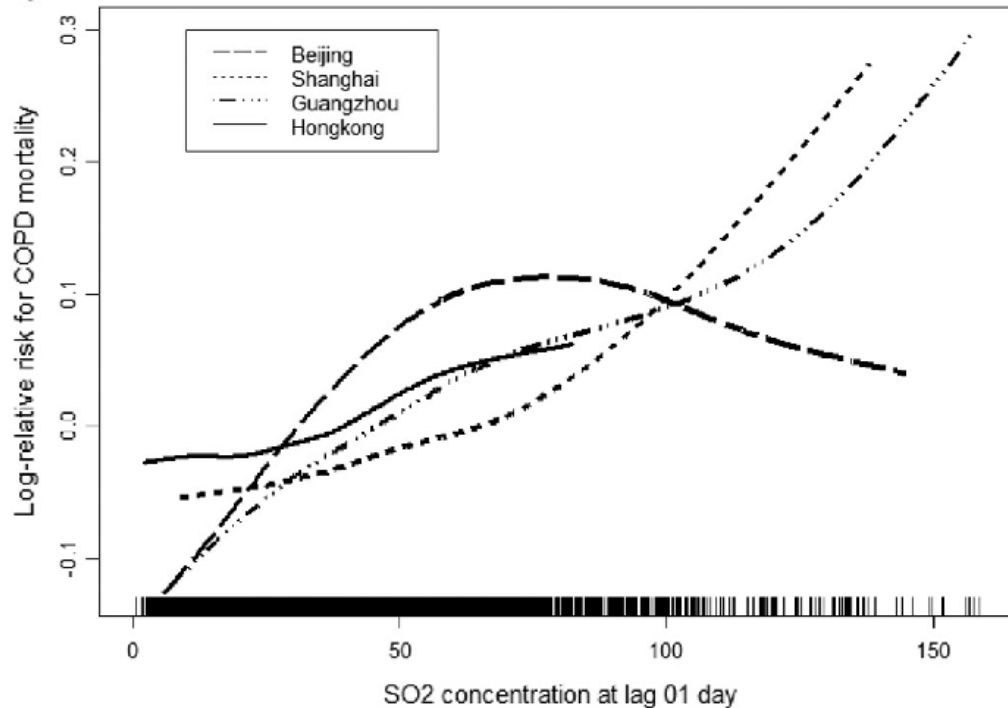
Figure 5-10 Percent increase in chronic obstructive pulmonary disease mortality associated with a 10 µg/m³ (3.62 ppb) increase in 24-h avg sulfur dioxide concentrations at various single and multiday lags in four Chinese cities.

Of the studies evaluated, only [Bellini et al. \(2007\)](#) (in a multicity study conducted in Italy) examined potential seasonal differences in the SO₂-cause-specific mortality relationship. [Bellini et al. \(2007\)](#) reported that risk estimates for respiratory mortality were dramatically increased in the summer compared to the winter from 4.1 to 12.0% for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 0–1 day (the all-year and winter were similar). These results are consistent with the seasonal pattern of SO₂ associations observed in [Bellini et al. \(2007\)](#) for total and cardiovascular mortality. However, it remains unclear whether this seasonal pattern of SO₂-respiratory mortality associations is observed in other locations.

An uncertainty that often arises when examining the relationship between short-term air pollution exposures and cause-specific mortality is whether the lag structure of associations and the C-R relationship is consistent with what is observed for total mortality. [Meng et al. \(2013\)](#) addressed both the lag structure of associations and the C-R

relationship in a study of short-term air pollution exposures and COPD mortality in four Chinese cities. Although not explicitly part of the China Air Pollution and Health Effects Study (CAPES) study, [Meng et al. \(2013\)](#) focused on four CAPES cities over the same time period as [Chen et al. \(2012b\)](#). In comparison to [Chen et al. \(2012b\)](#), who found a steady decline in risk estimates at single-day lags of 0 to 7 days with the largest effect at lag 0–1, [Meng et al. \(2013\)](#) observed a steady decline over single lag days, but some indication of larger associations, although highly uncertain, at longer multiday lags (i.e., 0–4 and 0–7 days) ([Figure 5-10](#)). Note that [Chen et al. \(2012b\)](#) did not examine multiday lags longer than 0–1 days, but the magnitude of the association for all respiratory mortality [3.3% (95% CI: 2.1, 4.6) for a 10-ppb increase in 24-h avg SO₂ concentrations] is similar to that reported in [Meng et al. \(2013\)](#) for COPD [3.7% (95% CI: 2.4, 4.9)].

[Meng et al. \(2013\)](#) also examined the shape of the SO₂-COPD mortality C-R relationship. To examine the assumption of linearity, the authors modeled the relationship between air pollution exposures and COPD mortality using a natural spline with 3 df. [Meng et al. \(2013\)](#) then computed the difference between the deviance of the linear and spline models to assess whether there was evidence of nonlinearity in the SO₂-COPD relationship. As depicted in [Figure 5-11](#), the authors found no evidence that the spline model resulted in a better fit of the SO₂-mortality relationship compared to the linear model. However, the authors did not present confidence intervals for each of the C-R curves, which complicates the interpretation of the results.



COPD = chronic obstructive pulmonary disease; SO₂ = sulfur dioxide.
 Source: Adapted from [Meng et al. \(2013\)](#).

Figure 5-11 City-specific concentration-response curves for short-term sulfur dioxide exposures and daily chronic obstructive pulmonary disease mortality in four Chinese cities.

Overall, recent multicity studies report evidence of consistent positive associations between short-term SO₂ concentrations and respiratory mortality, which is consistent with those studies evaluated in the 2008 SO_x ISA. Unlike studies evaluated in the 2008 SO_x ISA, recent studies examined whether copollutants confound the relationship between short-term SO₂ concentrations and respiratory mortality. Overall, these studies reported evidence that the SO₂-respiratory mortality association was attenuated in models with NO₂ and PM₁₀, but the analyses are limited to Asian cities where the air pollution mixture and concentrations are different from those reported in other areas of the world. Additional analyses focusing on seasonal patterns of associations, lag structure of associations, and the C-R relationship are limited in number, but suggest evidence of: larger associations in the summer/warm season, larger and more precise associations at shorter lag periods (in the range of 0 and 1 days), and a log-linear, no-threshold C-R relationship, respectively. However, for both total and cause-specific mortality, the studies have conducted a rather limited exploration of potential alternatives to linearity when examining the shape of the C-R relationship, which in combination with the

potential measurement error due to uncharacterized spatial and temporal variability in SO₂ concentrations, complicates the interpretation of the SO₂-mortality C-R relationship ([Sections 3.4.2.2](#) and [3.4.2.3](#)).

5.2.1.9 Summary and Causal Determination

Strong evidence indicates that there is a causal relationship between short-term SO₂ exposure and respiratory effects, particularly for respiratory effects in the at-risk population of individuals with asthma ([Section 6.3.1](#)). This determination is based on the consistency of SO₂-induced bronchoconstriction in exercising individuals with asthma in controlled human studies, coherence of asthma-related effects among multiple lines of evidence, and biological plausibility for effects specifically related to asthma exacerbation. There is limited support for a relationship between short-term SO₂ exposure and other respiratory effects, including exacerbation of COPD, allergy exacerbation, respiratory infection, respiratory effects in healthy populations, and respiratory mortality. The limited and inconsistent evidence for these nonasthma-related respiratory effects does not substantially contribute to the causal determination.

The determination of a causal relationship is the same as the conclusion of the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). This conclusion is substantially based on controlled human exposure studies that showed lung function decrements and respiratory symptoms in adult individuals with asthma exposed to SO₂ for 5–10 minutes under increased ventilation conditions. These findings are consistent with the current understanding of biological plausibility described in the mode of action section ([Section 4.3.6](#)). Previous epidemiologic studies provided supporting evidence indicating associations between short-term increases in ambient SO₂ concentration and respiratory-related ED visits and hospital admissions as well as respiratory symptoms. The evidence for a causal relationship is detailed below using the framework described in the Preamble to the ISAs ([U.S. EPA, 2015b](#)). While new evidence adds to the existing body of evidence, the determination remains largely based on previous controlled human exposure studies. The key evidence as it relates to the causal framework is presented in [Table 5-21](#).

Table 5-21 Summary of evidence for a causal relationship between short-term sulfur dioxide exposure and respiratory effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Asthma exacerbation			
Consistent evidence from multiple, high-quality controlled human exposure studies rules out chance, confounding, and other biases	Decreased lung function following exposures of 5–10 min in exercising individuals with asthma	Section 5.2.1.2 Table 5-2	400–600 ppb
	A group of responders (defined as having ≥15% decrease in FEV ₁ after exposure to 0.6 or 1.0 ppm SO ₂) showed statistically significant decrements in FEV ₁ following 5–10 min of exposure to 0.3 ppm SO ₂	Section 5.2.1.2 Table 5-3	300 ppb
	Decreased lung function following exposures of 5–10 min in 5–30% of exercising individuals with asthma	Section 5.2.1.2 Table 5-2	200–300 ppb
	Increased respiratory symptoms following exposure of 5–10 min in exercising individuals with asthma	Section 5.2.1.2 Table 5-2	400–1,000 ppb
Generally supporting evidence from multiple epidemiologic studies at relevant SO ₂ concentrations	Increase in asthma hospital admissions and ED visits in single- and multicity studies, among all ages, children and older adults	Section 5.2.1.2	1-h max: 9.6–10.8 ppb 24-h avg: 1.03–36.9 ppb
	Limited evidence for respiratory symptoms in children with asthma with school and/or home SO ₂ measurements judged to well represent subjects' exposure	†Spira-Cohen et al. (2011) , †Velická et al. (2015) Sections 5.2.1.2, 3.5	24-h avg: median 4.0 ppb
Uncertainty regarding exposure measurement error	SO ₂ exposures estimated from fixed-site monitors may not capture spatiotemporal variability of SO ₂ across a community	Section 3.4.2	
Uncertainty regarding potential copollutant confounding	Some SO ₂ associations were relatively unchanged in magnitude in copollutant models with NO ₂ , PM _{2.5} , or PM ₁₀ . Others were attenuated. Uncertainty in extent to which exposure measurement error is comparable for SO ₂ and copollutants. SO ₂ showed a wide range of correlations with copollutants across studies ($r = 0.4$ – 0.9).	Attenuated: †Spira-Cohen et al. (2011) Sections 5.2.1.2, 3.4.3	

Table 5-21(Continued): Summary of evidence for a causal relationship between short term sulfur dioxide exposure and respiratory effects

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Evidence for key events in proposed mode of action	Neural reflexes and/or inflammation lead to bronchoconstriction	Section 4.3.6	
	Increased airway eosinophils in adults with asthma exposed to SO ₂ Enhanced allergic inflammation in rats previously sensitized with an allergen and then repeatedly exposed to SO ₂	Gong et al. (2001) , Li et al. (2007) , Li et al. (2014)	750–2,000 ppb
	Enhancement of allergic sensitization, allergic inflammation and airway obstruction in guinea pigs exposed to SO ₂ repeatedly over several days and subsequently sensitized and challenged with an allergen	Park et al. (2001) , Riedel et al. (1988)	100 ppb
	Allergic inflammation leads to increased airway responsiveness. Association with increased AHR in a population of adults with asthma and a high prevalence of atopy	Taggart et al. (1996)	24-h avg: max 39 ppb
Other respiratory effects			
Limited and inconsistent evidence across disciplines and outcomes	Inconsistent evidence for allergy exacerbation, COPD exacerbation, respiratory infection, respiratory diseases, hospital admissions and ED visits, and respiratory effects in healthy individuals	Section 5.2.1.3 , Section 5.2.1.4 , Section 5.2.1.5 , Section 5.2.1.6 , and Section 5.2.1.7	
Respiratory mortality			
Consistent epidemiologic evidence from multiple studies at relevant SO ₂ concentrations	Increases in respiratory mortality in multicity studies conducted in the U.S., Canada, Europe, and Asia	Section 5.2.1.8 and Section 5.2.1.3 Figures 5-8 and 5-16	Mean 24-h avg: U.S., Canada, Europe: 0.4–28.2 ^d ppb Asia: 0.7–>200 ppb Table 5-39
Uncertainty regarding potential confounding by copollutants	No copollutant models with PM _{2.5} . SO ₂ associations remained positive but decreased in magnitude with adjustment for PM ₁₀ or NO ₂ , suggesting confounding. Studies limited to areas with high SO ₂ concentrations, which complicates the interpretation of independent association for SO ₂ .	Sections 5.2.1.8, 3.4.3	

Table 5-21(Continued): Summary of evidence for a causal relationship between short term sulfur dioxide exposure and respiratory effects

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Uncertainty regarding exposure measurement error	SO ₂ exposures estimated from fixed-site monitors may not capture spatiotemporal variability of SO ₂ across a community.	Section 3.4.2	

Avg = average; COPD = chronic obstructive pulmonary disease; ED = emergency department; FEV₁ = forced expiratory volume in 1 second; max = maximum; NO₂ = nitrogen dioxide; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; r = correlation coefficient; SO₂ = sulfur dioxide.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)).

^bDescribes the key evidence and references, supporting or contradicting, contributing most heavily to causal determination and, where applicable, to uncertainties or inconsistencies. References to earlier sections indicate where the full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated (for experimental studies, below 2,000 ppb).

^dThe value of 28.2 represents the median concentration from [Katsouyanni et al. \(1997\)](#).

†Studies published since the 2008 ISA for Sulfur Oxides.

Evidence for Asthma Exacerbation

A causal relationship between short-term SO₂ exposure and respiratory effects is primarily supported by evidence from controlled human exposure studies of respiratory effects in adults with asthma. These studies consistently demonstrated that the majority of individuals with asthma experience a moderate or greater decrement in lung function, as defined by a $\geq 100\%$ increase in sRaw or $\geq 15\%$ decrease in FEV₁. This decrement is frequently accompanied by respiratory symptoms following exposures of 5–10 minutes, with elevated ventilation rates at concentrations of 0.4–0.6 ppm ([Johns et al., 2010](#); [Linn et al., 1990](#); [Linn et al., 1988](#); [Balmes et al., 1987](#); [Linn et al., 1987](#); [Horstman et al., 1986](#); [Linn et al., 1983b](#)). A fraction of the population with asthma (~5–30%) has also been observed to have decrements in lung function at lower SO₂ concentrations (0.2–0.3 ppm) ([Linn et al., 1990](#); [Linn et al., 1988](#); [Linn et al., 1987](#); [Bethel et al., 1985](#)). Although the degree of lung function decrements are considered moderate, they are less likely to be accompanied by respiratory symptoms at these lower concentrations ([Linn et al., 1990](#); [Linn et al., 1988](#); [Linn et al., 1987](#); [Roger et al., 1985](#); [Linn et al., 1983b](#)). A group of responders (defined as having $\geq 15\%$ decrease in FEV₁ after exposure to 0.6 or 1.0 ppm SO₂) showed statistically significant decrements in FEV₁ following 5–10 minute exposure to 0.3 ppm SO₂ ([Johns et al., 2010](#)) ([Table 5-3](#)). While SO₂-induced respiratory effects have been examined in individuals classified as having mild and moderate asthma, these individuals are relatively healthy. Thus, extrapolating to individuals with severe asthma is difficult because such individuals cannot be tested in an exposure chamber due to the severity of their disease. Therefore, it is unknown whether people with severe asthma are at increased risk of having respiratory effects due to short-term SO₂ exposure. The same may be said about children with asthma. There are no laboratory studies of children exposed to SO₂, but a number of studies have assessed airway responsiveness of children and adults exposed to the bronchoconstrictive stimuli methacholine. Based largely on those studies, school-aged children (~5–11 years of age), particularly boys and perhaps obese children, might be expected to have greater responses (i.e., larger decrements in lung function) following exposure to SO₂ than adolescents and adults.

The coherence of the epidemiologic findings ([Section 5.2.1.2](#)) is supporting evidence for a causal relationship. Epidemiologic evidence for lung function changes in adults and children with asthma is inconsistent. However, short-term increases in ambient SO₂ concentration are associated with increases in asthma hospital admissions and ED visits among all ages, children (i.e., <18 years of age) and older adults (i.e., 65 years of age and older) ([Figure 5-3](#)), as well as asthma symptoms in children ([Velická et al., 2015](#); [Spira-Cohen et al., 2011](#)). Epidemiologic associations between short-term increases in ambient

SO₂ concentration and respiratory mortality provide support for a potential continuum of effects between respiratory morbidity and respiratory mortality.

Most epidemiologic studies indicating associations between short-term SO₂ exposures and asthma exacerbation assigned exposure using SO₂ concentrations measured at fixed-site monitors. The use of fixed-site monitors to assign exposure, particularly to 1-h max SO₂, may introduce exposure measurement error if the spatiotemporal variability in SO₂ concentrations is not captured. The studies did not statistically correct for measurement error. A few recent results reduce the uncertainty with SO₂ measured or modeled at or near children's school or home ([Velická et al., 2015](#); [Spira-Cohen et al., 2011](#)). Additional uncertainty exists regarding potential copollutant confounding. In many studies, SO₂ was moderately to highly correlated with PM_{2.5}, larger sized PM, EC/BC, NO₂, and VOCs ($r = 0.4\text{--}0.9$). The few available results show association with sulfate. A small number of studies examined copollutant models. Some associations were relatively unchanged in magnitude after adjustment for a copollutant; others did not persist. However, inference from copollutant models is limited given potential differences in exposure measurement error for SO₂ compared to NO₂, CO, PM, and O₃ and in some studies, high copollutant correlations. Copollutant interactions are not well studied. Some controlled human exposure studies demonstrate increased asthma-related effects with coexposure to SO₂ and NO₂ or O₃. Limited epidemiologic evidence shows increased asthma-related effects with joint increases in SO₂ and copollutants but does not clearly show a joint association that is greater than a single-pollutant association.

There is supportive evidence for a relationship between short-term SO₂ exposure and both airway responsiveness and pulmonary inflammation. Limited epidemiologic evidence points to an association with increased AHR in a population of adults with asthma and a high prevalence of atopy ([Taggart et al., 1996](#)). [Gong et al. \(2001\)](#) demonstrated an increase in airway eosinophils in adults with asthma 2 hours after a 10-minute exposure to 0.75 ppm SO₂. This effect, along with bronchoconstriction, was attenuated by pretreatment with a leukotriene receptor antagonist. Other pharmacologic studies have demonstrated that inflammatory mediators play an important role in SO₂ exposure-induced bronchoconstriction in people with asthma ([Section 4.3.1](#)). Further support for an important role of airway inflammation, including allergic inflammation, is provided by animal toxicological studies of repeated SO₂ exposure in allergic animals that are used to model the asthmatic phenotype ([Li et al., 2014](#); [Li et al., 2007](#)). In addition, repeated exposure of naive animals promoted allergic sensitization and enhanced allergic inflammation ([Park et al., 2001](#); [Riedel et al., 1988](#)). Increases in bronchial obstruction also observed in these studies suggest that SO₂ exposure increased airway responsiveness in the animals subsequently made allergic to ovalbumin. These

latter studies point to a possible increased sensitivity to allergens following SO₂ exposure.

Evidence for Other Respiratory Effects

Epidemiologic studies demonstrate some associations of ambient SO₂ concentrations with hospital admissions and ED visits for all respiratory causes combined (Figure 5-9). While these results suggest that the respiratory effects of short-term SO₂ exposure could extend beyond exacerbation of asthma, evidence across disciplines is inconsistent and/or lacks biological plausibility for conditions such as allergy exacerbation (Section 5.2.1.3), COPD exacerbation (Section 5.2.1.4), and respiratory infection (Section 5.2.1.5). Where epidemiologic associations were found, potential copollutant confounding is uncertain. For COPD exacerbation, a controlled human exposure study demonstrated no effect of SO₂ exposure, and epidemiologic associations are inconsistent for lung function, respiratory symptoms, hospital admissions, and ED visits. Some evidence supports SO₂-associated increases in hospital admissions and ED visits due to respiratory infections. However, the lack of multiple studies examining the same respiratory infection outcome, inconsistent findings for self-reported infections in children, and the lack of evidence from controlled human exposure and animal toxicological studies produces uncertainty as to whether a relationship exists. Controlled human exposure studies in healthy individuals provide evidence for transient decreases in lung function with ≥1 ppm SO₂ exposures for 5–10 minutes under exercising or a forced oral breathing condition with no evidence for increased respiratory symptoms. Epidemiologic evidence is inconsistent for SO₂ associations with lung function, respiratory symptoms, and pulmonary inflammation in healthy children and adults.

Conclusion

The evidence integrated across disciplines supports a causal relationship between short-term SO₂ exposure and respiratory effects, particularly asthma exacerbation.

This determination is primarily based on decreased lung function and increased respiratory symptoms observed in controlled human exposure studies in adults with asthma. Epidemiologic studies of asthma hospital admissions and ED visits and asthma symptoms in children provide supporting evidence. There is also supportive evidence for a relationship of short-term SO₂ exposure with pulmonary inflammation and AHR. Evidence for an effect of SO₂ exposure on allergy exacerbation, COPD exacerbation, respiratory infection, respiratory effects in healthy populations, and respiratory mortality is inconsistent within and across disciplines and outcomes, and there is uncertainty related to potential confounding by copollutants. The limited and inconsistent evidence

for these nonasthma-related respiratory effects does not substantially contribute to the causal determination.

5.2.2 Long-Term Exposure

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) reviewed the epidemiologic and toxicological evidence for long-term exposure to SO₂ and respiratory effects and concluded that the evidence was inadequate to infer a causal relationship. Although some positive associations with asthma prevalence, bronchitis, symptoms, and lung function were observed among children, uncertainties made it difficult at that time to assess the evidence as a whole. Uncertainties related to assessing the consistency of findings across a diverse set of respiratory outcomes, the potential for exposure measurement error to influence results, and the lack of information available to assess the impact of copollutant confounding were cited in the document. The studies of long-term exposure to SO₂ and respiratory morbidity that were considered in the last review are found in Supplemental Table 5S-11 ([U.S. EPA, 2017c](#)). Animal toxicological studies of the effects of long-term exposure to SO₂, which were reviewed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), examined lung function, morphology, and host defense. Most of these studies involved SO₂ concentrations well above 2 ppm. Recent toxicological studies add to this database.

Both older and more recent epidemiologic and toxicological studies that evaluate the relationship between long-term SO₂ exposure and asthma ([Section 5.2.2.1](#)), allergy ([Section 5.2.2.2](#)), lung function ([Section 5.2.2.3](#)), respiratory infection ([Section 5.2.2.4](#)), other respiratory diseases ([Section 5.2.2.5](#)), and respiratory mortality ([Section 5.2.2.6](#)) are discussed below. Overall, the collective evidence is strengthened by recent epidemiologic studies reporting increases in asthma incidence among children and findings of animal toxicological studies that provide a pathophysiologic basis for the development of asthma.

Recent cohort studies of asthma incidence ([Nishimura et al., 2013](#); [Clark et al., 2010](#)) use a longitudinal design, a methodological enhancement over the cross-sectional studies of asthma prevalence available in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). In a recent study, [Ierodiakonou et al. \(2015\)](#), using a longitudinal design, provided the first epidemiological report relating SO₂ exposure to increases in AHR (decreases in PC₂₀) in children with asthma. Uncertainties related to exposure estimates based on SO₂ measurements from monitors combined by inverse distance weighting (see [Section 3.3.2.3](#)) may limit the inferences that can be made for these recent studies. The majority of other recent and earlier epidemiologic studies used cross-sectional designs evaluating prevalence. Results were generally positive, although the strength of the associations varied across studies.

The designs used (i.e., ecological, cross-sectional) limit the contribution of these studies to possible inferences about causality of relationships between long-term SO₂ exposure and respiratory effects. The caution expressed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) related to the limitation of attributing an independent effect to SO₂ (due to the relationship of SO₂ levels to PM levels) is still a concern. The evidence base does not include studies evaluating concentration-responses, and few studies provide copollutant model analyses. The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) found that animal toxicological studies did not provide sufficient evidence to assess the effects of long-term SO₂ exposure on lung function, morphology, or host defense. The one new subchronic animal toxicological study that is discussed in this review found effects of SO₂ exposure on airway responsiveness, airway remodeling, and allergic inflammation. Short-term toxicological studies also provide some evidence for these responses to SO₂ exposure.

5.2.2.1 Development and Severity of Asthma

Development of Asthma

Asthma is described by the National Heart, Lung, and Blood Institute ([NHLBI NAEP, 2007](#)) as a chronic inflammatory disease of the airways that develops over time. Pulmonary inflammation can increase airway responsiveness, resulting in bronchoconstriction (bronchial smooth muscle contraction), and in turn, episodes of shortness of breath, coughing, wheezing, and chest tightness. When symptoms progress to the stage that people seek medical treatment, a diagnosis of asthma can result. Epidemiologic studies of SO₂ used self- or parental report of a diagnosis to define asthma. Epidemiologic studies reviewed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) were limited to those with cross-sectional designs [Supplemental Table 5S-11 ([U.S. EPA, 2017c](#))]. The majority of these studies reported positive associations of long-term SO₂ exposure with asthma prevalence. A few recent longitudinal epidemiologic studies support associations with asthma incidence and provide coherent evidence for associations with respiratory symptoms in healthy populations. Uncertainty remains in the adequacy of SO₂ exposure estimates and copollutant confounding. However, some support for an effect of SO₂ exposure comes from a recent toxicological study providing evidence for increased airway responsiveness.

A strength of recent epidemiologic studies of asthma development is their longitudinal design (see [Table 5-22](#)). The follow-up of children over time to mark the first record of a physician diagnosis with no prior record of diagnosis can better characterize the temporal sequence between SO₂ exposure and the incidence of asthma. In this regard, longitudinal studies can better distinguish between onset of asthma and the exacerbation of asthma. In

a large multicity study (N = 4,320 from Chicago, IL, Bronx, NY, Houston, TX, San Francisco Bay Area, CA, and the territory of Puerto Rico), [Nishimura et al. \(2013\)](#) observed that for SO₂ exposures during the first year of life, the OR for asthma incidence was 0.95 (95% CI: 0.59, 1.47) per 5 ppb change. SO₂ exposure during the first 3 years of life produced an OR for asthma incidence of 1.16 (95% CI: 0.73, 1.84) per 5 ppb SO₂. SO₂ exposures were estimated using the inverse distance weighted average of the four monitors within 50 km of the subject's residence. Selection bias due to differential loss to follow-up is not an issue given the retrospective design.

Table 5-22 Selected epidemiologic studies of long-term exposure to SO₂ and the development of asthma and intervention studies/natural experiments.

Study/Population	Location Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates (95% CI) ^a
Longitudinal studies of the development of asthma				
†Nishimura et al. (2013) GALA II and SAGE II cohorts (Latinos and African Americans 8–21 yr) N = 4,320	Chicago, IL; Bronx, NY; Houston, TX; San Francisco Bay Area, CA; and the territory of Puerto Rico (2006–2011)	4.0	IDW avg of monitors within 50 km of residence; annual avg and concentration during first 3 yr of life Copolutant correlations: NR	0.95 (0.59–1.47)—annual avg 1.16 (0.74–1.84)—early life exposure Covariate adjustment: age, sex, ethnicity, and composite SES.
†Clark et al. (2010) British Columbia Birth Cohort (N = 37,401)	Southwest British Columbia 1999–2000	In utero Controls: 5.11 Cases: 5.22 1st yr of life: Controls: 5.22 Cases: 5.37	IDW avg of three monitors within 50 km of postal code centroid. Concentrations for in utero and 1st yr of life estimated. Copolutant correlations: NR	1.47 (1.30–1.89) (both in utero/1st yr of life) Covariate adjustment: native status, breast-feeding, maternal smoking, income quartile, birth weight, and gestational length

Table 5-22 (Continued) Selected epidemiologic studies of long term exposure to SO₂ and the development of asthma and intervention studies/natural experiments.

Study/Population	Location Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates (95% CI) ^a
<p>†Chiang et al. (2016a), Chiang et al. (2016b)</p> <p>Recruited 587 children aged between 11 and 14 yr from junior high schools in each of nine townships.</p> <p>N = 587</p> <p>Incidence rates for asthma (ICD-9; 493) were obtained from the Taiwan Health Insurance Database.</p>	<p>Taiwan, near a petrochemical complex which yields a diverse pollution mix.</p> <p>1999 to 2010</p>	<p>The 3-year average of the 99th percentile of SO₂ levels in high and low exposure areas after 2003 was 137.3 ppb and 32.0 ppb in the high and low exposure areas, respectively, between 2003 and 2006. From 2003 to 2010, there were 138 h with hourly SO₂ concentrations above 75 ppb each year in the high exposure areas and 2 h in low exposure areas.</p>	<p>Two air quality monitoring stations, part of the Taiwan Environmental Protection Administration (TEPA), provided the SO₂ levels in the HE and LE areas. One is located 8.1 km south of the complex, and the other 16.2 km east and south of the complex. Three exposure periods were reported since opening of the complex.</p> <p>Copollutant correlations NR.</p>	<p>The incidence rate of asthma in the HE group (18.5%) was significantly higher than that in the LE group (11.0%) in the first 4 yr after the complex began its operations. A difference in the incidence of asthma between the two groups emerged after 12 mo, and the maximum difference appeared at 40 mo. The hazard ratios of the incidences of asthma, during the different study periods were adjusted for group, age, gender, living near roads, incense burning and passive smoking exposure. In example for the third study period (1999–2010), HR (CI): 1.29 (0.91 to 1.83) for the difference between high and low exposure areas.</p>
Intervention studies and natural experiments				
<p>Peters et al. (1996b)</p> <p>Children</p> <p>N = 3,521</p>	<p>Hong Kong, China (Kwai Tsing and Southern districts)</p> <p>Period of study: 1989–1991</p>	<p>Annual avg (µg/m³):</p> <p>Southern</p> <p>1989: 11</p> <p>1990: 8</p> <p>1991: 7</p> <p>Kwai Tsing</p> <p>1989: 111</p> <p>1990: 67</p> <p>1991: 23</p>	<p>Pre- and post-regulation concentrations compared in natural experiment; SO₂ emissions were reduced by 80% post-regulation.</p>	<p>Associations between respiratory symptoms and living in polluted areas observed and greater decline in symptoms post-regulation.</p> <p>Covariate adjustment: age, gender, environmental tobacco smoking in the family home, housing and father's education.</p>
<p>†Wong et al. (1998)</p> <p>Children (9–12 yr)</p> <p>N = 423</p>	<p>Hong Kong, China (Kwai Tsing and Southern districts)</p> <p>Period of study: 1989–1991</p>	<p>Annual avg (µg/m³):</p> <p>Southern</p> <p>1989: 11</p> <p>1990: 8</p> <p>1991: 7</p> <p>Kwai Tsing</p> <p>1989: 111</p> <p>1990: 67</p> <p>1991: 23</p>	<p>Pre- and post-regulation concentrations compared in natural experiment; SO₂ emissions were reduced by 80% post-regulation.</p>	<p>Decreased bronchial responsiveness observed post-intervention.</p>

Table 5-22 (Continued) Selected epidemiologic studies of long term exposure to SO₂ and the development of asthma and intervention studies/natural experiments.

Study/Population	Location Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates (95% CI) ^a
†Iwasawa et al. (2009) Miyake adults (N = 823)	Miyakejima Island, Japan, near Mt. Oyama volcano 2004–2006	31, post volcano (range: 19–45) Inhabited areas were classified into one lower SO ₂ and three higher SO ₂ areas to gauge exposure.	Seven monitors in residential areas used to estimate 2 yr avg; natural experiment comparing symptom prevalence pre- and post-volcano eruption. Copollutant correlations: NR	Minor health effects on the respiratory system observed. Phlegm higher in higher exposure areas. Note: no consistent differences in lung function observed. Logistic regression model used. Covariate adjustment: sex, age, current smoking status, residential area, and hyper-susceptibility.
†Iwasawa et al. (2015) 120 Miyake school children	Feb. 2005 to Nov. 2011	Average concentrations (ppb) of SO ₂ decreased year-by-year and ranged from 11.3 to 2.47 in low area, from 32.2 to 12.2 in high area-1, and from 75.1 to 12.1 in high area-2.	Six monitors in residential areas used to estimate post-volcano eruption concentrations in different residential areas. Other volcanic gases were measured and considered to be unlikely to cause the health effects seen in the study.	Prevalence of respiratory symptoms (cough, phlegm, wheeze, shortness of breath) was increased in areas with higher post-volcano SO ₂ concentrations compared to areas with lower concentrations. Exposure-dependent increases in symptoms observed (no effects observed at concentrations lower than 30 ppb). Logistic regression model used. Covariate adjustment: age, sex, and hyper-susceptibility.
†Longo et al. (2008) †Longo (2009) Adults (≥20 yr) N = 115 exposed N = 110 unexposed	Kilauea volcano, Hawaii Apr. to Jun. 2004	24.5 (exposed) 0.7 (unexposed). The volcanic plume is carried over the exposed area of the island by the Pacific trade winds. The unexposed area is located at the extreme end of the island from the volcano.	Ambient and indoor SO ₂ concentrations measured using a network of 70 passive samplers over a 3-wk sample period. Copollutant correlations: NR	Cough on most days for 3 consecutive mo or more (acute bronchitis) per yr increased in areas with higher levels. Note: associations with other symptoms also reported. Logistic regression model used. Covariate adjustment: age, sex, race, smoking, dust, and body mass index.

Table 5-22 (Continued) Selected epidemiologic studies of long term exposure to SO₂ and the development of asthma and intervention studies/natural experiments.

Study/Population	Location Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates (95% CI) ^a
† Tam et al. (2016) 1,836 4th/5th graders mean age 10,1 yr	Kilauea volcano, Hawaii 2002 to 2005	SO ₂ , PM _{2.5} , and particulate acid in four exposure zones. Mean (SD) SO ₂ across zones ranged from 0.3 to 10.1 ppb.	SO ₂ measured by passive diffusion for 1- to 4-wk intervals to determine zone levels at representative sites in each zone.	Strongly acidic respirable particulates associated with cough. SO ₂ not evaluated specifically but included in the area mix which was not related to cough. Cross-sectional study with adjustments for age, race, sex, sitting height, BMI, premature birth, maternal smoking during pregnancy, current smokers in the home, and visible mold in the home.

Avg = average; BMI = body mass index; CI = confidence interval; HR = heart rate; ICD = International Classification of Diseases; IDW = inverse distance weighting; n = sample size; N = population number; NR = not reported; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; SD = standard deviation; SES = socioeconomic status; SO₂ = sulfur dioxide.

^aEffect estimates are standardized per 5-ppb increase in SO₂ concentrations unless otherwise noted.

†Studies published since the 2008 ISA for Sulfur Oxides.

In a study of the British Columbia Birth Cohort (n =3,394 asthma cases), [Clark et al. \(2010\)](#) used inverse distance weighted estimate-based concentrations from the three closest monitors within 50 km of the participant’s postal code to estimate SO₂ exposure. These authors observed an adjusted OR per 5 ppb of 1.48 (95% CI: 1.3, 1.9) due to average exposures during both pregnancy and the first year of life. Conducted in Southwest British Columbia, the study had 14 SO₂ monitors available to provide data. [Clark et al. \(2010\)](#) conducted a quartile analysis to explore the exposure-response relationship and observed that the trend across quartiles was not linear [i.e., for the first-year exposure model the second quartile was smaller (negative with confidence intervals less than 1.0) than the positive first and last quartiles], lessening the strength of the association. In this nested case-control study (n = 37,401), medical records of children aged 3–4 years (born 1999–2000) were reviewed for asthma diagnosis ([Clark et al., 2010](#)). Selection bias due to differential loss to follow-up is not an issue because of the records-based analysis used.

Asthma incidence for school children from the Taiwan Health Insurance Database was evaluated contrasting high and low air pollution areas near a petrochemical complex for three time periods after the opening of the complex. The areas were indexed by 3-year annual average levels of the 99th percentile of SO₂ levels and periods above 75 ppb ([Chiang et al., 2016a, b](#)). The hazard ratios (HRs) were positive with wide confidence

intervals for the three periods. Caution is required in making inferences about an SO₂ effect because (1) the air in the areas examined represent complicated mixes from petrochemical complexes, (2) the uncertainty for exposure error may be too high to include area comparisons rather than individual level comparisons, and (3) there is an absence of evaluation for potential asthma risk factors.

The use of questionnaires in these studies to ascertain parents' report of physician-diagnosed asthma, a strength of the study design ([Burr, 1992](#); [Ferris, 1978](#)), adds to the strength of inference about associations with SO₂. A limitation of these longitudinal studies include the potential for exposure measurement error related to the use of inverse distance weighting for SO₂ exposure estimates and comparison of high and low concentration areas (see [Section 3.3.2.3](#)). Validation of SO₂ exposures was not discussed for these studies. The standard increment used in the current ISA, 5 ppb for an annual average, is larger than the mean exposures in these studies, especially so for [Clark et al. \(2010\)](#) where the mean exposure (SD) is 1.98 (0.97) ppb. Additionally, the strongest associations observed in both studies were with NO₂ concentration. Correlations between pollutant concentrations were not reported by [Chiang et al. \(2016a\)](#); [Nishimura et al. \(2013\)](#), while [Clark et al. \(2010\)](#) noted that correlations between pollutant concentrations were generally high but did not provide quantitative data. These studies suggest a potential relationship between long-term SO₂ exposure and the development of asthma. However, these results do little to reduce uncertainty related to potential copollutant confounding.

These studies considered confounding by asthma risk factors, although the relation to SO₂ exposure and potential for confounding, is not well characterized for many of these risk factors. Adjustment was made for SES with individual characteristics such as parental education, income, or health insurance status ([Nishimura et al., 2013](#); [Clark et al., 2010](#)) ([Table 5-22](#)). [Clark et al. \(2010\)](#) additionally adjusted for neighborhood income levels, and [Nishimura et al. \(2013\)](#) adjusted for race/ethnicity. These studies also adjusted for maternal smoking, which has an unclear relationship to SO₂ exposure. Obesity has been identified as a potential risk factor for asthma in children ([Gilliland et al., 2003](#); [Gold et al., 2003](#)). In the cohorts examined by [Nishimura et al. \(2013\)](#), obesity was associated with poorer asthma control ([Borrell et al., 2013](#)). The role of obesity as an effect modifier of SO₂ was not examined in any of the studies of asthma. [Nishimura et al. \(2013\)](#) did not find that atopy or family history of atopy modified the SO₂ association.

Several recent studies presented in Supplemental Table 5S-12 ([U.S. EPA, 2017c](#)) also examined the association of long-term exposure to SO₂ with the prevalence of asthma in cross-sectional designs with various SO₂ exposure estimates as discussed in the table. While these studies involve uncertainties, with the exception of [Portnov et al. \(2012\)](#),

most reported positive associations ([Liu et al., 2016](#); [Deng et al., 2015a](#); [Liu et al., 2014a](#); [Dong et al., 2013c](#); [Dong et al., 2013b](#); [Kara et al., 2013](#); [Deger et al., 2012](#); [Portnov et al., 2012](#); [Akinbami et al., 2010](#); [Sahsuvárovlu et al., 2009](#)). These studies are consistent with similar studies in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). [Deng et al. \(2015a\)](#) examined copollutant models and reported that adjusting for PM₁₀ only slightly changes the SO₂ asthma risk. However, the SO₂ association did not persist after adjusting for NO₂. [Liu et al. \(2016\)](#) only examined multipollutant models, which can be unreliable because of multicollinearity among pollutants, and found that adjusting for both NO₂ and PM₁₀ attenuated the SO₂ association. No longitudinal study of asthma incidence evaluated copollutant models. Thus, within the recent epidemiologic evidence base, the studies provided inconsistent new data and do not reduce the uncertainty related to whether the effect was from SO₂ or another pollutant. Studies of asthma incidence strengthen the inference by addressing the temporality of exposure and response.

Supportive evidence for a relationship between long-term SO₂ exposure and the development of asthma is provided by cross-sectional studies of respiratory symptoms related to asthma. In the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), studies examining an array of respiratory symptoms related to SO₂ exposure are presented in Supplemental Table F-6 ([U.S. EPA, 2008d](#)) and others are noted in the text of the 2008 SO_x ISA ([U.S. EPA, 2008d](#); [Ware et al., 1986](#); [Chapman et al., 1985](#); [Dodge et al., 1985](#)). These cross-sectional studies used monitors for the SO₂ exposure estimate. While associations were generally positive, some inverse or null associations were also observed. Recent studies evaluating the relationship between long-term SO₂ exposure and the prevalence of asthma symptoms [Supplemental Table 5S-12 ([U.S. EPA, 2017c](#))] also found positive associations ([Altuğ et al., 2013](#); [Pan et al., 2010](#); [Arnedo-Pena et al., 2009](#)).

Additional epidemiologic evidence for a link between long-term exposure to SO₂ and the development of asthma may come from intervention or natural experiment studies (see [Table 5-22](#)). Physicians diagnose asthma, in part, based on the occurrence or exacerbation of asthma symptoms, such as cough and wheeze, and the level of AHR in the subjects. Decline in such symptoms and AHR in relation to a decline of a pollutant level may support a relationship between asthma development and exposure to pollutants such as SO₂. Decreases in respiratory symptoms, including any wheeze or symptoms of asthma, cough and sore throat, in 3,521 healthy children (mean age of 9.51 years) were associated with decreases in SO₂ concentrations in Hong Kong due to a government restriction of sulfur content of fuels as discussed in the 2008 SO_x ISA [see [Peters et al. \(1996b\)](#), within [U.S. EPA \(2008d\)](#)]. During the same period, [Wong et al. \(1998\)](#) examined the effect of the same decrease in SO₂ concentrations on AHR in children aged 9–12 who were nonwheezing and did not have asthma at study entry. In the cohort analysis, which compared measurements made before the intervention and 1 year

afterwards, AHR declined. The subjective health measures seen in [Peters et al. \(1996b\)](#) were corroborated by the objective data of the histamine challenge test in [Wong et al. \(1998\)](#). These results should be interpreted with caution given the uncertainty of whether changes in AHR and respiratory symptoms were independently related to SO₂ in light of the concomitant decline in sulfate respirable suspended particles (RSP) (<10 μm). Over the study period, SO₂ declined about 80% (from about 111 to 23 μg/m³) while annual mean sulfate concentrations in RSP fell from 12.5 to 7.7 μg/m³. It is difficult to determine whether one was more important than the other. However, these studies add to the information base relating long-term SO₂ exposure and asthma-related outcomes.

Recent cross-sectional studies that estimated long-term SO₂ exposure from volcano emissions in Japan and Hawaii were conducted ([Table 5-22](#)). [Iwasawa et al. \(2009\)](#) observed increased frequencies of phlegm and minor effects on the respiratory system among both adults and children residing near the Mt. Ōyama volcano in Japan across four inhabitant areas with varying SO₂ levels. [Iwasawa et al. \(2015\)](#) further followed the children yearly from 2006 to 2011, finding the prevalence of respiratory symptoms (cough, phlegm, wheeze, shortness of breath) to be related to the higher SO₂ exposure. Studies conducted near the Kīlauea volcano in Hawaii observed an adjusted increase in cough on most days for 3 consecutive months or more per year in children and adults ([Longo, 2009](#); [Longo and Yang, 2008](#); [Longo et al., 2008](#)). [Tam et al. \(2016\)](#) related cough to a mixture containing acidic respirable particulates, but not to SO₂ exposure directly, in children near the Kīlauea volcano. As a whole, these studies are supportive of a link between SO₂ exposure and respiratory symptoms. However, such studies compare areas of high volcano emissions to areas of lower emissions (indexed by SO₂ concentration), and thus, results may be confounded by copollutant exposures.

Severity of Asthma

[NHLBI NAEPP \(2007\)](#) identifies stages of asthma such as mild, moderate, moderate-persistent, and severe. When going from mild to severe, the likelihood of acute exacerbations increases. Stages of worsening of asthma are usually based on severity scores as used in the following studies [[Supplemental Table 5S-12, \(U.S. EPA, 2017c\)](#)]. [Rage et al. \(2009\)](#) examined severity of asthma in adults. Long-term SO₂ exposure was correlated with a higher asthma severity score. Ozone showed the strongest relationship while NO₂ was unrelated. In 17-year-old male military recruits, [Greenberg et al. \(2016\)](#) related asthma severity to SO₂ measured as low, intermediate, and high. The observed associations between asthma severity and air pollution support the notion that air pollutants may increase asthma severity. However, the uncertainty as to whether these effects are influenced by short-term exposure needs to be examined. [Deger et al. \(2012\)](#) observed an association with long-term SO₂ exposure among children with active asthma

and a more marked association among children with poor asthma control. No other pollutants were examined. Adjusting for child's age and sex, parental atopy, and environmental tobacco smoke exposure slightly decreased the association, and stratification according to age (<6 years and ≥ 6 years) showed that associations with SO₂ were mainly observed in the older age group. Adjusting for socioeconomic status (i.e., household income and maternal educational level) had limited influence on the results of the analyses (<5%).

AHR is a key component of asthma. In a recent study, long-term exposures to SO₂, estimated from monitors up to 50 km from subjects' ZIP code centroid, were associated with increased methacholine responsiveness (i.e., decreased PC₂₀) determined by FEV₁ decreasing by 20% or more ([Ierodiakonou et al., 2015](#)). A large number of comparisons were made between pollutants, exposure lags, lung function parameters, cities, and asthma medication groups possibly increasing probability that the few associations observed may be due to chance. The PC₂₀ percent change per interquartile range (2 ppb 4-month moving average) was -6% (95% CI, -11 to -1.5%) in 2,661 observations in the Childhood Asthma Management Program (CAMP), a randomized clinical trial involving eight cities in North America. The PC₂₀ standardized to per 5 ppb is -15% (-27.5 to -3.75%). The 4-month average SO₂ was not associated with changes in lung function measured before or after bronchodilator treatment. Health outcome results for 1-day and 1-week exposure periods are discussed earlier in [Section 5.2.1.2](#); only the 4-month moving average results are discussed here. The original health study, a longitudinal prospective cohort study with repeated measures but without a pollution component, was designed to examine the long-term safety and effectiveness of daily inhaled anti-inflammatory medication in children diagnosed with mild to moderate asthma and was sponsored by the NHLBI. The children were 5 to 12 years of age and hyperresponsive to methacholine at study entry. Recruitment occurred from late December 1993 to early September 1995 ([CAMP Research Group, 1999](#); [Cherniack et al., 1999](#)) at two health maintenance organizations and six academic institutions.

Monitoring data on 24-h avg concentrations of pollutants ozone, CO, NO₂, and SO₂ were obtained for each metropolitan area from the Aerometric Information Retrieval System for the U.S. cities and from the Air Quality and Reporting Unit for Toronto were linked to the ZIP code of the subject's address at study entry. There is uncertainty in the measurement estimate and a potential for exposure measurement error. Distance or proximity of sites to subjects' residence is not known. For long-term studies, bias can go in either direction. Thus, the evidence base for a relationship between long-term SO₂ exposure and AHR is limited.

Animal Toxicological Studies

A single animal study of chronic SO₂ exposure-related effects on lung morphology was discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Study characteristics are summarized in [Table 5-23](#). [Smith et al. \(1989\)](#) found that rats exposed to 1 ppm of SO₂ had an increased incidence of bronchiolar epithelial hyperplasia and increased numbers of nonciliated epithelial cells after 4 months of exposure. However, these effects were not present at 8 months of exposure, suggesting that repair and/or adaptation may have taken place.

Table 5-23 Study-specific details from animal toxicological studies.

Study	Species (strain); n; Sex; Lifestage/Age	Exposure Details (Concentration; Duration)	Endpoints Examined
Smith et al. (1989)	Rats (Sprague-Dawley); n = 12–15 per data point; M; young adult; normal or elastase-impaired	1 ppm (2.62 mg/m ³) SO ₂ whole body; 5 h/day, 5 days/wk for 4 or 8 mo 8-mo exposure group sacrificed immediately or 3 mo after exposure ended	Endpoints examined prior to sacrifice Lung function—residual volume, functional residual capacity, quasi-static compliance, residual volume/total lung capacity, N ₂ washout Morphological effects
Song et al. (2012)	Rats (Sprague-Dawley); n = 10/group; M; 4 wk old neonates	Sensitization by i.p. injection of 10 mg ovalbumin followed by booster injection of 10 mg ovalbumin after 7 days Challenge with 1% ovalbumin aerosol for 30 min daily for 4 wk beginning at 15 days Exposure to 2 ppm SO ₂ for 4 h/day for 4 wk beginning at 15 days Exposure groups: (1) Control (2) SO ₂ alone (3) Ovalbumin alone (4) Ovalbumin + SO ₂	Endpoints examined 24 h after challenge Lung function—whole body plethysmography (MCh challenge) BALF-IL-4, IFN-γ Serum-IL-4, IFN-γ Lung—histopathology In vitro culture of airway smooth muscle cells from experimentally treated animals—stiffness and contractility

BALF = bronchoalveolar lavage fluid; IFN-γ = interferon gamma; IL-4 = interleukin-4; i.p. = intraperitoneal; M = male; MCh = methacholine; n = sample size; N₂ = nitrogen; SO₂ = sulfur dioxide.

No studies on airway responsiveness or pulmonary inflammatory responses to long-term exposure to SO₂ concentrations of 2 ppm and lower were discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). One new animal toxicological study of subchronic SO₂ exposure has become available since the last review. Key findings are discussed here, and study characteristics are summarized in [Table 5-23](#). [Song et al. \(2012\)](#) found that Penh, a measure of bronchial obstruction, was enhanced following methacholine challenge in a model of allergic airways disease. In this model, rats were first sensitized and challenged with ovalbumin and then exposed to 2 ppm SO₂ for 4 hours/day for 28 days. Penh was not changed with exposure to SO₂ alone in naive rats. However, [Song et al. \(2012\)](#) observed hyperemia in the lung parenchyma and inflammation in the airways of naive rats exposed only to SO₂. SO₂ exposure also increased the inflammatory responses in rats made allergic to ovalbumin. Airway remodeling was found in ovalbumin-treated rats with and without exposure to SO₂. A more pronounced increase in the airway smooth muscle layer was found in the ovalbumin/SO₂ group compared to the ovalbumin group. The authors concluded that the effects of SO₂ on bronchial obstruction and airway remodeling

were dependent on ovalbumin sensitization and challenge. [Song et al. \(2012\)](#) also measured concentrations of IL-4 and IFN- γ in the BALF and serum of rats exposed to SO₂, with and without prior sensitization and challenge with ovalbumin. Concentrations of IL-4 in the BALF were increased in the ovalbumin and the SO₂ groups, with the greatest increase occurring in the combined ovalbumin/SO₂ group. An increase in IL-4 in serum occurred only in the ovalbumin/SO₂ group. Concentrations of IFN- γ in the BALF were decreased in the ovalbumin, SO₂, and ovalbumin/SO₂ groups. A decrease in serum IFN- γ was observed in the ovalbumin and ovalbumin/SO₂ groups.

IL-4 is a Th2 cytokine associated with allergic responses, while IFN- γ is a Th1 cytokine. An increase in the ratio of Th2 to Th1 cytokines indicates Th2 polarization (or possibly a Type 2 immune response mediated by Group 2 innate lymphoid cells), a key step in allergic sensitization. These findings provide evidence that repeated SO₂ exposure enhances allergic responses and airway remodeling. Furthermore, increases in bronchial obstruction suggest that SO₂ exposure increased airway responsiveness in this model of allergic airway disease. Furthermore, repeated SO₂ exposure in naive rats increased levels of the Th2 cytokine IL-4, decreased levels of the Th1 cytokine IFN- γ in the BALF, and increased airway inflammation suggesting that SO₂ exposure may induce allergic sensitization on its own. Because allergic sensitization, airway remodeling, and increased airway responsiveness are key events (or endpoints) in the proposed mode of action for the development of asthma ([Section 4.3.6](#)), these results suggest that long-term exposure to SO₂ may lead to the development of an asthma-like phenotype in this animal model involving newborn rats.

Summary of Asthma Development and Severity

Recent epidemiologic evidence from a limited number of longitudinal studies report associations between asthma incidence among children and long-term SO₂ exposures. Additional supportive evidence for a link between long-term SO₂ exposure and the development of asthma is provided by cross-sectional studies of asthma prevalence. The longitudinal studies help reduce the uncertainty associated with the temporality of exposure and response that is inherent in cross-sectional study designs. This evidence is coherent with animal toxicological evidence of inflammation, allergic sensitization and other allergic responses, airway remodeling, and increased airway responsiveness, which are key events (or endpoints) in the proposed mode of action for the development of asthma ([Section 4.3.6](#)). The animal toxicological evidence provides support for an independent effect of SO₂ and strengthens the link between long-term exposure to SO₂ and the development of asthma in children. Additional evidence supportive of this link comes from cross-sectional studies of respiratory symptoms and respiratory allergies among children and from natural experiments. **Thus, multiple lines of evidence suggest**

that long-term SO₂ exposure results in a coherent and biologically plausible sequence of events that culminates in the development of asthma, especially allergic asthma, in children.

The potential for a relationship between long-term SO₂ exposure and severity of asthma has been examined in a few studies. One study in adults correlated exposure with higher asthma severity scores. A study in children found a more marked association in those with poor asthma control. AHR, measured as PC₂₀, worsened with long-term SO₂ exposure in a multicity cohort of children. Thus, evidence of asthma control and increased AHR provides suggestive but limited support for this relationship.

5.2.2.2 Development of Allergy

There is some evidence for a potential relationship between long-term SO₂ exposure and indicators of respiratory allergies and inflammation among children. Several recent cross-sectional studies examined the prevalence of respiratory allergies using different markers for respiratory allergies including IgE antibodies, rhinitis, eczema, sensitization to pollen, and hay fever related to long-term SO₂ exposure ([Liu et al., 2016](#); [Chan et al., 2013](#); [Bhattacharyya and Shapiro, 2010](#); [Penard-Morand et al., 2010](#); [Parker et al., 2009](#); [Nordling et al., 2008](#)) [see Supplemental Table 5S-11 ([U.S. EPA, 2017c](#))]. Positive results were observed for children using these various indicators of allergy. Further, a weak relationship was found by [Dales et al. \(2008\)](#) between long-term SO₂ exposure and eNO, an indicator of inflammation [see Supplemental Table 5S-13 ([U.S. EPA, 2017c](#))].

Recent studies examine two-pollutant models for allergic rhinitis prevalence. Results for allergic rhinitis prevalence based on responses from the ISAAC (International Study of Asthma and Allergies in Childhood) questionnaire data in Changsha, China ([Chan et al., 2013](#)) did not find an association for SO₂ for site-specific background SO₂ and allergic rhinitis in children 3–6 years old, but did find an association for age-related accumulative exposure in a single pollutant model using the closest monitor to kindergartens. The two-pollutant model with PM₁₀ was attenuated. For SO₂ exposures during the first year of life in Shanghai, China, [Liu et al. \(2016\)](#) found an association with allergic rhinitis in children at age 6 years, which was attenuated when adjusted for other pollutants using district monitors. These findings suggest the possibility that chronic exposure to SO₂ may play a role in the development of allergic conditions based on results for various allergic markers. The cross-sectional design of these studies makes these relationships uncertain with regard to the temporal relationship between exposure and outcome. Further, the exposure estimates from monitors may not adequately characterize the spatial and temporal variation in SO₂ concentrations potentially leading to exposure measurement

error ([Section 3.4.2](#)). Although two pollutant models have begun to address the role of SO₂ exposure in the development of allergic rhinitis, the evidence base for a relationship between long-term SO₂ exposure and this response is limited.

5.2.2.3 Lung Function

Epidemiologic Studies

Longitudinal epidemiologic studies examine associations between long-term SO₂ exposure and decrements in lung function. Lung function increases through early adulthood with growth and development, then declines with aging ([Stanojevic et al., 2008](#); [Zeman and Bennett, 2006](#); [Thurlbeck, 1982](#)). Thus, a relationship between long-term SO₂ exposure and decreased lung function over time in school-age children into early adulthood would be an indicator of decreased lung development.

As discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), earlier cross-sectional studies ([Dockery et al., 1989](#); [Schwartz, 1989](#)) found no association between long-term SO₂ exposure and lung function in U.S. children. A longitudinal cohort study ([Frischer et al., 1999](#)) reported that long-term SO₂ exposure was associated with decrements in lung function in the summer but not in the winter. In Poland, a prospective cohort study of children ([Jedrychowski et al., 1999](#)) found decrements in lung function growth related to a polluted area where concentrations of both TSP and SO₂ were high compared to a cleaner area where concentrations of both TSP and SO₂ were low, thus not providing results specifically for SO₂. In a cross-sectional study in adults in Switzerland, [Ackermann-Lieblich et al. \(1997\)](#) observed an association between SO₂ concentration and lung function, but after controlling for PM₁₀, this association was no longer evident. In former East Germany from 1992 to 1999, [Frye et al. \(2003\)](#) reported improvements in lung function associated with declines in SO₂ concentrations in 2,493 children over three cross-sectional surveys. These studies are presented in Supplemental Table 5S-11 ([U.S. EPA, 2017c](#)).

Recent studies in children and adults add to this evidence base [see Supplemental Table 5S-14 ([U.S. EPA, 2017c](#))]. In the repeated measure prospective Taiwan Children's Health Study, [Hwang et al. \(2015a\)](#) examined lung function over a 2-year period from age 12 to 14 years. No association was found for SO₂ exposure and FEV₁ or FVC for boys and girls, but a deficit was observed for boys for FEF₂₅₋₇₅. A single measure longitudinal study in several U.S. cities observed an association between FEV₁ and SO₂ during the first year of life [-1.01 (95%CI: -3.25, 1.27) per 1 ppb increase]. In this study, [Neophytou et al. \(2016\)](#) examined the same cohort that [Nishimura et al. \(2013\)](#) did, as

discussed earlier in this section for asthma incidence in the same cities with the same SO₂ exposure method evaluating the same confounding factors plus obesity.

In a longitudinal repeated-measures study of children, [Linares et al. \(2010\)](#) reported a decline in FEV₁ related to long-term SO₂ exposure in the entire study group. This study included children from two schools in different locations relative to a petrochemical zone. In an analysis of the children by sex, in single- and copollutant models with PM₁₀ or O₃, the association was attenuated. In a cross-sectional study of children in 14 communities in Taiwan, [Lee et al. \(2011c\)](#) found a reduction in FEV₁ related to long-term SO₂ exposure with larger reductions related to NO₂ and CO exposure. [Yogev-Baggio et al. \(2010\)](#) related the effect of the interaction, NO_x × SO₂ “event,” to reduction in FEV₁ in children in Israel near a coal-fired power plant. In a cross-sectional study of 32,712 adults in England, [Forbes et al. \(2009c\)](#) reported associations of FEV₁ with exposure to SO₂, PM₁₀, and NO₂, but not O₃. [Dales et al. \(2008\)](#) found a weak decline in FEV₁ and FVC related to long-term SO₂ exposure in school children in Windsor, ON using a cross-sectional prevalence design.

The majority of the recent studies and earlier studies used cross-sectional designs. Some studies took into account potentially confounding covariates detailed in the Supplemental Table 5S-14 ([U.S. EPA, 2017c](#)). [Neophytou et al. \(2016\)](#) controlled for age, height, and calendar time, allowing for nonlinear effects, indicator variables for sex, race/ethnicity, and continuous variables for SES (composite score variable), and numbers of smokers in the household and also assessed effect modification by sex, obesity, SES, atopy, and parental asthma. The designs used in most of the recent studies (i.e., ecological, cross-sectional, single measure) limit the possible inferences about the relationship between long-term SO₂ exposure and lung function. The evidence does not include studies evaluating concentration-responses. The one study conducting a copollutant analysis found attenuation of the effect with adjustment for PM₁₀. **Thus, recent studies do not add information that changes conclusions made in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) that there is not clear evidence that long-term SO₂ exposure is related to lung function changes.**

Animal Toxicological Studies

A single long-term study with SO₂ exposure concentrations at or below 2 ppm was discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Study characteristics are summarized in [Table 5-23](#). [Smith et al. \(1989\)](#) found that rats exposed to 1 ppm SO₂ for 4 months had decreased residual volume and quasi-static compliance when treated with saline (control). Rats treated with elastase (a model of emphysema) and exposed to 1 ppm SO₂ for 4 months had a decreased ratio of residual volume to total lung capacity and decreased

alveolar plateau of the single-breath nitrogen (N₂) washout (N₂-slope), indicating a worsening of the emphysema. However, [Smith et al. \(1989\)](#) concluded that the effects of SO₂ on lung function measurements were very minor in the saline (control) group and likely due to chance alone (residual volume) or to unusually high control values (quasi-static compliance).

Summary of Lung Function

Several studies evaluated the relationship between long-term SO₂ exposure and decrements in lung function. Evidence supporting this relationship is limited because associations were inconsistent and because both PM and SO₂ were at high concentrations in the same areas, which does not allow determination of individual SO₂ effects. Potential confounding of long-term SO₂ exposure-related decrements in lung function and lung development by other pollutants, especially PM, was evaluated in only one study. This study found an attenuation of the effect in copollutant analyses. No changes in lung function were found in long-term animal toxicological studies at relevant SO₂ concentrations. The recent studies support conclusions made in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) that evidence does not strongly support an effect of long-term SO₂ exposure on decreases in lung function in children.

5.2.2.4 Respiratory Infection

Epidemiologic Studies

Studies have also examined the association of long-term exposure to SO₂ with infant bronchiolitis, otitis media, and pneumonia in children, hospital admission for community-acquired pneumonia in adults aged 65 years or more, and tuberculosis in adults. Infant bronchiolitis was examined in British Columbia by [Karr et al. \(2009\)](#). These study authors observed an association with lifetime exposure to SO₂ after adjustment for an array of confounders [Supplemental Table 5S-13 ([U.S. EPA, 2017c](#))]. The largest associations were observed with NO₂ and CO concentrations. [MacIntyre et al. \(2011\)](#) found no increased risk for otitis media in relation to long-term SO₂ exposure in a study of children up to the age of 2 years in British Columbia, while [Bhattacharyya and Shapiro \(2010\)](#) found a strong relationship with long-term SO₂ exposure in the U.S. National Health Interview Survey of 126,060 children ages 3–6 years. [Lu et al. \(2014\)](#) observed that the prevalence of pneumonia in children 3 to 6 years old was related to long-term SO₂ exposure. [Liu et al. \(2016\)](#) reported that doctor-diagnosed pneumonia in children 4–6 years old was related to SO₂ exposure during the first year of life. [Neupane](#)

[et al. \(2010\)](#) estimated long-term SO₂ exposure at the residence (2-yr avg for 2001–2002) for both the case and control subjects using both bicubic spline and inverse distance weighting methods, obtaining means of 4.65 and 5.80 ppb, respectively, but with a twofold greater range for the bicubic spline method. Adjusted estimates of associations for SO₂ with hospitalization from community-acquired pneumonia were positive for the bicubic spline method but not for inverse distance weighting. The incidence of tuberculosis was associated with an increase of SO₂ in adult males ([Hwang et al., 2014](#)) but not in a study in California ([Smith et al., 2016](#)). Although limited in number, by inconsistency, and by their cross-sectional design, these studies suggest a potential relationship between long-term exposure to SO₂ and respiratory infections due to various infectious agents.

Animal Toxicological Studies

No new animal studies of the effects of long-term SO₂ exposure on lung host defense have been conducted since the previous review. Several studies of short- and long-term exposure to SO₂ were reported in the 1982 AQCD ([U.S. EPA, 1982a](#)) and discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Short-term exposure studies found some effects of 0.1–1 ppm SO₂ on the clearance of labeled particles. Long-term exposure studies found decreased tracheal mucus flow at a concentration of 1 ppm SO₂, but no effects on susceptibility to bacterial infection or alterations in the pulmonary immune system at concentrations of 2 ppm or less.

Summary of Respiratory Infection

Evidence for prevalence of infant bronchiolitis and/or respiratory infections consists of generally positive associations found in cross-sectional studies. Thus, they provide a limited evidence base in number and design. While some animal toxicological studies reported alterations in specific host defense mechanisms, there is no evidence to support increases in bacterial or viral infections in animals exposed to SO₂ at relevant concentrations.

5.2.2.5 Development of Other Respiratory Diseases: Chronic Bronchitis, Chronic Obstructive Pulmonary Disease, and Acute Respiratory Distress Syndrome

Chronic bronchitis consists of symptoms, including daily cough and/or congestion or phlegm for 3 months in a row. While these symptoms may have started with acute exacerbation, they are likely to represent chronic indolent symptoms. As discussed in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), earlier cross-sectional studies observed positive

relationships between long-term SO₂ exposure estimates derived from fixed-site monitors and chronic bronchitis as presented in Supplemental Table 5S-13 ([U.S. EPA, 2017c](#)). Recent cross-sectional studies of the association of long-term exposure to SO₂ with the prevalence of bronchitis also observed positive relationships after adjustment for potential confounders. In addition, a recent COPD incidence study in a national English cohort ([Atkinson et al., 2015](#)), discussed in Supplemental Table 5S-13 ([U.S. EPA, 2017c](#)), reported a positive association in an adjusted HR model with SO₂ exposure averaged over 3 years determined by dispersion models. Assessment of model validity using national network sites and separate verification sites yielded poor *R*² values for SO₂ of 0 and 0.39, respectively. Other limitations of this study include a short follow-up time and the failure to confirm the 36% of incident hospital admissions for COPD by a general practitioner diagnosis.

A relationship between Acute Respiratory Distress Syndrome (ARDS) and long-term SO₂ exposure has recently been studied ([Ware et al., 1986](#)) as discussed in Supplementary Table 5S-13 ([U.S. EPA, 2017c](#)). SO₂ and PM_{2.5} were not associated with ARDS.

5.2.2.6 Respiratory Mortality

Recent studies provide some evidence that respiratory mortality may be more consistently associated with long-term exposure to SO₂ than other cause-specific mortality outcomes ([Section 5.5.2](#) and [Figure 5-27](#)). There is uncertainty in the small, positive associations between long-term exposure to SO₂ and respiratory mortality observed in these studies because the exposure assessment and statistical methods are not adequate for studying a highly spatially and temporally heterogeneous pollutant like SO₂. Additionally, there is little evidence of respiratory health effects in adults in relation to long-term SO₂ exposure that could provide coherence with the observed associations with respiratory mortalities.

5.2.2.7 Summary and Causal Determination

Overall, the evidence is suggestive of, but not sufficient to infer, a causal relationship between long-term SO₂ exposure and respiratory effects, based on the evidence across disciplines for development of asthma in children. This conclusion represents a change from “inadequate to infer a causal association” for respiratory effects as stated in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Evidence for development of other respiratory diseases, lung function decrements, and respiratory mortality is relatively weak and does not substantially influence the causal determination.

The change in causal determination is based on evidence from recent epidemiologic studies and support from animal toxicological studies. A limited number of longitudinal epidemiologic studies report associations between asthma incidence among children and long-term SO₂ exposures. The longitudinal studies help to reduce the uncertainty associated with temporality that is inherent in the previous cross-sectional studies. The evidence from longitudinal studies is coherent with animal toxicological evidence of allergic sensitization, airway remodeling, and increased airway responsiveness, which are key events (or endpoints) in the proposed mode of action for the development of asthma. The animal toxicological evidence provides support for an independent effect of SO₂ and a possible relationship between long-term exposure to SO₂ and the development of asthma in children. Some evidence of a link between long-term exposure to SO₂ and respiratory symptoms and/or respiratory allergies among children further supports this relationship. The potential for SO₂ to serve as an indicator for other pollutants or mixture related to PM is an uncertainty that applies to the new body of epidemiologic evidence across the respiratory effects examined.

The key evidence supporting the causal determination is detailed below using the framework described in Table I of the Preamble to the ISAs ([U.S. EPA, 2015b](#)) and is presented in [Table 5-24](#).

Table 5-24 Summary of evidence for suggestive of, but not sufficient to infer, a causal relationship between long-term sulfur dioxide exposure and respiratory effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Development and severity of asthma			
Evidence from epidemiologic studies is generally supportive but not entirely consistent	Evidence for increases in asthma incidence in cohorts of children in U.S. and Canada. Adequate adjustment for confounding by asthma risk factors. Some inconsistency regarding time window	Nishimura et al. (2013) Clark et al. (2010)	Mean (SD) across five cities 4.0 (3.4) ppb 1.98 (0.97) ppb
	Supporting cross-sectional studies of asthma prevalence among children but uncertainty regarding the temporal sequence between exposure and the development of asthma	Section 5.2.2.1	
	Supporting evidence for respiratory symptoms and markers of respiratory allergies among children in cross-sectional studies	Sections 5.2.2.1 and 5.2.2.2	
	Supporting evidence from intervention studies and natural experiments	Section 5.2.2.1	
	Evidence for increases in asthma severity as indicated by asthma severity score, degree of asthma control, and AHR	Section 5.2.2.1	
Uncertainty regarding potential for measurement error in exposure estimates	SO ₂ concentrations assigned to subjects based on combining monitoring data by IDW in asthma incidence studies and monitors in cross-sectional studies may not adequately represent exposure	Nishimura et al. (2013) Clark et al. (2010) Section 3.5	
Uncertainty regarding potential confounding by copollutants	No copollutant models or correlations analyzed in asthma incidence studies. In cross-sectional study, SO ₂ association persists with PM ₁₀ adjustment, not NO ₂ . r with SO ₂ across lifestages: 0.60–0.83.	Table 5-22 and Deng et al. (2015a) Section 3.4.3	

Table 5-24 (Continued): Summary of evidence for suggestive of, but not sufficient to infer, a causal relationship between long term sulfur dioxide exposure and respiratory effects

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Limited animal toxicological evidence provides coherence and biological plausibility	Th2 polarization (or other Type 2 immune responses) and airway inflammation following repeated exposure of naive newborn rats for 28 days Evidence for enhanced inflammation, airway remodeling, and increased airway responsiveness following repeated exposure of allergic newborn rats for 28 days	Song et al. (2012)	2,000 ppb
Coherence with evidence from short-term animal toxicological studies	Inflammation and morphologic responses indicative of airway remodeling following repeated exposures of naive rats over several days	Li et al. (2007) Li et al. (2014)	2,000 ppb
	Enhancement of allergic sensitization, allergic inflammation, and bronchial obstruction in guinea pigs exposed repeatedly over several days and subsequently sensitized and challenged with an allergen	Riedel et al. (1988) Park et al. (2001)	100 ppb 100 ppb
	Enhanced inflammation and allergic responses in rats previously sensitized with an allergen and then repeatedly exposed	Li et al. (2007) Li et al. (2014)	2,000 ppb
Some evidence for key events in proposed mode of action	Inflammation, allergic sensitization, AHR, airway remodeling	Section 4.3.6	
Development of allergy			
Limited epidemiologic evidence but uncertainty regarding SO ₂ independent effects	Generally positive associations with different markers for allergies in cross-sectional studies in children. Uncertainty in temporality and exposures estimated from fixed-site monitors; copollutant confounding examined on a limited basis remains uncertain	Section 5.2.2.2	

Table 5-24 (Continued): Summary of evidence for suggestive of, but not sufficient to infer, a causal relationship between long term sulfur dioxide exposure and respiratory effects

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Lung function			
Inconsistent epidemiologic evidence among children from quality studies and uncertainty regarding SO ₂ independent effects	In cohort studies, associations inconsistent with adjustment for PM and by season	Neophytou et al. (2016) Jedrychowski et al. (1999) Frischer et al. (1999)	
	Inconsistent results from cross-sectional studies	Dockery et al. (1989) Schwartz (1989) Ackermann-Lieblich et al. (1997) Frye et al. (2003)	
Respiratory infection			
Limited epidemiologic evidence; uncertainty regarding SO ₂ independent effects	Generally positive associations in cross-sectional studies. Uncertainty in temporality, exposures estimated from monitors in the community, and copollutant confounding	Section 5.2.2.4	
Limited animal toxicological evidence	Altered clearance of particles and decreased tracheal mucus flow	U.S. EPA (1982a)	0.1–1 ppm
Lack of evidence for key events in proposed mode of action	Changes in specific host defense mechanisms but no evidence of greater infectivity		
Development of other respiratory diseases			
Limited epidemiologic evidence but uncertainty regarding SO ₂ independent effects	Generally positive associations for chronic bronchitis in cross-sectional studies. Uncertainty in temporality, exposures estimated from monitors in the community, and copollutant confounding	Section 5.2.2.5	

Table 5-24 (Continued): Summary of evidence for suggestive of, but not sufficient to infer, a causal relationship between long term sulfur dioxide exposure and respiratory effects

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Respiratory mortality			
Generally consistent epidemiologic evidence	Small, positive associations between long-term exposure to SO ₂ and respiratory mortality in several cohorts, even after adjustment for common potential confounders	Hart et al. (2011) , Nafstad et al. (2004) , Elliott et al. (2007) , Cao et al. (2011) , Carey et al. (2013) , Dong et al. (2012) , Katanoda et al. (2011)	2.4–41.4
No coherence between respiratory morbidity and respiratory mortality	No evidence for a relationship between long-term exposure and respiratory mortality to support the observed associations with respiratory morbidity	Section 5.2.2.6	

AHR = airway hyperresponsiveness; IDW = inverse distance weighting; NO₂ = nitrogen dioxide; PM = particulate matter; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; *r* = correlation coefficient; SD = standard deviation; SO₂ = sulfur dioxide Th2 = T-derived lymphocyte helper 2.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)).

^bDescribes the key evidence and references, supporting or contradicting, contributing most heavily to causal determination, and where applicable, to uncertainties or inconsistencies. References to earlier sections indicate where the full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated (for experimental studies, ≤2,000 ppb).

Evidence for the Development of Asthma

A limited number of longitudinal studies demonstrate associations between ambient SO₂ concentrations measured in the first year of life and/or over the first 3 years of life in children and asthma incidence such as [Clark et al. \(2010\)](#) and [Nishimura et al. \(2013\)](#) ([Section 5.5.2.1](#)). Results are fairly consistent between studies with one based on several different locations across the U.S., another over a large area in Canada, and one in Taiwan, involving a large number of participants. Uncertainties and the potential for exposure measurement error related to the use of inverse distance weighting and area comparisons in these studies may limit inferences that can be made ([Section 3.3.2.3](#)). Additional supportive evidence for a link between long-term SO₂ exposure and the development of asthma is provided by cross-sectional studies of asthma prevalence, respiratory symptoms, and markers of respiratory allergies among children ([Section 5.2.2.2](#)). Findings of studies evaluating respiratory symptoms are supportive of the development of asthma; however, they may also reflect other respiratory conditions.

Intervention and natural experiment studies also indicate a possible relationship between long-term exposure to SO₂ and the development of asthma.

Epidemiologic studies of asthma development in children have not clearly characterized potential confounding by other pollutants or mixtures of pollutants. This uncertainty was present in the previous review, and there is no new information from incidence studies to help reduce this uncertainty. No studies of asthma incidence have evaluated copollutant models to address copollutant confounding, making it difficult to evaluate the independent effect of SO₂ within the epidemiologic evidence base for incidence. A recent cross-sectional study of asthma prevalence provides inconsistent results from copollutant models, persistence of the SO₂ association with PM₁₀ not NO₂. In studies that examined both SO₂ and PM_{2.5}, positive associations were observed between PM_{2.5} concentrations and asthma development; the effects were similar in magnitude to those for SO₂ ([Nishimura et al., 2013](#); [Clark et al., 2010](#)). Correlations between SO₂ and PM_{2.5} were not reported in these studies. Thus, results from these two studies do not reduce the uncertainty related to potential copollutant confounding.

The uncertainties in the epidemiologic evidence base is reduced, in part, by the biological plausibility provided by findings from experimental studies that demonstrate SO₂-induced effects on key events or endpoints that are part of the proposed mode of action for the development of asthma [i.e., allergic sensitization, airway remodeling and increased airway responsiveness ([Section 4.3.6](#))]. An experimental study in newborn rats, which were not previously sensitized and challenged with an allergen (i.e., naive animals), found that repeated acute SO₂ exposures over several weeks led to airway inflammation and Th2 polarization (or other Type 2 immune responses), important steps in allergic sensitization [([Song et al., 2012](#)); see [Section 5.2.2.1](#)]. Repeated SO₂ exposure in the newborn rats, which were previously sensitized and challenged with an allergen (i.e., allergic animals), resulted in enhanced allergic airway inflammation and some evidence of airway remodeling and increased airway responsiveness. Additional evidence comes from experimental studies in adult animals involving short-term exposure to SO₂ over several days. In naive rats, airway inflammation and morphologic responses indicative of airway remodeling were seen ([Section 5.2.1.7](#)). Furthermore, enhancement of allergic sensitization and other inflammatory responses were observed along with evidence for increased airway responsiveness in guinea pigs exposed repeatedly to SO₂ for several days and subsequently sensitized and challenged with an allergen ([Section 5.2.1.7](#)). Similarly, SO₂ exposure enhanced airway inflammation in rats previously sensitized with an allergen ([Section 5.2.1.2](#)).

Evidence for the Severity of Asthma

A few studies provide evidence for a potential relationship between long-term SO₂ exposure and the severity of asthma, as indicated by asthma severity scores, asthma control, and AHR ([Section 5.2.2.1](#)).

Evidence for the Development of Allergies

Epidemiologic evidence from a few long-term studies provides a link between long-term SO₂ exposure and respiratory allergies and allergic rhinitis among children ([Section 5.2.2.2](#)). However, uncertainties remain given the cross-sectional design of these studies. Two-pollutant models have begun to address the role of SO₂ exposure in the development of allergic rhinitis.

Evidence for Lung Function

Several studies evaluated the relationship between long-term SO₂ exposure and decrements in lung function ([Section 5.2.2.3](#)). Evidence supporting this relationship is limited because associations were inconsistent and because both PM and SO₂ were at high concentrations in the same areas, precluding determination of individual SO₂ effects. Potential confounding of long-term SO₂ exposure-related decrements in lung function and lung development by other pollutants, especially PM, was evaluated in only one study. This study found an attenuation of the effect in two-pollutant analyses. No changes in lung function were found in long-term animal toxicological studies at relevant SO₂ concentrations. The recent studies support conclusions made in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) that the available evidence was inadequate to infer a causal relationship between long-term exposure to SO₂ at ambient concentrations and changes in lung function.

Evidence for Respiratory Infection

Respiratory infection related to long-term SO₂ exposure is discussed in [Section 5.2.2.4](#). A limited number of the cross-sectional studies examined indicate associations between long-term SO₂ exposure and bronchitis or respiratory infection due to various infectious agents; findings were generally positive. While some animal toxicological studies reported alterations in specific host defense mechanisms, there is no evidence to support increases in bacterial or viral infections in animals exposed to SO₂ at relevant concentrations.

Evidence for the Development of Other Respiratory Diseases

Evidence for prevalence of bronchitis and/or COPD consists of generally positive associations found in cross-sectional studies ([Section 5.2.2.5](#)).

Evidence for Respiratory Mortality

Small positive associations between long-term exposure to SO₂ and respiratory mortality among adults were found in several cohort studies after adjustment for common potential confounders ([Section 5.2.2.6](#)). There is little evidence of respiratory health effects in adults in relation to long-term SO₂ exposure that could provide coherence with the observed associations with respiratory mortality among adults.

Conclusion

Taken together, epidemiologic and animal toxicological studies provide evidence that is suggestive of, but not sufficient to infer, a causal relationship between long-term SO₂ exposure and respiratory effects (see [Table 5-24](#)). The strongest evidence is provided by coherence of findings of epidemiologic studies showing associations between long-term SO₂ exposure and increases in asthma incidence among children and findings of animal toxicological studies that provide a pathophysiologic basis for the development of asthma. These latter studies demonstrated that repeated SO₂ exposure over several weeks resulted in Th2 polarization (or other Type 2 immune responses) and airway inflammation, key steps in allergic sensitization, in naive newborn animals. In addition, repeated SO₂ exposure over several weeks resulted in enhanced airway inflammation and some evidence of airway remodeling and increased airway responsiveness in allergic newborn animals. Toxicological studies involving repeated exposure to SO₂ over several days provide additional evidence of these effects. However, because the toxicological evidence in animals is limited, particularly for long-term exposure, some uncertainty remains regarding an independent effect of long-term SO₂ exposure on the development of asthma. In addition, potential confounding by other pollutants is unexamined, and largely unavailable, for epidemiologic studies of asthma among children. **However, multiple lines of evidence suggest that long-term SO₂ exposure results in a coherent and biologically plausible sequence of events that culminates in the development of asthma, especially allergic asthma, in children.**

5.3 Cardiovascular Effects

5.3.1 Short-Term Exposure

5.3.1.1 Introduction

The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) reviewed studies published through 2006 and concluded that “the evidence as a whole is inadequate to infer a causal relationship” between short-term exposure to SO₂ and cardiovascular health effects. Specifically, the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) found a lack of consistency with regard to short-term exposure to SO₂ and markers of heart rate variability (HRV), cardiac repolarization, discharges of implantable cardioverter defibrillators (ICDs), blood pressure (BP), blood markers of cardiovascular disease risk, the triggering of a myocardial infarction, or ED visits or hospital admission for cardiovascular diseases. Recent epidemiologic studies add to the evidence for effects of SO₂ exposure on a broader array of cardiovascular effects and mortality. Still, substantial uncertainties remain concerning exposure measurement error, the limited mechanistic evidence to describe a role for SO₂ in the initiation of key events in a proposed mode of action, and potential confounding by copollutants.

This section reviews the published studies pertaining to the cardiovascular effects of short-term exposure (i.e., up to 1 month) to SO₂ in humans and animals. Emphasis has been placed on studies published since the 2008 ISA for Sulfur Oxides, with the existing body of evidence serving as the foundation. The majority of the recent evidence is from epidemiologic studies, which examined the association of SO₂ exposure with MI, cerebrovascular disease, and other cardiovascular effects. With few exceptions, most epidemiologic studies model the association of 24-h avg SO₂ concentration with cardiovascular outcomes. There are no toxicological studies evaluating cardiovascular effects following 5–10 minute exposures to SO₂.

The previous ISA included a small number of animal toxicological studies of blood pressure ([Section 5.3.1.5](#)), HR and HRV ([Section 5.3.1.10](#)), arrhythmia frequency ([Section 5.3.1.3](#)) and controlled human exposure studies that examined effects on the autonomic nervous system ([Section 5.3.1.10](#)) from short-term exposure to SO₂. Since the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), no controlled human exposure studies and few animal toxicological studies have investigated the effects of short-term SO₂ exposure on the cardiovascular system. Results from the experimental studies included in the past and current reviews that evaluated cardiovascular effects of short-term SO₂

exposures of less than 2,000 ppb are summarized in the relevant outcome section and additional study details are summarized in Supplemental Table 5S-15 ([U.S. EPA, 2017c](#)).

Studies examining cardiovascular effects of sulfite exposure (via intraperitoneal injection, intravenous injection, etc.) are not included in this section because these studies generally involve exposures to sulfite that are higher than what is expected to occur following inhalation of SO₂ at ambient relevant concentrations. Some studies using prolonged exposures to 300 ppb and higher concentrations of SO₂ reported measurable changes in the concentrations of sulfite/S-sulfonate in plasma and tissues. A positive correlation was found between the concentration of inhaled SO₂ and plasma sulfite/S-sulfonate levels in humans exposed continuously to SO₂ (300–6,000 ppb) ([Gunnison and Palmes, 1974](#)). Similarly, a recent report in mice exposed to 5,000–20,000 ppb SO₂ for 7 days found a concentration-dependent increase in sulfite/S-sulfonate levels in lung, heart, and brain compared to controls ([Meng et al., 2005b](#)). These studies suggest that prolonged exposure to SO₂ at concentrations higher than typically found in ambient air may increase circulating sulfite, but these changes would be expected to be far less following ambient exposures of shorter duration. The literature on the distribution and metabolism of sulfite is discussed in [Sections 4.2.3](#) and [4.2.4](#). The potential role of sulfite in the induction of systemic effects, such as effects of the cardiovascular system, is discussed in [Section 4.3.4](#).

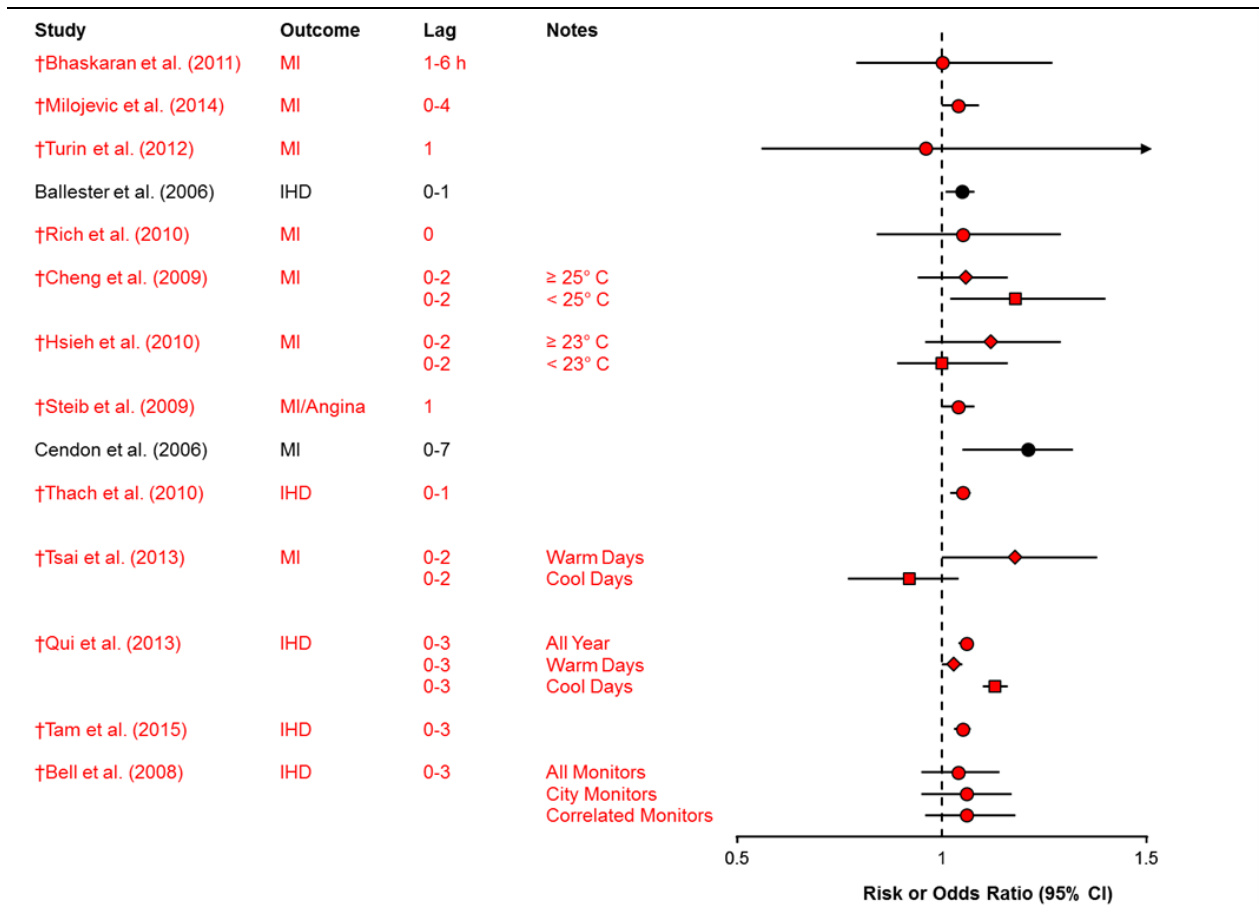
To clearly characterize the evidence underlying causality, the discussion of the evidence is organized into groups of related outcomes [myocardial infarction and ischemic heart disease (IHD) ([Section 5.3.1.2](#)), arrhythmia and cardiac arrest ([Section 5.3.1.3](#)), cerebrovascular disease ([Section 5.3.1.4](#)), hypertension ([Section 5.3.1.5](#)), venous thromboembolism (VTE) ([Section 5.3.1.6](#)), heart failure ([Section 5.3.1.7](#)), aggregated cardiovascular disease ([Section 5.3.1.8](#)), and cardiovascular mortality ([Section 5.3.1.9](#))]. Evidence for subclinical effects (e.g., heart rate variability, blood biomarkers of cardiovascular effects) of short-term exposure to SO₂ that potentially underlie the triggering or indication of various clinical events are discussed in [Section 5.3.1.10](#), and may provide biological plausibility for multiple outcomes.

5.3.1.2 Myocardial Infarction and Ischemic Heart Disease

Several lines of evidence are discussed in evaluating the relationship between short-term SO₂ exposure and MI. An MI, or heart attack, occurs as a consequence of IHD, resulting in insufficient blood flow to the heart that overwhelms myocardial repair mechanisms and leads to muscle tissue death. ICD codes for MI are classified within the group of IHDs; thus, studies in which IHD is evaluated will include any patients diagnosed with an

MI. Finally, acute MI may be characterized by ST-segment depression, a nonspecific marker of myocardial ischemia. The evaluation of evidence supporting a relationship between short-term SO₂ exposure and the triggering of an MI includes hospitalization and ED visits for MI or IHD and ST-segment amplitude changes.

The epidemiologic data available for review by the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) did not indicate an association between SO₂ and risk of MI. A number of additional studies based on administrative data of hospital admissions or ED visits or on clinical data are now available in [Figure 5-12](#). The air quality characteristics of the city, or across all cities, and the exposure assignment approach used in each MI-related hospital admission and ED visit study evaluated in this section are presented in [Table 5-25](#). The recent clinical registry studies provide inconsistent evidence for an association between MI and ambient SO₂, while multicity and single-city hospital admission and ED visit studies provide generally consistent evidence of an association. However, potential copollutant confounding and limited mechanistic evidence are still key uncertainties that make it difficult to interpret the results of these studies. Additionally, most studies examined 24-h avg exposure metrics for SO₂, which may not adequately capture the spatial and temporal variability in SO₂ concentrations ([Section 3.4.2](#)).



CI = confidence interval; IHD = ischemic heart disease; MI = myocardial infarction.

Note: Studies in red are recent studies. Studies in black were included in the 2008 ISA for Sulfur Oxides. All-year associations = circles; summer/warm-days associations = diamonds; winter/cold-days associations = squares. Relative risks are standardized to a 10-ppb or 40-ppb increase in sulfur dioxide for 24-hour avg and 1-hour max metrics, respectively. Lag times are reported in days, unless otherwise noted. Corresponding quantitative results are reported in Supplemental Table 5S-16 ([U.S. EPA, 2017c](#)). All results are from single pollutant models.

Figure 5-12 Results of studies of short-term sulfur dioxide exposure and hospital admissions for ischemic heart disease.

Table 5-25 Mean and upper percentile concentrations of sulfur dioxide from ischemic heart disease hospital admission and emergency department visit studies.

Study	Location Years	Exposure Assignment	Metric	Mean/Median Concentration ppb	Upper Percentile of Concentrations ppb
†Bhaskaran et al. (2011)	15 conurbations in England and Wales (2003–2006)	Fixed-site monitor from each conurbation (aggregated when more than one monitor)	1-h max	Mean: 1.9	75th: 3.4
†Milojevic et al. (2014)	230 acute hospitals in England and Wales (2003–2009)	Nearest monitor within 50-km distance from residence location	24-h avg	Median: 1.2	75th: 2.3
†Turin et al. (2012)	Takashima County, Japan (1988–2004)	Nearest monitor to Takashima County (20 km)	24-h avg	Mean: 3.9	75th: 4.8
Ballester et al. (2006)	14 Spanish cities (1995–1999)	Citywide average for each city	24-h avg	Mean: 2.9–15.6 across cities	90th: 4.8–28.8 across cities
†Rich et al. (2010)	New Jersey (2004–2006)	Closest of 14 monitor (those >10 km from monitor excluded)	24-h avg	NR	NR
†Cheng et al. (2009)	Kaohsiung, Taiwan (1996–2006)	Average across six monitoring stations	24-h avg	Mean: 9.33	75th: 11.69 Max: 31.26
†Hsieh et al. (2010)	Taipei, Taiwan (1996–2006)	Average across six monitoring stations	24-h avg	Mean: 4.36	75th: 5.48 Max: 17.82
†Stieb et al. (2009)	Seven Canadian cities (1992–2003)	Citywide average for each city	24-h avg	Mean: 2.6–10.0 across cities	75th: 3.3–13.4 across cities
Cendon et al. (2006)	São Paulo, Brazil (1998–1999)	Average across 13 monitoring stations	24-h avg	Mean: 5.6	95th: 12.1
†Thach et al. (2010)	Hong Kong, China (1996–2002)	Average across eight monitoring stations	24-h avg	Mean: 6.8	NR
†Tsai et al. (2012)	Taipei, Taiwan (1999–2009)	Average across six monitoring stations	24-h avg	Mean: 3.94	75th: 5.01 Max: 12.7

Table 5-25 (Continued): Mean and upper percentile concentrations of sulfur dioxide from ischemic heart disease hospital admission and emergency department visit studies

Study	Location Years	Exposure Assignment	Metric	Mean/Median Concentration ppb	Upper Percentile of Concentrations ppb
†Qiu et al. (2013a)	Hong Kong, China (1998, 2007)	Average across 14 monitoring stations	24-h avg	Mean: 7.4	NR
†San Tam et al. (2015)	Hong Kong, China (2001–2010)	Average across 13 monitoring stations	24-h avg	Mean: 7.6	75th: 9.3 Max: 51.9
†Bell et al. (2008)	Taipei, Taiwan (1995–2002)	Average across 13 monitoring stations; 5 within city limits; or 6 with correlations >0.75	24-h avg	Mean: 4.7	Max: 26.9

Avg = average; max = maximum; NR = not reported.

†Studies published since the 2008 ISA for Sulfur Oxides.

Some studies rely on clinical registries, which are generally less susceptible to misclassification of the outcome. Using data from the Myocardial Ischemia National Audit Project (MINAP) clinical registry, [Bhaskaran et al. \(2011\)](#) reported that hourly ambient SO₂ concentrations were not associated with risk of MI in a case-crossover study of 15 conurbations in England and Wales between 2003 and 2006. While no associations were reported in the population overall, there was some evidence of an association in subgroup analyses within older age groups (60–69, 70–79, and 80+) at inconsistent lag times. This study is unique because it included detailed data on the timing of MI onset in more than 79,000 patients, which allowed examination of the association with ambient SO₂ in the hours preceding MI. [Milojevic et al. \(2014\)](#) also used data from MINAP, from 2003 to 2009, and observed stronger evidence of an association between SO₂ concentrations and MI [4.3% (95% CI: –0.25, 8.8%) increase in risk of MI per 10-ppb increase in 24-h avg SO₂ at lag 0–4]. [Turin et al. \(2012\)](#) did not observe any association using data from the Takashima County Stroke and Acute Myocardial Infarction Registry in central Japan, although this study was likely underpowered to detect an association of the expected magnitude. None of the clinical registry studies examined copollutant models.

One prominent study from the previous 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) was conducted in 14 cities across Spain and found a 4.5% (95% CI: 1.3, 8.1%) increase in hospital admissions per 10-ppb shift in SO₂ for the composite endpoint of IHD, arrhythmias, and heart failure ([Ballester et al., 2006](#)). This association was still positive, but attenuated and no longer statistically significant after adjustment for CO or NO₂. It

was lessened in magnitude, but more precise (i.e., narrowed confidence intervals), with adjustment for TSP or O₃ in copollutant models (no quantitative results; results presented graphically). Several additional ED visit and hospital admission studies are now available. In a study of hospitalization in New Jersey, [Rich et al. \(2010\)](#) did not report strong evidence for an association between SO₂ and risk of hospital admissions for MI [OR: 1.05 (95% CI: 0.84, 1.29) per 10-ppb increase in 24-h avg SO₂ on the same day]. The inclusion of PM_{2.5} in a copollutant model did not reveal a positive association for SO₂ [OR: 0.91 (95% CI: 0.69, 1.21)]. In Kaohsiung, Taiwan, [Cheng et al. \(2009\)](#) reported an association between SO₂ concentrations and hospital admissions for MI, but only on days when the mean ambient temperature was <25°C. However, in copollutant models adjusting for PM₁₀, NO₂, or CO, SO₂ was no longer associated with increased admissions. Conversely, in Taipei, Taiwan, [Hsieh et al. \(2010\)](#) only observed an association between SO₂ and MI on warm days (≥23°C). Similar to the findings of [Cheng et al. \(2009\)](#), this association was no longer positive after adjustment for PM₁₀, NO₂, O₃, or CO in copollutant models. Most other studies have not considered copollutant models.

A study using data from 14 hospitals in seven Canadian cities found a 4.2% (95% CI: 0.4, 8.0%) increase in risk of ED visits for the composite endpoint of acute MI or angina per 10-ppb increase in SO₂ on the previous day ([Stieb et al., 2009](#)). Most ([San Tam et al., 2015](#); [Qiu et al., 2013a](#); [Tsai et al., 2012](#); [Thach et al., 2010](#); [Cendon et al., 2006](#); [Martins et al., 2006](#)) but not all ([Bell et al., 2008](#)) studies using data from individual cities have found associations between SO₂ concentrations and risk of hospital admissions or ED visits for ischemic heart disease or MI. None of the single-city studies evaluated potential copollutant confounding, and all of the studies in this section used fixed-site monitors to measure ambient SO₂. The limitations of these monitors in capturing spatial variation in SO₂ has been noted previously ([Section 3.4.2](#)).

ST-Segment Changes

ST-segment changes (either ST-segment elevation or depression) on the electrocardiogram are considered a nonspecific marker of myocardial ischemia. While the 2008 ISA for Sulfur Oxides did not review any epidemiologic studies of ambient SO₂ concentrations and markers of myocardial ischemia, one subsequent study reported an association. [Chuang et al. \(2008\)](#) conducted a repeated-measures study in adults with a history of coronary heart disease (CHD) and examined the association between ambient pollutants and ST-segment level changes. This study found an odds ratio of 3.0 (95% CI: 1.8, 5.5) for ST-segment depression of ≥0.1 mm per 10-ppb increase in SO₂ over the previous 24 hours. This finding was generally unchanged after additional control for PM_{2.5} and BC in copollutant models.

Summary of Myocardial Infarction and Ischemic Heart Disease

In summary, while evidence from epidemiologic studies suggests a potential association between ambient SO₂ concentrations and rates of hospital admissions or ED visits for MI or ischemic heart diseases in single-pollutant models, these associations may be the result of confounding by other pollutants. While three studies based on clinical data report inconsistent evidence regarding associations between ambient SO₂ concentrations and risk of MI, the majority of studies relying on MI hospital admission and ED visit data observed either seasonal or year-round associations with SO₂. However, some of these associations were either attenuated or no longer present after controlling for potential copollutant confounding ([Hsieh et al., 2010](#); [Cheng et al., 2009](#); [Ballester et al., 2006](#)), leaving uncertainties regarding the independent effect of short-term SO₂ exposure. In congruence with the evidence from hospital admission and ED visit studies, there was limited evidence from a single study indicating that SO₂ may be associated with ST-segment changes on the electrocardiogram in patients with a history of coronary heart disease. Most studies examined 24-h avg exposure metrics for SO₂, which may not adequately capture the spatial and temporal variability in SO₂ concentrations ([Section 5.2.1.2](#)). No experimental studies have been conducted to evaluate measures of ischemic heart disease or MI following short-term SO₂ exposure. Overall, despite some epidemiologic evidence of an association between short-term exposure to SO₂ and hospital admissions and ED visits for ischemic heart disease and MI, uncertainties regarding copollutant confounding continue to impede the determination of an independent SO₂ effect.

5.3.1.3 Arrhythmias and Cardiac Arrest

The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) concluded that the evidence available at the time did not suggest that SO₂ has an effect on cardiac arrhythmias. There continues to be essentially no epidemiologic or toxicological evidence suggestive of such a relationship.

[Metzger et al. \(2007\)](#) examined 518 patients with ICDs with 6,287 tachyarrhythmic event-days over a 10-year period in Atlanta, GA and found no association between SO₂ concentrations and the risk of tachyarrhythmias, either overall or in analyses limited to more severe tachyarrhythmic events, or stratified by season or the presence of a recent past arrhythmic event (results for this study and other studies in this section can be found in [Table 5-26](#)). A similar study in London, England also found limited evidence of an association between SO₂ concentrations and arrhythmic risk ([Anderson et al., 2010](#)). [Anderson et al. \(2010\)](#) reported an increase in risk of ICD activations corresponding to an

increase in ambient SO₂, but the confidence interval was wide [OR: 1.35 (95% CI: 0.75, 2.41) per 10-ppb increase in SO₂ at lag days 0–1]. Similarly, a study in Boston, MA of ICD activations was even less precise, offering no evidence of an association [32.0% (95% CI: –48.5, 336.2%) increase in ICD activations per 10-ppb increase in SO₂ concentrations at lag 1] ([Link et al., 2013](#)). Additionally, a multicity study in Canada ([Stieb et al., 2009](#)) and a large single-city study in Taipei, Taiwan ([Tsai et al., 2009](#)) have reported finding no association between SO₂ and ED visits for arrhythmias, while a large single-city study in Shanghai, China reported a positive association that was attenuated and no longer positive in a copollutant model adjusted for NO₂ ([Zhao et al., 2014](#)).

The majority of out-of-hospital cardiac arrests (OHCA) are due to cardiac arrhythmias. [Dennekamp et al. \(2010\)](#) considered the association between ambient pollutants and OHCA among 8,434 cases identified through the Victorian Cardiac Arrest Registry in Melbourne, Australia and found null and/or imprecise associations (i.e., wide 95% CIs) between SO₂ concentrations and risk of OHCA. A similar approach was used by [Silverman et al. \(2010\)](#) with data from 8,216 OHCAs in New York City. Quantitative results for SO₂ were not provided, but graphs showed a null association between OHCA and year-round SO₂ concentrations. [Silverman et al. \(2010\)](#) also presented season-specific analyses graphically, demonstrating that out-of-hospital cardiac arrests were positively but imprecisely (i.e., wide 95% CI) associated with SO₂ concentrations during the warm season. Two additional case-crossover studies of OHCA in Perth, Australia ([Straney et al., 2014](#)) and Helsinki, Finland ([Rosenthal et al., 2013](#)) observed null associations with ambient SO₂. In contrast, [Kang et al. \(2016\)](#) observed an association between 24-h avg SO₂ and OHCA in Seoul, South Korea at individual lag days 0 through 3 (no quantitative results; results presented graphically).

One animal toxicological study ([Nadziejko et al., 2004](#)) evaluated arrhythmia frequency in rats following short-term SO₂ exposure and reported no significant changes in spontaneous arrhythmias (irregular, delayed, or premature beats).

Table 5-26 Epidemiologic studies of arrhythmia and cardiac arrest.

Study	Location and Years Sample Size	Mean and Upper Concentration SO ₂ (ppb)	Exposure Assessment	Selected Effect Estimates ^a 95% CI
†Metzger et al. (2007)	Atlanta, GA 1993–2002 (n = 518)	1-h max: 15.5 90th percentile: 36 Max: 149	Central monitor	All tachyarrhythmic events (OR); year round Lag 0: 1.00 (0.94, 1.08) Warm season Lag 0: 1.06 (0.98, 1.25) Cold season Lag 0: 0.97 (0.91, 1.05) Cardiac pacing or defibrillation (OR): Lag 0: 0.98 (0.88, 1.09) Defibrillation (OR): Lag 0: 1.01 (0.98, 1.24)
†Anderson et al. (2010)	London, U.K. 1995–2003 [n = 705 (5,462 device activations)]	24-h avg: 1.03 75th percentile: 1.15 Max: 2.67	Citywide avg	ICD activations (OR); Lag 01: 1.35 (0.75, 2.41) Lag 05: 1.71 (0.69, 4.27) Correlations: PM ₁₀ : 0.48, PM _{2.5} : 0.42, BS: 0.35, SO ₄ ²⁻ : 0.19, PNC: 0.29, NO ₂ : 0.60, NO: 0.44, NO _x : 0.49, O ₃ : -0.36
†Link et al. (2013)	Boston, MA 2006–2010 [n = 176 (328 atrial fibrillation episodes ≥30 sec)]	24-h avg: 3.2 75th percentile: 4	Citywide avg	ICD activations (percent change); Lag 1: 32.0 (-48.5, 336.2) Correlations: CO: -0.06 to 0.75, NO ₂ : 0.05 to 0.69, O ₃ : -0.52 to -0.18, PM ₁₀ : 0.27 to 0.55, PM _{2.5} : 0.01 to 0.67
†Stieb et al. (2009)	Seven Canadian cities 1992–2003 (n = 45,160 ED visits)	24-h avg: 2.6 to 10 across cities 75th percentile: 3.3 to 13.4 across cities	Citywide avg for each city	Dysrhythmia ED visits (percent change); Lag 0: -1.4 (-6.0, 3.4) Lag 1: 0.8 (-6.4, 8.6) Lag 2: -5.0 (-9.2, -0.6) Correlations: PM ₁₀ : 0.52, NO ₂ : 0.43, CO: 0.24, O ₃ : 0.09
†Tsai et al. (2009)	Taipei, Taiwan 2000–2006 (n = 21,581 ED visits)	24-h avg: 3.93 75th percentile: 5.02 Max: 12.7	Citywide avg	Arrhythmia ED visits (OR); ≥23°C: 1.04 (0.88, 1.23) <23°C: 1.04 (0.88, 1.27) Correlations: PM ₁₀ : 0.52, NO ₂ : 0.43, CO: 0.24, O ₃ : 0.09
†Zhao et al. (2014)	Shanghai, China 2010–2011 (n = 56,940 outpatient visits)	24-h avg: 11.1 75th percentile: 14.1 Max: 49.6	Central monitor	Arrhythmia outpatient visits (percent change); Lag 0: 1.06 (1.04, 1.07)

Table 5-26 (Continued): Epidemiologic studies of arrhythmia and cardiac arrest.

Study	Location and Years Sample Size	Mean and Upper Concentration SO ₂ (ppb)	Exposure Assessment	Selected Effect Estimates ^a 95% CI
†Dennekamp et al. (2010)	Melbourne, Australia 2003–2006 (n = 8,434 OHCA)	24-h avg: 0.49 75th percentile: 0.76	Central monitor	OHCA (percent change); Lag 0: -10.0 (-40.3, 64.0) Lag 1: 6.9 (-34.9, 75.6) Lag 2: 0.8 (-39.0, 66.7) Lag 01: -0.7 (-34.9, 75.6)
†Silverman et al. (2010)	New York City, NY 2003–2006 (n = 8,216 OHCA)	24-h avg: 6.3 (median) 75th percentile: 9.6 95th percentile: 18	Citywide avg	No quantitative results; results presented graphically. Null association between OHCA and year-round SO ₂ concentrations. OHCA positively but imprecisely (i.e., wide 95% CI) associated with ambient SO ₂ during the warm season
†Straney et al. (2014)	Perth, Australia 2000–2010 (n = 8,551 OHCA)	1-h avg: 0.4 (median) 75th percentile: 0.9 95th: 3.5	Nearest monitor	OHCA (OR); Lag 0: 0.91 (0.71, 1.17)
†Rosenthal et al. (2013)	Helsinki, Finland 1998–2006 (n = 2,134 OHCA)	24-h avg: 1.5	Citywide avg	OHCA (OR); Lag 0: 0.93 (0.58, 1.44) Lag 1: 0.68 (0.42, 1.08) Lag 2: 1.08 (0.68, 1.66) Lag 3: 1.00 (0.63, 1.55) Lag 03: 0.86 (0.42, 1.55)
†Kang et al. (2016)	Seoul, South Korea 2006–2013 (n = 28,315 OHCA)	24-h avg: 2.1 75th percentile: 2.5 Max: 8.1		No quantitative results; results presented graphically. Positive, statistically significant associations at single day lags 0 through 3. Null associations at lags 4 and 5.

Avg = average; BS = black smoke; CI = confidence interval; CO = carbon monoxide; ED = emergency department; ICD = implantable cardioverter defibrillators; max = maximum; n = sample size; N = population number; NO = nitric oxide; NO₂ = nitrogen dioxide; NO_x = the sum of NO and NO₂; O₃ = ozone; OHCA = out-of-hospital cardiac arrests; OR = odds ratio; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; PNC = particle number concentration; SO₂ = sulfur dioxide; SO₄²⁻ = sulfate.

^aEffect estimates are standardized to a 10-ppb or 40-ppb increase in SO₂ concentration for 24-h avg and 1-h max metrics, respectively.

All Lag times are in days, unless otherwise noted.

†Studies published since the 2008 ISA for Sulfur Oxides.

In summary, studies of patients with implantable cardioverter defibrillators, hospital admissions for arrhythmias, and out-of-hospital cardiac arrests do not provide evidence to support the presence of an association between ambient SO₂ concentrations and arrhythmias. Most of these studies have been focused on other pollutants and, therefore, have not explored whether such an association might exist in certain subgroups. Additionally, the majority of studies used fixed-site monitors to estimate ambient SO₂ exposure, which have noted limitations in capturing spatial variation in SO₂ that generally lead to attenuation and loss of precision in the effect estimates ([Section 3.4.4](#)). One toxicological study also found no evidence for arrhythmias following short-term SO₂ exposure.

5.3.1.4 Cerebrovascular Diseases and Stroke

Results among the studies reviewed in the 2008 ISA for Sulfur Oxides were inconsistent with regard to the association between ambient SO₂ concentrations and hospital admissions or ED visits for cerebrovascular diseases or stroke (a specific form of cerebrovascular disease). Many additional studies are now available for consideration (study details and results presented in [Table 5-27](#) and [Figure 5-13](#)). In Edmonton, AB, [Szyszkowicz \(2008\)](#) reported that risk of ED visits for ischemic stroke was linked to SO₂ concentrations, but this association was observed only in subgroup analyses stratified by sex, season, and age. A subsequent study in Vancouver, BC, found that SO₂ was associated with risk of ED visits for ischemic stroke in the population overall [OR: 2.09 (95% CI: 1.23, 3.52) per 10-ppb increase in SO₂ at lag 3] ([Szyszkowicz et al., 2012a](#)). The association was generally unchanged after adjustment for O₃ in a copollutant model, and attenuated, although still positive, after adjustment for CO [OR: 1.73 (95% CI: 1.00, 3.10)]. [Chen et al. \(2014b\)](#) also observed an association between SO₂ and ischemic stroke at longer lags in Edmonton, AB. In Brazil, [Costa Nascimento et al. \(2012\)](#) observed a 7.8% (95% CI: 0.0, 16.5%) increase in risk of hospital admissions of stroke per 10-ppb increase in 24-h avg SO₂ at lag 0. [Zheng et al. \(2013\)](#) reported a small association between SO₂ concentrations and risk of hospital admission for cerebrovascular disease [1.7% increase (95% CI: 0.5, 2.8%) per 10-ppb increase in 24-h avg SO₂ at lag 2] in Lanzhou, a heavily polluted city in China with a high observed mean daily concentration of SO₂ (30.19 ppb) over the 5-year study period. The association was as strong, or stronger, after adjustment for PM₁₀ [1.8% increase (95% CI: 0.4, 3.2%)] or NO₂ [2.6% increase (95% CI: 1.4, 3.7%)] in copollutant models. In central Japan, [Turin et al. \(2012\)](#) found that the risk of hemorrhagic stroke was associated with SO₂ concentrations, but found no association with other types of stroke. However, the 95% CI for the

hemorrhagic stroke association was wide, and copollutant confounding was not considered.

Table 5-27 Mean and upper percentile concentrations of sulfur dioxide from cerebrovascular disease and stroke-related hospital admission and emergency department visit studies.

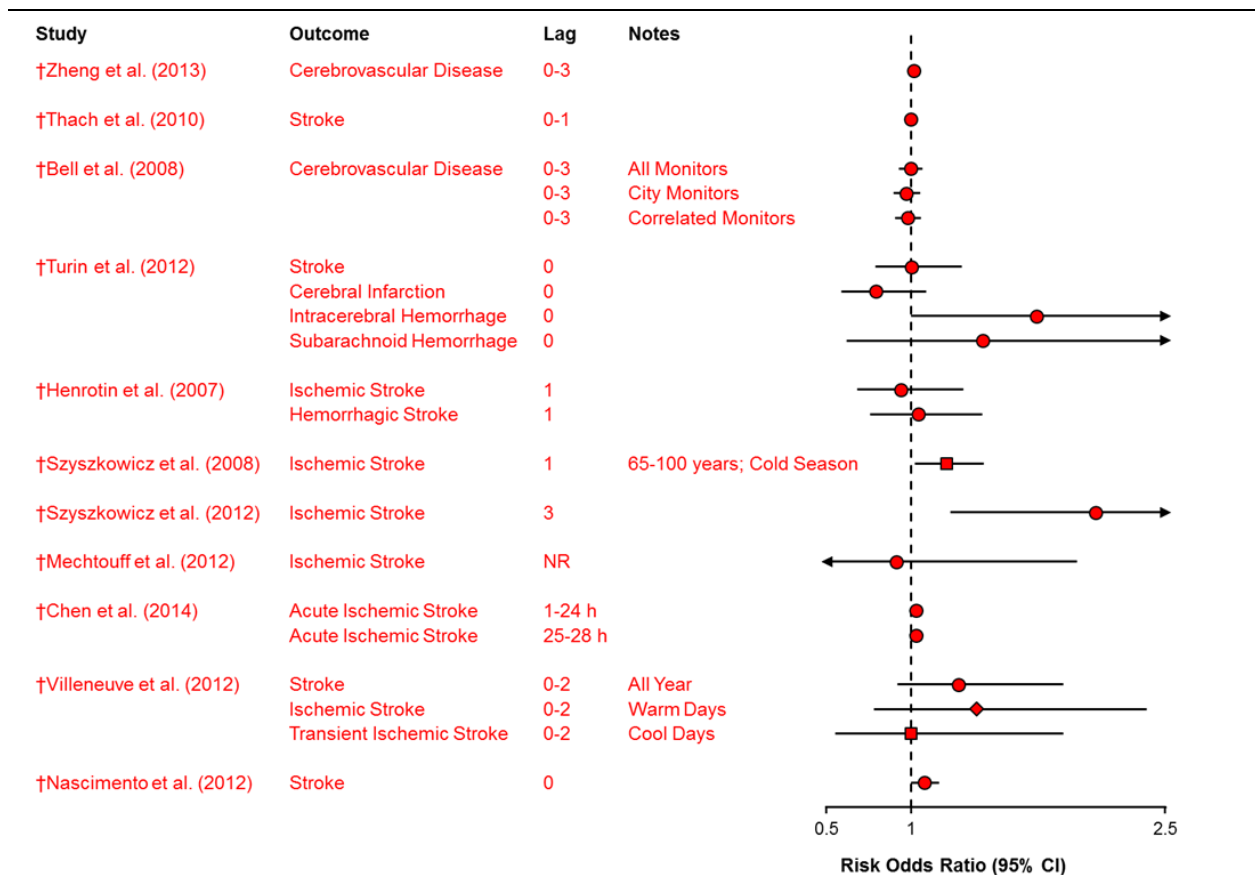
Study	Location Years	Exposure Assignment	Metric	Mean/Median Concentration ppb	Upper Percentile of Concentrations ppb
† Zheng et al. (2013)	Lanzhou, China (2001–2005)	Average across four monitoring stations	24-h avg	Mean: 30.19	75th: 40.46 Max: 141.60
† Thach et al. (2010)	Hong Kong, China (1996–2002)	Average across eight monitoring stations	24-h avg	Mean: 6.79	NR
† Bell et al. (2008)	Taipei, Taiwan (1995–2002)	Average across 13 monitoring stations; 5 within city limits; or 6 with correlations >0.75	24-h avg	Mean: 4.7	Max: 26.9
† Turin et al. (2012)	Takashima County, Japan (1988–2004)	Nearest monitor to Takashima County (20 km)	24-h avg	Mean: 3.9	75th: 4.8
Henrotin et al. (2007)	Dijon, France (1994–2004)	Fixed-site monitor	24-h avg	Mean: 2.63	75th: 3.44 Max: 24.81
† Szyszkowicz (2008)	Edmonton, AB (1992–2002)	Average across three monitoring stations	24-h avg	Mean: 2.6	NR
† Szyszkowicz et al. (2012a)	Vancouver, BC (1999–2003)	Average across 11 monitoring stations	24-h avg	Mean: 2.5	NR
† Mechtouff et al. (2012)	Lyon, France (2006–2007)	Average across five monitoring stations	24-h avg	Mean: 2.02	75th: 2.67 Max: 22.52
† Chen et al. (2014b)	Edmonton, AB (1998–2002)	Average across three monitoring stations	1-h avg	Mean: 2.0	95th: 6.7
† Villeneuve et al. (2012)	Edmonton, AB (2003–2009)	Average across three monitoring stations	24-h avg	Mean: 1.5	75th: 1.9

Table 5-27 (Continued) Mean and upper percentile concentrations of sulfur dioxide from cerebrovascular disease and stroke related hospital admission and emergency department visit studies.

Study	Location Years	Exposure Assignment	Metric	Mean/Median Concentration ppb	Upper Percentile of Concentrations ppb
† Costa Nascimento et al. (2012)	São Paulo, Brazil (2007–2008)	Fixed-site monitor	24-h avg	NR	NR

Avg = average; max = maximum; NR = not reported.

†Studies published since the 2008 ISA for Sulfur Oxides.



CI = confidence interval; NR = not reported.

Note: Studies in red are recent studies. Studies in black were included in the 2008 ISA for Sulfur Oxides. All-year associations = circles; summer/warm-days associations = diamonds; winter/cold-days associations = squares. Relative risks are standardized to a 10 or 40-ppb increase in sulfur dioxide for 24-hour avg and 1-hour max metrics, respectively, but not standardized for other metrics [e.g., (Chen et al., 2014b)]. Lag times are reported in days, unless otherwise noted. Corresponding quantitative results are reported in Supplemental Table 5S-17 (U.S. EPA, 2017c). All results are from single pollutant models.

Figure 5-13 Results of studies of short-term sulfur dioxide exposure and hospital admissions for cerebrovascular disease and stroke.

In contrast to the studies that reported some evidence of an association between SO₂ concentrations and cerebrovascular disease, a number of studies observed null or imprecise associations. In an effort to reduce uncertainty related to the use of fixed-site monitors, Bell et al. (2008) estimated SO₂ exposure over the entire Taipei, Taiwan area (average of 13 monitors), within Taipei City only (average of 5 monitors), and using a subset of monitors where all pairs of monitors had SO₂ correlations greater than 0.75 (6 monitors). Using three exposure metrics, the authors did not observe an association between SO₂ and risk of hospital admission for cerebrovascular diseases. Contrary to other studies that reported associations between SO₂ concentrations and hospital

admissions and ED visits for stroke in Canada ([Chen et al., 2014b](#); [Szyszkowicz et al., 2012a](#); [Szyszkowicz, 2008](#)), [Villeneuve et al. \(2012\)](#) reported null and/or imprecise (i.e., wide 95% CI) associations between SO₂ and all stroke, ischemic stroke, and hemorrhagic stroke in Edmonton, AB. Studies in Hong Kong ([Thach et al., 2010](#)), Dijon, France ([Henrotin et al., 2007](#)), and Lyon, France ([Mechtouff et al., 2012](#)) also observed null associations between SO₂ concentrations and rates of hospital admission for stroke.

Overall, findings for the association between SO₂ and cerebrovascular diseases continue to be inconsistent across studies. As for other outcomes, associations reported from single pollutant models in some locations may be at least partly due to confounding by other pollutants.

5.3.1.5 Blood Pressure and Hypertension

Based on the data available at the time, the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) concluded that the overall evidence was insufficient to determine that SO₂ has an effect on blood pressure. Recent evidence provides limited and inconsistent evidence for changes in blood pressure associated with short-term exposure to SO₂.

Epidemiologic Studies

A number of longitudinal studies measured BP in subjects in Beijing before, during, and after the 2008 Beijing Olympics when citywide air pollution control measures substantially reduced ambient levels of most criteria pollutants. [Huang et al. \(2012\)](#) measured blood pressure repeatedly on up to four occasions in 40 participants with pre-existing cardiovascular disease in Beijing, including one measurement during the 2008 Beijing Olympics when citywide air pollution control measures reduced ambient SO₂ concentrations by up to 50%. [Huang et al. \(2012\)](#) found a small decrement in diastolic blood pressure per IQR increase in prior 30-minute exposure to SO₂ [−0.9 mm Hg (95% CI: −2.0, 0.2 mm Hg)], but observed a null association between ambient SO₂ and systolic blood pressure. Focusing on healthy young adults, [Rich et al. \(2012\)](#) and [Zhang et al. \(2013\)](#) observed associations between SO₂ and blood pressure in repeated-measures studies conducted before, during, and after the 2008 Beijing Olympics (no quantitative results; results presented graphically). Using the same protocol, [Zhang et al. \(2013\)](#) and [Rich et al. \(2012\)](#) observed a positive association between 24-h avg SO₂ and systolic blood pressure, but an inverse association between 24-h avg SO₂ and diastolic blood pressure. The negative association between SO₂ and diastolic blood pressure was relatively unchanged after adjustment for PM_{2.5}, EC, or sulfate, while the

association between SO₂ and systolic blood pressure was also robust to sulfate, but attenuated, although still positive, after adjustment for PM_{2.5} or EC ([Zhang et al., 2013](#)).

In another repeated measures study, [Kim et al. \(2016b\)](#) observed positive associations between short-term SO₂ concentrations and systolic blood pressure, diastolic blood pressure, and mean arterial pressure among 560 older adults living in Seoul, South Korea. A pair of cross-sectional studies reported conflicting evidence of an association. Examining data from 7,578 participants in the Taiwanese Survey on Prevalence of Hyperglycemia, Hyperlipidemia, and Hypertension, [Chuang et al. \(2010\)](#) concluded that there is “no significant association” between SO₂ concentrations and blood pressure (no quantitative results presented). However, in a cross-sectional analysis of data from 9,238 participants in the Taiwan Community-based Integrated Screening program, [Chen et al. \(2012d\)](#) found a 4.0 mm Hg (95% CI: 3.0 to 5.0 mm Hg) increase in diastolic blood pressure per 10-ppb increase in SO₂ concentrations 2 days earlier, and a 1.6 mm Hg (95% CI: 0.15, 3.1 mm Hg) decrease in systolic blood pressure related to SO₂ concentrations 3 days earlier.

In addition to longitudinal and cross-sectional studies, a few new studies examined ED visits for hypertension. In Beijing, [Guo et al. \(2010\)](#) observed a 10.0% (95% CI: 1.1, 19.7%) increase in risk of ED visits for hypertension per 10-ppb increase in 24-h avg SO₂ on the same day. The association was attenuated, but still positive, in a copollutant model adjusting for PM₁₀ [6.7% (95% CI: -3.4, 17.9%) increase at lag 0] and no longer present in a copollutant model adjusting for NO₂ [-0.8% (95% CI: -12.8, 13.0%) change at lag 0]. Inconsistent results were reported in two studies of ED visits for hypertension in Canada. In a case-crossover study in Calgary and Edmonton, [Brook and Kousha \(2015\)](#) reported positive associations between ED visits for hypertension and 24-h avg SO₂ concentrations for males [OR: 2.50 (95% CI: 1.00, 5.87) per 10-ppb increase] and females [OR: 2.59 (95% CI: 1.12, 5.61) per 10-ppb increase]. Conversely, in Edmonton, [Szyszkowicz et al. \(2012b\)](#) observed that ED visits for hypertension were both positively and negatively associated with SO₂ depending on the lag time examined.

Experimental Studies

Several experimental studies examined hypertension and blood pressure following SO₂ exposure. Study characteristics are summarized in Supplemental Table 5S-15 ([U.S. EPA, 2017c](#)). One controlled human exposure study reported no change in mean arterial pressure following SO₂ exposure ([Routledge et al., 2006](#)). Two animal toxicological studies have examined blood pressure following SO₂ exposure ([Halinen et al., 2000b](#); [Halinen et al., 2000a](#)). In both studies, SO₂ was administered intratracheally to hyperventilated guinea pigs in cold, dry air. These studies reported increases in blood

pressure following cold, dry air exposure with and without SO₂ and did not determine whether there were any effects on blood pressure caused by SO₂ that may not be attributable to cold, dry air exposure.

Summary of Blood Pressure

In summary, epidemiologic studies evaluating the association between ambient SO₂ concentrations and blood pressure remain inconsistent with most relying on fixed-site monitors and few examining the potential for copollutant confounding. Experimental studies provide no additional evidence for SO₂-induced changes in blood pressure. The most informative studies to date found no evidence of within-person changes in blood pressure despite relatively large changes in SO₂ concentrations during the Beijing Olympics. Experimental studies do not demonstrate effects of SO₂ on blood pressure. As such, the current evidence does not support the presence of an association between ambient SO₂ and blood pressure.

5.3.1.6 Venous Thromboembolism

VTE is a term that includes both deep vein thrombosis (DVT) and pulmonary embolism (PE). DVT occurs when a blood clot develops in the deep veins, most commonly in the lower extremities. A part of the clot can break off and travel to the lungs, causing a PE, which can be life threatening.

There were no epidemiologic studies of VTE available for the 2008 ISA for Sulfur Oxides. One recent study covering the metropolitan region of Santiago, Chile, found a 10.8% (95% CI: 3.3, 15.7%) and 8.5% (95% CI: 4.0, 13.2%) increased rate of hospital admission for venous thrombosis and pulmonary embolism, respectively, per 10-ppb increase in 24-h avg SO₂ concentrations ([Dales et al., 2010](#)). Copollutant models were not evaluated. Given the limited epidemiologic evidence, the association between ambient SO₂ concentrations and venous thromboembolism is unclear.

5.3.1.7 Heart Failure

Results among the studies reviewed in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) were inconsistent with regard to the association between ambient SO₂ concentrations and hospital admissions or ED visits for heart failure. A small number of additional studies are now available, including a multicity study of seven Canadian cities ([Stieb et al., 2009](#)). [Stieb et al. \(2009\)](#) observed an imprecise association (i.e., wide 95%

CI) between 24-h avg SO₂ concentrations on the previous day and ED visits for heart failure [3.0% (95% CI: -1.9, 8.2%) increase in risk of ED visits per 10-ppb increase in SO₂]. Similarly, in Guangzhou, China, [Yang et al. \(2014a\)](#) observed a 14.5% increase (95% CI: 6.1, 23.2%) in emergency ambulance dispatches for heart failure per 10-ppb increase in 24-h avg SO₂ concentrations on the same day. This association was slightly attenuated, but still positive and statistically significant in copollutant models adjusting for PM₁₀ [13.1% (95% CI: 3.3, 23.4%)] and NO₂ [11.3% (95% CI: 1.7, 21.5%)]. In contrast, [Yang \(2008\)](#) did not observe evidence of a positive association between ambient SO₂ exposure and heart failure in Taipei, Taiwan.

In summary, the available epidemiologic evidence is limited and inconsistent and, therefore, does not support the presence of an association between ambient SO₂ concentrations and hospital admissions or ED visits for heart failure.

5.3.1.8 Aggregated Cardiovascular Disease

Many epidemiologic studies consider the composite endpoint of all cardiovascular diseases, which typically includes all diseases of the circulatory system (e.g., heart diseases and cerebrovascular diseases). This section summarizes the results of epidemiologic studies evaluating the association between ambient SO₂ concentrations and ED visits or hospitalizations for all cardiovascular diseases. [Table 5-28](#) presents study details and air quality characteristics of the city, or across all cities, from the U.S. and Canadian cardiovascular-related hospital admission and ED visit studies evaluated in the 2008 ISA for Sulfur Oxides and those more recent.

Table 5-28 Mean and upper percentile concentrations of sulfur dioxide from cardiovascular-related hospital admission and emergency department visit studies: U.S. and Canadian studies from the 2008 ISA for Sulfur Oxides and recent studies.

Study	Location Years	Type of Visit (ICD 9/10)	Metric	Mean Concentration ppb	Upper Percentile of Concentrations ppb
U.S.					
Gwynn et al. (2000)	Buffalo and Rochester, NY (1988–1990)	Hospital admissions: circulatory (401–405, 410–417)	24-h avg	12.2	Max: 37.7
†Ito et al. (2011)	New York City, NY (2000–2006)	Hypertensive diseases (402, I11); MI (410, I21–I22); IHD (414, I25); dysrhythmias (427, I48); heart failure (428, I50); and stroke (430–439, I60–I69)	24-h avg	7.4	
Koken et al. (2003)	Denver, CO (1993–1997)	Discharge data from Agency for Healthcare Research and Quality database: Acute MI (410.00–410.92), atherosclerosis (414.00–414.05), pulmonary heart failure (416.0–416.9), dysrhythmia (427.0–427.9), CHF (428.0)	24-h avg	5.7	Max: 18.9
Low et al. (2006)	New York City, NY (1995–2003)	Ischemic stroke (433–434), undetermined stroke (436); monitored intake in 11 hospitals (ED or clinic visits). Excluded stroke patients admitted for rehabilitation	24 h avg	10.98	Max: 96.0
Metzger et al. (2004)	Atlanta, GA (1993–2000)	ED visits: IHD (410–414); acute MI (410); dysrhythmias (427); cardiac arrest (427.5); CHF (428); peripheral and cerebrovascular disease (433–437, 440, 443–444, 451–453); atherosclerosis (440); stroke (436)	1-h max:	11.0 (median)	90th: 39

Table 5-28 (Continued): Mean and upper percentile concentrations of sulfur dioxide from cardiovascular related hospital admission and emergency department visit studies: U.S. and Canadian studies from the 2008 ISA for Sulfur Oxides and recent studies.

Michaud et al. (2004)	Hilo, HI (1997–2001)	ED visits Heart (410–414, 425–429)	24-h avg	1.92 (all hourly measurements)	Max: 447 (all hourly measurements)
Moolgavkar (2003) Moolgavkar (2000)	Cook County, IL; Los Angeles County, CA; Maricopa County, AZ (1987–1995)	Hospital admissions: CVD (390–429); cerebrovascular disease (430–448)	24-h avg	Cook: 6 (median) Los Angeles: 2 (median) Maricopa: 2 (median)	Cook: Max: 36 Los Angeles: Max: 16 Maricopa: Max: 14
Morris et al. (1995)	Los Angeles, CA; Chicago, IL; Philadelphia, PA; New York City, NY; Detroit, MI; Houston, TX; Milwaukee, WI (1986–1989)	Hospital admissions: CHF (428)	1-h max	Los Angeles: 10 Chicago: 25 Philadelphia: 29 New York City: 32 Detroit: 25 Houston: 18 Milwaukee: 17	NR
Peel et al. (2007)	Atlanta, GA (1993–2000)	ED visits: IHD (410–414), dysrhythmia (427), CHF (428), peripheral vascular and cerebrovascular disease (433–437, 440, 443, 444, 451–453)	1-h max	16.5 (17.1)	90th: 39
†Rich et al. (2010)	New Jersey (2004–2006)	Hospital Admissions: transmural infarction (410.0, 410.1, 410.2, 410.3, 410.4, 410.5, 410.6), nontransmural infarction (410.7)	24-h avg	NR	NR
Schwartz and Morris (1995)	Detroit, MI (1986–1989)	Hospital discharge: IHD (410–414), CHF (428), dysrhythmia (427)	24-h avg	25.4	90th: 44.0
Schwartz (1997)	Tuscon, AZ (1988–1990)	Hospital discharge: CVD (390–429)	24-h avg	4.6	90th: 10.1
Tolbert et al. (2007)	Atlanta, GA (1993–2004)	ED visits: CVD (410–414, 427, 428, 433–437, 440, 443–445, 451–453)	1-h max	14.9	Max: 149.0

Table 5-28 (Continued): Mean and upper percentile concentrations of sulfur dioxide from cardiovascular related hospital admission and emergency department visit studies: U.S. and Canadian studies from the 2008 ISA for Sulfur Oxides and recent studies.

Ulirsch et al. (2007)	Southeast Idaho (1994–2000)	Hospital admissions and medical visits: CVD (390–429)	NR	3.0	90th: 7.9, 7.7 Max: 30.3, 30.3 (two time series examined)
Wellenius et al. (2005b)	Birmingham, AL; Chicago, IL; Cleveland, OH; Detroit, MI; Minneapolis, MN; New Haven, CT; Pittsburgh, PA; Seattle, WA (1986–1999)	Hospital admissions: ischemic stroke, primary diagnosis of acute but ill-defined cerebrovascular disease or occlusion of the cerebral arteries; HS, primary diagnosis of intracerebral hemorrhage. (ICD codes not provided)	24-h avg	6.22 (median)	90th: 16.17
Wellenius et al. (2005a)	Allegheny County, PA (1987–1999)	Hospital admissions: CHF (428)	24-h avg	14.78 (9.88)	95th: 33.93
Canada					
Burnett et al. (1997)	Metropolitan Toronto (East York, Etobicoke, North York, Scarborough, Toronto, York) (1992–1994)	Hospital discharge: IHD (410–414); cardiac dysrhythmias (427); heart failure (428); all cardiac (410–414, 427, 428)	1-h max	7.9	Max: 26
Burnett et al. (1999)	Metropolitan Toronto (East York, Etobicoke, North York, Scarborough, Toronto, York) (1980–1994)	IHD (410–414); cardiac dysrhythmias (427); CHF (428); all cardiac (410–414, 427, 428)	24-h avg	5.35	Max: 57
Fung et al. (2005)	Windsor, ON (1995–2000)	CHF (428), IHD (410–414), dysrhythmias (427) and all cardiac	1-h max	27.5 (16.5)	Max: 129
Stieb et al. (2000)	Saint John, NB (1992–1996)	ED visits: angina pectoris, MI, dysrhythmia/conduction disturbance, CHF, all cardiac	24-h avg	6.7 (5.6)	95th: 18 Max: 60

Table 5-28 (Continued): Mean and upper percentile concentrations of sulfur dioxide from cardiovascular related hospital admission and emergency department visit studies: U.S. and Canadian studies from the 2008 ISA for Sulfur Oxides and recent studies.

†Szyszkowicz (2008)	Edmonton, AB (1992–2002)	ED visits: acute ischemic stroke (434 and 436)	24-h avg	2.6	NR
†Szyszkowicz et al. (2012a)	Vancouver, BC (1999–2003)	ED visits (discharge diagnosis): transient ischemic attack, cerebrovascular incident, seizure	24-h avg	2.5	NR
†Szyszkowicz et al. (2012b)	Edmonton, AB (1992–2002)	ED visits: hypertension (401.9)	24 h avg	2.6	Max: 16.3
Villeneuve et al. (2006a)	Edmonton, AB (1992–2002)	ED visits: stroke	24-h avg	All year: 2.6 (1.9)	All year 75th: 4.0

Avg = average; CHF = congestive heart failure; CVD = cardiovascular disease; ED = emergency department; HS = hemorrhagic stroke; ICD = International Classification of Diseases; IHD = ischemic heart disease; ISA = Integrated Science Assessment; max = maximum; MI = myocardial infarction; NR = not reported.

†Studies published since the 2008 ISA for Sulfur Oxides.

The majority of epidemiologic studies reviewed in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) found a positive association between ambient SO₂ concentrations and rates of hospital admission or ED visits for all cardiovascular diseases. One prominent study from the previous ISA was a study conducted in 14 cities across Spain, which observed a 3.5% (95% CI: 0.5, 6.7%) increased risk of hospital admissions for all cardiovascular diseases per 10-ppb increase in SO₂ at lag 0–1 [([Ballester et al., 2006](#)); study details and results for this and other studies in this section are presented in [Table 5-29](#), and [Figure 5-14](#)]. The study authors indicate (results not reported) that the association with SO₂ was attenuated after adjustment for CO or NO₂ in copollutant models. Most studies published since the 2008 ISA for Sulfur Oxides also observed positive associations between SO₂ and ED visits or hospitalizations for all CVD, although only a few considered potential copollutant confounding. For example, a case-crossover study in Beijing found that SO₂ averaged over eight monitoring sites was associated with risk of ED visits for all cardiovascular diseases in a single-pollutant model [OR: 1.04 (95% CI: 1.01, 1.06) per 10-ppb increase in SO₂ on the same day] ([Guo et al., 2009](#)). The association remained comparable in copollutant models adjusting for either PM_{2.5} [OR: 1.03 (95% CI: 0.99, 1.06)] or NO₂ [OR: 1.03 (95% CI: 1.00, 1.07)]. Similarly, in Shanghai, [Chen et al. \(2010b\)](#) reported a small increase in risk of hospital admissions for CVD per 10-ppb increase in 24-h avg SO₂ at lag 5 [1.7% (95% CI: 0.5, 3.0%)] and lag

0–6 [1.3% (5% CI: 0.0, 3.2%)]. The association at lag 5 was similar after adjusting for NO₂ or PM₁₀, while copollutant models for lag 0–6 were not presented.

A number of other studies considering single-pollutant models also reported generally consistent associations between SO₂ concentrations and hospital admissions or ED visits for CVD. A study in New York City ([Ito et al., 2011](#)) observed an association between SO₂ concentrations that was stronger in the warm season [OR: 1.026 (95% CI: 1.021, 1.031) per 10-ppb increase in 24-h avg SO₂] than in the cold season [OR: 1.018 (95% CI: 0.998, 1.049)]. Two studies in São Paulo, Brazil ([Filho et al., 2008](#); [Martins et al., 2006](#)) also found associations in single pollutant models (no quantitative results; results presented graphically). Another study found an increase in the risk of daily hospital admissions per IQR increase in 24-h avg SO₂ in the heavily polluted city of Lanzhou, China ([Zheng et al., 2013](#)). However, this association was less clinically relevant when standardized to a 10-ppb increase in 24-h avg SO₂. In contrast, a large study in Beijing, China reported that CVD ED visits were not associated with SO₂ concentrations on the same day ([Zhang et al., 2015b](#)). The authors also examined a number of other single-day lags and cumulative lags and found little evidence of an association. One important limitation is that these studies did not examine potential copollutant confounding particularly for PM_{2.5} and sulfate.

Table 5-29 Mean and upper percentile concentrations of sulfur dioxide from cardiovascular-related hospital admission and emergency department visit studies.

Study	Location Years	Exposure Assignment	Metric	Mean/Median Concentration ppb	Upper Percentile of Concentrations ppb
†Ito et al. (2011)	New York City, NY (2000–2006)	Average across five monitoring sites	24-h avg	Mean: 7.4	NR
Metzger et al. (2004)	Atlanta, GA (1993–2000)	Fixed-site monitor	1-h max	Median: 11	90th: 39
Moolgavkar (2003)	Los Angeles, CA (1987–1995)	Fixed-site monitor	24-h avg	NR	NR
Schwartz (1997)	Tuscon, AZ (1998–1990)	Fixed-site monitor	24-h avg	Mean: 4.6	75th: 5.9 90th: 10.1
Burnett et al. (1997)	Toronto, Canada (summer 1992–1994)	Average across four to six monitoring sites	1-h max	Mean: 7.9	75th: 11 Max: 26

Table 5-29 (Continued): Mean and upper percentile concentrations of sulfur dioxide from cardiovascular related hospital admission and emergency department visit studies.

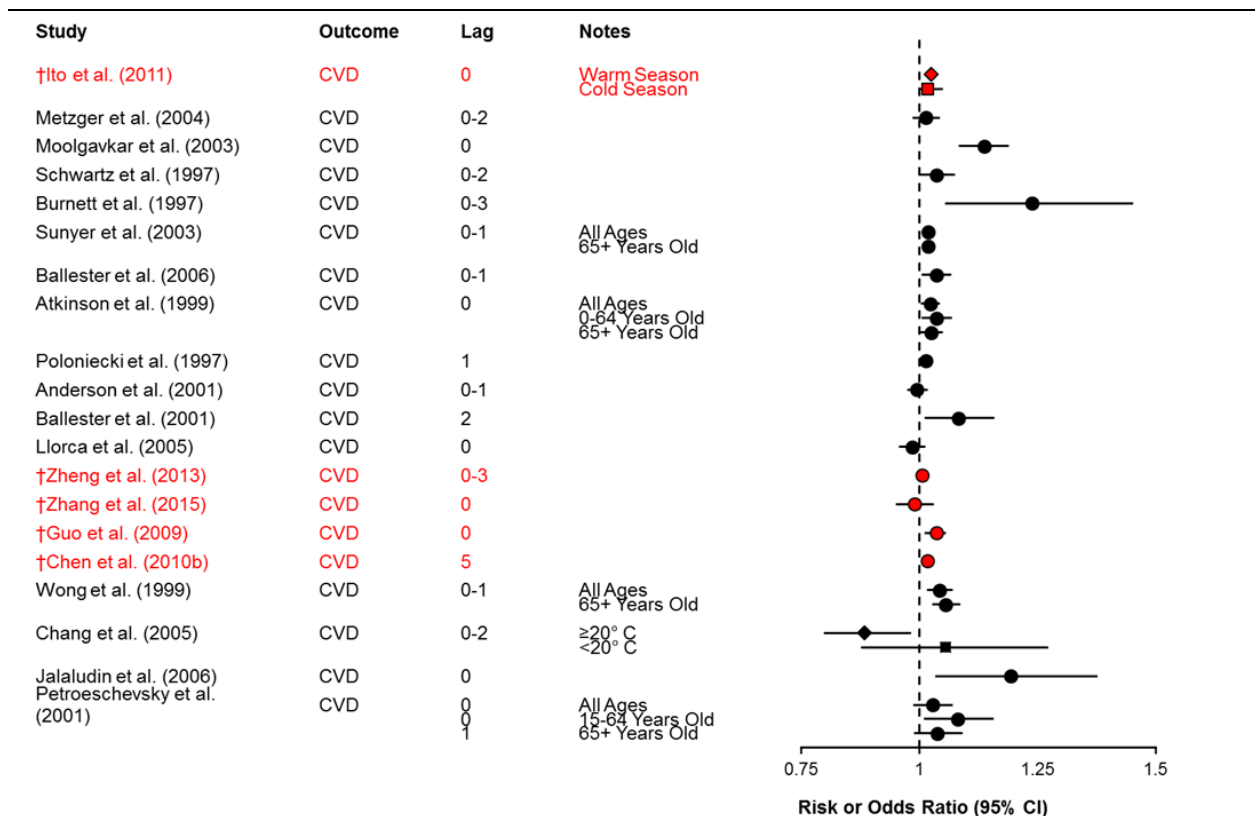
Study	Location Years	Exposure Assignment	Metric	Mean/Median Concentration ppb	Upper Percentile of Concentrations ppb
Sunyer et al. (2003)	Seven European cities (1990–1996)	Fixed-site monitors in each city	24-h avg	Median: 1.9–8.0 across cities	90th: 5.3–29.4 across cities
Ballester et al. (2006)	14 Spanish cities (1995–1999)	Citywide average for each city	24-h avg	Mean: 2.9–15.6 across cities	90th: 4.8–28.8 across cities
Atkinson et al. (1999)	London, England (1992–1994)	Average across five monitoring sites	24-h avg	Mean: 8.1	90th: 11.8 Max: 31.4
Poloniecki et al. (1997)	London, England (1987–1994)	Fixed-site monitor	24-h avg	Median: 6	90th: 21 Max: 114
Anderson et al. (2001)	Birmingham, England (1994–1996)	Average across five monitoring sites	24-h avg	Mean: 7.2	90th: 12.3 Max: 59.8
Ballester et al. (2001)	Valencia, Spain (1994–1996)	Average across 14 monitoring sites	24-h avg	Mean: 9.8	Max: 26.1
Llorca et al. (2005)	Torrelavega, Spain (1992–1995)	Average across three monitoring sites	24-h avg	Mean: 5.1	NR
†Filho et al. (2008)	São Paulo, Brazil (2001–2003)	Average across 13 monitoring sites	24-h avg	Mean: 5.3	Max: 16.4
†Martins et al. (2006)	São Paulo, Brazil (1996–2001)	Average across six monitoring sites	24-h avg	Mean: 6.5	Max: 28.7
†Zheng et al. (2013)	Lanzhou, China (2001–2005)	Average across four monitoring sites	24-h avg	Mean: 30.2	75th: 40.5 Max: 141.6
†Zhang et al. (2015b)	Beijing, China (2009–2011)	Average across 11 monitoring stations	24-h avg	Mean: 10.7	75th: 13.4 Max: 89.5
†Guo et al. (2009)	Beijing, China (2004–2006)	Average across eight monitoring sites	24-h avg	Mean: 18.8	75th: 23.7 Max: 111.8

Table 5-29 (Continued): Mean and upper percentile concentrations of sulfur dioxide from cardiovascular related hospital admission and emergency department visit studies.

Study	Location Years	Exposure Assignment	Metric	Mean/Median Concentration ppb	Upper Percentile of Concentrations ppb
†Chen et al. (2010b)	Shanghai, China (2005–2007)	Average across six monitoring sites	24-h avg	Mean: 21.4	75th: 27.5 Max: 89.7
Wong et al. (1999)	Hong Kong, China (1994–1995)	Average across seven monitoring sites	24-h avg	Median: 6.5	75th: 9.5 Max: 26.1
Chang et al. (2005)	Taipei, Taiwan (1997–2001)	Average across six monitoring sites	24-h avg	Mean: 4.3	75th: 5.5 Max: 14.6
Jalaludin et al. (2006)	Sydney, Australia (1997–2001)	Average across 14 monitoring sites	24-h avg	Mean: 1.07	75th: 1.39 Max: 3.94
Petroeshevsky et al. (2001)	Brisbane, Australia (1987–1994)	Average across two monitoring sites	24-h avg	Mean: 13.9	Max: 49.7

Avg = average; max = maximum; NR = not reported.

†Studies published since the 2008 ISA for Sulfur Oxides.



CI = confidence interval; CVD = cardiovascular disease.

Note: Studies in red are recent studies. Studies in black were included in the 2008 ISA for Sulfur Oxides. All-year associations = circles; summer/warm-days associations = diamonds; winter/cold-days associations = squares. Relative risks are standardized to a 10 or 40-ppb increase in sulfur dioxide for 24-hour avg and 1-hour max metrics, respectively. Lag times are reported in days, unless otherwise noted. Corresponding quantitative results are reported in Supplemental Table 5S-18 (U.S. EPA, 2017c). All results are from single pollutant models.

Figure 5-14 Studies of hospital admissions and emergency department visits for all cardiovascular disease.

Overall, consistent associations between ambient SO₂ concentrations and rates of hospital admissions or ED visits for all cardiovascular diseases have been observed. Although associations are evident in single-pollutant models in many locations, there was limited assessment of potential copollutant confounding. Therefore, this association may at least partly be the result of confounding by correlated pollutants. Additionally, most studies examined 24-h avg exposure metrics for SO₂, which may not adequately capture the spatial and temporal variability in SO₂ concentrations (Section 5.2.1.2).

5.3.1.9 Cardiovascular Mortality

Studies evaluated in the 2008 SO_x ISA that examined the association between short-term SO₂ exposure and cause-specific mortality found consistent positive associations with cardiovascular mortality using a 24-h avg exposure metric. Across studies, there was evidence that the magnitude of the SO₂-cardiovascular mortality relationship was similar or slightly larger than total mortality. Recent multicity studies conducted in Asia ([Chen et al., 2012b](#); [Kan et al., 2010](#)) and Italy ([Bellini et al., 2007](#)), and a meta-analysis of studies conducted in Asia ([Atkinson et al., 2012](#)) provide evidence that is consistent with those studies evaluated in the 2008 SO_x ISA ([Section 5.5.1.3](#), [Figure 5-18](#)). The associations between short-term SO₂ concentrations and cardiovascular mortality are further supported by studies focusing on stroke mortality ([Yang et al., 2014b](#); [Chen et al., 2013](#)). In a study conducted in eight of the CAPES cities, [Chen et al. \(2013\)](#) reported associations for SO₂ and stroke similar to those for all cardiovascular mortality across all of the CAPES cities ([Section 5.5.1.3](#), [Figure 5-18](#)). The magnitude of the association for stroke mortality observed in [Chen et al. \(2013\)](#) is supported by multiple systematic reviews and meta-analyses of stroke mortality ([Shah et al., 2015](#); [Yang et al., 2014b](#)). Both studies reported similar results, with [Yang et al. \(2014b\)](#) reporting a 2.5% increase in stroke mortality (95% CI: 1.8, 3.1) for a 10-ppb increase in 24-h avg SO₂ concentrations in a meta-analysis of mortality studies conducted in Asia, Europe, and North America and [Shah et al. \(2015\)](#) reporting a 2.2% increase in stroke mortality (95% CI: 1.4, 3.1) for a 10-ppb increase in SO₂ concentrations (averaging time was not reported) in a meta-analysis of studies conducted worldwide. However, when interpreting the results of [Yang et al. \(2014b\)](#), it is important to note that when examining regional associations in SO₂-related stroke (i.e., Asia vs. Europe and North America), which combined both mortality and hospital admission outcomes, the magnitude of the association was much smaller, 0.8% (95% CI: -0.2, 1.7), than those observed in studies conducted in Asia, 2.1% (95% CI: 1.2, 3.2). This could be attributed to the relatively low variability and overall low SO₂ concentrations observed in both Europe and North America compared to Asia ([Section 5.5.1.2](#) and [Table 5-39](#)).

Previous studies evaluated in and prior to the 2008 SO_x ISA that examined the association between short-term SO₂ exposures and cardiovascular mortality focused exclusively on single-pollutant analyses. Therefore, questions arose with regard to the independent effect of SO₂ on cardiovascular mortality and whether associations remained robust in copollutant models. A few recent multicity studies conducted in China ([Chen et al., 2012b](#)) and across Asia ([Kan et al., 2010](#)) examined both of these questions. [Chen et al. \(2012b\)](#) found that the SO₂-cardiovascular mortality association was attenuated, but remained positive in copollutant models with PM₁₀ [1.0% (95% CI: 0.08, 1.9) for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 0–1] and NO₂ [0.5% (95% CI:

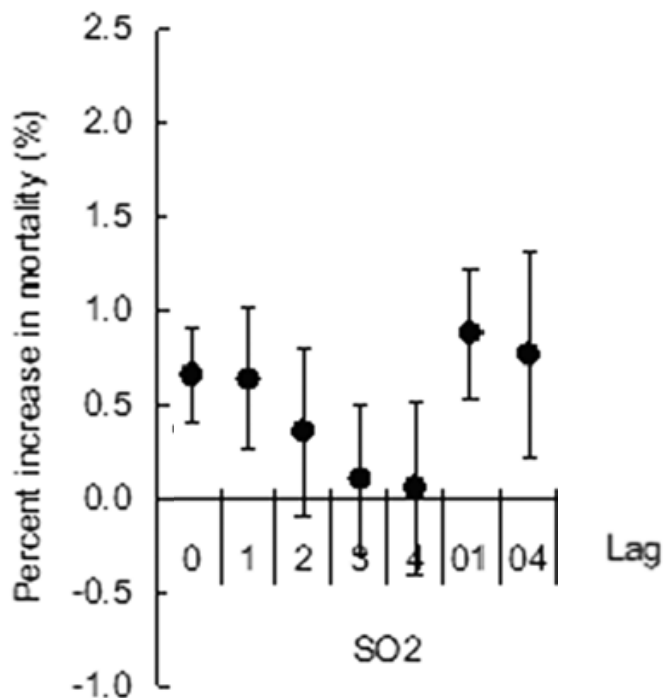
-0.5, 1.4)]. These results are similar to those reported by [Chen et al. \(2012b\)](#) when examining the SO₂-total mortality association in models with NO₂ (i.e., ~80% reduction), but a larger degree of attenuation was observed in models with PM₁₀ for cardiovascular mortality (i.e., ~40% reduction for total mortality and 50% reduction for cardiovascular mortality) ([Section 5.5.1.4](#)). [Kan et al. \(2010\)](#), as part of the PAPA study, also examined potential copollutant confounding (i.e., NO₂, PM₁₀, and O₃) but only in each city individually. The authors found that, although the SO₂-cardiovascular mortality association remained positive in copollutant models, there was evidence of an attenuation of the association in models with PM₁₀ and NO₂ ([Figure 5-19](#)). In an analysis of stroke mortality in eight of the CAPES cities, [Chen et al. \(2013\)](#) reported a pattern of associations similar to that of [Chen et al. \(2012b\)](#) and [Kan et al. \(2010\)](#) in copollutant models with PM₁₀ and NO₂. In single-pollutant models, the authors reported a 2.3% (95% CI: 1.4, 3.2) increase in stroke mortality for a 10 ppb increase in 24-h avg SO₂ concentrations at lag 0-1. However, in copollutant models, [Chen et al. \(2013\)](#) observed that SO₂-stroke mortality associations were attenuated in models with PM₁₀, ~40% reduction [1.9% (95% CI: 0.3, 3.5)] and NO₂, ~80% reduction [0.0% (95% CI: -1.8, 1.9)]. Overall, the studies that examined potential copollutant confounding on the SO₂-cardiovascular mortality relationship report results consistent with what was observed for total mortality. However, the overall assessment of copollutant confounding remains limited, and it is unclear how the results observed in Asia translate to other locations, specifically due to the unique air pollution mixture and higher concentrations observed in Asian cities.

Of the multicity studies evaluated, potential seasonal differences in SO₂-cardiovascular mortality associations were only assessed in a study conducted in Italy ([Bellini et al., 2007](#)) with additional information from U.S.-based single-city studies conducted in Philadelphia ([Sacks et al., 2012](#)) and New York City ([Ito et al., 2011](#)). In a study of 15 Italian cities, [Bellini et al. \(2007\)](#) reported larger SO₂-cardiovascular mortality associations in the summer [9.4% increase (April-September)], compared to both winter [1.6% increase (October-March)] and all-year analyses (2.9% increase), which are consistent with the pattern of associations observed for total and respiratory mortality. These results are supported by [Ito et al. \(2011\)](#) in a study conducted in New York City that found that when examining single-day lags of 0 to 3 days, the SO₂-cardiovascular mortality association was consistently positive during the warm season, ranging from a 1.2 to 3.5% increase across lags. The authors reported no evidence of an association in winter and all-year analyses. Within this analysis, [Ito et al. \(2011\)](#) reported rather poor monitor-to-monitor temporal correlations for SO₂, which would indicate potential exposure error and subsequently attenuation and imprecision in the risk estimate ([Sections 3.4.2, 3.4.4](#)). [Sacks et al. \(2012\)](#) provide additional support to the limited evidence indicating differences in the seasonal pattern of SO₂-cardiovascular mortality

associations. However, as detailed in [Section 5.5.1.4](#), [Sacks et al. \(2012\)](#) demonstrated that across models that use various approaches to control for seasonality and the potential confounding effects of weather, the magnitude of seasonal SO₂-cardiovascular mortality associations may vary depending on the modeling approach employed. Therefore, although [Bellini et al. \(2007\)](#) and [Ito et al. \(2011\)](#) provide initial evidence indicating potentially larger cardiovascular mortality associations in the summer, the results of [Sacks et al. \(2012\)](#) suggest that the evidence remains unclear whether the seasonal pattern of SO₂-cardiovascular mortality associations is consistent across statistical modeling choices and study locations.

An uncertainty that often arises when evaluating studies that examine the relationship between short-term air pollution exposures and cause-specific mortality is whether analyses of statistical modeling parameters, the lag structure of associations, and the C-R relationship provide results that are consistent with what is observed for total mortality. [Chen et al. \(2013\)](#) examined each of these issues in a study of stroke mortality, with additional supporting evidence from the full CAPES ([Chen et al., 2012b](#)). When examining alternative approaches to controlling for seasonality, [Chen et al. \(2013\)](#) found that increasing the df employed from 4 to 10 df per year did not substantially change the SO₂-stroke mortality association. However, when altering the lag structure of the temperature term included to control for the potential confounding effects of weather, [Chen et al. \(2012b\)](#) reported an attenuation of the association, although it did remain positive. As detailed in [Section 5.5.1.4](#), this could be the result of including only one temperature term in the model.

When examining the lag structure of associations, [Chen et al. \(2013\)](#) reported results for stroke mortality that are consistent with those observed for all cardiovascular mortality. As depicted in [Figure 5-15](#), there is evidence of a steady decline in the SO₂-stroke mortality association at longer individual lag days, with the strongest association occurring for a moving average of lag 0–1 day. A similar pattern of associations was observed for cardiovascular mortality by [Chen et al. \(2012b\)](#) in the full CAPES ([Figure 5-20](#)), as well as the PAPA study ([Kan et al., 2010](#)) ([Figure 5-21](#)). These results are further confirmed in a systematic review and meta-analysis of studies of stroke mortality conducted by [Yang et al. \(2014b\)](#), which found the strongest associations at lag 0 and 1 in a subgroup analysis of single-day lags of 0 to 2 days.

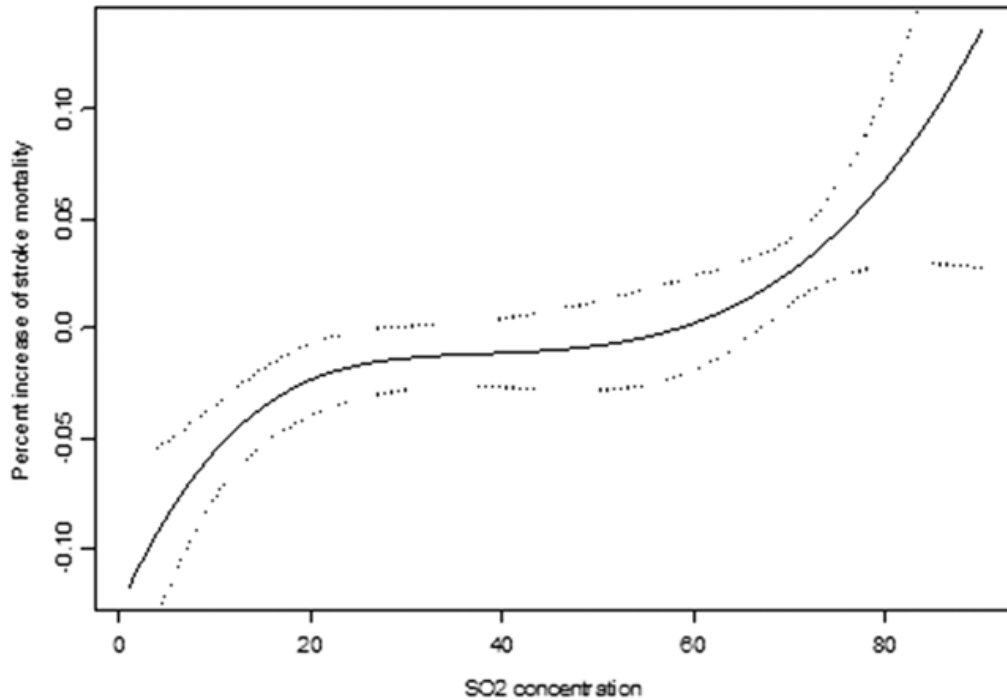


01 = lag 0-1 days; 04 = lag 0-4 days; SO₂ = sulfur dioxide.

Source: Adapted from [Chen et al. \(2013\)](#). Reprinted with permission of WoltersKluwerHealth.

Figure 5-15 Percent increase in stroke mortality associated with a 10 µg/m³ (3.62 ppb) increase in sulfur dioxide concentrations using different lag structures.

[Chen et al. \(2013\)](#) also examined the shape of the SO₂-stroke mortality C-R relationship. To examine the assumption of linearity, the authors fit both a linear and spline model to the SO₂-stroke mortality relationship. [Chen et al. \(2013\)](#) then computed the deviance between the two models to determine any evidence of nonlinearity. An examination of the deviance did not indicate that the spline model improved the overall fit of the SO₂-stroke mortality relationship ([Figure 5-16](#)).



SO₂ = sulfur dioxide.

Note: The solid line represents the mean estimate and the dotted lines are 95% confidence intervals.

Source: Adapted from [Chen et al. \(2013\)](#). Reprinted with permission of Wolters Kluwer Health.

Figure 5-16 Pooled concentration-response curves for sulfur dioxide and daily stroke mortality in eight Chinese cities for a 10 µg/m³ (3.62 ppb) increase in 24-hour avg concentrations at lag 0–1 day.

Overall, recent multicity studies report evidence of consistent positive associations between short-term SO₂ concentrations and cardiovascular mortality, which is consistent with those studies evaluated in the 2008 SO_x ISA. Unlike studies evaluated in the 2008 SO_x ISA, recent studies examined whether copollutants confound the relationship between short-term SO₂ concentrations and cardiovascular mortality. Overall, these studies reported evidence that the SO₂-cardiovascular mortality association was attenuated in models with NO₂ and PM₁₀, but the analyses are limited to Asian cities where the air pollution mixture and concentrations are different from those reported in other areas of the world. A few studies examined potential seasonal patterns in associations, and found initial evidence of larger SO₂-cardiovascular mortality associations in the summer/warm season. However, seasonal associations may be influenced by study location and the statistical modeling choice employed. In examining other uncertainties, a limited number of analyses suggest that: (1) when examining model specification, associations remain robust when alternating the df used to control for

seasonality; (2) when examining lag structure, associations are larger and more precise (i.e., narrower confidence intervals) within the first few days after exposure in the range of 0 and 1 days; and (3) when examining the C-R relationship, there is a log-linear, no threshold C-R relationship. However, for both total and cause-specific mortality, studies have conducted a rather limited exploration of potential alternatives to linearity when examining the shape of the C-R relationship, which in combination with the potential measurement error due to uncharacterized spatial and temporal variability in SO₂ concentrations, complicates the interpretation of the SO₂-mortality C-R relationship ([Sections 3.4.2.2](#) and [3.4.2.3](#)).

5.3.1.10 Subclinical Effects Underlying Cardiovascular Diseases

The following subsections review studies of subclinical effects that serve as useful measures of physiological and biochemical responses that could provide mechanistic evidence to describe a role for SO₂ in the manifestation of cardiovascular diseases. These subclinical effects are not widely validated markers of specific clinical cardiovascular outcomes, but could potentially underlie the development, progression, or indication of various clinical events and provide biological plausibility for multiple outcomes.

Heart Rate and Heart Rate Variability

The 2008 ISA for Sulfur Oxides concluded that the overall evidence available at the time was insufficient to conclude that SO₂ has an effect on cardiac autonomic control as assessed by indices of HRV. HRV provides a noninvasive marker of cardiac autonomic nervous system function. The rhythmic variation in the intervals between heart beats can be quantified in either the time domain or the frequency domain ([TFESC and NASPE, 1996](#)). Common time-domain measures of HRV include the standard deviation of all normal-to-normal intervals (SDNN, an index of total HRV) and the root-mean-square of successive differences (rMSSD, an index influenced mainly by the parasympathetic nervous system). In the frequency domain, HRV is usually divided into the high frequency (HF) and low frequency (LF) components, as well as the ratio of the LF to HF components (LF:HF) ([TFESC and NASPE, 1996](#)). Changes in indices of HRV have been associated with increased risk of cardiovascular events in prospective cohort studies ([Eguchi et al., 2010](#); [La Rovere et al., 2003](#); [Kikuya et al., 2000](#); [Tsuji et al., 1996](#); [Tsuji et al., 1994](#)).

Epidemiology

A number of additional epidemiologic studies are now available for review. In a cross-sectional study in South Korea, [Min et al. \(2009\)](#) reported negative associations between ambient SO₂ concentrations and indices of HRV (SDNN, and the LF and HF components) among 256 smokers, but no association among the 767 nonsmokers (no quantitative results; result presented graphically). In another cross-sectional study, [Min et al. \(2008b\)](#) reported a -7.6% (95% CI: -14.7, 0.1%) change in SDNN and a -23.1% (95% CI: -35.4, -6.5%) change in LF per 10-ppb increase in 24-h avg SO₂ among 1,349 participants in South Korea. The amount of overlapping participants between these two studies is unclear.

The above studies are limited by their cross-sectional approach that compares measures of HRV across individuals assessed on different days. In contrast, longitudinal or repeated-measure study provides an estimate of the average association between SO₂ and measures of HRV within individuals. [Huang et al. \(2012\)](#) measured HRV repeatedly in 40 participants with pre-existing cardiovascular disease in Beijing in the summer of 2007 and again in the summer of 2008, including one measurement period during the 2008 Beijing Olympics when citywide air pollution control measures substantially reduced ambient concentrations of most criteria pollutants. In this study, SO₂ concentrations during the Olympics were reduced by nearly 30% versus the previous month and nearly 50% versus the same period the previous summer ([Huang et al., 2012](#)). Despite these large changes in SO₂ concentrations, overall only small associations were observed between SO₂ concentrations and HRV indices, limited to a 4.8% reduction (95% CI: -9.1, -0.3%) in the LF component and an unexpected 4.1% increase (95% CI: -2.2, 10.9%) in the HF component of HRV per interquartile range (NR) increase in SO₂ in the previous 12 hours ([Huang et al., 2012](#)). In subgroup analyses, SDNN was significantly positively associated with SO₂ concentrations among those with higher levels of C-reactive protein (CRP; a marker of inflammation), those with diabetes, and males. These results are difficult to understand given that a higher SDNN is generally thought to be associated with lower risk of cardiovascular events. The findings were also inconsistent with another study that observed a negative association between SDNN and ambient SO₂ concentrations. A repeated measure study in Shanghai, China reported a 4.36% reduction (95% CI: -5.85, -2.86%) in SDNN per IQR increase (NR) in 4-hour moving average exposure to SO₂ ([Sun et al., 2015](#)). This association was attenuated but still statistically significant in copollutant models adjusting for BC [-2.91% (95% CI: -4.66, -1.13%)] and O₃ [-3.24% (95% CI: -4.83, -1.62%)]. The association was attenuated and no longer statistically significant but still negative in copollutant models adjusting for NO₂ [-0.56% (95% CI: -2.38, 1.30%)] and CO [-1.25% (95% CI: -3.02, 0.55%)]. In another study in Beijing before, during, and after the 2008 Olympics, [Rich et](#)

[al. \(2012\)](#) observed small but statistically significant increases in heart rate associated with ambient SO₂ concentrations on the previous day (no quantitative results; result presented graphically). In expanded results from the same protocol, [Zhang et al. \(2013\)](#) found that the association was similar in copollutants models adjusting for CO, NO₂, O₃, EC, or OC, but was attenuated and no longer positive after adjustment for PM_{2.5} or SO₄²⁻. [Zhang et al. \(2013\)](#) also reported a strong association between LF:HF and ambient SO₂ concentrations on the previous day. This association was relatively unchanged after adjustment for CO, NO₂, O₃, EC, OC, or PM_{2.5} in copollutant models and attenuated but still positive after adjustment for SO₄²⁻. In contrast, a panel study in Taipei, Taiwan used Holter monitors to continuously monitor HRV in 46 participants and observed no associations between ambient SO₂ and SDNN, r-MSSD, LF component, or HF component (quantitative results not reported) ([Chuang et al., 2007](#)). Although new studies are available, findings are mixed and they do not support the presence of an association between ambient SO₂ and measures of HRV.

Experimental Studies

Several experimental studies examined heart rate and HRV following SO₂ exposure. Study characteristics are summarized in Supplemental Table 5S-15 ([U.S. EPA, 2017c](#)). Animal studies have reported no changes in heart rate following SO₂ exposures of 1,000–5,000 ppb in guinea pigs and 1,200 ppb in rats ([Nadziejko et al., 2004](#); [Halinen et al., 2000b](#); [Halinen et al., 2000a](#)).

Controlled human exposure studies evaluated heart rate during and following SO₂ exposure. [Tunnicliffe et al. \(2001\)](#) reported no change in heart rate in healthy adults or adults with asthma during exposure to 200 ppb SO₂ for 1 hour at rest. However, [Routledge et al. \(2006\)](#) found an increase in heart rate measured by a statistically significant decrease in the RR interval from electrocardiographic (ECG) recordings 4 hours after SO₂ exposure in healthy adults. Statistically significant changes in heart rate were not observed in SO₂-exposed older adults with stable angina and coronary artery disease immediately after and 4 hours after exposure. [Tunnicliffe et al. \(2001\)](#) did not obtain ECG measures following exposure and, thus, would not have been able to observe the increase in heart rate reported by [Routledge et al. \(2006\)](#).

[Tunnicliffe et al. \(2001\)](#) and [Routledge et al. \(2006\)](#) reported changes in different measures of HRV in adults either during or shortly after SO₂ exposure. [Tunnicliffe et al. \(2001\)](#) reported that HF power, LF power, and total power were higher with SO₂ exposures compared to air exposure in the healthy subjects, but that these indices were reduced during SO₂ exposure in the subjects with asthma (statistical significance only in total power in healthy adults). The LF:HF ratios were unchanged in both groups. [Routledge et al. \(2006\)](#) reported a reduction in SDNN, rMSSD, percentage of successive

RR interval differences exceeding 50 ms (pNN₅₀), and HF power, the latter of which did not reach statistical significance, in healthy adults 4 hours after SO₂ exposure. Baroreflex sensitivity was also reduced 4 hours after SO₂ exposure determined by changes in α -HF and α -LF. There were no changes in HRV among the patients with coronary heart disease; however, this lack of response may be due to a drug treatment effect because a large portion of these patients were taking beta-blockers. The changes in HRV observed in [Tunnicliffe et al. \(2001\)](#) and [Routledge et al. \(2006\)](#) indicate the potential for SO₂ to affect the autonomic nervous system (see [Section 4.3.1](#)).

Summary of Heart Rate and Heart Rate Variability

The current epidemiologic evidence does not support the presence of an association between ambient SO₂ and measures of HRV. No changes in heart rate were observed in experimental animal studies while changes in HRV observed in human clinical studies may indicate the potential for SO₂ to affect the autonomic nervous system (see [Section 4.3.1](#)). Overall, studies evaluating the effect of ambient SO₂ concentrations and measures of HRV and heart rate remain limited.

QT Interval Duration

The time between the start of the Q wave and the end of the T wave on an ECG (QT interval) provides an electrocardiographic marker of ventricular repolarization. Prolongation of the QT interval is associated with increased risk of life-threatening ventricular arrhythmias. In an analysis of data from the Boston-area Normative Aging Study, [Baja et al. \(2010\)](#) observed a small and imprecise (i.e., wide confidence intervals) association between heart-rate-corrected QT interval and 10-hour moving average of SO₂ concentrations among older, generally white men (no quantitative results; result presented graphically). The only prior study available for comparison from the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) also found that SO₂ concentrations were positively associated with increased QT interval duration amongst a small sample of 56 men in Erfurt, Germany [3.75 ms increase (95% CI: 1.21, 6.28 ms) per 0.61-ppb increase in 24-h avg SO₂] ([Henneberger et al., 2005](#)). There was little variability between daily measured SO₂ concentrations, so the effect estimate is not standardized to prevent inflation of the confidence interval.

The two reviewed studies provide limited evidence of association between short-term SO₂ exposure and markers of ventricular repolarization. Neither of these studies evaluated potential copollutant confounding and coherence for an association between SO₂ exposure and arrhythmias is not provided by experimental studies ([Section 5.3.1.3](#)).

Insulin Resistance

There were no epidemiologic studies of diabetes or insulin deficiency available for the 2008 ISA for Sulfur Oxides. Recent studies reported contrasting findings regarding short-term associations between air pollutants and measures of insulin resistance and fasting glucose, which play key roles in the development of Type 2 diabetes mellitus. In a panel study of older adults in Korea, [Kim and Hong \(2012\)](#) observed 0.94 (95% CI: -0.02, 1.88) and 0.94 (95% CI: 0.01, 1.81) mean increases in the homeostatic model assessment index of insulin resistance [$\text{fasting insulin} \times (\text{fasting glucose} \div 22.5)$] per 10-ppb increase in 24-h avg SO₂ at lags 3 and 4, respectively. There were imprecise (i.e., wide 95% CI) or null associations at all other individual lag days examined, from 0 to 10. Another panel study, conducted in the heavily polluted Tangshan, China, reported an association between 24-h avg SO₂ concentrations and fasting glucose levels ([Chen et al., 2015b](#)). However, this association is unlikely to be clinically relevant when standardized to a 10-ppb increase in 24-h avg SO₂ [0.045 mmol/L (95% CI: 0.039, 0.050 mmol/L) increase at lag 0–3]. Conversely, [Kelishadi et al. \(2009\)](#) reported the lack of an association between 24-h avg SO₂ and insulin resistance in a cross-sectional study of 374 Iranian children aged 10–18 years.

In summary, the available epidemiologic evidence is limited and inconsistent, and does not support the presence of an association between ambient SO₂ concentrations and measures of insulin resistance.

Biomarkers of Cardiovascular Risk

Several epidemiologic and toxicological studies have explored the potential relationship between SO₂ and biomarkers of cardiovascular risk. In particular, markers of inflammation have been evaluated in a number of epidemiologic and toxicological studies published since the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) ([Table 5-30](#)). Relatively few studies have evaluated the potential link between SO₂ and other circulating markers of cardiovascular risk, including markers of coagulation, vascular injury, or lipid oxidation.

Table 5-30 Epidemiologic studies of biomarkers of cardiovascular effects.

Study	Location and Years Sample Size	Mean and Upper Concentration SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
†Dubowsky et al. (2006)	St. Louis, MO Mar–Jun 2002 (n = 44)	24-h avg: 6.7 75th percentile: 7.4 Max: 27	Fixed-site	CRP (percent change) Lag 04: -36.1 (-65.2, -2.8) IL-6 (percent change) Lag 04: -16.5 (-38.7, 6.5) White blood cells (cells/μL) Lag 04: 10.0 (0.4, 19.6)
†Steinvil et al. (2008)	Tel Aviv, Israel 2002–2006 (n = 3,659)	24-h avg: 2.8 75th percentile: 3.5	Citywide avg	CRP (percent change) men; women Lag 0: 0 (-38, 38); -13 (56, 28) Lag 1: -19 (-50, 25); -13 (-63, 38) Lag 2: 6 (-38, 44); -25 (-69, 31) Fibrinogen (mg/dL) men; women Lag 0: -20.0 (-40.0, 0.6); -23.8 (-51.3, 3.8) Lag 1: -21.3 (-42.5, 0.0); -13.1 (-41.3, 14.4) Lag 2: -15.0 (-37.5, 6.9); 17.5 (-11.9, 46.9) WBC (cells/μL) men; women Lag 0: 231 (-419, 875); -169 (-1,000, 656) Lag 1: 44 (-631, 713); -544 (-1,381, 294) Lag 2: -125 (-819, 563); -481 (-1,356, 388)
†Thompson et al. (2010)	Toronto, ON 1999–2003 (n = 45)	24-h avg: 3.57	Fixed-site	No quantitative results; results presented graphically. Increase in IL-6 associated with 4- and 5-day moving avg SO ₂ concentrations. Null association between SO ₂ and fibrinogen Correlations: CO: 0.43, NO ₂ : 0.44, O ₃ : -0.19, PM _{2.5} : 0.45
†Gandhi et al. (2014)	Piscataway, NJ 2005–2009 (n = 49)	24 h avg: 2.4 75th percentile: 3.2 Max: 13.8	Fixed-site	Change in plasma nitrate (nM): Lag 0: 53.6 (-4.5, 111.4) Lag 1: 45.0 (0.9, 90.9) Lag 2: 48.2 (-13.2, 110.0)

Table 5-30 (Continued): Epidemiologic studies of biomarkers of cardiovascular effects.

Study	Location and Years Sample Size	Mean and Upper Concentration SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
†Lee et al. (2011b)	Allegheny County, PA 1997–2001 (n = 1,696)	7-day avg: 8.4 75th percentile: 10.1 Max: 25.4	Citywide avg	No quantitative results presented. “...SO ₂ ... associations (with CRP) were negligible for both the entire population and nonsmokers only.”
†Hildebrandt et al. (2009)	Erfurt, Germany 2001–2002 (n = 38)	24-h avg: 1.35 Max: 14.2	Fixed-site	No quantitative results presented. “No significant associations” between SO ₂ and inflammatory (fibrinogen, E-selectin) or coagulation (D-dimer, prothrombin) markers.
Baccarelli et al. (2007a)	Lombardia, Italy 1995–2005 (n = 1,218)	24-h avg median: 2.4 75th percentile: 4.5 Max: 96.7	Citywide avg	Effect estimates not provided. SO ₂ not correlated with anticoagulation proteins (plasma fibrinogen, functional AT, functional protein C, protein C antigen, functional protein S, or free protein S).
Baccarelli et al. (2007b)	Lombardia, Italy 1995–2005 (n = 1,213)	24-h avg Median: 2.4 75th percentile: 4.5 Max: 96.7	Citywide avg	Homocysteine difference, fasting (percent change). Lag 24 h: 0.2 (–6.3, 6.7) Lag 0–6 days: 0.2 (–4.3, 4.7) Homocysteine difference, post-methionine-load (percent change) Lag 24 h: 2.6 (–3.2, 8.6) Lag 0–6 days: 2.6 (–1.5, 6.7)
Wellenius et al. (2007)	Boston, MA 2002–2003 (n = 28)	24-h avg: 4.8	Citywide avg	No quantitative results presented. “No significant associations were observed between (NO ₂) and B-type natriuretic peptide levels at any of the lags examined.”
†Goldberg et al. (2008)	Montreal, QC 2002–2003 (n = 31)	NR	Fixed-site	Oxygen saturation (mean difference) Lag 0: –0.104 (–0.320, 0.110) Lag 1: –0.277 (–0.497, –0.058) Lag 0–2: –0.210 (–0.536, 0.116)
†Brüske et al. (2011)	Augsburg, Germany 2003–2004 (n = 200)	24-h avg: 1.15 75th percentile: 1.26 Max: 2.4	Fixed-site	No quantitative results; results presented graphically. Inverse associations were observed for SO ₂ with Lp-PLA ₂ at lag days 2 and 3 and positive associations were estimated with Lp-PLA ₂ lag days 4 and 5. Correlations: PNC: 0.77, PM _{2.5} : 0.42, PM ₁₀ : 0.43, CO: 0.63, NO ₂ : 0.51, NO: 0.60, O ₃ : –0.45.

Table 5-30 (Continued): Epidemiologic studies of biomarkers of cardiovascular effects.

Study	Location and Years Sample Size	Mean and Upper Concentration SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
† Zhang et al. (2013)	Beijing, China Jun–Oct, 2008 (n = 125)	24-h avg Before: 7.45 During: 2.97 After: 6.81	Fixed-site	No quantitative results; results presented graphically. Positive association between SO ₂ and fibrinogen (lag 6). Inverse association between SO ₂ and WBC count (lag 5).
† Lin et al. (2015)	Beijing, China 2007–2008 (n = 36 school children)	NR	Monitor located nearby school	<i>Urinary 8-oxodG</i> (Geometric mean ratio by SO ₂ exposure percentile) <30th (<2.1 ppb): referent 30th–60th (2.1–6.4 ppb): 1.26 (0.93, 1.70) 60th–90th (6.4–49.1 ppb): 1.66 (1.15, 2.41) >90th (>49.1 ppb): 2.31 (1.54, 3.46) <i>Urinary Malondialdehyde</i> <30th: referent 30th–60th: 1.21 (1.05, 1.40) 60th–90th: 1.40 (1.15, 1.69) >90th: 1.40 (1.08, 1.83)
† Khafaie et al. (2013)	Pune City, India 2005–2007 (n = 1,392)	24-h avg: 8.3	Citywide avg	No quantitative results; results presented graphically. SO ₂ was associated with increases in CRP at lags 0, 1, 2, 4, 5, 0–7, 0–14, and 0–30.

AT = atascadero; avg = average; CI = confidence interval; CO = carbon monoxide; CRP = C-reactive protein; IL-6 = interleukin-6; Lp-PLA₂ = lipoprotein-associated phospholipase A2; max = maximum; n = sample size; N = population number; NO = nitric oxide; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; PNC = particle number concentration; SO₂ = sulfur dioxide; WBC = white blood cell.

^aEffect estimates are standardized to a 10-ppb or 40-ppb increase in SO₂ concentration for 24-h avg and 1-h max metrics, respectively.

†Studies published since the 2008 ISA for Sulfur Oxides.

Note: All lag times are in days, unless otherwise noted.

Epidemiologic Studies

The epidemiologic data available for review by the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) did not suggest a consistent link between SO₂ and biomarkers of cardiovascular risk, including markers of inflammation and coagulation. Results from more recent studies continue to be inconsistent. [Dubowsky et al. \(2006\)](#) investigated associations between ambient pollutants and markers of systemic inflammation in a panel (repeated-measures) study of 44 seniors in St. Louis, MO and found that higher ambient SO₂ concentrations were associated with lower levels of CRP and white blood cells, but not IL-6 (results for this study, and other studies in this section can be found in [Table 5-30](#)). Similarly, during the Beijing Olympics, SO₂ was inversely associated with white blood cell counts, although positively associated with fibrinogen ([Zhang et al., 2013](#)). The negative associations observed in these two studies are unexpected and difficult to explain. In contrast, among 45 nonsmoking adults, [Thompson et al. \(2010\)](#) found a positive association between SO₂ and IL-6, but not fibrinogen. In another panel study examining pollutant levels before, during, and after the Beijing Olympics, [Lin et al. \(2015\)](#) reported positive associations between SO₂ concentrations and urinary markers of oxidative stress, malondialdehyde and 8-oxodG, in children.

In a cross-sectional analysis of data from a panel study of 49 young adults in New Jersey, [Gandhi et al. \(2014\)](#) observed that plasma nitrite levels, a marker for endothelial dysfunction, were associated with an increase in 24-h avg SO₂ concentrations on the same day. [Khafaie et al. \(2013\)](#) observed a positive association between SO₂ and CRP in a cross-sectional study of patients with Type 2 diabetes in Pune City, India, whereas a study of 1,696 pregnant women ([Lee et al., 2011b](#)), and one of 38 male patients with chronic pulmonary disease ([Hildebrandt et al., 2009](#)) observed null associations between SO₂ and CRP. In a cross-sectional analysis of 3,659 participants in Tel-Aviv, [Steinvil et al. \(2008\)](#) observed inconsistent and/or imprecise (i.e., wide 95% CI) associations between SO₂ and CRP, white blood cells, or fibrinogen among men and women. Observed associations were both positive and negative depending on the length of the lags, making interpretation of the results difficult.

Ambient SO₂ concentrations are reportedly not associated with blood coagulation ([Baccarelli et al., 2007a](#)), plasma homocysteine ([Baccarelli et al., 2007b](#)), markers of vascular injury ([Hildebrandt et al., 2009](#)), or markers of functional status in patients with heart failure ([Wellenius et al., 2007](#)). Conversely, SO₂ concentrations were inversely associated with blood oxygen saturation in patients with heart failure ([Goldberg et al., 2008](#)) and positively associated with lipoprotein-associated phospholipase A2 (Lp-PLA₂) in survivors of myocardial infarction ([Brüske et al., 2011](#)).

Experimental Studies

Experimental studies examined biomarkers of cardiovascular risk following SO₂ exposure, including markers of inflammation, coagulation, and oxidative injury. Study characteristics are summarized in Supplemental Table 5S-15 ([U.S. EPA, 2017c](#)). No changes were reported in serum C-reactive protein or markers of coagulation (fibrinogen, D-dimer, platelet aggregation, blood count, or differential white blood cell count) in healthy humans and patients with stable angina and coronary artery disease exposed to SO₂ ([Routledge et al., 2006](#)). An animal toxicological study examined the hematological effects of short-term SO₂ exposure on blood biomarkers. Acute exposure of rats to 870 ppb SO₂ for 24 hours resulted in increased hematocrit, sulfhemoglobin, and osmotic fragility as well as decreased whole blood and packed cell viscosities ([Baskurt, 1988](#)). These results indicate a systemic effect of inhaled SO₂ and are consistent with an oxidative injury to red blood cells. Furthermore, prolonged exposure to as low as 0.3 ppm SO₂ results in measurable amounts of circulating sulfite and its metabolite S-sulfonate in humans (see [Section 4.3.4](#)). The relationship between circulating sulfite/S-sulfonate and cardiovascular effects of inhaled SO₂ has not yet been explored in human subjects.

Summary of Blood Markers of Cardiovascular Risk

There is inconsistent evidence from epidemiological studies regarding any potential link between SO₂ and other circulating markers of cardiovascular risk. Few experimental studies of markers of inflammation or oxidative stress in humans or animals were available; however, the presence of circulating sulfite/S-sulfonate has been reported. In addition, there is limited evidence for systemic oxidative stress under conditions of prolonged exposure. Overall, evidence from available studies is not coherent between disciplines and does not clearly support an effect of ambient SO₂ concentrations and markers of cardiovascular disease, including inflammation.

5.3.1.11 Summary and Causal Determination

Overall, the available evidence is inadequate to infer a causal relationship between short-term exposure to SO₂ and cardiovascular health effects. Multiple epidemiologic studies report positive associations between short-term ambient SO₂ concentrations and cardiovascular outcomes, such as cardiovascular mortality, myocardial infarction and ischemic heart disease, and aggregated cardiovascular outcomes; however, substantial uncertainties remain regarding exposure measurement error and copollutant confounding. Specifically, the majority of studies reporting positive associations evaluated averaged SO₂ concentrations over multiple monitors and used a 24-h avg exposure metric, which may not adequately capture the spatial and temporal variability in SO₂ concentrations

([Sections 3.4.2.2](#) and [3.4.2.3](#)). Further, among studies adjusting for copollutants, the observed associations are generally attenuated, complicating the determination of an independent SO₂ association. Support for observed epidemiological associations from experimental data is weak, particularly as some experimental evidence for cardiovascular effects relied on high concentrations and prolonged exposures. These experimental studies also provide some evidence to support key events in a proposed mode of action for cardiovascular effects from SO₂ exposure, including changes in the autonomic nervous system and oxidative stress. Overall, these experimental studies provide only limited evidence of biological plausibility and key events in a proposed mode of action; however, important limitations remain, as well as issues of coherence across disciplines.

This determination is consistent with that of the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). The majority of epidemiologic studies reviewed in the 2008 ISA for Sulfur Oxides examined hospital admissions or ED visits for aggregated categories of cardiovascular disease or for mortality from cardiovascular causes. These studies generally reported positive associations in single pollutant models, but analyses designed to assess copollutant confounding were limited. Relatively few studies evaluated specific cardiovascular outcomes such as MI, arrhythmia, cerebrovascular disease, and heart failure, and those that were available did not support an association with short-term SO₂ exposure. Controlled human exposure studies demonstrated the potential for SO₂ exposure to exert an effect on the autonomic nervous system but there was a lack of supporting animal toxicological data. The available animal toxicological studies did not report effects on HR, HRV, arrhythmia, or blood pressure following short-term SO₂ exposures [Table 5S-15 ([U.S. EPA, 2017c](#))]. In addition, limited and inconsistent mechanistic evidence, including evidence pertaining to key events in a proposed mode of action, offered only limited insight for the role of SO₂ in the triggering of cardiovascular diseases. Although multiple recent epidemiologic studies add to the evidence available for the current review, the additional studies do not substantially reduce uncertainties related to copollutant confounding. Moreover, there continues to be a limited experimental evidence to provide biological plausibility to strengthen the inference of causality for SO₂-related cardiovascular effects.

The evidence for cardiovascular effects, with respect to the causal determination for short-term exposure to SO₂ is detailed below using the framework described in the Preamble to the ISAs [([U.S. EPA, 2015b](#)), Tables I and II]. The key evidence, supporting or contradicting, as it relates to the causal framework is summarized in [Table 5-31](#).

Table 5-31 Summary of evidence, which is inadequate to infer a causal relationship between short-term sulfur dioxide exposure and cardiovascular effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Triggering a myocardial infarction			
Although most epidemiologic studies examining MI or all CVD report positive associations, results are generally attenuated after adjustment for copollutant confounding	Increases in hospital admissions and ED visits for IHD, MI, and all CVD in adults in multiple studies, including multicity studies However, a number of studies report associations with ED visits and hospital admissions were attenuated after adjustment with CO, NO ₂ , or PM ₁₀ .	Section 5.3.1.2 Section 5.3.1.8 Supplemental Figures 5S-3, 5S-4, and 5S-5 (U.S. EPA, 2017c)	24-h avg: 1.2–15.6 ppb 24-h avg: 1.9–30.2 ppb
Uncertainty regarding exposure measurement error	Studies examining the association between short-term SO ₂ exposures and cardiovascular effects generally rely on single or the average of multiple monitors. Because SO ₂ generally has low to moderate spatial correlations across urban geographical scales such studies are subject to exposure measurement error.	Section 3.4.2.2	
Uncertainty due to lack of coherence with other lines of evidence	Lack of evidence from epidemiologic panel studies and experimental studies for clinical cardiovascular effects		
Lack of evidence to identify key events in the proposed mode of action	Lack of mechanistic evidence for key events leading to extrapulmonary effects	Section 4.3	
	Limited and inconsistent evidence of increased systemic inflammation in epidemiologic studies	Section 5.3.1.10	

Table 5-31 (Continued): Summary of evidence, which is inadequate to infer a causal relationship between short term sulfur dioxide exposure and cardiovascular effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Other cardiovascular effects			
Inconclusive evidence from epidemiologic, controlled human exposure and toxicological studies	Epidemiologic studies report generally null associations between SO ₂ and risk of cardiac arrest and arrhythmias. One experimental study provides no evidence of arrhythmia.	Section 5.3.1.3	
	Inconsistent epidemiologic evidence for an association between SO ₂ and risk of cerebrovascular disease and stroke, and increased blood pressure and hypertension	Sections 5.3.1.4 and 5.3.1.5	
	Insufficient quantity of studies evaluating decompensation of heart failure and venous thrombosis and pulmonary embolism	Sections 5.3.1.6 and 5.3.1.7	
	Changes in HR and HRV reported in controlled human exposure but coherence with animal toxicological and epidemiologic studies is limited	Tunncliffe et al. (2001) Routledge et al. (2006) Section 5.3.1.10	200 ppb, 1 h at rest (humans)
Some evidence to identify key events in the proposed mode of action	Some evidence for activation of neural reflexes in humans leading to altered HRV	Section 4.3.1 Figure 4-2	200 ppb, 1 h at rest (humans)
	Some evidence of systemic oxidative stress based on measured sulfhemoglobin and mitochondrial changes	(Baskurt, 1988) Section 4.3.4	0.87 ppm, 24 h (rats) 1.34 ppm, 4 h/day for 30 days (rats)

Table 5-31 (Continued): Summary of evidence, which is inadequate to infer a causal relationship between short term sulfur dioxide exposure and cardiovascular effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Cardiovascular mortality			
Consistent epidemiologic evidence but uncertainty regarding SO ₂ independent effect	Multicity studies consistently observe associations with cardiovascular mortality, including stroke with 24-h avg SO ₂ at lags primarily of 0–1 days. Analysis of potential confounding by copollutants primarily limited to PM ₁₀ and NO ₂ reported evidence of attenuation of associations. No studies included copollutant analyses with PM _{2.5} .	Section 5.3.1.9 Chen et al. (2012b) Chen et al. (2013) Kan et al. (2010) Bellini et al. (2007) Atkinson et al. (2012)	24-h avg: 2.5–38.2

Avg = average; CO = carbon monoxide; CVD = cardiovascular disease; ED = emergency department; HR = heart rate; HRV = heart rate variability; IHD = ischemic heart disease; MI = myocardial infarction; NO₂ = nitrogen dioxide; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; SO₂ = sulfur dioxide.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)).

^bDescribes the key evidence and references, supporting or contradicting, that contribute most heavily to causal determination. References to earlier sections indicate where full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated.

Recent epidemiologic studies of specific cardiovascular outcomes add to the overall evidence for the effect of short-term SO₂ exposure on the cardiovascular system with a number of these studies evaluating effects related to triggering an MI ([Section 5.3.1.2](#)). Several recent epidemiologic studies of MI hospitalizations and ED visits consistently report associations in single pollutant models but associations are not always robust in copollutant models indicating that the associations may be due to confounding ([Hsieh et al., 2010](#); [Cheng et al., 2009](#); [Ballester et al., 2006](#)). The small number of studies based on clinical MI data, rather than hospitalizations, report inconsistent evidence regarding associations between ambient SO₂ concentrations and risk of MI ([Milojevic et al., 2014](#); [Turin et al., 2012](#); [Bhaskaran et al., 2011](#)). The only study that examined the association of hourly ambient SO₂ concentrations prior to MI onset reported no association, although there was some evidence of a positive association in a sensitivity analysis of older adults ([Bhaskaran et al., 2011](#)). Although [Chuang et al. \(2008\)](#) reported an association between short-term SO₂ exposure and ST-segment changes, a nonspecific marker of myocardial ischemia, in patients with a history of coronary heart disease that generally remained unchanged after additional control for PM_{2.5} and BC in copollutant models; the evidence overall, was not generally consistent.

Findings from recent studies of the association of short-term exposure to SO₂ with hospital admissions or ED visits for cerebrovascular diseases or stroke are inconsistent and, associations reported from single pollutant models in some locations may be due to confounding by copollutants ([Section 5.3.1.4](#)). Epidemiologic studies evaluating the association between ambient SO₂ concentrations and blood pressure remain inconsistent with most relying on centrally located monitors that do not capture the spatial variability of SO₂ and few examining the potential for copollutant confounding ([Section 5.3.1.5](#)). Although a small number of studies were conducted to examine the association of short-term SO₂ exposure with other clinical outcomes, including heart failure ([Section 5.3.1.7](#)) and VTE ([Section 5.3.1.6](#)), findings from these studies do not support an effect of short-term exposure to SO₂. There is also a lack of epidemiologic evidence supporting an effect of short-term SO₂ exposure on arrhythmia ([Section 5.3.1.3](#)), although associations between short-term SO₂ exposure and markers of ventricular repolarization abnormalities that are risk factors for arrhythmia have been observed ([Baja et al., 2010](#); [Henneberger et al., 2005](#)) ([Section 5.3.1.10](#)).

Consistently positive associations have been reported in epidemiologic studies of short-term SO₂ exposure and cardiovascular mortality ([Section 5.3.1.9](#)). These include studies reviewed in the 2008 ISA for Sulfur Oxides and recent multicity studies that generally report an association similar or slightly larger in magnitude for cardiovascular mortality compared to total mortality. Studies that report results from copollutants models generally report attenuation of the association between short-term SO₂ exposure and cardiovascular mortality after adjustment for PM₁₀ and NO₂.

Few experimental studies have evaluated the effects of SO₂ exposure on the cardiovascular system. The strongest evidence comes from controlled human exposure studies, for which copollutant confounding is not a concern, that short-term exposure to SO₂ can affect the autonomic nervous system of healthy adults and adults with asthma ([Routledge et al., 2006](#); [Tunnicliffe et al., 2001](#)) ([Section 5.3.1.10](#)). These studies report changes in HR and HRV following SO₂ exposure in adults, which is indicative of potential cardiovascular effects being mediated by the neural reflex pathway; however, these changes were not reported in animal studies, nor did epidemiological evidence support the presence of associations, particularly after adjusting for copollutant confounding ([Section 5.3.1.10](#)). Animal studies also provide limited evidence for the role of systemic oxidative stress as a key event in a proposed mode action ([Section 4.3](#)). In particular, studies observed changes in sulfhemoglobin ([Baskurt, 1988](#)). In addition, studies of long-term exposure found lipid peroxidation in the brain ([Qin et al., 2012](#)), and mitochondrial changes in the heart and brain ([Qin et al., 2016](#); [Qin et al., 2012](#)) ([Section 4.3.4](#)). While evidence of the diffusion of sulfite into the circulation and tissues following exposure to SO₂ has been reported and could play a role in the induction of

systemic effects, this evidence is limited by the lack of studies directly investigating the relationship between circulated sulfite and cardiovascular effects ([Section 4.3.4](#)). Furthermore, changes observed in animal studies were not reported in human experimental data, or epidemiological studies, and the available data relied on prolonged exposure periods and higher concentrations (0.87–1.34 ppm). Overall, these experimental studies provide limited evidence of biological plausibility and key events in a proposed mode of action; however, there remain important issues, such as limited coherence between disciplines, as well as limitations in the data based on exposure concentration and duration.

Despite numerous additional epidemiologic studies reporting positive associations between short-term SO₂ exposure and cardiovascular effects, a key uncertainty that remains since the 2008 ISA for Sulfur Oxides is the potential for confounding by other pollutants, specifically those from a common source that are highly correlated with SO₂. The majority of hospital admission or ED visit studies have not evaluated whether the reported associations with SO₂ are robust to adjustment for other pollutants. Those studies that do examine associations with SO₂ adjusted for PM [Figure 5S-3 and Table 5S-19 ([U.S. EPA, 2017c](#))], NO₂ [Figure 5S-4 and Table 5S-20 ([U.S. EPA, 2017c](#))], or other correlated pollutants [Figure 5S-5 and Table 5S-21 ([U.S. EPA, 2017c](#))] report that, in general, associations were either attenuated or no longer present after controlling for potential copollutant confounding ([Hsieh et al., 2010](#); [Cheng et al., 2009](#); [Ballester et al., 2006](#)). A limited number of studies examined copollutant confounding on the SO₂-cardiovascular mortality relationship, which included analyses on stroke mortality, and provided evidence that the SO₂ association was reduced in copollutant models with NO₂ and PM₁₀ ([Chen et al., 2013](#); [Chen et al., 2012b](#); [Kan et al., 2010](#)). Finally, while copollutant models are a common statistical tool used to evaluate the potential for copollutant confounding, their interpretation can be limited ([Section 5.1.2](#)). Without consistent and reproducible experimental evidence that is coherent with the effects observed in epidemiologic studies, uncertainty still exists concerning the role of correlated pollutants in the associations observed with SO₂. Thus, uncertainty remains regarding the extent to which SO₂ exposure is independently associated with cardiovascular outcomes or if SO₂ is a marker for the effects of another correlated pollutant or mix of pollutants.

In conclusion, the evidence overall is inadequate to infer a causal relationship between short-term SO₂ exposure and cardiovascular health effects. This conclusion does not represent a change from the conclusion of the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). Multiple epidemiologic studies report positive associations between short-term ambient SO₂ concentrations and cardiovascular outcomes, but these associations are generally attenuated after adjustment for copollutants. There is limited experimental

evidence in humans or animals evaluating exposure to SO₂ and the results of these studies only provide limited coherence for the positive associations observed in the epidemiologic studies. Further, while there is some evidence of key events in a proposed mode of action, important limitations remain in the available evidence regarding the biological plausibility of effects observed in epidemiological studies. The evidence from epidemiologic and experimental studies is of insufficient consistency and continues to have limited coherence, and thus, is inadequate to infer between short-term SO₂ exposure and cardiovascular effects.

5.3.2 Long-Term Exposure

5.3.2.1 Introduction

Studies of the effects of long-term exposure to SO₂ on the cardiovascular system were not available for inclusion in the 1982 AQCD ([U.S. EPA, 1982a](#)). The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) reviewed a limited body of toxicological and epidemiologic studies published through 2006 and concluded that the available evidence was “too limited to make any conclusions” between the effects of long-term exposure to SO₂ and cardiovascular health. These recent studies do not change the conclusion from the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)).

The 2008 ISA for Sulfur Oxides included one epidemiologic study, which reported an increased risk of cardiovascular events in association with long-term exposure to SO₂ in post-menopausal women (50–79 years old) without previous CVD from 36 U.S. metropolitan areas. In this study, [Miller et al. \(2007\)](#) found that PM_{2.5} was most strongly associated with cardiovascular events [MI, revascularization, angina, congestive heart failure (CHF), CHD death], compared to the other pollutants evaluated [HR: 1.24 (95% CI: 1.04, 1.48) per 10 µg/m³], followed by SO₂ [1.07 (95% CI: 0.95, 1.20) per 5 ppb]. Exposures to air pollution were estimated by assigning the annual (for the year 2000) mean air pollutant concentration measured at the monitor nearest to the subject’s five-digit residential ZIP code centroid. The effect estimate for SO₂ was strengthened in a multipollutant model that was adjusted for several other pollutants including PM_{2.5}. However, correlations among pollutants were not described and exposure measurement error may have introduced a bias ([Section 3.4.4.2](#)). Consequently, the extent to which this study supports an independent effect of SO₂ on the cardiovascular system is limited. Several recent epidemiologic studies of the association of long term SO₂ exposure with subclinical and clinical cardiovascular outcomes add to the available body of evidence.

Experimental animal studies with long-term exposures below 2,000 ppb were not available for inclusion in the 2008 ISA for Sulfur Oxides. Although a small number of studies using exposures above 2,000 ppb were included, they did not contribute substantially to conclusions because the concentrations of SO₂ used in these studies were unlikely to be relevant to ambient concentrations of SO₂. No new toxicological studies in humans or animals have been published since the 2008 ISA for Sulfur Oxides. Overall, the biological plausibility and independence of the SO₂ effect observed in epidemiologic studies remains an important uncertainty.

This section reviews the published studies of the cardiovascular effects of long-term exposure to SO₂ (i.e., longer than 1 month). To clearly characterize the evidence underlying causality, the discussion of the evidence is organized into groups of related outcomes [ischemic heart disease and myocardial infarction ([Section 5.3.2.2](#)), cerebrovascular disease and stroke ([Section 5.3.2.3](#)), hypertension ([Section 5.3.2.4](#)), other cardiovascular effects ([Section 5.3.2.5](#)), and cardiovascular mortality ([Section 5.3.2.6](#))]. Evidence for subclinical effects (e.g., blood biomarkers of cardiovascular effects) of long-term exposure to SO₂ are discussed in [Section 5.3.2.7](#) and serve to inform biological plausibility across multiple clinical cardiovascular events and outcomes.

Similar to [Section 5.3.1](#), studies examining cardiovascular effects of sulfite exposure (via intraperitoneal injection, intravenous injection, etc.) are not included in this section because these studies generally involve exposures to sulfite that are higher than what is expected to occur following inhalation of SO₂ at ambient relevant concentrations. Studies in humans and animals suggest that prolonged exposure to SO₂ may result in measurable changes in the concentrations of sulfite in plasma and tissues, but these changes would be expected to be far less following concentrations of SO₂ typically found in ambient air. The literature describing the distribution and metabolism of sulfite is discussed in [Sections 4.2.3](#) and [4.2.4](#). The potential role of sulfite in the induction of systemic effects, such as effects of the cardiovascular system, is discussed in [Section 4.2.4](#).

5.3.2.2 Myocardial Infarction and Ischemic Heart Disease

IHD generally develops due to a buildup of plaques in the arterial walls (i.e., atherosclerosis) that impede the blood flow and oxygen delivery to the heart. This restricted oxygen delivery or ischemia from excess plaque, plaque rupture, and clot formation can lead to an MI. Several epidemiologic studies provide evidence of a relationship between long-term exposure to SO₂ and ischemic heart disease and incident or fatal MI ([Table 5-32](#)). However, uncertainty remains regarding the influence of exposure measurement error on the effect estimates observed in epidemiologic studies

([Section 3.4.4.2](#)) and the ability of these studies to distinguish the independent effect of long-term SO₂ exposure from the effect of correlated copollutant exposures ([Section 3.4.3](#)).

Table 5-32 Epidemiologic studies of the association of long-term exposure to sulfur dioxide with cardiovascular disease.

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI) ^a
†Lipsett et al. (2011)	California Teachers Study Cohort N = 124,614 (n = 43 MI events) California Jun 1996– Dec 2005	SO ₂ IQR: 0.43 mean: 1.72	Geocoded residential address linked to pollutant surface developed using IDW (fixed-site monitors concentrations from 1995–2005 used to model exposure as a time-dependent function) Correlation of SO ₂ with: ozone, <i>r</i> = -0.17 PM _{2.5} , <i>r</i> = 0.02 PM ₁₀ , <i>r</i> = 0.54 NO ₂ , <i>r</i> = 0.67 CO, <i>r</i> = 0.80	MI incidence SO ₂ : HR 1.97 (0.07, 60) Covariates: age, race, smoking, second-hand smoke, BMI, lifetime physical activity, nutritional factors, alcohol, marital status, menopausal status, hormone replacement therapy, hypertension medication and aspirin, and family history of MI/stroke Copollutant adjustment: none
†Atkinson et al. (2013)	National GP Patient Cohort N = 810,686 (n = 13,956 MI events) England 2003	IQR: 0.83 mean (SD): 1.47	Annual average SO ₂ concentration for 2002 at a 1 by 1-km resolution derived from dispersion models and linked to residential post codes Correlation of SO ₂ with: NO ₂ , <i>r</i> = 0.86	MI incidence HR: 1.34 (1.13, 1.50) Covariates: age, sex, smoking, BMI, diabetes, hypertension, and index of multiple deprivation Copollutant adjustment: none

Table 5-32 (Continued): Epidemiologic studies of the association of long term exposure to sulfur dioxide with cardiovascular disease.

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI) ^a
†Rosenlund et al. (2006)	SHEEP cohort n = 1,397 cases and 1,870 controls Stockholm, Sweden 1992–1994	Cases median: 9.6 5th–95th: 2.6–18.2 Controls median: 9.3 5th–95th: 7.7–17.5	Dispersion models to estimate SO ₂ from heating at residential address. Residential history available for 30 yr exposure estimate. Correlation of 30 yr SO ₂ with: 30 yr NO ₂ , <i>r</i> = 0.73 30 yr CO, <i>r</i> = 0.49	First MI OR: 0.99 (0.9, 1.1) per 5 ppb Covariate adjustment: age, sex, hospital catchment area, smoking, diabetes, physical inactivity, and SES Copollutant adjustment: none
†Ancona et al. (2015)	Rome, Italy 2001–2010 N = 85559 (n = 491 men and 356 women)	2.5 µg/m ³ SO _x SD: 0.9	Lagrangian particle dispersion model (SPRAY Ver. 5) used SO _x as exposure marker for petrochemical refinery emissions PM ₁₀ : 0.81 H ₂ S: 0.78	IHD ^b HR men: 0.87 (0.74, 1.02) HR women: 0.83 (0.64, 1.07) CVD ^b HR men: 1.01 (0.93, 1.0) HR women: 1.02 (0.92, 1.12)
Miller et al. (2007)	WHI Cohort U.S. 1994–1998	NR	Annual avg (2000): nearest monitor to residence ZIP code centroid	Cardiovascular events (MI, revascularization, angina, CHF, CHD death) HR: 1.07 (0.95, 1.20) Covariates: age, ethnicity, education, household income, smoking, diabetes, hypertension, systolic blood pressure, BMI, and hypercholesterolemia HR: 1.13 (0.98, 1.30) after simultaneous adjustment for PM _{2.5} , PM _{10–2.5} , CO, NO ₂ , and O ₃

Table 5-32 (Continued): Epidemiologic studies of the association of long term exposure to sulfur dioxide with cardiovascular disease.

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI) ^a
† Qin et al. (2015) † Dong et al. (2013a)	N = 24,845 Random selection (18–74 yr) from households in 33 communities in 11 districts of northeastern China	Mean: 20.3 Median: 18 IQR: 7.5	3-yr avg (2006–2008) SO ₂ concentration for each district NO ₂ , <i>r</i> = 0.38 O ₃ , <i>r</i> = 0.87 PM ₁₀ , <i>r</i> = 0.70	CVD OR: 1.08 (0.93, 1.26) Note: associations stronger among males OR BMI <25 kg/m ² 1.11 (0.97, 1.27) OR BMI ≥25 kg/m ² 1.12 (0.99, 1.25) Covariate adjustment: age, race, education, income, smoking, drinking, exercise, diet, sugar, family history of CVD or stroke, and district Copollutant adjustment: none

Avg = average; BMI = body mass index; CHD = coronary heart disease; CHF = congestive heart failure; CI = confidence interval; CO = carbon monoxide; CVD = cardiovascular disease; GP = general practice; H₂S = hydrogen sulfide; HR = heart rate; IDW = inverse distance weighting; IHD = ischemic heart disease; IQR = interquartile range; MI = myocardial infarction; n = sample size; N = population number; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OR = odds ratio; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; PM_{10-2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm and greater than a nominal 2.5 μm; *r* = correlation coefficient; SD = standard deviation; SES = socioeconomic status; SHEEP = Stockholm Heart Epidemiology Programme; SO₂ = sulfur dioxide; SO_x = sulfur oxides; WHI = Women's Health Initiative.

^aEffect estimates are standardized per 5-ppb increase in SO₂ concentrations.

^bEffect estimate per 2.88 μg/m³ increase in SO_x concentration (as reported by author in original publication).

†Studies published since the 2008 ISA for Sulfur Oxides.

[Lipsett et al. \(2011\)](#) analyzed the association of incident MI with long-term exposure to SO₂, other gases (NO₂, CO, O₃), and PM. These authors studied a cohort of California public school teachers aged 20–80 years old (N = 124,614). Each participant's geocoded residential address was linked to pollutant surfaces that were determined by inverse distance weighted interpolation of pollutant concentrations measured at fixed-site monitors during the period 1996–2005. The average of monthly SO₂ concentrations was modeled as a time-dependent function for subjects with at least 12 months of exposure. Those living outside the radial range for which the monitor was intended to provide representative data were excluded from the analysis. This “representative range” was 3 km for neighborhood SO₂ monitors and 5 km for the urban/regional SO₂. There was no evidence of an association between SO₂ and incident MI [HR 1.97 (95% CI: 0.07, 60) per 5 ppb].

[Atkinson et al. \(2013\)](#) examined the association of incident cardiovascular disease with SO₂. These authors studied patients (aged 40–89 years) registered with 205 general

practices across England. The authors report that approximately 98% of the population is registered with a general practitioner minimizing the potential for selective participation. Predicted annual average SO₂ concentrations within 1-km × 1-km grids, estimated using dispersion models, were assigned to participants based on their residential postal code. Cardiovascular disease outcomes included in the analysis were MI, stroke, arrhythmias, and heart failure. Authors reported an association of SO₂ with MI in a fully adjusted model [HR: 1.34 (95% CI: 1.13, 1.50) per 5 ppb]. The performance of the dispersion model used to estimate SO₂ concentration was characterized as moderate to poor depending on the study year. Failure of the model to capture the spatial variability of SO₂ could lead to bias toward or away from the null ([Section 3.4.4.2](#)).

[Rosenlund et al. \(2006\)](#) conducted a population case-control study to examine the association of first MI with long-term exposure to air pollution in Stockholm, Sweden. In this study, residential histories were used to estimate 30-yr avg SO₂ concentration from residential heating sources using dispersion models. Although a positive association of SO₂ and other pollutants (NO₂, CO, PM₁₀) with fatal MI was observed in this study, no association between nonfatal MI and long-term SO₂ exposure was reported. [Panasevich et al. \(2013\)](#) reported higher TNF- α levels among those with a genetic polymorphism of a TNF- α gene (*TNF308G/A*) as well as an increased risk of MI in the same population ([Section 5.3.2.5](#)).

Weak or inverse associations of both cardiovascular and ischemic heart disease were reported in a study relying on a Lagrangian particle dispersion model (see [Section 3.3.2.4](#)) to estimate SO_x emissions (gaseous and particulate component) from a refinery ([Ancona et al., 2015](#)). Exposure model performance statistics were not reported. Null associations of cardiovascular hospitalizations with PM₁₀, which was highly correlated with SO_x ($r = 0.81$) in this study, were observed. Because SO_x was used as a marker for refinery emissions, which contains multiple toxics including VOCs, the study was not designed to evaluate the independent effect of SO₂. In addition to the study by [Miller et al. \(2007\)](#), which was included in the previous review, two analyses examined the association of long-term SO₂ exposure with a relatively broadly defined outcome that included several cardiovascular diseases ([Qin et al., 2015](#); [Dong et al., 2013a](#)). These studies, which were conducted among Chinese adults, reported imprecise increases in the risk of cardiovascular disease and results suggest the potential for age and body weight to modify the association with long-term SO₂ exposure. Neither of these analyses adjusted for copollutant confounding despite the potentially high correlation of SO₂ with PM_{2.5} (correlations reported with PM₁₀ were 0.70). Further, the district-level SO₂ concentrations used to indicate exposure may not have adequately captured the spatial variability of long-term SO₂ exposure.

Overall, these epidemiologic data do not provide support for an association of long-term SO₂ exposure with IHD or more broadly defined categories of cardiovascular disease. There is uncertainty related the independent effect of SO₂ on the cardiovascular system. Further, the exposure assessment techniques applied in the studies were subject to varying degrees of error depending on the method. Uncharacterized spatial variability in the exposure estimate has the potential to bias the health effect estimate ([Section 3.4.2](#)).

5.3.2.3 Cerebrovascular Diseases and Stroke

[Lipsett et al. \(2011\)](#) evaluated the association of incident stroke with long-term exposure to SO₂, other gases (NO₂, NO_x, CO, ozone), and PM ([Table 5-33](#)). The authors found no evidence of an association between SO₂ and incident stroke [HR: 6.21 (95% CI: 0.4, 88)]. A positive association of SO₂ with incident stroke of 1.13 (95% CI: 1.00, 1.34) per 5 ppb was reported by [Atkinson et al. \(2013\)](#) in patients across England (study methods in [Section 5.3.2.2](#)).

Two analyses of a random selection of adults (N = 24,845) ranging from 18 to 74 years old from households in 33 Chinese communities examined the association between long-term SO₂ exposure and stroke. Monitor concentrations within each district were used to derive 3-yr avg concentrations that were assigned to participants. The mean concentration among study participants was 20 ppb. [Dong et al. \(2013a\)](#) reported an increased risk of stroke [OR: 1.09 (1.01, 1.18) per 5 ppb] with the strongest associations in males. [Qin et al. \(2015\)](#) further evaluated effect modification by obesity and reported an increased risk of stroke among participants with body mass index (BMI) greater or equal to 25 kg/m² [OR: 1.18 (1.05, 1.32) per 5 ppb]. Neither of these studies considered copollutant confounding despite moderate to high ($r = 0.38$ to 0.87) correlations with the other pollutants that were evaluated (PM₁₀, NO₂, or ozone). The district level SO₂ concentrations may not have adequately captured the spatial variability of SO₂.

Table 5-33 Epidemiologic studies of the association of long-term exposure to sulfur dioxide with stroke.

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI)
†Lipsett et al. (2011)	California Teachers Study Cohort N = 124,614 (n = 56 stroke events) California Jun 1996– Dec 2005	SO ₂ IQR: 0.43 mean: 1.72	Geocoded residential address linked to pollutant surface developed using IDW (fixed-site monitors concentrations from 1995–2005 used to model exposure as a time-dependent function) Correlation of SO ₂ with: ozone, $r = -0.17$ PM _{2.5} , $r = 0.02$ PM ₁₀ , $r = 0.54$ NO ₂ , $r = 0.67$ CO, $r = 0.80$	Stroke incidence SO ₂ : HR 6.21 (0.4, 88) Covariates: age, race, smoking, second-hand smoke, BMI, lifetime physical activity, nutritional factors, alcohol, marital status, menopausal status, hormone replacement therapy, hypertension medication and aspirin, and family history of MI/stroke Copollutant adjustment: none
†Atkinson et al. (2013)	National GP Patient Cohort England 2003 N = 836,557 (n = 13,956 stroke events)	IQR: 0.83 mean (SD): 1.47	Annual average SO ₂ concentration for 2002 at a 1 by 1-km resolution derived from dispersion models and linked to residential post codes Correlation of SO ₂ with: NO ₂ , $r = 0.86$	Stroke incidence HR: 1.13 (1.00, 1.34) Covariates: age, sex, smoking, BMI, diabetes, hypertension, and index of multiple deprivation Copollutant adjustment: none

Table 5-33 (Continued): Epidemiologic studies of the association of long term exposure to sulfur dioxide with stroke

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI)
†Dong et al. (2013a)	N = 24,845 Random selection (18–74 yr) from households in 33 communities in 11 districts of northeastern China	Mean: 20 median: 18 IQR: 7.5	3-yr avg (2006–2008) SO ₂ concentration for each district NO ₂ , <i>r</i> = 0.38 O ₃ , <i>r</i> = 0.87 PM ₁₀ , <i>r</i> = 0.70	Prevalent stroke OR: 1.09 (1.01, 1.18) Note: associations stronger among males Covariate adjustment: age, sex, educational level, occupation, family income, BMI, hypertension, family history of stroke, family history of CVD, smoking status, drinking, diet, and exercise
†Qin et al. (2015)	N = 24,845 Random selection (18–74 yr) from households in 33 communities in 11 districts of northeastern China	Mean: 20.3 IQR: 7.5	3-yr avg (2006–2008) SO ₂ concentration for each district NO ₂ , <i>r</i> = 0.38 O ₃ , <i>r</i> = 0.87 PM ₁₀ , <i>r</i> = 0.70	Stroke BMI <25 kg/m ² : OR: 1.03 (0.92, 1.14) BMI 25 kg/m ² : OR: 1.18 (1.05, 1.32) Sex-stratified analyses also presented Covariate adjustment: age, race, education, income, smoking, drinking, exercise, diet, sugar, family history of CVD or stroke, and district

Table 5-33 (Continued): Epidemiologic studies of the association of long term exposure to sulfur dioxide with stroke

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI)
†Johnson et al. (2010)	Edmonton, Alberta Canada Jan 2003– Dec 2007	SO ₂ mean: 1.3	IDW average monitor SO ₂ concentration assigned at postal code centroid level Correlation of 5-yr avg SO ₂ with: NO ₂ , $r = 0.40$ O ₃ , $r = 0.41$ CO, $r = -0.19$	Ecological analysis of stroke incidence rates: Stroke ED visits Q1 RR: 1.0 (reference) Q2 RR: 0.91 (0.83, 1.00) Q3 RR: 0.89 (0.81, 0.98) Q4 RR: 0.84 (0.73, 0.96) Q5 RR: 0.93 (0.89, 0.98) ^a Results for HS, non-HS, and TIA also presented Covariate adjustment: age, sex, and household income Copollutant adjustment: none

Avg = average; BMI = body mass index; CI = confidence interval; CO = carbon monoxide; CVD = cardiovascular disease; ED = emergency department; GP = general practice; HR = heart rate; HS = hemorrhagic stroke; IDW = inverse distance weighting; IQR = interquartile range; MI = myocardial infarction; n = sample size; N = population number; NO₂ = nitrogen dioxide; non-HS = nonhemorrhagic e; O₃ = ozone; OR = odds ratio; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; Q1 = 1st quartile; Q2 = 2nd quartile; Q3 = 3rd quartile; Q4 = 4th quartile; Q5 = 5th quartile r = correlation coefficient; RR = relative risk; SD = standard deviation; SO₂ = sulfur dioxide; TIA = transient ischemic attack.

^aEffect estimates are standardized per 5-ppb increase in SO₂ concentrations.

†Studies published since the 2008 ISA for Sulfur Oxides.

An inverse association between SO₂ concentration and stroke incidence was observed in an ecological analysis of long-term exposure to ambient pollution conducted in Edmonton ([Johnson et al., 2010](#)), while an association of SO₂ with stroke prevalence was observed in a study of 33 Chinese communities [OR: 1.21 (95% CI 1.01, 1.46)] ([Dong et al., 2013a](#)).

In summary, the epidemiologic studies do not provide evidence to support an effect of long-term SO₂ exposure on stroke morbidity. Findings are not generally consistent across studies and there are uncertainties related to the potential for exposure measurement error and confounding by copollutants.

5.3.2.4 Blood Pressure and Hypertension

Several cross-sectional analyses conducted in China where the mean long-term SO₂ concentration is 18.7 ppb report positive associations with hypertension and increased blood pressure. [Dong et al. \(2013d\)](#) observed an association with hypertension [OR: 1.17 (95% CI: 1.06, 1.28) per 5-ppb increase in SO₂ concentration] among adults greater than 55 years of age in 33 Chinese communities. The absolute change in diastolic and systolic

blood pressure in the study population overall was 0.46 mm Hg (95% CI: 0.15, 0.75) and 1.18 mm Hg (95% CI: 0.68, 1.69) per 5-ppb increase in SO₂ concentration, respectively. [Zhao et al. \(2013\)](#) reported a greater effect of SO₂ on blood pressure among the overweight and obese in this population. A similar trend was also observed with other pollutants (i.e., ozone and NO₂). In a study of children 5–17 years old from elementary schools in seven Chinese cities, [Dong et al. \(2014\)](#) reported associations with arterial blood pressure hypertension in males [OR: 1.17 (95% CI 1.08, 1.27)] and females [OR 1.19 (95% CI 1.10, 1.28)] per 5-ppb increase in 4-yr avg SO₂ concentration. In an extended analysis of this cohort, [Dong et al. \(2015\)](#) reported large associations for SO₂ concentration in overweight and obese children. Although an array of risk factors was considered in the analysis as potential confounders ([Table 5-34](#)), no adjustment for copollutants was presented nor were copollutant correlations reported.

Table 5-34 Epidemiologic studies of the association of long-term exposure to sulfur dioxide with hypertension.

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI)
† Dong et al. (2013d) Cross-sectional	N = 24,845 Random selection (18–74 yr) from households in 33 communities in 11 districts of northeastern China	Mean: 20.3 IQR: 7.5	3-yr avg (2006–2008) SO ₂ concentration for each district NO ₂ , <i>r</i> = 0.38 O ₃ , <i>r</i> = 0.87 PM ₁₀ , <i>r</i> = 0.70	OR: 1.07 (1.03, 1.12) SBP: 0.21 mm Hg (0.07, 0.34) DBP: 0.53 mm Hg (0.31, 0.76) Covariate adjustment: age, race, education, income, smoking, drinking, exercise, diet, sugar, family history of hypertension, and district
† Zhao et al. (2013) Cross-sectional	N = 24,845 Random selection (18–74 yr) from households in 33 communities in 11 districts of northeastern China	Mean: 20.3 IQR: 7.5	3-yr avg (2006–2008) SO ₂ concentration for each district NO ₂ , <i>r</i> = 0.38 O ₃ , <i>r</i> = 0.87 PM ₁₀ , <i>r</i> = 0.70	OR normal: 1.03 (0.99–1.08) OR overweight: 1.10 (1.05–1.15) OR obese: 1.10 (0.99–1.23) Covariate adjustment: race, education, income, smoking, drinking, exercise, diet, sugar, family history of hypertension, and district

Table 5-34 (Continued): Epidemiologic studies of the association of long term exposure to sulfur dioxide with hypertension

Study	Cohort, Location, and Study Period	Mean ppb	Exposure Assessment	Effect Estimates (95% CI)
† Dong et al. (2014) † Dong et al. (2015) Cross-sectional	N = 9,354 Children (5–17 yr) Seven cities northeastern China 2012–2013	Mean: 18.7. IQR: 8.8	4-yr avg concentration for one fixed-site monitor within 1 km of participant's home Correlations NR	Hypertension in males: OR 1.17(1.08, 1.27) Hypertension in females: OR 1.19 (1.10, 1.28) per 5 ppb DPB (all children) 0.43 (0.26, 0.61) SBP (all children) 0.71 (0.50, 0.91) Normal weight: 0.89 (0.83, 0.96) Overweight: 1.36 (1.18, 1.56) Obese: 1.66 (1.46, 1.89) per 5 ppb Covariate adjustment: age, sex, BMI, parental education, LBW, premature birth, income, passive smoking exposure, home coal use, exercise time, area residence per person, family history of hypertension, and district

Avg = average; BMI = body mass index; CI = confidence interval; DPB = diastolic blood pressure; IQR = interquartile range; LBW = low birth weight; n = sample size; N = population number; NO₂ = nitrogen dioxide; NR = not reported; O₃ = ozone; OR = odds ratio; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; r = correlation coefficient; SBP = systolic blood pressure; SO₂ = sulfur dioxide.

†Studies published since the 2008 ISA for Sulfur Oxides.

5.3.2.5 Other Cardiovascular Effects

Few studies have evaluated other cardiovascular effects associated with long-term SO₂ concentrations. [Atkinson et al. \(2013\)](#) examined the association of arrhythmias and heart failure with long-term SO₂ exposure. Study methods are described in [Section 5.3.2.2](#). Authors reported a positive association of SO₂ with heart failure [HR: 1.27 (95% CI: 1.06–1.59) per 5 ppb] and with arrhythmia [HR: 1.13 (95% CI 1.00, 1.27)] in the fully adjusted model. No association of annual SO₂ concentration with hospital admissions for heart failure was reported in a study of county-level air pollution indicator concentrations ([Bennett et al., 2014](#)).

5.3.2.6 Cardiovascular Mortality

The recent evidence for associations between long-term SO₂ exposure and total mortality ([Section 5.5.2](#)) is generally consistent with the evidence in the 2008 ISA for Sulfur Oxides. Several studies report associations between long-term SO₂ exposure and cardiovascular mortality ([Figure 5-27](#)); however, there is no consistent trend toward positive associations for cardiopulmonary or cardiovascular causes of death overall. Additionally, confounding by copollutants is not ruled out ([Section 3.4.3](#)) and uncertainties remain regarding the influence of exposure measurement error ([Section 3.4.2](#)). Together, these uncertainties limit the interpretation of the causal nature of the associations observed in the available epidemiologic studies of long-term mortality.

5.3.2.7 Subclinical Effects Underlying Cardiovascular Diseases

Carotid intima-media thickness (cIMT) is a measurement of thickness of the inner layers of the wall of the artery and can be used to indicate the presence of subclinical atherosclerosis. Other markers of preclinical atherosclerosis include pulse wave velocity and augmentation index, both of which indicate arterial stiffening. In an analysis of the Atherosclerosis Risk in Young Adults study, which is a prospective cohort study ([Lenters et al., 2010](#)), no association of SO₂ concentration with cIMT was observed; however, there was a weak increase in aortic pulse wave velocity reported. SO₂ concentration at the home address for the year 2000 was assigned to participants of this study. The correlations of SO₂ with NO₂, black smoke, and PM_{2.5} reported in this study were low, ranging from $r = 0.09$ to 0.12 . The correlations of SO₂ with metrics of traffic intensity were also low ($r = -0.06$ to 0.06). In another study, [Weng et al. \(2015\)](#) reported that annual average SO₂ concentration was correlated with brachial-ankle pulse wave velocity in univariate analyses but not after adjustment for PM₁₀ and other potential confounders. This study was based on data from 127 heart disease patients undergoing hemodialysis in Taoyuan, Taiwan.

Inflammation and oxidative stress have been shown to play a role in the progression of chronic cardiovascular disease. [Forbes et al. \(2009b\)](#) examined the association of predicted annual average SO₂ concentration with CRP and fibrinogen among the English population. Multilevel linear regression models were used to determine pooled estimates across three cross-sectional surveys conducted during different years. Each participant's postal code of residence was linked to predicted annual average SO₂ concentration derived from dispersion models. SO₂ was not associated with increased CRP or fibrinogen in these data. A study conducted among men and women (45–70 years) in

Stockholm reported an association of 30-yr avg source-specific heating-related SO₂ concentration, estimated using dispersion models, with increases in IL-6; however, SO₂ was not associated with CRP, TNF- α , fibrinogen, or plasminogen activator inhibitor-1 in this study ([Panasevich et al., 2009](#)). Associations between long-term NO₂ concentration, which were moderately correlated with SO₂ ($r = 0.53$), and increased plasma IL-6 were also observed in this study. A study conducted among older adults in Taiwan reported no changes in blood pressure, total cholesterol, fasting glucose, hemoglobin A1c, IL-6, and neutrophils in association with increasing SO₂ concentration while associations between these endpoints and other pollutants were observed ([Chuang et al., 2011](#)).

Overall, the body of evidence is limited and there is no consistent positive trend in the associations observed between SO₂ and subclinical atherosclerosis or circulating markers of inflammation. These findings are consistent with the general lack of mechanistic evidence for key events in the proposed mode of action leading to extrapulmonary effects.

5.3.2.8 Summary and Causal Determination

Overall, the evidence is inadequate to infer a causal relationship between long-term exposure to SO₂ and cardiovascular health effects.

Although a number of epidemiologic studies report positive associations between long-term exposure to SO₂ concentrations and cardiovascular disease and stroke ([Section 5.3.2.3](#)), the evidence for any one outcome is limited and inconsistent. As discussed in [Section 3.4.2.2](#), fixed-site monitors may not capture the spatial variability in SO₂ concentration. Dispersion models generally capture SO₂ variability on near-source spatial scales (up to tens of km) but exposure estimates from such models are subject to other uncertainties ([Section 3.3.2.4](#)). Bias stemming from exposure measurement error can be either toward or away from the null. No studies corrected for such error, complicating the interpretation of findings from studies of long-term exposure of SO₂ ([Section 3.4.4.2](#)). There is also uncertainty regarding the potential for copollutant confounding ([Section 3.4.3](#)). Primary pollutants such as NO₂ and CO typically show moderate to high correlations with SO₂ ([Tables 5-32](#), [5-33](#), and [5-34](#)), and there is a lack of experimental evidence to provide coherence or biological plausibility for an independent effect of SO₂ on cardiovascular health. Several epidemiologic studies evaluated the association between SO₂ concentration and subclinical atherosclerosis or circulating markers of inflammation; however, there is no consistent positive trend in the associations observed between SO₂ and these potential key events in a mode of action.

The available evidence examining the relationship between long-term exposure to SO₂ and cardiovascular effects was evaluated using the framework described in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)). The key evidence, supporting or contradicting, as it relates to the causal framework is summarized in [Table 5-35](#). **In conclusion, the evidence lacks coherence and is of insufficient consistency, and thus, is inadequate to infer a causal relationship between long-term exposure to SO₂ and cardiovascular health effects.**

Table 5-35 Summary of evidence, which is inadequate to infer a causal relationship between long-term sulfur dioxide exposure and cardiovascular effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Some epidemiologic studies report positive associations but results are not generally consistent.	Positive associations of SO ₂ with MI, CVD events, or stroke events	†Lipsett et al. (2011)	1.72 ppb (mean)
		†Atkinson et al. (2013)	1.47 ppb (mean)
		Miller et al. (2007)	NR
	Null/inverse associations observed with MI and stroke	Rosenlund et al. (2006)	9.6 ppb (median)
†Johnson et al. (2010)		1.3 ppb (mean)	
Limited coherence with evidence for cardiovascular mortality	No consistent positive trend observed in long term studies of cardiovascular mortality.	Section 5.3.2.4	
Uncertainty due to confounding by correlated pollutants	Correlations of SO ₂ with CO and NO ₂ vary by location but are generally moderate to high.	Section 3.4.3 Tables 5-32, 5-33, 5-34	

Table 5-35 (Continued): Summary of evidence, which is inadequate to infer a causal relationship between long term sulfur dioxide exposure and cardiovascular effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Uncertainty due to exposure measurement error	Centrally located monitors may not capture spatial variability of SO ₂ concentrations.	Miller et al. (2007) Section 3.4.2	
	SO ₂ estimates from dispersion models in specific studies show poor to moderate agreement with measured concentrations.	† Atkinson et al. (2013) † Forbes et al. (2009a)	
	Exposure measurement error can introduce bias toward or away from the null in studies of long-term exposure	Section 3.4.4.2	
Uncertainty due to lack of coherence with other lines of evidence	Lack of experimental human or animal studies evaluating cardiovascular effects of long-term SO ₂ exposure		
Weak evidence to identify key events in the mode of action	Limited of mechanistic evidence for key events leading to extrapulmonary effects. Limited and inconsistent evidence of increased subclinical atherosclerosis and systemic inflammation (e.g., IL-6, CRP) in epidemiologic studies.	Sections 4.3, 5.3.2.7	

CO = carbon monoxide; CRP = C-reactive protein; CVD = cardiovascular disease; IL-6 = interleukin-6; MI = myocardial infarction; NO₂ = nitrogen dioxide; NR = not reported; SO₂ = sulfur dioxide.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)).

^bDescribes the key evidence and references, supporting or contradicting, that contribute most heavily to causal determination. References to earlier sections indicate where full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated.

†Studies published since the 2008 ISA for Sulfur Oxides.

5.4 Reproductive and Developmental Effects

5.4.1 Introduction

This section covers studies of health endpoints with exposures to SO₂ occurring during or around pregnancy and/or the first years of life. This includes not only pregnancy and

birth outcomes (including infant mortality) occurring close in time to the exposure, but also developmental outcomes potentially occurring years later. Exposures occurring in pregnancy and early life may alter development, and have effects not immediately identifiable but evident at later points. These studies are characterized in this section as they contribute to the weight of evidence for effects of SO₂ on reproductive health and development. Evidence regarding fertility, reproduction, and pregnancy are discussed in [Section 5.4.2](#), with a series of birth outcomes [fetal growth ([Section 5.4.3.1](#)), preterm birth (PTB) ([Section 5.4.3.2](#)), birth weight ([Section 5.4.3.3](#)), birth defects ([Section 5.4.3.4](#)), fetal mortality ([Section 5.4.3.5](#)), and infant mortality ([Section 5.4.3.6](#))] discussed in [Section 5.4.3](#). Studies of developmental outcomes are discussed in [Section 5.4.4](#), with a focus on respiratory developmental outcomes in [Section 5.4.4.1](#).

The 2008 SO_x ISA ([U.S. EPA, 2008d](#)) concluded the evidence was inadequate to infer a causal relationship with reproductive and developmental effects. Epidemiologic studies included in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)) examined impacts on reproductive outcomes including preterm birth, birth weight, intra-uterine growth retardation, birth defects, infant mortality, and neonatal respiratory hospitalizations. While positive associations were observed in the previous SO_x ISA ([U.S. EPA, 2008d](#)), there was little biologic plausibility for these associations provided by supporting toxicological literature. Interpretation of those results was also limited by the lack of control for potential confounding by copollutants, the small number of studies, and uncertainty regarding exposure. Overall, the number of studies examining associations between exposure to ambient SO₂ and reproductive and developmental outcomes has increased substantially since publication of the 2008 SO_x ISA, yet evidence for an association with individual outcomes remains relatively limited and key uncertainties have not been reduced.

The body of literature characterizing the reproductive health effects of exposure to SO₂ has grown considerably since the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), with over 50 recent epidemiologic studies. However, the number of studies for any particular outcome remains relatively limited. Among the recent epidemiologic studies, birth outcomes (e.g., small for gestational age, preterm birth, and birth weight) predominate. Several new studies of congenital anomalies are now available in addition to the single study included in the 2008 SO_x ISA. Recent studies of other outcomes, such as fetal mortality, infant mortality, fertility, and conditions related to pregnancy have also been published. Key epidemiologic studies are summarized in [Table 5-36](#). In toxicological research, a single study published at relevant exposure levels (1,500 ppb or lower) investigated reproductive and developmental changes in exposed female rats and their offspring, finding altered estrus cyclicity with fewer cycles over time, altered birth outcomes of increased litter size, and decreased postnatal body weight in offspring whose dams were exposed to SO₂. This study is summarized in [Table 5-37](#). The majority of the remaining

animal toxicological evidence for reproductive and development effects is for exposure at 5,000 ppb or greater, doses which are beyond the scope of this document.

Table 5-36 Key reproductive and developmental epidemiologic studies for sulfur dioxide.

Study	Location (Sample Size) Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
Fetal growth				
Liu et al. (2003)	Vancouver (n = 229,085) 1986–1998	4.9	Monitors at census subdivision level	IUGR (those with birth weight fall below the 10th percentile, by sex and gestational week, of all singleton live births in Canada between 1986 and 1998, term) M1: 1.07 (1.01, 1.13) Last mo: 1.00 (0.94, 1.06) T1: 1.07 (1.00, 1.14) T2: 0.98 (0.91, 1.04) T3: 1.03 (0.96, 1.10) Robust to NO ₂ , CO, and O ₃ in copollutant models
Brauer et al. (2008)	Vancouver (n = 70,249) 1999–2002	2.2	Inverse distance weighting of three closest monitors within 50 km, 14 SO ₂ monitors	SGA (those with birth weights below the 10th percentile of the cohort, stratified by sex, for each week of gestation) EP: 1.02 (1.00, 1.03) No copollutant models ^b
Rich et al. (2009)	New Jersey (n = 178) 1999–2003	T1: 5.7 T2: 5.6 T3: 5.5	Nearest monitor (within 10 km)	VSGA (growth ratio <0.75) T1: 1.00 (0.92, 1.08) T2: 1.04 (0.96, 1.13) T3: 1.05 (0.97, 1.14) No copollutant models ^b for SO ₂

Table 5-36 (Continued): Key reproductive and developmental epidemiologic studies for sulfur dioxide

Study	Location (Sample Size) Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
†Le et al. (2012)	Detroit, MI (n = 112,609) 1990–2001	5.8	Nearest monitor (ZIP code within 4 km of one of three monitors)	SGA (infants whose birth weights fell below the 10th percentile by sex and gestational week, based on study population's distribution, term) T1, adjusted for CO, NO ₂ , and PM ₁₀ Q1: reference Q2: 1.18 (0.92, 1.51) Q3: 1.01 (0.83, 1.23) Q4: 1.05 (0.87, 1.28) T2, adjusted for CO, NO ₂ , and PM ₁₀ Q1: reference Q2: 1.30 (1.01, 1.69) Q3: 1.12 (0.91, 1.37) Q4: 1.11 (0.90, 1.36) T3, adjusted for CO, NO ₂ , and PM ₁₀ Q1: reference Q2: 1.17 (0.94, 1.45) Q3: 1.24 (1.02, 1.50) Q4: 1.31 (1.06, 1.60) No copollutant models ^b
Preterm birth				
Liu et al. (2003)	Vancouver, BC (n = 229,085) 1986–1998	4.9	Monitors at census subdivision level	M1: 0.95 (0.88, 1.03) Last mo: 1.09 (1.01, 1.19) Robust to NO ₂ , CO, and O ₃ in copollutant models
Sagiv et al. (2005)	Pennsylvania (n = 187,997) 1997–2001	7.9	Monitors at county level	Last 6 wk: 1.05 (1.00, 1.10) 3-day lag: 1.02 (0.99, 1.05) No copollutant models ^b
†Zhao et al. (2011)	Guangzhou, China (n = 7,836 preterm births) 2007	20	City average from monitors	Same day: 1.04 (1.02, 1.06) 1-day lag: 1.01 (0.99, 1.04) 2-day lag: 1.02 (0.99, 1.04) 3-day lag: 1.02 (0.99, 1.04) Robust to NO ₂ and PM ₁₀ in copollutant models
†Mendola et al. (2016a)	U.S. (n = 223,502) 2002–2008	3.99	CMAQ-monitor fused surface with inverse distance weighting, spatial modeling scales not provided; delivery hospital referral region	Week 34 Asthma: 1.32 (1.05, 1.70) No asthma: 1.02 (0.90, 1.14) Week 35 Asthma: 1.17 (1.02, 1.34) No asthma: 0.98 (0.92, 1.05) Last 6 wk of pregnancy Asthma: 0.90 (0.81, 1.00) No asthma: 0.81 (0.77, 0.92) EP Asthma: 0.93 (0.83, 1.03) No asthma: 0.92 (0.87, 0.97) No copollutant models ^b

Table 5-36 (Continued): Key reproductive and developmental epidemiologic studies for sulfur dioxide

Study	Location (Sample Size) Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
Low birth weight				
Ha et al. (2001)	Seoul, South Korea (n = 276,763) 1996–1997	T1: 13 T3: 12	Monitors averaged to city	T1: 1.05 (1.02, 1.08) T1, adjusted for T3: 1.06 (0.98, 1.13) T3: 0.96 (0.92, 0.99) T3, adjusted for T1: 1.02 (0.94, 1.10) No copollutant models ^b
Lee et al. (2003)	Seoul, South Korea (n = 388,105) 1996–1998	12.1	Monitors averaged to city	EP: 1.02 (0.99, 1.05) T1: 1.05 (1.02, 1.09) T2: 0.97 (0.92, 1.00) T3: 1.12 (1.03, 1.20) No copollutant models ^b
Liu et al. (2003)	Vancouver, BC (n = 229,085) 1986–1998	4.9	Monitors at census subdivision level	M1: 1.11 (1.01, 1.22) Last mo: 0.98 (0.89, 1.08) Robust to NO ₂ , CO, and O ₃ in copollutant models
Dugandzic et al. (2006)	Nova Scotia (n = 74,284) 1988–2000	10	Nearest monitor (postcode within 25 km)	T1: 1.20 (1.05, 1.38) T2: 0.99 (0.91, 1.09) T3: 0.95 (0.86, 1.04) No copollutant models ^b
†Morello-Frosch et al. (2010)	California (n = 3,545,177) 1996–2006	2.1	Nearest monitor (census block centroid within 3, 5, or 10 km)	EP 3 km: 1.10 (0.95, 1.34) 5 km: 1.05 (0.95, 1.16) 10 km: 1.05 (1.00, 1.10) No copollutant models ^b
†Ebisu and Bell (2012)	Northeastern and mid-Atlantic U.S. (n = 1,207,800) 2000–2007	6.1	County average from monitors	EP: 1.05 (1.01, 1.09) No copollutant models ^b
†Kumar (2012)	Chicago, IL (n = 398,120) 2000–2004	Nearest monitor: 4.7 County average: 4.6	Nearest monitor (census tract within 3 miles) County average from monitors	Nearest monitor EP: 1.19 (0.90, 1.57) County average EP: 1.05 (0.91, 1.20) No copollutant models ^b

Table 5-36 (Continued): Key reproductive and developmental epidemiologic studies for sulfur dioxide

Study	Location (Sample Size) Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
Birth Weight				
† Darrow et al. (2011) Distributed lag, 1-h max SO ₂	Atlanta, GA (n = 400,556) 1994–2004	M1: 10.7 T3: 9.5	Population weighted spatial model based on monitors, five-county area, 1-h max	M1: 0.625 (–2.625, 3.875) T3: –6.500 (–12.500, –0.667) Non-Hispanic white T3: –8.667 (–15.333, –2.000) Non-Hispanic black T3: –3.167 (–9.833, 3.667) Hispanic T3: –9.5 (–19.000, –0.167) No copollutant models ^b
† Geer et al. (2012)	Texas (n = 1,548,904) 1998–2004	2.3	County average from monitors	EP: –15.594 (–25.344, –5.844) Robust to NO ₂ and O ₃ in copollutant models
Fetal and infant mortality				
† Hwang et al. (2011)	Taiwan (n = 9,325 cases) 2001–2007	5.7	Inverse distance weighting of monitors to township or district, 72 monitors	Among preterm deliveries EP: 1.16 (1.00, 1.34) M1: 1.22 (1.00, 1.34) M2: 1.22 (1.00, 1.34) M3: 1.16 (1.00, 1.34) Among term deliveries EP: 0.95 (0.82, 1.10) M1: 1.00 (0.90, 1.16) M2: 1.00 (0.90, 1.16) M3: 0.95 (0.86, 1.16) No copollutant models ^b
† Faiz et al. (2012)	New Jersey (n = 994) 1998–2004	5.9	Nearest monitor (within 10 km, 1 of 16 monitors)	EP: 1.32 (0.95, 1.84) T1: 1.23 (1.02, 1.51) T2: 1.21 (0.89, 1.53) T3: 1.47 (1.05, 1.69) No copollutant models ^b
† Faiz et al. (2013)	New Jersey (n = 1,277) 1998–2004	5.8	Nearest monitor (within 10 km, 1 of 16 monitors)	2-day lag 1.12 (1.02, 1.24) Adjusted PM _{2.5} : 1.18 (1.00, 1.40) Adjusted NO ₂ : 1.15 (1.00, 1.32) Adjusted CO: 1.05 (0.93, 1.20) No copollutant models ^b

Table 5-36 (Continued): Key reproductive and developmental epidemiologic studies for sulfur dioxide

Study	Location (Sample Size) Years	Mean SO ₂ ppb	Exposure Assessment	Selected Effect Estimates ^a (95% CI)
Woodruff et al. (2008)	U.S. (n = 6,639 cases) 1999–2002	3 (median)	Monitors, averaged to county Exposures for 2 mo after birth	All causes 0.93 (0.84, 1.04) Respiratory 1.09 (0.89, 1.36) Adjusted PM ₁₀ , CO, O ₃ : 1.13 (0.79, 1.60) Adjusted PM _{2.5} , CO, O ₃ : 1.21 (0.79, 1.84) No copollutant models ^b
Developmental				
Dales et al. (2006)	Atlanta, GA (n = 8,586 cases) 1986–2000	4.3	Monitors, averaged to city	Neonatal hospitalization for respiratory disease 2-day lag 2.59 (1.05, 4.39) Adjusted for O ₃ , NO ₂ , CO 1.95 (0.54, 3.68) Adjusted for O ₃ , NO ₂ , CO, PM ₁₀ 1.57 (0.25, 3.29) No copollutant models ^b
†Clark et al. (2010)	British Columbia (n = 3,482 cases) 1999–2000	2	Inverse distance weighting 3 nearest monitors (of 14) within 50 km	Asthma EP: 1.45 (1.28, 1.84) 1st year of life: 1.45 (1.28, 1.84) No copollutant models ^b

CI = confidence interval; CMAQ = Community Multiscale Air Quality; CO = carbon monoxide; EP = entire pregnancy; IUGR = intra-uterine growth restriction; max = maximum; M1 = Month 1; M2 = Month 2; M3 = Month 3; n = sample size; N = population number NO₂ = nitrogen dioxide; O₃ = ozone; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; Q1 = 1st quartile; Q2 = 2nd quartile; Q3 = 3rd quartile; Q4 = 4th quartile; SGA = small for gestational age; SO₂ = sulfur dioxide; T1 = 1st trimester; T2 = 2nd trimester; T3 = 3rd trimester; VSGA = very small for gestational age.

^aRelative risk per 5-ppb change in SO₂, unless otherwise noted.

^bStudy did not include two-pollutant model results (i.e., SO₂ and one other air pollutant), but may have included multipollutant model results (i.e., model that includes 3 or more air pollutants).

†Studies published since the 2008 ISA for Sulfur Oxides.

Table 5-37 Study specific details from animal toxicological studies of the reproductive and developmental effects of sulfur dioxide.

Study and Species	Concentration SO ₂ Exposure	Measured Outcome(s)
Mamatsashvili (1970b) Rat	0.057 or 1.5 ppm for 72 days	Estrus cyclicity duration (F0 and F1), litter size, offspring growth (body weight)

SO₂ = sulfur dioxide.

Several recent articles have reviewed methodological issues relating to the study of outdoor air pollution and adverse birth outcomes ([Chen et al., 2010a](#); [Woodruff et al., 2009](#); [Ritz and Wilhelm, 2008](#); [Slama et al., 2008](#)). Some of the key challenges to interpretation of birth outcome study results include: (1) the difficulty in assessing exposure as most studies use existing monitoring networks to estimate individual exposure to ambient air pollution; (2) the need for detailed exposure data and potential residential movement of mothers during pregnancy; (3) the inability to control for potential confounders such as other risk factors that affect birth outcomes (e.g., smoking), evaluating the exposure window (e.g., trimester) of importance; and (4) the limited evidence on the physiological modes of action for these effects ([Ritz and Wilhelm, 2008](#); [Slama et al., 2008](#)). An additional limitation is the failure for many studies of reproductive and developmental outcomes to adjust for co-occurring air pollutants. As ozone, PM_{2.5}, and NO_x have all been associated with reproductive and developmental health outcomes, the lack of adjustment makes interpretation of isolated SO₂ effects more difficult.

5.4.2 Fertility, Reproduction, and Pregnancy

Infertility affects approximately 11% of all women ages 15–44 in the U.S. ([Chandra et al., 2013](#)), and can have negative psychological impacts and affect quality of life; infertility and subfertility may also potentially signal poorer physiological health. Those with fertility problems are at higher risk for adverse pregnancy and birth outcomes if they do become pregnant ([Hansen et al., 2005](#); [Helmerhorst et al., 2004](#); [Jackson et al., 2004](#)). Outcomes studied in this area include fecundity (the ability to conceive frequently, quantified as length of time to pregnancy) and fertility (the ability to have a live birth). Studies in this area frequently use populations undergoing assisted reproductive treatment, as these populations have a large amount of data collected on them during

treatment and defined menstrual cycles and start points. In cohorts recruited from the general population, exact timing can be difficult to determine due to reliance on participant recall, particularly if they are surveyed well after initiation of pregnancy attempts. Many pregnancies are unplanned, which also adds a level of complication to quantifying fertility. Researchers may also investigate potential mechanistic links between pregnancy conditions and biomarkers and later birth outcomes; such as pregnancy-related hypertension, which is a leading cause of perinatal and maternal mortality and morbidity ([Lee et al., 2012](#)).

Four recent studies have examined the effects of SO₂ on measures of fertility; all use different populations and outcomes and observed mainly null effects for SO₂ exposures. Recent studies examined semen quality parameters in cohorts of men from Chongqing, China ([Zhou et al., 2014](#)) and Poland ([Radwan et al., 2015](#)) and observed decreases in normal morphology with increases in SO₂ exposure; however, all other quality metrics showed null associations. [Slama et al. \(2013\)](#) examined fecundity rate ratios (FRs) with SO₂ exposures before and after the initiation of unprotected intercourse in a Czech Republic population. Exposures prior to intercourse initiation (long-term, ~30 or 60 days) had slightly reduced FRs; however, SO₂ was highly correlated with PM_{2.5} and NO₂ in this population (i.e., $r > 0.7$) and stronger reductions in fertility were observed with those pollutants. [Legro et al. \(2010\)](#) examined odds of live birth in a population undergoing in vitro fertilization and observed null associations for SO₂ with all exposure windows from medication start to birth (short-term windows during in vitro fertilization, long term from transfer to pregnancy).

Mixed effect estimates are observed with SO₂ exposure across other pregnancy-related outcomes. Recent studies examined increased blood pressure during pregnancy or pregnancy-related hypertensive disorders, including pre-eclampsia. Several studies observed no associations between SO₂ exposure during the first trimester and changes in late pregnancy blood pressure ([Lee et al., 2012](#)) or hypertensive disorders ([Michikawa et al., 2015](#)); however, a study in Florida observed increased hypertension with higher SO₂ exposure during the first trimester ([Xu et al., 2014](#)). [Mendola et al. \(2016b\)](#) observed a positive association between pre-eclampsia and SO₂ exposure among people with asthma, but not among people without asthmas; the interaction between exposure to SO₂ and asthma was statistically significant for the first trimester exposure window. A small Iranian study found no association between pre-eclampsia and SO₂ above versus below median concentrations ([Nahidi et al., 2014](#)). [Assibey-Mensah et al. \(2015\)](#) observed no effect of SO₂ on hypertensive disorders in Beijing comparing 2008 Olympic period with same calendar days in 2009. In fact, there was an inverse relationship between SO₂ exposure in the third trimester and hypertensive disorders.

In other pregnancy-related outcomes, no associations were observed in the Allegheny County, PA population for short-term near-birth exposures and C-reactive protein, an inflammatory biomarker linked to increased risk of preterm birth ([Lee et al., 2011b](#)). [Michikawa et al. \(2016\)](#) observed positive associations with SO₂ exposure and placenta previa in a Japanese population, although the associations were smaller and less consistent than those observed for ozone or suspended particulate matter (SPM). Increases in SO₂ exposure during the preconception period and the first trimester were associated with increased odds of gestational diabetes mellitus ([Robledo et al., 2015](#)). [Assibey-Mensah et al. \(2015\)](#) examined other fetal-placental conditions, and observed no associations with SO₂ exposure in the first or second trimester, but reported a positive association with fetal-placental conditions and third trimester SO₂ exposures in Beijing comparing the 2008 Olympic period with the same calendar days in 2009. As part of the Consortium on Safe Labor, a retrospective cohort study that included 16 counties across the U.S., [Wallace et al. \(2016\)](#) observed positive associations between premature rupture of membranes and SO₂ exposure averaged over the whole pregnancy, but not for shorter exposure windows (i.e., days or hours before rupture).

No recent animal studies evaluating fertility and pregnancy were identified. An older study in laboratory animals exposed to SO₂ demonstrated reproductive toxicity in adult female rodents and their offspring. Adult female albino rats were exposed to either 0.057 ppm or 1.5 ppm SO₂ by inhalation for 72 days [[Mamatsashvili, 1970b](#)]; see [Table 5-37](#)]. During the first month of treatment at 1.5 ppm, substantial alterations in stages of the estrus cycle were seen including significant decreases in duration of diestrus and metaestrus. During the 2nd and 3rd month of exposure, prolongation of estrus cyclicity was found with exposure to 1.5 ppm SO₂, leading to fewer estrus cycles during the study period. This change was not permanent as by 7 months after exposure ceased, estrus cyclicity returned to normal. Exposure of adult female rodents to SO₂ caused disruption of estrus cyclicity that was not permanent as it returned to normal after cessation of SO₂ exposure.

While studies of fertility, reproduction, and pregnancy are limited in number, generally, SO₂ exposures appear to have no association with these outcomes. A group of studies examining hypertensive disorders during pregnancy report inconsistent results, with the majority observing no association with SO₂ exposure. Similarly, studies examining endpoints related to fertility and other pregnancy conditions are generally inconsistent, with the majority observing no association, and few studies examining any one specific outcome. Additionally, these studies do not provide evidence to help reduce uncertainty related to exposure measurement error, copollutant confounding, or biological mechanism by which SO₂ could cause these effects. These studies are summarized in Supplemental Table 5S-22 ([U.S. EPA, 2017c](#)).

5.4.3 Birth Outcomes

This section discusses several categories of birth outcomes, including fetal growth ([Section 5.4.3.1](#)), preterm birth ([Section 5.4.3.2](#)), birth weight ([Section 5.4.3.3](#)), birth defects ([Section 5.4.3.4](#)), fetal mortality ([Section 5.4.3.5](#)), and infant mortality ([Section 5.4.3.6](#)).

5.4.3.1 Fetal Growth

Fetal growth can be difficult to quantify; typically, small for gestational age (SGA) or intra-uterine growth restriction (IUGR) are used. These designations, often used interchangeably, are defined as infants with a birth weight below the 10th percentile for gestational age, usually with consideration for sex and race as well. There are a number of limitations in using SGA/IUGR as a metric of poor fetal growth. One is that a percentile-based measurement will always quantify a certain percentage of the infant population as growth restricted whether or not this is truly the case ([Wollmann, 1998](#)). For example, in term infants, it is unlikely that 10% are actually growth restricted. Whereas in preterm infants, it is likely that more than 10% are growth restricted; therefore, SGA cases would be overestimated in term infants and underestimated in preterm infants. In addition, exact definitions shift between studies and some studies use alternate definitions of SGA/IUGR. For example, some studies use the birth weight distribution of their study population for defining SGA ([Le et al., 2012](#)), which will naturally not be identical for every study population, and others use country standards, likely to be more stable over time ([Brauer et al., 2008](#); [Liu et al., 2003](#)). An alternate approach to categorizing growth restriction is to use ultrasound images during gestation ([Woodruff et al., 2009](#)). This approach has the benefit of examining all fetuses with ultrasounds, being less subjective to population definition, and distinguishing true growth restriction from merely small-sized infants. However, not all women receive prenatal care and ultrasounds, leading to the possibility of selection bias.

Several studies report positive associations between fetal growth and SO₂, although timing of exposure is inconsistent. A recent study conducted in Australia examined ultrasound measures in midgestation in association with SO₂ exposures during early pregnancy ([Hansen et al., 2008](#)). [Hansen et al. \(2008\)](#) observed decreases in biparietal diameter and abdominal circumference with increases in SO₂ during the first 4 months of pregnancy [5-ppb SO₂ increase in the 1st month: -4.25 mm (-6.81, -1.69) biparietal diameter; -9.31 mm (-19.31, 0.69) abdominal circumference]. Recent studies using the traditional definition of SGA/IUGR had mixed results. In Vancouver, increases in ORs for SGA were observed with entire pregnancy exposures ([Brauer et al., 2008](#)) and with

1st month and 1st trimester exposures ([Liu et al., 2003](#)). [Rich et al. \(2009\)](#) used an alternate definition of SGA—having a growth ratio (infant birth weight divided by median study cohort birth weight) below 0.75 for very SGA (VSGA), and between 0.75–0.85 for SGA—and observed elevated ORs with 1st trimester exposures for SGA, and 2nd and 3rd trimester exposures for VSGA. Other studies did not observe positive associations between fetal growth and SO₂. In a study conducted in Italy, ([Capobussi et al., 2016](#)) observed a null association for SGA when SO₂ exposure was estimated for the entire pregnancy, but modest positive associations when exposure was averaged across the first or second trimester. In a study conducted in Calgary, Edmonton, and Montreal, [Liu et al. \(2007\)](#) found lowered ORs for IUGR with exposures in Months 1 to 5 of pregnancy and no associations in Months 6 to 9. Of the two recent studies in the U.S., [Le et al. \(2012\)](#) observed generally null associations for SGA and 1st and last month exposures; ORs with trimester exposure windows were null, although ORs became elevated for the 2nd and 3rd trimesters after adjustment for CO, NO₂, and PM₁₀.

In summary, there is inconsistent evidence for increased odds of fetal growth restriction with exposure to SO₂ during pregnancy, and the evidence lacks consistency in fetal growth definition/metric and in exposure timing. Mean SO₂ exposures for these studies are generally low, although all studies examine average daily SO₂ concentrations. Additionally, these studies do not provide evidence to help reduce uncertainty related to exposure measurement error, copollutant confounding, or the biological mechanism by which SO₂ could cause these effects. No recent animal studies evaluating fetal growth were identified. Studies examining the association between SO₂ and fetal growth can be found in Supplemental Table 5S-23 ([U.S. EPA, 2017c](#)).

5.4.3.2 Preterm Birth

Preterm birth, delivery that occurs before 37 weeks of completed gestation, is a marker for fetal underdevelopment and a risk factor for further adverse health outcomes (e.g., infant mortality, neurodevelopmental problems, growth issues) ([Mathews and MacDorman, 2010](#); [Saigal and Doyle, 2008](#); [IOM, 2007](#); [Gilbert et al., 2003](#)). PTB is characterized by multiple etiologies (spontaneous, premature rupture of membranes, or medically induced), and identifying exact causes of PTB is difficult. It is likely that some mechanistic pathways are shared among the three groups; however, isolated causes are also likely to exist. Few, if any, studies distinguish among these three groups in examining associations between air pollution and PTB.

Given the uncertainty surrounding modes of action leading to PTB, many of the studies reviewed here consider both short- and long-term exposure periods. For example,

exposure periods may include all of gestation or a particular trimester for long-term exposure windows, or weeks or days leading up to birth for short-term exposure windows. With near-birth exposure periods, development will be at different points for term and preterm infants (e.g., exposure 2 weeks before birth is at 34 weeks for a 36-week PTB, and 38 weeks for a 40-week term birth), which suggests the possibility of different modes of action for increases in risk observed with near-birth exposures compared to exposures in specific periods of fetal development.

There is evidence supporting a relationship between SO₂ and preterm birth, primarily with exposure near-birth and including both older and newer studies. Among a U.S. birth cohort, [Mendola et al. \(2016a\)](#) examined PTB and exposure to SO₂ during different periods before and during pregnancy, observing generally null results among women with and without asthma, except for when exposure was limited to weeks near birth (specifically Weeks 34 and 36) for which positive associations were observed among women with asthma, but not for women without asthma. Studies in Europe and Asia reported increased ORs/RRs of PTB with exposures across pregnancy, although not consistently between studies ([Dibben and Clemens, 2015](#); [Zhao et al., 2011](#); [Leem et al., 2006](#); [Bobak, 2000](#); [Xu et al., 1995](#)). In a recent time-series analysis, [Zhao et al. \(2011\)](#) found increased RRs with SO₂ exposure lagged 0–3 days from birth, but SO₂ was also highly correlated with PM₁₀ (Pearson correlation coefficient = 0.75) and NO₂ (Pearson correlation coefficient = 0.84) in the study area. [Dibben and Clemens \(2015\)](#) used a pollution-climate model to assign SO₂ concentrations with high spatial resolution as well as incorporating daily activity data into the exposure and observed null associations with PTB and modest, positive associations with very preterm birth among births in Scotland. [Qian et al. \(2015\)](#) observed weak negative or null associations between SO₂ exposures and PTB across a range of different exposure windows among a birth cohort in Wuhan, China.

In the U.S. and Canada, studies of SO₂ and PTB in Pennsylvania ([Sagiv et al., 2005](#)) and Vancouver ([Liu et al., 2003](#)) found increased ORs with near-birth exposures [6 week prebirth RR = 1.05 (1.00, 1.10) [Sagiv et al. \(2005\)](#); last month OR = 1.09 (1.01, 1.19) per 5-ppb increase [Liu et al. \(2003\)](#)]. More recently, in a Detroit, MI cohort, [Le et al. \(2012\)](#) found similar associations for exposures in the last month of pregnancy [OR 4th to 1st quartile: 1.07 (1.01, 1.14)]. Another Vancouver cohort, examining entire pregnancy exposure, only observed increases [OR = 1.03 (0.93, 1.15) per 5-ppb SO₂ increase] with PTB <30 weeks ([Brauer et al., 2008](#)). Recent time-series and case-crossover studies in Atlanta, GA and Brisbane, Australia observed null associations for both 1st month and near-birth exposures using 1-h max SO₂ [exposure during last week of pregnancy RR per 5-ppb increase = 0.99 (0.98, 1.01)] ([Darrow et al., 2009](#)) and SO₂ concentrations 24–48 hours preceding the onset of labor ([Li et al., 2016](#)). Finally, a cross-sectional study

of PTB across the U.S. reported that SO₂ showed “nonsignificant” effects with PTB for exposures during the month of birth ([Trasande et al., 2013](#)). In contrast, a recent study conducted in Italy observed negative associations between SO₂ exposure averaged across the entire pregnancy as well as each trimester and PTB, suggesting the SO₂ exposure was associated with longer gestation ([Capobussi et al., 2016](#)).

In summary, there is some evidence for an association between exposure to SO₂ and preterm birth particularly with near-birth exposure windows. Studies examining PTB primarily used average daily SO₂. The one study that examined 1-h max SO₂ found no associations for PTB. Recent studies do not provide evidence to help reduce uncertainty related to exposure measurement error, copollutant confounding, or the biological mechanism by which SO₂ could cause preterm birth. No recent animal studies evaluating preterm birth were identified. Studies are characterized in Supplemental Table 5S-24 ([U.S. EPA, 2017c](#)).

5.4.3.3 Birth Weight

Birth weight is a measure of fetal growth and an important indicator of future infant and child health. Birth weight is determined by gestational age and intra-uterine growth, as well as maternal, placental, fetal, and environmental factors. Vulnerability to environmental insults affecting birth weight may occur throughout pregnancy. Implantation or formation of the placenta may be disrupted in the earliest weeks of pregnancy, leading to decreased fetal nutrition throughout pregnancy; or inflammation might result in constriction of the umbilical cord during the later trimesters resulting in poor fetal nutrition. As the largest gains in birth weight occur during the last weeks of gestation, this may be a particularly vulnerable period for birth weight outcomes. Information on birth weight is routinely collected for vital statistics; given that measures of birth weight do not suffer the same uncertainties as gestational age or growth restriction, it is one of the most studied outcomes within air pollution and reproductive health. Birth weight may be examined as a continuous outcome or a dichotomous outcome as low birth weight (LBW) (less than 2,500 g or 5 lbs, 8 oz).

Studies examining LBW have found elevated ORs with exposures in the first trimester or first month ([Dugandzic et al., 2006](#); [Lee et al., 2003](#); [Liu et al., 2003](#); [Ha et al., 2001](#)) and with entire pregnancy exposures ([Capobussi et al., 2016](#); [Dibben and Clemens, 2015](#); [Yorifuji et al., 2015a](#); [Ebisu and Bell, 2012](#); [Kumar, 2012](#); [Morello-Frosch et al., 2010](#)). In the two studies that examined distance to monitor, using concentrations from closer monitors lead to stronger effect estimates ([Kumar, 2012](#); [Morello-Frosch et al., 2010](#)).

Some studies examining entire pregnancy exposure have also observed null associations between SO₂ and LBW ([Brauer et al., 2008](#); [Bell et al., 2007](#)).

Studies examining continuous birth weight (Δ g) have inconsistent results. In a northeast U.S. population, [Bell et al. \(2007\)](#) observed no association with change in birth weight for entire pregnancy exposure [-2.711 g (-13.253, 7.831) per 5 ppb SO₂], including in a stratified analysis of white and black mothers. [Kumar \(2012\)](#) reported results that shifted around the null based on distance from monitor in Chicago; some effects were positive, and some negative but all had wide confidence intervals. And, in a cross-sectional study across the county, [Trasande et al. \(2013\)](#) reported only “nonsignificant” effects for SO₂. One recent California cohort study reported increases in birth weight with increases in SO₂ exposure in entire pregnancy and first trimester, although effects were reduced with use of closer monitors ([Morello-Frosch et al., 2010](#)). A recent Texas study observed decreases in birth weight with county average SO₂ exposure for the entire pregnancy [-15.594 g (-25.344, -5.844)] ([Geer et al., 2012](#)). A study in Beijing during the summer Olympics of 2008 found increased SO₂ in the 8th month of pregnancy associated with decrements in birth weight; however, SO₂ was highly correlated with PM_{2.5} and CO, which showed similar patterns of effect ([Rich et al., 2015](#)). Finally, a recent study in Atlanta found decreases in birth weight with increases in 3rd trimester 1-h max SO₂ ([Darrow et al., 2011](#)). This effect was stronger in non-Hispanic white and Hispanic mothers than non-Hispanic black mothers ([Darrow et al., 2011](#)).

No recent animal studies evaluating birth weight-related outcomes were identified. In laboratory animals from an older study, exposure to SO₂ affected birth outcomes in adult female rodents and their offspring. Adult female albino rats were exposed to either 0.057 ppm or 1.5 ppm SO₂ by inhalation for 72 days ([Mamatsashvili, 1970b](#)). At birth, litter sizes were significantly increased in number from dams that were exposed to SO₂ versus control dams ([Table 5-37](#)).

In summary, there is some evidence that LBW may be associated with SO₂, while evidence for an association with change in birth weight is inconsistent. Overall, the results of studies of LBW and birth weight remain inconsistent and these do not provide evidence to help reduce uncertainty related to exposure measurement error, copollutant confounding, or the biological mechanism by which SO₂ could cause these effects. Studies for both LBW and change in birth weight can be found in Supplemental Table 5S-25 ([U.S. EPA, 2017c](#)).

5.4.3.4 Birth Defects

Birth defects are structural and functional abnormalities that can cause physical disability, intellectual disability, and other health problems. They are a leading cause of infant mortality and developmental disability in the U.S. (Mai et al., 2016). Since 2008, there have been several studies examining birth defects and SO₂ during pregnancy, particularly during Weeks 3–8 of gestation, which is thought to be highly vulnerable to insults resulting in birth defects. Because birth defects as a whole are rare and specific birth defects are rarer, these studies often have effect estimates with very wide confidence intervals. Individual studies often look at different types of birth defects, meaning the body of work examining any one birth defect may still be limited. Cardiac birth defects and oral cleft defects are the most commonly studied anomalies. However, results (even for these defects) are inconsistent across studies. For example, odds of ventricular septal defects have been found to be increased (Gianicolo et al., 2014; Stingone et al., 2014; Agay-Shay et al., 2013; Gilboa et al., 2005), decreased (Hwang et al., 2015b; Dadvand et al., 2011a, b; Rankin et al., 2009), and null (Strickland et al., 2009) with increases in SO₂ exposure. Odds of cleft lip with or without cleft palate have been found to be increased (Zhu et al., 2015), decreased (Hwang and Jaakkola, 2008; Gilboa et al., 2005), or null (Dolk et al., 2010; Rankin et al., 2009) with increases in SO₂ exposure. A single study of limb deformities found increased odds with exposure to SO₂ during Weeks 9–12 of pregnancy (Lin et al., 2014). Two studies examining repeating chromosomal defects found no association or correlation between trisomy 21 or any sperm disomy and SO₂ (Chung et al., 2014; Jurewicz et al., 2014). Studies of any congenital anomaly in Israel and China have reported inverse associations with increasing SO₂ (Farhi et al., 2014; Liang et al., 2014).

In summary, results for birth defects are either inconsistent across studies or limited in number of studies. No recent animal studies evaluating birth defects were identified. Studies of birth defects and SO₂ are characterized in Supplemental Table 5S-26 (U.S. EPA, 2017c).

5.4.3.5 Fetal Mortality

Fetal mortality or stillbirth is the intra-uterine death of a fetus. In most areas fetal deaths are only reported after 20 weeks of completed gestation; this leads to potential bias, as the population at risk of fetal death is any conception but the actual measured population is only those fetuses reaching at least 20 weeks gestational age. A single recent case-control study of spontaneous abortion occurring before 14 weeks of gestation found no associations with SO₂ exposures determined by time-weighted concentrations for

residence and workplace ([Moridi et al., 2014](#)). A recent large California cohort found no associations between stillbirth and increasing SO₂ exposure ([Green et al., 2015](#)). In recent studies of a New Jersey population examining both long-term and short-term exposure windows, ORs for fetal death were elevated with a 2-day lag [OR per 5-ppb increase in SO₂: 1.12 (1.02, 1.24)] and with exposures across pregnancy and in each trimester, particularly the 3rd trimester [OR per 5-ppb increase in SO₂: 1.47 (1.05, 1.69)] ([Faiz et al., 2013](#); [Faiz et al., 2012](#)). [Hwang et al. \(2011\)](#) examined fetal mortality among term and preterm deliveries in Taiwan, finding elevated associations for exposures during the 1st trimester only among preterm deliveries. Other studies have also found increased associations between SO₂ and fetal mortality, although mean SO₂ concentrations were higher in these studies ([Hou et al., 2014](#); [Pereira et al., 1998](#)). [Pereira et al. \(1998\)](#) observed elevated RRs in a São Paulo, Brazil time series with short-term exposure. A recent study by [Enkhmaa et al. \(2014\)](#) found very strong correlations between seasonal SO₂ and fetal death, and [Hou et al. \(2014\)](#) found elevated ORs with long-term exposures around the time of conception. However, the [Hou et al. \(2014\)](#) models were unadjusted for confounding factors, and confidence intervals were very wide. In the study by [Enkhmaa et al. \(2014\)](#), other pollutants also showed very strong correlations with fetal death and the pollutants were highly correlated with one another (i.e., $r > 0.7$).

In summary, although few in number, studies of fetal mortality and SO₂ show elevated associations for both short- and long-term exposures. However, these studies are limited by the uncertainties associated reproductive and developmental outcomes identified in the 2008 SO_x ISA. No recent animal studies evaluating fetal mortality were identified. Studies are characterized in Supplemental Table 5S-27 ([U.S. EPA, 2017c](#)).

5.4.3.6 Infant Mortality

Studies of infant mortality and SO₂ are limited in number. In a U.S. study, [Woodruff et al. \(2008\)](#) observed increased ORs for respiratory-related post-neonatal infant mortality with long-term (2 months) exposure increases in county-level SO₂ concentrations [OR = 1.09 (0.89, 1.36) per 5-ppb increase]. This association remained after adjusting for other pollutants. A time-series study in Seoul, South Korea observed increased RRs for all cause post-neonatal infant mortality with short-term SO₂ exposure, although exact timing of exposure was unclear ([Son et al., 2008](#)). No recent animal studies evaluating postnatal mortality were identified. Studies are characterized in Supplemental Table 5S-27 ([U.S. EPA, 2017c](#)).

5.4.4 Developmental Outcomes

5.4.4.1 Respiratory Outcomes

Recent studies examined asthma onset in association with early life exposure to SO₂. [Clark et al. \(2010\)](#), [Liu et al. \(2016\)](#), [Deng et al. \(2015b\)](#), and [Deng et al. \(2015a\)](#) observed elevated ORs for asthma with SO₂ exposure during pregnancy and the first year of life. [Nishimura et al. \(2013\)](#) observed elevated ORs for asthma with SO₂ exposure in the first 3 years of life, but not the first year of life alone. Asthma onset is covered in further detail in [Section 5.2.1.2](#).

In a time-series study, [Dales et al. \(2006\)](#) investigated neonatal hospitalizations due to respiratory causes in Atlanta, GA; they observed elevated ORs with 2-day lagged SO₂ exposure. After adjustment for gaseous copollutants, confidence intervals for associations with gaseous pollutants and PM₁₀ were very large, but effect estimates remained elevated. Hospitalizations due to respiratory causes are covered in [Section 5.2.1.6](#).

In summary, there is some evidence for an association between gestational and early life exposure to SO₂ and respiratory health effects that extend into early childhood, although evidence is limited and exposure windows are uncertain. Key studies are summarized in [Table 5-36](#).

5.4.4.2 Other Developmental Effects

Studies examining other developmental exposures are limited in number. [Peel et al. \(2011\)](#) examined SO₂ exposure with apnea and bradycardia in a subpopulation of infants in Atlanta and observed no association for either health outcome. [Huang et al. \(2015a\)](#) observed no associations between prenatal and early life SO₂ exposures and atopic dermatitis among infants in Taiwan. [Poursafa et al. \(2016\)](#) examined the association between SO₂ exposure during pregnancy and markers of endothelial dysfunction (i.e., ICAM-1, vascular adhesion molecule-1, endothelin-1) in cord blood. They observed a positive association with endothelin-1, but not for other markers of endothelial dysfunction. Among a Japanese cohort, prenatal exposure to SO₂ was associated with verbal and fine motor delays assessed at ages 2.5 and 5.5 years ([Yorifuji et al., 2015b](#)). In an older study from the animal toxicology literature, adult female albino rats were exposed to either 0.057 ppm or 1.5 ppm SO₂ by inhalation, 12 hours/day for 72 days ([Mamatsashvili, 1970b](#)). Changes in offspring postnatal growth or body weight over time were reported with 1.5-ppm exposure.

Sulfur dioxide-dependent synaptic injury was measured in adolescent male rats exposed to 1.24 ppm SO₂ for 6 hours/day for 90 days ([Yun et al., 2013](#)). Nonsignificant morphological changes were seen in the hippocampal synaptic junctions using transmission electron microscopy. In the hippocampus, the synaptic vesicle membrane protein synaptophysin (SYP) was significantly downregulated as was ERK1/2 phosphorylation. Phosphorylation is an important contributor to synaptic plasticity. Thus, SO₂ exposure in the adolescent rat contributes to downregulation of synaptic vesicle protein SYP and decreased ERK1/2 phosphorylation, indicative of disruption at the hippocampal synapse.

5.4.5 Summary and Causal Determination

Overall the evidence is inadequate to infer a causal relationship between exposure to SO₂ and reproductive and developmental outcomes. This is consistent with the 2008 ISA for Sulfur Oxides, which also concluded the evidence was inadequate to infer a causal relationship with reproductive and developmental effects. All available evidence, including more than 50 recent studies, examining the relationship between exposure to SO₂ and reproductive and developmental effects was evaluated using the framework described in the Preamble to the ISAs ([U.S. EPA, 2015b](#)). The key evidence as it relates to the causal framework is summarized in [Table 5-38](#).

There are several well-designed, well-conducted epidemiologic studies, many described in papers published since the previous ISA, that indicate an association between SO₂ and reproductive and developmental health outcomes; the bulk of the evidence exists for adverse birth outcomes. For example, several high quality studies reported positive associations between SO₂ exposures during pregnancy and fetal growth metrics ([Le et al., 2012](#); [Rich et al., 2009](#); [Brauer et al., 2008](#); [Liu et al., 2003](#)), preterm birth ([Mendola et al., 2016a](#); [Le et al., 2012](#); [Zhao et al., 2011](#); [Sagiv et al., 2005](#); [Liu et al., 2003](#)), birth weight ([Ebisu and Bell, 2012](#); [Darrow et al., 2011](#); [Morello-Frosch et al., 2010](#); [Liu et al., 2003](#)), and fetal and infant mortality ([Faiz et al., 2012](#); [Hwang et al., 2011](#); [Woodruff et al., 2008](#)). However, the evidence is not entirely consistent, and has not substantially reduced any of the uncertainties connected with the associations observed between exposure to SO₂ and birth outcomes that were identified in the previous ISA.

One uncertainty is timing of exposure, wherein associations remain inconsistent among studies and across outcomes. For example, some studies observe the strongest associations when exposure is averaged over the entire pregnancy, while others observe the strongest association when exposure is averaged over either the first, second, or third trimester. As an exception to this, studies of PTB generally observed positive associations

between near-birth exposures (e.g., last month of gestation, same, or 3-day lag from birth) ([Mendola et al., 2016a](#); [Le et al., 2012](#); [Zhao et al., 2011](#); [Sagiv et al., 2005](#); [Liu et al., 2003](#)).

Table 5-38 Summary of evidence inadequate to infer a causal relationship between sulfur dioxide exposure and reproductive and developmental effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Overall reproductive and developmental effects—inadequate to infer a causal relationship			
Evidence from multiple epidemiologic studies of preterm birth is generally supportive but key uncertainties remain	Consistent positive associations observed with near-birth exposures to SO ₂ and preterm birth after adjustment for common potential confounders. Associations not evaluated in copollutant models.	Sagiv et al. (2005)	Mean: 7.9 ppb
		† Le et al. (2012)	Mean: 5.8 ppb
		† Mendola et al. (2016a)	Mean: 4.0 ppb
		Section 5.4.3.2	
	Associations robust in copollutant models	Liu et al. (2003)	Mean: 4.9 ppb

Table 5-38 (Continued): Summary of evidence inadequate to infer a causal relationship between sulfur dioxide exposure and reproductive and developmental effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Limited and inconsistent epidemiologic evidence for other birth outcomes	Several studies show positive associations with fetal growth metrics, although definitions vary across studies, and timing of exposure is inconsistent. Associations not evaluated in copollutant models	Section 5.4.3.1	Means: 4.9–5.8 ppb
	Several high quality studies show associations between SO ₂ exposure and low birth weight but not for change in birth weight. Timing of exposure is inconsistent across studies. Only one study uses 1-h max for exposure determination.	Section 5.4.3.3	Means: 2.1–13.2 ppb
	Limited and inconsistent epidemiologic evidence for associations with various birth defects	Section 5.4.3.4	Reported means: 1.9–6
	Limited number of studies of SO ₂ and fetal death, positive associations observed across studies, although timing of exposure and outcome definitions are inconsistent. Limited evidence for an association with SO ₂ in respiratory related infant mortality.	Section 5.4.3.6	Mean: 5.7 ppb Mean: 5.8 ppb Mean: 5.9 ppb Mean: 3 ppb
	Limited evidence for positive associations between prenatal/early life exposures and childhood respiratory outcomes	Section 5.4.4.1	Means: 2–4.3 ppb
Limited evidence for key events in proposed mode of action	Altered menstrual function, fetal growth, and birth weight outcomes with impaired postnatal growth for in utero exposed pups	Mamatsashvili (1970a)	57 or 1,427 ppb
Lack of evidence from epidemiologic studies to support an association of SO ₂ exposure with detrimental effects on fertility or pregnancy	A limited number of studies on fertility and pregnancy outcomes show no associations with SO ₂ .	Section 5.4.4.1	Mean 8.4–59 ppb

Table 5-38 (Continued): Summary of evidence inadequate to infer a causal relationship between sulfur dioxide exposure and reproductive and developmental effects.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Uncertainty regarding potential confounding by copollutants	Limited adjustment for copollutants, with no clear directionality or trends for effect estimate shifts after adjustment	†(Faiz et al. (2013); Slama et al. (2013); Le et al. (2012))	
Uncertainty regarding exposure measurement error	Fixed-site monitors subject to some degree of exposure error. Uncharacterized spatial and temporal heterogeneity may introduce exposure error in long-term effects and bias could be toward or away from the null.	Chapter 3 Section 3.4.4.2	
Uncertainty regarding exposure timing for specific outcomes	Associations of exposure to SO ₂ at particular windows during pregnancy are inconsistent among studies and across outcomes		

Max = maximum; SO₂ = sulfur dioxide.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in the Preamble to the ISAs (U.S. EPA, 2015b).

^bDescribes the key evidence and references contributing most heavily to causal determination and where applicable to uncertainties and inconsistencies. References to earlier sections indicate where full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated (for experimental studies, below 2,000 ppb).

†Studies published since the 2008 ISA for Sulfur Oxides.

Another uncertainty centers on spatial and temporal variability in SO₂ exposures. SO₂ is a temporally and spatially heterogeneous pollutant; it is difficult to accurately estimate for “long-term” exposures, and there is the potential for exposure measurement error in long-term SO₂ exposures to bias estimates toward or away from the null (Section 3.4.4.2). None of the epidemiologic studies made corrections or adjustments for exposure measurement error or accounted for the potential for bias away from the null, the potential for which has been demonstrated in simulation studies (see Section 3.4.4.2). Current epidemiologic methods are not able to disentangle whether associations are due to extended exposure to moderate concentrations of SO₂ or repeated short-term exposure to peaks in SO₂ concentration.

Potential confounding by copollutants may explain some of the observed associations and cannot be ruled out. SO₂ is part of a mix of ambient air pollution; SO₂ shares sources with particulate matter and is chemically linked to sulfate. Few studies evaluate or provide information that would inform the independent effect of SO₂ in the context of the greater

air pollution mixture, and of those that do, no clear trends for the effects of copollutant adjustment are apparent ([Faiz et al., 2013](#); [Slama et al., 2013](#); [Le et al., 2012](#)).

There is insufficient information on potential modes of action of SO₂ on reproductive outcomes at relevant exposure levels for this ISA ([Chapter 4](#)). In a single older study from [Mamatsashvili \(1970a\)](#), SO₂ inhalation exposure in laboratory rodents demonstrated reproductive changes in exposed females and their offspring, altered birth outcomes, and developmental effects. The specific outcomes affected after SO₂ exposure included altered estrus cycle length of F0 and F1 generations, decrements in offspring body weight gain or growth after in utero exposure, and changes in litter size. The majority of the remaining animal toxicological evidence for reproductive and developmental effects is for exposure at 5,000 ppb or greater, doses which are beyond the scope of this document.

Since the 2008 ISA for Sulfur Oxides, researchers have begun evaluating more health outcomes, including fertility, effects on pregnancy (e.g., pre-eclampsia, gestational diabetes), and developmental effects. For each of these individual outcomes the literature base is small, but new studies are quickly accumulating. However, at present there is little coherence or consistency among epidemiologic and toxicological studies for these outcomes. In general, it is challenging to synthesize study findings on the wide variety of health outcomes collected under the reproductive and developmental effects heading. Given the wide variety of potential mechanisms or adverse outcome pathways that could affect this breadth of outcomes, coherence is unlikely to be reached given the limited literature base.

Overall, many uncertainties remain when evaluating the evidence for these health endpoints; therefore, the evidence is inadequate to infer a causal relationship between exposure to SO₂ and reproductive and developmental outcomes.

5.5 Mortality

5.5.1 Short-Term Exposure

5.5.1.1 Introduction

The 2008 SO_x ISA concluded that the collective evidence is “suggestive of a causal relationship” between short-term SO₂ exposure and mortality. Overall, the number of studies that examined the relationship between short-term SO₂ exposure and mortality was sparse and there was limited data available to inform the potential for copollutant

confounding. Since the completion of the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), epidemiologic literature that has examined the association between short-term SO₂ exposure and mortality has expanded. However, similar to the collection of studies evaluated in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), most of the recent studies do not focus specifically on the SO₂-mortality relationship but instead on PM or O₃. Of the studies identified, a limited number have been conducted in the U.S., Canada, and Europe, with the majority being conducted in Asia due to the increased focus on examining the effect of air pollution on health in developing countries.

Studies included in the 1982 AQCD that examined the association between short-term SO_x exposure, mainly SO₂, and total mortality were limited to historical data on high air pollution episodes ([U.S. EPA, 1982a](#)). These studies were unable to decipher whether the associations observed were due to particle pollution or SO₂. Additional studies evaluated in the 1986 Second Addendum to the 1982 AQCD ([U.S. EPA, 1986](#)) further confirm the findings of these initial studies, but were still unable to address uncertainties and limitations related to examining the effect of SO₂ exposure on mortality, especially at lower concentrations.

In the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), a larger body of literature was available to assess the relationship between short-term SO₂ exposures and mortality; however, these studies were still limited in that they primarily focused on PM, with SO₂ only being examined in single-pollutant models. These studies found that excess risk estimates for total mortality due to short-term SO₂ exposure from multicity studies and meta-analyses generally ranged from 0.4 to 2.0% for a 10-ppb increase in 24-h avg SO₂ concentrations. These associations were primarily observed at mean 24-h avg SO₂ concentrations <15 ppb. Studies that examined cause-specific mortality found evidence of risk estimates larger in magnitude for respiratory and cardiovascular mortality compared to total mortality with the largest associations for respiratory mortality. The larger SO₂-respiratory mortality associations observed in the epidemiologic literature were coherent with the scientific evidence providing stronger support for SO₂ effects on respiratory morbidity compared to cardiovascular morbidity ([U.S. EPA, 2008d](#)).

An examination of potential copollutant confounding of the SO₂-mortality relationship was sparse. Studies evaluated in the 2008 SO_x ISA found that SO₂-mortality risk estimates from copollutant models were robust but imprecise (i.e., wide confidence intervals). An additional study that examined the potential interaction between copollutants [i.e., SO₂ and BS] did not find evidence of interaction when stratifying days by high and low concentrations of BS ([Katsouyanni et al., 1997](#)). Of the studies evaluated, only the Air Pollution and Health: A European Approach (APHEA) study examined seasonality and potential effect modifiers of the SO₂-mortality relationship.

APHEA provided initial evidence that mortality effects are larger during the warm season and that geographic location may influence city-specific SO₂-mortality risk estimates, respectively ([Katsouyanni et al., 1997](#)). The consistent, positive SO₂-mortality associations observed across studies were supported by an intervention study conducted in Hong Kong that examined the health impact of converting to fuel oil with low sulfur content and found evidence suggesting that a reduction in SO₂ concentrations leads to a reduction in mortality ([Hedley et al., 2002](#)).

As detailed in previous ISAs [e.g., [U.S. EPA \(2013b\)](#)], this section focuses primarily on multicity studies because they examine the association between short-term SO₂ exposure and mortality over a large geographic area using a consistent statistical methodology, which avoids the potential publication bias often associated with single-city studies ([U.S. EPA, 2008d](#)). However, where applicable single-city studies are evaluated that encompass a long study-duration, provide additional evidence indicating that a specific population or lifestage is at increased risk of SO₂-related mortality, or address a limitation or uncertainty in the SO₂-mortality relationship not represented in multicity studies. The remaining studies identified are not evaluated in this section due to issues associated with study design or insufficient sample size and are detailed in Supplemental Table 5S-28 ([U.S. EPA, 2017c](#)).

The organization of the material on short-term SO₂ exposure and mortality is as follows. [Section 5.5.1.2](#) evaluates studies that examined the association between short-term SO₂ exposure and mortality, with the remaining sections addressing key limitations and uncertainties in the SO₂-mortality relationship that were evident at the completion of the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Subsequent sections evaluate whether there is evidence of confounding (i.e., copollutants and seasonal/temporal) ([Section 5.5.1.3](#)), effect modification (i.e., sources of heterogeneity in risk estimates across cities or within a population) ([Section 5.5.1.4](#)), modification of the SO₂-mortality association including seasonal heterogeneity ([Section 5.5.1.4](#)), and the SO₂-mortality C-R relationship and related issues, such as the lag structure of associations ([Section 5.5.1.5](#)).

5.5.1.2 Associations between Short-Term Sulfur Dioxide Exposure and Mortality in All-Year Analyses

Multicity studies and meta-analyses evaluated in the 2008 SO_x ISA reported consistent, positive associations between short-term SO₂ exposure and total mortality in all-year analyses ([U.S. EPA, 2008d](#)). Although only a small number of multicity studies have been conducted since the completion of the 2008 SO_x ISA. These studies, as well as a meta-analysis of studies conducted in Asia ([Atkinson et al., 2012](#)), build upon and provide additional evidence for an association between short-term SO₂ exposure and total

mortality ([Figure 5-17](#)). Air quality characteristics and study specific details for the studies evaluated in this section are provided in [Table 5-39](#).

Table 5-39 Air quality characteristics of multicity studies and meta-analyses evaluated in the 2008 SO_x ISA and recently published multicity studies and meta-analyses.

Study	Location	Years	Mortality Outcome(s)	Averaging Time	Mean Concentration ppb	Upper Percentile Concentrations ppb
North America						
Dominici et al. (2003)	72 U.S. cities (NMMAPS) ^a	1987–1994	Total	24-h avg	0.4–14.2	---
Burnett et al. (2004)	12 Canadian cities	1981–1999	Total cardiovascular respiratory	24-h avg	0.9–9.6	---
† Moolgavkar et al. (2013)	85 U.S. cities (NMMAPS) ^b	1987–2000	Total	24-h avg	---	---
Europe						
Katsouyanni et al. (1997)	12 European cities (APHEA-1)	1980–1992	Total	24-h avg	5.0–28.2 ^c	90th: 17.2–111.8
Biggeri et al. (2005)	Eight Italian cities (MISA-1)	1990–1999	Total cardiovascular respiratory	24-h avg	2.5–15.6	95th: 6.0–50.1 Max: 7.1–111.0
Hoek (2003)	Netherlands	1986–1994	Total cardiovascular respiratory	24-h avg	3.5–5.6	---
† Berglind et al. (2009)	Five European cities ^d	1992–2002	Total	24-h avg	1.0–1.6 ^c	---
† Bellini et al. (2007)	15 Italian cities (MISA-2)	1996–2002	Total cardiovascular respiratory	24-h avg	---	---

Table 5-39 (Continued): Air quality characteristics of multicity studies and meta analyses evaluated in the 2008 SO_x ISA and recently published multicity studies and meta analyses.

Study	Location	Years	Mortality Outcome(s)	Averaging Time	Mean Concentration ppb	Upper Percentile Concentrations ppb
Asia						
†Kan et al. (2010); Wong et al. (2008b); Wong et al. (2010)	Four Asian cities (PAPA)	1996–2004 ^e	Total cardiovascular respiratory	24-h avg	5.0–17.1	75th: 6.0–21.5 Max: 23.4–71.7
†Chen et al. (2012b)	17 Chinese cities (CAPES)	1996–2010 ^f	Total cardiovascular respiratory	24-h avg	6.1–38.2	75th: 6.5–56.1 Max: 25.2–298.5
†Chen et al. (2013)	Eight Chinese cities	1996–2008 ^f	Stroke	24-h avg	6.1–32.1	---
†Meng et al. (2013)	Four Chinese cities	1996–2008 ^g	COPD	24-h avg	6.8–19.1	---
Meta-analyses						
Stieb et al. (2003)	Meta-analysis	1958–1999 ^b	Total	24-h avg	0.7–75.2	---
HEI (2004)	Meta-analysis (South Korea, China, Taiwan, India, Singapore, Thailand, Japan)	1980–2003 ^h	Total	24-h avg	~10–>200	---
†Atkinson et al. (2012)	Meta-analysis (Asia)	1980–2007 ⁱ	Total cardiovascular respiratory COPD	24-h avg	---	---
†Shah et al. (2015)	Meta-analysis	1948–Jan 2014	Stroke	NR	6.2 ^j	Max: 30.2

Table 5-39 (Continued): Air quality characteristics of multicity studies and meta analyses evaluated in the 2008 SO_x ISA and recently published multicity studies and meta analyses.

Study	Location	Years	Mortality Outcome(s)	Averaging Time	Mean Concentration ppb	Upper Percentile Concentrations ppb
†Yang et al. (2014b)	Meta-analysis (Asia, Europe, and North America)	1996–2013	Stroke	24-h avg	Asia: 11.4 ^c Europe: 5.2 ^c North America: 4.2 ^b	75th: Asia: 18.6 Europe: 2.3 North America: 7.6

APHEA = Air Pollution and Health: A European Approach study; avg = average; CAPES = China Air Pollution and Health Effects Study; COPD = chronic obstructive pulmonary disease; ISA = Integrated Science Assessment; max = maximum; MISA = meta-analysis of the Italian studies on short-term effects of air pollution; NMMAPS = The National Morbidity Mortality Air Pollution Study; NR = not reported; PAPA = Public Health and Air Pollution in Asia; SO₂ = sulfur dioxide; SO_x = sulfur oxides.

^aOf the 90 cities included in the NMMAPS analysis only 72 had SO₂ data.

^bOf the 108 cities included in the analyses using NMMAPS data, only 85 had SO₂ data.

^cMedian concentration(s).

^dSO₂ data was not available for Barcelona; therefore, the SO₂ results only encompass four cities.

^eThe study period varied for each city, Bangkok: 1999–2003, Hong Kong: 1996–2002, and Shanghai and Wuhan: 2001–2004.

^fStudy period varied for each city and encompassed 2 to 7 yr, Hong Kong was the only city that had air quality data prior to 2000.

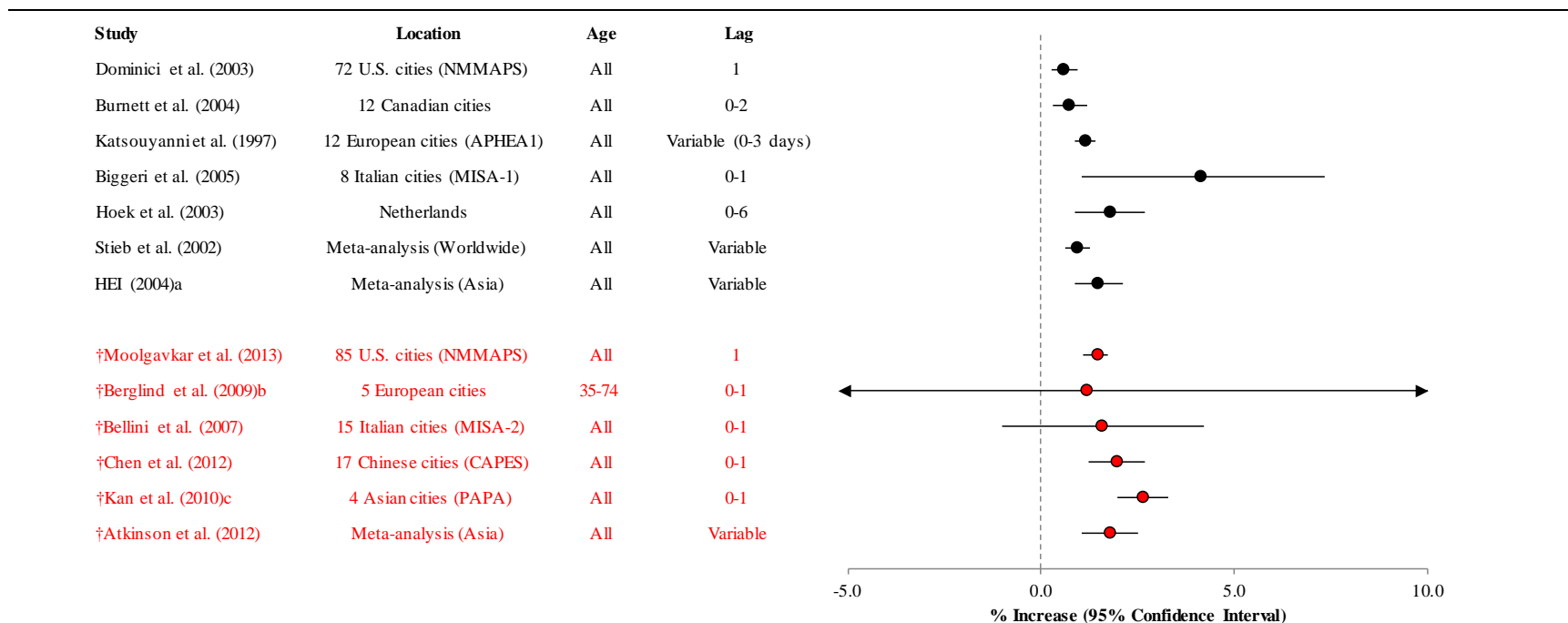
^gStudy period varied from 2 to 7 yr. Hong Kong was the only city that had air quality data prior to 2001.

^hStudies included within this meta-analysis were published during this time period.

ⁱYear defined represents the year in which studies were published that were included in the meta-analysis.

^jThe mortality time series of studies included in the meta-analysis spanned these years.

† = Studies published since the 2008 ISA for Sulfur Oxides.



APHEA = Air Pollution and Health: A European Approach study; CAPES = China Air Pollution and Health Effects Study; ISA = Integrated Science Assessment; MISA = meta-analysis of the Italian studies on short-term effects of air pollution; NMMAPS = The National Morbidity Mortality Air Pollution Study; PAPA = Public Health and Air Pollution in Asia.

^aMeta-analysis of Asian cities: South Korea, China, Hong Kong, Taipei, India, Singapore, Thailand, Japan (HEI, 2004).

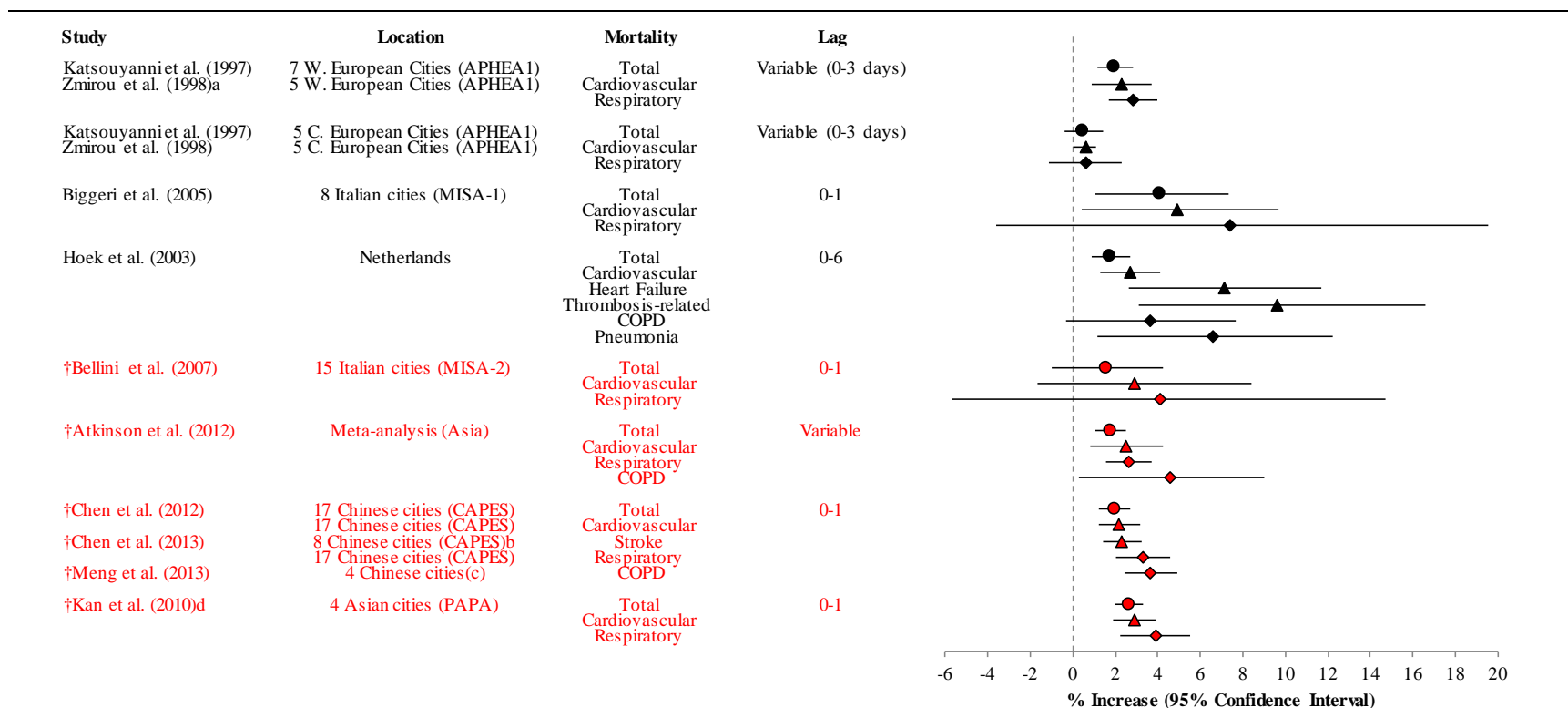
^bStudy was of myocardial infarction survivors therefore only included individuals 35+ (Berglind et al., 2009).

^cKan et al. (2010) reported results that were also found in (Wong et al., 2010; Wong et al. (2008b)).

Note: † and red text/circles = recent studies published since the 2008 ISA for Sulfur Oxides; black text/circles = U.S. and Canadian studies evaluated in the 2008 ISA for Sulfur Oxides. Corresponding quantitative results are reported in Supplemental Table 5S-29 (U.S. EPA, 2017c).

Figure 5-17 Percent increase in total mortality from multicentric studies and meta-analyses evaluated in the 2008 ISA for Sulfur Oxides (black circles) and recently published multicentric studies (red circles) for a 10-ppb increase in 24-hour avg sulfur dioxide concentrations.

When focusing on specific causes of mortality, some studies evaluated in the 2008 SO_x ISA reported similar risk estimates across mortality outcomes [e.g., ([Zmirou et al. \(1998\)](#)); ([Katsouyanni et al. \(1997\)](#))], while others indicated larger risk estimates for respiratory mortality ([Figure 5-18](#)). However, a study conducted in the Netherlands by [Hoek \(2003\)](#) suggested that specific cardiovascular mortality outcomes have larger risk estimates compared to all cardiovascular, total, and respiratory-related mortality outcomes. Recent multicity mortality studies provide additional support indicating larger risk estimates for respiratory mortality compared to total and cardiovascular mortality. Additionally, the results from the studies depicted in [Figure 5-18](#) lend additional support to the body of evidence indicating SO₂-induced respiratory effects presented in the 2008 SO_x ISA, as well as [Section 5.2](#) of this ISA. Unlike the results reported in [Hoek \(2003\)](#), recent studies do not provide evidence indicating associations larger in magnitude for SO₂-related cardiovascular mortality compared to other mortality outcomes.



APHEA = Air Pollution and Health: A European Approach study; CAPES = China Air Pollution and Health Effects Study; COPD = chronic obstructive pulmonary disease; ISA = Integrated Science Assessment; MISA = meta-analysis of the Italian studies on short-term effects of air pollution; PAPA = Public Health and Air Pollution in Asia.

Total mortality = circle; cardiovascular-related mortality = triangle; and respiratory-related mortality = diamond.

^a(Zmirou et al., 1998) reported on only five of the seven cities included in (Katsouyanni et al., 1997), which had cause-specific mortality data and were included in the analysis.

^b(Chen et al., 2013) examined stroke only in the China Air Pollution and Health Effects Study cities that had stroke data.

^c(Meng et al., 2013) was not part of CAPES, but the four cities included had data for the same years as CAPES.

^d(Kan et al., 2010) reported results which were also presented in (Wong et al., 2008b) and (Wong et al., 2010).

Note: † and red text/circles = recent studies published since the 2008 ISA for Sulfur Oxides; black text/circles = U.S. and Canadian studies evaluated in the 2008 ISA for Sulfur Oxides. Corresponding quantitative results are reported in Supplemental Table 5S-30 (U.S. EPA, 2017c).

Figure 5-18 Percent increase in total, cardiovascular, and respiratory mortality from multicity studies evaluated in the 2008 ISA for Sulfur Oxides (black) and recently published multicity studies (red) for a 10-ppb increase in 24-h avg sulfur dioxide concentrations.

5.5.1.3 Potential Confounding of the Sulfur Dioxide-Mortality Relationship

A limitation of the studies evaluated in the 2008 SO_x ISA, was the relatively sparse analyses of the potential confounding effects of copollutants on the SO₂-mortality relationship ([U.S. EPA, 2008d](#)). The 2008 SO_x ISA specifically stated that the “potential confounding and lack of understanding regarding the interaction of SO₂ with copollutants” was one of the major limitations of the scientific literature that contributed to the conclusion that the evidence is “suggestive of a causal relationship” between short-term SO₂ exposures and mortality. Copollutant analyses conducted in recent studies further attempt to identify whether SO₂ has an independent effect on mortality. In addition to examining potential copollutant confounding, some studies have also examined whether the covariates included in statistical models employed to examine short-term SO₂ exposures and mortality adequately control for the potential confounding effects of season/temporal trends and weather.

Examination of Potential Copollutant Confounding

In the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), the analysis of potential copollutant confounding was limited to studies conducted by [Dominici et al. \(2003\)](#) within the U.S. as part of the National Morbidity Mortality Air Pollution Study (NMMAPS), [Katsouyanni et al. \(1997\)](#) in Europe as part of the APHEA-1 study, [Hoek \(2003\)](#) in the Netherlands, and [Burnett et al. \(2004\)](#) in 12 Canadian cities. Copollutant models in these studies focused on the effect of PM₁₀, BS, or NO₂ on the SO₂-mortality relationship. The SO₂-mortality risk estimate was found to either increase ([Hoek, 2003](#)) or to be slightly attenuated ([Dominici et al., 2003](#); [Katsouyanni et al., 1997](#)) in models with BS or PM₁₀; while risk estimates were reduced, but still remained positive in models with NO₂ ([Burnett et al., 2004](#)). Additionally, there was limited evidence from [Burnett et al. \(2000\)](#) of attenuation of the SO₂ association when PM_{2.5} was included in the model. Recent multicity studies conducted in the U.S. and Asia have also examined whether there is evidence of copollutant confounding; however, similar to the literature base considered in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), the evaluation of copollutant confounding on the SO₂-mortality relationship has remained limited.

In a study of 108 U.S. cities using data from the NMMAPS for 1987–2000 (of which 85 had SO₂ data), [Moolgavkar et al. \(2013\)](#) used a subsampling approach where 4 randomly selected cities were removed from the 108 cities over 5,000 bootstrap cycles to examine associations between short-term air pollution concentrations and total mortality. This approach was used instead of the two-stage Bayesian hierarchical approach employed in

the original NMMAPS analysis, which assumes that city-specific risk estimates are normally distributed around a national mean ([Dominici et al., 2003](#)). In a single-pollutant model using 100 df (~7 df/year, which is consistent with NMMAPS) to control for temporal trends, [Moolgavkar et al. \(2013\)](#) found a 1.5% (95% CI: 1.1, 1.7) increase in total (nonaccidental) mortality at lag 1 for a 10-ppb increase in 24-h avg SO₂ concentrations. In a copollutant analysis, the SO₂-mortality risk estimate remained robust and was similar in magnitude to the single pollutant result upon the inclusion of PM₁₀ [1.3% (95% CI: 0.4, 2.0)]. An analysis of the influence of NO₂ on SO₂-mortality risk estimates was not conducted. The results of [Moolgavkar et al. \(2013\)](#) provide additional support for an SO₂-mortality association, as observed in [Dominici et al. \(2003\)](#), through an analysis that included more cities and used a different statistical approach than previously employed in multicity studies.

Additional multicity studies in Asia conducted more extensive analyses of potential copollutant confounding by examining the effect of gaseous pollutants, in addition to PM₁₀, on the SO₂-mortality relationship. In a study of 17 Chinese cities as part of CAPES, [Chen et al. \(2012b\)](#) examined associations between short-term SO₂ exposures and multiple mortality outcomes. The potential confounding effects of other pollutants on the SO₂-mortality relationship was assessed in copollutant models with PM₁₀ and NO₂. Within the cities examined, SO₂ was found to be moderately correlated with PM₁₀ ($r = 0.49$) and NO₂ ($r = 0.65$). The results from copollutant models ([Table 5-40](#)) indicate that although SO₂ risk estimates remained positive, they were attenuated by approximately 39–54% in models with PM₁₀ and 65–79% in models with NO₂. These results are consistent with those observed in [Chen et al. \(2013\)](#), which focused on stroke mortality in a subset of the CAPES cities (i.e., eight cities) and also reported a similar reduction in SO₂ risk estimates in models with PM₁₀ and NO₂.

Table 5-40 Percent increase in total, cardiovascular, and respiratory mortality for a 10-ppb increase in 24-h avg sulfur dioxide concentrations at lag 0–1 in single and copollutant models.

	Copollutant	Total Mortality % Increase (95% CI)	Cardiovascular Mortality % Increase (95% CI)	Respiratory Mortality % Increase (95% CI)
SO ₂	---	1.98 (1.24, 2.69)	2.19 (1.24, 3.15)	3.31 (2.05, 4.59)
	+PM ₁₀	1.10 (0.45, 1.76)	1.00 (0.08, 1.92)	2.03 (0.89, 3.17)
	+NO ₂	0.42 (-1.56, 1.00)	0.47 (-0.47, 1.42)	1.16 (-0.03, 2.37)

CI = confidence interval; NO₂ = nitrogen dioxide; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; SO₂ = sulfur dioxide.

Source: Adapted from [Chen et al. \(2012b\)](#).

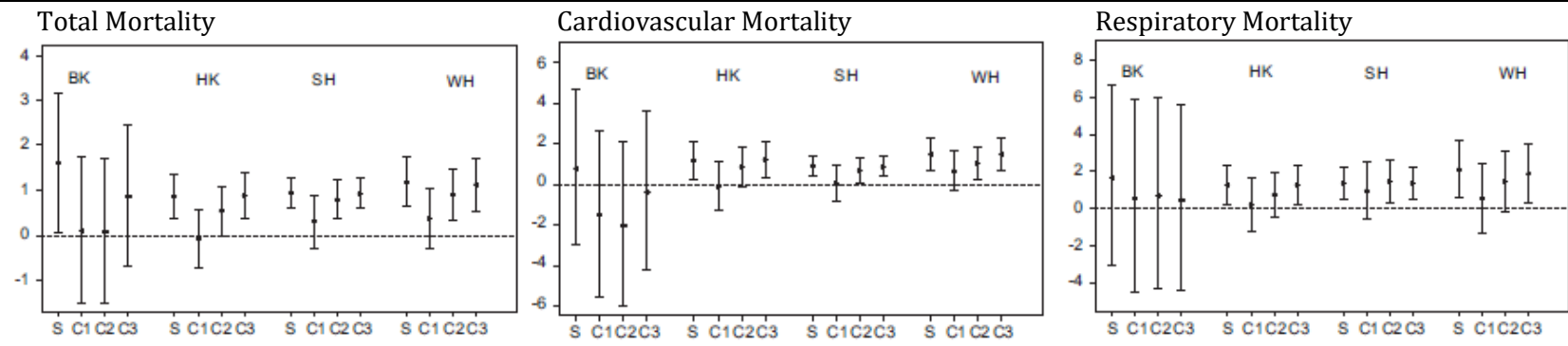
[Kan et al. \(2010\)](#) examined the association between short-term SO₂ exposures and mortality within four Asian cities as part of the PAPA study. Although the authors did not examine copollutant models in a combined four-city analysis, they did on a city-to-city basis. Similar to [Chen et al. \(2012b\)](#), in single pollutant models across cities and mortality outcomes, there was evidence of a consistent positive association ([Figure 5-19](#)). Of note is the highly imprecise estimate (i.e., wide confidence interval) for Bangkok, but it is speculated that the variability in risk estimates for Bangkok could be attributed to the lack of variability in SO₂ concentrations in this city compared to the Chinese cities (standard deviation in SO₂ concentrations of 1.8 ppb; Chinese cities: 4.6–9.7 ppb) ([Kan et al., 2010](#)). Across mortality outcomes and cities, SO₂-mortality risk estimates were attenuated, and in many cases null in copollutant models with NO₂. However, only in Shanghai and Wuhan were SO₂ correlations with NO₂ greater than 0.60 ($r = 0.64$ and 0.76 , respectively). Similarly, SO₂ was also found to be moderately correlated with PM₁₀ in Shanghai ($r = 0.67$) and Wuhan ($r = 0.65$), but SO₂ mortality risk estimates, although attenuated, remained positive across cities. In copollutant models with O₃, SO₂ mortality risk estimates were almost unchanged compared to single-pollutant results.

Recent multicity studies add to the limited number of studies that have examined the potential confounding effects of copollutants on the SO₂-mortality relationship. Within the only recent U.S. study, [Moolgavkar et al. \(2013\)](#) reported that SO₂-mortality risk estimates remained robust in copollutant models with PM₁₀, which is consistent with [Dominici et al. \(2003\)](#), but these studies did not evaluate potential confounding by gaseous pollutants. Studies that examined gaseous pollutants, including [Chen et al. \(2012b\)](#) and [Kan et al. \(2010\)](#) along with [Burnett et al. \(2004\)](#), found that in models with NO₂, SO₂ risk estimates were reduced to a large extent, but remained positive. However,

the overall assessment of copollutant confounding remains limited, and it is unclear how the results observed in Asia translate to other locations, specifically due to the unique air pollution mixture and higher concentrations observed in Asian cities.

Modeling Approaches to Control for Weather and Temporal Confounding

Mortality risk estimates may be sensitive to model specification, which includes the selection of weather covariates to include in statistical models to account for the potential confounding effects of weather in short-term exposure studies. As such, some recent studies have conducted sensitivity analyses to examine the influence of alternative approaches to control for the potential confounding effects of weather on mortality risk estimates.



BK = Bangkok; C1 = sulfur dioxide + nitrogen dioxide; C2 = sulfur dioxide + PM₁₀; C3 = sulfur dioxide + ozone; HK = Hong Kong; S = single-pollutant model; SH = Shanghai; WH = Wuhan.

Source: Figure adapted from [Kan et al. \(2010\)](#). Reprinted with permission of Elsevier.

Figure 5-19 Percent increase in total, cardiovascular, and respiratory mortality associated with a 10 µg/m³ (3.62 ppb) increase in 24-h avg sulfur dioxide concentrations, lag 0–1, in single and copollutants models in Public Health and Air Pollution in Asia cities.

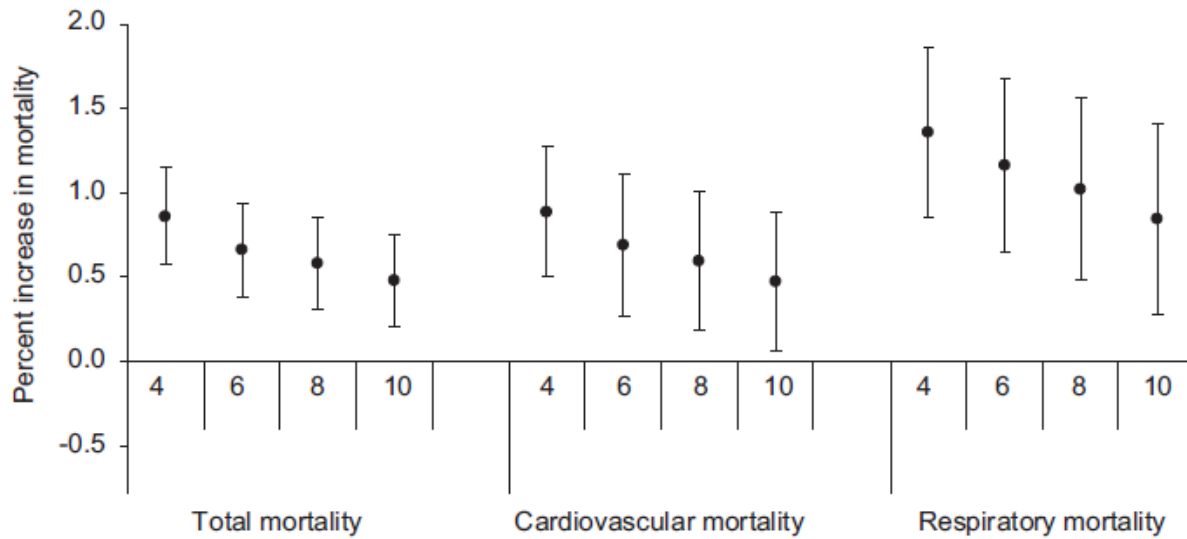
As part of CAPES, [Chen et al. \(2012b\)](#) examined the influence of alternative lag structures for controlling the potential confounding effects of temperature on the SO₂-mortality relationship by varying the lag structure of the temperature variable (i.e., lag 0, lag 0–3, or lag 0–7). The authors found that although the SO₂-mortality associations remained positive and statistically significant across alternative lag structures, risk estimates were attenuated as the number of lag days specified increased. The attenuation observed when using a temperature variable lagged from 0–3 to 0–7 days could be due to [Chen et al. \(2012b\)](#) only including one temperature term in the statistical model. This approach differs from that used in some of the seminal multicity studies (e.g., NMMAPS, APHEA) that include a temperature term averaged over multiple days (e.g., average of lag 1–3 days). A second temperature term is often included in models, in addition to a same-day temperature term, to account for (1) the potential delayed effects of temperature on mortality and (2) potential residual confounding due to temperature.

Temporal

In addition to examining the influence of model specification on mortality risk estimates through the use of alternative weather covariates, recent studies have also examined whether air pollution-mortality risk estimates are sensitive to the df per year employed to control for temporal trends.

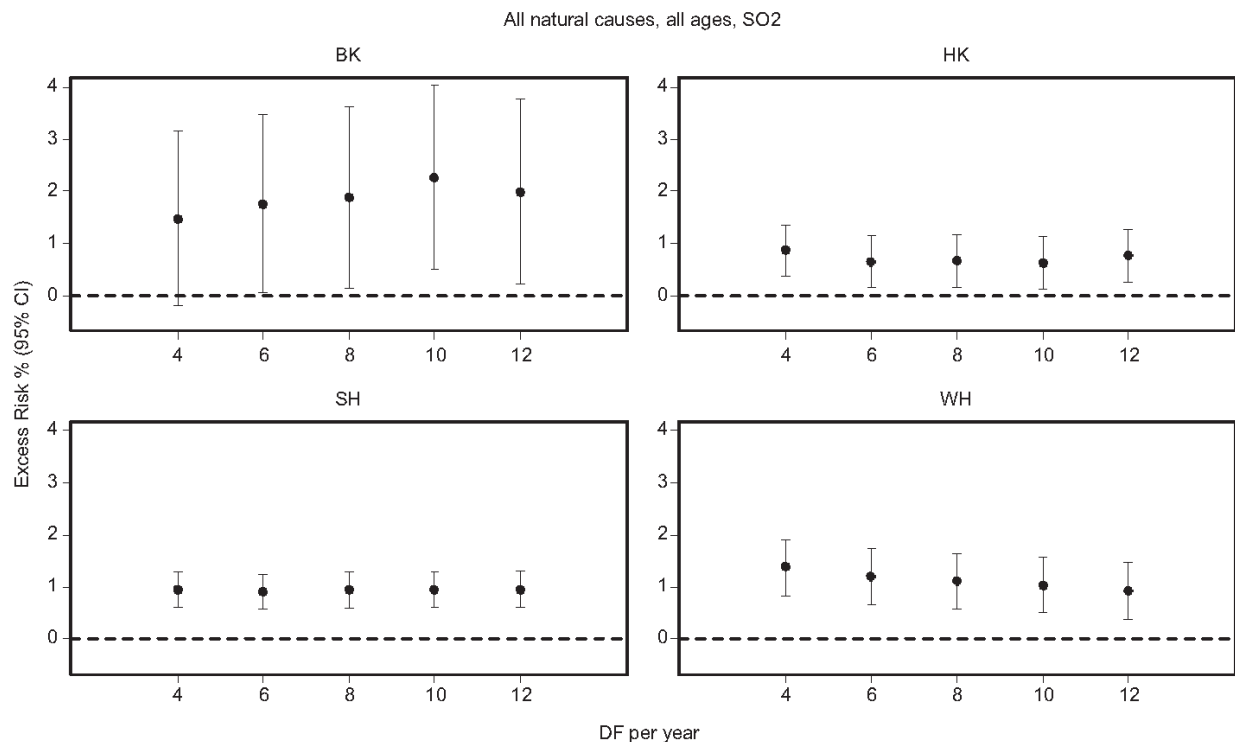
Within CAPES, [Chen et al. \(2012b\)](#) examined the influence of increasing the number of degrees of freedom per year (i.e., 4, 6, 8, and 10 df per year) to control for temporal confounding on SO₂-mortality risk estimates. The authors found that as the number of df per year increased, the percent increase in both total and cause-specific mortality attributed to SO₂ was slightly attenuated but remained positive across the range of df examined ([Figure 5-20](#).)

The results of [Chen et al. \(2012b\)](#) are consistent with those reported by [Kan et al. \(2010\)](#) in an analysis of each individual city within the PAPA study. In models using 4, 6, 8, 10, or 12 df per year, the authors reported relatively similar SO₂-mortality risk estimates across cities. However, as depicted in [Figure 5-20](#), and in some cities in [Figure 5-21](#), using 4 df per year likely leads to inadequate control for temporal trends based on the higher risk estimate observed compared to increasing the degrees of freedom.



Source: (Chen et al., 2012b). Reprinted with permission of Elsevier.

Figure 5-20 Percent increase in daily mortality associated with a $10 \mu\text{g}/\text{m}^3$ (3.62 ppb) increase in 24-h avg sulfur dioxide concentrations at lag 0–1 days using various degrees of freedom per year for time trend, China Air Pollution and Health Effects Study cities, 1996–2008.



BK = Bangkok; CI = confidence interval; DF = degrees of freedom; HK = Hong Kong; SH = Shanghai; SO₂ = sulfur dioxide; WH = Wuhan.

Source: (Kan et al., 2010). Reprinted with permission of Elsevier.

Figure 5-21 Percent increase in total mortality associated with a 10 µg/m³ (3.62 ppb) increase in 24-h avg sulfur dioxide concentrations at lag 0–1 in Public Health and Air Pollution in Asia cities, using different degrees of freedom per year for time trend.

Unlike [Chen et al. \(2012b\)](#) and [Kan et al. \(2010\)](#), which conducted systematic analyses of the influence of increasing the df per year to control for temporal trends on the SO₂-mortality relationship, [Moolgavkar et al. \(2013\)](#) only compared models that used 50 df (~3.5 df per year) or 100 df (~7 df per year). Similar to both [Chen et al. \(2012b\)](#) and [Kan et al. \(2010\)](#), the authors reported relatively similar SO₂-mortality risk estimates in both models [1.6% (95% CI: 0.9, 1.9) for a 10-ppb increase in 24-h avg SO₂ concentrations at lag 1 in the 50-df model and 1.5% (95% CI: 1.1, 1.7) in the 100 df model].

Overall, the studies that examined the effect of alternative approaches to control for the potentially confounding effects of weather and temporal trends report relatively consistent SO₂-mortality risk estimates across models. The results of these studies are further supported by an analysis conducted by [Sacks et al. \(2012\)](#), which examined

whether the different modeling approaches (to control for both weather and temporal trends) used in a number of multicity studies (e.g., NMMAPS, APHEA) resulted in similar risk estimates when using the same data set. In all-year analyses focusing on cardiovascular mortality, SO₂-mortality risk estimates remained relatively stable across models using different weather covariates and a varying number of df per year (ranging from 4 to 8 df per year across models) to control for temporal trends. Although the results of [Sacks et al. \(2012\)](#) are consistent with [Chen et al. \(2012b\)](#), [Kan et al. \(2010\)](#), and [Moolgavkar et al. \(2013\)](#) in all-year analyses, seasonal analyses indicate that differences in model specification may be more important when examining effects by season for some pollutants, such as SO₂.

5.5.1.4 Modification of the Sulfur Dioxide-Mortality Relationship

Individual- and Population-Level Factors

To date, a limited number of studies have examined potential factors that may increase the risk of SO₂-related mortality. In the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), only [Katsouyanni et al. \(1997\)](#) examined potential effect measure modifiers and within the APHEA-2 study reported that geographic location may influence city-specific SO₂-mortality risk estimates. Similar to the 2008 SO_x ISA, only few recent multicity studies [i.e., ([Chen et al. \(2012b\)](#); [Berglund et al. \(2009\)](#); [Wong et al. \(2008b\)](#))] conducted extensive analyses of potential effect measure modifiers of the SO₂-mortality relationship as detailed in [Chapter 6](#). These studies along with some single-city studies focusing on SO₂ and mortality provide limited evidence for potential differences in the risk of SO₂-related mortality by lifestyle, sex, and socioeconomic status (SES).

Season and Weather

A limited number of studies have examined whether there is evidence of seasonal differences or that certain weather patterns modify the SO₂-mortality relationship. In the 2008 SO_x ISA, only [Zmirou et al. \(1998\)](#) examined whether there are seasonal differences in SO₂-mortality risk associations in a subset of the APHEA-1 cities. The authors found some indication of larger associations in the summer months compared to the winter months.

Since the completion of the 2008 SO_x ISA, only a few recent studies have examined whether there are seasonal differences in SO₂-mortality associations, and these studies reported results consistent with [Zmirou et al. \(1998\)](#). In a study of 15 Italian cities, meta-analysis of the Italian studies on short-term effects of air pollution (MISA-2),

[Bellini et al. \(2007\)](#) is the only multicity study that examined whether there were seasonal differences in SO₂-mortality risk estimates. The authors found a similar pattern of associations across mortality outcomes with SO₂-mortality risk estimates being larger in the summer compared to the winter (total mortality: summer 3.2% vs. winter 1.4%; respiratory mortality: summer 12.0% vs. winter 4.1%; cardiovascular mortality: summer 9.4% vs. winter 1.6%). These results are consistent, with the only U.S.-based study that examined seasonal patterns in SO₂-mortality associations. In a study conducted in New York City focusing on cardiovascular mortality, [Ito et al. \(2011\)](#) reported larger risk estimates in the warm season [2.9% (95% CI: -1.2, 7.1)] compared to the cold season [0.0% (95% CI: -1.7, 1.8)] for a 10-ppb increase in 24-h avg SO₂ concentrations.

Instead of examining whether only specific seasons modify the SO₂-mortality association, [Vanos et al. \(2013\)](#) focused on weather patterns, referred to as synoptic weather types, in a study of 10 Canadian cities. Distinct weather types were identified by combining a number of variables including temperature, dew point temperature, sea level pressure, cloud cover, and wind velocity. Across the nine different synoptic weather types examined for SO₂, [Vanos et al. \(2013\)](#) reported that mortality risk estimates in all age analyses tended to be larger in magnitude for dry versus moist weather types, particularly in warmer seasons.

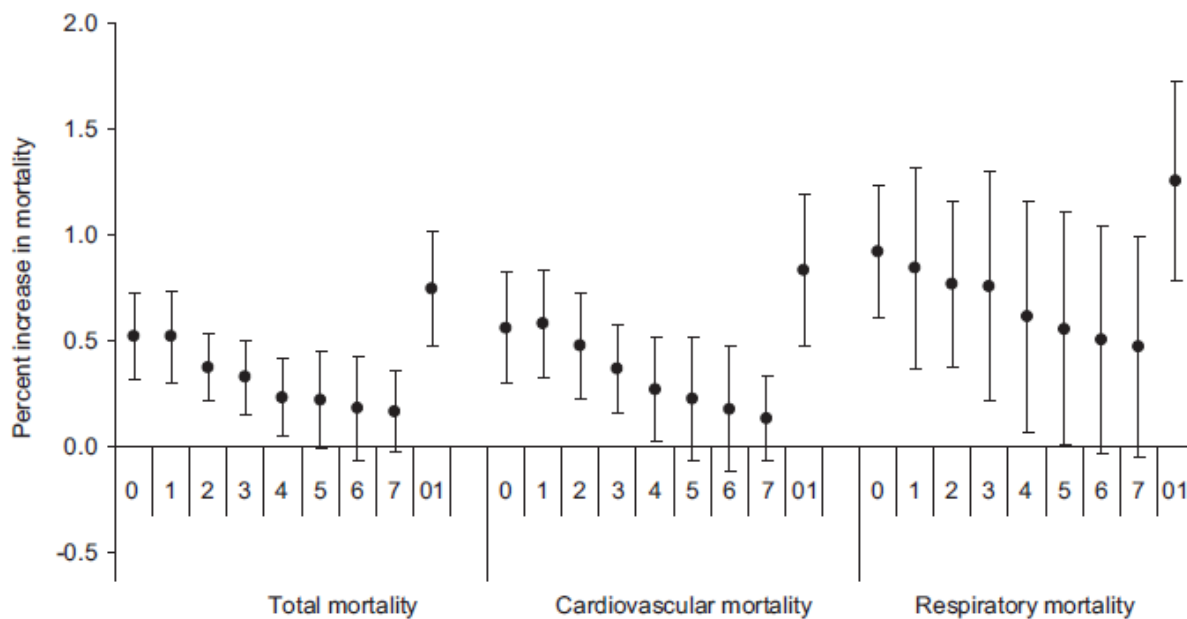
Overall, the limited number of studies that conducted seasonal analyses reported initial evidence indicating larger SO₂-mortality associations during the summer season. Additionally, there is preliminary evidence that specific weather patterns in combination with certain seasons may modify the SO₂-mortality association.

5.5.1.5 Sulfur Dioxide-Mortality Concentration-Response Relationship and Related Issues

Lag Structure of Associations

Of the studies evaluated in the 2008 SO_x ISA, the majority selected lag days a priori and did not extensively examine the lag structure of associations for short-term SO₂ exposures and mortality. These studies primarily focused on single- or multiday lags within the range of 0–3 days. However, in a study in the Netherlands, [Hoek \(2003\)](#) conducted more extensive analyses to examine whether there was evidence of immediate or delayed SO₂-mortality effects. The authors provided preliminary evidence of larger SO₂-mortality risk estimates at a multiday lag of 0–6 days compared to a single-day lag (i.e., lag 1 day). Recent multicity studies have conducted additional analyses further examining the lag structure of associations for short-term SO₂ exposures and mortality.

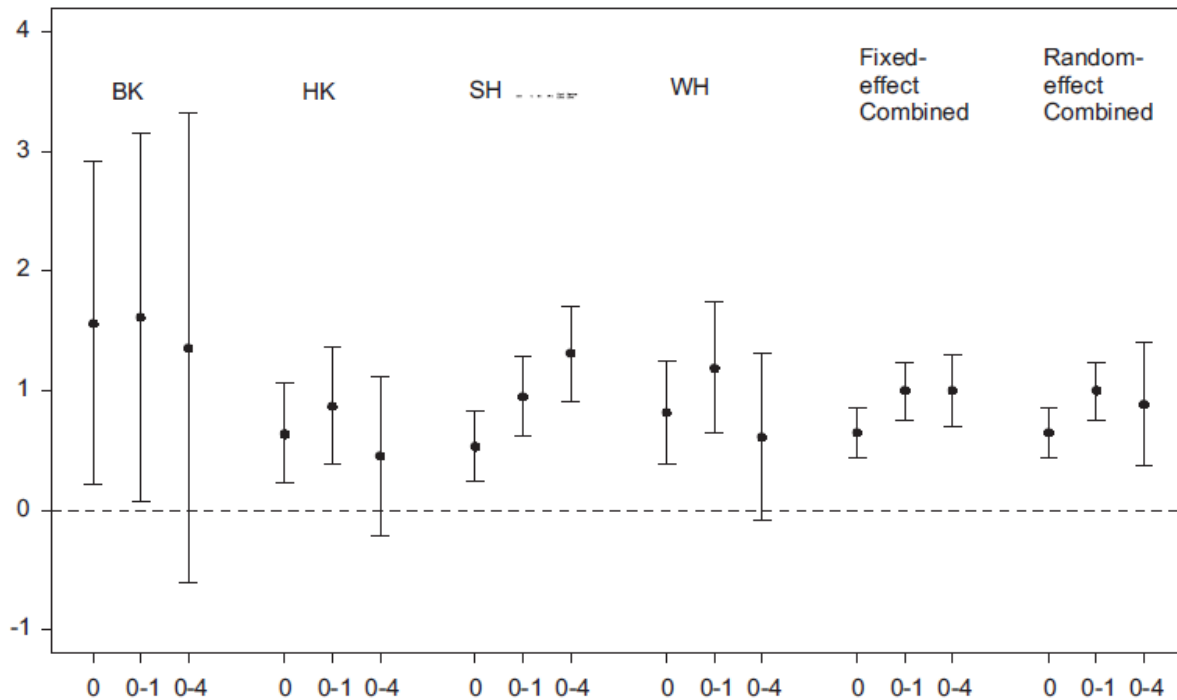
[Chen et al. \(2012b\)](#), within CAPES, examined individual lag days (lag day 0 to 7) and a multiday lag of 0–1 days. As depicted in [Figure 5-22](#), the authors found evidence of immediate SO₂ effects on mortality that slowly declined over time with the multiday lag of 0–1 days exhibiting the largest risk estimate across mortality outcomes.



Source: ([Chen et al., 2012b](#)). Reprinted with permission of Elsevier.

Figure 5-22 Percent increase in daily mortality associated with a 10 µg/m³ (3.62 ppb) increase in 24-h avg sulfur dioxide concentrations, using various lag structures for sulfur dioxide in the China Air Pollution and Health Effects Study cities, 1996–2008.

[Kan et al. \(2010\)](#) also examined the lag structure of associations for the SO₂-mortality relationship within the PAPA study, but did not examine an extensive number of alternative lags, instead focusing on lag 0 and moving averages of 0–1 and 0–4 days ([Figure 5-23](#)). Unlike [Chen et al. \(2012b\)](#), which focused on the combined risk estimate across all cities, [Kan et al. \(2010\)](#) examined the lag structure of associations both within individual cities and in a combined analysis across all PAPA cities. The results of both the individual city and combined analysis are consistent with those observed by [Chen et al. \(2012b\)](#) in CAPES (i.e., the effect largest in magnitude across the lag days examined occurred primarily at lag 0–1 days) ([Figure 5-22](#)).



BK = Bangkok; HK = Hong Kong; SH = Shanghai; WH = Wuhan.
 Source: [Kan et al. \(2010\)](#). Reprinted with permission of Elsevier.

Figure 5-23 Percent increase in total mortality associated with a 10 µg/m³ (3.62 ppb) increase in 24-h avg sulfur dioxide concentrations for different lag structures in individual Public Health and Air Pollution in Asia cities and in combined four city analyses.

[Bellini et al. \(2007\)](#) took a slightly different approach to examining the lag structure of associations in a study of 15 Italian cities (MISA-2) by focusing on whether there was evidence of mortality displacement. The authors reported larger SO₂-mortality effects at lag 0–15 days (3.8% for a 10-ppb increase in 24-h avg SO₂ concentrations) compared to a lag of 0–1 days (1.6%), which supports no evidence of mortality displacement. Additional information on the lag structure can be observed by examining the percent increase in mortality associated with short-term SO₂ exposures at each individual lag day of the lag 0–15-day model. The individual lag day results remained positive up to approximately lag day 10, which is consistent with the results from [Chen et al. \(2012b\)](#) ([Figure 5-22](#)). However, examining associations at single-day lags over a week, such as 10 days, may be uninformative due to potential inadequate control for weather variables at these longer durations. Additionally, these longer lags may not be biologically plausible due to controlled human exposure and animal toxicological studies

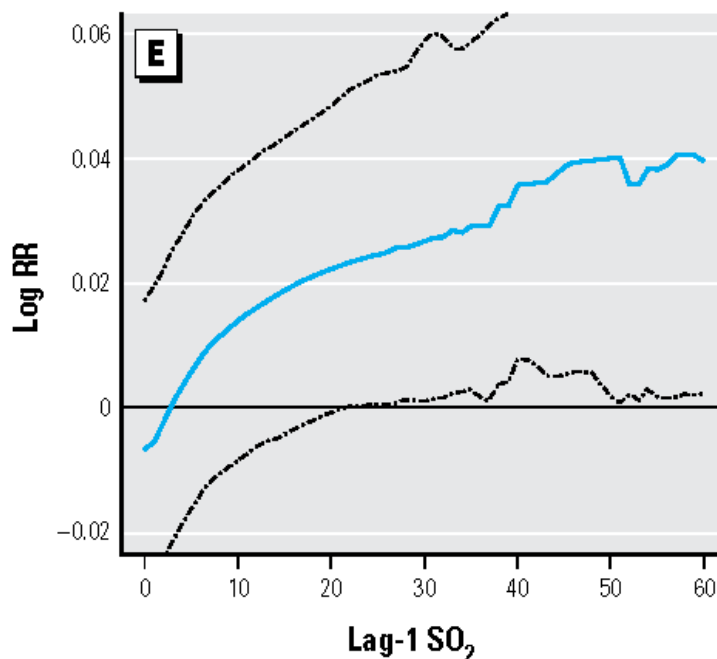
demonstrating that effects attributed to SO₂ exposure are rather immediate ([Section 5.2.1.2](#)).

Overall, the limited analyses that have examined the lag structure of associations for short-term SO₂ exposures and mortality suggest that the greatest effects occur soon after exposure (lag 0–1). However, the studies evaluated indicate that positive associations may persist longer although the magnitude of those effects diminishes over time.

Concentration-Response Relationship

The studies evaluated in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)), as well as prior assessments, have not conducted formal analyses of the SO₂-mortality C-R relationship. Although limited in number, a few recent studies published since the completion of the 2008 SO_x ISA have conducted analyses to examine the shape of the SO₂-mortality C-R relationship and whether a threshold exists in the combined C-R relationship across multiple cities, or in an evaluation of single-city C-R relationships in the context of a multicity study. However, these studies have not conducted extensive analyses examining alternatives to linearity in the shape of the SO₂-mortality C-R relationship.

Using a subsampling approach, [Moolgavkar et al. \(2013\)](#) examined the shape of the C-R relationship between short-term air pollution exposures and mortality in the NMMAPS data set by applying a nonlinear function (i.e., natural splines with 6 df) to each pollutant. As demonstrated in [Figure 5-24](#), the analysis conducted by [Moolgavkar et al. \(2013\)](#) provides support for a linear, no threshold relationship between short-term SO₂ exposures and total mortality.



RR = relative risk; SO₂ = sulfur dioxide.

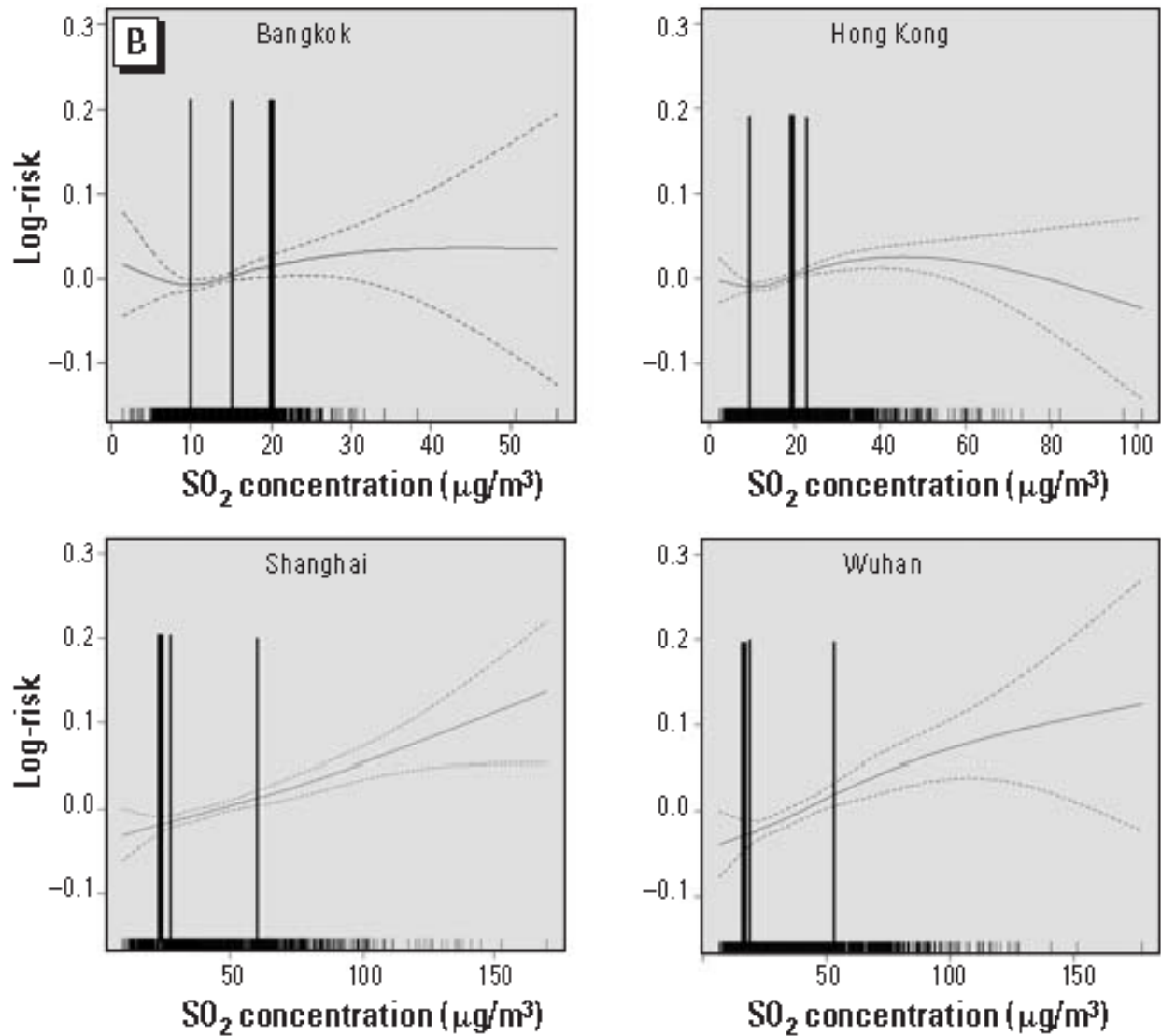
Note: Pointwise means and 95% confidence intervals adjusted for size of the bootstrap sample ($d = 4$).

Source: Reprinted from Environmental Health Perspectives; [Moolgavkar et al. \(2013\)](#).

Figure 5-24 Flexible ambient concentration-response relationship between short-term sulfur dioxide (ppb) exposure (24-h avg concentrations) and total mortality at lag 1.

In the four-city PAPA study, [Kan et al. \(2010\)](#) also examined the SO₂-mortality C-R relationship, but only focused on the shape of the C-R curve in each individual city. The C-R curve for the SO₂-mortality relationship was assessed by applying a natural spline smoother with 3 df to SO₂ concentrations. To examine whether the SO₂-mortality relationship deviates from linearity, the deviance between the smoothed (nonlinear) pollutant model and the unsmoothed (linear) pollutant model was examined. When examining the deviance, the authors only reported evidence for potential nonlinearity in Hong Kong. However, across the cities, there is evidence of a linear, no threshold, relationship within the range of SO₂ concentrations where the data density is the highest, specifically within the IQR ([Figure 5-25](#)). The linear relationship is most pronounced in Shanghai and Wuhan, with evidence of an inverted U-shape for Bangkok and Hong Kong. It should be noted, there is an overall lack of confidence in the shape of the C-R curve at the high end of the distribution of SO₂ concentrations in Bangkok and Shanghai due to the lower data density within this range of concentrations observed in both cities. A difficulty apparent in comparing the results across cities within [Kan et al. \(2010\)](#) is the

drastically different range of SO₂ concentrations in Bangkok and Hong Kong compared with Shanghai and Wuhan. However, the cities with similar distributions of SO₂ concentrations also have similar shapes to their respective SO₂-mortality C-R curves.



SO₂ = sulfur dioxide.

Note: x-axis is the average of lag 0–1 24-h avg SO₂ concentrations (µg/m³). Solid lines indicate the estimated mean percent change in daily mortality, and the dotted lines represent twice the standard error. Thin vertical lines represent the interquartile range of SO₂ concentrations within each city, while the thin vertical bar represents the World Health Organization guideline of 20 µg/m³ for a 24-h avg time of SO₂.

Source: Reprinted from Environmental Health Perspectives; (Wong et al., 2008b).

Figure 5-25 Concentration-response curves for total mortality (degrees of freedom = 3) for sulfur dioxide in each of the four Public Health and Air Pollution in Asia cities.

Both [Moolgavkar et al. \(2013\)](#) and [Kan et al. \(2010\)](#) examined the shape of the SO₂-mortality C-R relationship by focusing on all-cause (total) mortality. Additional information on the shape of the C-R curve can be assessed in studies that focused on cause-specific mortality as discussed in [Section 5.2.1.8](#) (respiratory mortality) and [Section 5.3.1.9](#) (cardiovascular mortality). In studies of multiple Chinese cities, [Meng et al. \(2013\)](#) and [Chen et al. \(2013\)](#) examined the shape of the C-R relationship for mortality and short-term air pollution exposures on COPD and stroke mortality, respectively. In both studies the authors conducted similar analyses of linearity by examining the deviance between linear and spline models. [Meng et al. \(2013\)](#) and [Chen et al. \(2013\)](#) both found no evidence of a deviation in linearity in the SO₂-COPD mortality and SO₂-stroke mortality relationship, respectively ([Figures 5-11](#) and [5-16](#)).

To date, studies have conducted a rather limited exploration of potential alternatives to linearity when examining the shape of the C-R relationship, which in combination with the potential measurement error due to the spatial and temporal variability in SO₂ concentrations ([Sections 3.4.2.2](#) and [3.4.2.3](#)), complicates the interpretation of the SO₂-mortality C-R relationship. With these limitations in mind, studies that examined the C-R relationship provide initial evidence that indicates a log-linear, no threshold relationship between short-term SO₂ concentrations and mortality, specifically within the range of SO₂ concentrations where the data density is highest. Some differences in the shape of the curve were observed on a city-to-city basis, which is consistent with the mortality C-R results that have been reported for other criteria air pollutants.

5.5.1.6 Summary and Causal Determination

Recent multicity studies evaluated since the completion of the 2008 SO_x ISA continue to provide consistent evidence of positive associations between short-term SO₂ exposures and total mortality. Although the body of evidence is larger, key uncertainties and data gaps still remain, which contribute to the conclusion that the evidence for short-term SO₂ exposures and total mortality is suggestive of, but not sufficient to infer, a causal relationship. This conclusion is consistent with that reached in the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Recent multicity studies evaluated have further informed key uncertainties and data gaps in the SO₂-mortality relationship identified in the 2008 SO_x ISA including confounding, modification of the SO₂-mortality relationship, potential seasonal differences in SO₂-mortality associations, and the shape of the SO₂-mortality C-R relationship. However, questions remain regarding whether SO₂ has an independent effect on mortality, which can be attributed to: (1) the limited number of studies that examined potential copollutant confounding, (2) the relative lack of copollutant analyses with PM_{2.5}, (3) and the evidence indicating attenuation of

SO₂-mortality associations in copollutant models with NO₂ and PM₁₀. Additionally, all of the studies evaluated averaged SO₂ concentrations over multiple monitors and used a 24-h avg exposure metric when assigning exposure, which may not adequately capture the spatial and temporal variability in SO₂ concentrations ([Sections 3.4.2.2](#), and [3.4.2.3](#)). While correlations between 24-h avg and 1-h max SO₂ concentrations are high ($r > 0.75$) at most monitors, lower correlations may occur at some monitors and in individual studies, which can add uncertainty to the ability of 24-h avg metrics to capture peak SO₂ concentrations. This section describes the evaluation of evidence for total mortality, with respect to the causal determination for short-term exposures to SO₂ using the framework described in Table II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)). The key evidence, as it relates to the causal framework, is summarized in [Table 5-41](#).

Table 5-41 Summary of evidence, which is suggestive of, but not sufficient to infer, a causal relationship between short-term sulfur dioxide exposure and total mortality.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Consistent epidemiologic evidence from multiple, high quality studies at relevant SO ₂ concentrations	Increases in mortality in multicity studies conducted in the U.S., Canada, Europe, and Asia	Section 5.5.1.2 Figure 5-15	Mean 24-h avg: U.S., Canada, South America, Europe: 0.4–28.2 ^d ppb Asia: 0.7–>200 ppb Table 5-39
Uncertainty regarding potential confounding by copollutants	The magnitude of SO ₂ associations remained positive, but were reduced in copollutant models with PM ₁₀ and NO ₂ . No studies examined copollutant models with PM _{2.5} . SO ₂ generally exhibits low to moderate correlations with other NAAQS pollutants at collocated monitors, and attenuation of SO ₂ -mortality association may be a reflection of spatial variability among the pollutants.	Sections 5.5.1.3, 3.4.3	
Uncertainty regarding exposure measurement error	U.S. studies that examine the association between short-term SO ₂ exposures and mortality rely on single or the average of multiple monitors in an area and SO ₂ generally has low to moderate spatial correlations across urban geographical scales	Section 3.4.2.2	

Table 5-41 (Continued): Summary of evidence, which is suggestive of, but not sufficient to infer, a causal relationship between short term sulfur dioxide exposure and total mortality.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Uncertainty due to limited coherence and biological plausibility with cardiovascular and respiratory morbidity evidence	Generally supportive, but not entirely consistent epidemiologic evidence for ischemic events such as triggering a myocardial infarction. Inconclusive epidemiologic and experimental evidence for other cardiovascular endpoints. Uncertainties with respect to the independent effect of SO ₂ on cardiovascular effects contributing to limited coherence and biological plausibility for SO ₂ -related cardiovascular mortality, which comprises ~35% of total mortality ^e	Section 5.3.1.11 Table 5-31	
	Consistent evidence of asthma exacerbations from controlled human exposure studies demonstrating respiratory effects (i.e., respiratory symptoms and decreased lung function) in response to typically 5–10-min exposures, with generally supportive evidence from short-term SO ₂ exposure epidemiologic studies demonstrating asthma-related morbidity, specifically hospital admissions and ED visits. Uncertainty as to the biological mechanism that explains the continuum of effects leading to SO ₂ -related respiratory mortality, which comprises ~8% of total mortality ^e	Section 5.2.1.9 Table 5-21	

Avg = average; ED = emergency department; NAAQS = National Ambient Air Quality Standards; NO₂ = nitrogen dioxide; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 μm; PM_{2.5} = in general terms, particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 μm; SO₂ = sulfur dioxide.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)).

^bDescribes the key evidence and references, supporting or contradicting, contributing most heavily to causal determination and, where applicable, to uncertainties or inconsistencies. References to earlier sections indicate where full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated.

^dThe value of 28.2 represents the median concentration from [Katsouyanni et al. \(1997\)](#).

^eStatistics taken from [American Heart Association \(2011\)](#).

Collectively, the evidence from recent multicity studies of short-term SO₂ exposures and mortality consistently demonstrate positive SO₂-mortality associations in single-pollutant models. In the limited number of studies that conducted copollutant analysis, correlations between SO₂ and other pollutants were low ($r < 0.4$) to moderate ($r = 0.4-0.7$). Although SO₂-mortality associations remain positive in copollutant models with PM₁₀ and NO₂, they were often attenuated to a large degree, questioning the independent effect of SO₂ on mortality. However, SO₂ is more spatially variable than other pollutants as reflected in the generally low to moderate spatial correlations across urban geographical scales ([Section 3.4.2.2](#)); therefore, the attenuation in SO₂ associations in copollutant models may be a reflection of the different degree of exposure error across pollutants ([Section 3.4.3](#)). It is important to note, the majority of recent studies that examined potential copollutant confounding have been conducted in Asian countries where correlations between pollutants may be higher, possibly limiting the generalizability of results to other study areas where SO₂ concentrations along with the concentrations of other air pollutants are much lower. This is reflected in the results of [Moolgavkar et al. \(2013\)](#) in a U.S. multicity study where there was very little evidence of attenuation of the SO₂-mortality association in copollutant models with PM₁₀, whereas the multicity studies conducted in Asian cities showed a rather pronounced reduction in SO₂ associations. In addition to copollutant analyses, recent studies examined the influence of the extent of temporal adjustment and the lag structure for weather covariates on the SO₂-mortality association. When examining, the extent of temporal adjustment, multiple studies reported similar SO₂-mortality associations across a range of degrees of freedom per year. Only [Chen et al. \(2012b\)](#) examined the lag structure for weather covariates, specifically temperature, and found evidence of a difference in SO₂-mortality associations as the number of lag days increased, but this could be attributed to the analysis being based on only one covariate for temperature.

An examination of factors that may contribute to increased risk of SO₂-related mortality, as discussed in [Chapter 6](#), found evidence indicating that older adults (≥ 65 years of age) may be at increased risk with very limited evidence of potential differences by sex and socioeconomic status. In the 2008 SO_x ISA, initial evidence suggested potential seasonal differences in SO₂-mortality associations, particularly in the summer months. A recent multicity study conducted in Italy along with single-city studies conducted in the U.S. add to this initial body of evidence suggesting larger associations during the summer or warm months. Preliminary evidence indicates that not only season, but season in combination with specific weather patterns, may modify the SO₂-mortality association. Additionally, an examination of different modeling approaches provides evidence that the magnitude of the seasonal association may depend on the modeling approach employed to control for the potential confounding effects of weather ([Sacks et al., 2012](#)).

Those studies that examined the lag structure of associations for the SO₂-mortality relationship generally observed that there is evidence of an immediate effect (i.e., lag 0 to 1 days) of short-term SO₂ exposures on mortality. Multicity studies conducted in the U.S. and Asia have examined the shape of the C-R relationship and whether a threshold exists in both a multi- and single-city setting. These studies have used different statistical approaches and consistently demonstrated a linear relationship with no evidence of a threshold within the range of SO₂ concentrations where the data density is highest. The evidence of linearity in the SO₂-mortality C-R relationship is further supported by studies of cause-specific mortality as detailed in [Section 5.3.1.9](#) (respiratory mortality) and [Section 5.2.1.9](#) (cardiovascular mortality). However, to date, studies have not conducted extensive analyses exploring alternatives to linearity when examining the shape of the SO₂-mortality C-R relationship.

Overall, recent epidemiologic studies build upon and support the conclusions of the 2008 SO_x ISA for total mortality. However, the biological mechanism that could lead to mortality as a result of short-term SO₂ exposures has not been clearly characterized. This is evident when evaluating the underlying health effects (i.e., cardiovascular effects in [Section 5.3](#) and respiratory effects in [Section 5.2](#)) that could lead to cardiovascular (~35% of total mortality) and respiratory (~9% of total mortality) mortality, the components of total mortality most thoroughly evaluated ([Hoyert and Xu, 2012](#)). For cardiovascular effects, the evidence is “inadequate to infer a causal relationship” with exposure to short-term SO₂ concentrations. An evaluation of epidemiologic studies that examined the relationship between short-term SO₂ exposure and cardiovascular effects found a number positive associations but the evidence was not entirely consistent. Within the collective body of evidence for cardiovascular effects, important uncertainties remain especially regarding disentangling whether there is an independent effect of SO₂ on cardiovascular effects, which is the same uncertainty in total mortality studies. Overall, this evidence complicates the interpretation of the relationship between SO₂ and cardiovascular mortality.” For respiratory effects the evidence indicates a causal relationship for short-term SO₂ exposures. The strongest evidence for respiratory effects is from studies examining SO₂-related asthma exacerbations, specifically controlled human exposure studies demonstrating respiratory effects (i.e., respiratory symptoms and decreased lung function) ([Section 5.2.1.2](#)) in people with asthma in response to short-term, generally 5–10-minutes, SO₂ exposures. The results from controlled human exposure studies are generally supported by epidemiologic studies reporting respiratory-related morbidity including hospital admissions and ED visits, specifically for asthma. However, the biological mechanism that explains the continuum of effects that could lead to respiratory-related mortality remains unclear. Additionally, it is important to note epidemiologic studies that examine the association between short-term SO₂ exposures and mortality rely on single or the average of multiple monitors over an area to

assign exposure. Therefore, the exposure assessment approach used in the mortality studies may contribute to exposure measurement error and underestimate associations observed due to the spatially heterogeneous distribution of SO₂ concentrations over a wide area ([Section 3.4.2.2](#)). In conclusion, the consistent positive associations observed across various multicity studies is limited by the uncertainty due to whether SO₂ is independently associated with total mortality, the representativeness of monitors and the 24-h avg SO₂ exposure metric in capturing the spatial and temporal variability in exposure to SO₂ ([Sections 3.4.2.2](#) and [3.4.2.3](#)), and the uncertainty in the biological mechanism that could lead to SO₂-induced mortality ([Section 4.3](#)). **Collectively, this body of evidence is suggestive, but not sufficient to conclude there is a causal relationship between short-term SO₂ exposure and total mortality.**

5.5.2 Long-Term Exposure

In past reviews, a limited number of epidemiologic studies assessed the relationship between long-term exposure to SO₂ and mortality in adults. The 2008 SO_x ISA concluded that the scarce amount of evidence was “inadequate to infer a causal relationship” ([U.S. EPA, 2008d](#)). The 2008 SO_x ISA identified concerns as to whether the observed associations were due to SO₂ alone, or if sulfate or other particulate SO_x, such as H₂SO₄, or PM indices could have contributed to these associations. The possibility that the observed effects may not be due to SO₂, but other constituents that come from the same source as SO₂, or that PM may be more toxic in the presence of SO₂ or other components associated with SO₂, could not be ruled out. Overall, a lack of consistency across studies, inability to distinguish potential confounding by copollutants, and uncertainties regarding the geographic scale of analysis limited the interpretation of the causal relationship between long-term exposure to SO₂ and mortality. Consistent with the conclusion of the 2008 SO_x ISA, the collective evidence informing the association between long-term SO₂ exposure and mortality continues to be limited. Despite the improved consistency of the associations between long-term exposure to SO₂ and both respiratory and total mortality with the inclusion of recent cohort studies, these studies do not address uncertainties identified in the 2008 SO_x ISA.

This section includes a review of the evidence for an association between long-term exposure to SO₂ and mortality, integrating evidence presented in previous NAAQS reviews with evidence that is newly available to this review. The evidence in this section will focus on epidemiologic studies because experimental studies of long-term exposure and mortality are generally not conducted. However, this section will draw from the morbidity evidence presented for different health endpoints across the scientific disciplines (i.e., animal toxicological, controlled human exposure studies, and

epidemiology) to support the association observed for cause-specific mortality. Studies are discussed by geographic region, with U.S. studies discussed in [Section 5.5.2.1](#), European studies in [Section 5.5.2.2](#), and Asian studies in [Section 5.5.2.3](#). [Section 5.5.2.4](#) describes studies that evaluated the SO₂-mortality relationship over small geographic scales. A brief summary of the studies included in these sections can be found in [Table 5-42](#).

Table 5-42 Summary of studies of long-term exposure and mortality.

Study	Location Years	Mean SO ₂ ppb	Exposure Assessment	Correlation with Other Pollutants	Selected Effect Estimates (95% CI) ^a
†Hart et al. (2011) Cohort: N = 53,814 Deaths (all-causes): 4,806 Deaths (respiratory): 317 Deaths (COPD): 209 Deaths (lung cancer): 800	U.S. (SO ₂ : 1985–2000; follow-up: 1985–2000)	4.8	Annual average exposures based on residential address from model using spatial smoothing and GIS-based covariates; current calendar year and long-term average from 1985–2000		All causes: 1.09 (1.03, 1.15) Respiratory: 1.10 (0.89, 1.35) COPD: 0.93 (0.71, 1.22) Lung cancer: 1.11 (0.98, 1.27)
Krewski et al. (2000) HSC Cohort: N = 8,111 Deaths (all-causes, HSC): 1,239 ACS Cohort: N = 559,049 Deaths (all-causes, ACS): 43,361	U.S. HSC: (SO ₂ : 1977–1985; follow-up: 1974–1991) ACS: (SO ₂ : 1980; follow-up: 1982–1989)	HSC: 1.6–24.0 ACS: 9.3	HSC: mean levels from fixed-site monitors ACS: City-specific annual mean	HSC: PM _{2.5} : 0.85 SO ₄ : 0.85 NO ₂ : 0.84	All causes: HSC: 1.05 (1.02, 1.09) ACS: 1.06 (1.05, 1.07) Lung cancer: HSC: 1.03 (0.91, 1.16)
Pope et al. (2002) Cohort: N = 539,000 Deaths: NR	U.S. (SO ₂ : 1982–1998; follow-up: 1982–1998)	6.7–9.7	Average across monitoring stations in each metropolitan area for each study year using daily average (i.e., 24-h avg) concentrations, averaged over 1 yr (1980) and the entire study period (1982–1998)		All causes: 1.03 (1.02, 1.05)
†Lipfert et al. (2009) Cohort: N = 67,938 Deaths (all-causes): 44,653	U.S. (SO ₂ : 1999; follow-up: 1976–2001)	4.3	County-level estimates from AER plume-in-grid air quality model; based on 1999 emissions inventory from point and area sources for 36 x 36-km grid squares	Subject-weighted: EC: 0.68 NO _x : 0.65 SO ₄ ²⁻ : 0.79	All causes: 1.02 (1.01, 1.03)

Table 5-42 (Continued): Summary of studies of long term exposure and mortality.

Study	Location Years	Mean SO ₂ ppb	Exposure Assessment	Correlation with Other Pollutants	Selected Effect Estimates (95% CI) ^a
†Krewski et al. (2009) Cohort: N = 513,450 Deaths: NR	U.S. (SO ₂ : 1980; follow-up: 1982–2000)	9.6	City-specific annual mean		All causes: 1.02 (1.02, 1.03) Lung cancer: 1.00 (0.98, 1.02)
Lipfert et al. (2006a) Cohort: N = 70,000 Deaths: NR	U.S. (SO ₂ : 1999–2001; follow-up: 1997–2001)	16.3	County-level “peak” concentrations	Subject-weighted: PM _{2.5} : 0.71 NO ₂ : 0.41 Peak O ₃ : 0.21 Peak CO: 0.41 SO ₄ ²⁻ : 0.77 OC: 0.34 EC: -0.13	All causes: 0.99 (0.97, 1.01)
Abbey et al. (1999) Cohort: N = 6,338 Deaths (all-causes): 1,575 Deaths (lung cancer): 30	U.S. (SO ₂ : 1966–1992; follow-up: 1977–1992)	5.6 IQR: 3.7	ZIP code-level mo averages cumulated and averaged over time	Mean concentration: PM ₁₀ : 0.31 O ₃ : 0.09 SO ₄ : 0.68 When exceeding 100 ppb (O ₃) or 100 µg/m ³ (PM ₁₀) PM ₁₀ : -0.05 O ₃ : 0.13	All causes: Men: 1.07 (0.92, 1.25) Women: 1.00 (0.88, 1.14) Lung cancer: Men: 2.52 (1.34, 4.77) Women: 4.40 (2.34, 8.33)
Beelen et al. (2008b) Cohort: N = 120,852 Deaths (all-causes): 17,610 Deaths (respiratory): 1,016 Deaths (lung cancer): 1,888	Netherlands (SO ₂ : 1976–1985, 1987–1996; follow-up: 1987–1996)	5.2 SD: 1.9	IDW to regional background monitors at baseline residential address		All causes: 0.94 (0.80, 1.10) Respiratory: 0.92 (0.64, 1.31) Lung cancer: 0.99 (0.73, 1.35)
Nafstad et al. (2004) Cohort: N = 16,209 Deaths (all causes): 4,227 Deaths (respiratory): 200 Deaths (lung cancer): 382	Norway (SO ₂ : 1974–1995; follow-up: 1972–1998)	3.6	Model results (per square kilometer) for some year/urban locations, supplemented with background monitoring data		All causes: 0.97 (0.95, 1.01) Respiratory: 1.04 (0.91, 1.19) Lung cancer: 1.00 (0.91, 1.11)
Filleul et al. (2005) Cohort: N = 14,284 Deaths (all causes): 2,396	France (SO ₂ : 1974–1976; follow-up: 1974–2000)	3.0–8.2	3-yr mean concentrations for 24 areas in seven different cities	BS: 0.29 TSP: 0.17 NO -0.01 NO ₂ -0.10	All causes: 1.01 (0.99, 1.04) Lung cancer: 0.99 (0.90, 1.09)

Table 5-42 (Continued): Summary of studies of long term exposure and mortality.

Study	Location Years	Mean SO ₂ ppb	Exposure Assessment	Correlation with Other Pollutants	Selected Effect Estimates (95% CI) ^a
† Bentayeb et al. (2015) Cohort: N = 20,327 Deaths (all-causes): 1,967 Deaths (respiratory): 284 Deaths (CVD): 165	France (SO ₂ : 1989–2008; follow-up: 1989–2013)	2.3	Annual concentrations from CHIMERE chemical-transport model, 2-km resolution	O ₃ : -0.13 PM _{2.5} : 0.58 PM ₁₀ : 0.57 PM _{10-2.5} : 0.30 NO ₂ : 0.56	All causes: 1.23 (0.98, 1.52) Respiratory: 0.76 (0.43, 1.33) CVD: 0.85 (0.44, 1.67)
† Hansell et al. (2016) Cohort: N = 367,658 Deaths (all causes): 47,775 Deaths (respiratory): 5,300 Deaths (COPD): 2,413 Deaths (lung cancer): 3,154 Deaths (CVD): 23,923	England (SO ₂ : 1971, 1981, 1991; follow-up: 1971–2009)	1971: 32.4 1981: 16.4 1991: 11.2	LUR models for annual concentrations in 1971, 1981, and 1991		1991 All causes: 1.09 (1.05, 1.15) Respiratory: 1.20 (1.09, 1.33) COPD: 1.43 (1.23, 1.66) Lung cancer: 1.29 (1.12, 1.47) CVD: 1.05 (0.99, 1.13)
† Carey et al. (2013) Cohort: N = 823,442 Deaths (all causes): 81,636 Deaths (respiratory): 10,408 Deaths (lung cancer): 5,192	England (SO ₂ : 2002; follow-up: 2003–2007)	1.5 SD: 0.8 IQR: 0.8	Annual mean for 1-km grid cells from air dispersion models (poor validation results for SO ₂)	PM ₁₀ : 0.45 NO ₂ : 0.37 O ₃ : -0.41	All causes: 1.26 (1.19, 1.34) Respiratory: 1.67 (1.42, 1.97) Lung cancer: 1.34 (1.06, 1.58)

Table 5-42 (Continued): Summary of studies of long term exposure and mortality.

Study	Location Years	Mean SO ₂ ppb	Exposure Assessment	Correlation with Other Pollutants	Selected Effect Estimates (95% CI) ^a
† Ancona et al. (2015) Cohort: N = 85,559 Deaths (all causes): 5,878 Deaths (CVD): 2,240 Deaths (respiratory): 384	Rome, Italy (SO _x : 2001–2010; follow-up: 2001–2010)	2.5 µg/m ³ SO _x SD: 0.9	Lagrangian particle dispersion model (SPRAY Ver. 5) used SO _x as exposure marker for petrochemical refinery emissions	PM ₁₀ : 0.81 H ₂ S: 0.78	All causes: Men: 1.04 (0.92, 1.18) Women: 0.93 (0.81, 1.07) CVD: Men: 1.08 (0.89, 1.31) Women: 1.00 (0.81, 1.25) IHD: Men: 1.05 (0.79, 1.41) Women: 1.25 (0.89, 1.75) Respiratory: Men: 1.31 (0.88, 1.95) Women: 0.64 (0.32, 1.28)
† Cao et al. (2011) Cohort: N = 70,947 Deaths (all causes): 8,319	China (SO ₂ : 1991–2000; follow-up: 1991–2000)	27.7	Annual average by linking fixed-site monitoring data with residential ZIP code		All causes: 1.02 (1.02, 1.03) CVD: 1.02 (1.00, 1.03) Respiratory: 1.04 (1.02, 1.06) Lung cancer: 1.06 (1.03, 1.08)
† Chen et al. (2016) Cohort: N = 39,054 Deaths (all causes): 1,353 Deaths (lung cancer): 140	China (SO ₂ : 1998–2009; follow-up: 1998–2009)	25.5	1-yr avg and time-varying exposure from monitoring stations calculated from 24-h avg		Lung cancer: 1.02 (1.01, 1.03)
† Dong et al. (2012) Cohort: N = 9,941 Deaths (all causes): 505 Deaths (respiratory): 72	China (SO ₂ : 1998–2009; follow-up: 1998–2009)	23.9 SD: 5.7	1-yr avg from five monitors		Respiratory: 1.05 (0.96, 1.16)
† Zhang et al. (2011) Cohort: N = 9,941 Deaths (all causes): 256	Shenyang, China (SO ₂ : 1998–2009; follow-up: 1998–2009)	23.9	1-yr avg and yearly deviations in each of five monitoring stations calculated from 24-h avg		All causes: 0.93 (0.90, 0.99)

Table 5-42 (Continued): Summary of studies of long term exposure and mortality.

Study	Location Years	Mean SO ₂ ppb	Exposure Assessment	Correlation with Other Pollutants	Selected Effect Estimates (95% CI) ^a
† Katanoda et al. (2011) Cohort: N = 63,520 Deaths (all causes): 6,687 Deaths (respiratory): 690 Deaths (lung cancer): 518	Japan (SO ₂ : 1974–1983; follow-up: 1983–1995)	2.4–19.0	Annual mean concentrations from monitoring station near each of eight study areas	Pearson: SPM: 0.47	Respiratory: 1.20 (1.15, 1.24) COPD: 1.15 (0.94, 1.41) Pneumonia: 1.20 (1.16, 1.25) Lung cancer: 1.12 (1.03, 1.22)
Elliott et al. (2007) Cohort: N = 662,343 Deaths (all causes): 52,792 Deaths (respiratory): 8,471 Deaths (lung cancer): 3,473	Great Britain (SO ₂ : 1966–1970, 1990–1994; follow-up: 1982–1986, 1994–1998)	12.2–41.4	4-yr exposure windows from annual average concentrations from monitoring sites located in residential areas		All causes: 1.02 (1.02, 1.02) Respiratory: 1.06 (1.06, 1.07) Lung cancer: 1.00 (0.99, 1.01)
† Bennett et al. (2014) Ecologic study at Ward level	Warwickshire, U.K. (SO ₂ : 2010; mortality data: 2007–2012)	NR	Single recorded level for each ward from 2010		Heart failure: 1.11 (0.988, 1.22)
† Wang et al. (2009) No individual data	Brisbane, Australia (SO ₂ : 1996–2004; follow-up: 1996–2004)	5.4	1-h max from 13 monitoring stations aggregated to annual means used with IDW		Cardiopulmonary: 1.26 (1.03, 1.54)
† Wang et al. (2014a) Hypothetical cohort	China (SO ₂ : 2004–2010; life table: 2010)	46.31	Annual average across monitoring stations in 85 city regions		Life expectancy: 10-µg/m ³ increase in SO ₂ correlated with 0.28–0.47 yr decrease in life expectancy

ACS = American Cancer Society; AER = Atmospheric and Environmental Research; avg = average; BS = black smoke; CHIMERE = regional chemistry transport model; CI = confidence interval; CO = carbon monoxide; COPD = chronic obstructive pulmonary disease; CVD = cardiovascular disease; EC = elemental carbon; GIS = geographic information systems; H₂S = hydrogen sulfide; HSC = Harvard Six Cities; IDW = inverse distance weighting; IHD = ischemic heart disease; IQR = interquartile range; LUR = land use regression; max = maximum; ; n = sample size; N = population number; NO = nitric oxide; NO₂ = nitrogen dioxide; NO_x = the sum of NO and NO₂; NR = not reported; O₃ = ozone; OC = organic carbon; PM_{2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 2.5 µm; PM₁₀ = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm; PM_{10-2.5} = particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm and greater than a nominal 2.5 µm; SD = standard deviation; SO₂ = sulfur dioxide; SO₄ = sulfate; SO₄²⁻ = sulfate ion; SO_x = sulfur oxide; SPM = suspended particulate matter; TSP = total suspended solids.

^aEffect estimates are standardized per 5-ppb increase in SO₂ concentrations.

†Studies published since the 2008 ISA for Sulfur Oxides.

5.5.2.1 U.S. Cohort Studies

A number of longitudinal cohort studies have been conducted in the U.S. and have found small, statistically significant positive associations between long-term exposure to SO₂ and total mortality ([Hart et al., 2011](#); [Lipfert et al., 2009](#); [Pope et al., 2002](#); [Krewski et al., 2000](#)). The body of evidence is smaller and less consistent when these studies examine cause-specific mortality, although [Hart et al. \(2011\)](#) observed positive, yet imprecise associations (i.e., wide 95% confidence intervals) with respiratory, lung cancer, and cardiovascular mortality. In the Trucking Industry Particle Study, [Hart et al. \(2011\)](#) used the work records for over 50,000 men employed in four U.S. trucking companies to identify all-cause and cause-specific mortality. Occupational exposures were assigned based on job title, while exposures to ambient air pollution (i.e., PM₁₀, SO₂, and NO₂ averaged over the study period) were determined using spatial smoothing and geographic information system (GIS)-based covariates using residential address. All three pollutants were independently associated with all-cause mortality, with central estimates the highest for the association with NO₂ and lowest for the association with PM₁₀. Both NO₂ and SO₂ were positively associated with lung cancer, cardiovascular disease, and respiratory disease mortality, and negatively associated with COPD mortality. Correlation coefficients between SO₂ and other measured air pollutants were not reported, making it difficult to evaluate for the potential of copollutant confounding on the associations attributed to SO₂. There was no evidence of confounding by occupational exposures (based on job-title).

The Harvard Six Cities study is a prospective cohort study of the effects of air pollution with the main focus on PM components in six U.S. cities and provides limited evidence for an association between mortality and exposure to SO₂. Cox proportional hazards regression was conducted with data from a 14- to 16-year follow-up study of 8,111 adults in the six cities. [Dockery et al. \(1993\)](#) reported that lung cancer and cardiopulmonary mortality were more strongly associated with the concentrations of inhalable and fine PM and sulfate particles than with the levels of TSP, SO₂, NO₂, or acidity of the aerosol. [Krewski et al. \(2000\)](#) conducted a sensitivity analysis of the Harvard Six Cities study and examined associations between gaseous pollutants (i.e., O₃, NO₂, SO₂, and CO) and mortality, observing positive associations between SO₂ and total mortality and cardiopulmonary deaths. In this data set SO₂ was highly correlated with PM_{2.5} ($r = 0.85$), sulfate ($r = 0.85$), and NO₂ ($r = 0.84$), making it difficult to attribute the observed associations to an independent effect of SO₂.

[Pope et al. \(1995\)](#) investigated associations between long-term exposure to PM and the mortality outcomes in the American Cancer Society (ACS) cohort and provides limited

evidence for an association between exposure to SO₂ and mortality. Ambient air pollution data from 151 U.S. metropolitan areas in 1981 were linked with individual risk factors in 552,138 adults who resided in these areas when enrolled in the prospective study in 1982. Death outcomes were ascertained through 1989. Gaseous pollutants were not analyzed in the original analysis. Extensive reanalysis of the ACS data, augmented with additional gaseous pollutants data, showed positive associations between mortality and SO₂, but not for the other gaseous pollutants ([Jerrett et al., 2003](#); [Krewski et al., 2000](#)). [Pope et al. \(2002\)](#) extended analysis of the ACS cohort with double the follow-up time (to 1998) and triple the number of deaths compared to the original study ([Pope et al., 1995](#)). Both PM_{2.5} and SO₂ were associated with all the mortality outcomes, although only SO₂ was associated with the deaths attributable to “all other causes.” The association of SO₂ with mortality for “all other causes” makes it difficult to interpret the effect estimates due to a lack of biological plausibility for this association. More recently, [Krewski et al. \(2009\)](#) conducted an extended reanalysis of the study conducted by [Pope et al. \(2002\)](#), including examination of ecologic covariates (e.g., education attainment, housing characteristics, income) and evaluation of exposure windows. The inclusion of ecologic covariates generally resulted in increased risk estimates, with the greatest effect on mortality from IHD. The authors also evaluated individual time-dependent exposure profiles to examine whether there is a critical exposure time window most strongly associated with mortality from ambient air pollution. The time window immediately preceding death (1–5 years) produced the strongest effects for mortality associated with exposure to SO₂, while later time windows (6–10 years and 11–15 years) generally showed null associations between SO₂ and mortality.

[Lipfert et al. \(2000\)](#) conducted an analysis of a national cohort of ~70,000 male U.S. military veterans who were diagnosed as hypertensive in the mid-1970s and were followed up for about 21 years (up to 1996) and provides scant evidence for an association between exposure to SO₂ and mortality. This cohort was 35% black and 57% of the cohort were current smokers (81% of the cohort had been smokers at one time). PM_{2.5}, PM₁₀, particulate matter with a nominal aerodynamic diameter less than or equal to 10 µm and greater than a nominal 2.5 µm (PM_{10-2.5}), TSP, sulfate, CO, O₃, NO₂, SO₂, and lead (Pb) were examined in these analyses. The county of residence at the time of entry to the study was used to estimate exposures. Four exposure periods (from 1960 to 1996) were defined, and deaths during each of the three most recent exposure periods were considered. The results for SO₂ as part of their preliminary screening were generally null. [Lipfert et al. \(2000\)](#) noted that Pb and SO₂ were not found to be associated with mortality, and thus, were not considered further. They also noted that the pollution effect estimates were sensitive to the regression model specification, exposure periods, and the inclusion of ecological and individual variables. The authors reported that indications of concurrent mortality risks were found for NO₂ and peak O₃. In a subsequent analysis,

[Lipfert et al. \(2006b\)](#) examined associations between traffic density and mortality in the same cohort, extending the follow-up period to 2001. As in their previous study ([Lipfert et al., 2000](#)), four exposure periods were considered but included more recent years, and reported that traffic density was a better predictor of mortality than ambient air pollution variables with the possible exception of O₃. The log-transformed traffic density variable was only weakly correlated with SO₂ ($r = 0.32$) and PM_{2.5} ($r = 0.50$) in this data set. [Lipfert et al. \(2006a\)](#) further extended analysis of the veterans' cohort data to include the U.S. EPA's Speciation Trends Network (STN) data, which collected chemical components of PM_{2.5}. They analyzed the STN data for 2002, again using county-level averages. PM_{2.5} and gaseous pollutant data for 1999 through 2001 were also analyzed. As in the previous study ([Lipfert et al., 2006b](#)), traffic density was the most important predictor of mortality, but associations were also observed for elemental carbon, vanadium, nickel, and nitrate. Ozone, NO₂, and PM₁₀ also showed positive but weaker associations. Once again, no associations were observed between long-term exposure to SO₂ and mortality. [Lipfert et al. \(2009\)](#) re-examined these associations, this time averaging the exposure variables over the entire follow-up period (1976–2001). For this exposure period, they observed positive associations between SO₂ and mortality. When the data set was stratified by county-level traffic density, the SO₂ association with mortality was stronger in the counties with high density traffic, and attenuated to near null in the counties with lower traffic density. The fact that the association between long-term exposure to SO₂ and mortality is only observed in areas where traffic density has been characterized as high, along with the moderate to strong correlations between SO₂ and other traffic-related pollutants (e.g., PM_{2.5}, NO₂, NO_x, EC) in these analyses, makes it difficult to discern whether these associations are truly attributable to SO₂, or could be due to some other traffic-related pollutant or mixture of pollutants.

[Abbey et al. \(1999\)](#) investigated associations of long-term ambient concentrations of PM₁₀, sulfate, SO₂, O₃, and NO₂ with mortality in a cohort of 6,338 nonsmoking California Seventh-Day Adventists. Monthly indices of ambient air pollutant concentrations at 348 monitoring stations throughout California were interpolated to ZIP codes according to home or work location of study participants, cumulated, and then averaged over time. They reported associations between PM₁₀ and total mortality for males and nonmalignant respiratory mortality for both sexes. SO₂ was positively associated with total mortality for males but not for females. Generally, null associations were observed for cardiopulmonary deaths and respiratory mortality for both males and females.

Overall, the majority of the limited evidence informing the association between long-term exposure to SO₂ and mortality from U.S. cohort studies was included in the 2008 SO_x ISA. A recent cohort study of male truck drivers ([Hart et al., 2011](#)) provided some

additional evidence for an association between long-term exposure to SO₂ and both respiratory mortality and total mortality, while updates to the ACS ([Krewski et al., 2009](#)) and Veterans ([Lipfert et al., 2009](#)) cohort studies provides some limited evidence for an association with total mortality, although none of these recent studies help to resolve the uncertainties identified in the 2008 SO_x ISA related to copollutant confounding or the geographic scale of the analysis.

5.5.2.2 European Cohort Studies

A number of European cohort studies examined the association between SO₂ concentrations and both total mortality and cause-specific mortality and found generally inconsistent results. [Beelen et al. \(2008b\)](#) analyzed data from the Netherlands Cohort Study on Diet and Cancer with 120,852 subjects. Traffic-related pollutants (BS, NO₂, SO₂, PM_{2.5}), and four types of traffic-exposure estimates were analyzed. While the local traffic component was estimated for BS, NO₂, and PM_{2.5}, no such attempt was made for SO₂. Traffic intensity on the nearest road was associated with all-cause mortality and a larger RR was observed for respiratory mortality. Results were similar for BS, NO₂, and PM_{2.5}, but no associations were observed for SO₂.

Several studies noting declining SO₂ concentrations during the follow-up period (from the mid-1970s through the mid-1990s) did not observe positive associations with mortality. [Nafstad et al. \(2004\)](#) linked data from 16,209 males (aged 0 to 49 years) living in Oslo, Norway with data from the Norwegian Death Register and with estimates of the average annual air pollution levels at the participants' home addresses. PM was not considered in this study because measurement methods changed during the study period. Exposure estimates for NO_x and SO₂ were constructed using models based on subject addresses, emission data for industry, heating, and traffic, and measured concentrations. While NO_x was associated with total, respiratory, lung cancer, and ischemic heart disease deaths, SO₂ did not show any associations with mortality. In this study, SO₂ levels were reduced by a factor of 7 during the study period (from 5.6 ppb in 1974 to 0.8 ppb in 1995), whereas NO_x did not show any clear downward trend. [Filleul et al. \(2005\)](#) linked daily measurements of SO₂, TSP, BS, NO₂, and nitric oxide (NO) with data on mortality for 14,284 adults who resided in 24 areas from seven French cities enrolled in the Air Pollution and Chronic Respiratory Diseases survey in 1974. Models were run before and after exclusion of six area monitors influenced by local traffic as determined by an NO:NO₂ ratio of >3. Before exclusion of the six areas, none of the air pollutants was associated with mortality outcomes. After exclusion of these areas, analyses showed associations between total mortality and TSP, BS, NO₂, and NO but not SO₂ or acidimetric measurements. In this study, SO₂ levels declined by a factor of two to three

(depending on the city) between the 1974 through 1976 period and the 1990 through 1997 period. The changes in air pollution levels over the study period complicate interpretation of reported effect estimates.

[Carey et al. \(2013\)](#) examined the associations between long-term exposure to ambient air pollutants and total and cause-specific mortality in a national English cohort (n = 835,607). The authors used air dispersion models to estimate annual mean air pollution concentrations for 1-km grid cells for a single year prior to the follow-up period. Model validation using national air quality monitors and networks demonstrated good agreement for NO₂ and O₃, moderate agreement for PM₁₀ and PM_{2.5}, but relatively poor agreement for SO₂ ($R^2 = 0-0.39$). The authors observed positive associations with total mortality for all of the air pollutants, and these associations were stronger for PM_{2.5}, NO₂, and SO₂ with respiratory and lung cancer mortality. Associations were generally not observed between cardiovascular mortality and any of the pollutants. Although the authors observed positive associations between SO₂ and mortality (especially respiratory mortality), these associations are difficult to interpret due to the poor validation of the dispersion model for SO₂. [Ancona et al. \(2015\)](#) used a Lagrangian particle dispersion model (see [Section 3.3.2.4](#) for details) to estimate annual means of SO_x (as an exposure marker for emissions from a petrochemical refinery) in Rome, Italy and associations with all-cause and cause-specific mortality among men and women. The authors did not present any validation results for their dispersion model. Predicted concentrations of SO_x were highly correlated with predicted concentrations of PM₁₀ ($r = 0.81$), and because SO_x was used as an exposure marker for petrochemical refinery emissions, it would likely be correlated with other stack or fugitive refinery emissions, including PM_{2.5} and VOCs. The authors observed associations for all-cause mortality and CVD mortality that were near the null value for both men and women. When restricted to IHD mortality, the association remained near the null value for men, but was elevated among women. Conversely, slightly increased risks were observed for respiratory mortality and mortality due to digestive diseases among men, while the risks for these were attenuated among women. Due to the unknown validity of the dispersion model and the high correlations with additional copollutants, it is difficult to interpret these associations.

Overall, the results of the European cohort studies provide very little evidence for an association between long-term exposure to SO₂ and mortality. The majority of these studies were included in the 2008 SO_x ISA ([Beelen et al., 2008b](#); [Filleul et al., 2005](#); [Nafstad et al., 2004](#)). Only the study by [Carey et al. \(2013\)](#) provided new evidence for this review. None of the studies used copollutant models or accounted for potential confounding or effect measure modification by other ambient air pollutants, including sulfate. The study by [Carey et al. \(2013\)](#) had the potential to inform uncertainties related to the geographic scale of the exposure assessment; however, the poor validation results

of the dispersion model used to estimate the SO₂ concentrations for 1-km grid cells makes it difficult to interpret these results.

5.5.2.3 Asian Cohort Studies

Four recent cohort studies have been conducted in China to examine the association between long-term exposure to SO₂ and mortality ([Chen et al., 2016](#); [Dong et al., 2012](#); [Cao et al., 2011](#); [Zhang et al., 2011](#)) and observed inconsistent results. Each of these studies used annual area-wide average concentrations from fixed-site monitoring stations to assign exposure. Notably, the mean SO₂ concentrations in these study areas were much higher than concentrations observed in other locations (see [Table 5-42](#)). [Cao et al. \(2011\)](#) observed generally modest positive associations with all-cause, respiratory and lung cancer mortality. [Chen et al. \(2016\)](#) observed a positive association with lung cancer mortality, although the correlation between SO₂ and PM₁₀ was high ($r > 0.94$), and it is possible that copollutant confounding could at least partially explain this relationship. [Dong et al. \(2012\)](#) observed a modest, positive association with respiratory mortality, while [Zhang et al. \(2011\)](#) observed modest negative associations with all-cause mortality.

[Katanoda et al. \(2011\)](#) conducted a cohort study in Japan investigating the association between long-term exposure to PM_{2.5}, NO₂, and SO₂ and lung cancer and respiratory mortality. The authors used annual mean concentrations from fixed-site monitoring stations near each of eight study areas. The authors observed positive associations of long-term exposure to PM_{2.5}, NO₂, and SO₂ with lung cancer and respiratory mortality, with the strongest effect observed for the SO₂ associations.

Overall, these recent Asian cohort studies provide some new evidence of an association between long-term exposure to SO₂ and mortality; however, they generally report similar associations for other ambient air pollutants, and do not evaluate for potential bias due to copollutant confounding (using copollutants models, reporting correlation coefficients between SO₂ and other measured pollutants, or other methods). Generally, these recent studies do not help to resolve the uncertainties identified in the 2008 SO_x ISA related to copollutant confounding or the geographic scale of the analysis.

5.5.2.4 Cross-Sectional Analysis Using Small Geographic Scale

[Elliott et al. \(2007\)](#) examined associations of BS and SO₂ with mortality in Great Britain using a cross-sectional analysis. However, unlike the earlier ecological cross-sectional mortality analyses in the U.S. in which mortality rates and air pollution levels were compared using large geographic boundaries (i.e., metropolitan statistical areas or

counties), [Elliott et al. \(2007\)](#) compared the mortality rates and air pollution concentrations using a much smaller geographic unit, the electoral ward, with a mean area of 7.4 km² and a mean population of 5,301 per electoral ward. Of note, SO₂ levels declined from 41.4 ppb in the 1966 to 1970 period to 12.2 ppb in 1990 to 1994. This type of analysis does not allow adjustments for individual risk factors, but the study did adjust for socioeconomic status data available for each ward from the 1991 census. Social deprivation and air pollution were more highly correlated in the earlier exposure windows. They observed positive associations for both BS and SO₂ with mortality outcomes. The estimated effects were stronger for respiratory illness than other causes of mortality for the most recent exposure period and most recent mortality period (when pollution levels were lower). The adjustment for social deprivation reduced the effect estimates for both pollutants. Simultaneous inclusion of BS and SO₂ reduced effect estimates for BS but not SO₂. [Elliott et al. \(2007\)](#) noted that the results were consistent with those reported in the [Krewski et al. \(2000\)](#) reanalysis of the ACS study. Similarly, [Bennett et al. \(2014\)](#) observed a positive association between ward-level SO₂ concentrations measured in 2010 and ward-level data on heart failure mortality from 2007–2012 in Warwickshire, U.K. Stronger associations were observed for estimated benzene exposure in this population, while estimated PM exposure was inversely associated with heart failure mortality. These analyses are ecological, but the exposure estimates in the smaller area compared to that in the U.S. cohort studies may have resulted in less exposure measurement error, and the large underlying population appears to be reflected in the narrow confidence bands of effect estimates.

In a recent cross-sectional analysis, [Wang et al. \(2009\)](#) examined the long-term exposure to gaseous air pollutants (i.e., NO₂, O₃, and SO₂) and cardio-respiratory mortality in Brisbane, Australia. Pollutant concentrations were estimated for small geographic units, statistical local areas, using inverse distance weighting. The authors observed a positive association between cardio-respiratory mortality and SO₂, but generally null associations for NO₂ and O₃.

The results of these cross-sectional studies are inconsistent, with much higher mortality effects attributed to SO₂ in Brisbane, Australia ([Wang et al., 2009](#)) and Warwickshire, U.K. ([Bennett et al., 2014](#)) than throughout Great Britain ([Elliott et al., 2007](#)). While each of these studies took a geospatial approach to their analyses, the cross-sectional nature of the study designs and the lack of control for potential bias due to copollutant confounding limit the interpretation of their results.

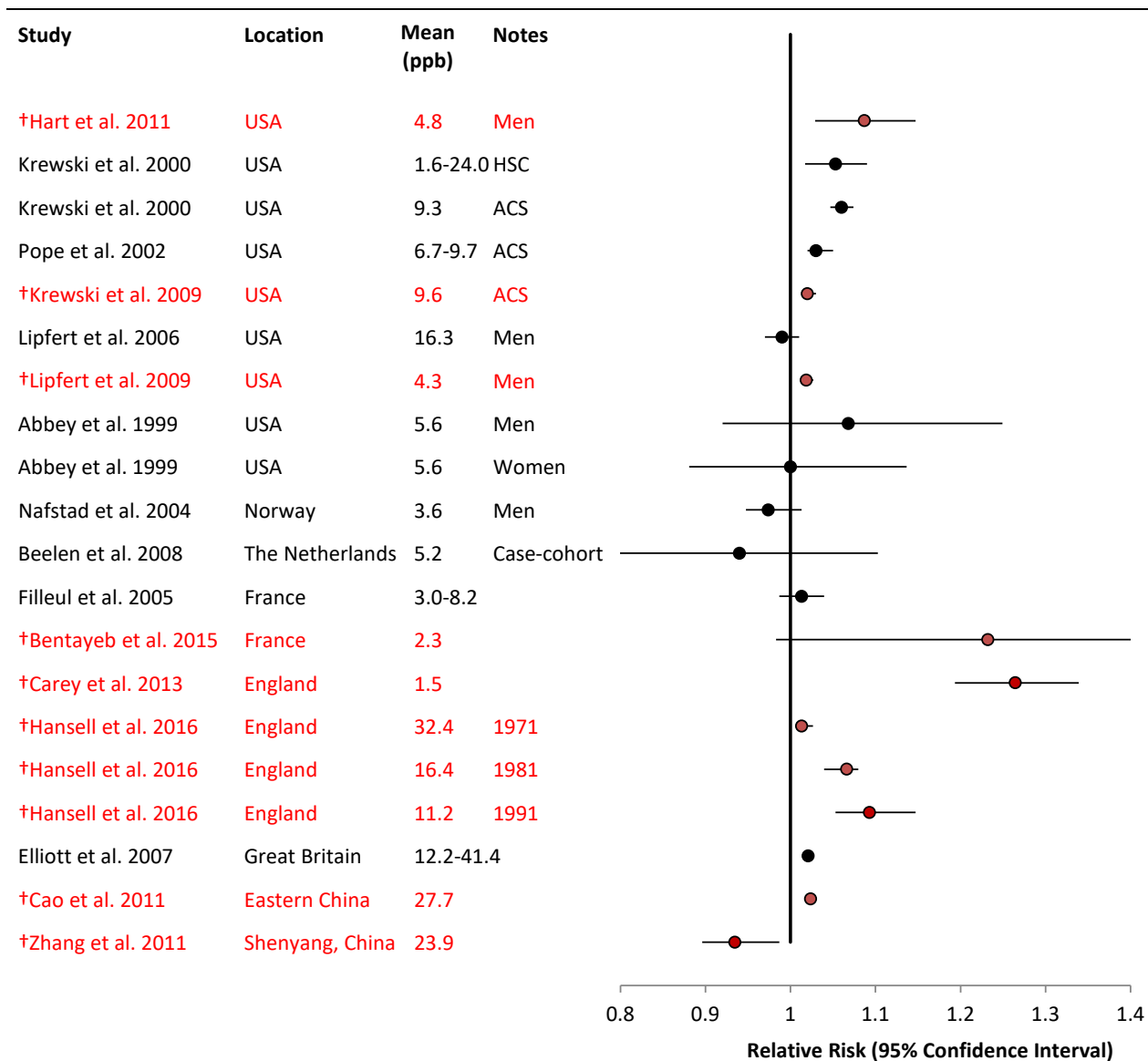
5.5.2.5 Summary and Causal Determination

Overall, the majority of the limited evidence informing the association between long-term exposure to SO₂ and mortality was included in the 2008 SO_x ISA. The 2008 SO_x ISA identified concerns regarding (1) the consistency of the observed associations; (2) whether the observed associations were due to SO₂ alone, or if sulfate or other particulate SO_x or PM indices could have contributed to these associations; and (3) the geographic scale of the exposure assessment. Specifically, the 2008 SO_x ISA noted the possibility that the observed effects may not be due to SO₂, but other co-occurring pollutants that come from the same source as SO₂, or that PM may be more toxic in the presence of SO₂ or other components associated with SO₂, could not be ruled out. None of the epidemiologic studies made corrections or adjustments for exposure measurement error or accounted for the potential for bias away from the null, the potential for which has been demonstrated in simulation studies (see [Section 3.4.4.2](#)). A lack of consistency across studies, inability to distinguish potential confounding by copollutants, and uncertainties regarding the geographic scale of analysis limited the interpretation of the causal relationship between long-term exposure to SO₂ and mortality.

[Figure 5-26](#) presents total mortality effect estimates associated with long-term exposure to SO₂. The overall range of effects spans 0.93 to 1.26 per 5-ppb increase in the annual (or longer period) average SO₂ concentration. The analyses of the Harvard Six Cities and the ACS cohort data, which likely provide effect estimates that are most useful for evaluating possible health effects in the U.S., observed effect estimates of 1.02 to 1.06, while the effect estimate from the recent cohort study of truck drivers was 1.09. Note that each of the U.S. cohort studies has its own advantages and limitations. The Harvard Six Cities data have a small number of exposure estimates, but the study cities were carefully chosen to represent a range of air pollutant exposures. The ACS cohort had far more subjects, but the population was more highly educated than the representative U.S. population. Because educational status appeared to be an important effect modifier of air pollution effects in both studies, the overall effect estimate for the ACS cohort may not be generalizable to the broader population. The evidence from the cohort studies conducted in Europe and Asia is generally similar to that observed from the U.S. cohort studies. That is, the magnitude of the effect estimates is generally similar, although there is greater inconsistency in the direction of the association. Also, the effect estimate observed by [Carey et al. \(2013\)](#) is much higher than that observed in any of the other studies. Generally, these results are consistent with a recent study ([Wang et al., 2014a](#)) that evaluated the correlation between life expectancy and SO₂ concentrations in 85 major city regions in China. After accounting for a surrogate for socioeconomic status, they observed that city regions with higher SO₂ concentrations were correlated with lower life expectancies.

[Figure 5-27](#) presents the cause-specific mortality effect estimates associated with long-term exposure to SO₂. The overall range of effects spans 0.93 to 4.40 per 5-ppb increase in the annual (or longer period) average SO₂ concentration. Generally, there was a trend toward more positive associations for respiratory and lung cancer mortality compared to cardiopulmonary, cardiovascular, and other causes of death. Specifically, recent studies examining respiratory mortality provide some evidence that this cause of death may be more consistently associated with long-term exposure to SO₂ than other causes of death. This is consistent with both the short- and long-term exposure to SO₂ that are associated with respiratory effects.

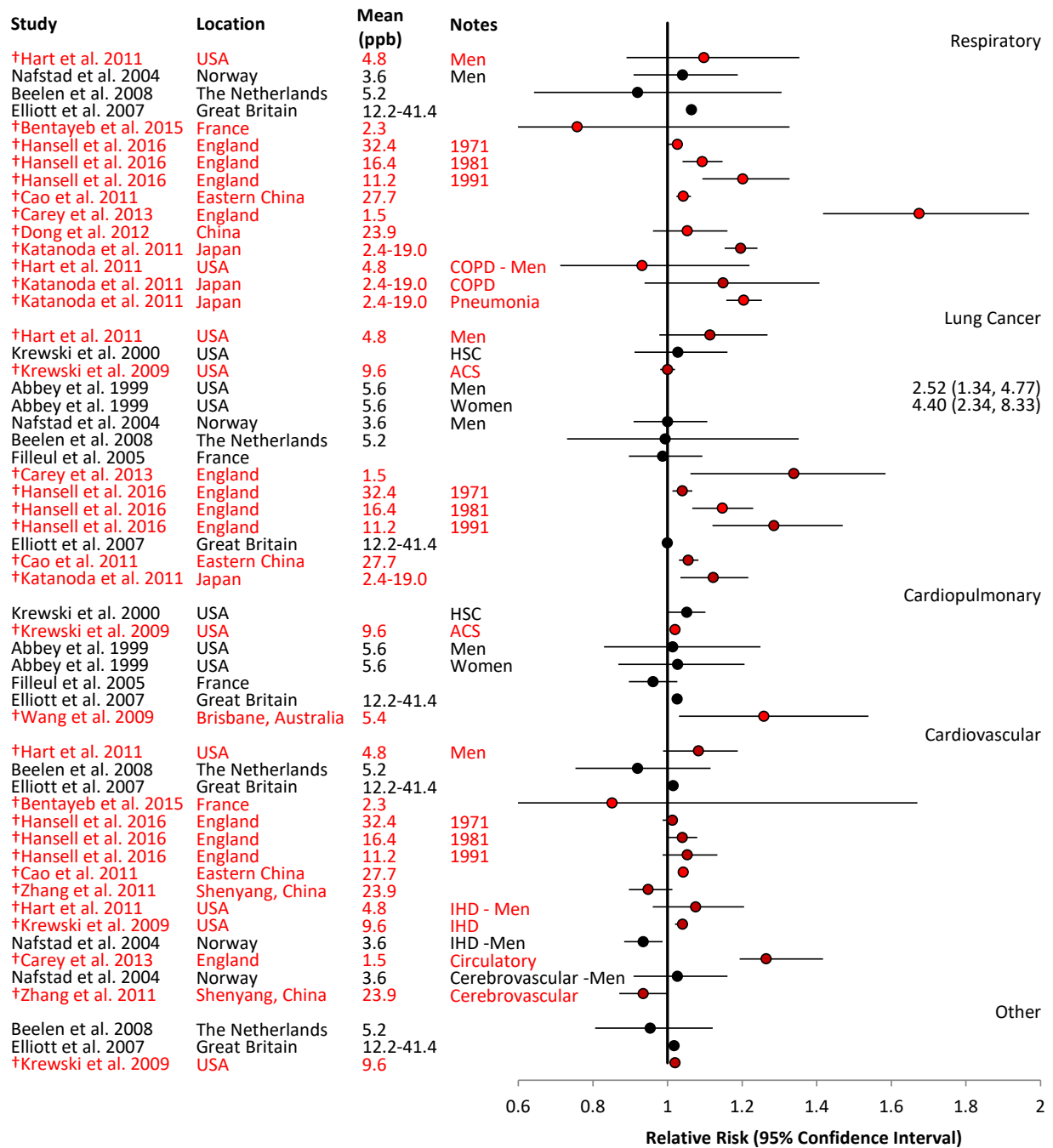
The recent evidence is generally consistent with the evidence in the 2008 SO_x ISA. The biggest notable difference is in the improved consistency in the association between long-term exposure to SO₂ and both respiratory and total mortality that comes from the inclusion of recent cohort studies. However, none of these recent studies help to resolve the uncertainties identified in the 2008 SO_x ISA related to copollutant confounding or the geographic scale of the analysis. All available evidence for mortality due to long-term exposure to SO₂ was evaluated using the framework described in Table II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)). The key evidence as it relates to the causal framework is summarized in [Table 5-43](#). The overall evidence is inadequate to infer a causal relationship between long-term exposure to SO₂ and total mortality among adults.



ACS = American Cancer Society Study; HSC = Harvard Six Cities Study.

Note: studies in red are recent studies. Studies in black were included in the 2008 ISA for Sulfur Oxides. Relative risks are standardized to a 5-ppb increase in sulfur dioxide concentrations. Corresponding quantitative results are reported in Supplemental Table 5S-31 ([U.S. EPA, 2017c](#)).

Figure 5-26 Relative risks (95% confidence interval) of sulfur dioxide-associated total mortality.



ACS = American Cancer Society Study; COPD = chronic obstructive pulmonary disease; HSC = Harvard Six Cities Study; IHD = ischemic heart disease.

Note: Studies in red are recent studies. Studies in black were included in the 2008 ISA for Sulfur Oxides. Relative risks are standardized to a 5-ppb increase in sulfur dioxide concentrations. Corresponding quantitative results are reported in Supplemental Table 5S-32 (U.S. EPA, 2017c).

Figure 5-27 Relative risks (95% confidence interval) of sulfur dioxide-associated cause-specific mortality.

Table 5-43 Summary of evidence, which is inadequate to infer a causal relationship between long-term sulfur dioxide exposure and total mortality.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Some epidemiologic studies report positive associations but results are not entirely consistent	Small, positive associations between long-term exposure to SO ₂ and mortality in the HSC cohort, the ACS cohort, and the Veterans cohort, even after adjustment for common potential confounders	Krewski et al. (2000) †Krewski et al. (2009) Jerrett et al. (2003) Krewski et al. (2000)	Mean: 1.6–24.0 ppb <hr/> City-specific annual mean: 9.3–9.6 ppb
	Recent cohort studies in the U.S. observe increases in total mortality and mortality due to lung cancer and cardiovascular and respiratory disease, but exposure assessment and statistical methods were not adequate for study of SO ₂	†Hart et al. (2011)	<hr/> County-level mean from air quality model: 4.3 ppb <hr/> Annual average at residential address from model: 4.8 ppb
Some epidemiologic studies report no associations	No association observed in European cohort studies for total, respiratory, or cardiovascular mortality	Beelen et al. (2008b)	IDW to regional monitors: 5.2 ppb
		Nafstad et al. (2004)	Model/monitor hybrid: 3.6 ppb
		Filleul et al. (2005)	3-yr mean: 3.0–8.2 ppb
Uncertainty due to potential confounding from correlated pollutants	When reported, correlations with copollutants were generally moderate (0.4–0.7) to high (>0.7). Confounding of observed associations by other pollutants or pollutant mixtures cannot be ruled out	Table 5-42	
Uncertainty regarding exposure measurement error	SO ₂ has low (<0.4) to moderate (0.4–0.7) spatial correlations across urban geographical scales. The geographical scale for estimating exposure used in these studies may be too large for a highly spatially heterogeneous pollutant such as SO ₂ .	Section 3.4.2	
	Exposure measurement error in long-term SO ₂ exposure can lead to bias toward or away from the null	Section 3.4.4.2	

Table 5-43 (Continued): Summary of evidence, which is inadequate to infer a causal relationship between long term sulfur dioxide exposure and total mortality.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
No coherence with evidence for respiratory and cardiovascular morbidity	No evidence for long-term exposure and respiratory health effects in adults to support the observed associations with respiratory mortality	Section 5.2.2.6	
	No evidence for long-term exposure and cardiovascular health effects in adults to support the observed associations with cardiovascular mortality	Section 5.3.2.6	

ACS = American Cancer Society; HSC = Harvard Six Cities; IDW = inverse distance weighting; SO₂ = sulfur dioxide.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)).

^bDescribes the key evidence and references contributing most heavily to causal determination and, where applicable, to uncertainties or inconsistencies. References to earlier sections indicate where full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated (for experimental studies, below 5,000 ppb).

†Studies published since the 2008 ISA for Sulfur Oxides.

5.6 Cancer

5.6.1 Introduction

The 2008 SO_x ISA summarized the literature on SO₂ concentrations and lung cancer as “inadequate to infer a causal relationship” ([U.S. EPA, 2008d](#)). Multiple studies across the U.S. and Europe investigated the relationship of SO₂ concentrations with lung cancer incidence and mortality. Many studies reported generally null associations, including some that were limited by a small number of cancer cases, but some studies demonstrated positive associations. The body of literature characterizing the carcinogenic, genotoxic, and mutagenic effects of exposure to SO₂ has grown since the 2008 SO_x ISA. The animal toxicology literature of SO₂ exposure is dominated by studies of SO₂ acting as a cocarcinogen or tumor promoter, with one study of SO₂ inhalation associated with an increased rate of lung tumor formation in lung tumor-susceptible female rodents. Genotoxicity and mutagenicity studies show mixed results with null studies in a *Drosophila* model and positive micronuclei (MN) findings in a mouse inhalation model of SO₂ exposure. These recent studies have not informed the uncertainties identified in

the previous review, including uncertainties due to exposure measurement error, potential copollutant confounding, and limited mechanistic evidence or biological plausibility.

The cancer section of the ISA characterizes epidemiologic associations of SO₂ exposure with cancer incidence or cancer mortality, as well as the animal toxicology carcinogenicity studies ([Section 5.6.2](#)). Subsections discuss the evidence relating to lung cancer ([Section 5.6.2.1](#)), bladder cancer ([Section 5.6.2.2](#)), and other cancers ([Section 5.6.2.3](#)). Laboratory studies of mutagenicity or genotoxicity are discussed in [Section 5.6.3](#). Supplemental Tables provide detailed summaries of the respective new epidemiologic [Table 5S-33 ([U.S. EPA, 2017c](#))] and genotoxic/mutagenic [Table 5S-34 ([U.S. EPA, 2017c](#))] literature.

5.6.2 Cancer Incidence and Mortality

5.6.2.1 Lung Cancer Incidence and Mortality

International studies exploring the associations between SO₂ concentrations and lung cancer incidence have provided inconsistent results. No recent studies on SO₂ concentration and lung cancer incidence in the U.S. have been published. Large studies conducted using the Netherlands Cohort Study on Diet and Cancer examined the association between SO₂ concentration and lung cancer incidence ([Brunekreef et al., 2009](#); [Beelen et al., 2008a](#)). Null associations were reported in both analyses of the full cohort and a case-cohort design. None of the analyses adjusted for copollutants. An ecological study in Israel examining lung cancer incidence among men also reported null results for the association with SO₂ concentrations ([Eitan et al., 2010](#)). Results were relatively unchanged when adjusting for PM₁₀. No association was observed between SO₂ concentrations and lung cancer hospitalizations among men or women in southern France in an ecological study that did not control for copollutants ([Pascal et al., 2013](#)). However, an ecological analysis performed among women in Taiwan demonstrated a positive association between SO₂ concentration and lung cancer incidence ([Tseng et al., 2012](#)). This positive association remained in a regression model adjusted for other pollutants (CO, NO₂, NO, O₃, and PM₁₀; none of these air pollutants exhibited an association with lung cancer incidence). The association was present in analyses for both types of lung cancer examined, adenocarcinomas and squamous cell carcinomas. Thus, overall, multiple ecologic studies have been performed examining SO₂ concentrations and lung cancer incidence with inconsistent findings, and analyses using a large cohort study reported no association between SO₂ concentrations and lung cancer incidence but had no control of copollutant confounders. Each of these studies used SO₂ concentrations

measured at fixed-site monitors to assign exposure. [Beelen et al. \(2008a\)](#) and [Brunekreef et al. \(2009\)](#) used inverse distance weighting between the fixed-site monitor location and residential address, and combined this with the output of land use regression (LUR) models for urban contributions. [Eitan et al. \(2010\)](#) generated spatially interpolated surfaces for a 7-year period, while the other ecological studies relied on annual averages from the fixed-site monitors. None of the studies corrected for exposure measurement error.

Studies in the U.S. have reported inconsistent findings for the association between SO₂ concentrations and lung cancer mortality (see [Section 5.5.2](#) and [Figure 5-27](#)). No association between SO₂ concentrations and lung cancer mortality was present in a report by the Health Effects Institute ([Krewski et al., 2009](#)). Estimates stratified by high school education (less than high school education, high school education, or greater) were also examined, and no association was present in either subgroup. In addition to the entire time period of the study, the researchers also examined 5-year increments, none of which demonstrated an association. However, a recent study of men in the trucking industry found a slight positive association between SO₂ concentrations and lung cancer mortality ([Hart et al., 2011](#)). With the inclusion of PM₁₀ and NO₂ in the model, the 95% CI included the null but the point estimate was in the positive direction and only slightly attenuated.

Recent studies have also been performed in Asia and Europe examining the relationship between SO₂ and lung cancer mortality. In China, a positive association was observed between SO₂ and lung cancer mortality ([Chen et al., 2016](#); [Cao et al., 2011](#)). In the study by [Cao et al. \(2011\)](#), this association was relatively unchanged with adjustment for either TSP or NO_x. A study in Japan also reported a positive association between SO₂ and lung cancer mortality ([Katanoda et al., 2011](#)). However, the estimate was reduced when additional potential confounders (smoking of parents during subjects' childhood, consumption of nonyellow or nongreen vegetables, occupation, and health insurance) were controlled for and no copollutant assessment was performed. Positive associations were also observed for suspended PM, PM_{2.5}, and NO₂ concentrations. When examining subgroups, the association was highest among male smokers. The point estimate was similar to the overall estimate for male former smokers, but the 95% confidence interval was wide due to the small size of the study population. The estimate was lowest among female never smokers. The number of male never smokers and female smokers were too small to assess individually. A study in the U.K. also demonstrated a positive association between SO₂ concentration and lung cancer mortality ([Carey et al., 2013](#)). The association was slightly attenuated when education was included in the model instead of income. However, a large study using the Netherlands Cohort Study on Diet and Cancer reported no association between SO₂ concentration and lung cancer mortality ([Brunekreef](#)

[et al., 2009](#)). This study was mentioned above and also did not demonstrate an association between SO₂ concentration and lung cancer incidence. No copollutant models were examined. In summary, consistent with studies conducted in the U.S. examining SO₂ concentrations and cancer mortality, recent studies performed in Asia and Europe also had inconsistent findings. Many of these studies used SO₂ concentrations measured at fixed-site monitors to assign exposure, and none of the studies corrected for exposure measurement error. [Brunekreef et al. \(2009\)](#) used inverse distance weighting between the fixed-site monitor location and residential address, and combined this with the output of LUR models for urban contributions. [Hart et al. \(2011\)](#) used spatial smoothing, and [Carey et al. \(2013\)](#) used a dispersion model constructed with emissions data to assign exposure.

A study in Italy used a Lagrangian dispersion model to estimate SO_x concentrations as a marker for refinery plant emissions [exposure assessment technique summarized in [Section 3.3.2.4 \(Ancona et al., 2015\)](#)]. The relationship between these estimates and cancer mortality and hospitalizations were investigated. No association was observed for lung cancer among men or women; however, these results are difficult to interpret. The estimated SO_x concentrations were highly correlated with estimates of PM₁₀, which is expected as SO_x was being treated as a marker for petrochemical refinery emissions. This makes interpretation difficult as copollutant models were not shown for lung cancer and additionally the validity of the model is unknown.

A recent meta-analysis ([Chen et al., 2015a](#)) combined the results of five studies of SO₂ and lung cancer and found an overall OR of 1.03 (95% CI: 1.02, 1.05), although one of the five studies [[Cao et al., 2011](#)]; characterized above] accounted for nearly 80% of the weight contributing to the overall OR and was the only study of the five to observe a positive and statistically significant association between SO₂ exposure and lung cancer. Three of the remaining studies included in the meta-analysis observed null associations between SO₂ and lung cancer.

Sulfur Dioxide Lung Carcinogenesis, Cocarcinogenic Potential, and Tumor Promotion in Laboratory Animal Models

The toxicological evidence for effects of sulfur dioxide in carcinogenicity, mutagenicity, or genotoxicity is characterized below.

Direct evidence of carcinogenicity was studied evaluating incidence of lung tumors in a lung adenoma-susceptible mouse strain, (the LX mouse), with chronic exposure to sulfur dioxide at 500 ppm, 5 minutes/day, 5 days/week for 2 years ([Peacock and Spence, 1967](#)). SO₂-exposed female mice had an increase in the number of lung tumors subgrouped as (1) adenomas and (2) primary carcinomas versus controls. Males had a smaller increase

in adenomas versus controls and similar levels of primary carcinomas compared to controls.

Evidence exists for SO₂ to be a cocarcinogen ([Pauluhn et al., 1985](#)); SO₂ and benzo[a]pyrene (B[a]P), coexposure increased the incidence of lung tumor formation in rodents versus B[a]P exposure alone. Chronic coexposure to SO₂ and B[a]P resulted in increased incidence of upper respiratory tract neoplasia in rats ([Laskin et al., 1976](#)) and hamsters ([Pauluhn et al., 1985](#)) over B[a]P exposure alone. SO₂ exposure shortened the induction period for spontaneous squamous cell lung tumor formation after B[a]P exposure ([Laskin et al., 1976](#)); rats were exposed 5 days a week, 6 hours/day for their lifetime to 10 ppm SO₂ alone via inhalation or 4 ppm SO₂ + 10 mg/m³ B[a]P (1 hour B[a]P/day). SO₂ exposure also shortened the induction time for methylcholanthrene-induced carcinogenesis.

Multiple studies explored SO₂ as a cocarcinogen or promoter after particulate-induced tumorigenesis. In a study of SPM--induced tumorigenesis (proliferative lesions of pulmonary endocrine cells) in the rat, SO₂ did not exacerbate SPM-dependent hyperplasia when rats were exposed to the mixture of SPM and SO₂ ([Ito et al., 1997](#)). Adult male rats were exposed to SO₂ for 11 months, 16 hours/day ± SPM for 4 weeks, once/week by intratracheal injection. SO₂ did not act as a tumor promoter or cocarcinogen in this model. In a separate study of diesel exhaust particle- (DEP-) dependent lung tumorigenesis, SO₂ was able to promote DEP-dependent tumorigenesis ([Ohyama et al., 1999](#)). Adult male rats were intratracheally instilled with diesel exhaust particle extract-coated carbon black particles (DEcCBP) and exposed to 4 ppm SO₂ for 10 months. Eighteen months after starting the experiment, the animals were examined for respiratory tract tumors and DNA adducts were measured in lung tissue. Lung tumors and DNA adducts were seen in animals with coexposure to SO₂ and DEcCBP but not in animals only exposed to DEcCBP. SO₂ acted as a tumor promoter in animals exposed to DEcCBP. In a separate investigation, hamsters were exposed to diesel engine exhaust (separately with and without particles) and a mixture of SO₂ and NO₂ with or without exposure to the carcinogen diethyl-nitrosamine to investigate the potential cocarcinogenic effect of exposure to the dioxides mixture and diesel engine exhaust in the respiratory tract ([Heinrich et al., 1989](#)). These adult male hamsters were exposed for 19 hours/day, 5 days/week for 6, 10.5, 15, or 18 months to diesel exhaust, filtered diesel exhaust (without particles), a dioxide mixture of NO₂ (5 ppm) and SO₂ (10 ppm), or clean air. Two exposure groups from each of the aforementioned test groups were also given a single subcutaneous injection of diethylnitrosamine (DEN) (3 mg or 6 mg/kg body weight). Exposure to the dioxide mixture by itself did not elevate tumor rate (tumor induction), nor did it exacerbate DEN-dependent effects (tumor promotion) in the hamster. In summary, a comparison of multiple studies of SO₂ coexposure with particles

reported mixed results in various models of carcinogenicity, cocarcinogenic potential, or tumor promotion.

Oncogene and tumor suppressor genes also appear to be affected by SO₂ exposure, especially with coexposure to B[a]P. Synergistic expression of *c-fos* and *c-jun* with SO₂ and B[a]P coexposure was observed in rodent lungs ([Qin and Meng, 2006](#)). SO₂ and B[a]P coexposure in male Wistar rats (26.5 ppm SO₂ inhalation, 6 hours/day for 7 days; 3 mg B[a]P instilled) statistically significantly downregulated expression of tumor suppressor genes *p16* and *myc*, and increased expression of oncogenes *c-myc*, *H-ras*, and *p53*. Others have reported that SO₂ exposure alone could induce *p53* expression in rats ([Bai and Meng, 2005](#)).

5.6.2.2 Bladder Cancer Incidence and Mortality

Several studies on the relationship between SO₂ concentrations and bladder cancer incidence and mortality have been published since the 2008 SO_x ISA ([U.S. EPA, 2008d](#)). Positive associations were observed in studies of bladder cancer mortality but not bladder cancer incidence. An ecological study in southern France reported on the relationship between SO₂ concentrations and hospitalizations for bladder cancer without examination of copollutant models ([Pascal et al., 2013](#)). Null associations were observed for men and women. Another ecological study in Israel examining bladder cancer incidence also reported sex-stratified results ([Eitan et al., 2010](#)). Neither sex demonstrated an association between SO₂ concentrations and bladder cancer in models with and without adjustment for PM₁₀. However, an association was observed in a study examining the relationship between SO₂ and bladder cancer mortality ([Liu et al., 2009a](#)). [Liu et al. \(2009a\)](#) investigated the association between SO₂ and bladder cancer mortality using controls with mortality due to causes unrelated to neoplasm or genitourinary-related disease and matched by sex, year of birth, and year of death. A positive association was observed between SO₂ concentration in the second and third tertiles of exposure and bladder cancer mortality. For further investigations, the authors created a three-level exposure variable combining NO₂ and SO₂ concentrations: the lowest tertile of SO₂ and NO₂ concentrations (≤ 4.32 ppb and ≤ 20.99 ppb, respectively), the highest tertile of SO₂ and NO₂ concentrations (> 6.49 ppb and > 27.33 ppb, respectively), and other categorizations/combinations. The ORs were 1.98 (95% CI 1.36, 2.88) for the highest level of NO₂ and SO₂ and 1.37 (95% CI 1.03, 1.82) for the middle level categorizations. Although the point estimates are higher than those observed for SO₂ alone [see Supplemental Table 5S-33, ([U.S. EPA, 2017c](#))], the 95% confidence intervals overlap, and therefore, conclusions that NO₂ and SO₂ combined contribute to higher odds of mortality than either alone cannot be drawn. Finally, a study using SO_x concentration

estimated using a Lagrangian dispersion model reported no association between SO_x concentration and bladder cancer mortality or hospitalizations among men or women ([Ancona et al., 2015](#)). However, results of this study are difficult to interpret because of unknown validity of the model (see [Section 3.3.2.4](#)) and high correlation with PM₁₀ and hydrogen sulfide (H₂S).

5.6.2.3 Incidence of Other Cancers

Recent studies of SO₂ concentrations and other cancer types have been published since the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), but provide limited information on associations with SO₂. An ecological study in southern France investigated the relationships between SO₂ and hospitalizations for breast cancer, acute leukemia, myeloma, and non-Hodgkin lymphoma ([Pascal et al., 2013](#)). Null associations were observed in sex-stratified analyses among men and women, with the exception of a positive association between SO₂ and acute leukemia among men. However, the authors urge caution when interpreting the results due to a small number of male acute leukemia cases. This study did not examine copollutant confounding. Another ecologic study used Surveillance, Epidemiology, and End Results data to examine the correlation between SO₂ concentrations and breast cancer incidence ([Wei et al., 2012](#)). A positive relationship was detected, but there was no control for potential confounders of other air pollutants (of which CO, NO_x, and VOCs, but not PM₁₀, also demonstrated a positive correlation with breast cancer incidence). Both of these studies are limited by their ecologic nature and the lack of individual-level data. A cross-sectional study was conducted in South Korea that looked at the association between symptom scores for prostate cancer and emissions data for SO_x (measured in kg/year/person) and a number of other air pollutants ([Shim et al., 2015](#)). In logistic regression models adjusted for age, the authors observed positive associations between men living in areas with greater emissions of SO_x and symptom scores for prostate cancer. Similar results were observed for NO_x, CO, PM₁₀, VOCs, and NH₃. The lack of control for potential confounding by other air pollutants or risk factors (e.g., smoking, SES) limit the interpretation of these results.

A cohort study examined the relationship between SO_x concentrations, estimated using a Lagrangian dispersion model, and hospitalizations and mortality for various cancer types ([Ancona et al., 2015](#)). No associations were found between SO_x concentrations and either hospitalizations or mortality due to cancers of the stomach, colon/rectum, liver, kidney, brain, or breast. Positive associations were observed for SO_x concentration and mortality due to pancreatic and larynx cancers among women but not men. The 95% confidence interval showed a large degree of imprecision in the estimates for cancer of the larynx. The association with pancreatic cancer was not robust to adjustment with H₂S or PM₁₀.

When examining the association between estimated SO_x concentration and hospitalizations, a positive, but imprecise, association was observed for cancer of the larynx among women and an inverse association was noted for cancers of lymphatic and hematopoietic tissue.

5.6.2.4 Summary of Cancer Incidence and Mortality

Similar to studies of SO₂ concentrations and lung cancer in the previous ISA ([U.S. EPA, 2008d](#)), recent studies of SO₂ concentrations and lung cancer have provided inconsistent results ([Carey et al., 2013](#); [Pascal et al., 2013](#); [Tseng et al., 2012](#); [Cao et al., 2011](#); [Hart et al., 2011](#); [Katanoda et al., 2011](#); [Eitan et al., 2010](#); [Brunekreef et al., 2009](#); [Beelen et al., 2008a](#)), with two of the largest studies reporting null results ([Brunekreef et al., 2009](#); [Krewski et al., 2009](#)). Studies of bladder cancer appear to find no association between SO₂ concentrations and bladder cancer incidence ([Pascal et al., 2013](#); [Eitan et al., 2010](#)), but a study of SO₂ concentration and bladder cancer mortality reported a positive association ([Liu et al., 2009a](#)). Limited information is available regarding other cancers. Animal toxicology models of SO₂ inhalation exposure show SO₂ acting as a promoter or cocarcinogen, with one study showing increased lung tumor formation in a lung tumor-prone animal model.

5.6.3 Genotoxicity and Mutagenicity

Multiple studies of genotoxicity or mutagenesis with SO₂ in vivo or SO₂ in vitro exposure have been reported in the literature and are detailed below in Supplemental Table 5S-34 ([U.S. EPA, 2017c](#)).

After inhalation exposure to SO₂, mouse bone marrow micronuclei formation was significantly elevated in both males and females after exposure to SO₂ (5.4, 10.7, 21.4, or 32.1 ppm SO₂, 4 hours/day for 7 days) ([Meng et al., 2002](#)). The polychromatophilic erythroblasts of the bone marrow (MNPCE) were formed in significantly increased numbers with SO₂ exposure. Another study recapitulated these findings; subacute exposure to SO₂ (10.7 ppm SO₂, 6 hours/day for 5 days) induced a significant increase in MNPCE with this effect attenuated by exogenous antioxidant SSO pretreatment ([Ruan et al., 2003](#)).

The rate of DNA single strand breaks induced by B[a]P exposure in fetal hamster lung cells (50 ppm for 2 weeks) ([Pool et al., 1988b](#)) and rat liver cells (2.5, 5, 9.9, or 19.9 ppm, 4 hours/day for 7 days) ([Pool et al., 1988a](#)) was significantly attenuated by concomitant exposure to SO₂ (50 ppm for 2 weeks).

Genotoxicity testing of *Drosophila* sperm for sex-linked recessive lethals after feeding larvae 0.04 M or 0.08 M sodium sulfite in a 1% glucose solution was performed, and no increase was found above background. One caveat is that sulfite can interact with glucose, making the exposure assessment more complicated.

Multiple studies of genotoxicity or mutagenesis with SO₂ in vivo or in vitro exposure have been reported in the literature and are summarized in Supplemental Table 5S-34 ([U.S. EPA, 2017c](#)). Mixed results of genotoxicity or mutagenicity have been reported after SO₂ exposure including positive associations with SO₂ inhalation exposure in the mouse MN assay.

5.6.4 Summary and Causal Determination

The overall evidence for long-term SO₂ exposure and cancer is inadequate to infer a causal relationship. This conclusion is based on the inconsistent evidence from epidemiologic studies, as well as mixed evidence within the animal toxicology and mode of action framework for mutagenesis and genotoxicity. In past reviews, a limited number of epidemiologic studies had assessed the relationship between long-term SO₂ concentrations and cancer incidence and mortality. The 2008 ISA for Sulfur Oxides concluded that the evidence was “inconclusive” ([U.S. EPA, 2008d](#)). Recent studies include evidence on lung cancer as well as new types of cancer, evaluating both incidence and mortality. However, the additional recent evidence has not informed any of the uncertainties identified in the previous review, including uncertainties due to exposure measurement error, potential copollutant confounding, and limited mechanistic evidence or biological plausibility. All available evidence for cancer due to long-term SO₂ concentrations was evaluated using the framework described in Table II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)). The key evidence as it relates to the causal framework is summarized in [Table 5-44](#).

Table 5-44 Summary of evidence, which is inadequate to infer a causal relationship between long-term sulfur dioxide exposure and cancer.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Among a small body of evidence, evidence from epidemiologic studies is inconsistent	Generally null associations from studies of cancer incidence, with some observed increases in lung cancer and bladder cancer mortality in studies conducted in the U.S., Europe, and Asia	Section 5.6.2	Means varied across studies including areas estimating mean concentrations of SO ₂ as low as 1.49 ppb to as high as 27.87 ppb. Associations observed with bladder cancer mortality at levels as low as 4.39–6.09 ppb.
Uncertainty due to exposure measurement error	Fixed-site monitors used in cancer studies may not capture spatial variability of SO ₂ concentrations.	Section 3.4.2.2	
	Exposure measure measurement error in long-term SO ₂ exposure assessment can bias toward or away from the null	Section 3.4.4.2	
Uncertainty due to confounding by correlated copollutants	Correlations of SO ₂ with other pollutants vary by study or are not examined. Some pollutants are moderately to highly correlated with SO ₂ but are not always taken into account as potential confounders	Section 3.4.3	
Uncertainty due to limited coherence with toxicological evidence	Studies in a tumor-susceptible mouse model, females had increased numbers of lung adenomas and carcinomas.	Peacock and Spence (1967)	500,000 ppb
		Laskin et al. (1976)	10,000 ppb
	Studies of facilitation of metastasis and coexposures with known carcinogens show mixed SO ₂ related effects.	Pauluhn et al. (1985)	172,000 ppb
		Ohyama et al. (1999)	4,000 ppb
		Heinrich et al. (1989)	5,000 or 10,000 ppb
		Ito et al. (1997)	4,000 ppb
		Section 5.6.2.1	

Table 5-44 (Continued): Summary of evidence, which is inadequate to infer a causal relationship between long term sulfur dioxide exposure and cancer.

Rationale for Causal Determination ^a	Key Evidence ^b	Key References ^b	SO ₂ Concentrations Associated with Effects ^c
Some evidence identifies key events within the MOA from mutagenesis and genotoxicity	Mixed evidence of mutagenicity and genotoxicity formation in animal cells exposed to SO ₂	Meng et al. (2002) , Ruan et al. (2003) , Pool et al. (1988b) Section 5.6.3	5,000, 10,700, 21,400, 32,100 ppb

MOA = mode of action; SO₂ = sulfur dioxide.

^aBased on aspects considered in judgments of causality and weight of evidence in causal framework in Tables I and II of the Preamble to the ISAs ([U.S. EPA, 2015b](#)).

^bDescribes the key evidence and references contributing most heavily to causal determination and, where applicable, to uncertainties or inconsistencies. References to earlier sections indicate where full body of evidence is described.

^cDescribes the SO₂ concentrations with which the evidence is substantiated (for experimental studies, below 5,000 ppb).

Overall, there is inconsistent evidence for an association between long-term SO₂ exposure and cancer from epidemiologic and toxicological studies. Some of the epidemiologic studies observed positive associations while others did not. Some of these studies with positive associations were relatively unchanged with the inclusion of various cofounders and copollutants, although many did not evaluate the potential for copollutant confounding. Cohort studies have reported null associations between SO₂ concentrations and lung cancer incidence. Similarly, some ecological studies also reported no associations; although, an ecological study in Taiwan among women did report an association between SO₂ concentrations and lung cancer incidence that was relatively unchanged when including other pollutants. Positive associations were also observed in a study of SO₂ concentrations and bladder cancer mortality but not in ecological studies of bladder cancer incidence. The study of bladder cancer mortality examined the relationship between bladder cancer mortality and joint exposure to high levels of NO₂ and SO₂, but no copollutant assessment was performed controlling for NO₂ or other air pollutants. None of the epidemiologic studies made corrections or adjustments for exposure measurement error, or accounted for the potential for bias away from the null, the potential for which has been demonstrated in simulation studies (see [Section 3.4.4.2](#)).

Animal toxicological studies employing SO₂ exposure with other known carcinogens provide some evidence, showing that inhaled SO₂ can increase tumor load in laboratory rodents. Toxicological data provided by a study in LX mice, lung adenoma susceptible animals, showed evidence of the direct carcinogenic potential of SO₂. Other studies in animal models show SO₂ as a cocarcinogen with B[a]P or as a tumor promoter with particulate-induced tumorigenesis. Nonetheless, toxicological data provide no clear evidence of SO₂ acting as a complete carcinogen and not all epidemiologic studies report positive associations.

Collectively, the inconsistent evidence from several toxicological and epidemiologic studies is inadequate to infer a causal relationship between long-term exposure to SO₂ and cancer incidence and mortality.

Annex for Chapter 5: Evaluation of Studies on Health Effects of Sulfur Oxides

Table A-1 Scientific considerations for evaluating the strength of inference from studies on the health effects of sulfur oxides.

Evaluation Factors
Study Design
Controlled Human Exposure:
Studies should clearly describe the primary and any secondary objectives of the study, or specific hypotheses being tested. Study subjects should be randomly exposed without knowledge of the exposure condition. Preference is given to balanced crossover (repeated measures) or parallel design studies that include control exposures (e.g., to clean filtered air). In crossover studies, a sufficient and specified time between exposure days should be employed to avoid carry over effects from prior exposure days. In parallel design studies, all arms should be matched for individual characteristics such as age, sex, race, anthropometric properties, and health status. In studies evaluating effects of disease, appropriately matched healthy controls are desired for interpretative purposes.
Animal Toxicology:
Studies should clearly describe the primary and any secondary objectives of the study, or specific hypotheses being tested. Studies should include appropriately matched control exposures (e.g., to clean filtered air, time matched). Studies should use methods to limit differences in baseline characteristics of control and exposure groups. Studies should randomize assignment to exposure groups and where possible conceal allocation from research personnel. Groups should be subjected to identical experimental procedures and conditions; animal care including housing, husbandry, etc. should be identical between groups. Blinding of research personnel to study group may not be possible due to animal welfare and experimental considerations; however, differences in the monitoring or handling of animals in all groups by research personnel should be minimized.
Epidemiology:
Inference is stronger for studies that clearly describe the primary and any secondary aims of the study, or specific hypotheses being tested. For short-term exposure, time-series, case crossover, and panel studies are emphasized over cross-sectional studies because they examine temporal correlations and are less prone to confounding by factors that differ between individuals (e.g., SES, age). Studies with large sample sizes and conducted over multiple years are considered to produce more reliable results. If other quality parameters are equal, multicity studies carry more weight than single-city studies because they tend to have larger sample sizes and lower potential for publication bias. For long-term exposure, inference is considered to be stronger for prospective cohort studies and case-control studies nested within a cohort (e.g., for rare diseases) than cross-sectional, other case-control, or ecologic studies. Cohort studies can better inform the temporality of exposure and effect. Other designs can have uncertainty related to the appropriateness of the control group or validity of inference about individuals from group-level data. Study design limitations can bias health effect associations in either direction.

Study Population/Test Model

Controlled Human Exposure:

In general, the subjects recruited into study groups should be similarly matched for age, sex, race, anthropometric properties, and health status. In studies evaluating effects of specific subject characteristics (e.g., disease, genetic polymorphism, etc.), appropriately matched healthy controls are preferred. Relevant characteristics and health status should be reported for each experimental group. Criteria for including and excluding subjects should be clearly indicated. For the examination of populations with an underlying health condition (e.g., asthma), independent, clinical assessment of the health condition is ideal, but self reporting of physician diagnosis generally is considered to be reliable for respiratory and cardiovascular disease outcomes.^a The loss or withdrawal of recruited subjects during the course of a study should be reported. Specific rationale for excluding subject(s) from any portion of a protocol should be explained.

Animal Toxicology:

Ideally, studies should report species, strain, substrain, genetic background, age, sex, and weight. Unless data indicate otherwise, all animal species and strains are considered appropriate for evaluating effects of SO₂ exposure. It is preferred that the authors test for effects in both sexes and multiple lifestages, and report the result for each group separately. All animals used in a study should be accounted for, and rationale for exclusion of animals or data should be specified.

Epidemiology:

Confidence in results is greater in studies that recruit the study population from the target population and examine a study population that is representative of the target population. Studies with high participation and low drop-out over time that is not dependent on exposure or health status are considered to have low potential for selection bias. Clear indication of criteria for including and excluding subjects can facilitate assessment of selection bias. For populations with an underlying health condition, independent, clinical assessment of the health condition is valuable, but self report of physician diagnosis generally is considered to be reliable for respiratory and cardiovascular outcomes. Comparisons of groups with and without an underlying health condition are more informative if groups are from the same source population. Selection bias can influence results in either direction or may not affect the validity of results but rather reduce the generalizability of findings to the target population.

Pollutant

Controlled Human Exposure:

The focus is on studies testing SO₂ exposure.

Animal Toxicology:

The focus is on studies testing SO₂ exposure.

Epidemiology:

The focus is on studies testing SO₂ exposure.

Exposure Assessment or Assignment

Controlled Human Exposure:

For this assessment, the focus will be on studies that use SO₂ concentrations less than or equal to 2 ppm ([Section 1.2](#)). Studies that use higher exposure concentrations may provide information relevant to mode of action, dosimetry, inter-species variation, or at-risk human populations. Controlled human exposure studies considering short-term, (e.g., generally exposures from 5–10 min, to 0.2–0.6 ppm SO₂, were emphasized) ([Section 1.2](#)).

Animal Toxicology:

For this assessment, the focus will be on studies that use SO₂ concentrations less than or equal to 2,000 ppb ([Section 1.2](#)). Studies that use higher exposure concentrations may provide information relevant to mode of action, dosimetry, inter-species variation, or at-risk human populations. Studies should characterize pollutant concentration, temperature, and relative humidity and/or have measures in place to adequately control the exposure conditions. The focus is on inhalation exposure. Noninhalation exposure experiments may provide information relevant to mode of action. In vitro studies may be included if they provide mechanistic insight or examine similar effects as in vivo, but are generally not included. All studies should include exposure control groups (e.g., clean filtered air).

Epidemiology:

Of primary relevance are relationships of health effects with the ambient component of exposure to SO₂. However, information about ambient exposure rarely is available for individual subjects; most often, inference is based on ambient concentrations. Studies that compare exposure assessment methods are considered to be particularly informative. Inference is stronger when the duration or lag of the exposure metric corresponds with the time course for physiological changes in the outcome (e.g., up to a few days for symptoms) or latency of disease (e.g., several years for cancer).

Given the spatial heterogeneity in ambient SO₂ and potentially variable relationships between personal exposures and ambient concentrations ([Sections 3.4.2.2](#) and [3.4.1](#)), validated methods that capture the extent of variability for the particular study design (temporal vs. spatial contrasts) and location carry greater weight. Fixed-site measurements, whether averaged across multiple monitors or assigned from the nearest or single available monitor, have well-recognized limitations in capturing spatial variation in air pollutants. Monitors impacted by large SO₂ sources are particularly subject to concentration fluctuations due to changes in emission rates and meteorological conditions and may not fully represent population exposure. Results based on fixed-site measurements can be informative if correlated with personal exposures, closely located to study subjects, highly correlated across monitors within a location, used in locations with well-distributed sources, or combined with time-activity information.

In studies of short-term exposure, temporal variability of the exposure metric is of primary interest. Metrics that may capture variation in ambient sulfur oxides and strengthen inference include concentrations in subjects' microenvironments and individual-level outdoor concentrations combined with time-activity data. Atmospheric models may be used for exposure assessment in place of or to supplement SO₂ measurements in epidemiologic analyses. Dispersion models (e.g., AERMOD) can provide valuable information on fine-scale temporal and spatial variations (within tens of km) of SO₂ concentrations, which is particularly important for assessing exposure near large stationary sources. Alternatively, grid-scale models (e.g., CMAQ) that represent SO₂ exposure over relatively large spatial scales (e.g., typically greater than 4 × 4-km grid size) often do not provide enough spatial resolution to capture acute SO₂ peaks that influence short-term health outcomes. Uncertainty in exposure predictions from these models is largely influenced by model formulations and the quality of model input data pertaining to emissions or meteorology, which tends to vary on a study-by-study basis.

For long-term exposures, models that capture within-community spatial variation in individual exposure may be given more weight for spatially variable ambient SO₂.

Exposure measurement error often attenuates health effect estimates or decreases the precision of the association (i.e., wider 95% CIs), particularly associations based on temporal variation in short-term exposure ([Section 3.4.4.1](#)). However, exposure measurement error can bias estimates away from the null, particularly for long-term exposures ([Section 3.4.4.2](#)).

Outcome Assessment/Evaluation

Controlled Human Exposure:

Endpoints should be assessed in the same manner for control and exposure groups (e.g., time after exposure, methods, endpoint evaluator) using valid, reliable methods. Blinding of endpoint evaluators is ideal, especially for qualitative endpoints (e.g., histopathology). For each experiment and each experimental group, including controls, precise details of all procedures carried out should be provided including how, when, and where. Time of the endpoint evaluations is a key consideration that will vary depending on the endpoint evaluated. Endpoints should be assessed at time points that are appropriate for the research questions.

Animal Toxicology:

Endpoints should be assessed in the same manner for control and exposure groups (e.g., time after exposure, methods, endpoint evaluator) using valid, reliable methods. Blinding of endpoint evaluators is ideal, especially for qualitative endpoints (e.g., histopathology). For each experiment and each experimental group, including controls, precise details of all procedures carried out should be provided including how, when, and where. Time of the endpoint evaluations is a key consideration that will vary depending on the endpoint evaluated. Endpoints should be assessed at time points that are appropriate for the research questions.

Epidemiology:

Inference is stronger when outcomes are assessed or reported without knowledge of exposure status. Knowledge of exposure status could produce artifactual associations. Confidence is greater when outcomes assessed by interview, self-reporting, clinical examination, or analysis of biological indicators are defined by consistent criteria and collected by validated, reliable methods. Independent, clinical assessment is valuable for outcomes such as lung function or incidence of disease, but report of physician diagnosis has shown good reliability.^a Outcomes assessed at time intervals that correspond with the time course for physiological changes (e.g., up to a few days for symptoms) are emphasized. When health effects of long-term exposure are assessed by acute events such as symptoms or hospital admissions, inference is strengthened when results are adjusted for short-term exposure. Validated questionnaires for subjective outcomes such as symptoms are regarded to be reliable,^b particularly when collected frequently and not subject to long recall. For biological samples, the stability of the compound of interest and the sensitivity and precision of the analytical method is considered.

If not based on knowledge of exposure status, errors in outcome assessment tend to bias results toward the null.

Potential Copollutant Confounding

Controlled Human Exposure:

Exposure should be well characterized to evaluate independent effects of SO₂.

Animal Toxicology:

Exposure should be well characterized to evaluate independent effects of SO₂.

Epidemiology:

Not accounting for copollutant confounding can produce artifactual associations; thus, studies that examine copollutant confounding carry greater weight. The predominant method is copollutant modeling, which is especially informative when measurement error is comparable for copollutants and correlations are not high. Interaction and joint effect models are examined to a lesser extent. Evaluating correlations between SO₂ and copollutants and comparing health associations between SO₂ and copollutants in single-pollutant models can add to the analysis of potential copollutant confounding, particularly when exposure measurement error is comparable among pollutants. Studies that examine SO₂ only in single-pollutant models provide minimal information on the potential for copollutant confounding. Copollutant confounding is evaluated based on the extent of observed correlations and relationships with health effects. Highly variable correlations have been observed between SO₂ and other criteria pollutants at collocated monitors ([Section 3.4.3](#)), ranging from negative to strong correlations, making evaluation of copollutant confounding necessary on a study-specific, rather than a general, basis.

Other Potential Confounding Factors

Controlled Human Exposure:

Preference is given to studies using experimental and control groups that are matched for individual level characteristics (e.g., body weight, smoking history, age) and time-varying factors (e.g., seasonal and diurnal patterns).

Animal Toxicology:

Preference is given to studies using experimental and control groups that are matched for individual level characteristics (e.g., body weight, litter size, food and water consumption) and time-varying factors (e.g., seasonal and diurnal patterns).

Epidemiology:

Factors are considered to be potential confounders if demonstrated in the scientific literature to be related to health effects and correlated with SO₂. Not accounting for confounders can produce artifactual associations; thus, studies that statistically adjust for multiple factors or control for them in the study design are emphasized. Less weight is placed on studies that adjust for factors that mediate the relationship between SO₂ and health effects, which can bias results toward the null. In the absence of information linking health risk factors to SO₂, a factor may be evaluated as a potential effect measure modifier, but uncertainty is noted as to its role as a confounder. Confounders vary according to study design, exposure duration, and health effect and may include, but are not limited to, the following:

For time-series and panel studies of short-term exposure:

- Respiratory effects—meteorology, day of week, season, medication use, allergen exposure (potential effect modifier)
- Cardiovascular effects—meteorology, day of week, season, medication use
- Total mortality—meteorology, day of week, season, long-term temporal trends

For studies of long-term exposure:

- Respiratory effects—socioeconomic status, race, age, medication use, smoking, stress
- Cardiovascular, reproductive, and development effects—socioeconomic status, race, age, medication use, smoking, stress, noise
- Total mortality—socioeconomic status, race, age, medication use, smoking, comorbid health conditions
- Cancer—socioeconomic status, race, age, occupational exposure

Statistical Methodology

Controlled Human Exposure:

Statistical methods should be clearly described and appropriate for the study design and research question (e.g., correction for multiple comparisons). Generally, statistical significance is used to evaluate the findings of controlled human exposure studies. Detection of statistical significance is influenced by a variety of factors including, but not limited to, the size of the study, exposure and outcome measurement error, and statistical model specifications. Sample size is not a criterion for exclusion; ideally, the sample size should provide adequate power to detect hypothesized effects (e.g., sample sizes less than three are considered less informative). Because statistical tests have limitations, consideration is given to both trends in data and reproducibility of results.

Animal Toxicology:

Statistical methods should be clearly described and appropriate for the study design and research question (e.g., correction for multiple comparisons). Generally, statistical significance is used to evaluate the findings of animal toxicology studies. Detection of statistical significance is influenced by a variety of factors including, but not limited to, the size of the study, exposure and outcome measurement error, and statistical model specifications. Sample size is not a criterion for exclusion; ideally, the sample size should provide adequate power to detect hypothesized effects (e.g., sample sizes less than three are considered less informative). Because statistical tests have limitations, consideration is given to both trends in data and reproducibility of results.

Epidemiology:

Multivariable regression models that include potential confounding factors are emphasized. However, multipollutant models (more than two pollutants) are considered to produce too much uncertainty due to copollutant collinearity to be informative. Models with interaction terms aid in the evaluation of potential confounding as well as effect modification. Sensitivity analyses with alternate specifications for potential confounding inform the stability of findings and aid in judgments of the strength of inference of results. In the case of multiple comparisons, consistency in the pattern of association can increase confidence that associations were not found by chance alone. Statistical methods that are appropriate for the power of the study carry greater weight. For example, categorical analyses with small sample sizes can be prone to bias results toward or away from the null. Statistical tests such as *t*-tests and Chi-squared tests are not considered sensitive enough for adequate inferences regarding pollutant-health effect associations. For all methods, the effect estimate and precision of the estimate (i.e., width of 95% CI) are important considerations rather than statistical significance.

AERMOD = American Meteorological Society/U.S. EPA Regulatory Model; CI = confidence interval; CMAQ = Community Multiscale Air Quality; SES = socioeconomic status; SO₂ = sulfur dioxide.

^a[Toren et al. \(1993\)](#); [Murgia et al. \(2014\)](#); [Weakley et al. \(2013\)](#); [Yang et al. \(2011\)](#); [Heckbert et al. \(2004\)](#); [Barr et al. \(2002\)](#); [Muhajarine et al. \(1997\)](#).

^b[Burney et al. \(1989\)](#).

CHAPTER 6 POPULATIONS AND LIFESTAGES POTENTIALLY AT INCREASED RISK FOR HEALTH EFFECTS RELATED TO SULFUR DIOXIDE EXPOSURE

6.1 Introduction

Interindividual variation in human responses to ambient air pollution exposure can result in some groups or lifestyles being at increased risk for health effects. The NAAQS are intended to protect public health with an adequate margin of safety. Protection is provided for both the population as a whole and those potentially at increased risk for health effects in response to exposure to a criteria air pollutant (e.g., SO₂) [see Preamble to the ISAs ([U.S. EPA, 2015b](#))]. The scientific literature has used a variety of terms to identify factors and subsequently populations or lifestyles that may be at increased risk of an air pollutant-related health effect, including *susceptible*, *vulnerable*, *sensitive*, and *at risk*, with recent literature introducing the term *response-modifying factor* ([Vinikoor-Imler et al., 2014](#)) [see Preamble to the ISAs ([U.S. EPA, 2015b](#))]. Acknowledging the inconsistency in definitions for these terms across the scientific literature and the lack of a consensus on terminology in the scientific community, this chapter takes an inclusive and all-encompassing approach and focuses on identifying those populations or lifestyles potentially “at risk” of an SO₂-related health effect.

As discussed in the Preamble to the ISAs ([U.S. EPA, 2015b](#)), risk of health effects from exposure to SO₂ may be modified as a result of intrinsic (e.g., pre-existing disease, genetic factors) or extrinsic factors (e.g., sociodemographic or behavioral factors), differences in internal dose (e.g., due to variability in ventilation rates or exercise behaviors), or differences in exposure to air pollutant concentrations (e.g., more time spent in areas with higher ambient concentrations). Some factors may lead to a reduction in risk and are recognized during the evaluation process, but for the purposes of identifying those populations or lifestyles at greatest risk to inform decisions on the NAAQS, the focus of this chapter is on characterizing those factors that may increase risk. While the emphasis of this chapter is for individual factors that may increase the risk of an SO₂-related health effect, it is recognized that in many cases, portions of the population are at increased risk of an SO₂-related health effect due to a combination of factors [e.g., residential location and socioeconomic status (SES)], but information on the interaction among factors remains limited. Thus, the following sections identify, evaluate, and characterize the overall confidence that individual factors potentially result in

increased risk for SO₂-related health effects [see Preamble to the ISAs ([U.S. EPA, 2015b](#))].

The objective of this chapter is to identify, evaluate, and characterize the overall confidence that various factors may increase the risk of an SO₂-related health effect in a population or lifestage, building on the conclusions drawn herein with respect to SO₂ exposure and health effects. The broad categories of factors evaluated in this chapter include pre-existing disease ([Section 6.3](#)), genetic factors ([Section 6.4](#)), and sociodemographic and behavioral factors ([Section 6.5](#)).

6.2 Approach to Evaluating and Characterizing the Evidence for At-Risk Factors

This chapter takes a systematic approach to identifying and evaluating factors that may increase the risk of a population or specific lifestage to an ambient SO₂-related health effect. This chapter is complementary to the characterization of health evidence presented in [Chapter 5](#), and the systematic approach is described in detail in the Preamble to the ISAs ([U.S. EPA, 2015b](#)). Briefly, in contrast to the overall evaluation of SO₂ exposures and health effects presented in [Chapter 5](#), this chapter specifically aims to identify and characterize the populations and lifestages at increased risk of an SO₂-related health effect. While [Chapter 5](#) includes discussion of some populations and lifestages in order to explicitly characterize the causal nature between SO₂ exposure and health effects based on the body of evidence (e.g., children, individuals with asthma), this chapter applies a systematic approach to evaluating evidence that can inform the identification of such populations and lifestages and applies a formal framework to transparently characterize the strength of this evidence [see Preamble to the ISAs ([U.S. EPA, 2015b](#))]. This chapter informs the NAAQS review with regard to identification of populations at risk.

The evidence evaluated in this chapter includes relevant studies discussed in [Chapter 5](#) of this ISA and builds on the evidence presented in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). Based on the approach developed in previous ISAs ([U.S. EPA, 2016d](#), [2013a](#), [b](#)), evidence is integrated across scientific disciplines and health effects, and where available, with information on exposure and dosimetry ([Chapter 3](#) and [Chapter 4](#)). Greater emphasis is placed on those health outcomes for which a “causal” or “likely to be causal” relationship is concluded in [Chapter 5](#) of this ISA, while information from studies of health outcomes for which the causal determination is “suggestive” is used as supporting evidence where appropriate. Studies examining health outcomes for which an “inadequate” relationship was concluded are not included in this chapter due primarily to the uncertainty in the independent association between exposure to SO₂ and the health

outcome; as a result, these studies are unable to provide information on whether certain populations are at increased risk of SO₂-related health effects. Conclusions are drawn based on the overall confidence that a specific factor may result in a population or lifestage being at increased risk of an SO₂-related health effect.

As discussed in the Preamble to the ISAs ([U.S. EPA, 2015b](#)), this evaluation includes evidence from epidemiologic, controlled human exposure, and toxicological studies in addition to considering relevant exposure-related information. With regard to epidemiologic studies, the evaluation focuses on those studies that include stratified analyses to compare populations or lifestages exposed to similar air pollutant concentrations within the same study design along with consideration of the strengths and limitations of each study. Other epidemiologic studies that do not stratify results but instead examine a specific population or lifestage can provide supporting evidence for the pattern of associations observed in studies that formally examine effect modification. Similar to the characterization of evidence in [Chapter 5](#), the greatest emphasis is placed on patterns or trends in results across studies. Experimental studies in human subjects or animal models that focus on factors, such as genetic background or health status, are evaluated because they provide coherence and biological plausibility of effects observed in epidemiologic studies. Also evaluated are studies examining whether factors may result in differential exposure to SO₂ and subsequent increased risk of SO₂-related health effects. Conclusions are made with respect to whether a specific factor increases the risk of an SO₂-related health effect based on the characterization of evidence framework detailed in [Table 6-1](#). A summary of the characterization of the evidence for each factor considered in this chapter is presented in [Section 6.6](#).

Table 6-1 Characterization of evidence for factors potentially increasing the risk for sulfur dioxide-related health effects.

Classification	Health Effects
Adequate evidence	There is substantial, consistent evidence within a discipline to conclude that a factor results in a population or lifestage being at increased risk of air pollutant-related health effect(s) relative to some reference population or lifestage. Where applicable, this evidence includes coherence across disciplines. Evidence includes multiple high-quality studies.
Suggestive evidence	The collective evidence suggests that a factor results in a population or lifestage being at increased risk of air pollutant-related health effect(s) relative to some reference population or lifestage, but the evidence is limited due to some inconsistency within a discipline or, where applicable, a lack of coherence across disciplines.
Inadequate evidence	The collective evidence is inadequate to determine whether a factor results in a population or lifestage being at increased risk of air pollutant-related health effect(s) relative to some reference population or lifestage. The available studies are of insufficient quantity, quality, consistency, and/or statistical power to permit a conclusion to be drawn.
Evidence of no effect	There is substantial, consistent evidence within a discipline to conclude that a factor does not result in a population or lifestage being at increased risk of air pollutant-related health effect(s) relative to some reference population or lifestage. Where applicable, the evidence includes coherence across disciplines. Evidence includes multiple high-quality studies.

6.3 Pre-existing Disease

Individuals with pre-existing disease may be considered at greater risk for some air pollution-related health effects because disease status and severity may put those individuals in a compromised biological state. The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) concluded that individuals with pre-existing respiratory diseases, especially those with asthma, were likely to be at greater risk for SO₂-related health effects. Recent epidemiologic studies evaluated effect modification of both respiratory and cardiovascular effects by a number of pre-existing diseases, including asthma, cardiovascular disease, diabetes, and obesity. While we evaluated the evidence from each of these studies, the focus of this section is on effect modification of respiratory disease by pre-existing asthma since the relationship between short-term SO₂ exposure and respiratory health effects was determined to be causal in [Chapter 5](#), with the strongest evidence coming from studies of asthma exacerbation. Since the relationship between short-term SO₂ exposure and cardiovascular effects (the only outcome for which evidence is available to inform pre-existing cardiovascular disease, diabetes, or obesity) was determined to be inadequate, we evaluated the evidence to determine if stratification by these pre-existing diseases could explain any of the inconsistencies in study results of cardiovascular effects. There was no indication that pre-existing disease status led to

heterogeneity in study results for cardiovascular effects, thus the results for those pre-existing diseases are not detailed in this section.

Of the recent epidemiologic studies evaluating effect modification of respiratory effects by pre-existing disease, most focused on asthma ([Section 6.3.1](#)). [Table 6-2](#) presents the prevalence of asthma and other respiratory diseases according to the Centers for Disease Control and Prevention's (CDC's) National Center for Health Statistics ([Blackwell et al., 2014](#)) and the National Health Interview Survey ([Bloom et al., 2012](#)) including the proportion of adults and children with a current diagnosis categorized by age and geographic region. The large proportions of the U.S. population affected by respiratory diseases indicates the potential public health impact, and thus, the importance of characterizing the risk of SO₂-related health effects for affected populations.

Table 6-2 Prevalence of respiratory diseases among adults and children by age and region in the U.S. in 2012.

		Adults ^a —By Age					Adults ^a —By Region			
Numbers in Thousands, (%)		All 18+	18–44	45–64	65–74	75+	North-east	Midwest	South	West
Total population		234,921	111,034	82,038	23,760	18,089	42,760	53,378	85,578	53,205
Asthma ^b	“Ever had”	29,660 (12.6)	14,929 (13.4)	10,380 (12.7)	2,863 (12.0)	1,489 (8.2)	5,686 (13.3)	7,080 (13.3)	10,071 (11.8)	6,822 (12.8)
	“Still has”	18,719 (8.0)	8,943 (8.1)	6,852 (8.4)	1,837 (7.7)	1,088 (6.0)	3,953 (9.2)	4,358 (8.2)	6,280 (7.3)	4,129 (7.8)
Chronic bronchitis		8,658 (3.7)	2,721 (2.5)	3,831 (4.7)	1,165 (4.9)	940 (5.2)	1,446 (3.4)	2,438 (4.6)	3,449 (4.0)	1,325 (2.5)
COPD		6,790 (2.9)	512 (0.5)	3,074 (3.7)	1,646 (6.9)	1,558 (8.6)	1,013 (2.4)	1,860 (3.5)	2,781 (3.2)	1,135 (2.1)
		Children ^c —By Age				Children ^c —By Region				
Numbers in Thousands, (%)		All <18	0–4	5–11	12–17	North-east	Midwest	South	West	
Total population		74,518	21,210	28,845	24,463	11,956	17,651	27,170	17,741	
Asthma ^b	“Ever had”	10,463 (14.0)	1,753 (8.3)	4,139 (14.3)	4,571 (18.7)	1,888 (15.8)	2,190 (12.4)	4,088 (15.0)	2,298 (13.0)	
	“Still has”	7,074 (9.5)	1,452 (6.8)	2,849 (9.9)	2,773 (11.3)	1,356 (11.3)	1,557 (8.8)	2,636 (9.7)	1,525 (8.6)	

COPD = chronic obstructive pulmonary disease.

^aSource: [Blackwell et al. \(2014\)](#); National Center for Health Statistics: Data from Table 3 of the Centers for Disease Control and Prevention report.

^bAsthma prevalence estimates are available for both “ever had asthma” and “still has asthma.”

^cSource: [Bloom et al. \(2012\)](#); Summary health statistics for U.S. children: National Health Interview Survey, Vital and Health Statistics, Series 10, Number 254. December, 2012.

6.3.1 Asthma

Approximately 8.0% of adults and 9.5% of children (age <18 years) in the U.S. currently have asthma ([Blackwell et al., 2014](#); [Bloom et al., 2012](#)), and it is the leading chronic illness affecting children ([Bloom et al., 2012](#)). Based on evidence from the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) and recent studies, a causal relationship exists between short-term SO₂ exposure and respiratory effects, based primarily on evidence from controlled human exposure studies demonstrating decrements in lung function in

individuals with asthma ([Sections 5.2.1.2](#) and [5.2.1.9](#)). This is nearly the same body of evidence evaluated in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), which also concluded that individuals with asthma are more sensitive to exposures to SO₂. Children with asthma may be particularly at risk compared to adults with asthma due to (1) their increased responsiveness to methacholine, a potential surrogate for SO₂ ([Section 5.2.1.2](#)), relative to adults; (2) children's increased ventilation rates relative to body mass compared to adults; and (3) the increased proportion of oral breathing observed among children, particularly boys, relative to adults ([Section 4.1.2](#)). Such oral breathing allows greater SO₂ penetration into the tracheobronchial region of the lower airways than nasal breathing ([Section 4.2.2](#)). In addition, children tend to spend more time outdoors (where SO₂ levels are higher, compared to indoor levels), and, consequently, have the potential for longer exposure to higher levels of SO₂. While there are a number of behavioral, environmental and physical characteristics that, in addition to SO₂, could contribute to asthma exacerbations, there is little or no empirical evidence for how these characteristics might interact with SO₂ and contribute to individuals with asthma being more at risk for health effects attributed to SO₂ than healthy individuals. This section briefly describes evidence from the experimental studies and supporting evidence from epidemiologic studies ([Table 6-3](#)).

Across experimental evidence, adults with asthma consistently have greater decrements in lung function with SO₂ exposure than those without asthma. Controlled human exposure studies have evaluated respiratory outcomes among adults at SO₂ concentrations ranging from 0.2 to 1 ppm and included exposures with and without exercise. In a controlled human exposure study examining SO₂ concentrations ranging from 0.2 to 0.4 ppm in repeated exposures, while exercising, of healthy individuals and individuals with mild, atopic and moderate/severe asthma, [Linn et al. \(1987\)](#) reported that individuals having moderate and severe asthma showed the greatest SO₂-dependent respiratory effects (airway resistance, FEV₁, symptoms). Other than asthma status, subject-level characteristics (e.g., weight, height, age and sex of those having asthma) did not influence the response. [Magnussen et al. \(1990\)](#) also reported increases in sRaw in subjects with asthma that were not observed in healthy controls, with SO₂ exposures incorporating exercise. In addition to controlled human exposure studies, a long-term exposure study conducted in ovalbumin (OVA)-sensitized rats as an asthma model demonstrated that 4 weeks of exposure to 2 ppm SO₂ resulted in increased airway resistance compared to normal rats ([Song et al., 2012](#)).

Table 6-3 Controlled human exposure, epidemiology, and animal toxicology studies evaluating pre-existing asthma and sulfur dioxide exposure.

Factor Evaluated	Reference Category	Direction of Effect Modification or Effect ^a	Outcome	Study Population ^b	Study Details	Study
Controlled human exposure						
Asthma (atopic)	Healthy	↑	Lung function (sRaw)	n = 4 healthy adults	0.2, 0.4, 0.6 ppm SO ₂ for 1 h with exercise; Exposures were repeated eight times	Linn et al. (1987)
Mild asthma		↑		n = 21 atopic adults		
Moderate/severe asthma		↑		n = 16 adults with mild asthma		
Asthma (atopic)	Healthy	↑	Lung function (FEV ₁)	n = 24 adults with moderate/severe asthma		
Mild asthma		↑				
Moderate/severe asthma		↑				
Asthma (atopic)	Healthy	↑	Respiratory symptoms during exposure			
Mild asthma		↑				
Moderate/severe asthma		↑				
Asthma	Healthy	↑	Lung function (sRaw)	n = 46 adults with bronchial asthma, 12 healthy adults	0.5 ppm SO ₂ for 10 min tidal breathing, 10 min of isocapnic hyperventilation (30 L/min); histamine challenge	Magnussen et al. (1990)
Asthma	Healthy	-	Lung function (FEV ₁ , FVC, MMEF)	n = 12 adults with asthma, 12 healthy adults	0.2 ppm SO ₂ for 1 h at rest	Tunnicliffe et al. (2003)

Table 6-3 (Continued): Controlled human exposure, epidemiology, and animal toxicology studies evaluating pre existing asthma and sulfur dioxide exposure

Factor Evaluated	Reference Category	Direction of Effect Modification or Effect ^a	Outcome	Study Population ^b	Study Details	Study
Epidemiology						
With asthma n = 84	Without asthma n = 422	-	Lung function (PEF)	n = 506 elementary school children ages 8–13 yr	Guadeloupe (French West Indies) December 2008–December 2009	Amadeo et al. (2015)
Toxicology						
Rat asthma model (OVA sensitization)	Normal rats	↑	AHR (methacholine)	Rats (Sprague-Dawley), n = 10 males/group (4 wk)	2 ppm SO ₂ for 4 h/day for 4 wk beginning at 15 days	Song et al. (2012)
		↑	IL-4 in BALF			
		-	IFN-γ in BALF			
		↑	Airway smooth muscle cell stiffness (in vitro)			
		↑	Airway smooth muscle cell contractility (in vitro)			

AHR = airway hyperresponsiveness; BALF = bronchoalveolar lavage fluid; FEV₁ = forced expiratory volume in 1 sec; FVC = forced vital capacity; IFN-γ = interferon gamma; IL-4 = interleukin 4; MMEF = maximum mid-expiratory flow; n = sample size; OVA = ovalbumin; PEF = peak expiratory flow; SO₂ = sulfur dioxide; sRAW = specific airway resistance.

^aUp facing arrow (↑) indicates that the effect of SO₂ is greater (e.g., larger lung function decrement, larger increase in airway inflammation) in the group with the factor evaluated than in the reference group. Down facing arrow (↓) indicates that the effect of SO₂ is smaller in the group with the factor evaluated than in the reference group. A dash (-) indicates no substantial difference in SO₂-related health effect between groups. In some studies, only a population with pre-existing disease was examined; therefore, the arrow or dash represents the direction of the effect in that population after exposure to SO₂ relative to exposure to filtered air.

^bUnless ages are indicated in the row for each study, the mean age or range was not reported in the study aside from indication of adult subjects.

Of the recent literature included in this ISA, one epidemiologic study included stratification by asthma status and did not find a difference for associations of short-term SO₂ exposure with changes in lung function [Table 6-3; (Amadeo et al., 2015)]. However, evidence presented in Section 5.2.1.2 from many studies generally demonstrates consistent positive associations between ambient SO₂ concentrations and asthma-related hospitalizations and ED visits. This is important supporting evidence because hospitalizations and ED visits related to asthma would generally not be studied in

individuals without asthma, and thus, are not amendable to stratification. In addition to these studies, some evidence from recent panel studies ([Dong et al., 2013c](#); [Sahsuvaroglu et al., 2009](#)) and studies reviewed in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) indicates that children with asthma experience respiratory symptoms associated with exposure to ambient SO₂.

In conclusion, evidence from controlled human exposure and animal toxicological studies is consistent in demonstrating decrements in lung function with SO₂ exposures in people with asthma. The evidence for healthy individuals does not show responses at exposures below 1 ppm (as summarized above and in [Section 5.2.1.2](#)). There is also clear biological plausibility including key events contributing to the mode of action ([Section 4.3](#)) linking SO₂ exposure to asthma exacerbation and supporting the observed effects from experimental studies. Furthermore, epidemiologic studies report associations between SO₂ exposure and ED visits and hospital admissions due to asthma, and that individuals with asthma experience respiratory symptoms associated with exposure to ambient SO₂.

Overall, there is adequate evidence from multiple, high-quality studies and coherence across scientific disciplines to conclude that people with pre-existing asthma are at increased risk of SO₂-induced respiratory effects.

6.4 Genetic Factors

Genetic variation in the human population is known to contribute to numerous diseases and differential physiologic responses. The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) discussed the biological plausibility of individuals with certain genotypes known to result in reduced function in genes encoding antioxidant enzymes being at increased risk for respiratory effects related to ambient air pollution. However, the evidence base was limited to two studies demonstrating individuals with polymorphisms in *GSTP1* and tumor necrosis factor to be at increased risk for SO₂-related asthma and decrements in lung function. A recently conducted study reviewed in this ISA examined effect measure modification by genotype ([Reddy et al., 2012](#)) and reported inconsistent results across *GSTM1* and *GSTP1* genotypes in a relatively small sample of children in South Africa. The *GSTM1* null genotype and the *GSTP1 Ile105Ile* and *Ile105Val* genotypes are associated with reduced antioxidant enzyme function; however, effect measure modification of these genotypes on SO₂-associated intra-day variability of FEV₁ showed conflicting results. **Despite biological plausibility characterized in the 2008 SO_x ISA, the limited and inconsistent evidence base is inadequate to determine whether genetic background contributes to increased risk for SO₂-related health effects.**

6.5 Sociodemographic and Behavioral Factors

Recent epidemiologic evidence evaluated effect modification of respiratory, cardiovascular, cancer, reproductive and developmental, and mortality effects associated with SO₂ exposure by a number of sociodemographic and behavioral factors, including lifestage, sex, socioeconomic status, race/ethnicity and smoking. While we evaluated the evidence from each of these studies within this ISA, the focus of this section is on effect modification of the association of SO₂ exposure with respiratory disease since the relationship determined to be causal is that for short-term SO₂ exposure and respiratory health effects (see [Chapter 5](#)). Since the evidence for relationships between SO₂ exposure and cardiovascular disease, cancer, and reproductive and developmental outcomes were determined to be inadequate in [Chapter 5](#), we evaluated the evidence to determine if stratification by sociodemographic or behavioral factors could explain any of the inconsistencies in study results. **There was no indication that sociodemographic and behavioral factors led to heterogeneity in study results for cardiovascular disease, cancer, or reproductive and developmental outcomes**, thus those results are not detailed in this section.

6.5.1 Lifestage

The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) discussed some evidence for increased risk of health effects related to SO₂ exposure among different lifestages (i.e., children and older adults). Lifestage refers to a distinguishable time frame in an individual's life characterized by unique and relatively stable behavioral or physiological characteristics associated with development and growth ([U.S. EPA, 2014c](#)). Differential health effects of SO₂ across lifestages theoretically could be due to several factors. With regard to children, the human respiratory system is not fully developed until 17–21 years of age for females and 19–23 years of age for males, and therefore, children could plausibly have intrinsic risk for respiratory effects due to potential perturbations in normal lung development ([Finkelstein and Johnston, 2004](#); [Hankinson et al., 1999](#)). In addition, children spend more time outdoors compared with adults, and, as a result, may experience greater SO₂ exposure ([Section 3.4.2.1](#)). Older adults (typically considered those 65 years of age or greater) have weakened immune function, impaired healing, decrements in pulmonary and cardiovascular function, and greater prevalence of chronic disease [([Rosenthal and Kavic, 2004](#)); [Table 6-2](#)], which may contribute to or worsen health effects related to SO₂ exposure. Also, exposure or internal dose of SO₂ may vary across lifestages due to varying ventilation rates, extent of oronasal breathing at rest, and time-activity patterns. The following sections present the evidence comparing lifestages

from the recent literature, which builds on the evidence presented in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)).

6.5.1.1 Children

According to the 2010 census, 24% of the U.S. population is less than 18 years of age, with 6.5% less than age 6 ([Howden and Meyer, 2011](#)). The large proportion of children within the U.S. demonstrates the public health importance of characterizing the risk of SO₂-related health effects among children. This is especially so because of the causal relationship between short-term SO₂ exposure and respiratory outcomes, with strong evidence demonstrating lung function decrements in individuals with asthma, which affects approximately 11.3% of children 12–17 years old and 9.5% of children less than 18 years old ([Table 6-2](#)). The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) presented evidence from epidemiologic studies indicating larger effect estimates for SO₂-related respiratory outcomes among children than for adults or all ages, including asthma-related ED visits and hospitalizations; however, recent evidence is not entirely consistent with the evidence considered previously ([Table 6-4](#)). Although [Son et al. \(2013\)](#) found children (0–14 years) to be at greater risk for SO₂-related asthma hospital admissions compared with adults, neither [Ko et al. \(2007b\)](#) nor [Alhanti et al. \(2016\)](#) observed differences between children and adults when examining associations of ambient SO₂ with asthma hospitalizations or ED visits. When examining evidence for different age groups of children, [Jalaludin et al. \(2008\)](#) observed that associations for respiratory-related ED visits among children ages 1–4 years were greater than for children ages 10–14 years; however, [Samoli et al. \(2011\)](#) and [Villeneuve et al. \(2007\)](#) did not find stronger associations for asthma-related hospital admissions or ED visits among younger children. Similarly, [Dong et al. \(2013c\)](#) did not find age-related differences among children for SO₂-associated asthma, and [Sahsuvaroglu et al. \(2009\)](#) found children ages 6–7 years had smaller risk for SO₂-associated nonallergic asthma compared to adolescents at 13–14 years.

Overall, the combined evidence from the previous and current ISA examining respiratory outcomes across lifestages is suggestive of increased risk in children compared to adults, given the inconsistencies across epidemiologic studies and limited toxicological evidence to inform plausibility. In addition, there is no evidence that younger (e.g., ≤7 years old) or older (e.g., 8–18 years old) children are more at risk for asthma exacerbation-related effects due to short-term SO₂ exposure than children in general. There are biological factors (e.g., increased ventilation rates relative to body mass among children and increased oral breathing that lead to greater SO₂ penetration and bronchial surface doses) that indicate a potential for children to receive a greater

internal dose in the bronchotracheal region, and for children with asthma to be at increased risk as compared to adults with asthma (Section 6.3). However, recent evidence, mainly from epidemiologic studies of respiratory ED visits and hospital admissions, does not consistently show increased risk among children (Table 6-4).

Table 6-4 Epidemiologic studies evaluating childhood lifestage and sulfur dioxide exposure.

Factor Evaluated	Reference Category	Direction of Effect Modification ^a	Outcome	Study Population	Study Details	Study
Short-term exposure						
Childhood ages 0–14 yr n = 60.1 admissions/day	All ages n = 104.9 admissions/day	↓	Hospital admissions for acute respiratory distress	14 hospitals	Hong Kong, China 1996–2002	Wong et al. (2009)
Childhood ages 0–14 yr n = 23,596	Adulthood ages 15–65 yr n = 21,204	-	Asthma hospital admissions	15 hospitals n = 69,176 admissions	Hong Kong, China 2000–2005	Ko et al. (2007b)
Childhood ages 0–14 yr n = 8.7 admissions/day	Adulthood ages 15–64 yr n = 4.3 admissions/day	↑	Asthma hospital admissions	Database accounting for 48% of South Korean population n = 19/d	Eight South Korean cities 2003–2008	Son et al. (2013)
Childhood ages 0–4 yr n = 72%	Childhood ages 5–14 yr n = 28%	-	Asthma hospital admissions	Three main children's hospitals approximately 85% of pediatric beds of metropolitan area of Athens n = 3,601	Athens, Greece 2001–2004	Samoli et al. (2011)
Childhood ages 2–4 yr n = 7,247	Childhood ages 5–14 yr n = 13,145	-	Asthma ED visits	Five hospitals servicing more than 80% of the metropolitan area n = 57,192 visits	Edmonton, Canada 1992–2002	Villeneuve et al. (2007)

Table 6-4 (Continued): Epidemiologic studies evaluating childhood lifestage and sulfur dioxide exposure

Factor Evaluated	Reference Category	Direction of Effect Modification ^a	Outcome	Study Population	Study Details	Study
Childhood ages 1–4 yr n = 109 admissions/day	Childhood ages 10–14 yr n = 25 admissions/day	↑	Respiratory-related ED visits	Daily number of ED visits in metropolitan Sydney from the New South Wales Health Department n = 174/d	Sydney, Australia 1997–2001	Jalaludin et al. (2008)
Childhood ages 5–18 yr n = 59.6 admissions/day	Adulthood ages 19–39 yr n = 41.1 admissions/day	-	Asthma ED visits	Daily number of ED visits in metropolitan area n = 62.8/day (Atlanta) n = 76.3/day (Dallas) n = 50.6/day (St. Louis)	Three U.S. cities (Atlanta, GA 1993–2009; Dallas, TX 2006–2009; St. Louis, MO 2001–2007)	Alhanti et al. (2016)
Long-term exposure						
Childhood ages 2–5 yr n = 7,508	Childhood ages 6–14 yr n = 23,541	-	Doctor-diagnosed asthma	n = 31,049 Children ages 2–14 yr	Seven northeastern cities study, Liaoning Province, northeast China 2008–2009	Dong et al. (2013c)
		↑	Respiratory symptoms (cough, phlegm, current wheeze)			
Younger children ages 6–7 yr n = 918	Older children ages 13–14 yr n = 549	↓	Nonallergic asthma	n~ 1,467 Children grades 1 (ages 6–7 yr) and 8 (ages 13–14 yr)	Hamilton, Canada 1994–1995	Sahsuvaroglu et al. (2009)

ED = emergency department; n = sample size.

^aUp facing arrow indicates that the effect of is greater (e.g., larger increase in hospital admission) in the group with the factor evaluated than in the reference group. Down facing arrow indicates that the effect of sulfur dioxide is smaller in the group with the factor evaluated than in the reference group. A dash indicates no substantial difference in sulfur dioxide-related health effect between groups.

6.5.1.2

Older Adults

According to the 2008 National Population Projections issued by the U.S. Census Bureau, approximately 12.9% of the U.S. population is age 65 years or older, and by 2030, this fraction is estimated to grow to 20% ([Vincent and Velkoff, 2010](#)). Thus, this lifestage represents a substantial proportion of the U.S. population demonstrating the public health importance of characterizing the potential for increased risk for health effects related to SO₂ exposure in this age group.

The 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) indicated that compared with younger adults, older adults (typically ages 65 years and older) may be at increased risk for SO₂-related respiratory emergency department visits and hospitalizations, but limited evidence was available to inform risk related to respiratory effects. Recently published studies evaluating risk in older adults compared to younger adults are characterized in [Table 6-5](#) and generally support conclusions from the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)). For example, [Villeneuve et al. \(2007\)](#) and [Son et al. \(2013\)](#) both reported that asthma-related ED visits and hospital admissions were more strongly associated with short-term ambient SO₂ exposure in individuals older than 75 years than adults 65–74 years or those younger than 65. However, the handful of recent studies evaluating asthma and nonasthma respiratory admissions or ED visits in adults older than 65 years of age reported inconsistent results compared to the earlier literature ([Alhanti et al., 2016](#); [Son et al., 2013](#); [Arbex et al., 2009](#); [Wong et al., 2009](#); [Ko et al., 2007b](#)). In addition to these studies of short-term SO₂ exposure, [Forbes et al. \(2009c\)](#) found older adults (45–74 and older than 75 years) to have larger decrements in lung function in association with average annual SO₂ concentrations, compared to adults aged 16–44.

Other recent studies comparing results in older and younger adults evaluated associations between short-term SO₂ exposures and mortality and generally observed inconsistent results ([Bravo et al., 2015](#); [Chen et al., 2012c](#); [Wong et al., 2008b](#)). [Chen et al. \(2012c\)](#) and [Wong et al. \(2008b\)](#) both observed evidence for increased risk of total mortality with short-term SO₂ exposures in adults older than 75 years compared to other age groups, which is consistent with age-specific evidence from studies of respiratory effects. However, [Bravo et al. \(2015\)](#) did not observe any differences in risk across age groups.

Taken together, the collective evidence builds on conclusions from the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) and is suggestive that older adults may be at increased risk for SO₂-related health effects compared to younger adults. The evidence from the current and previous ISA related to respiratory hospitalizations and ED visits indicates that older adults, particularly those older than 75 years, may be at

increased risk for SO₂-related health effects, although this evidence is not entirely consistent.

Table 6-5 Epidemiologic studies evaluating older adult lifestage and sulfur dioxide exposure.

Factor Evaluated	Reference Category	Direction of Effect Modification ^a	Outcome	Study Population	Study Details	Study
Older adulthood ages >65 yr n = 24,916	Younger adulthood ages 15–65 yr n = 21,204	-	Asthma hospital admissions	15 hospitals n = 69,176 admissions	Hong Kong, China 2000–2005	Ko et al. (2007b)
Older adulthood ages 65–74 yr n = 4,705	Younger adulthood ages 15–64 yr n = 32,815	-	Asthma ED visits	Five hospitals n = 57,912 visits	Edmonton, Canada 1992–2002	Villeneuve et al. (2007)
Older adulthood ages ≥75 yr n = 1,855		↑				
Adulthood ages 65+ yr n = 4.7 admissions/day	Adulthood ages 19–39 yr n = 41.1 admissions/day	-	Asthma ED visits	Daily number of ED visits in metropolitan area n = 62.8/day (Atlanta) n = 76.3/day (Dallas) n = 50.6/day (St. Louis)	Three U.S. cities (Atlanta, GA 1993–2009; Dallas, TX 2006–2009; St. Louis, MO 2001–2007)	Alhanti et al. (2016)
Older adulthood ages ≥65 yr n = 789	Younger adulthood ages 40–64 yr n = 980	↑	COPD ED visits	São Paulo Hospital, daily records for patients >40 yr n = 1,769	São Paulo, Brazil 2001–2003	Arbex et al. (2009)
Older adulthood ages 65–74 yr n = 5.8 admissions/day	Younger adulthood ages 15–64 yr n = 8.8 admissions/day	-	Asthma and allergic disease hospital admissions	Hospital admission database accounting for 48% of Korean population n = 37.7/day	Eight South Korean cities 2003–2008	Son et al. (2013)

Table 6-5 (Continued): Epidemiologic studies evaluating older adult lifestage and sulfur dioxide exposure.

Factor Evaluated	Reference Category	Direction of Effect Modification ^a	Outcome	Study Population	Study Details	Study
Older adulthood ages ≥75 yr n = 5.8 admissions/day	Younger adulthood ages 15–64 yr n = 8.8 admissions/day	↑				
Older adulthood ages ≥65 yr n = 59.6	All ages n = 91.5	-	COPD hospital admissions	14 hospitals	Hong Kong, China 1996–2002	Wong et al. (2009)
Older adulthood ages ≥65 yr n = 138.5	All ages n = 270.3	-	Respiratory disease hospital admissions			
Older adulthood ages ≥65 yr ^b	Adulthood, childhood ages 5–64 yr ^b	↑	Total mortality	Data from Municipal Centers for Disease Control and Prevention	17 Chinese cities	Chen et al. (2012c)
Older adulthood ages ≥75 yr	All ages (≥65 yr)	↑	Total mortality	Data from the Ministry of Public Health, Bangkok; the Census and Statistic Department, Hong Kong; the Shanghai Municipal Center of Disease Control and Prevention, Shanghai; and the Wuhan Centre for Disease Prevention and Control	Bangkok, Thailand; Hong Kong, Shanghai, and Wuhan, China 1996–2004	Wong et al. (2008b)
Older adulthood ages 65–74 yr n = 194,202	Younger adulthood ages 35–64 n = 315,435	↑	Mortality	N = 849,127	Sao Paulo, Brazil May 1996–December 2010	Bravo et al. (2015)
Older adulthood ages ≥75 yr n = 339,490	Younger adulthood ages 35–64 n = 315,435	↑				

COPD = chronic obstructive pulmonary disease; ED = emergency department; n = sample size.

^aUp facing arrow indicates that the effect of sulfur dioxide is greater (e.g., larger risk of hospital admission, larger decrement in heart rate variability) in the group with the factor evaluated than in the reference group. Down facing arrow indicates that the effect of sulfur dioxide is smaller in the group with the factor evaluated than in the reference group. A dash indicates no substantial difference in sulfur dioxide-related health effect between groups.

^bSample size not reported.

6.5.2

Sex

A vast number of health conditions and diseases have been shown to differ by sex, and there is some indication of differences by sex in the relationship between air pollution and health effects. The 2010 U.S. census indicates an approximately equal distribution of males and females in the U.S.: 49.2% male and 50.8% female ([Howden and Meyer, 2011](#)). However, the distribution varies by age, with a greater prevalence of females above 65 years of age compared to males. Thus, the public health implications of potential sex-based differences in SO₂-related health effects may vary among age groups within the population.

There are a number of studies evaluating sex-based differences in SO₂-associated health effects, as detailed in [Table 6-6](#). Studies of short-term SO₂ exposures and respiratory effects in children and adults did not consistently indicate differences by sex. [Ishigami et al. \(2008\)](#) found adult females to have increased respiratory symptoms with ambient SO₂ exposure compared to adult males; however, [Son et al. \(2013\)](#) found larger associations for asthma or allergic disease hospitalizations for males compared to females. No differences were found between men and women for SO₂-related COPD ED visits ([Arbex et al., 2009](#)). In children, SO₂-associated decrements in lung function were not different between boys and girls ([Linares et al., 2010](#); [Dales et al., 2009](#)), although [Samoli et al. \(2011\)](#) found boys to have higher associations between ambient SO₂ exposure and asthma hospital admissions. In a long-term SO₂ exposure study, [Deng et al. \(2015a\)](#) observed stronger associations with asthma incidence among boys compared to girls.

The collective body of evidence does not clearly indicate that SO₂-related health effects differ between males and females. Due to the inconsistent results demonstrated across epidemiologic studies and a lack of experimental studies examining sex-based differences, **the evidence is inadequate to determine whether males or females may be at increased risk for SO₂-related health effects.**

Table 6-6 Epidemiologic studies evaluating effect modification by sex and sulfur dioxide exposure.

Factor Evaluated	Reference Category	Direction of Effect Modification ^a	Outcome	Study Population	Study Details	Study
Short-term exposure						
Female 20% person h	Male 80% person h	↑	Respiratory symptoms (cough, scratchy throat, sore throat, breathlessness)	Healthy adult volunteers working on an active volcanic island after the evacuation order was lifted n = 955	Miyakejima Island, Japan 2005	Shigami et al. (2008)
Female n = 39	Male n = 114	-	Lung function (FEV ₁)	Elementary school children with asthma (no cigarette smoking in home) n = 182 children (ages 9–14 yr)	Windsor, Canada October–December 2005	Dales et al. (2009)
Female n = 235	Male n = 229	-	Lung function (FEV ₁ , FVC, PEF, FEV ₁ /FVC)	Children recruited from two schools with different roadway proximity n = 464 (6–14 yr)	Salamanca, Mexico 2004–2005	Linares et al. (2010)
Female n = 794	Male n = 875	-	COPD ED visits	São Paulo Hospital, daily records for patients >40 yr n = 1,769	São Paulo, Brazil 2001–2003	Arbex et al. (2009)
Female n = 7.4 admissions/ day	Male n = 8 admissions/ day	↓	Asthma hospital admissions	Database accounting for 48% of South Korean population n = 19/day	Eight South Korean cities 2003–2008	Son et al. (2013)
Female n = 7.1 admissions/ day	Male n = 8 admissions/ day	↓	Allergic disease hospital admissions			

Table 6-6 (Continued): Epidemiologic studies evaluating effect modification by sex and sulfur dioxide exposure.

Factor Evaluated	Reference Category	Direction of Effect Modification ^a	Outcome	Study Population	Study Details	Study
Female n = 1,332	Male n = 2,269	↓	Asthma hospital admissions	Three main children's hospitals; approximately 85% of pediatric beds of metropolitan area of Athens n = 3,601	Athens, Greece 2001–2004	Samoli et al. (2011)
Long-term exposure						
Female n = 1,153	Male n = 1,337	↓	Asthma incidence	Children from 36 different kindergartens n = 2,490	Changsha, China	Deng et al. (2015a)

COPD = chronic obstructive pulmonary disease; ED = emergency department; FEV₁ = forced expiratory volume in 1 sec; FVC = forced vital capacity; n = sample size; PEF = peak expiratory flow.

^aUp facing arrow indicates that the effect of SO₂ is greater (e.g., larger risk of hospital admission) in the group with the factor evaluated than in the reference group. Down facing arrow indicates that the effect of sulfur dioxide is smaller in the group with the factor evaluated than in the reference group. A dash indicates no substantial difference in sulfur dioxide-related health effect between groups.

6.5.3 Socioeconomic Status

SES is a composite measure that usually consists of economic status indicated by income, social status by education, and work status by occupation. Generally, persons with lower SES have been found to have a higher prevalence of pre-existing diseases, potential inequities in access to resources such as healthcare, and possibly increased nutritional deficiencies, which may increase their risk for SO₂-related health effects ([Wong et al., 2008a](#); [WHO, 2006](#)). According to U.S. census data, 15.9% (approximately 48.5 million) of Americans lived below the poverty threshold in 2011 as defined by household income, which is one metric used to define SES ([Bishaw, 2012](#)). The wide array of SES factors that can be used to describe or assign SES can complicate any synthesis of findings because definitions of SES vary across countries based on population demographics, bureaucracy, and the local economy. In addition, lower SES may coincide with proximity to pollution sources in some locations. As a result of these complexities, the ability to draw conclusions regarding SES as a factor for increased risk for health effects related to SO₂ exposure can be difficult.

A single study ([Cakmak et al., 2016](#)) evaluated the potential for SES (income or education) to modify the effect of long-term exposure to SO₂ on respiratory effects, specifically measures of lung function. The authors observed greater decrements in lung function for those in the lowest income and education groups when compared to those in the highest. Similarly, a single studies provides limited evidence for the potential for SES to modify the effect of short-term SO₂ exposure on mortality. [Chen et al. \(2012c\)](#) found low education to increase risk for mortality with short-term SO₂ exposure. Overall, the evidence for effect modification by SES on SO₂-related health outcomes is limited to a single study of respiratory health effects and one on mortality. **This limited evidence is inadequate to determine whether low SES increases risk for SO₂-related health effects.**

6.5.4 Smoking

Smoking is a common behavior as indicated by the 2010 National Health Interview Survey, which estimated that approximately 19.2% of the U.S. adult population report being current smokers and 21.5% report being former smokers ([Schiller et al., 2012](#)). Smoking is a well-documented risk factor for many diseases, but it is unclear whether smoking exacerbates health effects associated with air pollutant exposures, including SO₂.

[Dong et al. \(2012\)](#), [Forbes et al. \(2009c\)](#), and [Smith et al. \(2016\)](#) investigated effect modification of the relationship between long-term exposure to SO₂ and respiratory endpoints by smoking status. [Dong et al. \(2012\)](#) found that among the few respiratory deaths included in their retrospective cohort study, associations with long-term ambient SO₂ concentrations were only present with current smoking status. [Smith et al. \(2016\)](#) observed positive associations between long-term average SO₂ concentration and pulmonary tuberculosis among ever smokers, but not with never smokers. [Forbes et al. \(2009c\)](#), on the other hand, did not find current smoking to increase risk for lung function decrements with long-term SO₂ exposure compared to not smoking; however, former smoking did appear to increase risk in this study.

Overall, the inconsistent evidence is inadequate to determine whether smoking exacerbates SO₂-related health effects. A limited number of long-term exposure studies observed positive associations among current or former smokers, but not for never smokers for various respiratory health endpoints, including respiratory mortality. No studies evaluated smoking as an effect modifier of the relationship between short-term exposure to SO₂ and respiratory outcomes, for which there is the most confidence in the causal nature of the relationship.

This chapter characterized factors that may result in populations and lifestyles being at increased risk for SO₂-related health effects; a summary of at-risk factors and resulting evidence classifications is included in [Table 6-7](#). The evaluation of each factor focused on the consistency, coherence, and biological plausibility of evidence integrated across scientific disciplines: specifically, epidemiologic, controlled human exposure, and toxicological studies using the weight-of-evidence approach detailed in [Table 6-1](#). In evaluating and integrating evidence related to at-risk factors, it is important to consider additional information including exposure concentrations, dosimetry, modes of action, and/or the independence of relationships of SO₂ exposure with health effects as detailed in [Chapter 5](#). For many potential at-risk factors summarized in [Table 6-7](#), there was limited evidence of an influence on SO₂-related health effects.

Consistent with observations made in the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)), **the evidence is adequate to conclude that people with asthma are at increased risk for SO₂-related health effects**. Most of the evidence for this conclusion was presented in the previous ISA, but recent studies consistently indicate increased risk across studies. Furthermore, the evidence is based on findings for short-term SO₂ exposure and respiratory effects (specifically lung function decrements), for which a causal relationship exists ([Section 5.2.1.9](#)). There are a limited number of epidemiologic studies evaluating SO₂-related respiratory effects in people with asthma, but there is evidence for asthma-related hospital admissions and ED visits ([Section 5.2.1.2](#)). Further support for increased risk in individuals with asthma is provided by biological plausibility drawn from modes of action. Asthma prevalence in the U.S. is approximately 8–11% across age groups ([Blackwell et al., 2014](#); [Bloom et al., 2012](#)) and thus represents a substantial fraction of the population that may be at risk for respiratory effects related to ambient air SO₂ concentrations.

There is suggestive evidence of an increased risk of SO₂-related respiratory effects in children and older adults. Although the 2008 ISA for Sulfur Oxides ([U.S. EPA, 2008d](#)) discussed several studies indicating stronger associations between SO₂ and respiratory outcomes for these lifestyles, the recent evidence is less consistent. For children, studies comparing SO₂-associated respiratory outcomes reported mixed results, but known age-related factors such as higher ventilation rates ([Section 4.1.2.1](#)) and time-activity patterns ([Section 3.4.2.1](#)) provide plausibility for higher SO₂ exposure and/or dose in children. For adults, recent research generally finds similar associations for SO₂-related respiratory outcomes across age groups, although individuals over 75 years were more consistently at increased risk. In addition, there was limited toxicological evidence to support observations made across epidemiologic studies.

Table 6-7 Summary of evidence for potential increased sulfur dioxide exposure and increased risk of sulfur dioxide-related health effects.

Factor Evaluated	At-Risk Group	Health Evidence	Rationale for Classification
Adequate evidence			
Pre-existing disease	Individuals with asthma	Respiratory	Consistent evidence for increased risk for SO ₂ -related lung function decrements in controlled human exposure studies Support provided by epidemiologic studies of hospital admissions and ED visits for respiratory causes
Suggestive evidence			
Lifestage	Children	Respiratory	Evidence for increased risk among children provided in 2008 SO _x ISA; older studies provide biological plausibility; recent epidemiologic studies provide limited support, and are not entirely consistent
	Older adults	Respiratory	Evidence for increased risk for older adults provided in 2008 SO _x ISA; mixed results in recent epidemiologic studies for respiratory-related outcomes
		Mortality	Generally inconsistent evidence from a limited number of recent epidemiologic studies of short-term SO ₂ exposure and mortality
Inadequate evidence			
Genetic background	None identified	Respiratory	Single epidemiologic study shows decreased risk among <i>GSTM1</i> individuals and increased risk among <i>GSTP1</i> Ile/Ile (AA) individuals
Sex	None identified	Respiratory	Inconsistent differences in SO ₂ -related health effects, or no observed differences; studies limited in quantity
Socioeconomic status	None identified	Respiratory	Evidence limited to single study
		Mortality	Evidence limited to single study
Smoking	None identified	Respiratory	Single controlled human exposure study saw no effect on lung function among current smokers but an increase in risk among former smokers compared to never smokers
Evidence of no effect			
None			
ED = emergency department; ISA = Integrated Science Assessment; SO ₂ = sulfur dioxide.			

For all other at-risk factors considered based on available information, evidence was inadequate to determine whether those factors result in increased risk for SO₂-related health effects. Generally, there was a limited number of studies available evaluating SES, genetic background, and smoking. Many of these factors are interrelated and are known to impact health risks related to air pollution in general, but the scientific evidence available in the published literature specific to health effects associated with ambient SO₂ exposure is inadequate to determine whether these factors confer increased risk.

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